

Xi Liu, xl3504, Homework 5

Problem 1

since $X \in \{0, 1, \dots, N\} = \{a_0, a_1, \dots, a_N\}$

expectation of a discrete random variable X with the values a_1, a_2, \dots

$$\begin{aligned} E(X) &= \sum_i a_i P(X = a_i) \\ &= \sum_{n=0}^N a_n P(X = a_n) \\ &= \sum_{n=0}^N n P(X = n) \end{aligned}$$

	P(X = 0)	P(X = 1)	P(X = 2)	P(X = 3)	P(X = 4)	P(X = 5)	...	P(X = N)
row 1		1	1	1	1	1		1
row 2			1	1	1	1		1
row 3				1	1	1		1
row 4					1	1		1
row 5						1		1
...								...
row N								1

for each column of $P(X = n)$, its contribution to the expected value is $nP(X = n)$, in which n is equal to the sum of 1's in each column of $P(X = n)$

$$\sum_{n=0}^N n P(X = n) = \sum_{i=1}^N (\text{contribution from row } i)$$

using summation by parts

$$\begin{aligned}
E(X) &= \sum_{n=0}^N n P(X = n) \\
&= \sum_{n=1}^N P(X = n) + \sum_{n=2}^N P(X = n) \\
&\quad + \sum_{n=3}^N P(X = n) + \dots + \sum_{n=N}^N P(X = n) \\
&= \sum_{n=0}^{N-1} \left(\sum_{i=n}^N P(X = i) \right) \\
&= \sum_{n=0}^{N-1} P(X > n) \\
&\quad /* \text{ since } P(X > n) = \sum_{i=n}^N P(X = i) */
\end{aligned}$$

Problem 2

1.

$$a := x$$

$$F(a) = P(X \leq a)$$

$$= \int_{-\infty}^a f(x) dx$$

$$= \int_0^a (2Kxe^{-Kx^2}) dx$$

$$/* u := -Kx^2$$

$$du = -2Kx dx$$

$$- du = 2Kx dx$$

$$- \int e^u du = -e^u */$$

$$= \left[-e^{-Kx^2} \right]_0^a$$

$$= - \left[e^{-Kx^2} \right]_0^a$$

$$= - \left(e^{-Ka^2} - e^0 \right)$$

$$= - \left(e^{-Ka^2} - 1 \right)$$

$$= 1 - e^{-Ka^2}$$

$$= 1 - e^{-Kx^2}$$

$$F(x) = \begin{cases} 1 - e^{-Kx^2} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

2.
mean = $E[X]$

$$\begin{aligned} E[X] &= \int_{-\infty}^{\infty} x f(x) dx \\ &= \int_0^{\infty} x \left(2K x e^{-Kx^2} \right) dx \\ &= \int_0^{\infty} \left(2K x^2 e^{-Kx^2} \right) dx \\ &= 2K \int_0^{\infty} x^2 e^{-Kx^2} dx \\ &\quad /* \text{ see calculations in next page } */ \\ &= 2K(I_1) \\ &= 2K \left(\frac{1}{2K} I_2 \right) \\ &= I_2 \\ &= \boxed{\sqrt{\frac{\pi}{4K}}} \end{aligned}$$

$$I_1 := \int_0^{\infty} x^2 e^{-kx^2} dx = \int_0^{\infty} x \cdot x e^{-kx^2} dx \quad \begin{cases} u := x; & du = dx \\ dv := x e^{-kx^2} & \end{cases}$$

$$V = \int dv = \int x e^{-kx^2} dx \quad \begin{cases} u_2 := -kx^2 \\ du_2 = -2kx dx; & x dx = \frac{du_2}{-2k} \end{cases}$$

$$V = -\frac{1}{2k} e^{-kx^2}$$

$$\int u dv = uv - \int v du$$

$$\begin{aligned} \int u dv &= x \left(-\frac{1}{2k} e^{-kx^2} \right) - \int \left(-\frac{1}{2k} e^{-kx^2} \right) dx \\ &= -\frac{1}{2k} x e^{-kx^2} + \frac{1}{2k} \int e^{-kx^2} dx \end{aligned}$$

