#### THIS PAPER IS NOT TO BE REMOVED FROM THE EXAMINATION HALL



EC2020 ZA

BSc DEGREES AND GRADUATE DIPLOMAS IN ECONOMICS, MANAGEMENT, FINANCE AND THE SOCIAL SCIENCES, THE DIPLOMA IN ECONOMICS AND SOCIAL SCIENCES AND THE CERTIFICATE IN EDUCATION IN SOCIAL SCIENCES

#### **Elements of Econometrics**

Thursday 23 May 2019: 10.00 – 13.00

Time allowed: 3 hours

#### DO NOT TURN OVER UNTIL TOLD TO BEGIN

Candidates should answer **EIGHT** of the following **TEN** questions: **ALL** of the questions in Section A (8 marks each) and **THREE** questions from Section B (20 marks each). **Candidates are strongly advised to divide their time accordingly.** 

Extracts from statistical tables are given after the final question on this paper.

Graph paper is provided at the end of this question paper. If used, it must be detached and fastened securely inside the answer book.

A handheld calculator may be used when answering questions on this paper and it must comply in all respects with the specification given with your Admission Notice. The make and type of machine must be clearly stated on the front cover of the answer book.

If more questions are answered than requested, only the first answers attempted will be counted.

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### **SECTION A**

#### Answer all questions from this section

1. Consider the following regression model

$$Y_t = \beta_0 + \beta_1 X_t + \beta_2 t + \varepsilon_t, \qquad t = 1, ..., T.$$

Both  $\{Y_t\}_{t=1}^T$  and  $\{X_t\}_{t=1}^T$  are trend stationary processes. The errors  $\{\varepsilon_t\}_{t=1}^T$  are independent random variables with zero mean and constant variance.

- (a) **(4 marks)** Discuss the concept "trend stationarity" and contrast it to the concept "difference stationarity". In your answer make sure you also explain what stationarity means.
- (b) **(4 marks)** Provide a clear interpretation of the parameter  $\beta_1$ . You are told that you can obtain the estimator for  $\beta_1$  using "detrended" variables only. Discuss this statement.
- 2. Consider the following ADL(1,1) model relating the crime rate in a particular province,  $crime_t$ , to the clear-up rate (percentage of crimes resulting in a conviction)

$$\mathit{crime}_t = \alpha + \rho \mathit{crime}_{t-1} + \delta_1 \mathit{clearup}_t + \delta_2 \mathit{clearup}_{t-1} + u_t, \text{ with } |\rho| < 1$$

where  $u_t$  is white noise, an i.i.d. innovation that is uncorrelated to anything in the past.

- (a) (2 marks) Briefly indicate whether OLS will provide unbiased and consistent parameter estimators.
- (b) (2 marks) Derive the long run relationship between crime and clearup.
- (c) **(4 marks)** Rewrite the model above in terms of an error correction model (ECM) and interpret its coefficients.
- 3. Consider the simple linear regression model

$$Y_i = \beta X_i + u_i, \quad i = 1, ..., n.$$

We assume that the errors  $\{u_i\}_{i=1}^n$  are independent normal random variables with zero mean. The regressor  $\{X_i\}_{i=1}^n$  is non-stochastic (fixed under repeated sampling). You suspect that the errors exhibit heteroskedasticity.

- (a) **(3 marks)** Explain what we mean by the concept of heteroskedasticity. Enhance your answer with the help of a graphical illustration.
- (b) **(5 marks)** Suppose you want to test  $H_0: \beta = 1$  against  $H_1: \beta > 1$ . Discuss how you would conduct this test based on the OLS estimator, recognizing the presence of heteroskedasticity. Please provide a detailed answer.

4. Consider the simple linear regression model

$$Y_i = \beta_0 + \beta_1 X_i + u_i, \qquad i = 1, ..., n.$$

We assume that the errors  $\{u_i\}_{i=1}^n$  are independent random variables with zero mean. The regressor  $\{X_i\}_{i=1}^n$  is non-stochastic (fixed under repeated sampling). Under these conditions, the OLS estimator for  $\beta_1$ ,  $\hat{\beta}_1$ , is unbiased. (You are not asked to derive  $\hat{\beta}_1$ ).

- (a) (2 marks) Explain the concept of unbiasedness of an estimator.
- (b) **(4 marks)** Let us consider two other estimators for the slope  $\beta_1$ :

$$\hat{\beta}_{1}^{\circ} = \frac{\sum_{i=1}^{n} (Z_{i} - \bar{Z}) Y_{i}}{\sum_{i=1}^{n} (Z_{i} - \bar{Z}) X_{i}}; \quad \hat{\beta}_{1}^{*} = \frac{\sum_{i=1}^{n} (Z_{i} - \bar{Z}) Y_{i}}{\sum_{i=1}^{n} (Z_{i} - \bar{Z}) Z_{i}}$$

where  $Z_i = \sqrt{X_i}$  for all i and  $\bar{Z} = \frac{1}{n} \sum_{i=1}^n Z_i$ . Please indicate whether  $\hat{\beta}_1^{\circ}$  and  $\hat{\beta}_1^*$  are unbiased estimators for  $\beta_1$ . Clearly show your derivations.

- (c) **(2 marks)** Briefly indicate how you would choose between the three estimators,  $\hat{\beta}_1$ ,  $\hat{\beta}_1^{\circ}$  and  $\hat{\beta}_1^{*}$ .
- 5. Consider the OLS estimator for  $\beta$  in the linear regression model

$$Y_i = \beta X_i + \varepsilon_i, \qquad i = 1, ..., n,$$

where  $\{(Y_i, X_i)\}_{i=1}^n$  form an i.i.d. sample from a population and the errors are drawn from an **unknown** distribution with mean zero and variance  $\sigma^2$ . You are told that  $X_i$  and  $\varepsilon_i$  are **uncorrelated** (not necessarily independent therefore!).

- (a) (4 marks) Discuss the importance of convergence in probability.
- (b) (4 marks) Discuss the importance of convergence in distribution.

#### SECTION B

#### Answer three questions from this section.

- 6. Let us consider how workplace smoking bans affect the incidence of smoking. Below, we use data on 10,000 US indoor workers from 1991 to 1993 taken from "Do Workplace Smoking Bans Reduce Smoking", by Evans et al. (*American Economic Review*, 1999). Let *smoker* be a dummy variable indicating whether a worker smokes (1=yes, 0=no) and *smkban* a dummy variable indicating whether there is a ban on smoking in the workplace (1=yes, 0=no).
  - (a) (6 marks) The following OLS regression results were obtained

$$\widehat{\text{smoker}} = 0.290 - 0.078 \, \text{smkban}$$
 (6.1)  
 $n = 10000, R^2 = 0.0078, RSS = 1821.59$ 

The standard errors (SEs) are in parentheses. Interpret the parameter estimates of the coefficient on *smkban*. Provide the (approximate) 95% confidence interval for the coefficient on *smkban*. How can we use this confidence interval to test the hypothesis that  $\beta_{smkban} = 0$ ?

A further specification was considered that included other characteristics of the worker: the age (in years), gender (male/female), ethnicity (black/hispanic/white), and level of education (E1=highschool dropout, E2=highschool graduate, E3=some college, E4=college graduate, E5 Master degree or above). The following OLS regression results were obtained for this multiple linear regression model

$$\widehat{smoker} = \underbrace{0.201 - 0.045}_{(0.019)} \underbrace{smkban - 0.033female}_{(.009)} - \underbrace{0.001}_{(.0003)} \underbrace{age - 0.027}_{(.016)} - \underbrace{0.104}_{(.014)} \underbrace{hispanic}_{(.019)} + \underbrace{0.310}_{(.012)} + \underbrace{0.224}_{(.012)} + \underbrace{0.156}_{(.012)} \underbrace{E3}_{(.012)} + \underbrace{0.042}_{(.012)} \underbrace{E4}_{(.012)}$$

$$n = 10000, R^2 = 0.0526, RSS = 1736.81$$

The SEs are in parentheses.

- (b) **(3 marks)** Compare the coefficient estimates on *smkban* from the simple and multiple regression model in (6.1) and (6.2) and explain why the estimates differ.
- (c) **(5 marks)** Interpret the estimated parameter on E2 (highschool graduate) in (6.2) and indicate how you can obtain its p-value and what information the p-value provides.
- (d) **(6 marks)** Both the simple and multiple regression model suffer from heteroskedasticity. Explain why. What are the implications of heteroskedasticy for the parameter estimates and the standard errors in (6.1) and (6.2)? What can you do to resolve this problem? Explain your answer.

7. In question 6 we considered the linear regression model to study how workplace smoking bans affect the incidence of smoking. Here we consider the results from applying a probit regression of *smoker* (1=yes, 0=no) on *smkban* (1=yes, 0=no) and the other explanatory variables:

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smkban	1517626	.0289268	-5.25	0.000	208458	0950671
female	1106249	.0287785	-3.84	0.000	1670298	05422
age	0042031	.0011748	-3.58	0.000	0065057	0019006
black	07969	.0525369	-1.52	0.129	1826604	.0232804
hispanic	3327039	.0476677	-6.98	0.000	4261308	2392769
E1	1.094231	.0714121	15.32	0.000	.9542663	1.234197
E2	.8518588	.0594747	14.32	0.000	.7352906	.9684271
E3	.6492566	.0606989	10.70	0.000	.530289	.7682241
E4	.2224747	.0649939	3.42	0.001	.0950891	.3498603
cons	9842425	.0756055	-13.02	0.000	-1.132427	8360584

- (a) **(5 marks)** It is argued that using the probit regression model is better than using the linear probability model when explaining the binary variable *smoker*. Discuss the benefits/drawbacks of using the Probit model when trying to explain a binary variable.
- (b) **(5 marks)** Explain briefly how the Probit estimates are obtained and discuss the properties of the parameter estimates.

