ARIMA processes

Definition

- Definition
- Random walk

- Definition
- Random walk
- Autocovariance function

- Definition
- Random walk
- Autocovariance function
- ACF, PACF

A stochastic process whose characteristics do not change over time

A stochastic process whose characteristics do not change over time

Weak or wide-sense stationarity

A process (y_t) is said to be weakly stationary, if for each t and k:

$$\begin{cases} \mathbb{E}(y_t) = \mu \\ \operatorname{Cov}(y_t, y_{t+k}) = \gamma_k \end{cases}$$

A stochastic process whose characteristics do not change over time

Weak or wide-sense stationarity

A process (y_t) is said to be weakly stationary, if for each t and k:

$$\begin{cases} \mathbb{E}(y_t) = \mu \\ \operatorname{Cov}(y_t, y_{t+k}) = \gamma_k \end{cases}$$

Strong or strict-sense stationarity

A process (y_t) is said to be strictly stationary, if for each k joint distribution of a r.v. $(y_t, y_{t+1}, y_{t+2}, \dots, y_{t+k})$ does not depend on t

Independent observations

The quantities (y_t) are independent and equally distributed with finite expectation μ_y and finite variance σ_y^2

Independent observations

The quantities (y_t) are independent and equally distributed with finite expectation μ_y and finite variance σ_u^2

$$\mu_y = \mathbb{E}(y_t)$$

Independent observations

The quantities (y_t) are independent and equally distributed with finite expectation μ_y and finite variance σ_y^2

$$\mu_y = \mathbb{E}(y_t)$$

$$\gamma_0 = \operatorname{Cov}(y_t, y_t) = \operatorname{Var}(y_t) = \sigma_y^2$$

Independent observations

The quantities (y_t) are independent and equally distributed with finite expectation μ_y and finite variance σ_y^2

$$\mu_y = \mathbb{E}(y_t)$$

$$\gamma_0 = \operatorname{Cov}(y_t, y_t) = \operatorname{Var}(y_t) = \sigma_y^2$$

$$\gamma_k = \operatorname{Cov}(y_t, y_{t+k}) = 0, \text{ for } k \geq 1$$

Random Walk

$$\begin{cases} y_0 = \mu \\ y_t = y_{t-1} + u_t, \text{ for } t \ge 1 \end{cases}$$

Random Walk

$$\begin{cases} y_0 = \mu \\ y_t = y_{t-1} + u_t, \text{ for } t \ge 1 \end{cases},$$

Explicitly:
$$y_t = \mu + u_1 + u_2 + ... + u_t$$
.

Random Walk

$$\begin{cases} y_0 = \mu \\ y_t = y_{t-1} + u_t, \text{ for } t \ge 1 \end{cases},$$

Explicitly:
$$y_t = \mu + u_1 + u_2 + \ldots + u_t$$
.
$$\mu_y = \mathbb{E}(y_t)$$

Random Walk

$$\begin{cases} y_0 = \mu \\ y_t = y_{t-1} + u_t, \text{ for } t \ge 1 \end{cases}$$

Explicitly:
$$y_t = \mu + u_1 + u_2 + \ldots + u_t$$
.
$$\mu_y = \mathbb{E}(y_t)$$

$$\gamma_0 = \operatorname{Cov}(y_t, y_t) = \operatorname{Var}(y_t) = \operatorname{Var}(\mu + u_1 + \dots + u_t) = t\sigma_u^2$$

Random Walk

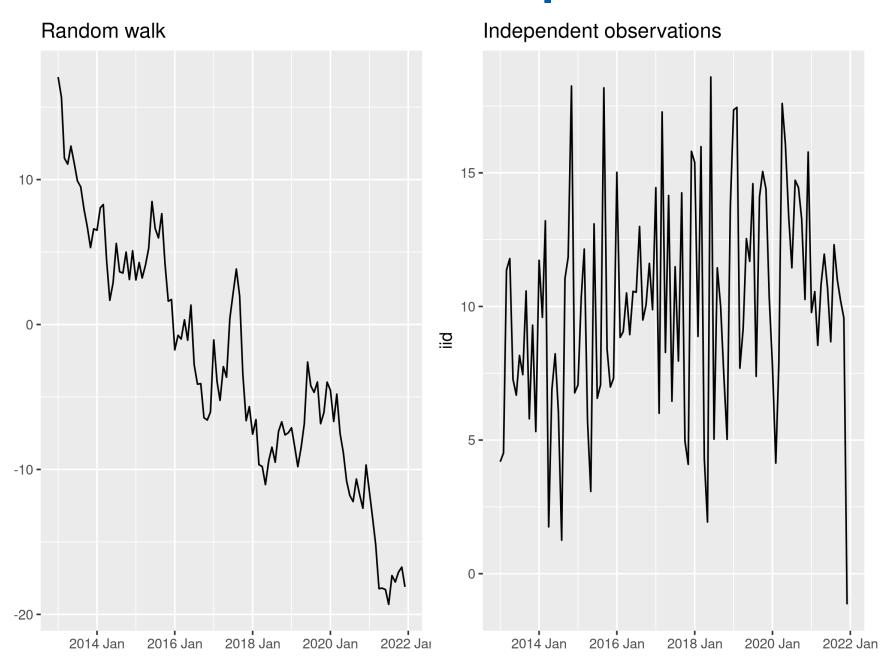
$$\begin{cases} y_0 = \mu \\ y_t = y_{t-1} + u_t, \text{ for } t \ge 1 \end{cases},$$

Explicitly:
$$y_t = \mu + u_1 + u_2 + \ldots + u_t$$
.
$$\mu_y = \mathbb{E}(y_t)$$

$$\gamma_0 = \operatorname{Cov}(y_t, y_t) = \operatorname{Var}(y_t) = \operatorname{Var}(\mu + u_1 + \dots + u_t) = t\sigma_u^2$$

$$\gamma_k = \text{Cov}(y_t, y_{t+k}) = \text{Cov}(y_t, y_t + u_{t+1} + \dots + u_{t+k}) = \text{Var}(y_t)$$

Random Walk vs Random Sample



Autocovariance function

Definition

For a stationary process (y_t) , the function $\gamma_k = \text{Cov}(y_t, y_{t+k})$ is called autocovariance

Autocovariance function

Definition

For a stationary process (y_t) , the function $\gamma_k = \text{Cov}(y_t, y_{t+k})$ is called autocovariance

Definition

For a stationary process (y_t) , the function $\rho_k = \operatorname{Corr}(y_t, y_{t+k})$ is called autocorrelation

ACF: autocorrelation function

$$\rho_k = \operatorname{Corr}(y_t, y_{y+j}) = \frac{\operatorname{Cov}(y_t, y_{y+j})}{\sqrt{\operatorname{Var}(y_t) \operatorname{Var}(y_{t+k})}} =$$

ACF: autocorrelation function

$$\rho_k = \operatorname{Corr}(y_t, y_{y+j}) = \frac{\operatorname{Cov}(y_t, y_{y+j})}{\sqrt{\operatorname{Var}(y_t)\operatorname{Var}(y_{t+k})}} = \frac{\gamma_k}{\sqrt{\gamma_0 \gamma_0}} = \frac{\gamma_k}{\gamma_0}$$

Partial Correlation

Definition

$$pCorr(U, D; R_1, R_2, ..., R_n) = Corr(U^*, D^*), \text{ where}$$

$$U^* = U - Best(U; R_1, R_2, ..., R_n),$$

$$D^* = D - Best(D; R_1, R_2, ..., R_n)$$

Partial Correlation

Definition

$$pCorr(U, D; R_1, R_2, ..., R_n) = Corr(U^*, D^*), \text{ where}$$

$$U^* = U - Best(U; R_1, R_2, ..., R_n),$$

$$D^* = D - Best(D; R_1, R_2, ..., R_n)$$

The values U^* and D^* are the versions of U and D uninfluenced by the covariates R_1, \ldots, R_n

$$Cov(U^*, R_i) = 0, \quad Cov(D^*, R_i) = 0.$$

PACF

Definition

For a stationary process (y_t) the function

$$\varphi_{kk} = \operatorname{pCorr}(y_t, y_{t+k}; y_{t+1}, \dots, y_{t+k-1})$$

is called partial autocorrelation

ACF and PACF: Intuition

For stationary process

ACF:

$$\rho_k = \operatorname{Corr}(y_t, y_{t+k})$$

Joint strength of relationship between y_t and y_{t+k}

• PACF:

$$\varphi_{kk} = \operatorname{pCorr}(y_t, y_{t+k}; y_{t+1}, \dots, y_{t+k-1})$$

Strength of relationship between y_t and y_{t+k} with the links through intermediate observations being broken

• Constants $\mathbb{E}(y_t)$, $\gamma_k = \text{Cov}(y_t, y_{t+k})$

- Constants $\mathbb{E}(y_t)$, $\gamma_k = \text{Cov}(y_t, y_{t+k})$
- The random sample is stationary

- Constants $\mathbb{E}(y_t)$, $\gamma_k = \text{Cov}(y_t, y_{t+k})$
- The random sample is stationary
- Random walk is non-stationary

