

w04 m-logit - laba - pola

multinomial logit.

exp : 2 alternatives.

$y_i = 0$ [base] $y_i = 1$

$$P(y_i = 1 | x_i) = \frac{\exp(\beta_0 + \beta_1 x_i)}{1 + \exp(\beta_0 + \beta_1 x_i)} = \frac{\exp(x_i^T \cdot \beta)}{1 + \exp(x_i^T \cdot \beta)}$$

View:
"score"

$$\text{utility}_1(x_i) = x_i^T \cdot \beta^{(1)}$$

$$\text{utility}_0(x_i) = x_i^T \cdot \beta^{(0)}$$

vector

$$\frac{\exp(0)}{\exp(x_i^T \cdot 0)} = 1$$

$$\beta^{(0)} = \begin{pmatrix} \beta_0^{(0)} \\ \vdots \\ \beta_k^{(0)} \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}$$

$$P(y_i = 1 | x_i) = \frac{\exp(\text{utility}_1)}{\exp(\text{utility}_0) + \exp(\text{utility}_1)}$$

$$P(y_i = 0 | x_i) = \frac{\exp(\text{utility}_0)}{\sum_{j=0} \exp(\text{utility}_j)}$$

a alternatives:

$$[\text{base}] \text{utility}_0 = x_i^T \cdot \beta^{(0)} = 0$$

$$\text{utility}_1 = x_i^T \cdot \beta^{(1)}$$

$$\text{utility}_2 = x_i^T \cdot \beta^{(2)}$$

\vdots

$$\text{utility}_{a-1} = x_i^T \cdot \beta^{(a-1)}$$

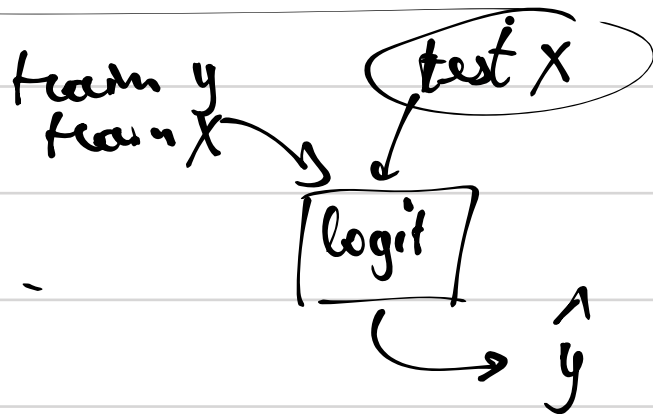
$$P(y_i = j | x_i) = \frac{\exp(\text{utility}_j)}{\sum_{s=0}^{a-1} \exp(\text{utility}_s)}$$

mult. logit

LDA - QDA

linear discriminant analysis LDA
quadratic // QDA

⑦ generative models.



you can predict y given x

LDA, QDA, MDA...

- * you can predict y given x
- * can generate new artificial data (both x and y)

LDA/
QDA

$y_i \in \{1, 2, 3 \dots K\}$ K alternatives

$$P(y_i = k) = \pi_k$$

$$x_i \in \mathbb{R}$$

$$(x_i | y_i = k) \sim N(\mu_k, \Sigma_k)$$

LDA: each class has its own μ_k
but $\Sigma_k = \Sigma$

QDA: each class has its own μ_k, Σ_k

ex.

$n = 100$ (obs)

$y_i \in \{1, 2, 3\}$

$$x_i = \begin{pmatrix} a_i \\ b_i \end{pmatrix} \in \mathbb{R}^2$$

how many par-s does LDA model

$$2 + 3 + 3 \cdot 2 = 11 \text{ par.}$$

a)

b)

QDA model

$$\mu_1 = \begin{pmatrix} \mu_{11} \\ \mu_{12} \end{pmatrix}$$

$$\mu_3 = \begin{pmatrix} \mu_{31} \\ \mu_{32} \end{pmatrix}$$

$$\Sigma = \begin{pmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix}$$

$$\pi_1 + \pi_2 + \pi_3 = 1$$

3 par-s in Σ

$$\begin{aligned} (x_i | y_i = 1) &\sim N(\mu_1; \Sigma) \\ (x_i | y_i = 2) &\sim N(\mu_2; \Sigma) \end{aligned}$$

17p

ex. Why is it called "linear"

$\hat{\mu}_1 = \begin{pmatrix} 2 \\ 1 \end{pmatrix}$
 $\hat{\mu}_2 = \begin{pmatrix} -1 \\ 2 \end{pmatrix}$
 $\hat{\mu}_3 = \begin{pmatrix} 0 \\ 1 \end{pmatrix}$