Hint: You may recall that for the Probit model, we will specify

$$\Pr(\mathsf{smoker} = 1) = \Phi(\beta_0 + \beta_1 \mathsf{smkban} + \beta_2 \mathsf{female} + \dots + \beta_8 E3 + \beta_9 E4)$$

where  $\Phi$  is the standard normal CDF (cumulative distribution function).

- (c) **(5 marks)** Explain how you can estimate the effect of the smoking ban on the probability of smoking for a 50-year old white, college graduated man. You are not expected to use your calculator, clarity of the computations required is enough.
- (d) (5 marks) Discuss how you could test the joint significance of the worker's characteristics (gender, age, ethnicity and level of education) using the likelihood ratio test. Clearly indicate the test statistic, its distribution, the rejection rule and the additional information you would need to implement it.

8. An economist is interested in estimating the production function for widgets which is postulated to follow a Cobb-Douglas specification:

$$Y_i = \exp(\beta_0) L_i^{\beta_L} K_i^{\beta_K} \exp(u_i),$$

where  $Y_i$  is a measure of output for firm i,  $L_i$  is labor,  $K_i$  is capital stock and  $u_i$  is an unobserved term that captures technological or managerial efficiency and other external factors (e.g., weather). The parameters to be estimated are  $(\beta_0, \beta_L, \beta_K)$ . Taking logs,

$$\ln Y_i = \beta_0 + \beta_L \ln L_i + \beta_K \ln K_i + u_i.$$

- (a) **(5 marks)** Provide the interpretation of the parameter  $\beta_L$ . What effect, if any, will changing the units of measurement of labor have on the parameter estimates for  $(\beta_0, \beta_L, \beta_K)$ ? Explain your answer.
- (b) **(5 marks)** Assume that you have a cross-section of firms and that more productive firms hire less workers (labor). Explain why OLS would not provide consistent estimates for  $(\beta_0, \beta_L, \beta_K)$ . Would it over- or underestimate  $\beta_L$  on average? Clearly explain your answer.

Instead of applying OLS, the economist decides to use the average wage paid by firm i,  $W_i$ , as an instrument for the (log) quantity of labor employed by that firm,  $\ln L_i$ .

- (c) **(7 marks)** Describe in detail how you would estimate the parameters of the production function using Two Stage Least Squares (TSLS). What restrictions would be necessary for this researcher to successfully use this instrumental variable in the estimation of the parameters  $(\beta_0, \beta_L, \beta_K)$  and what would you need to assume about capital stock?
- (d) (3 marks) If average wages per firm do not vary much by firm (potentially because of unionization or high mobility of the labor force), how would this affect the properties of the estimation procedure suggested in (c)? Explain your answer.

9. Let us consider the expectations augmented Phillips curve (see also Mankiw, 1994):

$$infl_t - infl_t^e = \beta_1 (unem_t - \mu_0) + e_t,$$

where  $\mu_0$  is the natural rate of unemployment (assumed to be constant over time) and  $infl_t^e$  is the expected rate of inflation formed in t-1.

This model suggests that there is a tradeoff between unanticipated inflation  $(infl_t-infl_t^e)$  and cyclical unemployment (difference between actual unemployment and the natural rate of unemployment). We assume that  $e_t$  (also called supply shock) is an i.i.d. random variable with zero mean.

(a) (2 marks) You are told that expectations are formed as follows

$$infl_t^e - infl_{t-1}^e = \lambda \left( infl_{t-1} - infl_{t-1}^e \right)$$

What name do we give such a process and how should we interpret  $\lambda$ ?

(b) (7 marks) Show that you can rewrite the model as

$$\Delta infl_t = \gamma_0 + \gamma_1 unem_t + \gamma_2 unem_{t-1} + v_t, \tag{9.1}$$

where  $\Delta \mathit{infl}_t = \mathit{infl}_t - \mathit{infl}_{t-1}$ . Clearly indicate the relation between  $(\gamma_0, \gamma_1, \gamma_2)$  and  $(\mu_0, \beta_1, \lambda)$  and show that

$$v_t = e_t - (1 - \lambda) e_{t-1}$$
.

*Hint:* If you want you may use the following shorthand notation in your derivations:  $y_t = \inf_t y_t^e = \inf_t x_t^e$  and  $x_t = unem_t$ .

- (c) (3 marks) Discuss what assumptions you would like to make about  $e_t$  (the supply shock) that will guarantee that the OLS estimator for the parameters in (9.1) is consistent. *Hint*: You may want to give the assumptions you need to make about  $v_t$  (composite error term) first.
- (d) **(3 marks)** Show how you can obtain a consistent estimator for  $\lambda$  using your consistent estimates for  $(\gamma_0, \gamma_1, \gamma_2)$ ? Provide a proof of its consistency. [*Note:* If you did not manage to get an explicit relation between  $\lambda$  and  $(\gamma_0, \gamma_1, \gamma_2)$ , consider  $\lambda = g(\gamma_0, \gamma_1, \gamma_2)$  where  $g(\cdot)$  is some continuous function].
- (e) (5 marks) One of the assumptions provided in (c) is rather unreasonable (*Hint*: future unemployment may be related to current supply shocks). Discuss how you could use IV (TSLS) to obtain a consistent estimator for the parameters in (9.1). Discuss what conditions you instruments need to satisfy and propose a suitable instrument in this setting.

10. This question is based on "Capital Accumulation and Growth: A New Look at the Empirical Evidence", by Bond et al. (*Journal of Applied Econometrics*, 2010). In this article, the authors are interested in a regression model for the (logarithm of) output-per-capita  $y_t$  in a given country and time period t similar to:

$$y_t = \alpha + \rho y_{t-1} + \beta x_t + \gamma t + \varepsilon_t, \quad |\rho| < 1, \tag{10.1}$$

where  $x_t$  is (the logarithm of) investment-per output. Imagine that investment rates are also affected by current output-per-capita, so that

$$x_t = \varphi_0 + \varphi_1 y_t + u_t. \tag{10.2}$$

Equations (10.1) and (10.2) then form a simultaneous equation model. Both errors  $\varepsilon_t$  and  $u_t$  have zero mean.

For (a) and (b), we start by assuming that  $\varepsilon_t$  and  $u_t$  are not serially correlated.

- (a) **(5 marks)** Obtain the reduced form equations for  $y_t$  and  $x_t$  (Note: the variables t and  $y_{t-1}$  are exogenous). Are equations (10.1) and (10.2) identified? Discuss.
- (b) **(5 marks)** Assume that  $x_t = \varphi_0 + u_t$  (that is  $\varphi_1 = 0$ ). Under what conditions is the OLS estimator for the parameters in the structural equation for output-per-capita (i.e., equation (10.1)) consistent? Discuss.
- (c) (5 marks) Assume that  $x_t = \varphi_0 + u_t$  (that is  $\varphi_1 = 0$ ),  $u_t$  is not serially correlated, and  $\gamma = 0$ . How would you test whether  $\varepsilon_t$  is serially correlated?
- (d) **(5 marks)** Assume that  $x_t = \varphi_0 + u_t$  (that is  $\varphi_1 = 0$ ),  $\rho = 0$  and  $\gamma = 0$ . Using the Dickey-Fuller test, the authors fail to reject the hypothesis of a unit root in  $y_t$  and  $x_t$  for most of the countries in their sample at usual significance levels (5%). (Note: we are no longer assuming  $\varepsilon_t$  or  $u_t$  to not be serially correlated). Explaining your answers, what can you say about the OLS estimator for and

applied to equation (10.1)

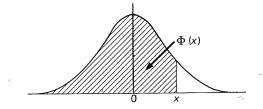
- i. if  $\varepsilon_t$  is non-stationary and strongly dependent?
- i. if  $\varepsilon_t$  is stationary and weakly dependent?