- Constants $\mathbb{E}(y_t)$, $\gamma_k = \text{Cov}(y_t, y_{t+k})$
- The random sample is stationary
- Random walk is non-stationary
- Autocovariance function

- Constants $\mathbb{E}(y_t)$, $\gamma_k = \text{Cov}(y_t, y_{t+k})$
- The random sample is stationary
- Random walk is non-stationary
- Autocovariance function
- Partial correlation correlation with the effect of a set of controlling random variables removed

- Constants $\mathbb{E}(y_t)$, $\gamma_k = \text{Cov}(y_t, y_{t+k})$
- The random sample is stationary
- Random walk is non-stationary
- Autocovariance function
- Partial correlation correlation with the effect of a set of controlling random variables removed
- In the time series, we removed the effect of intermediate observations

MA Process

MA Process: Plan

Definition and notations with lags

MA Process: Plan

- Definition and notations with lags
- Stationarity

MA Process: Plan

- Definition and notations with lags
- Stationarity
- Predictability

MA Process: Plan

- Definition and notations with lags
- Stationarity
- Predictability
- Reversibility

Definition

$$Ly_t = y_{t-1}$$

Definition

$$Ly_t = y_{t-1}$$

$$L^2 y_t = L \cdot L \cdot y_t = L \cdot y_{t-1} = y_{t-2}$$

Definition

$$Ly_t = y_{t-1}$$

$$L^{2}y_{t} = L \cdot L \cdot y_{t} = L \cdot y_{t-1} = y_{t-2}$$

$$\Delta y_t = y_t - y_{t-1} = (1 - L)y_t$$

Definition

$$Ly_t = y_{t-1}$$

$$L^{2}y_{t} = L \cdot L \cdot y_{t} = L \cdot y_{t-1} = y_{t-2}$$

$$\Delta y_t = y_t - y_{t-1} = (1 - L)y_t$$

$$\Delta_{12}y_t = y_t - y_{t-12} = (1 - L^{12})y_t$$

MA process

Definition

Process (y_t) , which can be represented as

$$y_t = \mu + u_t + \alpha_1 u_{t-1} + \ldots + \alpha_q u_{t-q},$$

where $\alpha_q \neq 0$ and (u_t) is white noise, is called the MA(q) process

MA — Moving Average

MA process

Definition

Process (y_t) , which can be represented as

$$y_t = \mu + u_t + \alpha_1 u_{t-1} + \ldots + \alpha_q u_{t-q},$$

where $\alpha_q \neq 0$ and (u_t) is white noise, is called the MA(q) process

MA — Moving Average

Example MA(1) process:

$$y_t = 5 + u_t + 0.3u_{t-1},$$

where (u_t) is some white noise

Notation with lags

MA with lag polynomial

Process (y_t) , which can be represented as

$$y_t = \mu + P(L)u_t,$$

where P(L) is a polynomial of degree q in lag L with P(0) = 1, and (u_t) is white noise, is called MA(q) a process

Notation with lags

MA with lag polynomial

Process (y_t) , which can be represented as

$$y_t = \mu + P(L)u_t,$$

where P(L) is a polynomial of degree q in lag L with P(0) = 1, and (u_t) is white noise, is called MA(q) a process

An example MA(2) process:

$$y_t = 5 + (1 - 0.2L + 0.3L^2)u_t,$$

where (u_t) is white noise

ACF and Forecasts

Traditionally MA(q) process is evaluated assuming joint normality of (y_t) .

ACF and Forecasts

Traditionally MA(q) process is evaluated assuming joint normality of (y_t) .

Zero $\rho_k = 0$ for k > q implies independence of y_t and y_{t+k} .

ACF and Forecasts

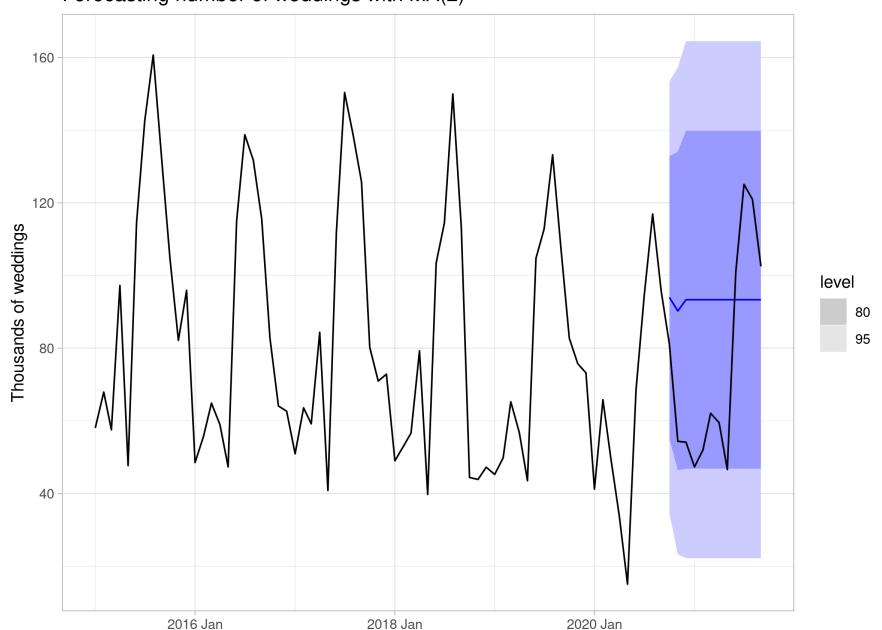
Traditionally MA(q) process is evaluated assuming joint normality of (y_t) .

Zero $\rho_k = 0$ for k > q implies independence of y_t and y_{t+k} . Forecasts more than q steps ahead are exactly the same.

$$(y_{T+q+1} \mid \mathcal{F}_T) \sim (y_{T+q+2} \mid \mathcal{F}_T) \sim (y_{T+q+3} \mid \mathcal{F}_T) \sim \dots$$

Predictions for MA(2)

Forecasting number of weddings with MA(2)



$MA(\infty)$

Definition

Process (y_t) , which can be represented as

$$y_t = \mu + u_t + \alpha_1 u_{t-1} + \alpha_2 u_{t-2} + \dots,$$

where (u_t) is white noise, an infinite number of $\alpha_i \neq 0$ and $\sum_{i=1}^{\infty} \alpha_i^2 < \infty$, is called the $MA(\infty)$ process

$MA(\infty)$

Definition

Process (y_t) , which can be represented as

$$y_t = \mu + u_t + \alpha_1 u_{t-1} + \alpha_2 u_{t-2} + \dots,$$

where (u_t) is white noise, an infinite number of $\alpha_i \neq 0$ and $\sum_{i=1}^{\infty} \alpha_i^2 < \infty$, is called the $MA(\infty)$ process

 $MA(\infty)$:

$$y_t = 5 + u_t + 0.5u_{t-1} + 0.5^2u_{t-2} + 0.5^3u_{t-3} + \dots$$

$MA(\infty)$

Definition

Process (y_t) , which can be represented as

$$y_t = \mu + u_t + \alpha_1 u_{t-1} + \alpha_2 u_{t-2} + \dots,$$

where (u_t) is white noise, an infinite number of $\alpha_i \neq 0$ and $\sum_{i=1}^{\infty} \alpha_i^2 < \infty$, is called the $MA(\infty)$ process

 $MA(\infty)$:

$$y_t = 5 + u_t + 0.5u_{t-1} + 0.5^2u_{t-2} + 0.5^3u_{t-3} + \dots$$

And this is not allowed:

$$y_t = 5 + u_t + \frac{1}{\sqrt{2}}u_{t-1} + \frac{1}{\sqrt{3}}u_{t-2} + \frac{1}{\sqrt{4}}u_{t-3} + \dots$$

Convergences

Theorem

If a $\sum_{i=0}^{\infty} \alpha_i^2 < \infty$ and (u_t) is a zero-mean stationary process, then the sequence of partial sums y_t^q of the form

$$y_t^q = \mu + \sum_{i=0}^q \alpha_i u_{t-i}$$

converges for $q \to \infty$ in mean, in probability, and in distribution

Nuance: the convergence of the weighted sum is guaranteed for the stationary (u_t)

Convergences

Theorem

If a $\sum_{i=0}^{\infty} \alpha_i^2 < \infty$ and (u_t) is a zero-mean stationary process, then the sequence of partial sums y_t^q of the form

$$y_t^q = \mu + \sum_{i=0}^q \alpha_i u_{t-i}$$

converges for $q \to \infty$ in mean, in probability, and in distribution

Nuance: the convergence of the weighted sum is guaranteed for the stationary (u_t)

Bonus

...and the resulting process (y_t) is stationary

Wald's Theorem

Theorem

If (y_t) is a stationary process, then it can be represented as:

$$y_t = \sum_{i=0}^{\infty} \alpha_i u_{t-i} + r_t,$$

where

- (u_t) white noise,
- $\sum \alpha_i^2 < \infty$,
- r_t is a linear predictable random process,
- $Cov(u_t, r_t) = 0$

