

Algorithmic Trading & Quantitative Strategies

Lecture 5 (4/2/2024)

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Today's Session

- I will send a project description on Thursday
- There will be a short multiple-answer quiz at the end of the lecture.
- Please send answers in the form
 - 1-a
 - 2-c
 - ...

To gp2642@nyu.edu. You will have 5 minutes

Recap on Last Quiz

1. A drawback of statistical models is
 - a. Low predictive power
 - b. Computational burden
 - c. Limited history
 - d. **Low information use**
2. The standard PCA is based on an underlying factor model
 - a. True
 - b. **False**
 - c. Uncertain
3. If two eigenvalues of a matrix are identical, its eigenvectors have
 - a. Indefinite norm
 - b. **Indefinite direction**
 - c. The jargon name "Dick" and "Jane" in the academic literature
 - d. Very important
4. Why can't you use Marchenko–Pastur for cleaning correlation matrices?
 - a. Unit variance assumption does not hold
 - b. **Zero correlation assumption does not hold**
 - c. Sample is too small
5. PPCA recommends to
 - a. Shrink the factor volatilities and correlations
 - b. Shrink the factor correlations
 - c. **Shrink the factor volatilities**
6. PPCA recommends to
 - a. Shrink the idio volatilities
 - b. **Shrink the big idio volatilities, increase the small ones**
 - c. Leave the damn things unchanged

Topics

- Statistical Models
 - Practical considerations: z-scoring and reducing turnover
- Fundamental Factor Models
- Next Time:
 - Recap on Fundamental Models
 - Advanced Portfolio Construction

Why Fundamental Models?

- Interpretable
- Extensible
- Good performance
- Connects to academic research and to alpha research

Primitives

- Returns
 - Which returns? Closing auctions. Based on equal weighted/EWMA/volume weighted averages
- Characteristics
 - "Features", "Terms", "Signals"
 - Continuous variables
 - Discrete/categorical

Process

- Data ingestion and quality assurance
 - Missingness
 - Turnover
- Winsorization
- Estimation universe selection
- Loadings generation
- Cross-sectional regression
- Time-series estimation

Statistical Models

Reminder: positive features of the spectrum of a matrix are:

- The spectrum is high and separated from the bulk. The spiked assumptions hold better
- Spike eigenvalues are separated. The eigenvectors are identifiable

Whitening with idio vol proxy helps

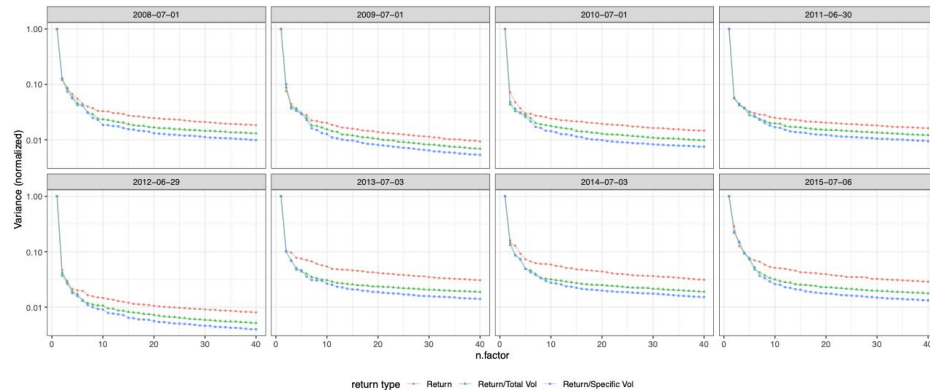


Figure 10.6: Variances of the eigenfactors (normalized to the variance of the first eigenfactor) for the first forty factors. Note that the scale of the y axis is logarithmic.