$$\begin{aligned} I_1 &= \left[-\frac{1}{2k} x e^{-kx^2} \right]_0^{\infty} + \frac{1}{2k} \int_0^{\infty} e^{-kx^2} dx \\ &= -\frac{1}{2k} (0 - 0) + \frac{1}{2k} \int_0^{\infty} e^{-kx^2} dx \\ &= \frac{1}{2k} \int_0^{\infty} e^{-kx^2} dx \end{aligned}$$

$$I_2 := \int_0^{\infty} e^{-kx^2} dx$$

$$x^2 + y^2 = r^2, \quad x = r \cos \theta$$

$$y = r \sin \theta$$

$$I_2^2 = \left(\int_0^{\infty} e^{-kx^2} dx \right) \left(\int_0^{\infty} e^{-ky^2} dy \right)$$

$$= \int_0^{\infty} \int_0^{\infty} e^{-kx^2 - ky^2} dx dy = \int_0^{\infty} \int_0^{\infty} e^{-k(x^2 + y^2)} dx dy$$

$$= \int_0^{\pi/2} \int_0^{\infty} e^{-kr^2} r dr d\theta \quad \begin{cases} u_3 := -kr^2 \\ du_3 = -2kr dr; & r dr = \frac{du_3}{-2k} \end{cases}$$

$$= \left(\frac{\pi}{2} \right) \left(-\frac{1}{2k} \right) \left[e^{-kr^2} \right]_0^{\infty} = \left(-\frac{\pi}{4k} \right) (0 - 1) = \frac{\pi}{4k}$$

$$I_2^2 = \frac{\pi}{4k} \quad ; \quad I_2 = \sqrt{\frac{\pi}{4k}}$$

alternatively, the calculation of I_2 can be done as follows:

a continuous random variable has a normal distribution with parameters μ and $\sigma^2 > 0$ if its probability density function is given by

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx \\ &= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx \\ /* \text{ set } \mu = 0; \quad K &= \frac{1}{2\sigma^2}; \quad \sigma = \frac{1}{\sqrt{2K}} */ \\ &= \frac{1}{\sqrt{2\pi(1/\sqrt{2K})^2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}\left(\frac{x-0}{1/(\sqrt{2K})}\right)^2} dx \\ &= \frac{1}{\sqrt{2\pi}(1/\sqrt{2K})} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(\sqrt{2K}x)^2} dx \\ &= \frac{\sqrt{2K}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(2Kx^2)} dx \\ &= \sqrt{\frac{K}{\pi}} \int_{-\infty}^{\infty} e^{-Kx^2} dx = 1 \end{aligned}$$

/* since $\int_{-\infty}^{\infty} f(x) dx = 1$ for a f that is the probability density

function of a continuous random variable X */

$$\int_{-\infty}^{\infty} e^{-Kx^2} dx = \sqrt{\frac{\pi}{K}}$$

/* since $f(x) = e^{-Kx^2}$ is an even function */

$$\int_{-\infty}^{\infty} e^{-Kx^2} dx = 2 \int_0^{\infty} e^{-Kx^2} dx$$

$$\sqrt{\frac{\pi}{K}} = 2 \int_0^{\infty} e^{-Kx^2} dx$$

$$\begin{aligned} \int_0^{\infty} e^{-Kx^2} dx &= \frac{1}{2} \sqrt{\frac{\pi}{K}} \\ &= \sqrt{\frac{1}{4}} \sqrt{\frac{\pi}{K}} \\ &= \sqrt{\frac{\pi}{4K}} = I_2 \end{aligned}$$

variance = $Var(X)$

$$Var(X) = E[X^2] - E[X]^2$$

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx$$

$$\begin{aligned} E[X^2] &= \int_{-\infty}^{\infty} x^2 f(x) dx \\ &= \int_0^{\infty} x^2 (2Kx e^{-Kx^2}) dx \\ &= 2K \int_0^{\infty} x^3 e^{-Kx^2} dx \end{aligned}$$