Predictable Process

Correct definition

A process (r_t) is called linearly predictable if

- (r_t) is stationary,
- $r_t = \beta_0 + \beta_1 r_{t-1} + \beta_2 r_{t-2} + \dots + \beta_p r_{t-p}$

Reversibility condition

Characteristic representation

The equation MA(q) of the process satisfies the reversibility condition if the characteristic polynomial $\phi(\lambda)$ has all roots $|\lambda_i| < 1$

Reversibility condition

Characteristic representation

The equation MA(q) of the process satisfies the reversibility condition if the characteristic polynomial $\phi(\lambda)$ has all roots $|\lambda_i| < 1$

Lag representation

The equation MA(q) of the process satisfies the reversibility condition if all roots of the lag polynomial P(L) are $|\ell_i| > 1$

Example of reversible notation MA(1)

$$y_t = 5 + u_t + 0.5u_{t-1}, \quad \sigma_u^2 = 4$$

Example of reversible notation MA(1)

$$y_t = 5 + u_t + 0.5u_{t-1}, \quad \sigma_u^2 = 4$$
$$\lambda^1 + 0.5 \cdot \lambda^0$$
$$\phi(\lambda) = \lambda + 0.5$$

Example of reversible notation MA(1)

$$y_t = 5 + u_t + 0.5u_{t-1}, \quad \sigma_u^2 = 4$$
$$\lambda^1 + 0.5 \cdot \lambda^0$$
$$\phi(\lambda) = \lambda + 0.5$$
$$\lambda_1 = -0.5$$

Example of MA(1) irreversible notation

$$y_t = 5 + u_t + 2u_{t-1}, \quad \sigma_u^2 = 1$$

Example of MA(1) irreversible notation

$$y_t = 5 + u_t + 2u_{t-1}, \quad \sigma_u^2 = 1$$
$$\lambda^1 + 2 \cdot \lambda^0$$
$$\phi(\lambda) = \lambda + 2$$

Example of MA(1) irreversible notation

$$y_t = 5 + u_t + 2u_{t-1}, \quad \sigma_u^2 = 1$$
$$\lambda^1 + 2 \cdot \lambda^0$$
$$\phi(\lambda) = \lambda + 2$$
$$\lambda_1 = -2$$

Nuance

Difference

Stationarity is a property of the (y_t) process itself.

Reversibility is a property of the equation (process notation) for (y_t) .

Nuance

Difference

Stationarity is a property of the (y_t) process itself.

Reversibility is a property of the equation (process notation) for (y_t) .

MA(q) has a single notation when reversible

• MA(q) — weighting of several white noises

- MA(q) weighting of several white noises
- MA(q) is a stationary process

- MA(q) weighting of several white noises
- MA(q) is a stationary process
- ACF vanishes sharply, PACF tends to zero

- MA(q) weighting of several white noises
- MA(q) is a stationary process
- ACF vanishes sharply, PACF tends to zero
- Reversibility condition: roots of the characteristic polynomial $|\lambda_i| < 1$ or roots of the lag polynomial $|\ell_i| > 1$.

ARMA equation

ARMA equation: Plan

Definition

ARMA equation: Plan

- Definition
- Non-uniqueness of solutions

Goal: A simple equation for a wide variety of processes

Goal: A simple equation for a wide variety of processes Problems:

Non-uniqueness of equation for one process

Goal: A simple equation for a wide variety of processes Problems:

Non-uniqueness of equation for one process

Requirement reversibility of the equation

Goal: A simple equation for a wide variety of processes Problems:

Non-uniqueness of equation for one process

Requirement reversibility of the equation

• $MA(\infty)$ has infinite number of parameters

Goal: A simple equation for a wide variety of processes Problems:

Non-uniqueness of equation for one process

Requirement reversibility of the equation

• $MA(\infty)$ has infinite number of parameters

Let's try adding lags y_t to the equation!

 $y_t - y_{t-1} = u_t - u_{t-1}$, where (u_t) is white noise

$$y_t - y_{t-1} = u_t - u_{t-1}$$
, where (u_t) is white noise

$$y_t - y_{t-1} = u_t - u_{t-1}$$
, where (u_t) is white noise

Solutions:

• $y_t = u_t$;

$$y_t - y_{t-1} = u_t - u_{t-1}$$
, where (u_t) is white noise

- $y_t = u_t$;
- $y_t = u_t 0.7$;

$$y_t - y_{t-1} = u_t - u_{t-1}$$
, where (u_t) is white noise

- $y_t = u_t$;
- $y_t = u_t 0.7$;
- $y_t = u_t 0.8$

$$y_t - y_{t-1} = u_t - u_{t-1}$$
, where (u_t) is white noise

- $y_t = u_t$;
- $y_t = u_t 0.7$;
- $y_t = u_t 0.8$

$$y_t - y_{t-1} = u_t - u_{t-1}$$
, where (u_t) is white noise

Solutions:

- $y_t = u_t$;
- $y_t = u_t 0.7$;
- $y_t = u_t 0.8$

Infinite number of solutions

ARMA equation

Definition

Equation

$$y_t = c + \beta_1 y_{t-1} + \ldots + \beta_p y_{t-p} + u_t + \alpha_1 u_{t-1} + \ldots + \alpha_q u_{t-q},$$

where (u_t) is white noise, we'll call an ARMA equation

ARMA — Autoregression and Moving Average

ARMA equation

Definition

Equation

$$y_t = c + \beta_1 y_{t-1} + \ldots + \beta_p y_{t-p} + u_t + \alpha_1 u_{t-1} + \ldots + \alpha_q u_{t-q},$$

where (u_t) is white noise, we'll call an ARMA equation

ARMA — Autoregression and Moving Average

Definition

An equation of the form $P(L)y_t=c+Q(L)u_t$, where (u_t) is white noise, P(L) and Q(L) are lag polynomials with P(0)=Q(0)=1, we'll call an ARMA equation

ARMA equation: Summary

An equation is not a process!

ARMA equation: Summary

An equation is not a process! Why?

One equation has many solutions

ARMA equation: Summary

An equation is not a process!

Why?

- One equation has many solutions
- One process can be described by several equations

Equation irreducibility

- Equation irreducibility
- Solution structure

- Equation irreducibility
- Solution structure
- ARMA process

Irreducibility of Equation

Definition

ARMA an equation of the form $P(L)y_t=c+Q(L)u_t$ is called irreducible, if the polynomials P(L) and Q(L) do not have common roots

Irreducibility of Equation

Definition

ARMA an equation of the form $P(L)y_t=c+Q(L)u_t$ is called irreducible, if the polynomials P(L) and Q(L) do not have common roots

Reducible Equation:

$$y_t - y_{t-1} = u_t - u_{t-1}$$
 or $(1 - L)y_t = (1 - L)u_t$

Irreducibility of Equation

Definition

ARMA an equation of the form $P(L)y_t = c + Q(L)u_t$ is called irreducible, if the polynomials P(L) and Q(L) do not have common roots

Reducible Equation:

$$y_t - y_{t-1} = u_t - u_{t-1} \text{ or } (1 - L)y_t = (1 - L)u_t$$

Irreducible equation:

$$y_t - y_{t-1} = u_t - 0.5u_{t-1}$$
 or $(1 - L)y_t = (1 - 0.5L)u_t$

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

•
$$y_0 = 0$$

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

•
$$y_0 = 0$$

$$y_1 = u_1, y_2 = u_2 + 0.5u_1, y_3 = u_3 + 0.5u_2 + 0.25u_1, \dots$$

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

•
$$y_0 = 0$$

$$y_1 = u_1, y_2 = u_2 + 0.5u_1, y_3 = u_3 + 0.5u_2 + 0.25u_1, \dots$$

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

•
$$y_0 = 0$$

$$y_1 = u_1$$
, $y_2 = u_2 + 0.5u_1$, $y_3 = u_3 + 0.5u_2 + 0.25u_1$, ...

•
$$y_0 = 2u_1$$

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

•
$$y_0 = 0$$

$$y_1 = u_1$$
, $y_2 = u_2 + 0.5u_1$, $y_3 = u_3 + 0.5u_2 + 0.25u_1$, ...

•
$$y_0 = 2u_1$$

$$y_1 = 2u_1$$
, $y_2 = u_2 + u_1$, $y_3 = u_3 + 0.5u_2 + 0.5u_1$, ...

Irreducible ARMA equation:

$$y_t = 0.5y_{t-1} + u_t$$
, where (u_t) is white noise

Let's try different initial conditions:

•
$$y_0 = 0$$

$$y_1 = u_1$$
, $y_2 = u_2 + 0.5u_1$, $y_3 = u_3 + 0.5u_2 + 0.25u_1$, ...

•
$$y_0 = 2u_1$$

$$y_1 = 2u_1$$
, $y_2 = u_2 + u_1$, $y_3 = u_3 + 0.5u_2 + 0.5u_1$, ...

The initial conditions also determine the past y_t !

