Université de Technologie de Compiegne

SY19

MACHINE LEARNING

First Assignment

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Abstract

This report is an assignment for the course SY19: Machine Learning. It is focused on two exercises: a regression problem (Chapter I) and classification problem (Chapter II). In each part, we will describe the methodology we used to build the best model for each dataset.

The content presented in this report is the result of the theoretical and practical courses of SY19 taught by Thierry Denoeux and a literature review. We put on a side note that we both have not followed SY09, and only one of us has already done SY02, which prevented us to go deeper in the explanations and analysis sometimes but we made sure to make researches every time it was needed.

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0.1 Abbreviations

MSE	Mean Squared Error
RSE	Resiual Standard Error
RSS	Residual Sum of Squares
AIC	Akaike information criterion
BIC	Bayesian information criterion
R^2	Adjusted R^2
LR	Linear Regression
CV	Cross Validation
LOOCV	Leave One Out Cross Validation

Chapter 1

Breast Cancer Recurring Time

1.1 Context

This part aims to build the best model to predict the recurring time of breast cancer based on about 30 features computed from a breast mass. This regression problem will take advantage of a given dataset describing about 200 patient cases.

1.2 Dataset Description

The very first step of our method consists in taking a look at the raw dataset to get precious hints on how each feature contributes to the recurring time. The dataset comprises 194 patient cases, each of which is described through 32 features and the cancer recurring time Time that we have to predict.

1.2.1 Time

Let's first describe the distribution of the variable Time. To do so, we can use the R functions boxplot (figure 1.1) and hist (figure 1.2). According to these figures, our dataset mostly represents short reappearing times (lower than 40), and there are very few patients whose variable Time is higher than 80. However, we do not know if this distribution is also representative of the whole population. If this is is the case, our model should be able to work on other datasets, if not, our model will be biaised.

1.2.2 Features Description

Each patient is represented with a set of 32 features extracted and computed from a digitized image of a breast mass. The data description we were given does not specify the units, but we do not need them for the following analysis. Here are the 32 features we are provided with:

- Lymph Node Status
- Mean, Standard Error and Mean of the three largest values (also called "Worst") of
 - Radius
 - Texture
 - Perimeter
 - Area
 - Smoothness
 - Compactness
 - Concavity
 - Concave points
 - Symmetry
 - Fractal dimension
- Tumor Size

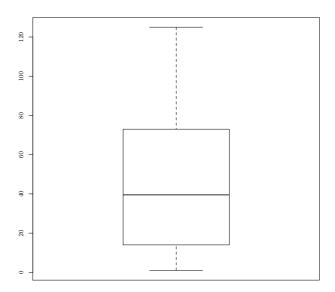


Figure 1.1: Box Plot

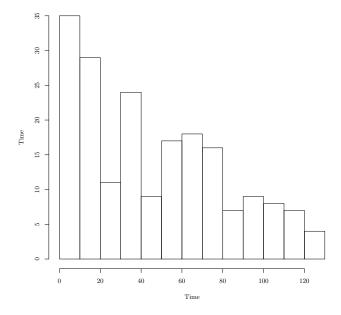


Figure 1.2: Histogram

Feature Correlation

Based on the definition of the parameters described above, we already know that many features are correlated. For instance :

- The mean of each parameters is smaller than the "worst" value;
- The radius, the perimeter and the area are most likely to be linked together;
- The compactness can be computed with the perimeter and the area thanks to the given formula : $Compactness = \frac{perimeter^2}{area-1}$

These dependent features might be a cause of model low performance.

1.2.3 Data Relevance

We should first check that every patient is relevant to our study, in other words, that there is no abnormal observation in the dataset. Cook's Distance is an interesting measure to verify this important criteria, it can be computed after a simple Linear Regression.

Cook's distance aims to study the influence of each observation on the regression coefficient estimates. To do so, this method uses a straight-forward

approach that consists in computing the difference between the original coefficient estimates $\hat{\beta}$ and the coefficient estimates without taking into account the *i*-th observation $\hat{\beta}_{(-i)}$. The difference is then normalized using the number of parameters and the standard deviation estimate. A value higher than 1 often indicates an outlier that should be removed from the dataset.

In R, we can use the following code to compute and plot the Cook's distance of each observation :

```
linreg = lm(Time ~ ., data=data_set)
cooks.distance(linreg)
```

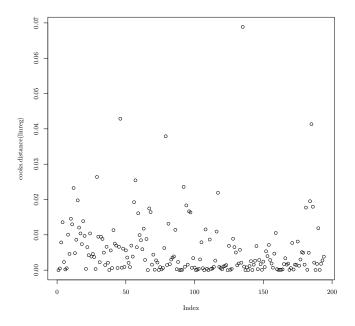


Figure 1.3: Cook's Distance

According to this plot (figure 1.3), no observation is located beyond the critical Cook's boundary of 1. This means that we can potentially use each and every patient case of our dataset to build our regression model.

1.2.4 Relation between "feature" and "time"

In this section, we will take a first look at the relationship between the variable Time and the features used to describe a patient case.

A first way to do it is to separately plot Time against each feature. An example of such plot is shown in figure 1.4. Unfortunately, most plots picture

very scattered points that do not seem to follow any specific model. The variance is so significant that we cannot even estimate the type of function that links a feature and the variable Time together. It could be a simple linear model with a high variance that we may be able to estimate, or more complex models with non-linearities and feature correlations.

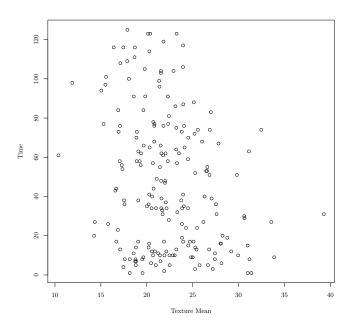


Figure 1.4: Plot of Time against Texture Mean

To have better clues on the type of model we should be dealing with, we can draw a QQ-Plot that plots the Studentized Residuals against the quantiles.

In R, we can use the library car to easily draw the QQ-Plot (figure 1.5):

```
library(car)
qqPlot(linreg, main="QQ Plot")
```

The bottom tail of the QQ-Plot seems to deviate from the linear line, which is a sign of the error's non-normality. This may mean that the error does not follow a normal model, or that the model is actually non-linear.

We can also analyze the Residuals-Fitted plot and the Scale-Location plot to get a better understanding on the model. To do, we can simply apply the function plot on the linear model (figure 1.6). It turns out that both plots show non-normal scatters of the residuals. In particular, the points displayed in the Residuals-Fitted plot seem to follow a fan shape. This is a sign of a non-constant variance, also called heteroscedasticity.

1.3 Measures to Compare Models

Before building any model, we have to properly define the measures we will use later to compare their performance.

1.3.1 Some Measures

A first way to assess the performance of a model F is to compute the Mean Squared Error (MSE). Let \hat{Time} be the estimate of the variable Time using the model F. MSE is defined as:

$$MSE(F) = mean_i(Time_i - Time_i)^2$$

We can also use adjusted \mathbb{R}^2 score.

1.3.2 Data Split

These measures should not be applied on a set whose data was also used to train the model. Indeed, this would include a biais that might distort our conclusions. To cope with this problem, we have to split the dataset into two disjointed sets .