/* see calculations in next page */

$$= 2K(I_3)$$

$$= 2K \left(\frac{1}{2K^2} \right)$$

$$= \frac{1}{K}$$

$$\begin{aligned}
I_3 &:= \int_0^\infty x^3 e^{-Kx^2} dx \\
&= \int_0^\infty x^2 \cdot x e^{-Kx^2} dx \\
u &:= x^2; \quad du = 2x dx; \quad dv := x e^{-Kx^2} dx \\
v &= \int dv = \int x e^{-Kx^2} dx \\
u_2 &:= -Kx^2; \quad du_2 = -2Kx dx; \quad x dx = -\frac{1}{2K} du_2 \\
v &= -\frac{1}{2K} \int e^{u_2} du_2 = -\frac{1}{2K} e^{-Kx^2} \\
\int u dv &= uv - \int v du \\
&= x^2 \left(-\frac{1}{2K} e^{-Kx^2} \right) - \int \left(-\frac{1}{2K} e^{-Kx^2} \right) (2x dx) \\
&= -\frac{1}{2K} x^2 e^{-Kx^2} + \int \frac{1}{K} x e^{-Kx^2} dx \\
I_3 &= \left[-\frac{1}{2K} x^2 e^{-Kx^2} \right]_0^\infty + \int_0^\infty \frac{1}{K} x e^{-Kx^2} dx
\end{aligned}$$

/* since $\forall a \in \mathbb{R}, \lim_{x \rightarrow \infty} \frac{e^x}{x^a} > 0$

the exponential function e^x is an asymptotic upper bound
on the polynomial function x^a */

$$\begin{aligned} I_3 &= -\frac{1}{2K}(0-0) + \int_0^\infty \frac{1}{K} x e^{-Kx^2} dx \\ &= \int_0^\infty \frac{1}{K} x e^{-Kx^2} dx \\ &= \frac{1}{K} \int_0^\infty x e^{-Kx^2} dx \end{aligned}$$

$$/* \quad u := -Kx^2; \quad du = -2Kx dx; \quad x dx = \frac{du}{-2K}$$

$$\frac{1}{K} \int x e^{-Kx^2} dx = \frac{1}{K} \cdot \frac{1}{-2K} e^{-Kx^2} = \frac{1}{-2K^2} e^{-Kx^2}$$

*/

$$\begin{aligned} I_3 &= \left[-\frac{1}{2K^2} e^{-Kx^2} \right]_0^\infty \\ &= -\frac{1}{2K^2}(0-1) \\ &= \frac{1}{2K^2} \end{aligned}$$

$$\begin{aligned} Var(X) &= E[X^2] - E[X]^2 \\ &= \frac{1}{K} - \left(\sqrt{\frac{\pi}{4K}} \right)^2 \\ &= \frac{1}{K} - \frac{\pi}{4K} \\ &= \boxed{\frac{4-\pi}{4K}} \end{aligned}$$

3.

let $F_Y(a)$ be the cumulative distribution function of Y

$$\begin{aligned} F_Y(y) &= P(Y \leq y) \\ &= P(KX^2 \leq y) \\ &= P(X \leq \sqrt{\frac{y}{K}}) \\ &= 1 - e^{-K(\sqrt{y/K})^2} \\ &= 1 - e^{-K(y/K)} \\ &= \boxed{1 - e^{-y}} \end{aligned}$$

let $f_Y(x)$ be the probability density function of Y

$$\begin{aligned} f_Y(x) &= \frac{d}{dy} (F_Y(y)) \\ &= \frac{d}{dy} (1 - e^{-y}) \\ &= (-1)(-e^{-y}) \\ &= e^{-y} \end{aligned}$$

$$f_Y(x) = \boxed{\begin{cases} e^{-y} & \text{if } y > 0 \\ 0 & \text{otherwise} \end{cases}}$$

Problem 3

1.

mean = $E[X]$

$$\begin{aligned} E[X] &= \int_{-\infty}^{\infty} x f(x) dx \\ &= \int_0^{\infty} x (\lambda e^{-\lambda x}) dx \\ &= \lambda \int_0^{\infty} x e^{-\lambda x} dx \end{aligned}$$