**END OF PAPER** 

# TABLE 4. THE NORMAL DISTRIBUTION FUNCTION

The function tabulated is  $\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{x} e^{-\frac{1}{2}t^2} dt$ .  $\Phi(x)$  is

the probability that a random variable, normally distributed with zero mean and unit variance, will be less than or equal to x. When x < 0 use  $\Phi(x) = \mathbf{1} - \Phi(-x)$ , as the normal distribution with zero mean and unit variance is symmetric about zero.



x	$\Phi(x)$	$\boldsymbol{x}$	$\Phi(x)$	×	$\Phi(x)$	x	$\Phi(x)$	x	$\Phi(x)$	x	$\Phi(x)$
			. ,		` ,		` ,				, ,
0.00	0.2000	0.40	0.6554	0.80	0.7881	1.30	0.8849	1.60	0.9452	2.00	0.97725
.OI	·504 <b>0</b>	·41	.6591	·81	.7910	.21	·8869	·61	.9463	.01	.97778
.02	·5080	.42	·6628	·82	.7939	.22	.8888	.62	9474	.02	·97831 ·97882
.03	.2120	·43	.6664	.83	.7967	.53	.8907	.63	9484	.03	
·04	.2160	·44	.6700	·8 <b>4</b>	.7995	.24	.8925	·64	·9495	·04	.97932
0.02	0.2199	0.45	0.6736	o·85	0.8023	1.25	0.8944	1.65	0.9505	2.05	0.97982
·06	.5239	·46	.6772	·86	·8051	.26	·8962	.66	.9515	∙06	·9803 <b>0</b>
.07	.5279	·47	·68o8	·8 <sub>7</sub>	·8o78	·27	·8980	·6 <sub>7</sub>	.9525	· <b>07</b>	.98077
·08	.5319	·48	·6844	.88	·8106	.28	·8997	.68	.9535	.08	.98124
.09	.5359	·49	.6879	.89	.8133	.29	.9015	.69	·954 <b>5</b>	.09	·98169
0.10	0.5398	0.20	0.6915	0.00	0.8159	1.30	0.0032	1.70	0.9554	2.10	0.98214
.11	.5438	.21	.6950	.01	·8186	.31	.9049	.71	.9564	·II	98257
·12	·5478	.52	.6985	.92	.8212	.32	.9066	.72	.9573	.12	·98300
.13	.2212	.53	.7019	.93	.8238	33	.9082	.73	.9582	.13	.98341
·14	5557	·54	.7054	·94	·8264	.34	.9099	<sup>.</sup> 74	.9591	•14	.98382
0.77	0.5596	0.55	0.7088	0.95	0.8289	1.35	0.9115	1.75	0.0299	2.15	0.98422
·16	.5636	o·55 ·56	7123	.96	.8315	.36	.9131	·76	.9608	.16	.98461
.17	.5675	·57	7123	·97	.8340	·37	9131	.77	.9616	.17	.98500
·18	·5714	.58	·7190	·98	·8365	.38	9162	·78	.9625	·18	.98537
.10	5753	.59	.7224	.99	.8389	.39	9177	.79	.9633	.10	.98574
-9	3733	39	/	99	030.9	37	<i>y-11</i>	• • •	7.00	Í	, , , , ,
0.30	0.5793	0.60	0.7257	1.00	0.8413	1.40	0.9192	1.80	0.9641	2.30	0.98610
.31	.5832	·61	.7291	.oı	·8 <b>43</b> 8	·4I	.9207	·81	-9649	.31	·98645
.22	·5871	·6 <b>2</b>	.7324	.02	·8461	42	.9222	·82	.9656	.22	.98679
.23	.5910	·6 <b>3</b>	.7357	.03	·8485	·43	·9236	·8 <sub>3</sub>	•9664	.53	.98713
·24	.5948	·6 <b>4</b>	.7389	.04	·8508	·44	.9251	·8 <sub>4</sub>	·9671	.24	·98745
0.25	0.5987	o·65	0.7422	1.05	0.8531	1.45	0.9265	r·85	0.9678	2.25	0.98778
·26	.6026	.66	.7454	.06	.8554	·46	.9279	⋅86	•9686	·26	·98809
.27	.6064	·6 <del>7</del>	.7486	.07	.8577	·47	.9292	·8 <sub>7</sub>	·969 <b>3</b>	·27	·9884 <b>0</b>
-28	.6103	.68	.7517	·08	.8599	·48	·9306	-88	-9699	·28	·98870
· <b>2</b> 9	.6141	·6 <b>9</b>	7549	.09	·8621	· <b>49</b>	.9319	.89	·97 <b>0</b> 6	.29	·9889 <b>9</b>
0.30	0.6179	0.70	0.7580	1.10	0.8643	1.20	0.9332	1.90	0.9713	2:30	0.98928
.31	.6217	·7I	.7611	·II	·8665	.21	.9345	·91	.9719	.31	.98956
.32	.6255	.72	.7642	·12	·8686	.52	9357	.92	.9726	.32	.98983
.33	.6293	.73	.7673	.13	·8708	.53	.9370	.93	.9732	.33	.99010
·34	.6331	·74	.7704	.14	8729	·54	.9382	·94	.9738	·34	.99036
0.32	0.6368	0.75	0.7734	1.12	0.8749	1.22	0.9394	1.95	0.9744	2.35	0.99061
.36	.6406	·76	·7764	.16	·8770	·56	·939 <del>4</del>	96	.9750	·36	.99086
.37	.6443	.77	7794	·17	·879 <b>0</b>	·57	·9418	.97	.9756	.37	.00111
·38	·648 <b>o</b>	·78	·7823	.18	.8810	·58	9429	.98	.9761	.38	.99134
.39	.6517	·79	.7852	.19	·8830	.59	.9441	.99	.9767	.39	.99158
0.40	0.6554	0.80	0.7881	1.30	o·8849	r·60	0.9452	2.00	0.9772	2.40	0.99180

## TABLE 4. THE NORMAL DISTRIBUTION FUNCTION

$\boldsymbol{x}$	$\Phi(x)$	$\boldsymbol{x}$	$\Phi(x)$	$\boldsymbol{x}$	$\Phi(x)$	$\boldsymbol{x}$	$\Phi(x)$	x	$\Phi(x)$	$\boldsymbol{x}$	$\Phi(x)$
2.40	0.99180	2.55	0.99461	2.70	0.99653	2.85	0.99781	3.00	0.99865	3.12	0.99918
·41	.99202	· <b>5</b> 6	<b>.</b> 99477	.71	·99664	∙86	.99788	·or	•99869	.16	.99921
.42	.99224	·57	.99492	.72	.99674	·8 <sub>7</sub>	.99795	.02	.99874	.17	99924
·43	.99245	·58	·99506	.73	·99683	-88	·99801	.03	.99878	٠18	.99926
·44	·99266	.29	.99520	·74	.99693	∙89	.99807	.04	·99882	.19	.99929
2.45	0.99286	2.60	0.99534	2.75	0.99702	2.90	0.99813	3.05	0.99886	3.30	0.99931
·46	.99302	·61	.99547	.76	.99711	.91	.99819	.06	.99889	21	199934
·47	·993 <del>24</del>	.62	·99560	.77	.99720	.92	.99825	.07	.99893	.22	.99936
·48	.99343	·63	.99573	.78	.99728	.93	•99831	.08	.99896	.23	.99938
· <b>4</b> 9	·99361	·6 <b>4</b>	.99585	.79	.99736	·94	·99836	.09	.99900	·24	.99940
2.20	0.99379	2.65	0.99598	2.80	0.99744	2.95	0.99841	3.10	0.99903	3.52	0.99942
.21	.99396	.66	•99609	·81	99752	·96	•99846	·II	.99906	· <b>26</b>	199944
.52	.99413	·6 <b>7</b>	·99621	· ·82	.99760	.97	·99851	.12	.99910	.27	99946
.53	·99430	∙68	·99632	.83	.99767	.98	.99856	.13	.99913	.28	.99948
·54	·99446	.69	.99643	·8 <sub>4</sub>	99774	.99	·99861	14	.99916	·29	.99950
2.55	0.99461	2.70	0.99653	2.85	0.99781	3.00	0.99865	3.12	0.99918	3.30	0.99952

The critical table below gives on the left the range of values of x for which  $\Phi(x)$  takes the value on the right, correct to the last figure given; in critical cases, take the upper of the two values of  $\Phi(x)$  indicated.

3:075	2.262 0.9994	3.731 0.99990 3.759 0.99991 3.791 0.99993 3.826 0.99993	3.916 0.99995
3.022 3.102 0.0003 3.022 0.0003	3.320 0.9994 3.320 0.9995	3 732 0.99991	3.976 0.99996 3.916 0.99997
3 103 0.9991	3 320 0.9996	3759 0.99992	3.970 0.99997
3 130 0.9992	3·389 0·9996 3·480 0·9997	3.791 0.00003	4.055 0.00008
3.174 0.0003	3.480 0.0008	3.826	4.173
3·174 0·9993 3·215 0·9994	3.615 0.9998	3.867 0.99994	4.055 0.99997 4.173 0.99999 4.417 1.00000

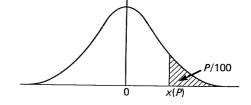
When x > 3.3 the formula  $1 - \Phi(x) = \frac{e^{-\frac{1}{x^2}}}{x\sqrt{2\pi}} \left[ 1 - \frac{1}{x^2} + \frac{3}{x^4} - \frac{15}{x^6} + \frac{105}{x^8} \right]$  is very accurate, with relative error less than  $945/x^{10}$ .

# TABLE 5. PERCENTAGE POINTS OF THE NORMAL DISTRIBUTION

This table gives percentage points x(P) defined by the equation

$$\frac{P}{{\rm 100}} = \frac{{\rm I}}{\sqrt{2\pi}} \int_{x(P)}^{\infty} \!\! e^{-\frac{1}{2}t^2} \, dt.$$

If X is a variable, normally distributed with zero mean and unit variance, P/100 is the probability that  $X \ge x(P)$ . The lower P per cent points are given by symmetry as -x(P), and the probability that  $|X| \ge x(P)$  is 2P/100.