• Training Set: About 66% of the dataset dedicated to the model fitting;

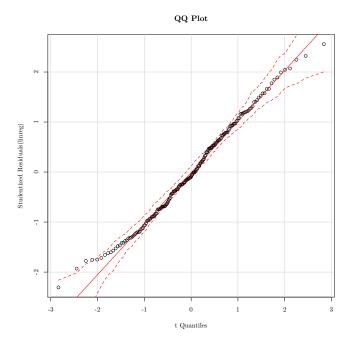


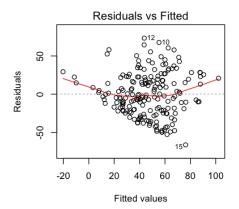
Figure 1.5: QQ-Plot

• Test Set: The remaining 34% only used at the end to provide some kind of objective measure of the model performance.

```
n = dim(data_set)[1]
train_id = sample(1:n, n * 2/3)

train_set = data_set[train_id,]
train_set.x = train_set[,-33]
train_set.y = train_set[,33]

test_set = data_set[-train_id,]
test_set.x = test_set[,-33]
test_set.y = test_set[,33]
```



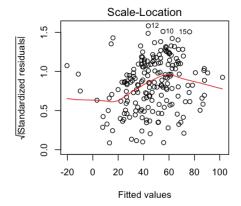


Figure 1.6

Once this data split is done, we can finally dive into the model building.

1.4 K-nearest neighbors (KNN)

We start our analysis with a very simple model called the KNN. Given a positive integer k and a test observation. The KNN model first identifies the k closest points to each point of the test observation then it estimates the response using the average of the k closest training responses.

1.4.1 Knn Model

The KNN model in R is done by calling the function reg of the package knn. As we will see in the following sections, For most prediction algorithms, we first have to build the prediction model on the training data, and then use the model to test our predictions. However, the KNN function does both in a single step. In order to find the best k we set a maximum number of neighbors to be considered (in our model it is 120), then we calculate the MSE for each k which is the mean of the squared difference between the real value of Time and the predicted one. All the steps are detailed in the code below.

Model Implementation

The graph below (figure 1.7) shows the \mathbf{MSE} plotted against the values of k in a range from 0 to 120. Graphically we notice that a minimum is reached between 10 and 20.

MSE against k neighbours

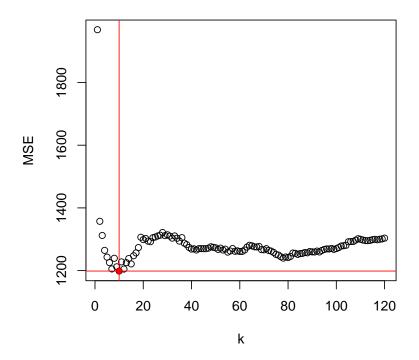


Figure 1.7: MSE against K neighbours (minimum in red)

We are looking for the k neighbors that has the smallest MSE so we use the function which.min that returns the index of the minimum value of MSE.

```
\begin{aligned} & model.knn.best.k = 11 \\ & model.knn.best.MSE = 1060.46 \end{aligned}
```

Now that we have the k that minimizes the MSE we call KNN algorithm with this best k and plot the predicted values against the real values as the residuals. The corresponding code is shown below:

The red line is the function y = x; so further are the points from this line the further are the predicted values (\hat{y}) from the real one (y).

y-hat (Predicted) against y

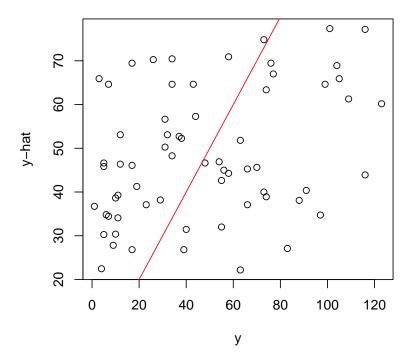


Figure 1.8: Predicted values against the real values

Distribution of Residuals in Knn best case

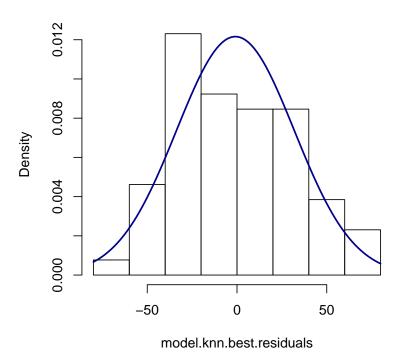


Figure 1.9: Distribution of Residuals in Knn best case

Model Analysis

We notice that the predicted \hat{y} diverge a lot from the real values y. We expected those results since the MSE is around 967 which is quite high.

This approach of finding the best k was quite optimistic. Actually we tried finding the best k while minimizing the MSE in **the test data**. Therefore the model is very specific to our test data which yields to a high bias. The solution is to find the best k among the **training data** and then use the best k in the test data.

To find the best **unbiased** k number of neighbors k we use the method of cross validation on the train data then we predict the response on the test.

1.4.2 The Validation Set Approach

There are two main methods in cross validation: the validation set approch and the cross validation leave one out (LOOCV) which is a particular case of the K-fold cross validation. As we do not have that much observations

(n=198) we can afford the computation of cross validation leave one out, but before we will argument our choice.

Cross Validation

The cross validation approach is based on dividing the provided data in 2 sets: a training set and a validation set. The model is fit on the training set, then the fitted model is used to predict responses of observations in the validation set. We validate our model using the best MSE.

Leave-One-Out Cross-Validation

Like the cross validation the LOOCV involves splitting the set of observations in two parts. The main difference is that we have a single observation (x_1, y_1) in the test data and the n-1 remaining is used for the train data. The MSE in this case is $MSE_1 = (y_1 - \hat{y})^2$. This provides an unbiased estimate for the test error since the size of the train model is approximately the one of all the data of the observation. However it is highly variable as it is based in one observation. The LOOCV repeats the procedure n times fitting each time a different set of observations. The LOOCV test MSE is computed with calculating the average of the n test error estimates.

$$CV_{(n)} = \frac{1}{n} \sum_{i=1}^{n} MSE_i$$

The LOOCV will always give the same results in contrast to the CV that depends on the randomness of how the data are split. Furthermore as it was stated before the CV runs the train approach on around the half of the size of the original data while the LOOCV repeats the validation set approach n times using n-1 observations. Hence the LOOCV yields to a not overestimated test error rate compared to the validation set approach. The only disadvantage of the LOOCV is it is computation time which can be very time consuming n is large.

An alternative to LOOCV that has a smaller computation time is k-Fold Cross Validation. This approach is based on dividing the training observations on k groups, each time one group will be considered as the test set and the k-1 left as the training set. Therefore the CV becomes:

$$CV_{(k)} = \frac{1}{k} \sum_{i=1}^{k} MSE_i$$

We can see that the k-Fold Cross-Validation fits the model k times instead of n, which reduces considerably the computation time. Our original data has only 198 observations, we can then afford the computation of the LOOCV.

Model implementation

After deducting the best k neighbors from the model we use it on the test observations to predict the values of Time. We then compute the MSE and plot the predicted values \hat{y} against the real ones y.

```
library("kknn")
  model.kknn = train.kknn(Time ~., data= train_set, kmax =
      30, ks = NULL, distance = 2, kernel = "optimal")
  model.kknn.best.k = model.kknn$best.parameters$k
  model.kknn.best = knn.reg(train=train_set.x,
      test=test_set.x, y=train_set.y, k=model.kknn.best.k)
  plot(test_set.y, model.kknn.best$pred, xlab='y',
      ylab='prediction')
   abline(0,1, col='red')
  residuals = test_set.y - model.kknn.best$pred
   errors = residuals^2
  model.kknn.best.MSE = mean(errors)
11
  hist(residuals, freq=FALSE, main="Distribution of Residuals
      in Knn LOOCV best case")
  residuals.mean = mean(residuals)
  residuals.stdev = sqrt(var(residuals))
  curve(dnorm(x, mean=residuals.mean, sd=residuals.stdev),
      col="darkblue", lwd=2, add=TRUE, yaxt="n")
```

 $\begin{aligned} & model.kknn.best.k = 30 \\ & model.kknn.best.MSE = 1146.05 \end{aligned}$

We run a the same plot again with the predicted reponses against the real ones 1.10 and the histogram of the residuals 1.11.

