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Positive semi-definite correlation matrix completion

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Abstract

We give an intuitive derivation for the correlation matrix completion algorithm suggested in [KG06]. This leads us to a more general formula for the completion. The presented extension is positive semi-definite by construction, but we also give a simplified algebraic proof for its universal validity.

1 Introduction

Since the nature of this note is to present an extension to [KG06], we skip the general motivation and background of the problem and refer the reader to the references [KG06, Kah07].

Given a set of $2n$ standard normal variates x_1, \dots, x_n , and y_1, \dots, y_n , and the constraint that the pairwise correlations

$$\langle x_i x_j \rangle = r_{ij} \quad (1)$$

$$\langle x_i y_i \rangle = \eta_i \quad (2)$$

are pre-specified, we seek a completion of the as yet under-specified (auto-)correlation matrix of

$$z := (x_1, \dots, x_n, y_1, \dots, y_n)^\top \quad (3)$$

which has the structure

$$\langle z \cdot z^\top \rangle = \begin{pmatrix} r_{11} & \dots & r_{1n} & \eta_1 & & ? \\ \vdots & \ddots & \vdots & & \ddots & \\ r_{n1} & \dots & r_{nn} & ? & & \eta_n \\ \eta_1 & & ? & 1 & & ? \\ & \ddots & & & \ddots & \\ ? & & \eta_n & ? & & 1 \end{pmatrix} = \begin{pmatrix} R & B \\ B^\top & C \end{pmatrix} \quad (4)$$

with $r_{ii} = 1$.

2 Pairwise Cholesky construction

We start our intuition with the suggestion that each of the y_i can be represented as a linear combination

of x_i and a further standard normal variate ϵ_i which is independent from all the x_j , as given by a pairwise Cholesky decomposition:

$$y_i = \eta_i x_i + \eta'_i \epsilon_i, \quad (5)$$

$$\langle x_i \epsilon_j \rangle = 0, \quad (6)$$

with

$$\eta'_i := \sqrt{1 - \eta_i^2}. \quad (7)$$

This immediately yields

$$\langle x_i y_j \rangle = r_{ij} \eta_j \quad (8)$$

whence we choose $B := RH$ with

$$H := \text{diag}(\eta_1, \dots, \eta_n) \quad (9)$$

as in [KG06]. Further, we have

$$c_{ij} = \langle y_i y_j \rangle = \eta_i r_{ij} \eta_j + \eta'_i \langle \epsilon_i \epsilon_j \rangle \eta'_j. \quad (10)$$

We note that for $\langle \epsilon_i \epsilon_j \rangle = 0$ we obtain the structure given in [KG06].

3 Positive semi-definiteness

Since the matrices B and C given in the previous section are derived from linear combinations of standard normal variates, the completed matrix

$$A := \langle z \cdot z^\top \rangle = \begin{pmatrix} R & B \\ B^\top & C \end{pmatrix} \quad (11)$$

is *by construction* symmetric positive semi-definite, which we denote as

$$A \succeq 0. \quad (12)$$

However, for the sake of completeness, we provide below a simple algebraic proof.

Given two matrices $R, E \in \mathbb{R}^{n \times n}$, with $R \succeq 0, E \succeq 0$, we set

$$A := \begin{pmatrix} R & RH \\ HR & C \end{pmatrix} \quad (13)$$

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with

$$C := HRH + H'EH', \quad (14)$$

$$H' := \text{diag}(\eta'_1, \dots, \eta'_n). \quad (15)$$

Since the spectrum of A is invariant with respect to the addition of one of its (scaled) rows to any other, and likewise for columns, Gaussian elimination gives us

$$\begin{pmatrix} R & RH \\ HR & C \end{pmatrix} \succeq 0 \quad (16)$$

$$\begin{pmatrix} R & RH \\ 0 & C - HRH \end{pmatrix} \succeq 0 \quad (17)$$

$$\begin{pmatrix} R & 0 \\ 0 & C - HRH \end{pmatrix} \succeq 0. \quad (18)$$

Since $R \succeq 0$, equation (18) holds if $C - HRH \succeq 0$. This, however, follows trivially since

$$C - HRH = HRH + H'EH' - HRH \quad (19)$$

$$= H'EH' \succeq 0. \quad (20)$$

4 Summary

We showed how, given a correlation structure R for n standard normal variates x_1, \dots, x_n , and given the correlations $\langle x_i y_i \rangle = \eta_i$ to a second set of standard normal variates y_1, \dots, y_n , one can *constructively* arrive at

$$b_{ij} = \langle x_i y_j \rangle = r_{ij} \eta_i \quad (21)$$

$$c_{ij} = \langle y_i y_j \rangle = \eta_i r_{ij} \eta_j + \eta'_i e_{ij} \eta'_j \quad (22)$$

for an arbitrary correlation matrix $E \in \mathbb{R}^{n \times n}$, $E \succeq 0$, as a possible choice for the completed correlation matrix $A = \begin{pmatrix} R & B \\ B^\top & C \end{pmatrix}$. We also proved

$$\begin{pmatrix} R & RH \\ HR & HRH + H'EH' \end{pmatrix} \succeq 0 \quad (23)$$

by the aid of straightforward Gaussian elimination of rows and columns.

It remains to be said that in practice one may wish to use the homogenous parametric form

$$e_{ij} = \beta + (1 - \beta) \delta_{ij} \quad (24)$$

for E , with $\beta \in \left[-\frac{1}{n-1}, 1\right]$ and $\delta_{(\cdot)}$ being the Kronecker symbol, for the sake of simplicity.

References

- [Kah07] C. Kahl. *Modeling and simulation of stochastic volatility in finance*. PhD thesis, Bergische Universität Wuppertal and ABN AMRO, 2007. Published by www.dissertation.com, www.amazon.com/Modelling-Simulation-Stochastic-Volatility-Finance/dp/1581123833/, ISBN-10: 1581123833.
- [KG06] C. Kahl and M. Günther. Complete the Correlation Matrix. Working paper, Bergische Universität Wuppertal, 2006. www.math.uni-wuppertal.de/~kahl/publications/CompleteTheCorrelationMatrix.pdf.