

ST/02 Week 19

Analysis of Variance (ANOVA)

One-way ANOVA

Given k independent random samples, i.e. for $j \in [k]$, i.i.d. random sample $\{X_{ij}\}_{i=1}^{n_j}$ from $N(\mu_j, \sigma^2)$. Test:

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k$$

H_1 : not all μ_j 's are the same

Some definitions:

1) The j -th sample mean: $\bar{X}_{\cdot j} := \frac{1}{n_j} \sum_{i=1}^{n_j} X_{ij}$

2) Overall sample mean: $\bar{X} := \frac{1}{n} \sum_{j=1}^k \sum_{i=1}^{n_j} X_{ij} = \frac{1}{n} \sum_{j=1}^k n_j \bar{X}_{\cdot j}$

where $n := \sum_{j=1}^k n_j$.

3) Total variation: $\sum_{j=1}^k \sum_{i=1}^{n_j} (X_{ij} - \bar{X})^2$ with $df = n - 1$

4) Between-groups variation: $B := \sum_{j=1}^k n_j (\bar{X}_{\cdot j} - \bar{X})^2$, $df = k - 1$

5) Within-groups variation (residual sum of squares):

$$W := \sum_{j=1}^k \sum_{i=1}^{n_j} (X_{ij} - \bar{X}_{\cdot j})^2, \quad df = n - k$$

6) Total variation / total sum of squares: $\sum_{j=1}^k \sum_{i=1}^{n_j} (X_{ij} - \bar{X})^2$

7) The ANOVA decomposition: Total variation = $B + W$.

8) Equivalent forms: Total variation = $\sum_{j=1}^k \sum_{i=1}^{n_j} X_{ij}^2 - n\bar{X}^2$

$$B = \sum_{j=1}^k n_j \bar{X}_{\cdot j}^2 - n\bar{X}^2$$

$$W = \sum_{j=1}^k \sum_{i=1}^{n_j} X_{ij}^2 - \sum_{j=1}^k n_j \bar{X}_{\cdot j}^2$$

$$= \sum_{j=1}^k (n_j - 1) S_j^2$$

Under H_0 , it can be shown that:

$$F := \frac{\sum_{j=1}^k n_j (\bar{X}_{\cdot j} - \bar{X})^2 / (k-1)}{\sum_{j=1}^k \sum_{i=1}^{n_j} (X_{ij} - \bar{X}_{\cdot j})^2 / (n-k)} = \frac{B / (k-1)}{W / (n-k)} \sim F_{k-1, n-k}$$

Reject H_0 if $f > F_{k-1, n-k, \alpha}$ ($F_{\alpha; k-1, n-k}$).

One-way ANOVA table:

Source	DF	SS	MS	F	p-value
Factor	$k-1$	B	$B/(k-1)$	$\frac{B/(k-1)}{W/(n-k)}$	p
Error	$n-k$	W	$W/(n-k)$		
Total	$n-1$	$B+W$			

Without further assumptions, we also have:

1) estimator of σ^2 : $\hat{\sigma}^2 := S := \sqrt{\frac{W}{n-k}}$

2) $\alpha \cdot 100\%$ - confidence interval for μ_j :

$$\bar{X}_{\cdot j} \pm t_{\frac{\alpha}{2}, n-k} \cdot \frac{S}{\sqrt{n_j}} \quad \text{for } j \in [k]$$

A system view of one-way ANOVA

$$X_{ij} = \mu + \beta_j + \varepsilon_{ij} \quad \text{for } i \in [n_j] \text{ and } j \in [k],$$

with $\varepsilon_{ij} \sim N(0, \sigma^2)$, all independent.

$\sum_j \beta_j = 0$ is required for identifiability.

μ is the average effect and β_j the j -th level treatment effect.

$$H_0: \beta_1 = \beta_2 = \dots = \beta_k = 0$$

Two-way ANOVA

$$X_{ij} = \mu + \delta_i + \beta_j + \varepsilon_{ij} \quad \text{for } i \in [r] \text{ and } j \in [c].$$

μ : average treatment effect

β_j : treatment (column) levels

δ_i : different block (row) levels

ε_{ij} : follow $N(0, \sigma^2)$, all independent

Conditions for identifiability: $\sum_i \gamma_i = \sum_j \beta_j = 0$.

