CSCI 3022 Intro to Data Science Regression nbs

Today we do 3 notebooks: (nb20, nb21, and an unnumbered notebook about prediction. (see "Predicts" on course schedule).

Broad strokes: For a linear model, we have:

- A list of assumptions + 3 about error
- Today: hypothesis tests based on those values and their standard errors.

- A few important statistics we calculate from the data: $\hat{\beta}_0$, $\hat{\beta}_1$, $\hat{\sigma^2}$, R^2

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

With 3 assumptions on ε :

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

2

$$\operatorname{Cov}[\varepsilon_i, \varepsilon_j] = 0 \qquad \forall i, j$$

Independence of errors

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

2

$$Cov[\varepsilon_i, \varepsilon_j] = 0 \quad \forall i, j$$

Independence of errors

3.

$$Var(\varepsilon_i) = \sigma^2 \qquad \forall i$$

Homoskedasticity of errors

Definition: Simple Linear Regression (SLR)

The Simple Linear Regression model is a model of the form

1.

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

With 3 assumptions on ε :

2

$$\mathsf{Cov}[\varepsilon_i, \varepsilon_j] = 0 \qquad \forall i, j$$

Independence of errors

3.

$$Var(\varepsilon_i) = \sigma^2 \qquad \forall i$$

Homoskedasticity of errors

4.

$$\varepsilon_i \sim N(0,1)$$

Simple Linear Regression Model

The β estimators in the model are:

- 1. $\hat{\beta}_0 = \bar{Y} \hat{\beta}_1 \bar{X}$
- 2. $\hat{\beta}_1 = \frac{Cov[X,Y]}{\sum_{i=1}^n (X_i \bar{X})^2} = \frac{\sum_{i=1}^n (X_i \bar{X})(Y_i \bar{Y})}{\sum_{i=1}^n (X_i \bar{X})^2}$

Important Terminology:

- \triangleright x: the independent variable, predictor, or explanatory variable (usually known). x is not random.
- \triangleright Y: The dependent variable or response variable. For fixed x, Y is random.
- \triangleright ε : The random deviation or random error term. For fixed x, ε is random. Has variance σ^2 .
- \triangleright β : the regression coefficients.
- ▶ r: the residuals or observed errors. Used to estimate σ^2 . The residuals are the differences between the observed and fitted y values: $\hat{\varepsilon_i} = r_i = \hat{e_I} = Y_i \hat{Y_i} = Y_i \hat{\beta_0} + \hat{\beta_1} X_i$ Mullen: Regression nbs

Error of a model

The goodness-of-fit of a regressive model is often decomposed into three components based on squared deviations. These are:

- 1. SSE: Sum of squared errors: (vertical) distances from the regression line to the data values. $\sum_i \left(\hat{Y} Y_i\right)^2$.
- 2. SST: Sum of squares, total: total deviation in Y. Looks like Var[Y]. $\sum_i \left(Y_i \bar{Y}\right)^2$
- 3. SSR: Sum of squares of regression line: the amount of variability tied to the model. $\sum_i \left(\hat{Y}_i \bar{Y}\right)^2$

The coefficient of determination is: $R^2 = 1 - \frac{SSE}{SST} = \frac{SSR}{SST}$ This coefficient is a number between 0 and 1 and is the proportion of observed y variation explained by the model.

Estimating SLR Parameters: Results

For a model of the form $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$; $\varepsilon \sim N(0, \sigma^2)$

1.
$$\hat{\beta_0} =$$

2.
$$\hat{\beta_1} =$$

Our estimate of the variance of the model is like a measure for an average of this summed errors SSE:

$$\hat{\sigma^2} = \frac{SSE}{n-2}.$$

Estimating SLR Parameters: Results

For a model of the form $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$; $\varepsilon \sim N(0, \sigma^2)$

$$1 \hat{\beta_0} = \bar{Y} - \hat{\beta_1} \bar{X}$$

$$2. \left(\hat{\beta}_{1}\right) = \frac{Cov[X,Y]}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}} = \frac{\sum_{i=1}^{n} (X_{i} - \bar{X})(Y_{i} - \bar{Y})}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}$$

Our estimate of the variance of the model is like a measure for an average of this summed errors SSE: $\hat{\sigma}^2 = \frac{SSE}{\pi^2}$.

$$\hat{\sigma^2} = \frac{SSE}{n-2}$$

Estimating SLR Parameters: Results

For a model of the form $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$; $\varepsilon \sim N(0, \sigma^2)$

1.
$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$$

2.
$$\hat{\beta}_1 = \frac{Cov[X,Y]}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

Our estimate of the variance of the model is like a measure for an average of this summed errors SSE:

$$\hat{\sigma^2} = \frac{SSE}{n-2}.$$

The parameters in SLR have distributions. From these distributions, we can conduct hypothesis tests (e.g., _____), compute confidence intervals, etc.

Distributions:

The parameters in SLR have distributions. From these distributions, we can conduct hypothesis tests (e.g., $H_0: \beta_1=0$), compute confidence intervals, etc.