ARMA Solutions

Theorem I

Any ARMA equation with at least one y_t lag has an infinite number of solutions

ARMA Solutions

Theorem I

Any ARMA equation with at least one y_t lag has an infinite number of solutions

Theorem II

In order to obtain a unique solution of an ARMA equation of the form $P(L)y_t=c+Q(L)u_t$, it suffices to specify the initial conditions in an amount equal to the power of P(L)

ARMA Solutions

Theorem I

Any ARMA equation with at least one y_t lag has an infinite number of solutions

Theorem II

In order to obtain a unique solution of an ARMA equation of the form $P(L)y_t=c+Q(L)u_t$, it suffices to specify the initial conditions in an amount equal to the power of P(L)

$$y_t = 0.6y_{t-1} + 0.08y_{t-2} + u_t$$

ARMA Solutions

Theorem I

Any ARMA equation with at least one y_t lag has an infinite number of solutions

Theorem II

In order to obtain a unique solution of an ARMA equation of the form $P(L)y_t=c+Q(L)u_t$, it suffices to specify the initial conditions in an amount equal to the power of P(L)

$$y_t = 0.6y_{t-1} + 0.08y_{t-2} + u_t$$
 and $y_0 = u_0, y_1 = u_0 + 4$

Correct theorem

- has exactly one stationary solution if the lag polynomial $P(\ell)$ has all roots $|\ell_i| \neq 1$;
- has no stationary solutions if the lag polynomial $P(\ell)$ has a root with $|\ell_i|=1$

Correct theorem

- has exactly one stationary solution if the lag polynomial $P(\ell)$ has all roots $|\ell_i| \neq 1$;
- has no stationary solutions if the lag polynomial $P(\ell)$ has a root with $|\ell_i|=1$
- $y_t = 0.5y_{t-1} + u_t, P(L) = 1 0.5L, \ell_1 = 2$:

Correct theorem

- has exactly one stationary solution if the lag polynomial $P(\ell)$ has all roots $|\ell_i| \neq 1$;
- has no stationary solutions if the lag polynomial $P(\ell)$ has a root with $|\ell_i|=1$
- $y_t = 0.5y_{t-1} + u_t, P(L) = 1 0.5L, \ell_1 = 2$: one stationary solution;

Correct theorem

- has exactly one stationary solution if the lag polynomial $P(\ell)$ has all roots $|\ell_i| \neq 1$;
- has no stationary solutions if the lag polynomial $P(\ell)$ has a root with $|\ell_i|=1$
- $y_t = 0.5y_{t-1} + u_t, P(L) = 1 0.5L, \ell_1 = 2$: one stationary solution;
- $y_t = y_{t-1} + u_t, P(L) = 1 L, \ell_1 = 1$: no stationary solutions

AR process

Definition

AR(p) process with equation

$$y_t = c + \beta_1 y_{t-1} + \ldots + \beta_p y_{t-p} + u_t,$$

where (u_t) is white noise and $\beta_p \neq 0$, is the solution of this equation in the form of $MA(\infty)$ with respect to (u_t)

AR process

Definition

AR(p) process with equation

$$y_t = c + \beta_1 y_{t-1} + \ldots + \beta_p y_{t-p} + u_t,$$

where (u_t) is white noise and $\beta_p \neq 0$, is the solution of this equation in the form of $MA(\infty)$ with respect to (u_t)

Definition with lags

AR(p) process with equation

$$P(L)y_t = c + u_t,$$

where (u_t) is white noise, P(L) has power p and P(0)=1, is the solution of this equation in the form of $MA(\infty)$ with respect to (u_t)

Definitions by different authors

In our definition of AR(p), the process is necessarily stationary.

Definitions by different authors

In our definition of AR(p), the process is necessarily stationary. Some authors do not include the requirement of stationarity in the definition of AR(p).

ARMA process

Definition

ARMA(p,q) process with equation

$$y_t = c + \beta_1 y_{t-1} + \ldots + \beta_p y_{t-p} + u_t + \alpha_1 u_{t-1} + \ldots + \alpha_q u_{t-q},$$

where (u_t) is white noise, $\beta_p \neq 0$, $\alpha_q \neq 0$ and the equation is irreducible, is the solution of this equation in the form of $MA(\infty)$ with respect to (u_t)

ARMA process

Definition

ARMA(p,q) process with equation

$$y_t = c + \beta_1 y_{t-1} + \ldots + \beta_p y_{t-p} + u_t + \alpha_1 u_{t-1} + \ldots + \alpha_q u_{t-q},$$

where (u_t) is white noise, $\beta_p \neq 0$, $\alpha_q \neq 0$ and the equation is irreducible, is the solution of this equation in the form of $MA(\infty)$ with respect to (u_t)

Definition with lags

ARMA(p,q) process with equation

$$P(L)y_t = c + Q(L)u_t,$$

where (u_t) is white noise, P(L) and Q(L) have powers p and q, respectively, are irreducible and P(0) = Q(0) = 1, is the solution of this equation in the form $MA(\infty)$ with respect to (u_t)

What about non-uniqueness?

The same ARMA(p,q) process (y_t) can be described by different equations!

What about non-uniqueness?

The same ARMA(p,q) process (y_t) can be described by different equations!

Reversibility

If the series (y_t) is an ARMA(p,q) process with the equation $P(L)y_t=c+Q(L)u_t$, then this equation will be unique if the MA part satisfies the reversibility condition.

What about non-uniqueness?

The same ARMA(p,q) process (y_t) can be described by different equations!

Reversibility

If the series (y_t) is an ARMA(p,q) process with the equation $P(L)y_t=c+Q(L)u_t$, then this equation will be unique if the MA part satisfies the reversibility condition.

Condition for ARMA reversibility of the equation:

- the characteristic polynomial $\phi_{MA}(\lambda)$ has all roots $|\lambda_i| < 1$;
- the lag polynomial Q(L) has all roots $|\ell_i| > 1$

The process is stationary by definition

- The process is stationary by definition
- AR and MA processes are special cases of the ARMA process

- The process is stationary by definition
- AR and MA processes are special cases of the ARMA process
- Theoretical ACF and PACF decrease exponentially

- The process is stationary by definition
- AR and MA processes are special cases of the ARMA process
- Theoretical ACF and PACF decrease exponentially
- The reversibility condition guarantees uniqueness

- The process is stationary by definition
- AR and MA processes are special cases of the ARMA process
- Theoretical ACF and PACF decrease exponentially
- The reversibility condition guarantees uniqueness
- An irreducible equation either has a unique stationary solution or it does not exist

ARIMA process

ARIMA process: Plan

• Stationarity of ARMA

ARIMA process: Plan

- Stationarity of ARMA
- Definition of ARIMA

ARIMA process: Plan

- Stationarity of ARMA
- Definition of ARIMA
- Differencing

• Process $y_t \sim ARMA(p,q)$ is stationary by definition:

• Process $y_t \sim ARMA(p,q)$ is stationary by definition:

$$\mathbb{E}(y_t) = \mu_y$$
, $\operatorname{Var}(y_t) = \gamma_0$, $\operatorname{Cov}(y_t, y_{t-k}) = \gamma_k$

- Process $y_t \sim ARMA(p,q)$ is stationary by definition: $\mathbb{E}(y_t) = \mu_y$, $Var(y_t) = \gamma_0$, $Cov(y_t, y_{t-k}) = \gamma_k$
- In the canonical notation ARMA(p,q) of the process $P(L)y_t=c+Q(L)u_t$ for the polynomial P(L) all roots $|\ell|>1$

- Process $y_t \sim ARMA(p,q)$ is stationary by definition: $\mathbb{E}(y_t) = \mu_y$, $Var(y_t) = \gamma_0$, $Cov(y_t, y_{t-k}) = \gamma_k$
- In the canonical notation ARMA(p,q) of the process $P(L)y_t=c+Q(L)u_t$ for the polynomial P(L) all roots $|\ell|>1$
- When estimating the ARMA(p,q) process by the maximum likelihood method, these restrictions are imposed a priori

Definition

The random process (y_t) is called the ARIMA(p,1,q) w.r.t. the white noise process (u_t) , if (y_t) is non-stationary, but Δy_t is a stationary ARMA(p,q) process w.r.t. the white noise (u_t)

Definition

The random process (y_t) is called the ARIMA(p,1,q) w.r.t. the white noise process (u_t) , if (y_t) is non-stationary, but Δy_t is a stationary ARMA(p,q) process w.r.t. the white noise (u_t)

Definition

The random process (y_t) is called the ARIMA(p,2,q) w.r.t. the white noise process (u_t) , if (y_t) and (Δy_t) are non-stationary, but $\Delta^2 y_t$ is a stationary ARMA(p,q) process w.r.t. the white noise (u_t)

Definition

The random process (y_t) is called the ARIMA(p,1,q) w.r.t. the white noise process (u_t) , if (y_t) is non-stationary, but Δy_t is a stationary ARMA(p,q) process w.r.t. the white noise (u_t)