P	x(P)	P	x(P)	P	x(P)	P	x(P)	$m{P}$	x(P)	P	x(P)
50	0.0000	5.0	1.6449	3.0	1.8808	2.0	2.0537	1.0	2.3263	0.10	3.0902
45	0.1257	4·8	1.6646	2.9	1.8957	1.0	2.0749	0.0	2.3656	0.00	3.1214
40	0.2533	4.6	1.6849	2.8	1.9110	1.8	2.0969	0.8	2.4089	0.08	3.1220
35	0.3823	4.4	1.7060	2.7	1.9268	1.7	2.1201	0.2	2.4573	0.07	3.1947
30	0.244	4.3	1.7279	2.6	1.9431	1.6	2.1444	0.6	2.2121	0.06	3.2389
25	0.6745	4.0	1.7507	2.5	1.9600	1.2	2.1701	0.2	2.5758	0.02	3.2905
20	0.8416	3.8	1.7744	2.4	1.9774	1.4	2.1973	0.4	2.6521	0.01	3.7100
15	1.0364	3.6	1.7991	2.3	1.9954	1.3	2.2262	0.3	2.7478	0.002	3.8906
10	1.5816	3.4	1.8250	2.2	2.0141	1.3	2.2571	0.3	2.8782	0.001	4.2649
5	1.6449	3.5	1.8522	2·I	2.0335	I.I	2.2904	0.1	3.0902	0.0002	4.4172

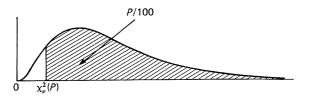
# TABLE 8. PERCENTAGE POINTS OF THE x²-DISTRIBUTION

This table gives percentage points  $\chi^2_{\nu}(P)$  defined by the equation

$$\frac{P}{100} = \frac{\mathrm{I}}{2^{\nu/2} \, \Gamma(\frac{\nu}{2})} \int_{\chi^2_{\nu}(P)}^{\infty} x^{\frac{1}{2}\nu - 1} \, e^{-\frac{1}{2}x} \, dx.$$

If X is a variable distributed as  $\chi^2$  with  $\nu$  degrees of freedom, P/100 is the probability that  $X \geqslant \chi^2_{\nu}(P)$ .

For  $\nu > 100$ ,  $\sqrt{2X}$  is approximately normally distributed with mean  $\sqrt{2\nu - 1}$  and unit variance.



(The above shape applies for  $\nu \geqslant 3$  only. When  $\nu < 3$  the mode is at the origin.)

P	99.95	99.9	99 <sup>.</sup> 5	99	97.5	95	90	8o	70	60
$\nu = \mathbf{I}$	o·o <sup>6</sup> 3927	0.021221	0.043927	0.031571	0.039821	0.003932	0.01579	0.06418	0.1485	0.2750
2	0.001000	0.002001	0.01003	0.02010	0.05064	0.1026	0.2107	0.4463	0.7133	1.022
3	0.01528	0.02430	0.07172	0.1148	0.2158	0.3518	0.5844	1.002	1.424	1.869
4	0.06392	0.00080	0.2070	0.2971	0.4844	0.7107	1.064	1.649	2.195	2.753
		,	•							
5	0.1281	0.5105	0.4117	0.5543	0.8312	1.145	1.610	2.343	3.000	3.655
ĕ	0.2994	0.3811	0.6757	0.8721	1.237	1.635	2.204	3.070	3.828	4.570
7	0.4849	0.5985	0.9893	1.239	1.690	2.167	2.833	3.822	4.671	5.493
8	0.7104	0.8571	1.344	1·646	2.180	2.733	3.490	4.294	5.527	6.423
9	0.9717	1.152	1.735	2.088	2.700	3.325	4·168	5.380	6.393	7:357
-				,						
10	1.265	1.479	2.156	2.558	3.247	3.940	4.865	6.179	7.267	8.295
11	1.587	1.834	2.603	3.023	3.816 *	4.575	5.578	6.989	8.148	9.237
12	1.934	2.214	3.074	3.221	4.404	5.226	6.304	7.807	9.034	10.18
13	2.302	2.617	3.565	4.102	5.009	5.892	7.042	8.634	9.926	11.13
14	2.697	3.041	4.075	4.660	5.629	6.571	7.790	9.467	10.82	12.08
							•			
15	3.108	3.483	4.601	5.229	6.262	7.261	8.547	10.31	11.72	13.03
16	3.236	3.942	5.142	5.812	6.908	7.962	9.312	11.12	12.62	13.98
17	3.980	4.416	5 697	6·408	7.564	8.672	10.00	12.00	13.23	14.94
18	4.439	4.902	6.265	7.015	8.531	9.390	10.86	12.86	14.44	15.89
19	4.912	5.407	6.844	7.633	8.907	10.13	11.65	13.72	15.35	16.85
-	#.aoQ	F-0.07	7.101	8.260	9.591	10.85	12:44	14.58	16.27	17.81
20 21	5·398 5·896	5·921 6·447	7·434 8·034	8.897	10.28	11.20	13.24	15.44	17.18	18.77
22	5 090 6·404	6.983	8.643	9.542	10.98	12.34	14.04	16.31	18.10	19.73
	6.924	7.529	9.260	10.50	11.60	13.09	14.85	17.19	19.02	20.69
23 24	7.453	8·085	9.886	10.86	12.40	13.85	15.66	18.06	19.94	21.65
24	7 453	0 003	9 000	10 00	12 40	-3 °5	-5 00		- 7 7 7	3
25	7.991	8.649	10.2	11.52	13.12	14.61	16.47	18.94	20.87	22.62
26	8.538	9.222	11.16	12.30	13.84	15.38	17.29	19.82	21.79	23.58
27	9.093	9.803	11.81	12.88	14.57	16.15	18.11	20.70	22.72	24.24
28	9.656	10.39	12.46	13.26	15.31	16.93	18.94	21.59	23.65	25.21
29	10.53	10.99	13.12	14.26	16.05	17.71	19.77	22:48	24.58	26.48
-										
30	10.80	11.20	13.79	14.95	16.79	18.49	20.60	23.36	25.21	27.44
32	11.08	12.81	15.13	16.36	18.39	20.07	22.27	25.12	27:37	29:38
34	13.18	14.06	16.20	17.79	19.81	21.66	23.95	26.94	29.24	31.31
36	14.40	15.32	17.89	19.23	21.34	23.27	25.64	28.73	31.15	33.25
38	15.64	16.61	19.29	20.69	22.88	24.88	27:34	30.24	32.99	32.19
40	-6.0-	T	20.77	22.16	24.43	26.21	29.05	32.34	34 <sup>.8</sup> 7	37.13
40	16.91	17.92	20.71				37·69	32 34 41.45	44·3 <b>I</b>	46.86
50 60	23.46	24.67	27.99	29.7I	32·36 40·48	34·76 43·19	46·46	50·64	53·81	56.62
60 70	30.34	31.74	35.23	37.48	48·76	51.74	55.33	59.90	63.32	66.40
70 80	37.47	39.04	43.28	45.44	57·15	60.39	55 55 64·28	69.31	72.92	76.19
00	44.79	46.22	51.17	53.54	3/ -3	oo 39	0- <b>--</b>	~y ~*	/ y	,~ ~7
90	52.28	54.16	59.20	61.75	65.65	69.13	73.29	78· <b>5</b> 6	82.51	85.99
100	20.00	61.92	67.33	70.06	74.22	77.93	82.36	87.95	92.13	95.81
		-		•	-					

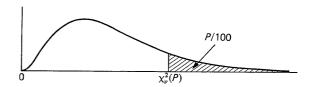
# TABLE 8. PERCENTAGE POINTS OF THE $\chi^2$ -DISTRIBUTION

This table gives percentage points  $\chi^2_{\nu}(P)$  defined by the equation

$$\frac{P}{100} = \frac{1}{2^{\nu/2} \Gamma(\frac{\nu}{2})} \int_{\chi_p^2(P)}^{\infty} x^{\frac{1}{2}\nu - 1} e^{-\frac{1}{2}x} dx.$$

If X is a variable distributed as  $\chi^2$  with  $\nu$  degrees of freedom, P/100 is the probability that  $X \geqslant \chi^2_{\nu}(P)$ .

For  $\nu > 100$ ,  $\sqrt{2X}$  is approximately normally distributed with mean  $\sqrt{2\nu - 1}$  and unit variance.



(The above shape applies for  $\nu \geqslant 3$  only. When  $\nu < 3$  the mode is at the origin.)