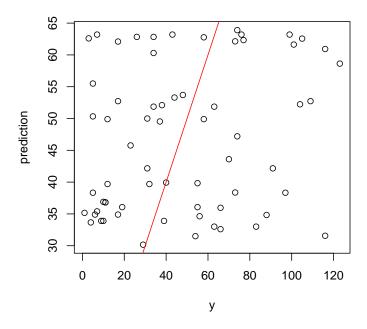


Figure 1.10: Predicted values against Real values with LOOCV $\,$

Distribution of Residuals in Knn LOOCV best case

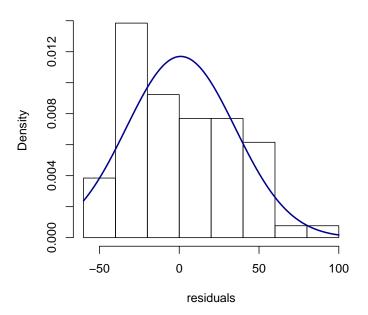


Figure 1.11: Distribution of Residuals in Knn with LOOCV

Model Analysis

One might argue that the prediction was not improved with the LOOCV since the test MSE obtained is higher that the MSE we obtained when we computed using the test set directly. However this result was expected since in the first approach we run our knn directly on the test so we underestimated the MSE, in other words the first MSE is specific to that test set of data (biased), in contrast to our the LOOCV model. In other words, in general the k=30 will guarantee a smaller MSE than the k=11 on any test observation. The LOOCV has also an advantage over the first approach because it will always give the same result of best k.

1.5 Simple Linear Regression

1.5.1 Idea

Our next attempt consists in using the same linear model we used in the feature analysis section. This model takes advantage of the simple assumption that Time depends on the other features in a linear way. A linear regression model

has the following form :

$$Y = \beta_0 + \sum_{i=1}^p \beta_i X_i$$

The regression coefficients β_i are choosen so that they minimize the MSE. Both analytical and optimization methods (Gradient Descent, Stochastical Gradient Descent, ...) can be used to find the best coefficient estimates.

To build the model in R, we can use the function 1m:

```
model.linreg = lm(Time ~ ., data=train_set)
```

1.5.2 Model Performance

This model has a MSE approximatively equal to 1285. The raw residuals distribution is pictured on figure 1.12, it shows a very spread out distribution that should be improved.

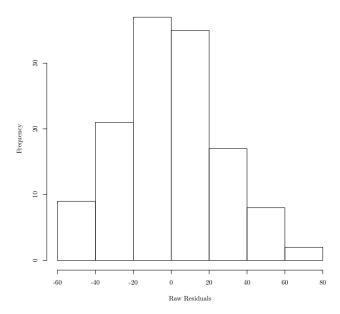


Figure 1.12

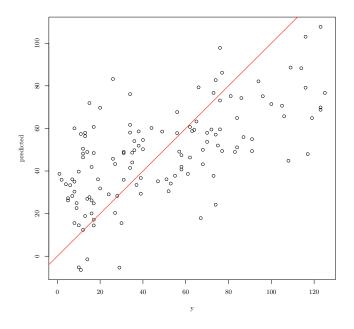


Figure 1.13

1.6 Linear Regression with Features Selection

1.6.1 Idea

A simple method to improve the performance of the previous Linear Regression is to select a subset of features that better describes the distribution of Time.

Once the simple linear model is fitted, we can use the function **summary** to display the value of each coefficient.

Coefficients:

Estimate Std. Error t value Pr(>|t|) (Intercept) 5.722e+01 1.587e+02 0.361 0.7191 Lymph_node -1.752e-01 7.034e-01 -0.2490.8038 radius_mean 3.001e+01 4.445e+01 0.675 0.5012 texture_mean -3.913e-01 2.043e+00 -0.1920.8485 perimeter_mean -0.570 0.5697 -3.794e+00 6.651e+00 area_mean -1.328e-01 1.354e-01 -0.981 0.3291 smoothness_mean -7.124e+02 9.050e+02 -0.7870.4331 compactness_mean -8.767e+01 3.554e+02 -0.247 0.8057 -0.973 -2.654e+02 2.728e+02 0.3332 concavity_mean concave_points_mean 1.236e+03 5.911e+02 2.090 0.0392 * symmetry_mean -3.199e+01 2.647e+02 -0.121 0.9041

```
fractal_dimension_mean
                          1.818e+03
                                     1.666e+03
                                                  1.091
                                                           0.2781
                         -2.897e+01
                                     9.953e+01
                                                 -0.291
                                                           0.7716
radius_se
                         -1.784e+00
                                     1.246e+01
                                                 -0.143
                                                           0.8865
texture_se
perimeter_se
                          8.817e+00
                                     1.289e+01
                                                  0.684
                                                           0.4955
area_se
                         -3.871e-01
                                     4.212e-01
                                                 -0.919
                                                           0.3604
smoothness_se
                          2.987e+03
                                     2.535e+03
                                                  1.178
                                                           0.2415
compactness_se
                          7.135e+02
                                     7.864e+02
                                                  0.907
                                                           0.3665
                                     7.170e+02
                                                  0.530
                                                           0.5975
concavity_se
                          3.798e+02
                                                           0.5926
concave_points_se
                         -7.782e+02
                                     1.450e+03
                                                 -0.537
                         -1.239e+03
                                                 -1.576
                                                           0.1182
symmetry_se
                                     7.861e+02
                         -5.979e+03
fractal_dimension_se
                                     6.280e+03
                                                 -0.952
                                                           0.3434
radius_worst
                          9.300e-01
                                     1.367e+01
                                                  0.068
                                                           0.9459
texture_worst
                         -1.307e+00
                                     1.964e+00
                                                 -0.666
                                                           0.5072
perimeter_worst
                         -1.026e+00 1.449e+00
                                                 -0.708
                                                           0.4805
area_worst
                          8.705e-02
                                     7.021e-02
                                                  1.240
                                                           0.2181
smoothness_worst
                         -3.761e+02
                                     4.088e+02
                                                 -0.920
                                                           0.3599
compactness_worst
                         -7.031e+01
                                     9.964e+01
                                                 -0.706
                                                           0.4821
concavity_worst
                         -3.011e+01
                                     7.311e+01
                                                 -0.412
                                                           0.6814
                                                 -0.930
concave_points_worst
                         -2.181e+02
                                     2.345e+02
                                                           0.3546
symmetry_worst
                          1.586e+02
                                     1.425e+02
                                                  1.113
                                                           0.2687
                          1.028e+03
                                     7.100e+02
fractal_dimension_worst
                                                  1.448
                                                           0.1509
                         -1.188e+00
                                     1.915e+00
                                                 -0.620
                                                           0.5364
Tumor_size
```

The last column contains the P-value of each coefficient, which is a measure to test the hypothesis that this particular coefficient is null. A P-value lower than 5% allows us to conclude that the coefficient is not equal to zero. In our case, the feature <code>concavity_points_mean</code> is not null, but we cannot make such assumptions for the other parameters. Therefore, we are not able to select a subset of interesting features based on the P-values.