/* see calculations in next page */

$$= \lambda(I_6)$$

$$= \lambda \left(\frac{1}{\lambda^2} \right)$$

$$= \boxed{\frac{1}{\lambda}}$$

$$\begin{aligned}
I_6 &:= \int_0^\infty x e^{-\lambda x} dx \\
u &:= x; \quad du = dx; \quad dv := e^{-\lambda x} dx \\
v &= \int dv = \int e^{-\lambda x} dx \\
u_2 &:= -\lambda x; \quad du_2 = -\lambda dx; \quad dx = \frac{du_2}{-\lambda} \\
v &= -\frac{1}{\lambda} e^{-\lambda x} \\
\int u dv &= uv - \int v du \\
&= x \left(-\frac{1}{\lambda} e^{-\lambda x} \right) + \int \left(-\frac{1}{\lambda} e^{-\lambda x} \right) dx \\
&= -\frac{1}{\lambda} x e^{-\lambda x} - \frac{1}{\lambda} \int e^{-\lambda x} dx \\
u &= -\lambda x; \quad du = -\lambda dx; \quad dx = \frac{du}{-\lambda} \\
I_6 &= \left[-\frac{1}{\lambda} x e^{-\lambda x} \right]_0^\infty + \left(\frac{1}{\lambda^2} \right) [e^{-\lambda x}]_0^\infty \\
&= \left(-\frac{1}{\lambda} \right) (0 - 0) + \left(\frac{1}{\lambda^2} \right) (0 - 1) \\
&= \frac{1}{\lambda^2}
\end{aligned}$$

2.

variance = $Var(X)$

$$Var(X) = E[X^2] - E[X]^2$$

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f_X(x)dx$$

$$\begin{aligned} E[X^2] &= \int_{-\infty}^{\infty} x^2 f(x) dx \\ &= \int_0^{\infty} x^2 (\lambda e^{-\lambda x}) dx \\ &= \lambda \int_0^{\infty} x^2 e^{-\lambda x} dx \end{aligned}$$

/* see calculations in next page */

$$= \lambda(I_7)$$

$$= \lambda \left(\frac{2}{\lambda^3} \right)$$

$$= \frac{2}{\lambda^2}$$

$$Var(X) = E[X^2] - E[X]^2$$

$$= \frac{2}{\lambda^2} - \left(\frac{1}{\lambda} \right)^2$$

$$= \frac{2}{\lambda^2} - \frac{1}{\lambda^2}$$

$$= \boxed{\frac{1}{\lambda^2}}$$

$$\begin{aligned}
I_7 &:= \int_0^\infty x^2 e^{-\lambda x} dx \\
u &:= x^2; \quad du = 2x dx; \quad dv := e^{-\lambda x} dx \\
v &= \int dv = \int e^{-\lambda x} dx \\
u_2 &:= -\lambda x; \quad du_2 = -\lambda dx; \quad dx = \frac{du_2}{-\lambda} \\
v &= -\frac{1}{\lambda} e^{-\lambda x}
\end{aligned}$$

$$\begin{aligned}
\int u dv &= uv - \int v du \\
&= x^2 \left(-\frac{1}{\lambda} e^{-\lambda x} \right) + \int \left(-\frac{1}{\lambda} e^{-\lambda x} \right) (2x dx) \\
&= -\frac{1}{\lambda} x^2 e^{-\lambda x} - \frac{2}{\lambda} \int x e^{-\lambda x} dx \\
&\quad /* \text{ from previous calculation of } I_6, \text{ we know} \\
&\quad \int x e^{-\lambda x} dx = -\frac{1}{\lambda} x e^{-\lambda x} + \frac{1}{\lambda^2} e^{-\lambda x} */ \\
&= -\frac{1}{\lambda} x^2 e^{-\lambda x} - \frac{2}{\lambda} \left(-\frac{1}{\lambda} x e^{-\lambda x} + \frac{1}{\lambda^2} e^{-\lambda x} \right) \\
&= -\frac{1}{\lambda} x^2 e^{-\lambda x} + \frac{2}{\lambda^2} x e^{-\lambda x} - \frac{2}{\lambda^3} e^{-\lambda x} \\
&= e^{-\lambda x} \left(-\frac{1}{\lambda} x^2 + \frac{2}{\lambda^2} x - \frac{2}{\lambda^3} \right) \\
I_7 &= \left[e^{-\lambda x} \left(-\frac{1}{\lambda} x^2 + \frac{2}{\lambda^2} x - \frac{2}{\lambda^3} \right) \right]_0^\infty \\
&= \left(0 - (1) \left(-0 + 0 - \frac{2}{\lambda^3} \right) \right) \\
&= \frac{2}{\lambda^3}
\end{aligned}$$

