

# Probability and Statistics: Final Exam

## Spring 2018

Name: \_\_\_\_\_

This exam is scheduled for 110 minutes. Notes and other outside materials are not permitted. Non graphing calculators are allowed; if you do not have any, numerical formulas are enough.

**Show all work to receive full credit, except where specified.** The exam is worth 60 points.

Problem Number	Problem Points	Points Earned
MC	10	
TF	5	
FR1	10	
FR2	10	
FR3	10	
FR4	8	
FR5	7	
Total	60	

## Multiple Choice

(2 points each) Circle the correct answer for each question. *You need not justify your answer, but one point of partial credit may be awarded.*

- D** 1 Consider the dataset below:

Handwritten notes for Question 1:  
 $90.25 = 2 + 0.25$   
 $90.75 = 15 + 0.75 \cdot 3$   
 $= 17.25$   
 $1, 2, 3, 3, 4, 7, 15, 18, 21, 25$   
 ~~$1, 21, 3, 18, 4, 7, 15, 3, 18, 2, 25$~~

What is the interquartile range for this dataset?

- (A) 17 (D) 15  
 (B) 18 (E) 16  
 (C) 14

**B**

- 2 Let  $X$  be a normally distributed random variable with mean  $\mu = 90$  and variance  $\sigma^2 = 9$ . What is a good approximation of  $P(85 \leq X \leq 92)$ ?

*Note: Tables for the cumulative distribution function  $\Phi$  of  $N(0, 1)$  are provided at the end of the exam*

- (A) 0.45 (D) 0.84  
 (B) 0.70 (E) 0.77  
 (C) 0.62

Handwritten notes for Question 2:  
 $\Phi\left(\frac{85-90}{3}\right) = \Phi(-1.67) \approx 0.0478$   
 $\frac{2}{3} = 0.67$

**A**

- 3 Let  $X$  be a binomial random variable with parameters  $n = 200$  and  $p = 0.1$ . According to Markov's inequality, what is an upper bound for  $P(X \geq 120)$ ?

- (A)  $\frac{1}{6}$  (D)  $\frac{2}{17}$   
 (B) 0.287 (E) 0.1173  
 (C)  $\frac{1}{8}$

Handwritten note for Question 3:  
 $E[X] = 20$

Handwritten note for Question 3:  
 $20 \geq 120 P(X \geq 120)$

Handwritten note for Question 3:  
 $\frac{20}{120} = \frac{1}{6}$

D

4 Consider two random variables  $X$  and  $Y$  with the following joint probability mass function

		$X$		
		0	1	2
$Y$	0	$\frac{1}{6}$	$\frac{1}{4}$	$\frac{1}{8}$
	1	$\frac{1}{8}$	$\frac{1}{6}$	$\frac{1}{6}$

What is  $E[X^3Y]$ ?

$$\frac{1}{6} + \frac{8}{6} = \frac{9}{6} = \frac{3}{2}$$

(A)  $\frac{1}{2}$

(D)  $\frac{3}{2}$

(B)  $\frac{5}{2}$

(E)  $\frac{7}{4}$

(C)  $\frac{9}{4}$

$$\frac{1}{\lambda} = 22 \quad \lambda = \frac{1}{22}$$

A

5 A continuous random variable  $X$  has an exponential distribution with a mean of 22. What is the variance of  $X$ ?

(A) 484

(D) 0.0020661157

(B) 22

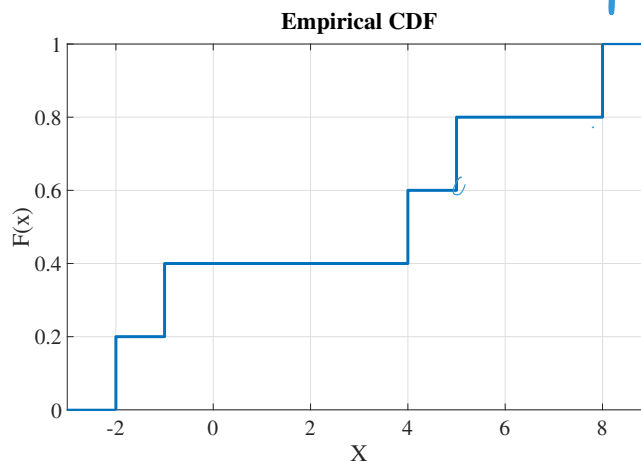
(E) 0.22

(C) 0.0454

## True or False

(1 point each) Indicate whether each statement is true or false. No partial credit will be given.