Feature subset selection algorithms exist to extract the best subset of features from the dataset. Such method builds a linear model based on multiple subsets of features and compute a performance score to compare them. Our dataset contains a quite small set of features to deal with, therefore we can use an exhaustive feature subset selection algorithm which will apply a linear regression on each and every subset available.

In R, the following function is available to fit the linear models:

We can then use the function plot to compare the models according to a given scale. In our case we use the BIC measurements.

$$BIC = \frac{1}{n}(RSS + \log(n)d\hat{\sigma}^2)$$

The BIC will be smaller when the RSS is small, so we select the model with the

```
plot(model.linreg.regsubsets, scale="bic")
```

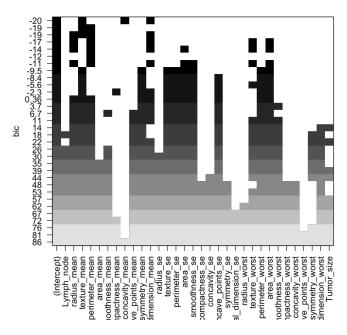


Figure 1.14: BIC-score of each feature subset

According to this plot, the best BIC is reached with a model that only uses the following features :

- texture_mean
- fractal_dimension_mean
- concavity_mean

1.6.2 Model Performance

The MSE of this model is approximatively equal to 1067, which is better than the full-featured model. The raw residual distribution, shown on figure 1.15, is not as spread out as the previous linear regression model.

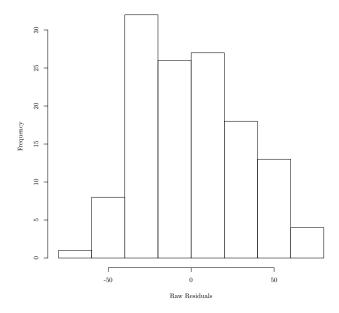


Figure 1.15

1.7 Linear Regression with Regularization

In this section we will discuss some methods that will help us shrink the model by reducing the number of parameters. We will use the **glmnet** package in order to build the ridge regression and the lasso in R.

1.7.1 Ridge Regression

The Linear Regression with least squares estimates the parameters $\beta_0, \beta_1...\beta_p$ that to minimize the term of the RSS.

$$RSS = \sum_{i=1}^{n} (y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij})^2$$

Ridge Regression works the same way as the least squares in the sense that it also tries to minimize the RSS but is also has another term $\lambda \sum j \beta_j^2$ called the **shrinkage penalty** where $\lambda \geq 0$ is the **tuning parameter**. The formula is:

$$RSS + \lambda \sum_{j} \beta_{j}^{2} \tag{1.1}$$

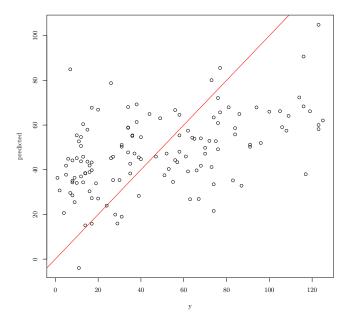


Figure 1.16

If $\lambda=0$ we are in the same case of a least squares estimates. The higher λ gets, the higher will be the penalty. Hence Ridge regression will try to minimize the parameters β_i in order to minimize the term 1.1.

When applying the penalty on the coefficients those with different scales (for instance a perimeter in m and the other one in Km) will be "treated" differently , because the penalized term is a sum of squares of all the coefficients. An alternative to get the penalty applied uniformly across the predictors is to standardize the independent predictors first with the function scale.

Model Implementation

The glmnet function takes for parameters the matrix x of predictors and vector y of responses. The model.matrix will help us transform out data sets into matrix. This function not only gives out a matrix but it also converts the qualitative variables into dummy variables.

```
train_set.x = model.matrix(Time~.,train_set)[,-1]
train_set.y = train_set$Time
scale(train_set.x, center = TRUE, scale = TRUE)

test_set.x =model.matrix(Time~., test_set)[,-1]
test_set.y = test_set$Time
```

```
7 scale(test_set.x, center = TRUE, scale = TRUE)
```

The glmnet function takes in parameter the train data, the parameter **alpha** indicates whether it is a Ridge or Lasso regression (alpha=0 for Ridge). By default glmnet chooses an automatic range of λ , however we chose a wide range with the function **grid** that takes the minimum value the maximum and the length and returns a grid. We chose $\lambda \in [10^{-2}, 10^{10}]$ to cover all possibilities.

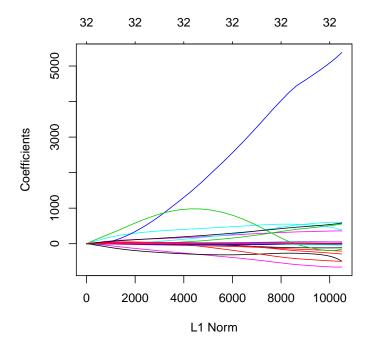


Figure 1.17: Coefficients β_j against L1 Norm

Each curve of the 1.17 corresponds to the ridge regression coefficient estimate for one of the 33 predictors, plotted as a function of L1Norm = . It is to be noted that the higher the L1 Norm, the closer to zero are the coefficients β_j .

In order to find the best tuning parameter λ we perform a cross validation on the training data. cv.glmnet function runs a 10 fold cross validation on the

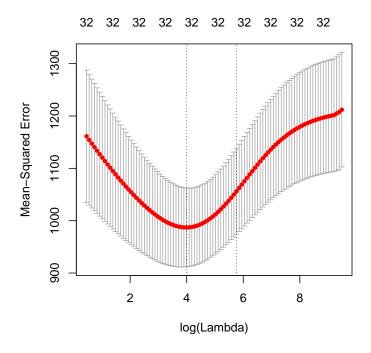


Figure 1.18: MSE against $\log(\lambda)$

The value λ that yields to the smallest MSE is shown in the graph 1.18 near $log(\lambda)=4$ and $log(\lambda)=5$

```
hist(residuals, freq=FALSE, main="Distribution of Residuals
                in Ridge")
residuals.mean = mean(residuals)
residuals.stdev = sqrt(var(residuals))
curve(dnorm(x, mean=residuals.mean, sd=residuals.stdev),
                col="darkblue", lwd=2, add=TRUE, yaxt="n")
```

model.ridge.best.lambda = 45.07model.ridge.best.MSE = 1063.3

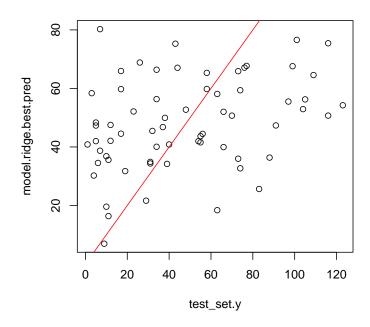


Figure 1.19: predicted Time \hat{y} against real responses y

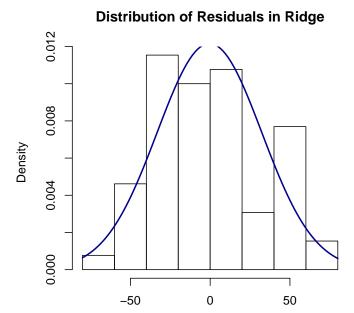


Figure 1.20: Residuals against Time(y)

residuals

Performing a Ridge regression didn't really improve out MSE, we notice that we still have the same shape when we plot the predicted response against the real values. In fact we still have the extreme values scattered away from the line y=x while the values in the middle are grouped around it. The residuals shows most residuals are among -20 and 40, however we can't neglect the important number of observation that have residuals below -20.