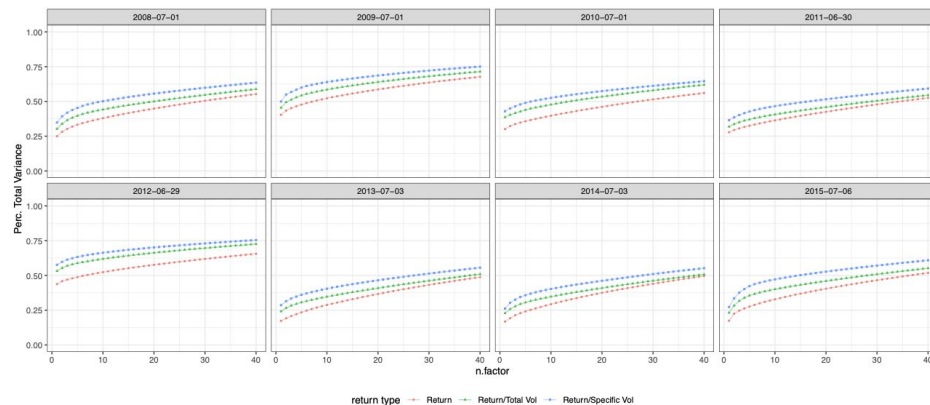


Figure 10.7: Cumulative percentage of variance described by the first n factors, for difference covariance matrices.

Additional Benefits of Idio Whitening: Returns

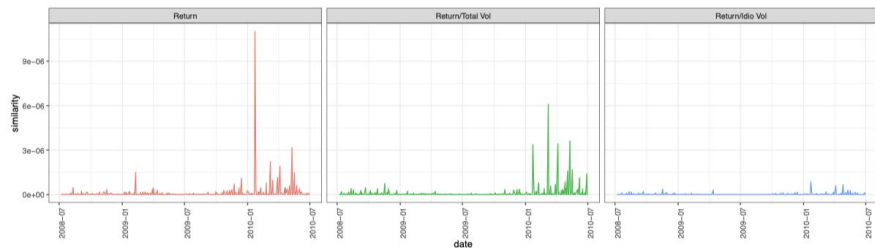
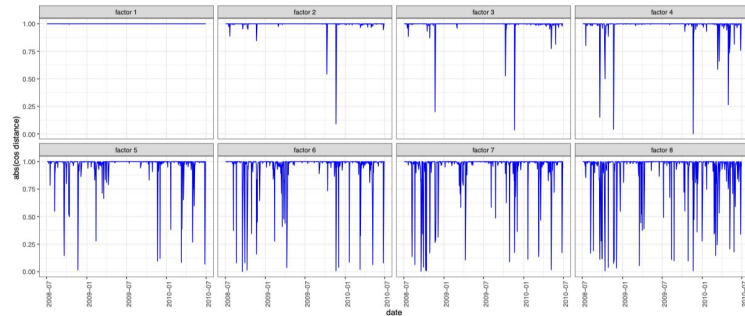
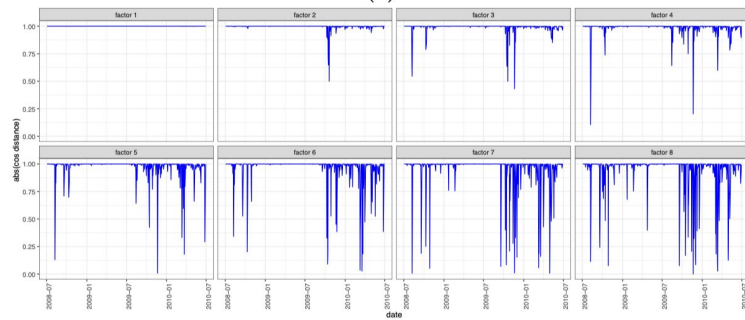


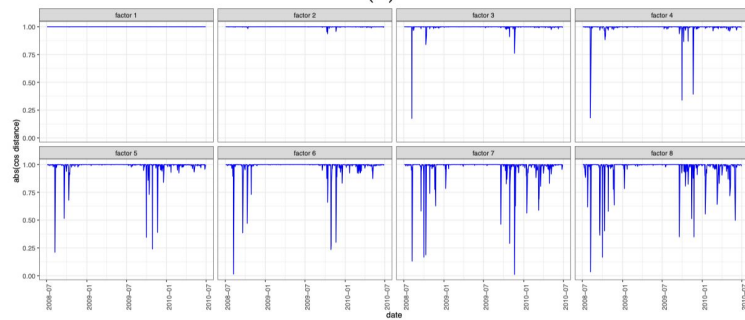
Figure 9.9: Distance between column subspaces of the first eight eigenfactors in consecutive periods. The eigenfactors are generated by PCAs on total returns, total returns/total vol, and total return/idio vol.