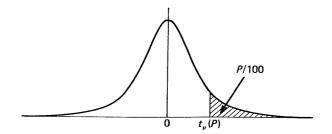
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	82 15·20 27 17·73 47 20·00 52 22·11 46 24·10 32 26·02 12 27·87
2 1.386 1.833 2.408 3.219 4.605 5.991 7.378 9.210 10.60 13 3 2.366 2.946 3.665 4.642 6.251 7.815 9.348 11.34 12.84 16 4 3.357 4.045 4.878 5.989 7.779 9.488 11.14 13.28 14.86 18 5 4.351 5.132 6.064 7.289 9.236 11.07 12.83 15.09 16.75 20	27 17.73 47 20.00 52 22.11 46 24.10 32 26.02 12 27.87
4 3.357 4.045 4.878 5.989 7.779 9.488 11.14 13.28 14.86 18 5 4.351 5.132 6.064 7.289 9.236 11.07 12.83 15.09 16.75 20	47 20·00 52 22·11 46 24·10 32 26·02 12 27·87
5 4:351 5:132 6:064 7:289 9:236 11:07 12:83 15:09 16:75 20	52 22·11 46 24·10 32 26·02 12 27·87
	46 24·10 32 26·02 12 27·87
B 5:24X 6:211 7:221 X:55X TO:64 TO:60 TILLE -4:00:55 A	26·02 27·87
221, 1-1, 12, 12, 12, 14, 14, 14, 10, 10, 10, 10, 10, 10, 10, 10, 10, 10	12 27.87
7 6.346 7.283 8.383 9.803 12.02 14.07 16.01 18.48 20.28 24	
<b>8</b> 7.344 8.351 9.524 11.03 13.36 15.51 17.53 20.09 21.95 26.	38 20.67
9 8.343 9.414 10.66 12.24 14.68 16.92 19.02 21.67 23.59 27	,.,
10 9:342 10:47 11:78 13:44 15:99 18:31 20:48 23:21 25:19 29:	59 31.42
<b>II</b> 10·34 11·53 12·90 14·63 17·28 19·68 21·92 24·72 26·76 31·	26 33.14
12 11·34 12·58 14·01 15·81 18·55 21·03 23·34 26·22 28·30 32·	34.82
<b>13</b> 12·34 13·64 15·12 16·98 19·81 22·36 24·74 27·69 29·82 34·	36.48
14 13·34 14·69 16·22 18·15 21·06 23·68 26·12 29·14 31·32 36·	12 38.11
<b>15</b> 14·34 15·73 17·32 19·31 22·31 25·00 27·49 30·58 32·80 37·	70 39.72
<b>16</b> 15·34 16·78 18·42 20·47 23·54 26·30 28·85 32·00 34·27 39·	
17 16·34 17·82 19·51 21·61 24·77 27·59 30·19 33·41 35·72 40·	
18 17·34 18·87 20·60 22·76 25·99 28·87 31·53 34·81 37·16 42·	
<b>19</b> 18·34 19·91 21·69 23·90 27·20 30·14 32·85 36·19 38·58 43·	32 45.97
20 19:34 20:95 22:77 25:04 28:41 31:41 34:17 37:57 40:00 45:	
<b>21</b> 20·34 21·99 23·86 26·17 29·62 32·67 35·48 38·93 41·40 46·	
<b>22</b> 21·34 23·03 24·94 27·30 30·81 33·92 36·78 40·29 42·80 48·	
<b>23</b> 22·34 24·07 26·02 28·43 32·01 35·17 38·08 41·64 44·18 49·	1
<b>24</b> 23:34 25:11 27:10 29:55 33:20 36:42 39:36 42:98 45:56 51:	8 53.48
<b>25</b> 24·34 26·14 28·17 30·68 34·38 37·65 40·65 44·31 46·93 52·0	2 54.95
<b>26</b> 25:34 27:18 29:25 31:79 35:56 38:89 41:92 45:64 48:29 54:0	
<b>27</b> 26·34 28·21 30·32 32·91 36·74 40·11 43·19 46·96 49·64 55·2	
<b>28</b> 27.34 29.25 31.39 34.03 37.92 41.34 44.46 48.28 50.99 56.8	
<b>29</b> 28·34 30·28 32·46 35·14 39·09 42·56 45·72 49·59 52·34 58·3	0 60.73
<b>30</b> 29·34 31·32 33·53 36·25 40·26 43·77 46·98 50·89 53·67 59·	0 62.16
<b>32</b> 31·34 33·38 35·66 38·47 42·58 46·19 49·48 53·49 56·33 62·4	9 65.00
<b>34</b> 33°34 35°44 37°80 40°68 44°90 48°60 51°97 56°06 58°96 65°2	5 67.80
<b>36</b> 35·34 37·50 39·92 42·88 47·21 51·00 54·44 58·62 61·58 67·9	9 70.59
<b>38</b> 37·34 39·56 42·05 45·08 49·51 53·38 56·90 61·16 64·18 70·7	0 73:35
<b>40</b> 39·34 41·62 44·16 47·27 51·81 55·76 59·34 63·69 66·77 73·4	
<b>50</b> 49·33 51·89 54·72 58·16 63·17 67·50 71·42 76·15 79·49 86·6	, ,
<b>60</b> 59.33 62.13 65.23 68.97 74.40 79.08 83.30 88.38 91.95 99.6	
70 69:33 72:36 75:69 79:71 85:53 90:53 95:02 100:4 104:2 112:	
80 79·33 82·57 86·12 90·41 96·58 101·9 106·6 112·3 116·3 124·8	128.3
90 89.33 92.76 96.52 101.1 107.6 113.1 118.1 124.1 128.3 137.2	140.8
100 99.33 102.9 106.9 111.7 118.5 124.3 129.6 135.8 140.2 149.4	-

## TABLE 10. PERCENTAGE POINTS OF THE t-DISTRIBUTION

This table gives percentage points  $t_{\nu}(P)$  defined by the equation

$$\frac{P}{\mathrm{100}} = \frac{\mathrm{I}}{\sqrt{\nu \pi}} \frac{\Gamma(\frac{1}{2}\nu + \frac{1}{2})}{\Gamma(\frac{1}{2}\nu)} \int_{t_{\nu}(P)}^{\infty} \frac{dt}{(\mathrm{I} + t^2/\nu)^{\frac{1}{2}(\nu + 1)}}.$$

Let  $X_1$  and  $X_2$  be independent random variables having a normal distribution with zero mean and unit variance and a  $\chi^2$ -distribution with  $\nu$  degrees of freedom respectively; then  $t=X_1/\sqrt{X_2/\nu}$  has Student's t-distribution with  $\nu$  degrees of freedom, and the probability that  $t \geq t_{\nu}(P)$  is P/100. The lower percentage points are given by symmetry as  $-t_{\nu}(P)$ , and the probability that  $|t| \geq t_{\nu}(P)$  is 2P/100.



The limiting distribution of t as  $\nu$  tends to infinity is the normal distribution with zero mean and unit variance. When  $\nu$  is large interpolation in  $\nu$  should be harmonic.

P	40	30	25	20	15	10	5	2.2	r	0.2	0.1	0.02
$\nu = \mathbf{r}$	0.3249	0.7265	1.0000	1.3764	1.963	3.078	6.314	12.71	31.82	63.66	318.3	636.6
2	0.2887	0.6172	0.8165	1.0607	1.386	1.886	2.920	4.303	6.965	9.925	22.33	31.60
3	0.2767	0.5844	0.7649	0.9785	1.220	1.638	2.353	3.185	4.241	5.841	10.31	12.92
4	0.2707	0.2686	0.7407	0.9410	1.190	1.233	2.135	2.776	3.747	4.604	7.173	8.610
5	0.2672	0.5594	0.7267	0.9192	1.126	1.476	2.012	2.21	3.362	4.032	5.893	6.869
6	0.2648	0.5534	0.7176	0.9057	1.134	1.440	1.943	2.447	3.143	3.707	5.208	5.959
7	0.2632	0.5491	0.411	o·896 <b>o</b>	1.119	1.412	1.895	2.362	2.998	3.499	4·785	5.408
8	0.5619	0.5459	0.7064	0.8889	1.108	1.392	1.860	2.306	2.896	3.352	4.201	5.041
9	0.3610	0.5435	0.7027	0.8834	1.100	1.383	1.833	2.262	2.821	3.520	4.297	4.781
	_			0.			0 .	0		( -		0
10	0.5605	0.2412	0.6998	0.8791	1.003	1.372	1.812	2.228	2.764	3.169	4.144	4.587
11	0.2596	0.2399	0.6974	0.8755	1.088	1.363	1.796	2.301	2.718	3.106	4.025	4.437
12	0.2590	0.5386	0.6955	0.8726	1.083	1.356	1.782	2.179	2.681	3.055	3.930	4.318
13	0.2586	0.2372	0.6938	0.8702	1.079	1.350	1.771	2.160	2.650	3.015	3.852	4.551
14	0.2582	0.2366	0.6924	0.8681	1.076	1.342	1.761	2.145	2.624	2.977	3.787	4.140
15	0.2579	0.5357	0.6912	0.8662	1.074	1.341	1.753	2.131	2.602	2.947	3.733	4.073
16	0.2576	0.2320	0.6901	0.8647	1.071	1.337	1.746	2.150	2.583	2.921	3.686	4.012
17	0.2573	0.5344	0.6892	0.8633	1.069	1.333	1.740	2.110	2.567	2.898	3.646	3.965
18	0.2571	0.2338	0.6884	0.8620	1.062	1.330	1.734	2.101	2.552	2.878	3.610	3.922
19	0.2569	0.2333	0.6876	0.8610	1.066	1.328	1.729	2.003	2.239	2.861	3.579	3.883
_,		- 5555				- 0	• •	, ,	•••		0 0	
20	0.2567	0.2329	0.6870	0.8600	1.064	1.325	1.725	2.086	2.528	2.845	3.552	3·85 <b>0</b>
21	0.2566	0.5325	0.6864	0.8591	1.063	1.323	1.721	2.080	2.218	2·831	3.527	3.819
22	0.2564	0.2321	0.6858	0.8583	1.001	1.321	1.212	2.074	2.208	2.819	3.202	3.792
23	0.2563	0.5317	0.6853	0.8575	1.060	1.319	1.714	2.069	2.200	2.807	3.485	3.768
24	0.2562	0.2314	0.6848	0.8569	1.029	1.318	1.411	2.064	2.492	2.797	3.467	3.745
			0.6844	0.8562	1.058	1.316	1.408	2.060	2.485	2.787	21450	3.725
25 26	0.2561	0.2312	0.6840	0.8557	1.028	1.312	1.706	2.056	2.479		3.450	3 725 3.707
26 27	0.2560	0.2300	0.6837	0.8551	_		1.703	2.052	2·479	2·779 2·771	3.435	3.690
27 28	0·2559 0·2558	0.2306	0.6834	0.8551	1·056	1.314 1.314	1.401	2.048	2·467	2.763	3·421 3·408	3.674
		0.2304	0.6830	0.8542	1.022	1.311	1.699	2.045	2.462	2.756	3.396	3.659
29	0.2557	0.5302	0 0030	0 0342	1 055	1 3.1	1 099	2 043	2 402	4 /30	3 390	3 039
30	0.2556	0.2300	0.6828	0.8538	1.052	1.310	1.697	2.042	2.457	2.750	3.382	3.646
32	0.2555	0.297	0.6822	0.8530	1.054	1.309	1.694	2.037	2.449	2.738	3.362	3.622
34	0.2553	0.5294	0.6818	0.8523	1.023	1.302	1.69 <b>1</b>	2.032	2.441	2.728	3.348	3.601
36	0.2552	0.2291	0.6814	0.8517	1.022	1.306	1.688	2.028	2.434	2.419	3.333	3.282
38	0.5251	0.288	0.6810	0.8512	1.021	1.304	1.686	2.024	2.429	2.712	3.319	3.566
_		06	(0	0.0			60 ·	0.000	0.400	a.== :	21555	a. = = =
40	0.2550	0.2286	0.6807	0.8507	1.020	1.303	1.684	2.021	2.423	2.704	3.307	3.221
50	0.2547	0.5278	0.6794	0.8489	1.047	1.500	1.676	2.009	2.403	2.678	3.561	3.496
60 700	0.2545	0.5272	0.6786	0.8477	1.042	1.596	1.671	2.000	2.390	2.660	3.535	3.460
120	0.2539	0.5258	0.6765	0.8446	1.041	1.589	1.658	1.980	2.358	2.617	3.190	3.373
∞	0.2533	0.2244	0.6745	0.8416	1.036	1.585	1.645	1.960	2.326	2.576	3.090	3.591