After predicting the responses with the best- λ , we can also to get the coefficients β_j for this model.

```
(Intercept)
                        62.5015644263887
          Lymph_node
                        -0.100938631782285
          radius_mean
                        -0.49162598035402
         texture mean
                        -0.546588424261081
      perimeter_mean
                        -0.0736547875729137
                        -0.00430679044764489
           area_mean
                        96.8117931071261
    smoothness_mean
   compactness_mean
                        4.59693750491085
       concavity_mean
                        -14.6877128128471
 concave_points_mean
                        -6.59290982908008
      symmetry_mean
                        33.6374215458585
fractal_dimension_m...
                        237.875435541426
             radius se
                        -0.554934245779119
            texture_se
                        -4.18461238448654
         perimeter_se
                        -0.148476465484444
                        -0.0120040219753369
              area_se
       smoothness_se
                        150.189441839833
      compactness_se
                        6.11836242882208
          concavity_se
                        -88.7802623855958
    concave_points_se
                        -146.9592219009
         symmetry_se
                        26.0064126060648
  fractal_dimension_se
                        401.735755303345
          radius worst
                        -0.171542834544902
         texture_worst
                        -0.33291548057428
      perimeter_worst
                        -0.0333700857746032
           area_worst
                        -0.00111259498281325
    smoothness_worst
                        49.8677352921429
   compactness_worst
                        4.02281481898061
      concavity_worst
                        -7.31642217426913
 concave_points_worst
                        5.14294581484783
      symmetry_worst
                        17.695513572159
fractal dimension wo...
                        85.6833237064771
           Tumor_size
                        -0.415565750547081
```

Figure 1.21: Ridge coefficients β_i for $(\lambda = 45.07)$

Model Analysis

So we notice that some coefficients are very close to 0 such as area_se (-0.012), but none of them is null. In fact Ridge Regression only shrinks the coefficients and does not perform any variable selection; that is why we are going to use Lasso in the following part.

After comparing both MSE of the least squares (1285) and the Ridge Regression's one (1063) we notice that Ridge improves over Linear Regression. The question is how does Ridge improve over the least squares? It all can be summarized in **biais-variance trade-off**. Let's take the following example where we we fit the model for each λ and predict the responses for $\lambda \in [10^{-1}, 10^4]$. For each value of λ we compute the MSE, the squared bias and the cube root of the variance. We computed the cube root in order to plot it with the bias in the same plot.

```
max = 100
  biais2<-rep(0,max)
  variance <-rep(0,max)</pre>
  mse \leftarrow rep(0, max)
   grid=10^seq(4,-1, length=max)
  for( i in 1:max)
  {
  fit.ridge = glmnet(x.train, y.train, lambda=grid[i],
      alpha=0)
  ridge.pred = predict(fit.ridge, s=grid[i] ,newx=x.test)
  mse[i] = mean((ridge.pred - y.test)^2) #MSE
  biais2[i] = (mean(ridge.pred - y.test))^2 #squared bias
  variance[i] = (var(ridge.pred))^(1/3) #variance
13
  plot(grid, variance, type='1', xlab="lambda",
      main="Biais-Variance trade-off")
  lines(grid, biais2,col='red')
  plot(MSE)
```

Bias-Variance trade-off

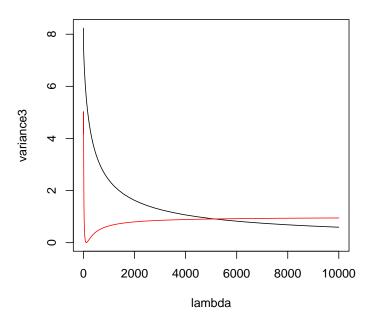


Figure 1.22: Bias Variance trade-off, squared bias (red), root squared variance (black)

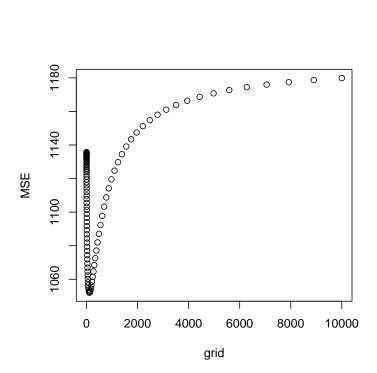


Figure 1.23: MSE against λ

At $\lambda=0$ (least squares), the variance is high but there is no bias. As λ increases the variance and bias decrease but at some point the variance continues to decrease while the bias starts increasing. Which means that the shrinkage of the coefficients by λ reduces the variances on the expense of the bias, which can be explained by the fact that λ in underestimating the coefficients by shrinking them. Let us compare now those results to the MSE curve.

$$MSE = (E[\Theta] - \Theta)^2 + Var(\Theta) = (Bias[\Theta])^2 + Var(\Theta)$$

Hence if the variance and the bias decrease significantly in the beginning the MSE will decrease too, and when the variance will decrease less and the bias will start increasing the MSE will increase too. We should also note that at at $\lambda = 0$ (least squares) the MSE is high.

To sum up, in linear regression we can have a low bias but a high variance, this is where the ridge regression improves over the least squares because it shrinks the variance on the expense of the bias.

1.7.2 Lasso Regression

As it was discussed on the previous section, Ridge regression's disadvantage is that it does not perform a variable selection. Lasso Regression however enables variable selection by minimizing the term:

$$RSS + \lambda \sum_{j} |\beta_{j}^{2}| \tag{1.2}$$

The only difference with the term of Ridge is that we now have $|\beta_j^2|$ instead of β_j^2 . This alternative will reduce the β_j that are close to zero to null, hoping we have a more interpretable model with reducing the number of variables. To call the lasso model we use the same function **glmnet** but with the parameter **alpha=1**.

Model Implementation

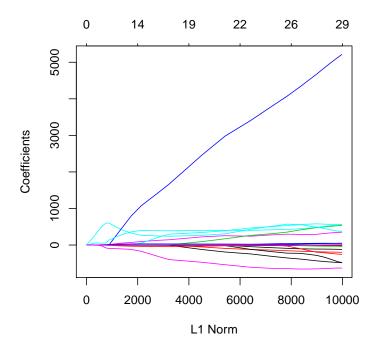


Figure 1.24: Coefficients against L1 Norm

We notice that for some values of λ the coefficients are null. We now perform the 10 fold cross validation on the training set to deduct the best tuning pa-

rameter. **cv.glmnet** does k-fold cross-validation for glmnet (by default k=10) it then produces a plot and returns a value for lambda.

After running the cross validation we get the best lambda that minimized the term 1.2. We now can run a prediction on the test set with the best lambda and compute the residuals and the Test MSE.

