

# **Trend-seasonal decomposition and exponential smoothing models**

# Data and Tasks

# Data and Tasks: Plan

- Time series is a data type
- Tasks for one row
- Tasks for multiple rows

# What is a time series?

## Time series

A sequence of observations ordered in time

0, 0, 5, 7, 102, 53, 23

## Time series

A sequence of random variables ordered in time

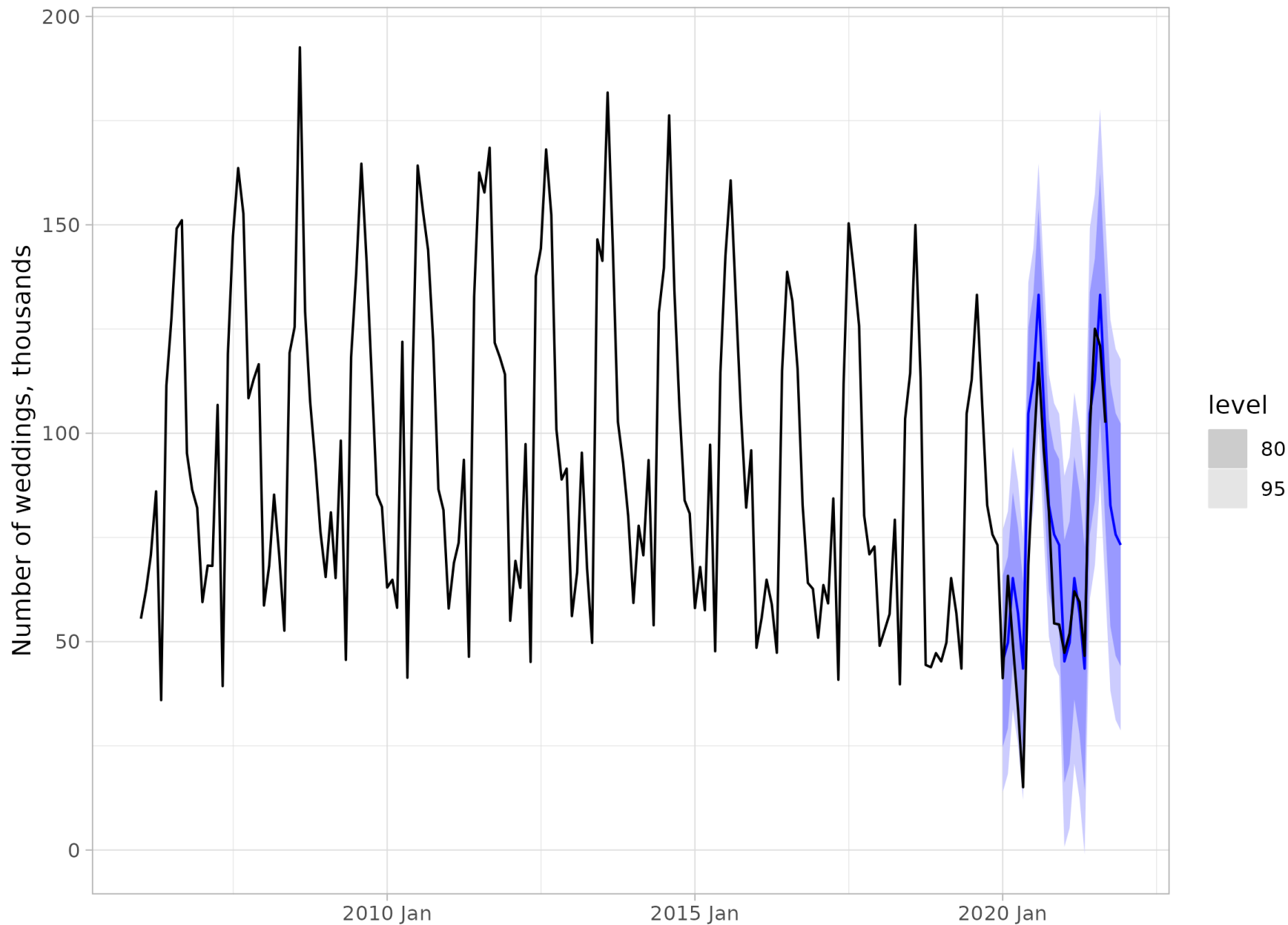
$y_1, y_2, y_3, y_4, \dots, y_T$

# Tasks for one series

- Predict the following values
- Restore missing values in the middle of a series
- Restore individual observations from aggregated ones
- Detect point of discord (or structural break)
- Decompose series to a trend and seasonal parts
- ...

# Forecasting

Seasonal naive forecast for number of weddings in Russia

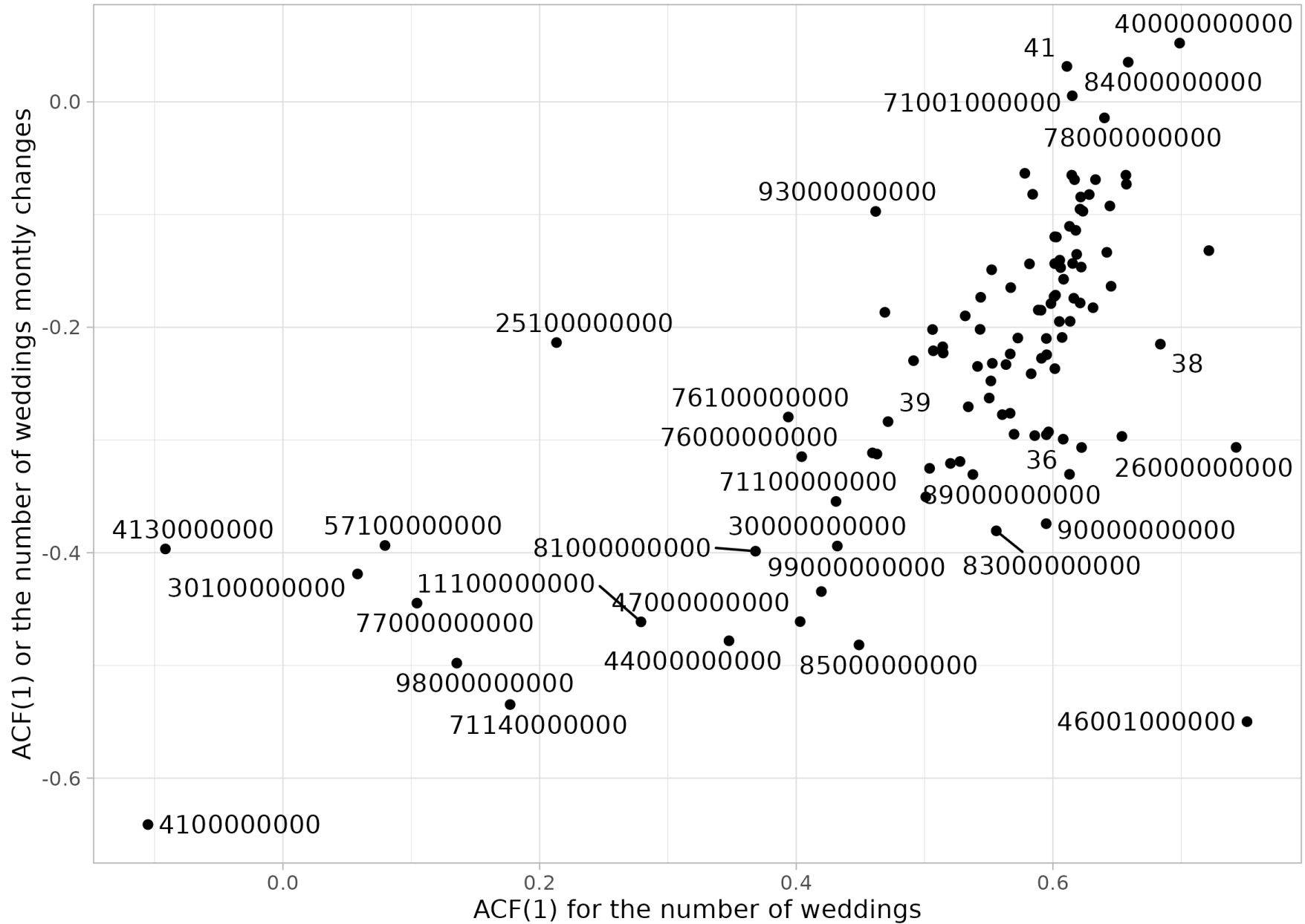


# Tasks for multiple series

- Use additional series when studying the target series
- Understand if series are related
- Measure cause and effect relationships
- Classify the new series into one of the existing classes
- Understand which series are close to each other
- Cluster series into an unknown set of clusters
- ...

# Measuring series proximity

Russian regions proximity by the number of weddings dynamics





# Models and algorithms

## Models

- Explicit assumptions about the values  $y_1, y_2, \dots, y_T$
- Estimation method: maximum likelihood, Bayesian approach
- Point and interval forecasts, hypothesis testing

ETS, ARIMA, ORBIT, PROPHET, ...

# Models and algorithms

## Algorithms

- Fuzzy assumptions about the values  $y_1, y_2, \dots, y_T$
- A special instruction for actions
- Point estimates without confidence intervals

STL, gradient boosting, random forest, ...

# Course Focus

Forecasting one-dimensional series using models

# Series Components

# Series Components: Plan

- Trend, cyclicity and seasonality
- Additive and multiplicative decomposition
- A formal definition?

# Looking for components

Additive series decomposition:

$$y_t = trend_t + seas_t + remainder_t$$

**Trend** — smoothly changing component of the series

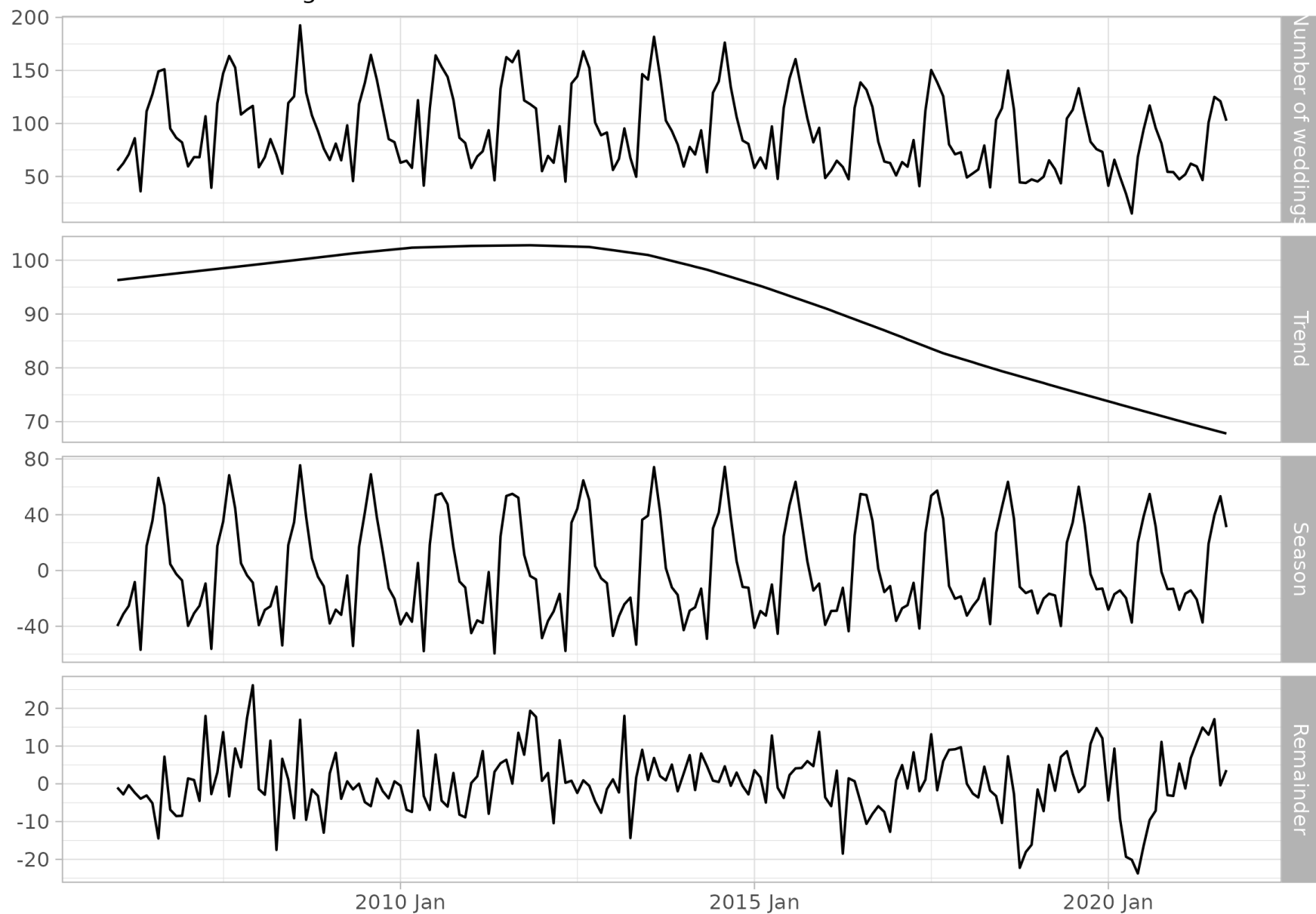
**Seasonal component** — a component with a clear frequency and stable intensity

**Random component** (remainder) — everything else

# Trend, seasonality and residual

STL decomposition the number of weddings in Russia

Number of weddings = Trend + Season + Remainder



# Strict definition?

There will be no single strict definition for the components!