1 Consider the empirical cumulative distribution function below.



$$F_n(x) = 1 = \frac{\# \leq 8}{n}$$

It corresponds to a dataset with 10 data points.

☒ True

☐ False

T

2 If  $X_1, X_2, \dots, X_n$  is a random sample from a distribution with finite mean  $\mu$  and finite variance  $\sigma^2$ , then

$$S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

where  $\bar{X}_n$  is the sample mean, is ALWAYS an unbiased estimator for the variance  $\sigma^2$ .

☒ True

☐ False

F

3 If  $X_1, X_2, \dots, X_n$  is a random sample from a distribution with finite mean  $\mu$  and finite variance  $\sigma^2$ , then

$$S_n = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2}$$

where  $\bar{X}_n$  is the sample mean, is ALWAYS an unbiased estimator for the standard deviation  $\sigma$ .

☐ True

neg

☐ False

$$0 \leq p(X^2 Y^4) \leq 1$$

T

4 For any two random variables  $X$  and  $Y$ ,  $E[X^2 Y^4]$  is always a positive number.

☐ True

☐ False

F

5 The method of least squares only applies to linear regression.

☐ True

☐ False

## Free Response

Be sure to show all your work neatly and indicate your final answer where appropriate.

1 (10 points) The number  $N_1$  of male customers entering a high end female clothing store per 1-hour time slot is a Poisson process with rate  $\lambda_1 = 1$ , and the number  $N_2$  of female customers entering that same clothing store per 1-hour time slot is a Poisson process with rate  $\lambda_2 = 2$ . In this problem, we will assume that  $N_1$  and  $N_2$  are independent Poisson processes.

(A) (2 Points) In Problem 2 of Homework 9, you showed that the total number of customers  $N = N_1 + N_2$  entering the store in any given 1-hour time slot is also a Poisson process. What is the rate of this Poisson process?

(B) (4 Points) Find the probability of the event " $N_{[0,1]} = 2$  and  $N_{[0,2]} = 5$ ".

(C) (4 Points) If you know that  $N_{[0,1]} = 2$ , what is the probability that  $N_{1[0,1]} = 1$ ?

$$A \quad \lambda_1 = 3$$

$$B. \quad P_{N_{[0,1]}}(2) \cdot P_{N_{[1,2]}}(3)$$

$$= \frac{3^2}{2!} e^{-3} \cdot \frac{3^3}{3!} e^{-3} = \frac{3^4}{4} e^{-6}$$

$$C. \quad P(N_1=1 \mid N_1=2)$$

$$= \frac{P(N_1=1) \cdot P(N_2=1)}{P(N_1=2)}$$

$$= \frac{e^{-1} \cdot 2 \cdot e^{-2}}{3^2 e^{-3}} = \frac{2e^{-3}}{9e^{-3}}$$

$$\frac{2!}{e} = \frac{2}{\frac{4}{9}}$$

- 2 (10 Points) An airline uses an Airbus airplane with 150 seats for a flight between Cleveland and New York City. For this flight, data shows that the probability that a person who bought a ticket actually comes and checks in for the flight is only  $p = 0.75$ . The airline thus decides to sell more tickets  $n$  than seats on the airplane:  $n > 150$ . Let  $Y$  be the random variable corresponding to the number of people who bought a ticket and checked in for the flight.

- (A) (3 Points) What is the exact probability mass function for  $Y$ ? Please write it explicitly.
- (B) (7 Points) Use the central limit theorem to compute an accurate approximation of the number  $n$  of tickets the airline can sell while being 95% sure that all the customers checking in will have a seat on the airplane.

