

# Assignment 1

Yun Siyu, 17307110448

2020 年 3 月 25 日

## 1 Problem 1

### 1.1 MLE

According to the MLE

$$\begin{aligned} L(\theta) &= \prod_{i=1}^{10} f(x_i|\theta) \\ &= \left(\frac{2\theta}{3}\right)^2 \cdot \left(\frac{\theta}{3}\right)^3 \cdot \left(\frac{2(1-\theta)}{3}\right)^3 \cdot \left(\frac{1-\theta}{3}\right)^3 \end{aligned} \quad (1)$$

So we can get

$$\begin{aligned} l(\theta) &= \log L(\theta) \\ &= 2\ln \frac{2\theta}{3} + 3\ln \frac{\theta}{3} + 3\ln \frac{2(1-\theta)}{3} + 2\ln \frac{1-\theta}{3} \end{aligned} \quad (2)$$

Therefore, the score function is,

$$\begin{aligned} U(\theta) &= \frac{dl(\theta)}{d\theta} \\ &= \frac{5}{\theta} - \frac{5}{1-\theta} \end{aligned} \quad (3)$$

We know that,

$$\frac{dU(\theta)}{d\theta} = -\frac{5}{\theta^2} - \frac{5}{(1-\theta)^2} < 0 \quad (4)$$

So we can make  $U(\theta) = 0$ . Thus maximum likelihood estimator of  $\theta$  is,

$$\hat{\theta}_{MLE} = \frac{1}{2}$$

### 1.2 MME

According to the MME

$$\frac{1}{10} \sum_{i=1}^{10} x_i = 1.5 \quad (5)$$

$$\begin{aligned} E(X) &= \sum_{i=0}^3 iP(X=i) \\ &= \frac{7-6\theta}{3} \end{aligned} \quad (6)$$

We know that,

$$\frac{1}{10} \sum_{i=1}^{10} x_i = E(X) \quad (7)$$

So we can calculate method of moments estimator of  $\theta$ ,

$$\hat{\theta}_{MME} = \frac{5}{12}$$

## 2 Problem 2

### 2.1 MLE

According to the MLE

$$\begin{aligned} L(\sigma) &= \prod_{i=1}^n f(X_i|\sigma) \\ &= \prod_{i=1}^n \frac{1}{2\sigma} e^{-\frac{|X_i|}{\sigma}} \\ &= \frac{1}{(2\sigma)^n} e^{-\frac{\sum_{i=1}^n |X_i|}{\sigma}} \end{aligned} \quad (8)$$

So we can get

$$\begin{aligned} l(\sigma) &= \log L(\sigma) \\ &= -n \ln 2\sigma - \frac{\sum_{i=1}^n |X_i|}{\sigma} \end{aligned} \quad (9)$$

Therefore, the score function is,

$$\begin{aligned} U(\theta) &= \frac{dl(\theta)}{d\theta} \\ &= -\frac{n}{\sigma} + \frac{\sum_{i=1}^n |X_i|}{\sigma^2} \end{aligned} \quad (10)$$

So we can make  $U(\theta) = 0$ . Thus maximum likelihood estimator of  $\theta$  is,

$$\hat{\sigma}_{MLE} = \frac{1}{n} \sum_{i=1}^n |X_i|$$

### 2.2 MME

We can calculate that,

$$\begin{aligned} E(X) &= \int_{-\infty}^{\infty} \frac{1}{2\sigma} x e^{-\frac{|x|}{\sigma}} dx \\ &= 0 \end{aligned} \quad (11)$$

So we need to use secondary moment,

$$\begin{aligned} E(X^2) &= \int_{-\infty}^{\infty} \frac{1}{2\sigma} x^2 e^{-\frac{|x|}{\sigma}} dx \\ &= 2\sigma^2 \end{aligned} \quad (12)$$

We know that,

$$\frac{1}{n} \sum_{i=1}^n X_i^2 = E(X) \quad (13)$$

So we can calculate method of moments estimator of  $\theta$ ,

$$\hat{\sigma}_{MME} = \sqrt{\frac{1}{2n} \sum_{i=1}^n X_i^2}$$

## 3 Problem 3

### 3.1 MLE

According to the MLE

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n f(x_i|\theta) \\ &= \left(\frac{1}{\theta}\right)^n \end{aligned} \quad (14)$$