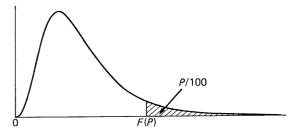
# TABLE 12(a). 10 PER CENT POINTS OF THE F-DISTRIBUTION

The function tabulated is  $F(P) = F(P|\nu_1, \nu_2)$  defined by the equation

$$\frac{P}{\text{100}} = \frac{\Gamma(\frac{1}{2}\nu_1 + \frac{1}{2}\nu_2)}{\Gamma(\frac{1}{2}\nu_1)\;\Gamma(\frac{1}{2}\nu_2)} \, \nu_1^{\frac{1}{2}\nu_1} \; \nu_2^{\frac{1}{2}\nu_2} \int_{F(P)}^{\infty} \frac{F^{\frac{1}{2}\nu_1 - 1}}{(\nu_2 + \nu_1 F)^{\frac{1}{2}(\nu_1 + \nu_2)}} \, dF,$$

for P=10, 5, 2.5, 1, 0.5 and 0.1. The lower percentage points, that is the values  $F'(P)=F'(P|\nu_1,\nu_2)$  such that the probability that  $F\leqslant F'(P)$  is equal to P/100, may be found by the formula

$$F'(P|\nu_1, \nu_2) = 1/F(P|\nu_2, \nu_1).$$

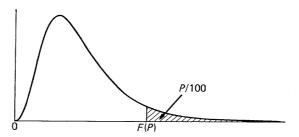


(This shape applies only when  $\nu_1 \geqslant 3$ . When  $\nu_1 < 3$  the mode is at the origin.)

$\nu_1 =$	r	2	3	4	5	6	7	8	10	12	24	∞
_			_			58.20	58.91	59.44	60.19	60.71	62.00	63.33
$\nu_2 = \mathbf{I}$	39.86	49.50	53.59	55.83	57.24	• .	9:349	9·367	9.392	9.408	9.450	9.491
2	8.526	9.000	9.162	9.243	9.293	9.326	: -		5.530	5.216	5.176	5.134
3	5.238	5.462	5.391	5.343	5.309	5.285	5.266	5.252	3.920	3.896	3.831	3.761
4	4.242	4.322	4.191	4.107	4.021	4.010	3.979	3.955	3 920	3 090	3 032	3.792
5	4.060	3.780	3.619	3.520	3.453	3.405	3.368	3.339	3.297	3.268	3.191	3.102
6	3.776	3.463	3.289	3.181	3.108	3.055	3.014	2.983	2.937	2.902	2.818	2.722
7	3.289	3.257	3.074	2.961	2.883	2.827	2.785	2.752	2.703	2.668	2.575	2.471
8	3.458	3.113	2.924	2.806	2.726	2.668	2.624	2.589	2.538	2.502	2.404	2.293
9	3.360	3.006	2.813	2.693	2.611	2.221	2.202	2.469	2.416	2.379	2.277	2.159
10	3.285	2:924	2.728	2.605	2.22	2.461	2.414	2.377	2.323	2.284	2.178	2.055
11		2.860	2.660	2.536	2.451	2.389	2.342	2.304	2.248	2.209	2.100	1.972
12	3.225	2.807	2.606	2·480	2.394	2.331	2.283	2.245	2.188	2.147	2.036	1.904
	3.177		_			2.283	2.234	2.192	2.138	2.097	1.983	1·846
13	3.136	2.763	2.260	2.434	2.347	•	2.193	2.124	2.095	2.054	1.938	1.797
14	3.103	2.726	2.22	2.395	2.307	2.243	2 193	2 154	2 093	2054	2 930	- 151
15	3.073	2.695	2.490	2.361	2.273	2.208	2.158	2.113	2.059	2.017	1.899	1.755
16	3.048	2.668	2.462	2.333	2.244	2.178	2.158	2.088	2.028	1.985	1.866	1.718
17	3.026	2.645	2.437	2.308	2.218	2.152	2.102	2.061	2.001	1.958	1.836	1.686
18	3.007	2.624	2.416	2.286	2.196	2.130	2.079	2.038	1.977	1.933	1.810	1.657
19	3.007	2.606	2.397	2.266	2.176	2.100	2.058	2.017	1.956	1.912	1.787	1.631
19	2 990	2 000	4 397	<b>4</b> 400	2 2 7 0		4 -3-		75			
20	2.975	2.589	2.380	2.249	2.158	2.091	2.040	1.999	1.937	1.892	1.767	1.607
21	2.961	2.575	2.365	2.233	2.142	2.075	2.023	1.982	1.920	1.875	1.748	1.586
22	2.949	2.561	2.351	2.219	2.128	2.060	2.008	1.967	1.904	1.859	1.731	1.267
23	2.937	2.549	2.339	2.207	2.112	2.047	1.995	1.953	1.890	1.845	1.716	1.249
24	2.927	2.538	2.327	2.195	2.103	2.032	1.983	1.941	1.877	1.832	1.702	1.233
	2.079	2.528	2:317	2·184	2.092	2.024	1.971	1.020	1.866	1.820	1.689	1.518
25 26	2.918	-		-	2.082	2.014	1.961	1.010	1.855	1.809	1.677	1.204
26	2.909	2.219	2:307	2·174 2·165	2.073	2.005	1 952	1.000	1.845	1.799	1.666	1.491
27	2.901	2.211	2.299	_	2.064	1.009	1.943	1.000	1.836	1.790	1.656	1.478
28	2 894	2.503	2.291	2.157	•	1.988	1.935	1.892	1.827	1.781	1.647	1.467
29	2.887	2.495	2.283	2.149	2.057	1 900	1 933	1 092	1027	- ,0-	- 5-17	- 1-7
30	2.881	2.489	2.276	2.142	2.049	1.980	1.927	1.884	1.819	1.773	1.638	1.456
32	2.869	2.477	2.263	2.129	2.036	1.967	1.913	1.870	1.802	1.758	1.622	1.437
34	2.859	2.466	2.252	2.118	2.024	1.955	1.901	1.858	1.793	1.745	1.608	1.419
36	2.850	2.456	2.243	2.108	2.014	1.945	1.891	1.847	1.781	1.734	1.292	1.404
38	2.842	2.448	2.234	2.099	2.005	1.935	1.881	1.838	1.772	1.724	1.584	1.390
40	0.00#	01446	2,226	21001	1.992	1.927	1.873	1.829	1.763	1.715	1.574	1.377
40 60	2.835	2:440	2.226	2.091	1.946	1.875	1.819	1.775	1.707	1.657	1.211	1.591
60	2.791	2.393	2.177	2.041		1.824	1.767	1.722	1.652	1.601	1.447	1.103
120	2.748	2.347	2.130	1.992	1.896	-		1.670	1.299	1.546	1.383	1.000
90	2.706	2.303	2.084	1.942	1.847	1.774	1.212	1 0/0	- 399	- 340	- 303	_ 000

# TABLE 12(b). 5 PER CENT POINTS OF THE F-DISTRIBUTION

If  $F=\frac{X_1}{\nu_1}\Big/\frac{X_2}{\nu_2}$ , where  $X_1$  and  $X_2$  are independent random variables distributed as  $\chi^2$  with  $\nu_1$  and  $\nu_2$  degrees of freedom respectively, then the probabilities that  $F\geqslant F(P)$  and that  $F\leqslant F'(P)$  are both equal to P/100. Linear interpolation in  $\nu_1$  and  $\nu_2$  will generally be sufficiently accurate except when either  $\nu_1>12$  or  $\nu_2>40$ , when harmonic interpolation should be used.