Definition

The random process (y_t) is called the ARIMA(p,2,q) w.r.t. the white noise process (u_t) , if (y_t) and (Δy_t) are non-stationary, but $\Delta^2 y_t$ is a stationary ARMA(p,q) process w.r.t. the white noise (u_t)

$$\Delta y_t = y_t - y_{t-1}$$
 and $\Delta^2 y_t = \Delta y_t - \Delta y_{t-1}$

Definition

The random process (y_t) is called the ARIMA(p,1,q) w.r.t. the white noise process (u_t) , if (y_t) is non-stationary, but Δy_t is a stationary ARMA(p,q) process w.r.t. the white noise (u_t)

Definition

The random process (y_t) is called the ARIMA(p,2,q) w.r.t. the white noise process (u_t) , if (y_t) and (Δy_t) are non-stationary, but $\Delta^2 y_t$ is a stationary ARMA(p,q) process w.r.t. the white noise (u_t)

$$\Delta y_t = y_t - y_{t-1}$$
 and $\Delta^2 y_t = \Delta y_t - \Delta y_{t-1}$

ARIMA — AutoRegressive Integrated Moving Average

ARIMA(p, 0, q) or ARIMA(p, 1, q) or ARIMA(p, 2, q)

ARIMA(p,0,q) or ARIMA(p,1,q) or ARIMA(p,2,q)

Analyse the graph:

ARIMA(p, 0, q) or ARIMA(p, 1, q) or ARIMA(p, 2, q)

 Analyse the graph: stationary process graph oscillates around its mean with a constant deviation

ARIMA(p, 0, q) or ARIMA(p, 1, q) or ARIMA(p, 2, q)

- Analyse the graph: stationary process graph oscillates around its mean with a constant deviation
- Evaluate all the models and choose the best one by cross-validation

ARIMA(p, 0, q) or ARIMA(p, 1, q) or ARIMA(p, 2, q)

- Analyse the graph: stationary process graph oscillates around its mean with a constant deviation
- Evaluate all the models and choose the best one by cross-validation Time consuming!

ARIMA(p, 0, q) or ARIMA(p, 1, q) or ARIMA(p, 2, q)

- Analyse the graph: stationary process graph oscillates around its mean with a constant deviation
- Evaluate all the models and choose the best one by cross-validation Time consuming!
- You cannot use AIC!

ARIMA(p,0,q) or ARIMA(p,1,q) or ARIMA(p,2,q)

- Analyse the graph: stationary process graph oscillates around its mean with a constant deviation
- Evaluate all the models and choose the best one by cross-validation Time consuming!
- You cannot use AIC!

```
\ln L(y_1, \ldots, y_n \mid \theta) and \ln L(y_2, \ldots, y_n \mid \theta, y_1) and \ln L(y_3, \ldots, y_n \mid \theta, y_1, y_2) incomparable!
```

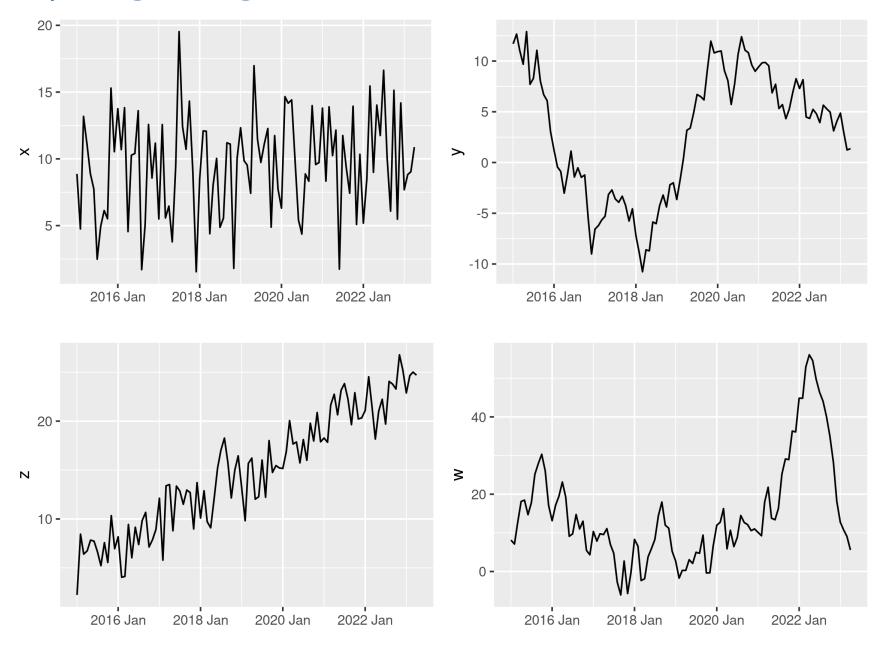
ARIMA(p,0,q) or ARIMA(p,1,q) or ARIMA(p,2,q)

- Analyse the graph: stationary process graph oscillates around its mean with a constant deviation
- Evaluate all the models and choose the best one by cross-validation Time consuming!
- You cannot use AIC! $\ln L(y_1,\ldots,y_n\mid\theta)$ and $\ln L(y_2,\ldots,y_n\mid\theta,y_1)$ and $\ln L(y_3,\ldots,y_n\mid\theta,y_1,y_2)$ incomparable!
- There are unit root tests!

ARIMA(p,0,q) or ARIMA(p,1,q) or ARIMA(p,2,q)

- Analyse the graph: stationary process graph oscillates around its mean with a constant deviation
- Evaluate all the models and choose the best one by cross-validation Time consuming!
- You cannot use AIC! $\ln L(y_1,\ldots,y_n\mid\theta)$ and $\ln L(y_2,\ldots,y_n\mid\theta,y_1)$ and $\ln L(y_3,\ldots,y_n\mid\theta,y_1,y_2)$ incomparable!
- There are unit root tests!
 ADF, KPSS, PP, ...

Analysing the graphs



ARIMA: Summary

• ARMA is for stationary series

ARIMA: Summary

- *ARMA* is for stationary series
- Sometimes Δy_t or $\Delta^2 y_t$ is stationary

ARIMA: Summary

- *ARMA* is for stationary series
- Sometimes Δy_t or $\Delta^2 y_t$ is stationary
- Choose between ARMA and ARIMA

SARIMA process

SARIMA process: Plan

• Seasonal ARMA

SARIMA process: Plan

- Seasonal ARMA
- Seasonal ARIMA

SARIMA process: Plan

- Seasonal ARMA
- Seasonal ARIMA
- Choosing between models

Seasonality and ARIMA

Using ARMA and ARIMA models, we can model seasonality!

Seasonality and ARIMA

Using ARMA and ARIMA models, we can model seasonality!

$$MA(12): y_t = c + u_t + a_1u_{t-1} + a_2u_{t-2} + \dots + a_{12}u_{t-12}.$$

$$ARIMA(12,1,0): \Delta y_t = c + u_t + b_1 \Delta y_{t-1} + \ldots + b_{12} \Delta y_{t-12}.$$

ARMA should be economical!

Let's focus on non-zero coefficients!

ARMA should be economical!

Let's focus on non-zero coefficients!

Definition

If the stationary ARMA model for y_t can be written with fewer parameters as

$$P_{non}(L)P_{seas}(L^{12})y_t = c + Q_{non}(L)Q_{seas}(L^{12})u_t,$$

where the degrees of the lag polynomials are $\deg P_{non} = p$, $\deg P_{seas} = P$, $\deg Q_{non} = q$, $\deg Q_{seas} = Q$, then it is also called SARMA(p,q)(P,Q)[12]

Examples

• SARMA(1,0)(0,2)[12]

$$(1 - b_1 L)y_t = c + (1 + d_1 L^{12} + d_2 L^{24})u_t;$$

Examples

• SARMA(1,0)(0,2)[12]

$$(1 - b_1 L)y_t = c + (1 + d_1 L^{12} + d_2 L^{24})u_t;$$

• SARMA(0, 2)(1, 0)[12]

$$(1 - f_1 L^{12})y_t = c + (1 + a_1 L + a_2 L^2)u_t;$$

Examples

• SARMA(1,0)(0,2)[12]

$$(1 - b_1 L)y_t = c + (1 + d_1 L^{12} + d_2 L^{24})u_t;$$

• SARMA(0, 2)(1, 0)[12]

$$(1 - f_1 L^{12})y_t = c + (1 + a_1 L + a_2 L^2)u_t;$$

• SARMA(1, 2)(2, 1)[12]

$$(1 - f_1 L^{12} - f_2 L^{24})(1 - b_1 L^1)y_t = c + (1 + a_1 L + a_2 L^2)(1 + d_1 L^{12})u_t$$

By analogy with the difference $\Delta y_t = y_t - y_{t-1}$, we can consider the seasonal difference $\Delta_{12}y_t = y_t - y_{t-12}$

By analogy with the difference $\Delta y_t = y_t - y_{t-1}$, we can consider the seasonal difference $\Delta_{12}y_t = y_t - y_{t-12}$

Definition

If the series $z_t = \Delta^d \Delta_{12}^D y_t$ is described by the stationary model SARMA(p,q)(P,Q)[12], then y_t is said to be described by the SARIMA(p,d,q)(P,D,Q)[12] model

By analogy with the difference $\Delta y_t = y_t - y_{t-1}$, we can consider the seasonal difference $\Delta_{12}y_t = y_t - y_{t-12}$

Definition

If the series $z_t = \Delta^d \Delta_{12}^D y_t$ is described by the stationary model SARMA(p,q)(P,Q)[12], then y_t is said to be described by the $SARIMA(p,\mathbf{d},q)(P,\mathbf{D},Q)[12]$ model

d is the number of times the first difference should be taken $\Delta=1-L$;

D is the number of times the seasonal difference should be taken $\Delta_{12}=1-L^{12}$;

By analogy with the difference $\Delta y_t = y_t - y_{t-1}$, we can consider the seasonal difference $\Delta_{12}y_t = y_t - y_{t-12}$

Definition

If the series $z_t = \Delta^d \Delta_{12}^D y_t$ is described by the stationary model SARMA(p,q)(P,Q)[12], then y_t is said to be described by the $SARIMA(p,\mathbf{d},q)(P,\mathbf{D},Q)[12]$ model

d is the number of times the first difference should be taken $\Delta=1-L$;

D is the number of times the seasonal difference should be taken $\Delta_{12}=1-L^{12}$; $y_t\sim SARIMA(0,0,2)(1,1,2)[12]$ means that

$$\Delta_{12}y_t \sim SARMA(0,2)(1,2)[12]$$

SARIMA(p, 0, q)(P, 0, Q) or SARIMA(p, 0, q)(P, 1, Q)[12]?