It is obviously that  $L(\theta)$  is a monotonically decreasing function. Therefore, if we want to maximize  $L(\theta)$ , we need to minimize  $\theta$ . We know that  $\theta$  is a parameter from a uniform distribution and  $X_1, \dots, X_n$  form a random sample from the distribution, so we can get,

$$\theta \geq \max\{X_1, \dots, X_n\}$$

Thus maximum likelihood estimator of  $\theta$  is,

$$\hat{\theta}_{MLE} = \max\{X_1, \dots, X_n\}$$

### 3.2 MME

According to the MME

$$\begin{aligned} E(X) &= \int_0^\theta \frac{x}{\theta} dx \\ &= \frac{\theta}{2} \end{aligned} \tag{15}$$

We know that,

$$\frac{1}{n} \sum_{i=1}^n X_i = E(X) \tag{16}$$

So we can calculate method of moments estimator of  $\theta$ ,

$$\hat{\theta}_{MME} = \frac{2}{n} \sum_{i=1}^n X_i$$

## 4 Problem 4

### 4.1 Central Limit Theorem

We know that,  $X_n$  can be decomposed into  $n$  independent random variables that obey the Bernoulli distribution with the parameter  $p$ ,

$$X_n = x_1 + \dots + x_n, x_i \sim b(1, p)$$

Therefore, our target can be changed as,

$$\frac{X_n}{n} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}$$

According to the Bernoulli distribution

$$E(x_i) = p$$

$$Var(x_i) = p(1 - p)$$

So, we can get,

$$E(\bar{x}) = p$$

$$Var(\bar{x}) = \frac{p(1 - p)}{n}$$

According to the Central Limit Theorem,

$$\begin{aligned} \frac{\bar{x} - p}{\sqrt{\frac{p(1-p)}{n}}} &\sim N(0, 1) \\ \bar{x} &\sim N\left(p, \frac{p(1-p)}{n}\right) \end{aligned} \tag{17}$$

Therefore,

$$\frac{X_n}{n} \sim N\left(p, \frac{p(1-p)}{n}\right)$$

## 4.2 Delta Method

According to the Delta Method,

$$\mu = p$$

$$\begin{aligned} g(\mu) &= g(p) \\ &= p(1-p) \end{aligned} \tag{18}$$

$$\begin{aligned} g(\bar{x}) &= g\left(\frac{X_n}{n}\right) \\ &= \delta_n \end{aligned} \tag{19}$$

We know that  $p \neq \frac{1}{2}$ , so,

$$g'(p) = 1 - 2p \neq 0$$

Therefore, according to the usual first order Delta method states,

$$\delta_n \sim N(p(1-p), \frac{p(1-p)(1-2p)^2}{n})$$

## 5 Problem 5

According to the Bernoulli distribution

$$\begin{aligned} E(x_i) &= \mu \\ Var(x_i) &= \sigma^2 \end{aligned}$$

So, we can get,

$$\begin{aligned} E(\bar{X}_n) &= \mu \\ Var(\bar{X}_n) &= \frac{\sigma^2}{n} \end{aligned}$$

According to the Central Limit Theorem,

$$\sqrt{n}(\bar{X}_n - \mu) \sim N(0, \sigma^2) \tag{20}$$

According to the Delta Method,

$$\begin{aligned} g(\mu) &= \mu^2 \\ g(\bar{X}_n) &= \bar{X}_n^2 \\ g'(\mu) &= 2\mu \end{aligned}$$

if  $\mu \neq 0$ , according to the usual first order Delta method states,

$$\bar{X}_n^2 \sim N(\mu^2, \frac{4\sigma^2\mu^2}{n})$$

if  $\mu = 0$ , the second derivative,

$$g'' = 2 \neq 0$$

Therefore, According to the usual second order Delta method states,

$$\begin{aligned} n(\bar{X}_n^2 - \mu^2) &\sim \sigma^2 \chi_1^2 \\ \bar{X}_n^2 &\sim \frac{\sigma^2}{n} \chi_1^2 + \mu^2 \end{aligned}$$

## 6 Problem 6

According to the problem 4,

$$\frac{\hat{p} - p}{\sqrt{\frac{p(1-p)}{n}}} \sim N(0, 1)$$

Using the asymptotic normal approximation,

$$\frac{\hat{p} - p}{\sqrt{\frac{\hat{p}(1-\hat{p})}{n}}} \sim N(0, 1)$$

Therefore,  $100(1-\alpha)\%$  confidence interval for  $p$ ,

$$[\hat{p} - z_{1-\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}, \hat{p} + z_{1-\frac{\alpha}{2}} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}]$$

We know that,  $\hat{p} = 0.2$ ,  $\alpha = 0.05$ , so the 95% confidence interval for  $p$ ,

$$[0.103, 0.297]$$