Some models and algorithms formally **define** these components



# Cyclical component

Sometimes the series can be decomposed further

$$y_t = trend_t + cycle_t + seas_t + remainder_t$$

**Cyclical component** — component with floating frequency and unstable intensity

**Trend** (in the narrow sense) — a smoothly changing monotonous component of a series

# Additive and multiplicative decomposition

Additive series decomposition:

$$y_t = trend_t + seas_t + remainder_t$$

Multiplicative series decomposition:

$$y_t = trend_t \cdot seas_t \cdot remainder_t$$

Let's transform one into another:

$$\ln y_t = \ln trend_t + \ln seas_t + \ln remainder_t$$

# Box-Cox Transformation

For  $y_t$ , whose fluctuations increases with  $y_t$ , it's reasonable to try the logarithm or the Box-Cox transformation.

Logarithm:  $y_t \rightarrow \ln y_t$

Box-Cox transformation:  $y_t \rightarrow bc_\lambda(y_t)$

(Generalized) Box-Cox transformation:

$$bc_\lambda(y_t) = \begin{cases} \ln y_t, & \text{if } \lambda = 0, \\ \text{sign}(y_t)(|y_t|^\lambda - 1)/\lambda, & \text{if } \lambda \neq 0 \end{cases}$$

How to select the parameter  $\lambda$ ?

- Some models contain it inside and estimate  $\lambda$  within themselves
- You can choose  $\lambda$  by yourself to stabilize the amplitude of oscillations of the series

# What to choose?

The formal definition of **depends on the model**

**STL algorithm**: one decomposition

$$y_t = trend_t + seas_t + remainder_t$$

**ETS(AAA)**: different decomposition

$$y_t = trend_t + seas_t + remainder_t$$

It is important to understand the **purpose of constructing** the decomposition

# Why decompose?

- Interesting **by itself**
- For **predicting** a series using component prediction
- To get **characteristics of the series**

## Why characteristics?

- To classify the new series into one of the given classes
- To identify unknown clusters in series

# Series Components: Summary

- Trend **smoothly changes** and includes a cyclical component
- The seasonal component has **clear periodicity** and **stable amplitude**
- The exact formalization of the components **depends on the model**

# Naive Models

# Naive Models: Plan

- White noise
- Independent observations
- Random walk



# White noise

## White noise

Time series  $u_t$  is white noise if:

- $\mathbb{E}(u_t) = 0$ ;
- $\text{Var}(u_t) = \sigma^2$ ;
- $\text{Cov}(u_s, u_t) = 0$  for  $s \neq t$

- An integral part of all models; most often, white noise is not modelled explicitly
- Often **independence** and **normality** are assumed  
ARCH, GARCH volatility models are based on the fact that  $u_t$  and  $u_s$  can be dependent!

# Independent observations

## Model

$$y_t = \mu + u_t,$$

where  $u_t$  is white noise,  $u_t \sim \mathcal{N}(0; \sigma^2)$

Estimators:

$$\hat{\mu}_{ML} = \bar{y}, \quad \hat{\sigma}_{ML}^2 = \frac{\sum (y_i - \bar{y})^2}{T}$$

Interval forecast  $h$  steps ahead:

$$[\bar{y} - 1.96\hat{\sigma}; \bar{y} + 1.96\hat{\sigma}]$$

# Random Walk

## Naive model

$$y_t = y_{t-1} + u_t,$$

where  $u_t$  is white noise,  $u_t \sim \mathcal{N}(0; \sigma^2)$ , starting  $y_1$  is given

Let's reformulate:  $y_t - y_{t-1} = \Delta y_t = u_t$

Estimators:

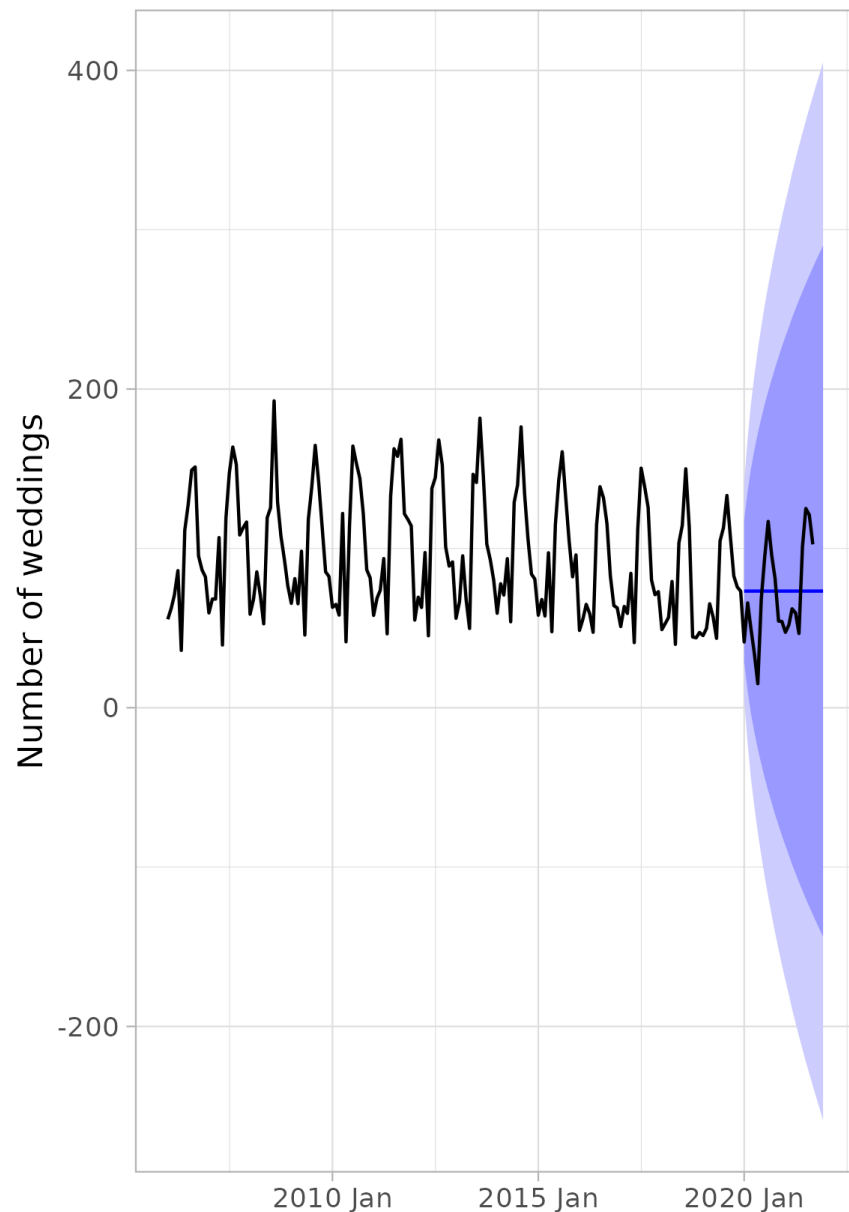
$$\hat{\sigma}_{ML}^2 = \frac{\sum (\Delta y_i - \overline{\Delta y})^2}{T - 1}$$

Interval forecast  $h$  steps ahead:

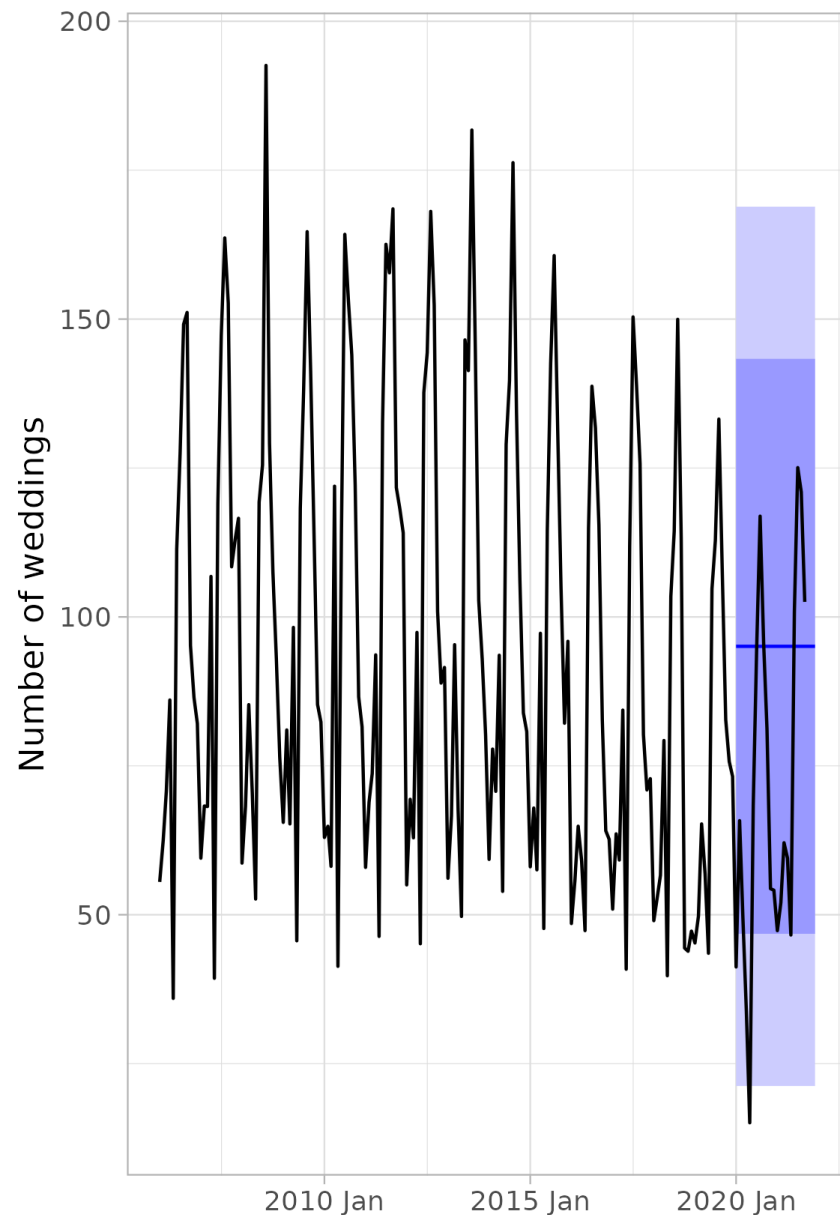
$$[y_T - 1.96\hat{\sigma}\sqrt{h}; y_T + 1.96\hat{\sigma}\sqrt{h}]$$