Analyse the graph!

- Analyse the graph!
- Evaluate all these models and choose the best one by cross-validation

- Analyse the graph!
- Evaluate all these models and choose the best one by cross-validation Time consuming!

- Analyse the graph!
- Evaluate all these models and choose the best one by cross-validation Time consuming!
- You cannot use AIC!

- Analyse the graph!
- Evaluate all these models and choose the best one by cross-validation Time consuming!
- You cannot use AIC! The conditional and unconditional likelihood functions contain different numbers of terms.

- Analyse the graph!
- Evaluate all these models and choose the best one by cross-validation Time consuming!
- You cannot use AIC! The conditional and unconditional likelihood functions contain different numbers of terms.
- There are unit root tests!

- Analyse the graph!
- Evaluate all these models and choose the best one by cross-validation Time consuming!
- You cannot use AIC! The conditional and unconditional likelihood functions contain different numbers of terms.
- There are unit root tests!
 And rules of thumb...

STL decomposition and the power of seasonality

Step 1. Find the STL expansion of the series (y_t)

$$y_t = trend_t + seas_t + remainder_t$$

STL decomposition and the power of seasonality

Step 1. Find the STL expansion of the series (y_t)

$$y_t = trend_t + seas_t + remainder_t$$

Step 2. Calculate the strength of seasonality

$$F_{seas} = \max \left\{ 1 - \frac{\text{sVar}(remainder)}{\text{sVar}(seas + remainder)}, 0 \right\}$$

STL decomposition and the power of seasonality

Step 1. Find the STL expansion of the series (y_t)

$$y_t = trend_t + seas_t + remainder_t$$

Step 2. Calculate the strength of seasonality

$$F_{seas} = \max \left\{ 1 - \frac{\text{sVar}(remainder)}{\text{sVar}(seas + remainder)}, 0 \right\}$$

Step 3. If the strength of seasonality is above the threshold, then move to $\Delta_{12}y_t=y_t-y_{t-12}$

SARIMA: Summary

Seasonal ARIMA is more compact

SARIMA: Summary

- Seasonal ARIMA is more compact
- The strength of seasonality from the STL expansion is used to decide if a seasonal difference $\Delta_{12}y_t$ is needed

Unit root tests: ADF test

ADF test: Plan

Test assumptions

ADF test: Plan

- Test assumptions
- Test algorithm

ADF test: Plan

- Test assumptions
- Test algorithm
- Three variations of the test

We want to answer the questions:

We want to answer the questions:

• Should the ARMA model be used for (y_t) or for (Δy_t) ?

We want to answer the questions:

- Should the ARMA model be used for (y_t) or for (Δy_t) ?
- How to include a constant in a model?

We want to answer the questions:

- Should the ARMA model be used for (y_t) or for (Δy_t) ?
- How to include a constant in a model?

We want to answer the questions:

- Should the ARMA model be used for (y_t) or for (Δy_t) ?
- How to include a constant in a model?

Name "unit root test":

We want to answer the questions:

- Should the ARMA model be used for (y_t) or for (Δy_t) ?
- How to include a constant in a model?

Name "unit root test":

$$\Delta = 1 - L = P(L)$$

The equation $1 - \ell = 0$ has a root $\ell = 1$

ADF test

ADF — Augmented Dickey Fuller test

ADF test

ADF — Augmented Dickey Fuller test

Three variations of the test: without a constant, with a constant, with a trend

$$\Delta y_t = c + \beta y_{t-1} + d_1 \Delta y_{t-1} + \dots + d_p \Delta y_{t-p} + u_t,$$

$$\Delta y_t = c + \beta y_{t-1} + d_1 \Delta y_{t-1} + \dots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

$$\Delta y_t = m + x_t$$
;

 (x_t) is a stationary AR(p) process with $\mathbb{E}(x_t)=0$;

$$y_t = y_0 + mt + \sum_{i=1}^{t} x_i$$

$$\Delta y_t = c + \beta y_{t-1} + d_1 \Delta y_{t-1} + \dots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

$$\Delta y_t = m + x_t$$
;

 (x_t) is a stationary AR(p) process with $\mathbb{E}(x_t) = 0$;

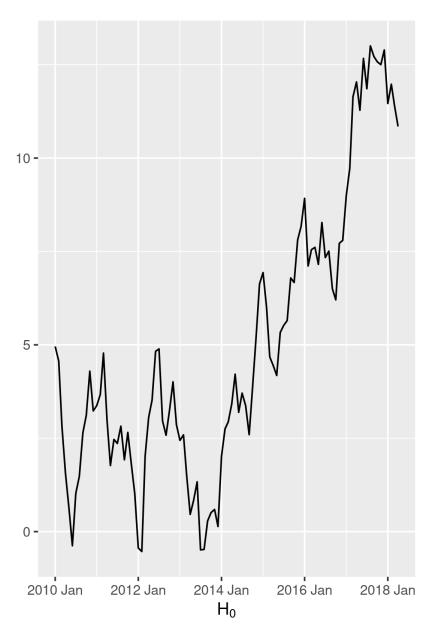
$$y_t = y_0 + mt + \sum_{i=1}^{t} x_i$$

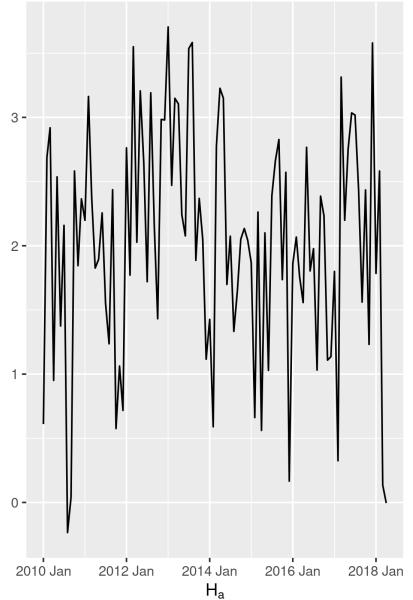
$$H_a$$
: $\beta < 0$

 (y_t) is a stationary AR(p+1) process

ADF with constant: H_0 and H_a

ADF with constatnt





Step 1. Evaluate regression

$$\widehat{\Delta y_t} = \hat{c} + \hat{\beta} y_{t-1} + \hat{d}_1 \Delta y_{t-1} + \ldots + \hat{d}_p \Delta y_{t-p}$$

Step 1. Evaluate regression

$$\widehat{\Delta y_t} = \hat{c} + \hat{\beta} y_{t-1} + \hat{d}_1 \Delta y_{t-1} + \ldots + \hat{d}_p \Delta y_{t-p}$$

Step 2. Calculate the *t*-statistics using the classic formula

$$ADF = \frac{\hat{\beta} - 0}{se(\hat{\beta})}$$

Step 1. Evaluate regression

$$\widehat{\Delta y_t} = \hat{c} + \hat{\beta} y_{t-1} + \hat{d}_1 \Delta y_{t-1} + \ldots + \hat{d}_p \Delta y_{t-p}$$

Step 2. Calculate the *t*-statistics using the classic formula

$$ADF = \frac{\hat{\beta} - 0}{se(\hat{\beta})}$$

Under true H_0 , the distribution of the ADF-statistic converges to the special DF distribution with DF^c !

Step 1. Evaluate regression

$$\widehat{\Delta y_t} = \hat{c} + \hat{\beta} y_{t-1} + \hat{d}_1 \Delta y_{t-1} + \ldots + \hat{d}_p \Delta y_{t-p}$$

Step 2. Calculate the *t*-statistics using the classic formula

$$ADF = \frac{\hat{\beta} - 0}{se(\hat{\beta})}$$

Under true H_0 , the distribution of the ADF-statistic converges to the special DF distribution with DF^c !