(a)



(b)



(c)

Additional Benefits of Idio Whitening: Returns

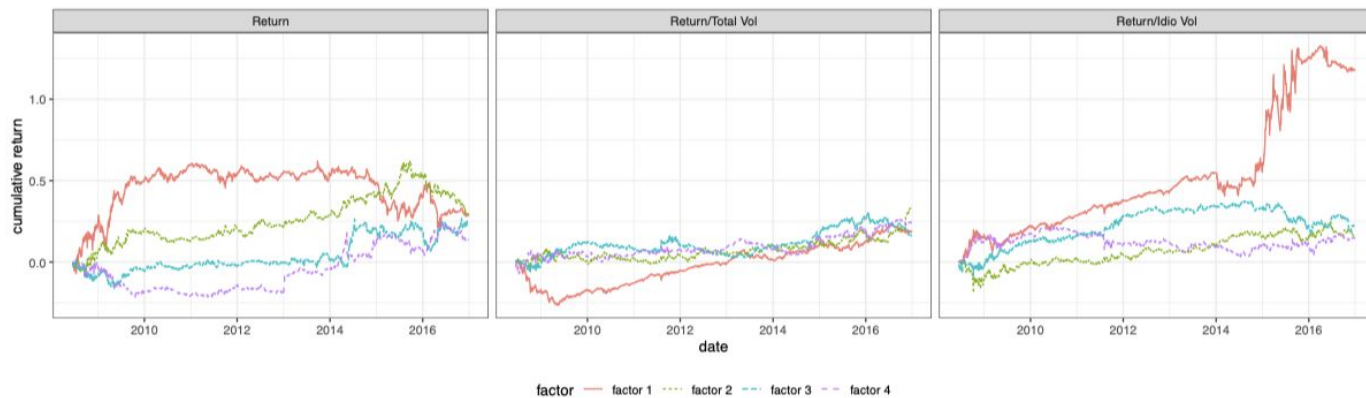


Figure 9.10: Factor returns for the first four eigenvectors. The eigenfactors are generated by PCAs on total returns, total returns/total vol, and total return/idio vol.

Reducing Turnover

Factors are turn over a lot. They can also change sign. They can be stabilized by rotating them from one period to the next.

$$\begin{aligned}\mathbf{B}_0 &:= \mathbf{B}_0 \\ \tilde{\mathbf{B}}_{t+1} &= \arg \min \|\mathbf{B}_t - \mathbf{Y}\|_F^2 \\ \text{s.t. } \mathbf{Y} &= \mathbf{B}_{t+1}\mathbf{X} \\ \mathbf{X}'\mathbf{X} &= \mathbf{I}_m \\ \mathbf{X} &\in \mathbb{R}^{m \times m}\end{aligned}$$

There is a clever solution for this problem.

Read the notes for details

$$\mathbf{A} := \mathbf{B}_t' \mathbf{B}_{t+1}$$

$$\mathbf{A} = \mathbf{U}\mathbf{S}\mathbf{V}'$$

$$\mathbf{X} = \mathbf{V}\mathbf{U}'$$

Putting it all Together

1. **Inputs:** $\mathbf{R} \in \mathbb{R}^{n \times T}$, $\tau_s \geq \tau_f > 0$, $p \in \mathbb{N}$, $m > 0$.

2. **Time-Series Reweighting:**

$$\mathbf{W}_{\tau_f} := \kappa \text{diag}(\exp(-T/\tau_f), \dots, \exp(-1/\tau_f))$$

$$\tilde{\mathbf{R}} = \mathbf{R} \mathbf{W}_{\tau_f}$$

3. **First stage PCA:** $\tilde{\mathbf{R}} := \tilde{\mathbf{U}} \tilde{\mathbf{S}} \tilde{\mathbf{V}}'$