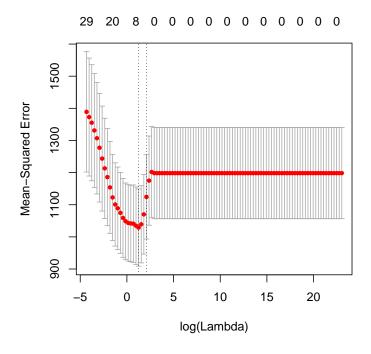


Figure 1.25: MSE against $log(\lambda)$

```
# predicted values against real values
plot(x=test_set.y, y=model.lasso.pred)
# y = x
abline(0,1, col='red')

# histogram of residuals and the normal distribution
hist(residuals, freq=FALSE, main="Distribution of Residuals in Lasso")
residuals.mean = mean(residuals)
residuals.stdev = sqrt(var(residuals))
curve(dnorm(x, mean=residuals.mean, sd=residuals.stdev), col="darkblue", lwd=2, add=TRUE, yaxt="n")
dev.off()
```

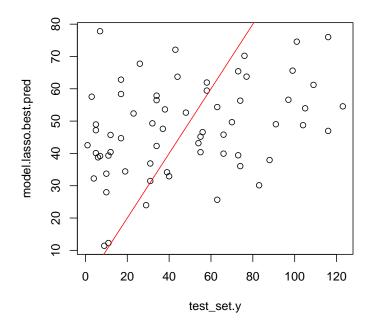


Figure 1.26: Predicted responses \hat{y} against real values y

Distribution of Residuals in Lasso

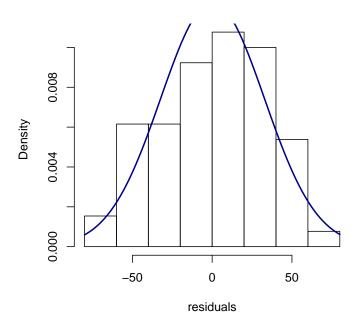


Figure 1.27: Distribution of Residuals in Lasso

```
predict(model.lasso, type="coefficients", s=model.lasso.best.lambda)[1:33,]
```

```
(Intercept)
                        79.0308812608241
          Lymph_node
          radius_mean
                        0
         texture_mean
                        -1.02108853748218
       perimeter_mean
                        -0.339703271620677
           area_mean
                        -0.00128975783963866
    smoothness_mean
                        0
                        0
   compactness_mean
                        0
       concavity_mean
                        0
 concave_points_mean
      symmetry_mean
                        0
fractal_dimension_m...
                        511.808342216284
             radius_se
                        -5.70256925538677
            texture_se
         perimeter_se
                        0
                        0
              area_se
                        0
       smoothness_se
      compactness_se
          concavity_se
                        -34.3617192851664
    concave_points_se
                        0
                        0
         symmetry_se
  fractal_dimension_se
                        0
                        0
          radius_worst
         texture_worst
                        0
      perimeter_worst
                        0
                        0
           area worst
    smoothness_worst
                        0
                        0
   compactness_worst
      concavity_worst
                        0
 concave_points_worst
                        0
      symmetry_worst
                        8.55146169876415
fractal_dimension_wo...
                        53.1202672054075
           Tumor_size
```

Figure 1.28: Lasso coefficients β_j for $(\lambda = 3.25)$

We can clearly see that Lasso performed a variable selection by setting the

unselected predictors' coefficients to 0

Model Analysis

Lasso selected only 8 predictors out of 33 with a slightly improved test MSE compared to Ridge that uses all the predictors. Therefore we selected 25% of variables The features selected to predict the Time are :

- texture_mean
- perimeter_mean
- smoothness_mean
- fractal_dimension_mean
- texture_se
- smoothness_se
- smoothness_worst
- symmetry_worst
- fractal_dimension_worst

Those are the features that with $\lambda=3.25$ their coefficients β was not null . Lasso selected only 10 predictors out of 33 with an slightly improved MSE compared to Ridge that uses all the 33 predictors.

1.8 Linear Regression with Dimension Reduction

1.8.1 Idea

In the Introduction, we mentioned the fact that some features are actually correlated. This means that they carry redundant information that make the model more complex, thus harder to fit. In this section, we will apply a Dimension Reduction method that decreases the number of features while keeping the information needed to predict Time.

Principal Component Analysis (PCA) is one such method. It consists in transforming the set of p features into a set of M orthogonal vectors (M < p) using linear transformations. The new set of vectors (called components) contains as much information as the initial set. "Information" is here defined in terms of variance, in other words, each new component should be able to explain as much variance as possible, while being orthogonal to the others.

Once a set of principal components is found, it can be used as an input to a simple linear regression to build what is called a Principal Component Regression (PCR) model. The number of components M should be taken so

that the model built with PCR has a minimum error value; once again, a 10 fold cross validation method can be used to determine the best M.

PCR is available in R with the package pls:

```
library(pls)
model.pcr = pcr(Time ~ ., data=train_set, scale=TRUE,
validation="CV")
```

We can then plot the MSE for each set of components :

```
validationplot(model.pcr, val.type = "MSEP")
```

According to figure 1.29, the model with only 4 components yields to the lowest MSE. However, since M=4 is found after running a random cross validation method, this number may be different if we run this algorithm another time.

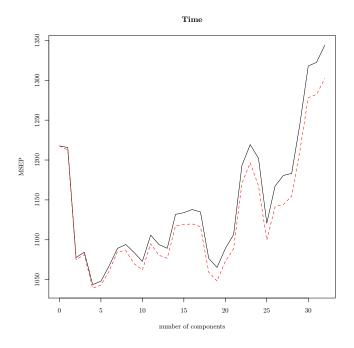


Figure 1.29: PCR

The function summary tells us that 4 components are enough to explain 75% of the features' variance, which seems to be enough for this dataset. Indeed, this model performs better that the other linear regression methods: the MSE is approximatively equal to 967.

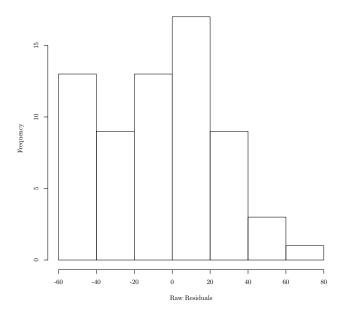


Figure 1.30: PCR Raw Residuals Distribution

1.9 Models Comparaison

Model	Mean Squared Error (MSE)		
KNN neighbours (with LOOCV)	1146		
Linear Regression	1285		
Best Subset	1067		
Ridge Regression	1063		
Lasso	1048		
PCR	967		

1.9.1 KNN neighbors and Linear Regression

The linear regression is a **parametric method**, it comes with the hypothesis that there is a relationship between the response Y and the p predictors X_i that can be expressed like $Y = \beta_1 X_1 + \beta_2 X + 2 + ... + \beta_p X_p$. Their main advantages is that it is easy to use and to interpret however when the relationship between the predictors and the response is not linear the the results might not be satisfying. The KNN neighbors however is a non parametric method. In other words it makes no assumptions on the response Y. If we compare between the **test MSE** obtained by the KNN (1146) and the one obtained by the Linear Regression (1285), we see that the KNN does a better job. These arguments the fact that



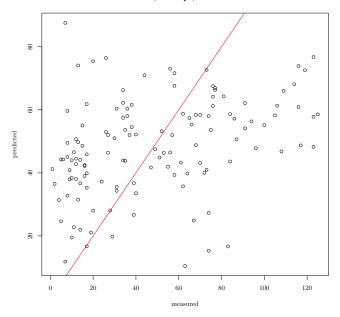


Figure 1.31: PCR Predicted values against Real values

the model is not really linear. A remark that was made when observing the QQ plots in the first section.

1.9.2 Best subset Selection and Linear Regression Regularizations

Ridge & Lasso

The lasso's advantage over Ridge Regression is the variable selection; the model is thus easier to interpret. The MSE in Lasso is also slightly improved. In fact Lasso works better when there isn't an important number of predictors which corresponds to our case since we only have 33. In order to compare between their variance bias trade off. We run the ridge regression with all the different values of λ but for Lasso we run the regression with the new data set that contains only 8 variables.