Step 3. We conclude:

If $ADF < DF^c$ then H_0 is rejected

$$\Delta y_t = \beta y_{t-1} + d_1 \Delta y_{t-1} + \dots + d_p \Delta y_{t-p} + u_t,$$

$$\Delta y_t = \beta y_{t-1} + d_1 \Delta y_{t-1} + \dots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

 (Δy_t) is a stationary AR(p) process with $\mathbb{E}(\Delta y_t) = 0$;

$$y_t = y_0 + \sum_{i=1}^t \Delta y_i$$

$$\Delta y_t = \beta y_{t-1} + d_1 \Delta y_{t-1} + \ldots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

 (Δy_t) is a stationary AR(p) process with $\mathbb{E}(\Delta y_t)=0$;

$$y_t = y_0 + \sum_{i=1}^t \Delta y_i$$

$$H_a$$
: $\beta < 0$

 (y_t) is a stationary AR(p+1) process with $\mathbb{E}(y_t)=0$;

$$\Delta y_t = \beta y_{t-1} + d_1 \Delta y_{t-1} + \ldots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

 (Δy_t) is a stationary AR(p) process with $\mathbb{E}(\Delta y_t) = 0$;

$$y_t = y_0 + \sum_{i=1}^t \Delta y_i$$

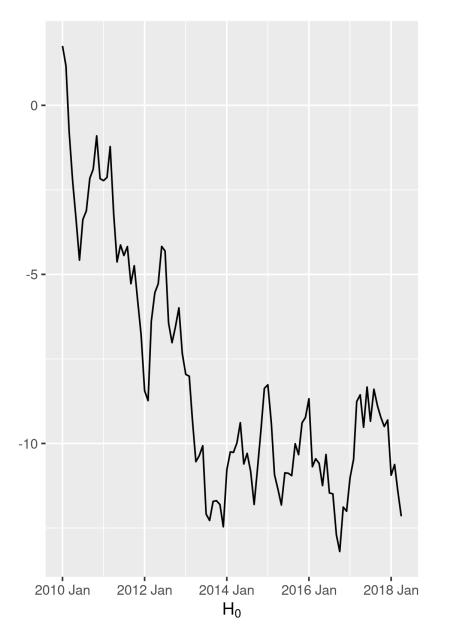
$$H_a$$
: $\beta < 0$

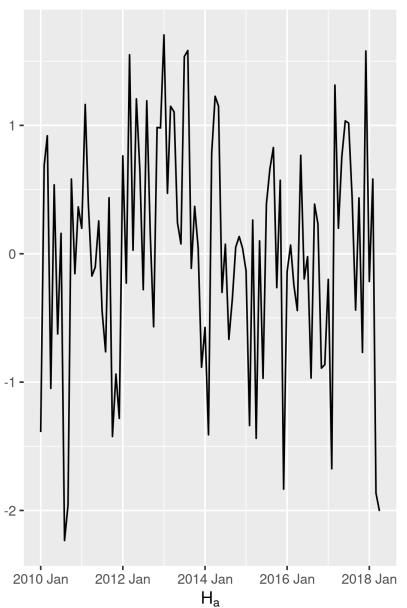
 (y_t) is a stationary AR(p+1) process with $\mathbb{E}(y_t)=0$;

The algorithm will have regression without a constant and another distribution DF^0

ADF without constant: H_0 and H_a

ADF without constatnt





$$\Delta y_t = c + gt + \beta y_{t-1} + d_1 \Delta y_{t-1} + \dots + d_p \Delta y_{t-p} + u_t,$$

$$\Delta y_t = c + gt + \beta y_{t-1} + d_1 \Delta y_{t-1} + \dots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

$$\Delta y_t = k_1 + k_2 t + x_t;$$

 (x_t) is a stationary AR(p) process with $\mathbb{E}(x_t)=0$;

$$y_t = y_0 + m_1 t + m_2 t^2 + \sum_{i=1}^t x_i$$

$$\Delta y_t = c + gt + \beta y_{t-1} + d_1 \Delta y_{t-1} + \ldots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

$$\Delta y_t = k_1 + k_2 t + x_t;$$

$$(x_t) \text{ is a stationary } AR(p) \text{ process with } \mathbb{E}(x_t) = 0;$$

$$y_t = y_0 + m_1 t + m_2 t^2 + \sum_{i=1}^t x_i$$

$$H_a: \beta < 0$$

$$y_t = m_1 + m_2 t + x_t;$$

$$(x_t) \text{ is a stationary } AR(p+1) \text{ process with } \mathbb{E}(x_t) = 0;$$

$$\Delta y_t = c + gt + \beta y_{t-1} + d_1 \Delta y_{t-1} + \ldots + d_p \Delta y_{t-p} + u_t,$$

$$H_0: \beta = 0$$

$$\Delta y_t = k_1 + k_2 t + x_t;$$

$$(x_t) \text{ is a stationary } AR(p) \text{ process with } \mathbb{E}(x_t) = 0;$$

$$y_t = y_0 + m_1 t + m_2 t^2 + \sum_{i=1}^t x_i$$

$$H_a: \beta < 0$$

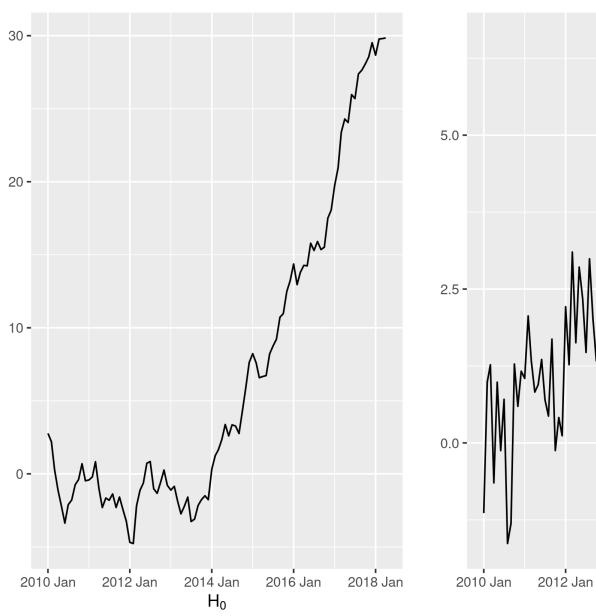
$$y_t = m_1 + m_2 t + x_t;$$

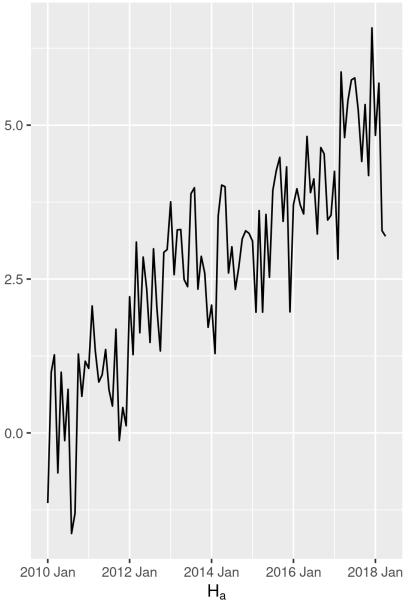
$$(x_t) \text{ is a stationary } AR(p+1) \text{ process with } \mathbb{E}(x_t) = 0;$$

The algorithm will have a regression with a constant and a trend and another distribution DF^{ct}

ADF with trend: H_0 and H_a

ADF with trend





ADF test: Summary

- Applicable for making a decision about the transition to Δy_t

ADF test: Summary

- Applicable for making a decision about the transition to Δy_t
- Three variants of the ADF test with different assumptions

Unit root tests: KPSS test

KPSS test: Plan

• Long-term variance

KPSS test: Plan

- Long-term variance
- Prerequisites for the test

KPSS test: Plan

- Long-term variance
- Prerequisites for the test
- Two variations of the test

KPSS test

KPSS — Kwiatkowski–Phillips–Schmidt–Shin test

KPSS test

KPSS — Kwiatkowski–Phillips–Schmidt–Shin test

Two variations of the test: with a constant, with a trend

Long-term variance

Definition

For a stationary process (y_t) , the quantity λ^2 is called long-term variance if

$$Var(\bar{y}) = \frac{\lambda^2}{T} + o(1/T)$$

or

$$\lim_{T \to \infty} T \operatorname{Var}(\bar{y}) = \lambda^2,$$

where $\bar{y} = (y_1 + ... + y_T)/T$.

Long-term variance

Definition

For a stationary process (y_t) , the quantity λ^2 is called long-term variance if

$$Var(\bar{y}) = \frac{\lambda^2}{T} + o(1/T)$$

or

$$\lim_{T \to \infty} T \operatorname{Var}(\bar{y}) = \lambda^2,$$

where $\bar{y} = (y_1 + ... + y_T)/T$.