# First predictions!

Random walk



Independent observations



# Seasonal random walk

## Seasonal naive model

$$y_t = y_{t-12} + u_t,$$

where  $u_t$  is white noise,  $u_t \sim \mathcal{N}(0; \sigma^2)$ ,  $y_1, \dots, y_{11}$  are given

Let's reformulate:  $y_t - y_{t-12} = \Delta_{12}y_t = u_t$

Estimators:

$$\hat{\sigma}_{ML}^2 = \frac{\sum (\Delta_{12}y_i - \overline{\Delta_{12}y})^2}{T - 12}$$

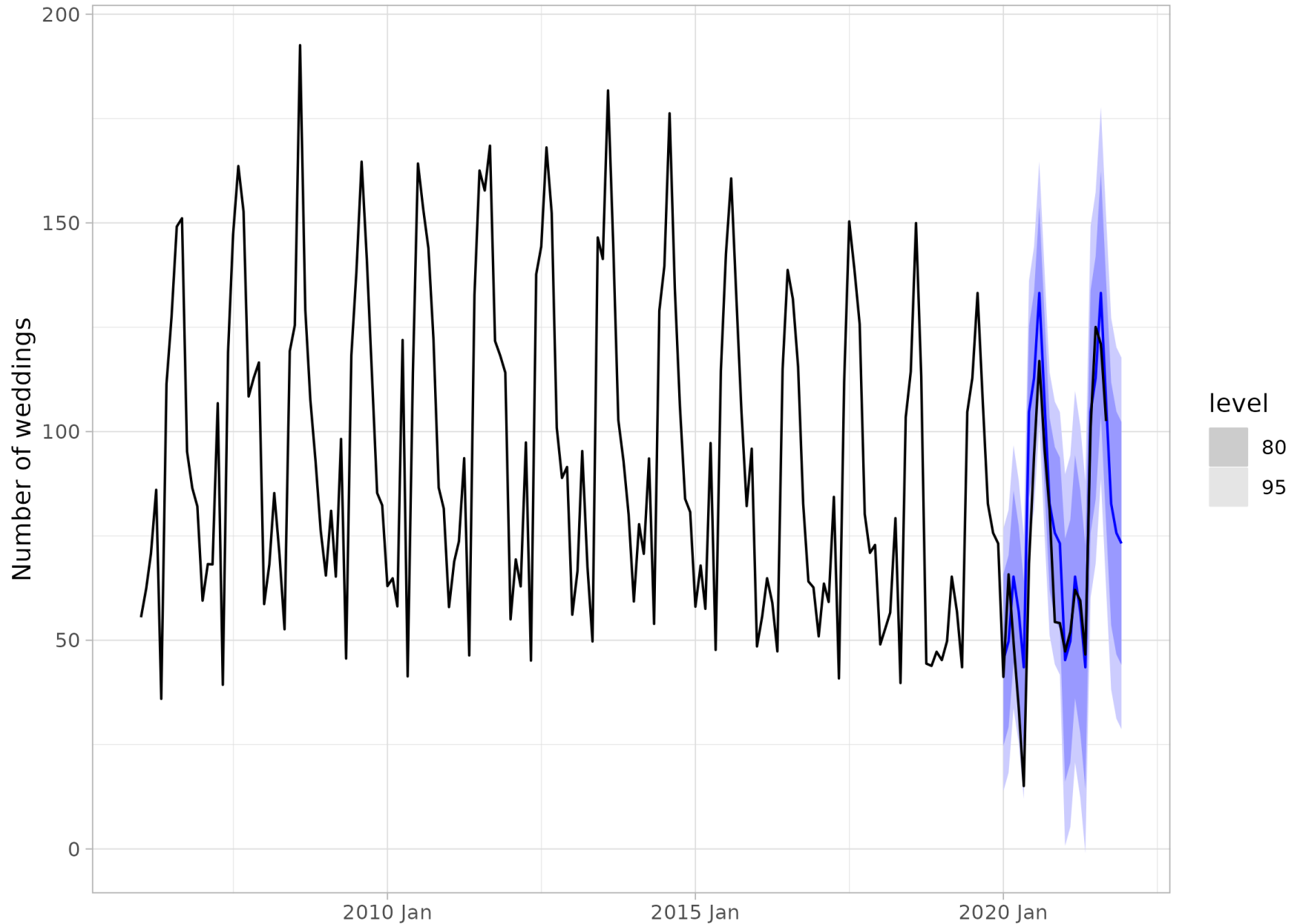
Interval forecast for  $h$  steps ahead:

$$\left[ y_{T-12+h\%12} - 1.96\hat{\sigma} \sqrt{\left\lceil \frac{h}{12} \right\rceil}; y_{T-12+h\%12} + 1.96\hat{\sigma} \sqrt{\left\lceil \frac{h}{12} \right\rceil} \right]$$

where  $\%$  - remainder from the division,  $\lceil x \rceil$  - ceiling function

# Not bad already!

Seasonal naive forecast for number of weddings in Russia



# Why do we need naive models?

- **Ideas** for complex model:  
the **stationary series** models are similar to the independent observations model;  
**non-stationary series** models are similar to a random walk
- **Benchmark for comparison:**  
when evaluating a complex model, it is very important to have a base of comparison
- **Averaging** with other models' forecasts:  
you can **average forecasts** of a complex model and a naive seasonal one!

# Naive Models: Summary

- White noise is what you don't want to simulate
- Independent observations and random walk
- Ideas, parts, and helpers for other models
- Base for comparison



# STL algorithm

# STL algorithm: Plan

- Local regression
- STL outer loop
- STL inner loop
- STL options

# STL

**STL** — Seasonal Trend decomposition with LOESS

STL — seasonality and trend decomposition using LOESS

**LOESS** — LOcal regrESSion

LOESS — local linear regression

# STL as a black box

## Input:

Row  $Y_t$

Algorithm parameters  $n_p, n_i, n_o, n_l, n_s, n_t$

## Output:

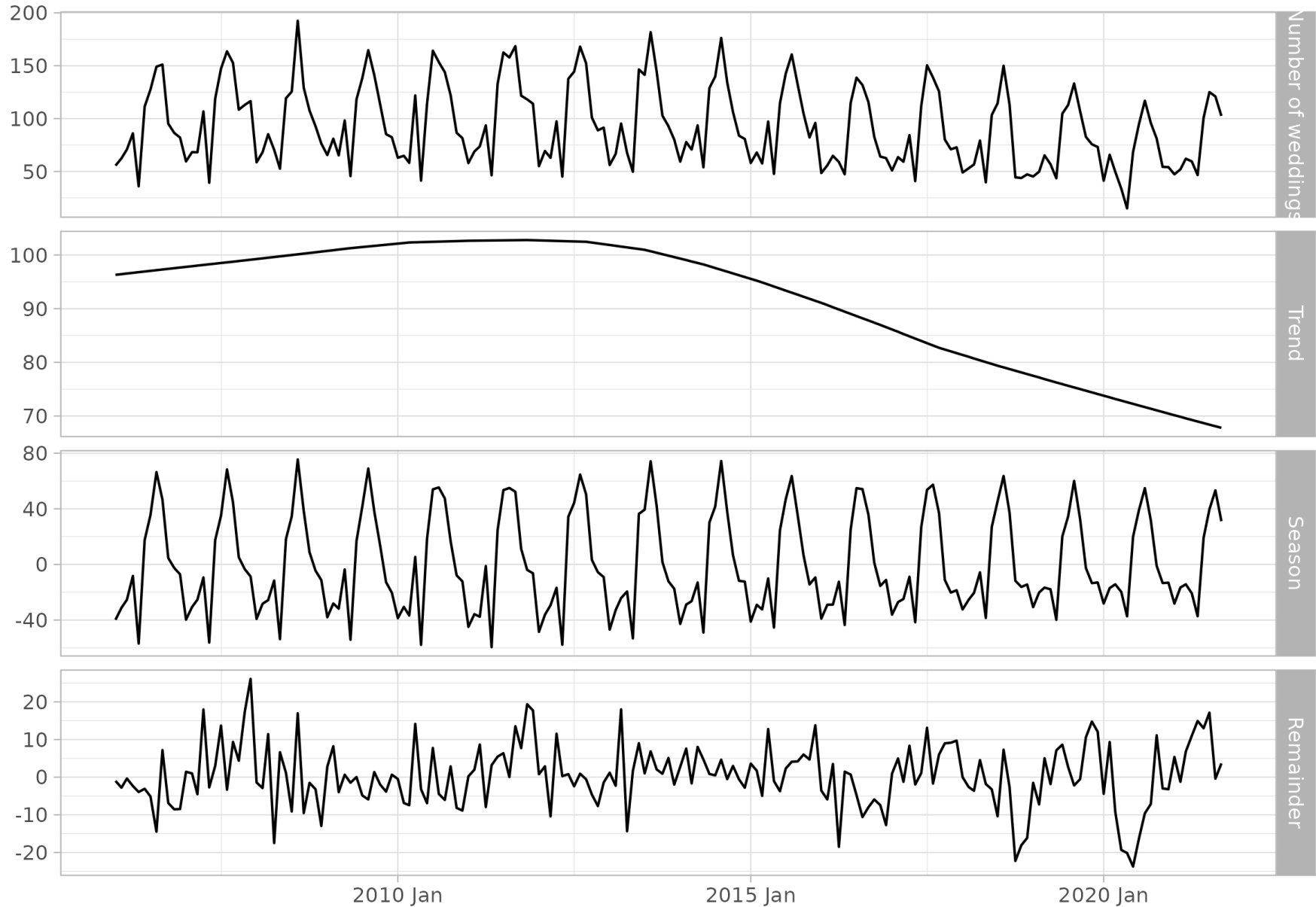
Decomposition  $Y_t = T_t + S_t + R_t$

Black box set up

# STL: result

STL decomposition the number of weddings in Russia

Number of weddings = Trend + Season + Remainder



# LOESS

- We want to build a forecast for the point  $x$
- Find **local estimates**  $\hat{\beta}_1(x), \hat{\beta}_2(x)$

$$\min \sum_i K_h(x_i - x)(y_i - \hat{\beta}_1 - \hat{\beta}_2 x_i)^2$$

- Predicting:

$$\hat{y} = \hat{\beta}_1(x) + \hat{\beta}_2(x)x.$$

## Kernel function

- The function  $K_h(x_i - x)$  decreases with increasing distance  $|x_i - x|$ ;
- The  $h$  parameter controls the width of the smoothing window

For example,  $h$  is the number of points  $x_i$  next to  $x$  that we take into account

# Nuances of local regression

- Select **degrees of the polynomial**

$$\min \sum_i K_h(x_i - x)(y_i - \hat{\beta}_1 - \hat{\beta}_2 x_i - \hat{\beta}_3 x_i^2)^2$$

- Select **kernel** function

$$K_h(d) = \frac{1}{\sqrt{2\pi}h} \exp\left(-d^2/2h^2\right)$$

- Select **window width**  $h$

# STL algorithm

Purpose: decomposition of  $Y_t = T_t + S_t + R_t$

The algorithm contains two loops: **outer** and **inner** loop

1. Initialize  $T_t = 0, R_t = 0$

**Outer** loop:

2. Calculate the weight of each observation,  $\rho_t$ :  
on the first pass,  $\rho_t = 1$  for each observation;  
on subsequent passes,  $\rho_t$  depends negatively on the new value of  $R_t$
3. Update the current decomposition  $Y_t = T_t + S_t + R_t$  taking into account new weights  $\rho_t$



# STL: inner loop

Goal: update the decomposition  $Y_t = T_t + S_t + R_t$ .