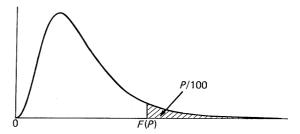


(This shape applies only when  $\nu_1 \geqslant 3$ . When  $\nu_1 < 3$  the mode is at the origin.)

$v_1 =$	r	2	3	4	5	6	7	8	10	12	24	ø
$\nu_2 = \mathbf{r}$	161.4	199.5	215.7	224.6	230.2	234.0	236.8	238.9	241.9	243.9	249.1	254.3
2	18.21	19.00	19.16	19.25	19.30	19.33	19.35	19.37	19.40	19.41	19.45	19.20
3	10.13	9.552	9.277	9.117	9.013	8.941	8.887	8.845	8.786	8.745	8.639	8.526
4	7.709	6.944	6.291	6.388	6.256	6.163	6.094	6.041	5.964	5.912	5.774	5.628
5	6.608	5.786	5.409	5.192	5.050	4.950	4.876	4.818	4.735	4.678	4.527	4.365
6	5.987	5.143	4.757	4.234	4.387	4.284	4.502	4.147	4·060	4.000	3.841	3.669
7	5.201	4.737	4.347	4.130	3.972	3.866	3.787	3.726	3.637	3.575	3.410	3.530
8	5.318	4.459	4.066	3.838	3.687	3.281	3.200	3.438	3.347	3·284	3.112	2.928
9	5.112	4.256	3.863	3.633	3.482	3:374	3.563	3.530	3.137	3.073	2.000	2.707
9	3 7	T ~3°	3 003	J ~JJ	3 404	3 3/4	3 493	3 230	3 -37	3 0/3	2 900	2 /0/
10	4.965	4.103	3.708	3.478	3.326	3.512	3.132	3.072	2.978	2.913	2.737	2.538
11	4.844	3.982	3.587	3.357	3.504	3.095	3.013	2.948	2.854	2.788	2.609	2.404
12	4.747	3.885	3.490	3.259	3.106	2.996	2.913	2.849	2.753	2.687	2.202	2.296
13	4.667	3·806	3.411	3.179	3.022	2.912	2.832	2.767	2.671	2.604	2.420	2.206
14	4.600	3.739	3.344	3.115	2.958	2.848	2.764	2.699	2.602	2.534	2*349	2.131
15	4.543	3.682	3.287	3.056	2.901	2.790	2.707	2.641	2.544	2.475	2.288	2.066
16	4.494	3.634	3.239	3.007	2.852	2.741	2.657	2.201	2.494	2.425	2.235	2.010
17	4.451	3.592	3.197	2.965	2.810	2.699	2.614	2.548	2.450	2.381	2.100	1.960
18	4.414	3.555	3.160	2.928	2.773	2.661	2.577	2.210	2.412	2.342	2.120	1.017
19	4.381	3.222	3.127	2.895	2.740	2.628	2.544	2.477	2.378	2.308	2.114	1.878
20	4.351	3.493	3.008	2.866	2.711	2.299	2.214	2.447	2.348	2.278	2.082	1.843
21	4.325	3.467	3.072	2.840	2.685	2.573	2.488	2.420	2.321	2.250	2.054	1.812
22	4.301	3.443	3.049	2.817	2.661	2.549	2.464	2.397	2.297	2.226	2.028	1.783
23	4.279	3.422	3.028	2.796	2.640	2.528	2.442	2.375	2.275	2.204	2.005	1.757
24	4.260	3.403	3.009	2.776	2.621	2.508	2.423	2.355	2.255	2.183	1.984	1.733
25	4.242	3.385	2.991	2.759	2.603	2.490	2.405	2.337	2.236	2.165	1.964	1.411
26	4.225	3.369	2.975	2.743	2.587	2.474	2.388	2.321	2.220	2.148	1.946	1.601
27	4.210	3.354	2.960	2.728	2.572	2.459	2.373	2.305	2.204	2.132	1.930	1.672
28	4.196	3.340	2.947	2.714	2.558	2.445	2.359	2.291	2.190	2.118	1.915	1.654
29	4.183	3.358	2.934	2.701	2.545	2.432	2.346	2.278	2.177	2.104	1.901	1.638
30	4.171	3.316	2.922	2.690	2.534	2.421	2.334	2.266	2.165	2.002	1.887	1.622
32	4.149	3.295	2.901	2.668	2.212	2.399	2.313	2.244	2.142	2.070	1.864	1.204
34	4.130	3.276	2.883	2.650	2.494	2.380	2.294	2.225	2.123	2.050	1.843	1.269
36	4.113	3.259	2.866	2.634	2.477	2.364	2.277	2.209	2.106	2.033	1.824	1.247
38	4.098	3.245	2.852	2.619	2.463	2.349	2.262	2.194	2.091	2.017	1.808	1.527
40	4.085	3.232	2.839	2.606	2.449	2.336	2.249	2·180	2.077	2.003	1.793	T:500
60	4.001	3.120	2.758	2.525	2.368	2.254	2.167	2.097	1.993	1.017	1.700	1.389
120	•		2.680		2.300		2.087	2.016		1.834	1.408	
±20 ∞	3.920	3.072	_	2:447	-	2.175	•	_	1.010			1.254
w	3.841	2.996	2.605	2.372	2.514	2.099	2.010	1.938	1.831	1.752	1.217	1.000

# TABLE 12(d). 1 PER CENT POINTS OF THE F-DISTRIBUTION

If  $F=\frac{X_1}{\nu_1}\Big/\frac{X_2}{\nu_2}$ , where  $X_1$  and  $X_2$  are independent random variables distributed as  $\chi^2$  with  $\nu_1$  and  $\nu_2$  degrees of freedom respectively, then the probabilities that  $F\geqslant F(P)$  and that  $F\leqslant F'(P)$  are both equal to P/too. Linear interpolation in  $\nu_1$  or  $\nu_2$  will generally be sufficiently accurate except when either  $\nu_1>12$  or  $\nu_2>40$ , when harmonic interpolation should be used.



(This shape applies only when  $\nu_1 \geqslant 3$ . When  $\nu_1 < 3$  the mode is at the origin.)

$\nu_1 =$	I	2	3	4	5	6	7	8	10	12	24	œ
$v_2 = \mathbf{I}$	4052	4999	5403	5625	5764	5859	5928	5981	6056	6106	6235	6366
2	98·5 <b>0</b>	99.00	99.17	99.25	99.30	99.33	99:36	99.37	99.40	99.42	99.46	99.50
3	34.13	30.82	29:46	28.71	28.24	27.91	27.67	27:49	27.23	27.05	26.60	26.13
4	21.30	18.00	16.69	15.98	15.2	15.51	14.98	14.80	14.22	14.37	13.93	13.46
5	16.26	13.27	12.06	11.39	10.97	10.67	10.46	10.39	10.05	9.888	9.466	9.020
6	13.75	10.92	9.780	9.148	8.746	8.466	8.260	8.102	7.874	7.718	7.313	6.880
7	12.25	9.547	8.451	7.847	7:460	7.191	6.993	6.840	6.620	6.469	6.074	5.650
8	11.26	8.649	7.591	7.006	6.632	6.371	6.178	6.029	5.814	5.667	5.279	4.859
9	10.26	8.022	6.992	6.422	6.057	5.802	5.613	5.467	5.257	5.111	4.729	4.311
10	10.04	7:559	6.552	5.994	5.636	5.386	5.200	5.057	4.849	4.706	4.327	3.909
11	9.646	7.206	6.217	5.668	5.316	5.069	4.886	4.744	4.239	4.392	4.021	3.602
12	9.330	6.927	5.953	5.412	5.064	4.821	4.640	4.499	4·296	4.122	3·780	3.361
13	9.074	6.401	5.739	5.202	4.862	4.620	4.441	4.305	4.100	3.960	3.282	3.162
14	8.862	6.212	5.264	5.032	4.695	4.456	4.278	4.140	3.939	3.800	3.427	3.004
15	8.683	6.359	5.417	4.893	4.556	4.318	4.142	4.004	3.805	3.666	3*294	2.868
16	8.231	6.226	5.292	4.773	4.437	4.303	4.026	3.890	3.691	3.223	3.181	2.753
17	8.400	6.115	5.182	4.669	4.336	4.103	3.927	3.491	3.293	3.455	3.084	2.653
18	8.285	6.013	5.092	4.579	4.248	4.012	3.841	3.705	3.208	3.371	2.999	2.566
19	8.185	5.926	5.010	4.200	4.141	3.939	3.765	3.631	3.434	3.597	2.925	2.489
20	8.096	5.849	4.938	4.431	4.103	3.871	3.699	3.564	3.368	3.531	2.859	2.421
21	8.017	5·780	4.874	4•369	4.042	3.812	3.640	3·506	3.310	3.173	2.801	2.360
22	7.945	5.719	4.817	4.313	3.988	3.758	3.587	3.453	3.258	3.151	2.749	2.302
23	7.881	5.664	4.765	4.264	3.939	3.210	3.239	3·406	3.511	3.074	2.702	2.256
24	7.823	5.614	4.718	4.518	3.895	3.667	3.496	3.363	3.168	3.032	2.659	2.311
25	7.770	5.568	4.675	4.177	3.855	3.627	3.457	3:324	3.129	2.993	2.620	2.169
26	7.721	5.526	4.637	4.140	3.818	3.291	3.421	3.288	3.094	2.958	2.585	2.131
27	7.677	5.488	4.601	4.106	3.785	3.228	3.388	3.256	3.062	2.926	2.22	2.097
28	7.636	5.453	4.268	4.074	3.754	3.528	3.328	3.226	3.032	2.896	2.222	2.064
29	7.598	5.420	4.538	4.042	3.725	3.499	3.330	3.198	3.005	2.868	2.495	2.034
30	7.562	5.390	4.210	4.018	3.699	3.473	3.304	3.173	2.979	2.843	2.469	2.006
32	7:499	5.336	4.459	3.969	3.652	3.427	3.258	3.127	2.934	2.798	2.423	1.956
34	7.444	5.289	4.416	3.927	3.611	3.386	3.518	3.087	2.894	2.758	2.383	1.911
36	7.396	5.248	4.377	3.890	3.574	3.321	3.183	3.025	2.859	2.723	2.347	1.872
38	7.353	5.311	4.343	3.858	3.242	3.319	3.125	3.031	2.828	2.692	2.316	1.837
40	7:314	5.179	4.313	3.828	3.214	3.291	3.124	2.993	2.801	2.665	2.288	1.805
6о	7.077	4.977	4.126	3.649	3.339	3.119	2.953	2.823	2.632	2:496	2.112	1.601
120	6.851	4.787	3.949	3.480	3.174	2.956	2.792	2.663	2.472	2.336	1.920	1.381
∞	6.635	4.605	3.782	3.319	3.012	2.802	2.639	2.211	2.321	2.182	1.791	1.000