4. **Idio Proxy Estimation:**

$$\mathbf{E} = \tilde{\mathbf{R}} - \tilde{\mathbf{U}}_p \tilde{\mathbf{S}}_p \tilde{\mathbf{V}}_p' \quad (\text{truncated SVD})$$

$$\sigma_i^2 = \sum_t [\mathbf{E}]_{i,t}^2 \quad (\text{idio vol proxies})$$

$$\mathbf{W}_\sigma := \text{diag}(\sigma_1^{-1}, \dots, \sigma_n^{-1})$$

5. **Idio Reweighting:**

$$\mathbf{W}_{\tau_s} := \kappa \text{diag}(\exp(-T/\tau_s), \dots, \exp(-1/\tau_s))$$

$$\hat{\mathbf{R}} := \mathbf{W}_\sigma \mathbf{R} \mathbf{W}_{\tau_s}$$

6. **Second Stage PCA:** $\hat{\mathbf{R}} := \hat{\mathbf{U}} \hat{\mathbf{S}} \hat{\mathbf{V}}'$

7. **Second-Stage Factor Model:** $\hat{\mathbf{r}} = \hat{\mathbf{U}}_m \mathbf{f} + \hat{\epsilon}$

where: $\mathbf{f} \sim N(0, \text{diag}(\ell(s_1^2), \dots, \ell(s_m^2)))$

$\epsilon \sim N(0, \bar{\lambda} \mathbf{I}_n)$

$$\bar{\lambda} = \frac{1}{n-m} \sum_{i=m+1}^n s_i^2$$

8. **Output: Final Factor Model:** $\mathbf{r} := \mathbf{B} \mathbf{f} + \epsilon$

where: $\mathbf{B} = \mathbf{W}_\sigma^{-1} \hat{\mathbf{U}}_m$

$\mathbf{f} \sim N(0, \text{diag}(\ell(s_1^2), \dots, \ell(s_m^2)))$

$\epsilon \sim N(0, \bar{\lambda} \mathbf{W}_\sigma^{-2})$

Fundamental Models

Takeaways for Fundamental Models

- At their core, very simple: **at their core**, alpha and risk modeling are very simple
- There are many details
 - Data
 - Shrinkage
 - Heteroskedastic and correlation correction
 - Regime Change

Steps

Steps:

1. Data ingestion: integrity, temporal continuity, missingness, outliers
2. Estimation universe: data quality, continuity, liquidity, relevance
3. Winsorization
4. Loadings Generation
5. Cross-Sectional Regression
6. Covariance Generation

Cross-Sectional Regression $\mathbf{r}_t = \mathbf{B}_t \mathbf{f}_t + \boldsymbol{\epsilon}_t, \quad t \in \mathbb{N}$

Assumptions:

1. Loadings are full rank
2. Homoskedastic idio returns are homoskedastic (jkg)
3. Factor returns and idio returns are independent
4. At least fourth moments are finite

Basics of X-sec regression

Gaussian Likelihood

$$\prod_{t=1}^T \frac{1}{(2\pi)^v |\boldsymbol{\Omega}_{\epsilon}|} \exp \left(-\frac{1}{2} (\mathbf{r}_t - \mathbf{B}\mathbf{f}_t)' \boldsymbol{\Omega}_{\epsilon}^{-1} (\mathbf{r}_t - \mathbf{B}\mathbf{f}_t) \right)$$

$$\begin{aligned} \min \quad & \left\| \boldsymbol{\Omega}_{\epsilon}^{-1/2} (\mathbf{R} - \mathbf{B}\mathbf{F}) \right\|^2 \\ \text{s.t. } & \mathbf{F} \in \mathbb{R}^{m \times T} \end{aligned} \quad \hat{\mathbf{f}}_t = (\mathbf{B}' \boldsymbol{\Omega}_{\epsilon}^{-1} \mathbf{B})^{-1} \mathbf{B}' \boldsymbol{\Omega}_{\epsilon}^{-1} \mathbf{r}_t$$

Note the similarity with PCA

Note that the estimation is exactly the same as factor mimicking portfolios

And this is why you should use squared loss

Shrinkage

- Factor volatilities are biased upwards because they are regression coefficients with errors, so they should be shrunk:

$$\text{var}(\hat{\mathbf{f}}_t) = \mathbf{\Omega}_f + \frac{1}{T}(\mathbf{B}'\mathbf{\Omega}_\epsilon^{-1}\mathbf{B})^{-1}$$