1. Remove the previously calculated trend from the series:

$$Y_t^{det} = Y_t - T_t.$$

2. Divide the detrended series into 12 series (one for each season)
3. Smooth each of the series individually with LOESS:

$$C^{jan} = LOESS_{\rho}(Y_{jan}^{det}), C^{feb} = LOESS_{\rho}(Y_{feb}^{det}), \dots$$

4. Extract the low-frequency component (double moving average + LOESS):

$$L_t = LOESS(MA(MA(C_t)))$$

# STL: inner loop

- 1-3. Remove the previously calculated trend from the series, break it down into 12 series and smooth each of them with LOESS.
4. Extract the low-frequency component  $L_t$ .
5. Get **new** seasonal component by removing the low-frequency component:

$$S_t^{new} = C_t - L_t$$

6. Get new trend component by removing new seasonal component from the original series and smoothing with LOESS:

$$T_t^{new} = LOESS_{\rho}(Y_t - S_t^{new})$$

# STL parameters

- $n_p$  — periodicity of seasonality, for example,  $n_p = 12$
- $n_o$  is the number of iterations of the outer loop:  
the larger the number  $n_o$ , the weaker the impact of outliers;  
 $n_o = 1$  is often sufficient
- $n_i$  — number of passes of the inner loop:  
 $n_i = 2$  is often enough to achieve convergence.

# STL smoothing parameters

- $n_l$  — low pass filter smoothing strength
- $n_s$  — seasonal contract smoothing strength
- $n_t$  — smoothing strength when highlighting a trend at the last step

## What to configure?

1. Be sure to specify the periodicity  $n_p$
2. Maybe play around with  $n_s$

# STL algorithm: Summary

- LOESS — local regression
- STL is a well-proven algorithm without an underlying model
- If you wish, you can play around with the smoothing parameters

# **Series Characteristics**

# Series Characteristics: Plan

- Sample autocorrelation
- Sample partial autocorrelation
- STL features

# Tasks for multiple series

- Classify the new series into one of the existing classes
- Understand which series are close to each other
- Cluster series into an unknown set of clusters

## How to solve?

1. Generate **features** for each series
2. Apply the algorithms for cross data to the obtained features

Classify using random forest;

Measure distance using the Mahalanobis metric;

Cluster using hierarchical clustering



# Creating features

Two **sets** of features:

- Sample ACF (**autocorrelation function**, AutoCorrelation Function)
- Sample PACF (**partial** autocorrelation function, Partial ACF)

From **one** series we get:

$ACF_1, ACF_2, ACF_3, \dots$

$PACF_1, PACF_2, PACF_3, \dots$

# ACF

## Sample ACF

Let's evaluate a set of paired regressions:

$$\hat{y}_t = \hat{\beta}_1 + \hat{\beta}_2 y_{t-1}, \quad ACF_1 = \hat{\beta}_2;$$

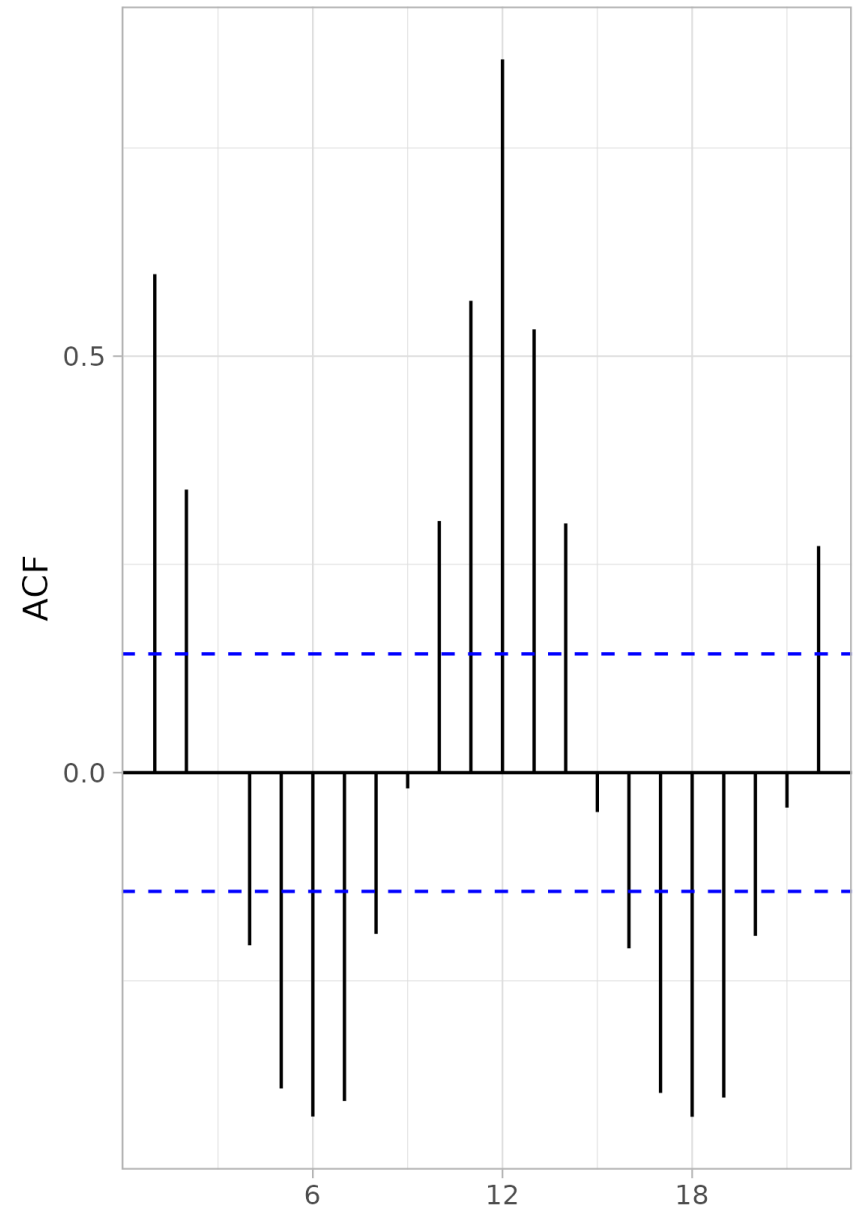
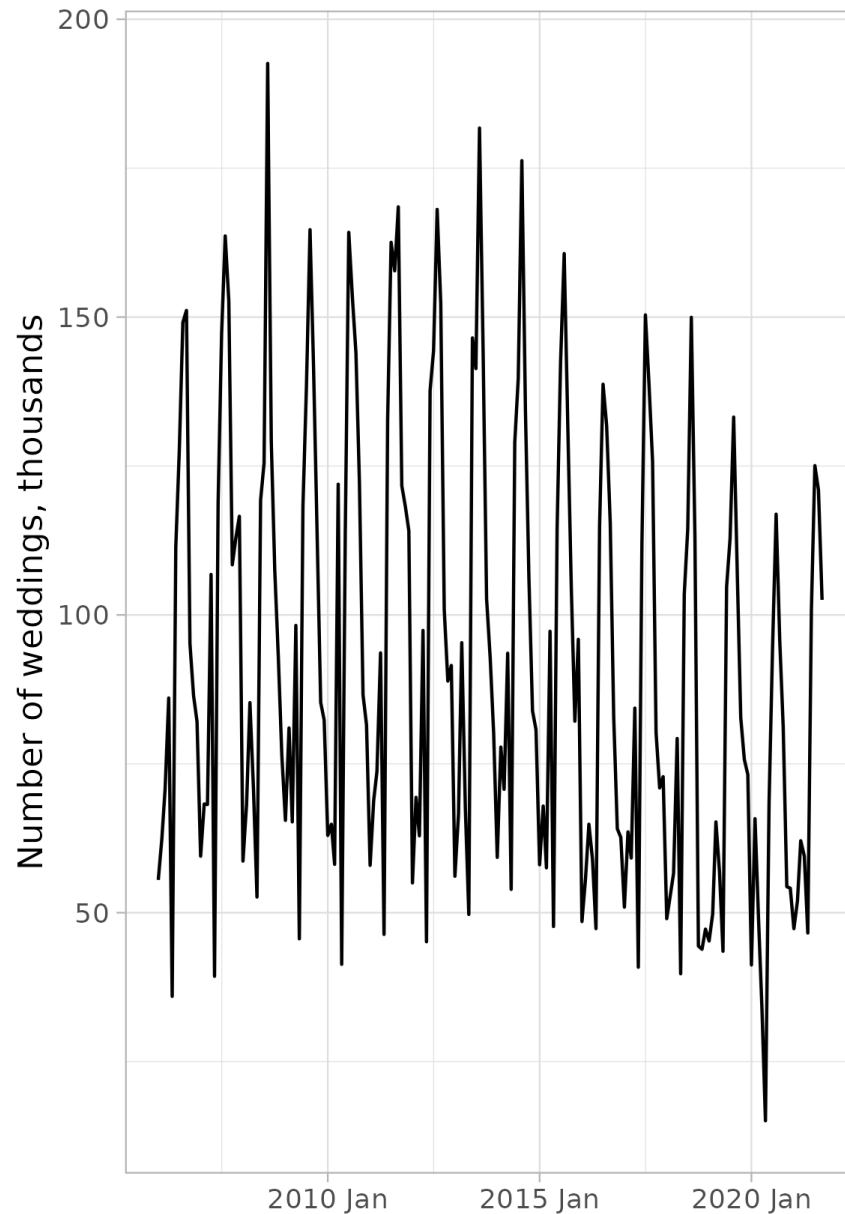
$$\hat{y}_t = \hat{\beta}_1 + \hat{\beta}_2 y_{t-2}, \quad ACF_2 = \hat{\beta}_2;$$

$$\hat{y}_t = \hat{\beta}_1 + \hat{\beta}_2 y_{t-k}, \quad ACF_k = \hat{\beta}_2$$

**Meaning**  $ACF_2$ : How many units is  $y_t$  above average on average if  $y_{t-2}$  is one unit above average.

# Series and its ACF

Number of weddings and ACF



# Why is ACF a correlation?

Classic definition

## Sample ACF

$ACF_k$  — sample correlation between series  $y_t$  and series  $y_{t-k}$

The difference between the definitions is small

# PACF

## Sample PACF

Let's evaluate a set of multiple regressions:

$$\hat{y}_t = \hat{\alpha} + \hat{\beta}_1 y_{t-1}, \quad PACF_1 = \hat{\beta}_1;$$