Durbin-Watson test statistic d: 1% significance points of  $d_{\rm L}$  and  $d_{\rm U}$ .

Γ		1 1	=1	k'=	. 2	7.					
	n	$d_{\rm L}$		1		1	=3	į	=4	1	<b>'</b> =5
F	15	<del></del>	$\frac{d_{\rm U}}{1 \ 1.0}$				$\frac{d_{\mathrm{U}}}{2}$		$d_{\mathtt{U}}$		$d_{\mathtt{U}}$
	16	1	4 1.0	1	1.25 1.25	1	9 1.4		9 1.7	1	39 1.96
	17	1	7 1.10	1	1.25	i	3 1.4	1	3 1.6	4	4 1.90
	18		) 1.12	1	1.25 $1.26$		7 1.4	1	7 1.63		8 1.85
	19		3 1.12		1.26		1 1.43		1 1.60	1	2 1.80
	20		5 1.15	1	1.27		4 1.4	1	5 1.58	1	6 1.77
i	21	1	1.16		1.27		7 1.4		8 1.57		0 1.74
ì	22		1.17				1.4		2 1.55	1	3 1.71
	23	1.02		i .			3 1.40	1	5 1.54	1	6 1.69
- 1	24		1.20	1			5 1.40	1	7 1.53	1	0 1.67
	25		1.21	Į.	,		3 1.41	1	1.53	1	2 1.66
- 1	26		1.22				1.41	1	1.52	Į.	5 1.65
i	27		1.23				1.41	1	1.52		3 1.64
- 1	28		1.24		- 1		1.41	1	1.51	i	ı
	29		1.25				1.41	1	1.51	1	
1	30		1.26				1.42		1.51		5 1.61
J	1		1.27	1.07			1.42		1.51		1.61
- 1	2		1.28	1.10	ı		1.42		1.51		1
3	- 1	1.17		1.10			1.43		1.51		1.60
3	1	1.18		1.11			1.43		1.51		1.59
$\frac{1}{3}$		1.19	- 1	1.13		1.07			1.51		1.59
3			1.32	1.14 1			1.44		1.51		1.59
3'		1.22					1.44		1.51		1.59
38			1.33	1.16 1 1.18 1			1.45	1.06			1.59
39		1.24		1.19 1		1.12		1.07			1.58
40	- 1	1.25		1.19 1		1.14		1.09			1.58
45	- 1	1.29		1.24 1		1.15		1.10			1.58
50	- 1		1.40	1.28 1		1.20		1.16		1.11	
55			1.43	1.32 1.	1	1.24		1.20		1.16	
60		1.38 1		1.35 1.	- 1	1.28	- 1	1.25			1.59
65	- 1	1.41 1		1.38 1.	1	1.32 1.35	ì	1.28		1.25	1
70	- 1		.49				1.53		1.57		1.61
75			.50		- 1		1.55		1.58		1.61
80			.52	1.42 1.	i		1.56		1.59		1.62
85	1		.53	1.46 1.5	- 1		.58		1.60		1.62
90	1		.54	1.47 1.5			.59		1.60		1.63
95	f		.55	1.49 1.5	j		.60		.61		1.64
100	1		.56	1.50 1.5	ı		.60		.62		1.64
							.00	1.70 1	ادن.	1.44	1.65

n = number of observations

k' = number of explanatory variables

Durbin-Watson test statistic d: 5% significance points of  $d_{\rm L}$  and  $d_{\rm U}$ .

			k'=1			k' = 2		k'=3			k'=4		k'=5	
	n	a	L	$d_{\mathtt{U}}$	$d_{\rm L}$		- 1		$d_{ m U}$	$d_{\scriptscriptstyle  m L}$				
	1:			1.36	0.95				.75	0.6				
	16	5 1.	10 1	.37	1		,		.73	0.7		3		
	17	7   1.1	13 1	.38	1.02		3		71	0.78				
	18		16 1	.39	1.05	1.53			69	0.82		l l		
	19			.40	1.08	1.53	3 0.9		68	0.86				
	20	1		.41	1.10	1.54	4   1.0		68	0.90				
	21			.42	1.13		1.0	3 1.	67	0.93		ı		
	22	1		.43	1.15			5 1.0	66	0.96				
	23	1		.44	1.17			8 1.6	56	0.99				
	24	1		.45	1.19		t	) 1.6	66	1.01	1.78	3		
	25	3		45	1.21	1.55	1		56	1.04	1.77	7   0.95		
	26	1.3		46	1.22	1.55	1		- 1	1.06	1.76	6 0.98		
1	27	1.3		47	1.24	1.56	1		- 1	1.08	1.76	$5 \mid 1.01$	1.86	
	28 29	1.33		48	1.26	1.56	1		1	1.10	1.75		1.85	
1	30	1.34		48	1.27	1.56	1		- 1	1.12	1.74			
	31	1.35		49	1.28	1.57	1			1.14	1.74		The state of the s	
	32	1.37		- 1	1.30	1.57	1.23			1.16	1.74	ı		
	33	1.38		T I	1.31 1.32	1.57	1.24			1.18	1.73		i	
	34	1.39			1.32	1.58 1.58	1.26		- 1	1.19	1.73	•	1	
- 1	35	1.40		1	1.34	1.58	1.27		- 1	1.21	1.73	1.15		
ſ	36	1.41		- 1	1.35	1.59	1.28 1.29	1.63	- 1	1.22	1.73	1.16	1.80	
1	37	1.42			1.36	1.59	1.29	1.65 1.66	1	1.24	1.73	1.18	1.80	
ı	38	1.43	1.5		1.37	1.59	1.32	1.66		1.25 1.26	1.72	1.19	1.80	
1	39	1.43	1.5	,	1.38	1.60	1.33			1.27	1.72 1.72	1.21	1.79	
4	40	1.44	1.5		1.39	1.60	1.34	1.66		1.29	1.72	1.22 1.23	1.79	
4	15	1.48	1.5	- 6	1.43	1.62	1.38	1.67	1	1.34	1.72	1.29	1.79	
5	50	1.50	1.59		1.46	1.63	1.42	1.67	Į	1.38	1.72	1.34	1.77	
	55	1.53	1.60	)   :	1.49	1.64	1.45	1.68	1	1.41	1.72	1.38	1.77	
	0	1.55	1.62	2   1	1.51	1.65	1.48	1.69	1	.44	1.73	1.41	1.77	
	5	1.57	1.63	3   1	1.54	1.66	1.50	1.70	,	.47	1.73	1.44	1.77	
	0	1.58	1.64		.55	1.67	1.52	1.70	1	.49	1.74	1.46	1.77	
7		1.60	1.65	ſ	57	1.68	1.54	1.71		.51	1.74	1.49	1.77	
8	- 1	1.61	1.66	1		1.69	1.56	1.72	1	.53	1.74	1.51	1.77	
8:		1.62	1.67	,		1.70	1.57	1.72	1	.55	1.75	1.52	1.77	
90		1.63	1.68	1		1.70	1.59	1.73	1	.57	1.75	1.54	1.78	
95	,	1.64	1.69	1		1.71	1.60	1.73		.58	1.75	1.56	1.78	
10	U	1.65	1.69	1.	.63	1.72	1.61	1.74	1.	.59	1.76	1.57	1.78	

n = number of observations

k' = number of explanatory variables