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# Statistics 202: Data Mining

Linear Discriminant Analysis

Based in part on slides from textbook, slides of Susan Holmes

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#### Nearest centroid rule

- Suppose we break down our data matrix as by the labels yielding  $(\mathbf{X}^j)_{1 < j < k}$  with sizes  $\operatorname{nrow}(\mathbf{X}^j) = n_j$ .
- A simple rule for classification is:

Assign a new observation with features  $\boldsymbol{x}$  to

$$\widehat{f}(\boldsymbol{x}) = \operatorname*{argmin}_{1 \le j \le k} d(\boldsymbol{x}, \boldsymbol{X}^j)$$

• What do we mean by distance here?

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#### Nearest centroid rule

- If we can assign a central point or centroid  $\widehat{\mu}_j$  to each  $\mathbf{X}^j$ , then we can define the distance above as distance to the centroid  $\widehat{\mu}_j$ .
- This yields the nearest centroid rule

$$\widehat{f}(\mathbf{x}) = \operatorname*{argmin}_{1 \le j \le k} d(\mathbf{x}, \widehat{\mu}_j)$$

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#### Nearest centroid rule

• This rule is described completely by the functions

$$h_{ij}(\mathbf{x}) = \frac{d(\mathbf{x}, \widehat{\mu}_j)}{d(\mathbf{x}, \widehat{\mu}_i)}$$

with  $\widehat{f}(\mathbf{x})$  being any i such that

$$h_{ij}(\mathbf{x}) \geq 1 \ \forall j.$$

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#### Nearest centroid rule

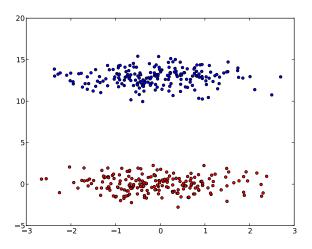
• If  $d(x,y) = ||x-y||_2$  then the natural centroid is

$$\widehat{\mu}_j = \frac{1}{n_j} \sum_{l=1}^{n_j} \boldsymbol{X}_l^j$$

- This rule just classifies points to the nearest  $\hat{\mu}_i$ .
- But, if the covariance matrix of our data is not I, this rule ignores structure in the data . . .

# Why we should use covariance

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#### Choice of distance

- Often, there is some background *model* for our data that is equivalent to a given procedure.
- For this nearest centroid rule, using the Euclidean distance effectively assumes that within the set of points  $X^j$ , the rows are multivariate Gaussian with covariance matrix proportional to I.
- It also implicitly assumes that the  $n_j$ 's are roughly equal.
- For instance, if one  $n_j$  was very small just because a data point is close to that  $\widehat{\mu}_j$  doesn't necessarily mean that we should conclude it has label j because there might be a huge number of points of label i near that  $\widehat{\mu}_i$ .

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#### Gaussian discriminant functions

- Suppose each group with label j had its own mean  $\mu_j$  and covariance matrix  $\Sigma_j$ , as well as proportion  $\pi_j$ .
- The Gaussian discriminant functions are defined as

$$h_{ij}(\mathbf{x}) = h_i(\mathbf{x}) - h_j(\mathbf{x})$$
  
$$h_i(\mathbf{x}) = \log \pi_i - \log |\Sigma_i|/2 - (\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i)/2$$

- The first term weights the prior probability, the second two terms are  $\log \phi_{\mu_i, \Sigma_i}(\mathbf{x}) = \log L(\mu_i, \Sigma_i | \mathbf{x})$ .
- Ignoring the first two terms, the  $h_i$ 's are essentially within-group Mahalanobis distances . . .

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#### Gaussian discriminant functions

- The classifier assigns x label i if  $h_i(x) \ge h_j(x) \forall j$ .
- Or,

$$f(\mathbf{x}) = \operatorname*{argmax}_{1 \leq i \leq k} h_i(\mathbf{x})$$

- This is equivalent to a Bayesian rule. We'll see more Bayesian rules when we talk about naïve Bayes . . .
- When all  $\Sigma_i$  and  $\pi_i$ 's are identical, the classifier is just nearest centroid using Mahalanobis distance instead of Euclidean distance.