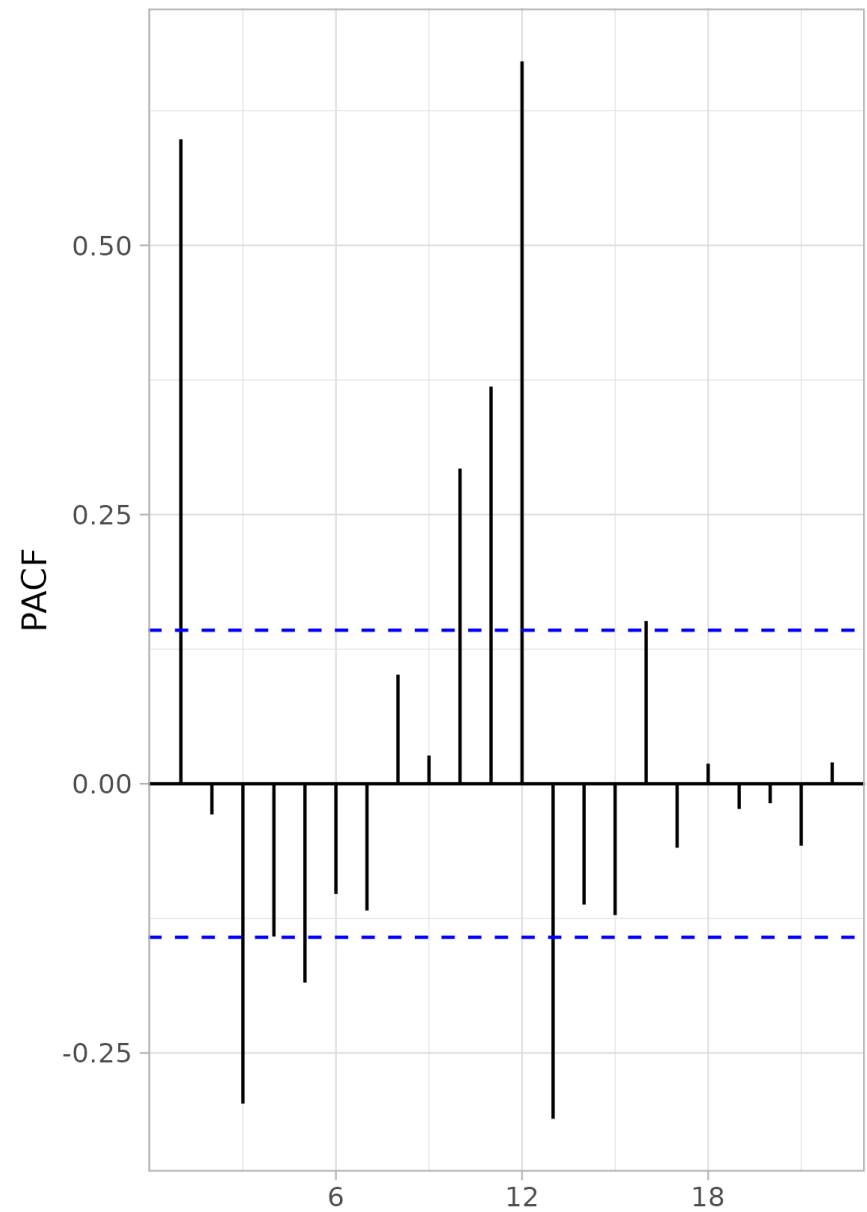
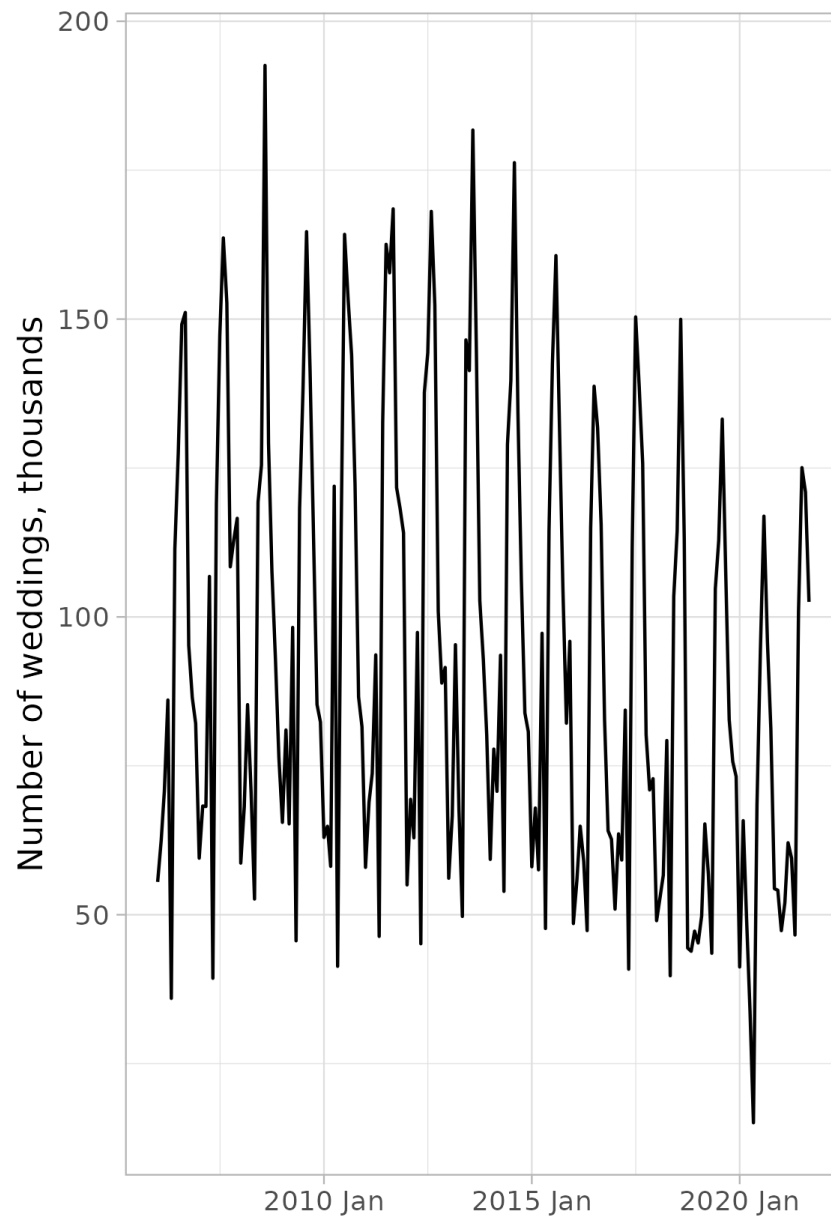
$$\hat{y}_t = \hat{\alpha} + \hat{\beta}_1 y_{t-1} + \hat{\beta}_2 y_{t-2}, \quad PACF_2 = \hat{\beta}_2;$$

$$\hat{y}_t = \hat{\alpha} + \hat{\beta}_1 y_{t-1} + \dots + \hat{\beta}_k y_{t-k}, \quad PACF_k = \hat{\beta}_k$$

**Meaning**  $PACF_2$ : how many units is  $y_t$  above average on average if  $y_{t-2}$  is one unit above average, and  $y_{t-1}$  is at the middle level

# Series and its PACF

Number of weddings and PACF



# Why is PACF a correlation?

## Classic definition

### Custom PACF

$PACF_4$  — sample correlation between  $a_t$  residuals and  $b_t$  residuals:

$a_t$  — regression residuals

$$y_t \mid 1, y_{t-1}, y_{t-2}, y_{t-3};$$

$b_t$  — regression residuals

$$y_{t-4} \mid 1, y_{t-1}, y_{t-2}, y_{t-3}$$

The difference between the definitions of **small**

# STL features

Output:

$$y_t = T_t + S_t + R_t$$

Let's measure:

- Strength of trend  $F_{trend}$
- Strength of seasonality  $F_{seas}$



# Strength of trend and seasonality

We got the decomposition:

$$y_t = trend_t + seas_t + remainder_t.$$

Definition idea:

For an ideal decomposition with uncorrelated components:

$$F_{trend} = \frac{sVar(trend)}{sVar(trend) + sVar(remainder)},$$

$$F_{seas} = \frac{sVar(seas)}{sVar(seas) + sVar(remainder)},$$

# Strength of trend and seasonality

We have the decomposition:

$$y_t = trend_t + seas_t + remainder_t.$$

In practice:

- Trend strength:

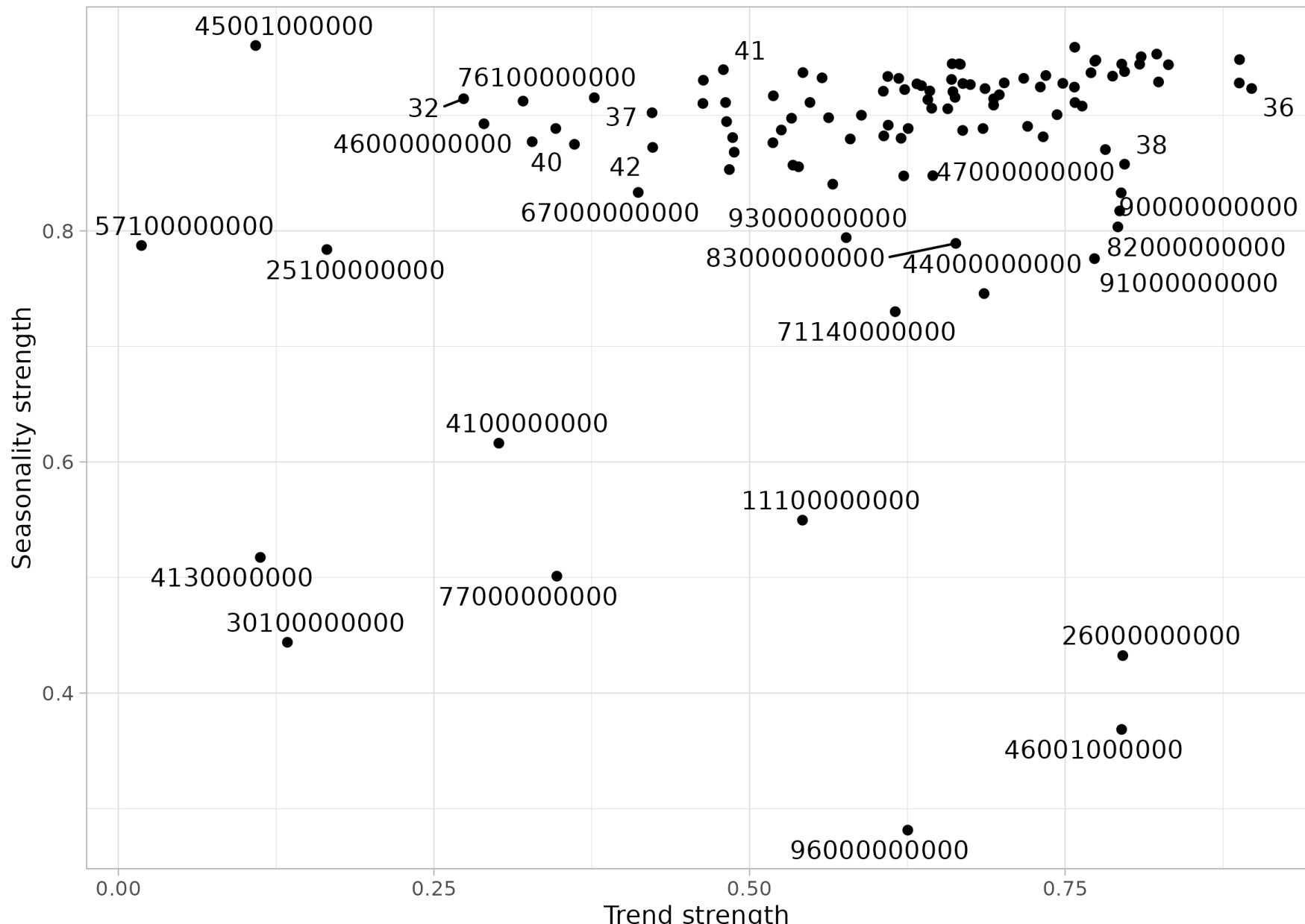
$$F_{trend} = \max \left\{ 1 - \frac{sVar(remainder)}{sVar(trend + remainder)}, 0 \right\}.$$

- Seasonality strength:

$$F_{seas} = \max \left\{ 1 - \frac{sVar(remainder)}{sVar(seas + remainder)}, 0 \right\}.$$

# Strength of trend and seasonality

Russian regions proximity by the strength of trend and seasonality



# Series Characteristics: Summary

- ACF — coefficients in **paired** regressions or correlations
- PACF — coefficients in **multiple** regressions or correlations
- STL allows you to measure **strength of trend and seasonality** in comparison to the residual component

# **ETS Model (Part I)**

# ETS Model: Plan

- ETS as a model
- Formulas for predictions
- Adding a trend
- Idea of damped trend

# How many ETS models in total?

ETS — **Error, Trend, Seasonality** (error, trend, seasonality)

**Error:** A, M

**Trend:** N, A, Ad, M, Md

**Seasonality:** N, A, M

A — **additive** component

M — **multiplicative** component

N — no component

d — **damping** for the trend

Formally: **30 options**

# Historical names

ETS(ANN) — simple exponential smoothing

ETS(AAA) — an additive Holt-Winters method

ETS(AAM) — the multiplicative Holt-Winters method

ETS(AAdM) — Holt-Winters method with a fading trend



# ETS(ANN) terminology

$y_t$  — the observed series;

$\ell_t$  — trend, cleaned series;

$u_t$  — a random error

$$y_t = \ell_{t-1} + u_t;$$

$$\ell_t = \ell_{t-1} + \alpha u_t, \text{ starts at } \ell_0;$$

$$u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent}$$

Parameters:  $\alpha, \sigma^2, \ell_0$

# Recognized?

ETS(ANN) is a generalization of **random walk**

$$\begin{cases} y_t = \ell_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \end{cases}$$