Note: Tables for the cumulative distribution function  $\Phi$  of  $N(0, 1)$  are provided at the end of the exam

(A)  $P_Y(Y) = \binom{n}{y} Y^{0.75} (n-Y)^{0.25}$

(B)  $E[Y] = 0.75n$

$\text{Var}(Y) = 0.1875n$

$P(Y \leq 150) \geq 0.95$

$P\left(\frac{Y - \mu}{\sigma} \leq \frac{150 - 0.75n}{\sqrt{0.1875n}}\right) \geq 0.95$

$\Phi\left(\frac{150 - 0.75n}{\sqrt{0.1875n}}\right) \geq 1.65$

$150 - 0.75n = 1.65 \cdot \sqrt{0.1875n}$   
 $= 0.75 + 1.65 \sqrt{0.1875}$

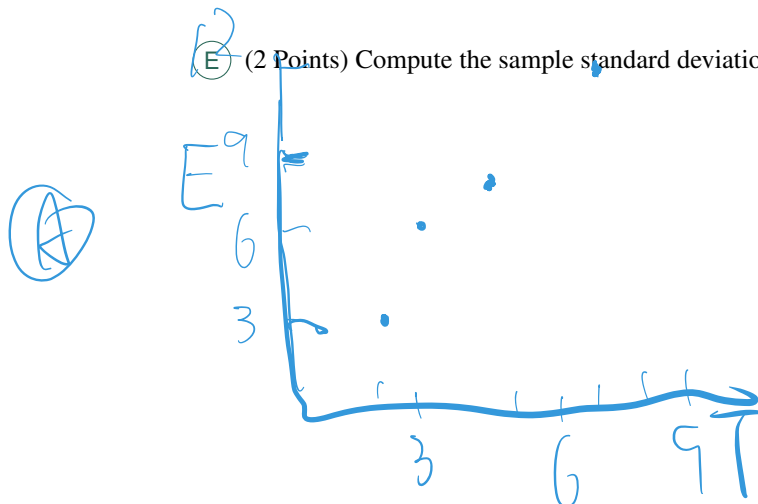


$$n=187$$

- 3 (10 points) A physics student measured the electrical conductivity in piece of copper as a function of temperature, and obtained the following data set, in appropriately nondimensionalized units:

Temperature	2	3	5	7
Electrical conductivity	3	6	7.5	12

- (A) (2 Points) Draw the scatterplot for this data set.
- (B) (4 Points) Compute the best linear fit for the data set.
- (C) (1 Points) Compute the residuals  $r_1, r_2, r_3, r_4$  for this data set, measuring the difference between the data points and the best linear fit.
- (D) (1 Points) Compute the sample mean of the residuals.
- (E) (2 Points) Compute the sample standard deviation of the residuals.



(B)

$$(2, 3), (3, 6), (5, 7.5), (7, 12)$$

$$\beta = \frac{145.5 - 121.125}{87 - \frac{n^2}{4}} = 1.6525$$

$$\bar{x}_n = \frac{17}{4}$$

$$\bar{y}_n = \frac{57}{8}$$

$$\alpha = 0.11$$

$$\hat{y}_n = 1.6525 \bar{x}_n + 0.11$$

$$\textcircled{C} \quad r_1 = 3 - 1.65 \cdot 2 - 0.11 = -0.41$$

$$r_2 = 6 - 1.65 \cdot 3 - 0.11 = 0.94$$

$$r_3 = -0.86$$

$$r_4 = 0.34$$

$$\textcircled{D} \quad \bar{r}_n = \frac{57}{8} - \frac{17}{4} \cdot 1.65 - 0.11$$

$$= \frac{1}{400} \approx 0$$

$$\textcircled{E} \quad \sigma^2 = (0.41)^2 + \dots + (0.34)^2$$

$$\approx 1.91$$

$$\text{Var}(\bar{\sigma}^2) = \frac{1.91}{4} = 0.477$$

$$\sigma = 0.69$$

- 4 (8 points) Suppose  $X_1, X_2, \dots, X_n$  is a random sample from a distribution with probability density function

$$f_X(x) = \begin{cases} \frac{1+\alpha x}{2} & \text{if } -1 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

where  $\alpha$  is the parameter of interest, with  $-1 \leq \alpha \leq 1$ .