3.
part 1

$$a := x$$

$$\begin{aligned} F(a) &= P(X \leq a) \\ &= \int_{-\infty}^a f(x) dx \\ &= \int_0^a (\lambda e^{-\lambda x}) dx \\ &= \lambda \int_0^a e^{-\lambda x} dx \end{aligned}$$

$$\begin{aligned} u &:= -\lambda x; & du &= -\lambda dx; & dx &= \frac{du}{-\lambda} \\ &= -[e^{-\lambda x}]_0^a \\ &= -(e^{-\lambda a} - 1) \\ &= 1 - e^{-\lambda a} \\ &= 1 - e^{-\lambda x} \end{aligned}$$

$$F(x) = \begin{cases} 1 - e^{-\lambda x} & \text{if } x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

complementary CDF:

$$\begin{aligned} P(X > x) &= 1 - (1 - e^{-\lambda x}) \\ &= e^{-\lambda x} \end{aligned}$$

part 2

$$\begin{aligned}P(X > s + t | X > t) &= \frac{P((X > s + t) \cap (X > t))}{P(X > t)} \\&\quad /* \text{ since } s > 0, t > 0 */ \\&= \frac{P(X > s + t)}{P(X > t)} \\&= \frac{1 - P(X \leq s + t)}{1 - P(X \leq t)} \\&= \frac{1 - (1 - e^{-\lambda(s+t)})}{1 - (1 - e^{-\lambda t})} \\&= \frac{e^{-\lambda(s+t)}}{e^{-\lambda t}} \\&= e^{-\lambda s} \\&= P(X > s)\end{aligned}$$

Problem 4

X is the random variable corresponding to the time of collapse (in million years)

1.

$$\begin{aligned}P(X \leq 1) &= 0.00000002 \\&= 1 - e^{-\lambda(1)} \\&= 1 - e^{-\lambda} \\e^{-\lambda} &= 1 - 0.00000002 = 0.99999998 \\-\lambda &= \ln(0.99999998) \\\lambda &= -\ln(0.99999998) \\\lambda &= \boxed{2.00000002 \cdot 10^{-8}}\end{aligned}$$

2.

$$\begin{aligned}E[X] &= \frac{1}{\lambda} \\&= \frac{1}{2.00000002 \cdot 10^{-8}} \\&= \boxed{49999999.5}\end{aligned}$$

3.

$$\begin{aligned} billion &= 10^9 \\ million &= 10^6 \\ 1 \text{ billion} &= (10^3)(10^6) = 1000 \text{ million} \\ 3 \text{ billion} &= 3000 \text{ million} \\ P(X \leq 3000) &= 1 - e^{-\lambda(3000)} \\ &= 1 - e^{-3000\lambda} \\ &= 1 - e^{-3000(2.00000002 \cdot 10^{-8})} \\ &= 1 - 0.99994000179 \\ &= \boxed{0.00005999821} \end{aligned}$$

4.

$$\begin{aligned} 10 \text{ billion} &= 10000 \text{ million} \\ P(X > 10000) &= e^{-\lambda(10000)} \\ &= e^{-(10000)\lambda} \\ &= e^{-(10000)(2.00000002 \cdot 10^{-8})} \\ &= \boxed{0.99980001999} \end{aligned}$$

Problem 5

let A be the event: “a machine has a failure within the first 5 years of operation”

let B be the event: “a machine goes permanently out of order”

$$P(A) = 0.3$$

$$P(A^C) = 1 - 0.3 = 0.7$$

$$P(B|A) = 0.75$$

$$P(B|A^C) = 0.4$$

1.