Substitute  $\alpha = 1$ :

$$y_t = \ell_t = \ell_{t-1} + u_t$$

# Estimation

Maximum likelihood method is used for estimation

Main idea: decompose the likelihood into a sum

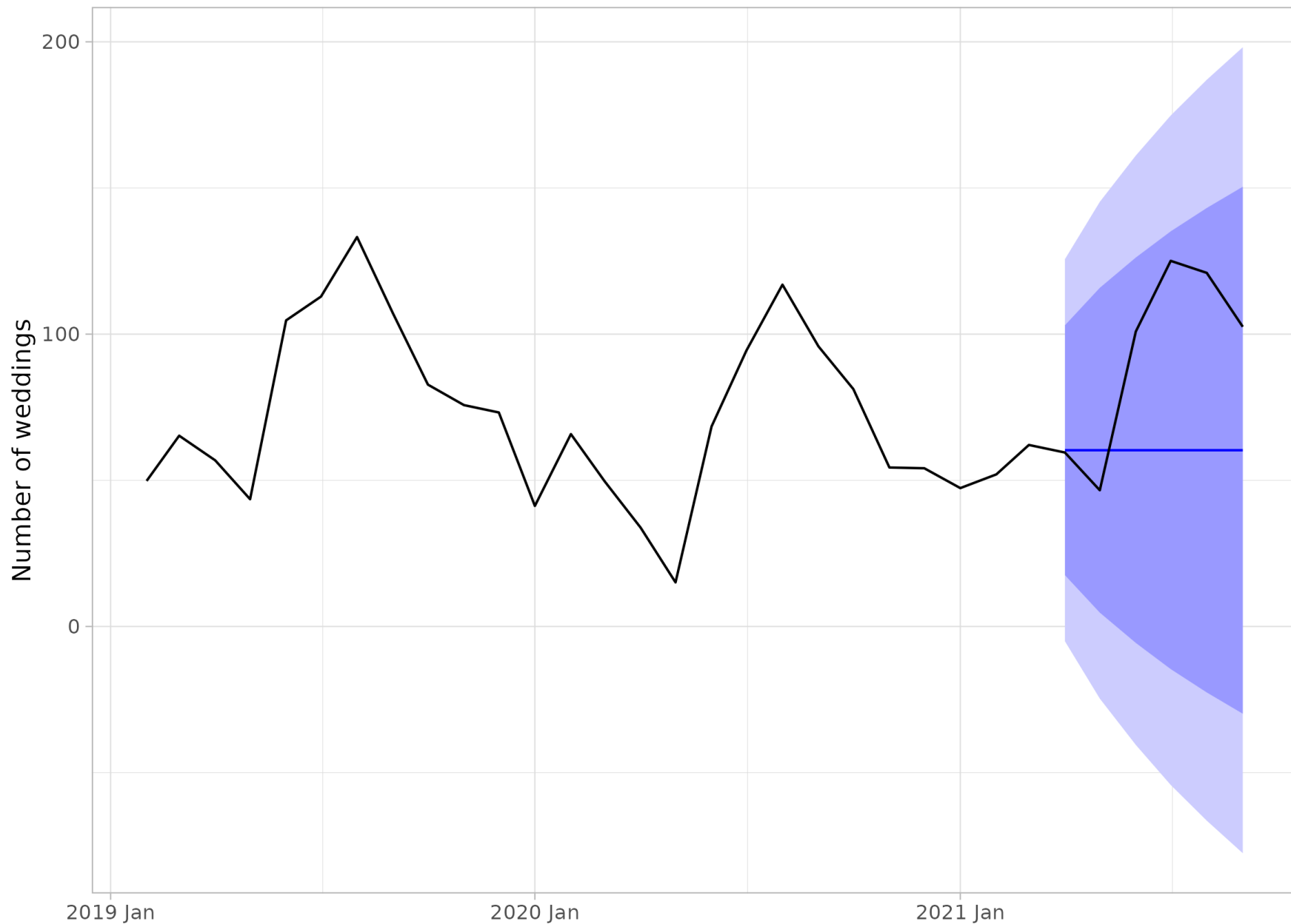
$$\ln L(y \mid \theta) = \ln L(y_1 \mid \theta) + \ln L(y_2 \mid y_1, \theta) + \dots + \\ + \ln L(y_T \mid y_{T-1}, \dots, y_1, \theta),$$

where  $\theta = (\alpha, \ell_0, \sigma^2)$

Unfortunately, there are no explicit formulas for the estimators

# Forecasting

ETS(ANN)



# Forecast 1 step ahead

Luckily, there are **recurrent formulas** for predictions

$$\begin{cases} y_t = \ell_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \end{cases}$$

$$y_{T+1} = \ell_T + u_{T+1}$$

$$(y_{T+1} \mid \mathcal{F}_T) \sim \mathcal{N}(\ell_T; \sigma^2)$$

# Forecast 2 steps ahead

$$\begin{cases} y_t = \ell_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \end{cases}$$

$$y_{T+2} = \ell_{T+1} + u_{T+2} = \ell_T + \alpha u_{T+1} + u_{T+2}$$

$$(y_{T+2} \mid \mathcal{F}_T) \sim \mathcal{N}(\ell_T; \sigma^2(\alpha^2 + 1))$$

# Predictive intervals

From distribution law

$$(y_{T+2} \mid \mathcal{F}_T) \sim \mathcal{N}(\ell_T; \sigma^2(\alpha^2 + 1))$$

we can derive a **predictive interval**

$$[\hat{\ell}_T - 1.96\hat{\sigma}\sqrt{\hat{\alpha}^2 + 1}; \hat{\ell}_T + 1.96\hat{\sigma}\sqrt{\hat{\alpha}^2 + 1}].$$

# What was discovered in the 1950s?

$$\begin{cases} y_t = \ell_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + \alpha u_t, \text{ starts at } \ell_0 \end{cases}$$

Let's rewrite the second equation:

$$\ell_t = \ell_{t-1} + \alpha(y_t - \ell_{t-1}) = \alpha y_t + (1 - \alpha)\ell_{t-1}$$

Simple exponential smoothing:

$$\hat{\ell}_1 = y_1$$

$$\hat{\ell}_t = \alpha y_t + (1 - \alpha)\hat{\ell}_{t-1}$$

$$\min_{\alpha} \sum (y_t - \hat{\ell}_t)^2$$



# Adding trend!

$y_t$  — the observed series;

$\ell_t$  — trend, cleaned series;

$b_t$  — current growth rate of the cleaned series;

$u_t$  — a random error

ETS(AAN):

A — **additive** error;

A — **additive** trend;

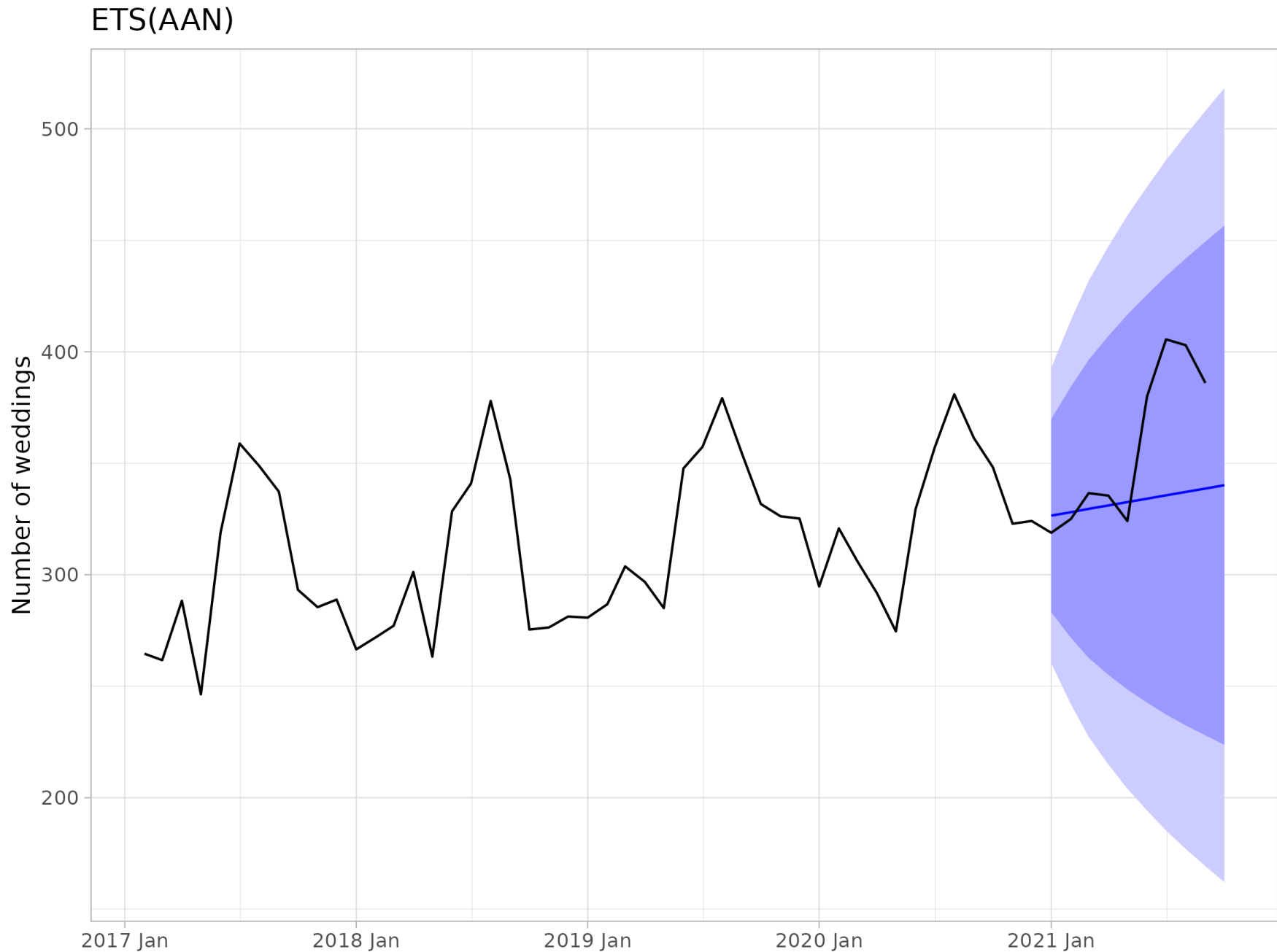
N — **no** seasonality

# ETS(AAN): equations

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + b_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + b_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \\ b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0; \end{array} \right.$$

Parameters:  $\alpha, \beta, \sigma^2, \ell_0, b_0$

# ETS(AAN): Forecasting



# Forecast 1 step ahead

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + b_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + b_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent.} \\ b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0 \end{array} \right.$$

$$y_{T+1} = \ell_T + b_T + u_{T+1}$$

$$(y_{T+1} \mid \mathcal{F}_T) \sim \mathcal{N}(\ell_T + b_T; \sigma^2)$$

# Forecast 2 steps ahead

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + b_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + b_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \\ b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0 \end{array} \right.$$

$$y_{T+2} = \ell_{T+1} + b_{T+1} + u_{T+2} = (\ell_T + b_T + \alpha u_{T+1}) + \\ + (b_T + \beta u_{T+1}) + u_{T+2}$$

$$(y_{T+2} \mid \mathcal{F}_T) \sim \mathcal{N}(\ell_T + 2b_T; \sigma^2((\alpha + \beta)^2 + 1))$$

# Problem with trend in ETS(AAN)

In the ETS(AAN) model **growth rate** of the  $\ell_t$  trend is defined by the formula

$$b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0$$

Consequently,

$$\mathbb{E}(b_t) = \mathbb{E}(b_{t-1}), \quad \mathbb{E}(b_{T+h} \mid b_T) = b_T$$

Long-term forecast of a positive indicator at  $b_T < 0$  will become negative

# Contradiction

In short-term we expect a change in the indicator:  
we **want a trend** in the model.