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}; \quad \hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

Distributions:

The parameters in SLR have distributions. From these distributions, we can conduct hypothesis tests (e.g., $H_0: \beta_1 = 0$), compute confidence intervals, etc.

hypothesis tests (e.g.,
$$\underline{H_0:\beta_1=0}$$
), compute confidence intervals, etc.
$$\hat{\beta_0}=\bar{Y}-\hat{\beta_1}\bar{X}; \quad \hat{\beta_1}=\frac{\sum_{i=1}^n(X_i-\bar{X})(Y_i-\bar{Y})}{\sum_{i=1}^n(X_i-\bar{X})^2} \quad \text{Normal, each darking points}$$
 Distributions:
$$\hat{\beta_0} \sim N \left(\frac{\beta_0}{\beta_0} \frac{\sigma^2}{n} + \frac{\sigma^2\bar{X}^2}{\sum_{i=1}^n(X_i-\bar{X})^2} \right) \quad \text{each has, its on } \hat{\beta_1} \sim N \left(\frac{\beta_1}{\beta_1} \frac{\sigma^2}{(X_i-\bar{X})^2} \right) \quad \text{event of the primary of t$$

... but of course, we don't know σ^2 , so we estimate with SSE/(n-2).

Confidence Intervals: The CIs for regression are two-sided, and because $\varepsilon \sim N(0, \sigma^2)$, we may use t statistics. Since we have written down the variances of the β s, we can also write down their standard errors:

$$s.e.(\hat{\beta}_0) = \sqrt{\frac{1}{n} + \frac{X^2}{\left(X_i - \bar{X}\right)^2}};$$

where we replace
$$\sigma$$
 with the estimate $s=\frac{s}{2}$

Confidence Intervals: The CIs for regression are two-sided, and because $\varepsilon \sim N(0, \sigma^2)$, we may use t statistics. Since we have written down the variances of the β s, we can also write down their standard errors:

$$s.e.(\hat{\beta_0}) = \sigma \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{(X_i - \bar{X})^2}}; \qquad s.e.(\hat{\beta_1}) = \sigma \sqrt{\frac{1}{(X_i - \bar{X})^2}}$$
 These lead to CIs of
$$\beta_i \in (\hat{\beta_i}) \pm t_{\alpha/2, n-2} \underbrace{\{s.e.(\hat{\beta_i})\}}$$

where we replace σ with the estimate $s = \frac{SSE}{n-2}$

Confidence Intervals: The CIs for regression are two-sided, and because $\varepsilon \sim N(0, \sigma^2)$, we may use t statistics. Since we have written down the variances of the β s, we can also write down their standard errors:

$$s.e.(\hat{\beta}_0) = \sigma \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{(X_i - \bar{X})^2}}; \qquad s.e.(\hat{\beta}_1) = \sigma \sqrt{\frac{1}{(X_i - \bar{X})^2}}$$

These lead to CIs of

$$\beta_i \in (\hat{\beta}_i \pm t_{\alpha/2, n-2} \cdot s.e.(\hat{\beta}_i))$$

where we replace σ with the estimate $s = \frac{SSE}{n-2}$ be solved by the corresponding critical t values for a one or two-tailed test.

Mullen: Regression nbs

Inferences about Y

There are more types on confidence intervals we may care about!

- 1. Last slide was how to perform inference on the **parameters** of the *line* β . We also might care about inference on values of Y!
- 2. A **confidence band** is how sure we are about the mean of Y at specific values of X, or E[Y|X].
- 3. A **prediction band** is how we estimate the distribution of new Y observations at specific values of X. It's the same as the confidence band, but *also* includes our estimate for σ^2 . This is also known as a *forecast*.

Idea: If we want to **guess** the average $y=\beta_0+\beta_1 x$, (for a specified x) we have to combine our uncertainties for the β s. If we want to describe all the y's for a single value of x, we also would need to include the uncertainty $s^2\approx\sigma^2$ that accompanies ε .

See: nb accompanying lecture: SLR Prediction and Confidence

The usual inference:

The most common inference for linear regression is to answer the question "Does x affect y?" This is a hypothesis test asking about the value of the *slope* of the regression line. We have a CI for this of

$$\beta_1 \pm t_{\alpha/2,n-2} \cdot s.e.(\hat{\beta}_i)$$

where

$$s.e.(\hat{\beta_1}) = \sigma \sqrt{\frac{1}{\left(X_i - \bar{X}\right)^2}}$$

The corresponding hypothesis test is $t=\frac{\hat{\beta_i}}{s.e.(\hat{\beta_i})}$ with n=2 degrees of freedom. Two big things to notice

- 1. The error grows as σ grows: noisy/random data is harder to estimate.
- 2. The denominator *looks a lot* like the "standard deviation of x." We get more **confident** in our estimates if the predictor variable locations are spread out!

Daily Recap

Today we learned

1. Regression Inference!

Moving forward:

- nb day Friday

Next time in lecture:

- More Regression! More predictor!