In order to create the new data_set of 8 predictors we used the function **cbind** that binds columns and **data.frame** for storing data tables. cbind did not keep the columns'names so we renamed them using the function **colnames**.

```
data_set$texture_se, data_set$concavity_se,
      data_set$symmetry_worst,
   data_set$fractal_dimension_worst, data_set$Time))
  colnames(newdata_set) <- c("texture_mean",</pre>
      "perimeter_mean", "area_mean", "fractal_dimension_mean",
      "texture_se",
   "concavity_se", "symmetry_worst", "fractal_dimension_worst",
      "Time" )
         = nrow(newdata_set)
         = ncol(newdata_set)
10
        = \operatorname{round}(2*n/3)
   napp
11
  ntst = n-napp
           = sample(1:n, napp)
13 train
14 newtrain_set = newdata_set[train,]
15 newtest_set
                 = newdata_set[-train,]
newtrain_set.x = model.matrix(Time~.,newtrain_set)[,-1]
newtrain_set.y = newtrain_set$Time
scale(newtrain_set.x, center = TRUE, scale = TRUE)
21 newtest_set.x =model.matrix(Time~., newtest_set)[,-1]
22 newtest_set.y = newtest_set$Time
  scale(newtest_set.x, center = TRUE, scale = TRUE)
```

Now that we have our new data_set, let's compare between the test MSE and variance bias trade off variance in function of the tuning parameter λ

```
#Number of itterations
               = 100
  #Ridge: squared Biais, root square variance and MSE
4 biais2_r = rep(0, max)
  variance3_r = rep(0, max)
  MSE_r
               = rep(0, max)
  #Lasso: squared Biais, root square variance and MSE
9 biais2_1
             = rep(0, max)
variance3_1 = rep(0,max)
11 MSE_1
              = rep(0, max)
_{\rm 13} #Grid of lamba between 10^-1 and 10^4
              = 10^seq(4,-1, length=max)
14 grid
15 for( i in 1:max)
16
      # Ridge Regression
17
      model.ridge = glmnet(train_set.x, train_set.y,
         lambda=grid[i], alpha=0)
      model.ridge.pred = predict(model.ridge, s=grid[i]
```

```
,newx=test_set.x)
      MSE_r[i] = mean((model.ridge.pred - test_set.y)^2) #MSE
20
      biais2_r[i] = (mean(model.ridge.pred - test_set.y))^2
21
          #squared bias
      variance3_r[i] = (var(model.ridge.pred))^(1/3) #sqrt
         cubic of variance
23
      #Lasso Regression with 8 predictors
24
      model.lasso = glmnet(newtrain_set.x , newtrain_set.y,
25
          lambda=grid[i], alpha=1)
      model.lasso.pred = predict(model.lasso, s=grid[i],
          newx=newtest_set.x)
      MSE_1[i] = mean((model.lasso.pred - test_set.y)^2) #MSE
27
      biais2_1[i] = (mean(model.lasso.pred - test_set.y))^2
28
         #squared bias
      variance3_1[i] = (var(model.lasso.pred))^(1/3) #sqrt
29
         cubic of variance
30 }
31 #Red for Ridge
32 #Variance
plot(grid, variance3_r, type="l", xlab="lambda", ylab="root
       squared Var and squared
      Biais",col="red",ylim=range(biais2_r, variance3_r,
      biais2_1, variance3_1),
      main="Bias(dashed)-Variance(line) trade-off of
      Ridge(Red) and Lasso(Blue)")
   lines(grid, biais2_r, col="red", lty=2)
37 #Blue for Lasso
38 #Variance
39 lines(grid, variance3_1, col="blue")
40 #Biais
1 lines(grid, biais2_1, col="blue", lty=2)
```

Bias(dashed)-Variance(line) trade-off of Ridge(Red) and Lasso(Blue)

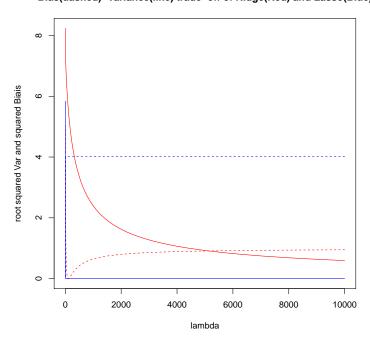


Figure 1.32: Bias(dashed line)-Variance(line) Trade off of Ridge(Red) and Lasso(Blue)

As it was stated in the subsection above 1.7.1 both Ridge and Lasso influence on the bias and the variance same way. It is clear from the curves above that Lasso decreases the variance more on the expense of the bias which explains why the Ridge's vaiance is higher than the Lasso's one but its bias is smaller. These leads us to conclude that there is neither the models performs better than the other. Lasso will be more pertinent to use if you have few predictors as it will set some coefficients to null and so the model will be easier to interpret. Nevertheless if we have lot of predictors then Ridge will be a better idea because all the coefficients will be small.

Regularization & Best subset

Both Lasso and Best Subset perform a variable selection, their MSE are close but with the Best Subset we only kept 3 variable which is 38% less predictors. How can we compare between those two approaches?

First Lasso adds up a new tuning parameter to penalize the coefficients which is why it is named shrinkage method. In contrast to the Best subset selection that doesn't add up anything. Therefore the regression coefficients obtained by Lasso are biased.

1.9.3 Dimension Reduction

Chapter 2

Phoneme Recognition

2.1 Context

In the context of speech recognition the aim is to predict which of the phonemes is pronounced my the subject. We have five phonemes to recognize:

```
g = \begin{cases} \text{"sh" as in "}she" \\ \text{"dcl" as in "}dark" \\ \text{"iy" as the vowel in "}she" \\ \text{"aa" as the vowel in "}dark" \\ \text{"ao" as the first vowel in "}water" \end{cases}
```

2.2 Dataset Description

The study involved 4509 speeches pronounced by 50 male speakers. The method used for speech recognition is the Log-periodogram, an estimate of the spectral density of the sound signal.

Our dataset is composed of 4509 log-periodograms (observations) of length 256 (features). The column labelled **speakers** identifies the different speakers. We notice that some are labeled train and some test. Hence we have a training set composed of 3340 observations (74%) and a test set that comprises 1169 observations (26%). The column **g** shows the responses. The frequencies of each phoneme for the 4509 speeches are shown in the table below.

Phonemes	aa	ao	dcl	iy	$_{ m sh}$
Train	519	759	562	852	648
Test	176	263	195	311	224
Total	695	1022	757	1163	872

Table 2.1: Frequencies of phonemes in the train and test data set

Our dataset tells us which speaker the periodogram is extracted from. We could input this information in the model, however, the system may not be aware of the speaker in a real-case scenario, therefore, we will not consider the speaker as a feature.

```
data_set = read.csv("phoneme.data.txt")
head(data_set)
data_set.nb_features = 256
```

2.3 Measure to Compare Models

Before building any model, we have to properly define the measures we will use later to compare their performance. In a classification problem, "accuracy" is often used. It is defined by :

$$A = \frac{\text{Number of Good Predictions}}{\text{Number of Cases}}$$

This is a very simple measure that may cause some issues when applied on classification problems that deal with skewed classes. For instance, let A and B, two classes. A contains 95% of the dataset, while B contains the remaining 5%. Let M a pretty bad classifier which classifies the whole dataset in class A. According to the definition stated above, this classifier has an accuracy of 95%, which is excellent. In this case, more specific performance measures have to be applied.