In long-term negative values are impossible:  
we **don't want a trend** in the model.

Solution: **damped** or **fading** trend.

# Extra parameters are expensive!

We want richer trend dynamics — we need **additional** parameters.

Additional parameters — risk **overfitting** of the model, **wider confidence intervals** for the remaining parameters.

Let's solve the problem with only **one** new parameter!



# Damped trend

We introduce the trend damping parameter  $\phi \in (0; 1)$  into the slope equation:

$$b_t = \phi b_{t-1} + \beta u_t, \text{ starts at } b_0$$

And for the rest of the equations:

$$\begin{cases} y_t = \ell_{t-1} + \phi b_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + \phi b_{t-1} + \alpha u_t, \text{ starts at } \ell_0 \end{cases}$$

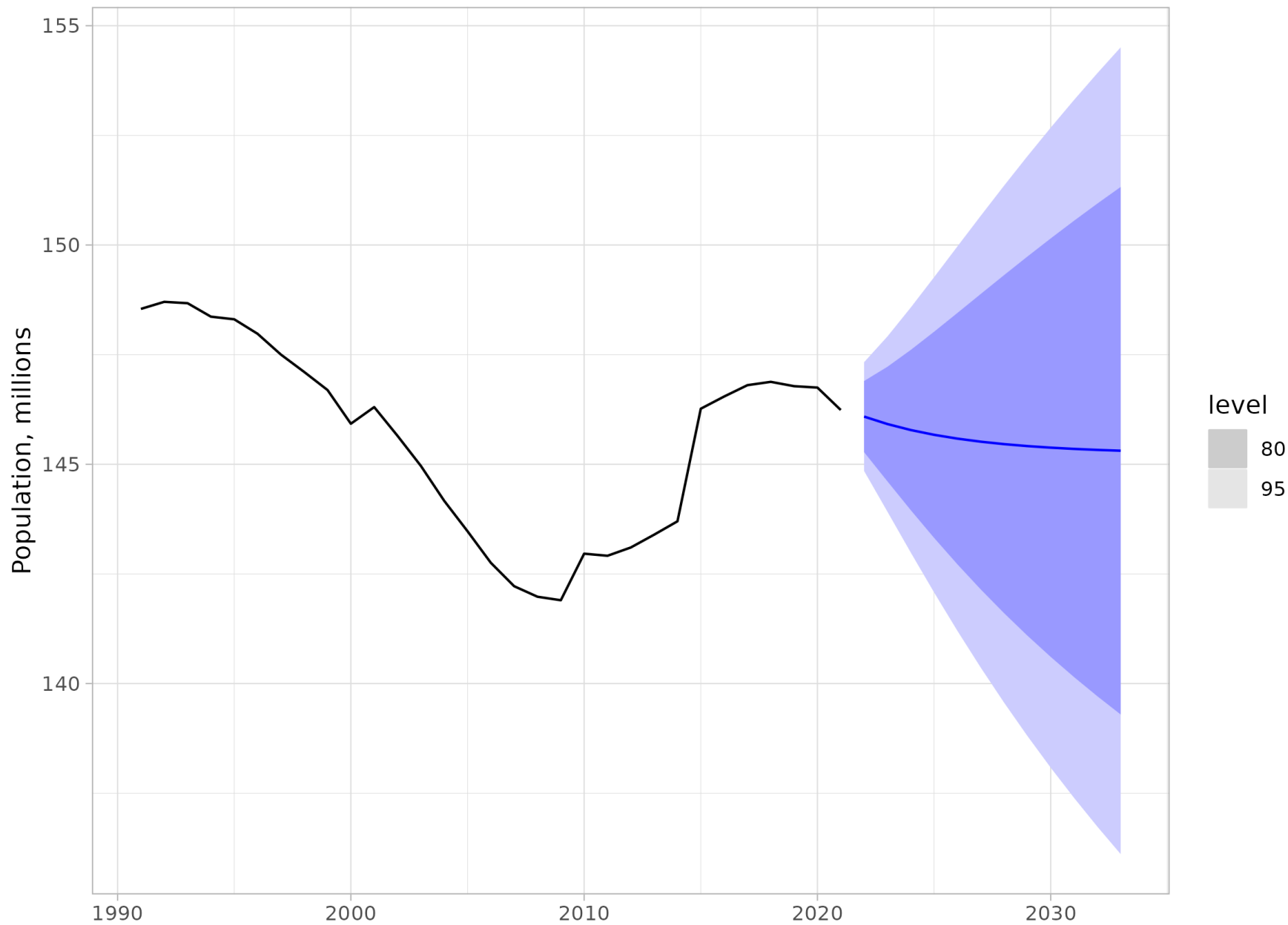
# General form of ETS(AAdN)

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + \phi b_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + \phi b_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ b_t = \phi b_{t-1} + \beta u_t, \text{ starts at } b_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \end{array} \right.$$

Parameters:  $\alpha, \sigma^2, \ell_0, b_0, \beta, \phi$

# ETS(AAdN): Forecasting

ETS(AAdN): population of Russia



# ETS Model: Summary

- Formulas for **exponential smoothing** have been around for a long time
- ETS — a wide class of modern models
- The slope of the trend line can change
- Damped trend: on a small forecasting horizon **there is a trend**, on a large horizon — **none**

# **ETS Model (Part II)**

# ETS Model: Plan

- Adding seasonality
- Formulas for predictions
- Decomposition into components
- Multiplicative components

# Adding seasonality!

$y_t$  — the observed series;

$\ell_t$  — trend, cleaned series;

$b_t$  — current growth rate of the cleaned series;

$s_t$  — seasonal component;

$u_t$  — a random error

ETS(AAA):

A — additive error;

A — additive trend;

A — additive seasonality

# ETS(AAA): equations

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + b_{t-1} + s_{t-12} + u_t; \\ \ell_t = \ell_{t-1} + b_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \\ b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0; \\ s_t = s_{t-12} + \gamma u_t; \text{ starts at } s_0, s_{-1}, \dots, s_{-11} \end{array} \right.$$

Parameters:  $\alpha, \beta, \gamma, \sigma^2, \ell_0, b_0, s_0, s_{-1}, \dots, s_{-11}$

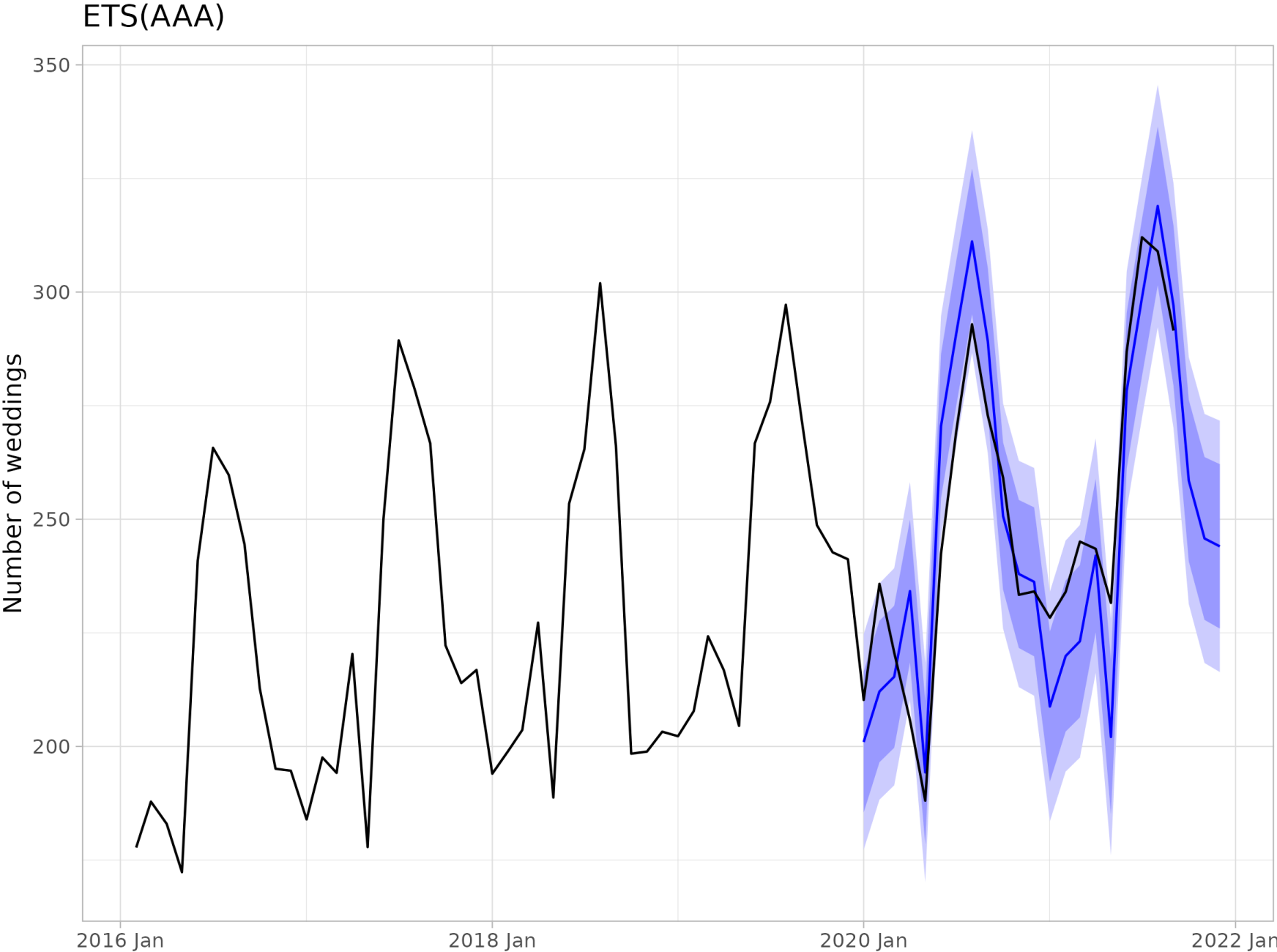
**Restriction:**  $s_0 + s_{-1} + \dots + s_{-11} = 0$

How many independent parameters are we estimating?

Correct answer: 17



# ETS(AAA): Forecasting



# Forecast 1 step ahead

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + b_{t-1} + s_{t-12} + u_t; \\ \ell_t = \ell_{t-1} + b_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent.} \\ b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0; \\ s_t = s_{t-12} + \gamma u_t \end{array} \right.$$

$$y_{T+1} = \ell_T + b_T + s_{T-11} + u_{T+1}$$

$$(y_{T+1} \mid \mathcal{F}_T) \sim \mathcal{N}(\ell_T + b_T + s_{T-11}; \sigma^2)$$

# Forecast 2 steps ahead

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + b_{t-1} + s_{t-12} + u_t; \\ \ell_t = \ell_{t-1} + b_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \\ b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0; \\ s_t = s_{t-12} + \gamma u_t \end{array} \right.$$

$$y_{T+2} = \ell_{T+1} + b_{T+1} + s_{T-10} + u_{T+2} = (\ell_T + b_T + \alpha u_{T+1}) + \\ + (b_T + \beta u_{T+1}) + s_{T-10} + u_{T+2}$$

$$(y_{T+2} \mid \mathcal{F}_T) \sim \mathcal{N}(\ell_T + 2b_T + s_{T-10}; \sigma^2((\alpha + \beta)^2 + 1))$$

# Decomposition for free!

Consider the output of ETS(AAA):

**Parameter estimates:**  $\hat{\alpha}, \hat{\beta}, \hat{\gamma}, \hat{\sigma}^2, \hat{\ell}_0, \hat{b}_0, \hat{s}_0, \hat{s}_{-1}, \dots, \hat{s}_{-11}$ .