$\hat{\Sigma} = \begin{pmatrix} 4 & -1 \\ -1 & 9 \end{pmatrix}$
 $\hat{\pi}_1 = 0,2$
 $\hat{\pi}_2 = 0,3$

draw a line $P(y_i = 1 | x) = P(y_i = 3 | x)$

$n=100$
 $x_i = \begin{pmatrix} a_i \\ b_i \end{pmatrix} \in \mathbb{R}^2$
 $y_i \in \{1, 2, 3\}$

$$P(y_i = 1 | x_i) = \frac{P(y_i = 1, x_i)}{f(x_i)} = \frac{P(y_i = 1) \cdot f(x_i | y_i = 1)}{\sum_{k=1}^3 P(y_i = k) \cdot f(x_i | y_i = k)}$$

$y_i \in \{1, 2, 3\}$
 $x_i \in \mathbb{R}^2$

$$P(y_i = 3 | x_i) = \dots = \frac{P(y_i = 3) \cdot f(x_i | y_i = 3)}{\sum_{k=1}^3 P(y_i = k) \cdot f(x_i | y_i = k)}$$

$$P(y_i = 1) \cdot f(x_i | y_i = 1) = P(y_i = 3) \cdot f(x_i | y_i = 3)$$

$$\ln \pi_1 + \ln f(x_i | y_i = 1) = \ln \pi_3 + \ln f(x_i | y_i = 3)$$

$$\ln 0,2 + \ln f_1 = \ln 0,5 + \ln f_3$$

$$(x_i | y_i = k) \sim \mathcal{N}(\mu_k; \Sigma)$$

$$f(x_i | y_i = k) = \frac{1}{\sqrt{(2\pi)^d \det \Sigma}} \cdot \exp\left(-\frac{1}{2}(x_i - \mu_k)^T \cdot \Sigma^{-1} \cdot (x_i - \mu_k)\right)$$

d - dimensions of x_i
 $d=2$

$$\ln 0,2 - \frac{1}{2}(x - \mu_1)^T \cdot \Sigma^{-1} \cdot (x - \mu_1) = \ln 0,5 - \frac{1}{2}(x - \mu_3)^T \cdot \Sigma^{-1} \cdot (x - \mu_3)$$

$$-\frac{1}{2}x^T \cdot \Sigma^{-1} \cdot x$$

$$-\frac{1}{2}x^T \cdot \Sigma^{-1} \cdot x$$

$$\ln 0.2 - \frac{1}{2} (x - \mu_1)^T \Sigma^{-1} (x - \mu_1) = \ln 0.5 - \frac{1}{2} (x - \mu_3)^T \Sigma^{-1} (x - \mu_3)$$

$\Sigma^{-1} = \mathcal{J}$ "precision" matrix
(inverse of cov)

$$2 \ln 0.2 + \underbrace{x^T \cdot \mathcal{J} \cdot \mu_1}_{\text{red}} + \underbrace{\mu_1^T \cdot \mathcal{J} \cdot x}_{\text{red}} - \mu_1^T \cdot \mathcal{J} \cdot \mu_1 =$$

$$= 2 \ln 0.5 + \underbrace{x^T \cdot \mathcal{J} \cdot \mu_3}_{\text{red}} + \underbrace{\mu_3^T \cdot \mathcal{J} \cdot x}_{\text{red}} - \mu_3^T \cdot \mathcal{J} \cdot \mu_3$$

$$x^T \cdot \mathcal{J} \cdot \mu_1 = \mu_1^T \mathcal{J} \cdot x \in \mathbb{R}^1 \text{ [scalar]}$$

$$(x_1 \dots x_d) \cdot \begin{bmatrix} \mathcal{J}^1 \\ \vdots \\ \mathcal{J}^d \end{bmatrix} \cdot \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_d \end{pmatrix} = (\mu_1 \dots \mu_d) \begin{bmatrix} \mathcal{J}^1 \\ \vdots \\ \mathcal{J}^d \end{bmatrix} \cdot \begin{pmatrix} x_1 \\ \vdots \\ x_d \end{pmatrix}$$

$\mathcal{J}^T = \mathcal{J}$ (symmetric)

$$\boxed{2(\mu_1^T - \mu_3^T) \cdot \mathcal{J}^1} x = \underbrace{2 \ln 0.5 - 2 \ln 0.2 + \mu_1^T \cdot \mathcal{J} \cdot \mu_1 - \mu_3^T \cdot \mathcal{J} \cdot \mu_3}_{\text{const}}$$

$$x = \begin{pmatrix} a \\ b \end{pmatrix}$$

$$2 \cdot \left(\begin{pmatrix} 2 \\ 1 \end{pmatrix} - \begin{pmatrix} 0 \\ 1 \end{pmatrix} \right)^T \cdot \begin{pmatrix} 4 & -1 \\ -1 & 9 \end{pmatrix}^{-1} = 2 \cdot (2 \ 0) \cdot \frac{1}{35} \cdot \begin{pmatrix} 9 & 1 \\ 1 & 4 \end{pmatrix} =$$

$$= \frac{2}{35} (18 \ 2) = \left(\frac{36}{35} \ \frac{4}{35} \right)$$

$$\left(\frac{36}{35} \ \frac{4}{35} \right) \cdot \begin{pmatrix} a \\ b \end{pmatrix} = \text{const}$$

$$\frac{36}{35} a + \frac{4}{35} b = \text{const.}$$

$$x = \begin{pmatrix} a \\ b \end{pmatrix}$$

$$p(y_i=1|x) > p(y_i=3|x)$$

$$p(y_i=3|x) > p(y_i=1|x)$$

$> a$

exer $n=3$ discrete $y_i \in \{1, 2\}$ cont $x_i \in \mathbb{R}^1$

i	y_i	x_i
1	1	3
2	2	2
3	2	5

estimate LDA model using max. likelihood.