Motivation

For independent observations with the constant variance

$$\operatorname{Var}(\bar{y}) = \frac{\sigma^2}{T}$$
, where $\sigma^2 = \operatorname{Var}(y_i)$

KPSS with constant

$$y_t = c + rw_t + x_t,$$

KPSS with constant

$$y_t = c + rw_t + x_t,$$

 $H_0: rw_t = 0$

 (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;

KPSS with constant

$$y_t = c + rw_t + x_t,$$

$$H_0: rw_t = 0$$

 (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;

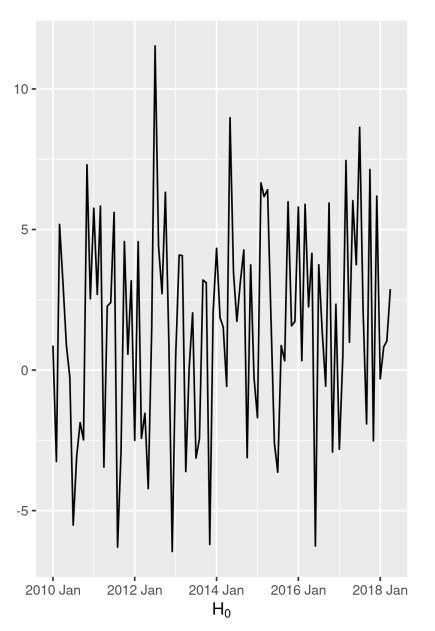
$$H_a$$
: $rw_t = rw_{t-1} + u_t$

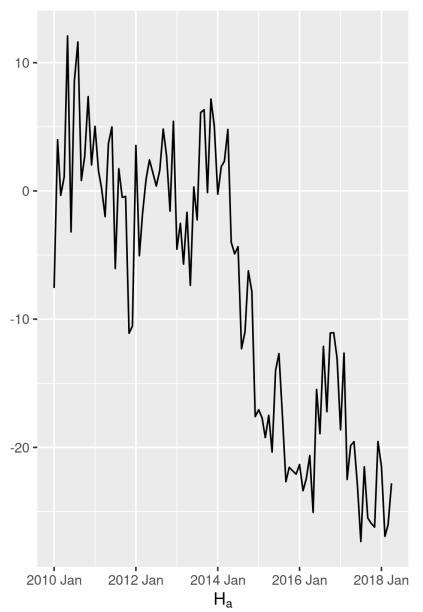
$$rw_0 = 0;$$

- (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;
- (u_t) is white noise independent of (x_t)

KPSS with constant: H_0 and H_a

KPSS with constant





Step 1. Evaluate regression on a constant

$$\widehat{y}_t = \hat{c}$$

Step 1. Evaluate regression on a constant

$$\widehat{y_t} = \hat{c}$$

Step 2. Calculate KPSS statistics

$$KPSS = \frac{\sum_{t=1}^{T} S_t^2}{T^2 \hat{\lambda}^2},$$

where S_t is the accumulated sum of residuals, $S_t = \hat{u}_1 + \ldots + \hat{u}_t$, and $\hat{\lambda}^2$ is a consistent estimator of the long-term variance.

Step 1. Evaluate regression on a constant

$$\hat{y_t} = \hat{c}$$

Step 2. Calculate KPSS statistics

$$KPSS = \frac{\sum_{t=1}^{T} S_t^2}{T^2 \hat{\lambda}^2},$$

where S_t is the accumulated sum of residuals, $S_t = \hat{u}_1 + \ldots + \hat{u}_t$, and $\hat{\lambda}^2$ is a consistent estimator of the long-term variance.

Under true H_0 , the distribution of the KPSS-statistic converges to a special distribution with $KPSS^c$!

Step 1. Evaluate regression on a constant

$$\hat{y_t} = \hat{c}$$

Step 2. Calculate KPSS statistics

$$KPSS = \frac{\sum_{t=1}^{T} S_t^2}{T^2 \hat{\lambda}^2},$$

where S_t is the accumulated sum of residuals, $S_t = \hat{u}_1 + \ldots + \hat{u}_t$, and $\hat{\lambda}^2$ is a consistent estimator of the long-term variance.

Under true H_0 , the distribution of the KPSS-statistic converges to a special distribution with $KPSS^c$!

Step 3. We conclude:

If $KPSS > KPSS^c$ then H_0 is rejected

$$y_t = c + bt + rw_t + x_t,$$

$$y_t = c + bt + rw_t + x_t,$$

 $H_0: rw_t = 0$

 (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;

$$y_t = c + bt + rw_t + x_t,$$

$$H_0$$
: $rw_t = 0$

 (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;

$$H_a$$
: $rw_t = rw_{t-1} + u_t$

$$rw_0 = 0;$$

 (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;

 (u_t) is white noise independent of (x_t)

$$y_t = c + bt + rw_t + x_t,$$

 $H_0: rw_t = 0$

 (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;

 H_a : $rw_t = rw_{t-1} + u_t$

 $rw_0 = 0;$

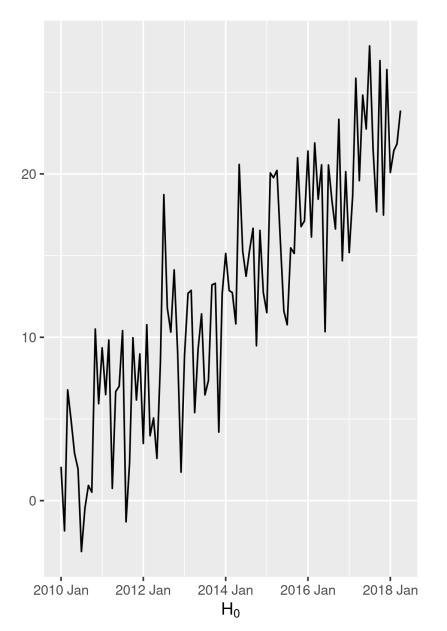
 (x_t) is a stationary process with $\mathbb{E}(x_t) = 0$;

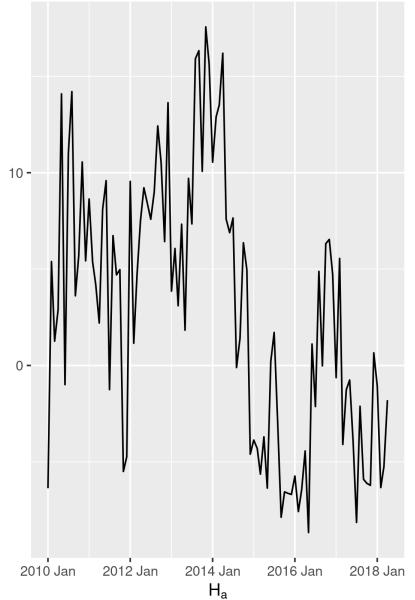
 (u_t) is white noise independent of (x_t)

The first step of the algorithm will have a regression on a constant and a trend and the statistic under null will have another special distribution $KPSS^{ct}$

KPSS with trend: H_0 and H_a

KPSS with trend





$$A. \quad y_t = a + bt + x_t$$

 (y_t) — trend stationary (stationary around the trend)

 (x_t) — a stationary process with $\mathbb{E}(x_t) = 0$

$$A. \quad y_t = a + bt + x_t$$

 (y_t) — trend stationary (stationary around the trend)

 (x_t) — a stationary process with $\mathbb{E}(x_t) = 0$

Recipe: Estimate regression a + bt with ARMA errors for (y_t) .

$$A. \quad y_t = a + bt + x_t$$

- (y_t) trend stationary (stationary around the trend)
- (x_t) a stationary process with $\mathbb{E}(x_t) = 0$

Recipe: Estimate regression a + bt with ARMA errors for (y_t) .

B.
$$y_t = a + \sum_{i=1}^t x_i \text{ or } y_t = a + bt + \sum_{i=1}^t x_i$$

- (x_t) a stationary process with $\mathbb{E}(x_t) = 0$
- (y_t) difference stationary (stationary in differences)

$$A. \quad y_t = a + bt + x_t$$

- (y_t) trend stationary (stationary around the trend)
- (x_t) a stationary process with $\mathbb{E}(x_t) = 0$

Recipe: Estimate regression a + bt with ARMA errors for (y_t) .

B.
$$y_t = a + \sum_{i=1}^t x_i \text{ or } y_t = a + bt + \sum_{i=1}^t x_i$$

- (x_t) a stationary process with $\mathbb{E}(x_t) = 0$
- (y_t) difference stationary (stationary in differences)

Recipe: evaluate ARMA for (Δy_t) .

$$A. \quad y_t = a + bt + x_t$$

 (y_t) — trend stationary (stationary around the trend)

 (x_t) — a stationary process with $\mathbb{E}(x_t) = 0$

Recipe: Estimate regression a + bt with ARMA errors for (y_t) .

B.
$$y_t = a + \sum_{i=1}^t x_i \text{ or } y_t = a + bt + \sum_{i=1}^t x_i$$

 (x_t) — a stationary process with $\mathbb{E}(x_t) = 0$

 (y_t) — difference stationary (stationary in differences)

Recipe: evaluate ARMA for (Δy_t) .

Both (y_t) are non-stationary!

KPSS test: Summary

• Applicable for making a decision about the transition to Δy_t

KPSS test: Summary

- Applicable for making a decision about the transition to Δy_t
- Two versions of the KPSS test with different assumptions