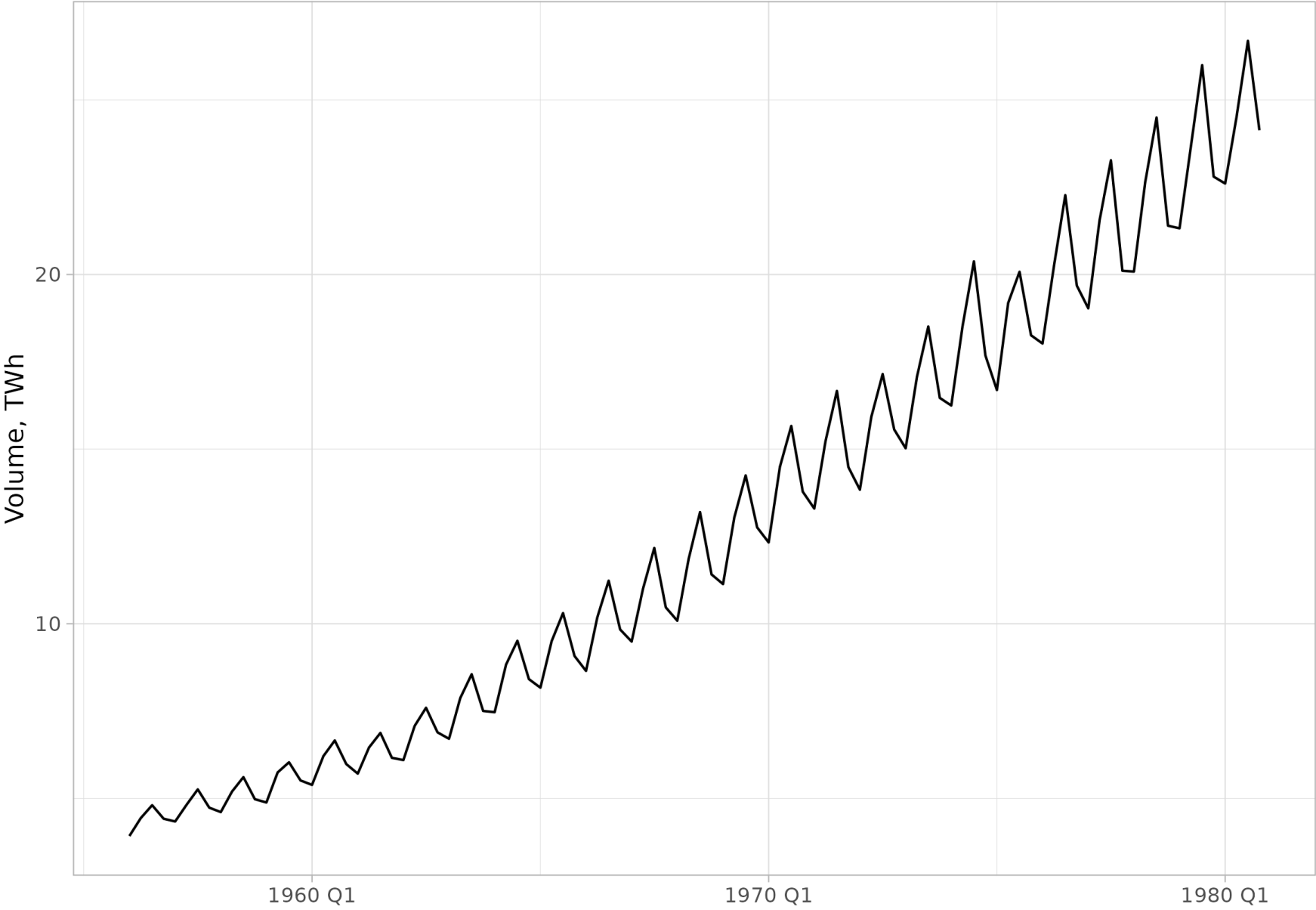
**Constraints:**  $\hat{s}_0 + \hat{s}_{-1} + \dots + \hat{s}_{-11} = 0$ .

**Estimated component values:**  $\hat{\ell}_t, \hat{b}_t, \hat{s}_t$ .

We automatically get **decomposition:**  $y_t = \hat{\ell}_t + \hat{s}_t + remainder_t$ .

# Oscillation amplitude can vary

Volume of electricity produced in Australia



# Various oscillation amplitude

Possible solutions:

- Switch to logarithms,  $y_t \rightarrow \ln y_t$
- Box-Cox transformation,  $y_t \rightarrow bc(y_t, \lambda)$
- Multiplicative components

# ETS(MNM): equations

ETS(MNM) for monthly data:

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} \cdot s_{t-12} \cdot (1 + u_t); \\ \ell_t = \ell_{t-1} \cdot (1 + \alpha u_t), \text{ starts at } \ell_0; \\ s_t = s_{t-12} \cdot (1 + \gamma u_t), \text{ starts at } s_0, \dots, s_{-11}; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent} \end{array} \right.$$

ETS(ANA):

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + s_{t-12} + u_t; \\ \ell_t = \ell_{t-1} + \alpha u_t, \text{ starts at } \ell_0; \\ s_t = s_{t-12} + \gamma u_t, \text{ starts at } s_0, \dots, s_{-11}; \end{array} \right.$$

# Units

Series  $y_t, \ell_t$  — initial units.

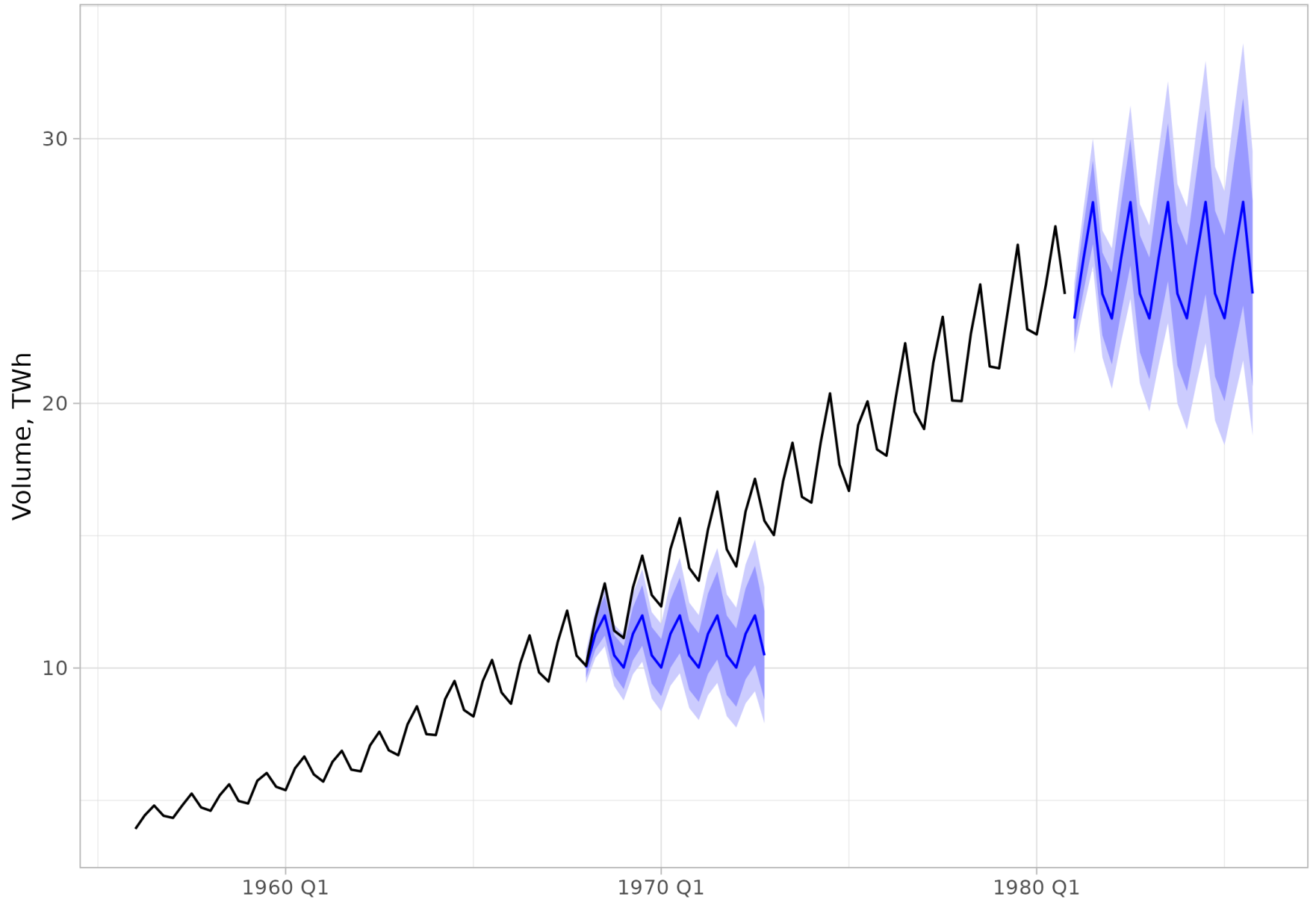
The  $s_t$  series is measured relative to one, for example,  $s_t = 0.9$  — 10% below the trend.

The  $u_t$  series is measured relative to zero, for example,  $u_t = -0.1$  — a 10% drop.



# ETS(MNM): Forecasting

ETS(MNM): Volume of electricity produced in Australia



# Which option to choose?

Different amplitude of fluctuations: indication of multiplicative models.

Automatic selection based on the AIC criterion works.

You can get the ETS(AAdA) model with seasonality.

Some of the multiplicative models can be numerically unstable or not implemented in the software.

# ETS Model: Summary

- The slope of the trend and seasonality may change
- Automatic decomposition into components
- Multiplicative models take into account **changing** oscillation amplitudes
- A lot of possible **combinations**

# Theta method

# Theta method: Plan

- An unexpected leader
- Author's version
- Special case of ETS

# Theta method

Appeared in 2000 and became a sensation at the competition M3 for predicting series.

Works for non-seasonal series.

Initially suggested without a statistical model.

# Author's version

1. Decompose the series into two **theta lines** ( $\theta = 0, \theta = 2$ )
2. Predict zero-line using linear regression
3. Predict the second line using ETS(ANN)
4. Average the forecasts

You can pre-delete seasonality and add it back in the end

# What is a theta line?

Zero theta line — regression on time:

$$\hat{y}_t = \hat{\beta}_1 + \hat{\beta}_2 t$$

Theta line for arbitrary theta:

$$\Delta^2 y_t^{new} = \theta \Delta^2 y_t$$



# Intuition

- The zero theta line catches the long-term trend of the series
- Theta line ( $\theta = 2$ ) catches the short-term trend:  
acceleration of theta line is  $\theta$  times stronger than of the initial series
- Averaging reduces the variance of predictions

# How is the theta line selected?

We take  $\theta = 2$ :

$$\Delta^2 y_t^{new} = 2\Delta^2 y_t$$

Or

$$y_t^{new} - 2y_{t-1}^{new} + y_{t-2}^{new} = 2(y_t - 2y_{t-1} + y_{t-2})$$

The new series  $y_t^{new}$  is completely determined by  $y_1^{new}, y_2^{new}$

We solve the optimization problem:

$$\sum_{t=1}^T (y_t - y_t^{new})^2 \rightarrow \min$$

# Statistical Model

Formal model appeared in 2003:

$$\begin{cases} y_t = \ell_t + b + u_t; \\ \ell_t = \ell_{t-1} + b + \alpha u_t; \\ \ell_1 = y_1 \end{cases}$$

Or:

$$\Delta y_t = b + (\alpha - 1)u_{t-1} + u_t$$

# Theta method — ETS variant

A special case of a more general model — ETS(AAN):

$$\left\{ \begin{array}{l} y_t = \ell_{t-1} + b_{t-1} + u_t; \\ \ell_t = \ell_{t-1} + b_{t-1} + \alpha u_t, \text{ starts at } \ell_1; \\ b_t = b_{t-1} + \beta u_t, \text{ starts at } b_0; \\ u_t \sim \mathcal{N}(0; \sigma^2) \text{ and are independent.} \end{array} \right.$$

Remove the trend stochasticity setting  $\beta = 0$ .

Nuances of initialization are possible.

# Theta method: Summary

- Works well for non-seasonal data
- A special variation of the ETS model