- Anderson's asymptotics for $T \rightarrow \infty$ doesn't apply. Recommended: simple Ledoit-Wolf. Let's not overkill.

$$\mathbf{\Omega}_{\text{shrink}}(\rho) = (1 - \rho_1)\hat{\mathbf{\Omega}}_f + \rho_2 \frac{\text{trace}(\hat{\mathbf{\Omega}}_f)}{m} \mathbf{I}_m$$

Separation Volatility-Correlation

Correlations change at a slower time-scale than volatilities

When people say, "in a crisis all correlations go to 1"... they should think twice [discuss!]

$$\Omega_f = \mathbf{V} \mathbf{C} \mathbf{V}$$

$$\text{diag}(\mathbf{V}_t^2) = \kappa_V \sum_{s=0}^T e^{-s/\tau_V} \hat{\mathbf{f}}_{t-s} \circ \hat{\mathbf{f}}_{t-s}$$

$$\mathbf{C} := \kappa_C \sum_{s=0}^T e^{-s/\tau_C} \mathbf{V}_{t-s}^{-1} \hat{\mathbf{f}}_{t-s} \hat{\mathbf{f}}_{t-s}' \mathbf{V}_{t-s}^{-1}$$

Dynamic Factor Volatility

Add a "modulating term".

The intuition is that cross-sectional factor returns are pretty large, and responsive.

If you get them wrong, adjust.

$$\begin{aligned}\mathbf{f}_t &= e^{x_t/2} \mathbf{C}_t^{1/2} \mathbf{V}_t \boldsymbol{\eta}_t & \mathbf{u}_t &:= \mathbf{V}_t^{-1} \mathbf{C}_t^{-1/2} \mathbf{f}_t \\ \boldsymbol{\eta}_t &\sim N(\mathbf{0}, \mathbf{I}_n) & \kappa &:= E(\log \|\boldsymbol{\eta}_t\|^2) \\ x_{t+1} &= \phi \xi_t + \sigma \gamma_t & \epsilon_t &:= \kappa - \log \|\boldsymbol{\eta}_t\|^2 \\ & & x_t &:= \xi_t \\ & & y_t &:= \log \|\mathbf{u}_t\|^2 - \kappa\end{aligned}$$

$$e^{\hat{x}_t/2} = \kappa_0 \exp \frac{1}{2} \left(\sum_{s=0}^{\infty} e^{-s/\tau_0} (\log \|\mathbf{u}_{t-s}\|^2 - \kappa) \right)$$

$$\simeq \kappa_0 \sum_{s=0}^{\infty} e^{-s/\tau_0} \frac{\|\mathbf{u}_{t-s}\|}{\sqrt{m}}$$

Autocorrelation Correction

You need this. Example: Asynchronous Returns.

$$(\mathbf{C}_l)_{i,j} = \text{cov}(\mathbf{f}_{t,i}, \mathbf{f}_{t-l,j})$$

Scholes correction:

$$\boldsymbol{\Omega}_{\mathbf{f}} = \hat{\boldsymbol{\Omega}}_{\mathbf{f}} + \frac{1}{2} \sum_{l=1}^{l_{\max}} (\mathbf{C}_l + \mathbf{C}'_l)$$

Newey-West correction:

$$\boldsymbol{\Omega}_{\mathbf{f}} = \hat{\boldsymbol{\Omega}}_{\mathbf{f}} + \frac{1}{2} \sum_{l=1}^{l_{\max}} \left(1 - \frac{l}{l_{\max}}\right) (\mathbf{C}_l + \mathbf{C}'_l)$$

Winsorization

- Ideally, choose the right estimation/investment universe and **never** winsorize
- In practice, the investment universe selection is imperfect
 - Bad data sources
 - Illiquid assets
 - Bankruptcies relistings
- So use a simple measure

$$d_{i,t} = \frac{|\log(1 + r_{i,t})|}{\text{median}(\log(1 + r_{i,t-1}), \dots, \log(1 + r_{i,t-T}))}$$