Two directions

↙
 $H_0: \beta_1 = \beta_2 = \dots = \beta_c = 0$

↘
 $H_0: \gamma_1 = \gamma_2 = \dots = \gamma_r = 0$

Some definitions:

1) i -th block sample mean: $\bar{X}_{i\cdot} := \frac{1}{c} \sum_{j=1}^c X_{ij}$

2) j -th treatment sample mean: $\bar{X}_{\cdot j} := \frac{1}{r} \sum_{i=1}^r X_{ij}$

3) overall sample mean: $\bar{X} := \frac{1}{n} \sum_{j=1}^c \sum_{i=1}^r X_{ij}$

4) total variation: Total SS := $\sum_{i=1}^r \sum_{j=1}^c (X_{ij} - \bar{X})^2$, $df = rc - 1$

5) between-blocks variation: $B_{row} := c \sum_{i=1}^r (\bar{X}_{i\cdot} - \bar{X})^2$, $df = r - 1$

6) between-treatments variation: $B_{col} := r \sum_{j=1}^c (\bar{X}_{\cdot j} - \bar{X})^2$, $df = c - 1$

7) residual sum of squares: Residual SS := $\sum_{i=1}^r \sum_{j=1}^c (X_{ij} - \bar{X}_{i\cdot} - \bar{X}_{\cdot j} + \bar{X})^2$
 $df = (r-1)(c-1)$

8) two-way ANOVA decomposition: Total SS = $B_{row} + B_{col} + \text{Residual SS}$

Equivalent forms:
$$\begin{aligned} \text{Total SS} &= \sum_{i=1}^r \sum_{j=1}^c X_{ij}^2 - rc \bar{X}^2 \\ B_{row} &= c \sum_{i=1}^r \bar{X}_{i.}^2 - rc \bar{X}^2 \\ B_{col} &= r \sum_{j=1}^c \bar{X}_{.j}^2 - rc \bar{X}^2 \\ \text{Residual SS} &= \text{Total SS} - B_{row} - B_{col} \end{aligned}$$

Case I: $H_0: \sigma_1 = \sigma_2 = \dots = \sigma_r = 0$

$$F := \frac{B_{row} / (r-1)}{\text{Residual SS} / [(r-1)(c-1)]} = \frac{(c-1)B_{row}}{\text{Residual SS}}$$

Under H_0 , $F \sim F_{r-1, (r-1)(c-1)}$

reject H_0 if $f > F_{\alpha; r-1, (r-1)(c-1)}$

Case II: $H_0: \beta_1 = \beta_2 = \dots = \beta_c = 0$

$$F := \frac{B_{col} / (c-1)}{\text{Residual SS} / [(r-1)(c-1)]} = \frac{(r-1)B_{col}}{\text{Residual SS}}$$

Under H_0 , $F \sim F_{c-1, (r-1)(c-1)}$.

Reject H_0 if $f > F_{\alpha; c-1, (r-1)(c-1)}$

Two-way ANOVA table

Source	DF	SS	MS	F	p-value
Row factor	$r-1$	B_{row}	$B_{row} / (r-1)$	$(c-1)B_{row} / \text{RSS}$	p_1
Column factor	$c-1$	B_{col}	$B_{col} / (c-1)$	$(r-1)B_{col} / \text{RSS}$	p_2
Residual	$(r-1)(c-1)$	Residual SS	$\frac{\text{Residual SS}}{(r-1)(c-1)}$		
Total	$rc-1$	Total SS			

Remark on residuals:

By model we have $X_{ij} = \mu + \alpha_i + \beta_j + \varepsilon_{ij}$,

While from data we have the decomposition:

$$X_{ij} = \bar{X} + (\bar{X}_{i.} - \bar{X}) + (\bar{X}_{.j} - \bar{X}) + (X_{ij} - \bar{X}_{i.} - \bar{X}_{.j} + \bar{X})$$

$$\Rightarrow \hat{\mu} = \bar{X}$$

$$\hat{\alpha}_i = \bar{X}_{i.} - \bar{X}$$

$$\hat{\beta}_j = \bar{X}_{.j} - \bar{X}$$

$$\hat{\varepsilon}_{ij} = X_{ij} - \bar{X}_{i.} - \bar{X}_{.j} + \bar{X}$$