However, our current classification problem do not deal with skewed class since all classes are almost equally-represented in the dataset. Therefore, we will be able to run this performance measure without any important issue. This measure will be aplied on the test set in order not to include any biais.

```
train_set = data_set[1:3340, 2:258]
train_set.x = train_set[,1:256]
train_set.y = train_set[,257]

test_set = data_set[3341:4509, 2:258]
test_set.x = test_set[,1:256]
test_set.y = test_set[,257]
```

2.4 Classification

2.4.1 LDA - Linear Discriminant Analysis

The first model we are going to use takes advantage of Bayes' Theorem. Let $f_k(x)$ the class-conditional density of X when Y = k, and π_k the prior probability of class k. Bayes' Theorem proposes the following statement:

$$\mathbb{P}(Y = k | X = x) = \frac{f_k(x)\pi_k}{\sum_{l=1}^{K} f_l \pi_l}$$

If the class distribution of the dataset is representative of the real distribution, we can estimate π_k by calculating the proportion of k class elements in the dataset. How can we now estimate $f_k(x)$? Linear Discriminant Analysis assumes that $f_k(x)$ is a Gaussian distribution whose covariance matrix \sum is diagonal with $\forall k, \Sigma_k = \Sigma$.

Model Implementation

The function lda works the same was as lm for linear regression. It returns the group means that are the average of each predictor in each class. The coefficients of linear discriminants output are used to form the LDA decision rule. The prior probability is the percentage of the response for each class in the observation.

```
library(MASS)
model.lda = lda(g ~ ., data=train_set)
summary(model.lda)
```

```
Prior probabilities of groups:

aa ao dcl iy sh

0.1553892 0.2272455 0.1682635 0.2550898 0.1940120
```

The prior probabilities tell us that in the train observation 15% of phonemes are aa, 23% ao, 17% dcl, 25% iy and 19% sh.

Now that we trained the model we can call the predict function to predict the responses of the test set of data. This function returns an element class, a list of the predicted phonemes. An element posterior with the posterior probability of the response of the k-th class. Last the linear discriminants are found in \mathbf{x} .

```
model.lda.predicted = predict(model.lda,newdata=test_set)
perf = table(test_set$g,model.lda.predicted$class)
perf
sum(diag(perf))/dim(test_set)[1]
```

The table perf of the predicted responses is shown below:

```
aa
          ao dcl
                     iy
                          sh
 g
    129
                 0
                           0
aa
           47
                      0
                 0
ao
      39 223
                      1
                           0
       0
            0 190
                      5
                           0
dcl
iу
                   309
                           0
                 2
                      0 224
sh
```

0.919589392643285

Here is an example of how this table is read, for the second line when the phoneme pronounced is aa, the speech recognition detects that 129 is detected as aa and 47 as ao, in other words only 27% is misclassified. In order to compute the total error rate we divide the sum of the diagonal terms (which are the true phonemes detected) and divide it by the number total of observations We find that in 92% of the cases the speech recognition detected right the phoneme. This classifier performs quite well, but it does not work well with the phonemes aa and ao which sound very similar.

Improvements with PCA

The dataset has a high number of features (256) that may prevent the model from fitting well. We can cope with this performance issue by applying a dimension reduction method, such as Principal Component Analysis (PCA). The functionprcomp is available in R to compute the principal components of the 256 features. Since PCA depends on the scaling of the inputs, it is important to scale and ceter each feature.

```
pca = prcomp(train_set.x, center = TRUE, scale. = TRUE)
```

The main question is now: how many principal components should we use to replace our main features? The best number of components M yields to the highest accuracy. The following code finds the best M for the test set.

```
accs = matrix(0, 50, 1)
   for (M in 2:50) {
      train_set.pca.x = as.data.frame(pca$x[,1:M])
      train_set.pca = as.data.frame(cbind(train_set.y,
          train_set.pca.x))
      test_set.pca.x = predict(pca, newdata = test_set.x)[,1:M]
      test_set.pca = as.data.frame(cbind(test_set.y,
         test_set.pca.x))
      model.lda.pca = lda(train_set.y ~ ., data =
          train_set.pca)
      model.lda.pca.predicted =
         predict(model.lda.pca,newdata=test_set.pca)
      perf = table(test_set$g,model.lda.pca.predicted$class)
13
      accs[M] = sum(diag(perf))/dim(test_set)[1]
14
   }
15
   which.max(accs[-1])
   plot(accs[-1])
```

The best M on the test set is 8 (figure 2.1), which yields to an accuracy of

92.2%. However, this method raises a bia is since we computed M based on the test set. A proper way to find M is using a k fold cross validation technique.

```
library(caret)
   accs = matrix(0, 50, 1)
   for (M in 2:50) {
      a.train_set.pca.x = as.data.frame(pca$x[,1:M])
      a.train_set.pca = as.data.frame(cbind(train_set.y,
          a.train_set.pca.x))
      folds = createFolds(train_set.pca$train_set.y)
      acc = 0;
10
      for (k in 1:10) {
11
         validation_indexes = folds[[k]]
         a.train_set.x =
13
             a.train_set.pca.x[-validation_indexes,]
         a.train_set = a.train_set.pca[-validation_indexes,]
14
15
         a.validation_set.x =
             a.train_set.pca.x[validation_indexes,]
         a.validation_set =
17
```

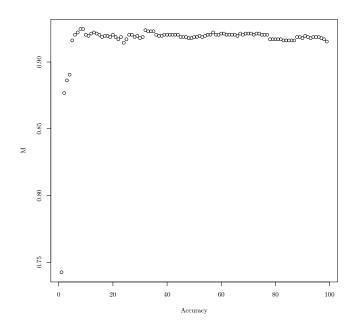


Figure 2.1: Best M for LDA + PCA on Test Set

```
a.train_set.pca[validation_indexes,]
18
         model.lda.pca = lda(a.train_set$train_set.y ~ ., data
19
             = a.train_set)
         model.lda.pca.predicted =
             predict(model.lda.pca,newdata=a.validation_set)
             table(a.validation_set$train_set.y,model.lda.pca.predicted$class)
         acc = acc + sum(diag(perf))/dim(a.validation_set)[1]
      }
      acc = acc / 10
26
      accs[M] = acc
27
   }
28
```

2.4.2 QDA- Quadratic Discriminant Analysis

Model Implementation

Model Analysis

2.4.3 Logistic Regression ROC??

Logistic Regression is a type of Linear Classification method that models the posterior probabilities with linear functions. When there are only two features to deal with, the model is definined by :

$$\log \frac{\mathbb{P}(Y=0|X=x)}{\mathbb{P}(Y=1|X=x)} = \beta_0 + \beta x$$

If $\beta_0 + \beta x$ is positive, Y is most likely to be equal to 0; otherwise, Y is most likely to be equal to 1.

In a multinomial scenario, like our case, more complex methods have to be applied. In particular, the library nnet provides us with a function called multinom that uses a feedforward Neural Network with a single hidden layer to fit a multinomial logistic regression model. This kind of Neural Networks often runs a backpropagation method along with an iterative optimization algorithm to estimate the weights of each neuron. That is why we have to specify the maximum number of iterations.

```
library("nnet")
model.lr = multinom(formula = g ~ . - g - speaker -
row.names, data=train, MaxNWts=2000, maxit=1000)
```

```
model.lr.predicted = predict(model.lr,newdata=test_set)
perf = table(test_set$g,model.lr.predicted)
```

```
perf
sum(diag(perf))/dim(test_set)[1]
```

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Model Analysis

2.5 Models Comparison

LDA and QDA