$$\begin{aligned} P(B) &= \sum_i P(B|A_i)P(A_i) \\ &= P(B|A)P(A) + P(B|A^C)P(A^C) \\ &= (0.75)(0.3) + (0.4)(0.7) \\ &= \boxed{0.505} \end{aligned}$$

2.

$$\begin{aligned} P(A^C|B) &= \frac{P(A^C \cap B)}{P(B)} \\ P(A^C \cap B) &= P(B \cap A^C) = P(B|A^C)P(A^C) \\ P(A^C|B) &= \frac{P(B|A^C)P(A^C)}{P(B)} \\ &= \frac{(0.4)(0.7)}{0.505} \\ &= \boxed{0.55445544554} \end{aligned}$$

3.

let $p_X(a)$ be the probability mass function for X

$p := P(A) = 0.3$

$1 - p = 1 - 0.3 = 0.7$

$$\binom{10}{a}(1-p)^{10-a}p^a = \binom{10}{a}(0.7)^{10-a}(0.3)^a$$

$$p_X(a) = \begin{cases} \binom{10}{a}(0.7)^{10-a}(0.3)^a & \text{if } a \in [0, 10] \cap \mathbb{N} \\ 0 & \text{otherwise} \end{cases}$$

4.

part 1

expected value = $E[X]$

let n be the number of trials

$$\begin{aligned} E[X] &= \sum_i a_i P(X = a_i) \\ &= \sum_{a=0}^{10} a \binom{10}{a} (0.7)^{10-a} (0.3)^a \\ &\quad /* \text{ see calculations in next page } */ \\ &= \boxed{3} \end{aligned}$$

alternatively:

$$\begin{aligned} E[X] &= np \\ &= 10(0.3) \\ &= \boxed{3} \end{aligned}$$

```

/* calculation of  $\sum_{a=0}^{10} a \binom{10}{a} (0.7)^{10-a} (0.3)^a$  */

#include <stdio.h>
#include <math.h>

int fac(int n)
{
    int ret = 1;
    for(int i = 2; i <= n; i++)
        ret = ret * i;
    return ret;
}

int ncr(int n, int r)
{
    return fac(n) / (fac(r) * fac(n - r));
}

void exp_x()
{
    double sum = 0;
    for(int a = 0; a <= 10; a++)
    {
        sum += a * ncr(10, a)
                * pow(0.7, 10 - a) * pow(0.3, a);
    }
    printf("E[X] = %0.10f\n", sum);
}

int main()
{
    exp_x();
    return 0;
}

```

part 2

variance = $Var(X)$

$$Var(X) = E[X^2] - E[X]^2$$

$$\begin{aligned} E[g(X)] &= \sum_i g(a_i)P(X = a_i) \\ E[X^2] &= \sum_i a_i^2 P(X = a_i) \\ &= \sum_{a=0}^{10} a^2 \binom{10}{a} (0.7)^{10-a} (0.3)^a \\ &\quad /* see calculations below */ \\ &= 11.1 \end{aligned}$$

/ the calculation of $\sum_{a=0}^{10} a^2 \binom{10}{a} (0.7)^{10-a} (0.3)^a$
uses the `exp_x_squared()` function,
which calls the same `fac()`
and `ncr()` described in part 1*/*

```
void exp_x_squared()  
{  
    double sum = 0;  
    for(int a = 0; a <= 10; a++)  
    {  
        sum += a * a * ncr(10, a)  
              * pow(0.7, 10 - a) * pow(0.3, a);  
    }  
    printf("E[X^2] = %0.10f\n", sum);  
}
```

$$\begin{aligned}
 Var(X) &= E[X^2] - E[X]^2 \\
 &= 11.1 - (3)^2 \\
 &= 11.1 - 9 \\
 &= \boxed{2.1}
 \end{aligned}$$

alternatively:

$$\begin{aligned}
 Var(X) &= np(1 - p) \\
 &= (10)(0.3)(1 - 0.3) \\
 &= (3)(0.7) \\
 &= \boxed{2.1}
 \end{aligned}$$