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#### Estimating discriminant functions

- In practice, we will have to estimate  $\pi_j, \mu_j, \Sigma_j$ .
- Obvious estimates:

$$\widehat{\pi}_{j} = \frac{n_{j}}{\sum_{j=1}^{k} n_{j}}$$

$$\widehat{\mu}_{j} = \frac{1}{n_{j}} \sum_{l=1}^{n_{j}} \boldsymbol{X}_{l}^{j}$$

$$\widehat{\Sigma}_{j} = \frac{1}{n_{j} - 1} (\boldsymbol{X}^{j} - \widehat{\mu}_{j} \mathbf{1})^{T} (\boldsymbol{X}^{j} - \widehat{\mu}_{j} \mathbf{1})$$

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#### Estimating discriminant functions

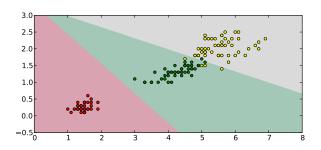
• If we assume that the covariance matrix is the same within groups, then we might also form the pooled estimate

$$\widehat{\Sigma}_{P} = \frac{\sum_{j=1}^{k} (n_j - 1) \widehat{\Sigma}_j}{\sum_{j=1}^{k} n_j - 1}$$

- If we use the pooled estimate  $\Sigma_j = \widehat{\Sigma}_P$  and plug these into the Gaussian discrimants, the functions  $h_{ij}(\mathbf{x})$  are *linear* (or affine) functions of  $\mathbf{x}$ .
- This is called Linear Discriminant Analysis (LDA).
- Not to be confused with the other LDA (Latent Dirichlet Allocation) . . .

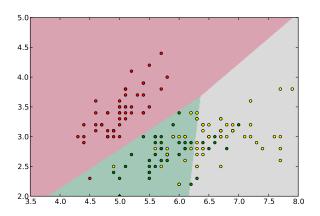
# Linear Discriminant Analysis using (petal.width, petal.length)

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# Linear Discriminant Analysis using (sepal.width, sepal.length)

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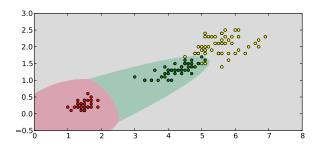
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#### Quadratic Discriminant Analysis

- If we use don't use pooled estimate  $\Sigma_j = \widehat{\Sigma}_j$  and plug these into the Gaussian discrimants, the functions  $h_{ij}(\mathbf{x})$  are *quadratic* functions of  $\mathbf{x}$ .
- This is called Quadratic Discriminant Analysis (QDA).

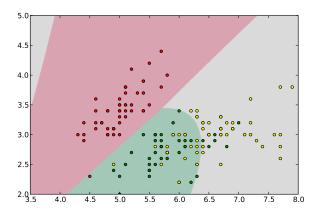
# Quadratic Discriminant Analysis using (petal.width, petal.length)

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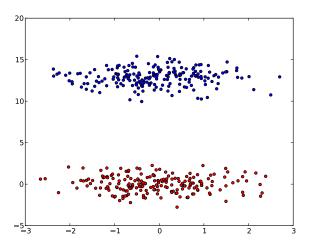
# Quadratic Discriminant Analysis using (sepal.width, sepal.length)

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#### Motivation for Fisher's rule

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#### Fisher's discriminant function

- Fisher proposed to classify using a linear rule.
- He first decomposed

$$(\mathbf{X} - \widehat{\mu}\mathbf{1})^T(\mathbf{X} - \widehat{\mu}\mathbf{1}) = \widehat{SS}_B + \widehat{SS}_W$$

Then, he proposed,

$$\widehat{v} = \operatorname{argmax}_{v:v^T \widehat{SS}_W v = 1} v^T \widehat{SS}_B v$$

- Having found  $\hat{v}$ , form  $\mathbf{X}^{j}\hat{v}$  and the centroid  $\eta_{j} = \text{mean}(\mathbf{X}^{j}\hat{v})$
- In the two-class problem k = 2, this is the same as LDA.

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#### Fisher's discriminant functions

- The direction  $\hat{v}_1$  is an eigenvector of some matrix. There are others, up to k-2 more.
- Suppose we find all k-1 vectors and form  $\mathbf{X}_j V^T$ , each one an  $n_i \times (k-1)$  matrix with centroid  $\eta_i \in \mathbb{R}^{k-1}$ .
- The matrix V determines a map from  $\mathbb{R}^p$  to  $\mathbb{R}^{k-1}$ , so given a new data point we can compute  $V\mathbf{x} \in \mathbb{R}^{k-1}$ .
- This gives rise to a classifier

$$\widehat{f}(\mathbf{x}) = \text{nearest centroid}(V\mathbf{x})$$

• This is LDA (assuming  $\pi_j = \frac{1}{k}$ ) ...

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#### Discriminant models in general

A discriminant model is generally a model that estimates

$$P(Y = j | \mathbf{x}), 1 \le j \le k$$

- That is, given that the features I observe are x, the probability I think this label is j ...
- LDA and QDA are actually generative models since they specify

$$P(X = \boldsymbol{x} | Y = j).$$

There are lots of discriminant models . . .