$$P(y_i=1) \cdot f(x_i=3 | y_i=1)$$

sel. n

$$\text{likelihood} = P(y_1=1, x_1=3) \cdot P(y_2=2, x_2=2) \cdot P(y_3=2, x_3=5)$$

not honest probabilities but a mix prob/density.

$$\ln L = \ln \pi_1 + \ln f(x_1=3 | y_1=1) + \ln \pi_2 + \ln f(x_2=2 | y_2=2) + \ln \pi_2 + \ln f(x_3=5 | y_3=2)$$

depends only on π_1

depends only on μ_1, μ_2, σ^2

$$\rightarrow \max_{\pi_1, \mu_1, \mu_2, \sigma^2}$$

$$\begin{aligned} \pi_1 + \pi_2 &= 1 \quad (1 \text{ free par.}) \\ (x|y=1) &\sim \mathcal{N}(\mu_1; \sigma^2) \\ (x|y=2) &\sim \mathcal{N}(\mu_2; \sigma^2) \end{aligned} \quad \left. \vphantom{\begin{aligned} \pi_1 + \pi_2 &= 1 \\ (x|y=1) &\sim \mathcal{N}(\mu_1; \sigma^2) \\ (x|y=2) &\sim \mathcal{N}(\mu_2; \sigma^2) \end{aligned}} \right\} 3 \text{ free par. - s: } \mu_1, \mu_2, \sigma^2$$

$$\max_{\pi_1} \ln \pi_1 + \ln(1 - \pi_1) + \ln(1 - \pi_1)$$

intuit: $\hat{\pi}_1 = \frac{1}{3}$ (1 obs n from class 1 and n=3)

$$\frac{1}{\hat{\pi}_1} - \frac{1}{1 - \hat{\pi}_1} \cdot 2 = 0$$

$$1 - \hat{\pi}_1 = 2\hat{\pi}_1$$

$$3\hat{\pi}_1 = 1$$

$$\hat{\pi}_1 = \frac{1}{3}$$

part II of the log-likelihood function

$$\ln f(x=3|y=1) + \\ + \ln f(x=2|y=2) + \\ + \ln f(x=5|y=2)$$

i	y _i	x _i
1	1	3
2	2	2
3	2	5

$$\ln f(x|y=k) = \ln \left[\frac{1}{\sqrt{2\pi}\sigma^2} \cdot \exp\left(-\frac{(x-\mu_k)^2}{2\sigma^2}\right) \right] = \\ = \underbrace{-\frac{1}{2} \ln(2\pi)}_{\text{part I}} - \underbrace{\frac{1}{2} \cdot \ln(\sigma^2)}_{\text{part I}} - \underbrace{\frac{1}{2} \frac{(x-\mu_k)^2}{\sigma^2}}_{\text{part II}}$$

part II of $\ln L = -\frac{1}{2} \ln(2\pi) \times 3 - \frac{1}{2} \ln(\sigma^2) \times 3$

$$- \frac{1}{2\sigma^2} \left[(3-\mu_1)^2 + (2-\mu_2)^2 + (5-\mu_2)^2 \right] \rightarrow \text{min}$$

$$\min_{\mu_1, \mu_2} (3-\mu_1)^2 + (2-\mu_2)^2 + (5-\mu_2)^2 = 0^2 + 1.5^2 + 1.5^2 =$$

$$\hat{\mu}_1 = 3$$

$$\hat{\mu}_2 = \frac{5+2}{2} = 3.5$$

$$= 2 \cdot 2.25 = 4.5$$

$$\max_{\sigma^2} -\frac{3}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \cdot [4.5] = h(\sigma^2)$$

$$h'(\sigma^2) = -\frac{3}{2} \cdot \frac{1}{\sigma^2} + \frac{4.5}{2} \cdot \frac{1}{(\sigma^2)^2}$$

$$3\hat{\sigma}^2 = 4.5$$

$$\hat{\sigma}^2 = \frac{4.5}{3}$$

//

MDA = mixture discrimin. analysis.

$$(x_i | y_i = k) \sim \alpha_1 \cdot \mathcal{N}(\mu_{k1}, \Sigma_{k1}) + \alpha_2 \cdot \mathcal{N}(\mu_{k2}, \Sigma_{k2})$$

mixture of normal distributions.