- (A) (2 Points) Show that for any integer  $i$  between 1 and  $n$ ,  $E[X_i] = \frac{\alpha}{3}$

- (B) (6 Points) Consider the estimator  $T$  for  $\alpha$  given by  $T = 3 \frac{X_1 + X_2 + \dots + X_n}{n}$ . Compute the bias, variance, and mean squared error of the estimator  $T$ .

$$\text{Var}(T) = 3 \left( \frac{X_1}{n} + \frac{X_2}{n} + \dots + \frac{X_n}{n} \right) = 9 \text{Var} \left[ \frac{X_1}{n} \right]$$

$$\begin{aligned} E[X_i] &= \int_{-1}^1 x f(x) dx \\ &= \frac{1}{2} \int_{-1}^1 (x + \alpha x^2) dx \\ &= \frac{1}{2} \left[ \frac{x^2}{2} + \alpha \cdot \frac{x^3}{3} \right]_{-1}^1 \\ &= \frac{1}{2} \left( \frac{\alpha \cdot 2}{3} \right) = \frac{\alpha}{3} \end{aligned}$$

$$\begin{aligned} E[T] &= E \left[ 3 \cdot \frac{X_1 + \dots + X_n}{n} \right] \\ &= 3 \left( E \left[ \frac{X_1}{n} \right] + \dots + E \left[ \frac{X_n}{n} \right] \right) \\ &= 3 \cdot \frac{E(X_1) + \dots + E(X_n)}{n} \\ &= \frac{3 \cdot n \cdot \frac{\alpha}{3}}{n} \\ &= \alpha \end{aligned}$$

unbiased

$$\text{Var}(T) = \frac{9}{n^2} \text{Var}(X_i)$$

$$= \frac{9}{n^2} \left[ \int_{-1}^1 \left( x^2 - \frac{1+\alpha x}{2} \right)^2 dx - \frac{\alpha^2}{9} \right] \cdot n$$

$$= \frac{9}{n} \left[ \frac{1}{2} \int_{-1}^1 (x^2 + \alpha x^3) dx - \frac{\alpha^2}{9} \right] = \frac{9}{n^2} \left[ \frac{1}{2} \left[ \frac{x^3}{3} + \frac{\alpha x^4}{4} \right]_{-1}^1 - \frac{\alpha^2}{9} \right]$$

$$= \frac{9}{n} \left( \frac{1}{2} \left( \frac{1}{3} + \frac{\alpha}{4} + \frac{1}{3} - \frac{\alpha}{4} \right) - \frac{\alpha^2}{9} \right)$$

$$= \frac{9}{n} \left( \frac{3-\alpha^2}{9} \right)$$

$$= \frac{3-\alpha^2}{n}$$

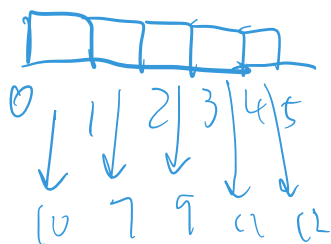
$$\textcircled{c} \text{MSE} = \text{Var}(T) + (E(T) - \theta)^2$$

$$= \frac{3-\alpha^2}{n}$$

- 5 (7 points) A fast food restaurant is trying to predict the number of customers per hour it will have to serve. To do so, it assumes that this number is well modeled by a Poisson distribution, with unknown parameter  $\lambda$ . One day, in order to determine  $\lambda$ , the restaurant manager records the number of customers over a time span of 5 hours, split into separate 1-hour time slots, and finds the following numbers: 10, 7, 9, 11, 12.

What is the maximum likelihood estimate of  $\lambda$ ?

$$P_{N_t} k = \frac{\lambda^k}{k!} e^{-\lambda}$$



$$L(\lambda) = \frac{\lambda^{10}}{10!} e^{-\lambda} \cdot \frac{\lambda^7}{7!} e^{-\lambda} \cdot \frac{\lambda^9}{9!} e^{-\lambda} \cdot \frac{\lambda^{11}}{11!} e^{-\lambda} \cdot \frac{\lambda^{12}}{12!} e^{-\lambda}$$

$$l(\lambda) = \ln(L(\lambda)) = \ln\left(\frac{\lambda^{10}}{10!}\right) - 5\lambda + \ln \frac{\lambda^7}{7!} \dots$$

$$= -5\lambda + \ln(\lambda^{10}) + \ln(10!^{-1})$$

$$= -5\lambda + 49 \ln \lambda + C \leftarrow \text{Constant}$$

$$\frac{dl(\lambda)}{d\lambda} = -5 + \frac{49}{\lambda} = 0$$

$$\lambda = \frac{49}{5}$$



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