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#### Logistic regression

- The logistic regression model is ubiquitous in binary classification (two-class) problems
- Model:

$$P(Y = 1|\mathbf{x}) = \frac{\alpha + e^{\mathbf{x}^T \beta}}{1 + e^{\alpha + \mathbf{x}^T \beta}} = \pi(\alpha, \beta, \mathbf{x})$$

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#### Logistic regression

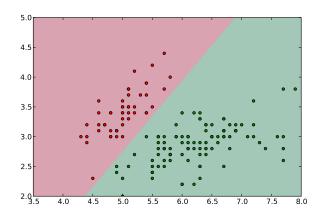
- Software that fits a logistic regression model produces an estimate of  $\beta$  based on a data matrix  $\boldsymbol{X}_{n \times p}$  and binary labels  $\boldsymbol{Y}_{n \times 1} \in \{0,1\}^n$
- It fits the model minimizing what we call the deviance

$$\mathsf{DEV}(\beta) = -2\sum_{i=1}^{n} (\boldsymbol{Y}_{i} \log \pi(\beta, \boldsymbol{X}_{i}) + (1 - \boldsymbol{Y}_{i}) \log(1 - \pi(\beta, \boldsymbol{X}_{i}))$$

- While not immediately obvious, this is a convex minimization problem, hence is fairly easy to solve.
- Unlike trees, the convexity yields a globally optimal solution.

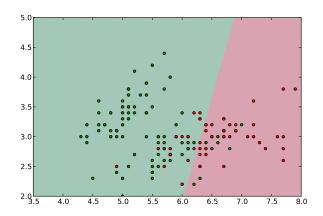
# Logistic regression, setosa vs virginica, versicolor using (sepal.width, sepal.length)

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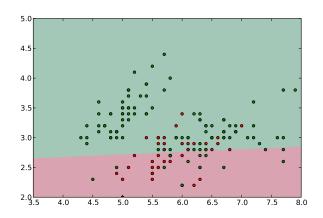
# Logistic regression, *virginica* vs *setosa*, *versicolor* using (sepal.width, sepal.length)

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# Logistic regression, *versicolor* vs *setosa*, *virginica* using (sepal.width, sepal.length)

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#### Logistic regression

Logistic regression produces an estimate of

$$P(\mathbf{y}=1|\mathbf{x})=\widehat{\pi}(\mathbf{x})$$

- Typically, we classify as 1 if  $\hat{\pi}(\mathbf{x}) > 0.5$ .
- This yields a 2 × 2 confusion matrix

	Predicted: 0	Predicted: 1
Actual: 0	TN	FP
Actual: 1	FN	TP

• From the  $2 \times 2$  confusion matrix, we can compute Sensitivity, Specificity, etc.

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#### Logistic regression

- However, we could choose the threshold differently, perhaps related to estimates of the prior probabilities of 0's and 1's
- Now, each threshold  $0 \le t \le 1$  yields a new confusion matrix
- This yields a 2 × 2 confusion matrix

	Predicted: 0	Predicted: 1
Actual: 0	TN(t)	FP(t)
Actual : 1	FN(t)	TP(t)

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#### ROC curve

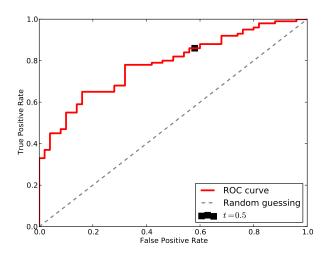
- Generally speaking, we prefer classifiers that are both highly sensitive and highly specific.
- These confusion matrices can be summarized using an ROC (Receiver Operating Characteristic) curve.
- This is a plot of the curve

$$(1 - \mathsf{Specificity}(t), \mathsf{Sensitivity}(t))_{0 \le t \le 1}$$

- Often, Specificity is referred to as TNR (True Negative Rate), and (1 Specificity(t)) as FPR.
- Often, Sensitivity is referred to as TPR (True Positive Rate), and (1 Sensitivity(t)) as FNR.

#### AUC: Area under ROC curve

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#### **ROC** curve

- Points in the upper left of the ROC curve are good.
- Any point on the diagonal line represents a classifier that is guessing randomly.
- A tree classifier as we've discussed seems to correspond to only one point in the ROC plot.
- But one can estimate probabilities based on frequencies in terminal nodes.

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#### ROC curve

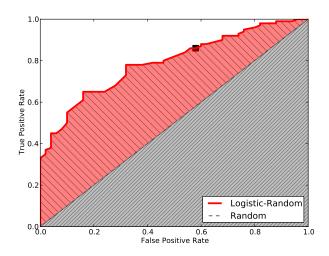
A common numeric summary of the ROC curve is

AUC(ROC curve) = Area under ROC curve.

- Can be interpreted as an estimate of the probability that the classifier will give a random positive instance a higher score than a random negative instance.
- Maximum value is 1.
- For a random guesser, AUC is 0.5

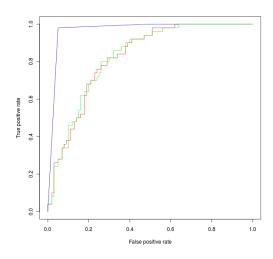
#### AUC: Area under ROC curve

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## ROC curve: logistic, rpart, lda

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