Bootstrap For Complex Data

Mark M. Fredrickson (mfredric@umich.edu)

Computational Methods in Statistics and Data Science (Stats 406)

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- Use the bootstrap distribution to form confidence intervals (several methods)

Example: Student academic test scores



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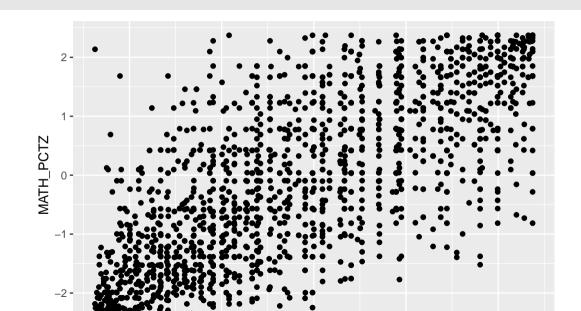
Differences between Hispanic and non-Hispanic families in social capital and child development: First-year findings from an experimental study

Adam Gamoran a,*, Ruth N. López Turley b, Alyn Turner a, Rachel Fish a

^a University of Wisconsin-Madison, , United States
^b Rice University, , United States

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Reading Test Scores



Test Score Correlation

Confidence Intervals

```
> boot.ci(cor_boot, type = "basic")
BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
Based on 100 bootstrap replicates
CALL:
boot.ci(boot.out = cor_boot, type = "basic")
Intervals :
Level Basic
95% (0.6924, 0.7443)
Calculations and Intervals on Original Scale
Some basic intervals may be unstable
```

Bootstrap for Dependence and

Structure

Bootstrap basics

Recall our usual setup for the bootstrap:

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We'll consider two cases:

- Bootstrap for time series (stochastic processes)
- Stratified bootstrap for different groups

Time series models

Suppose we are interested in the following **stochastic process**:

$$X(t) = \rho X(t-1) + e(t), \quad e(t) \sim F, E(e(t)) = 0, \forall t$$

and we are interested in estimating ρ .

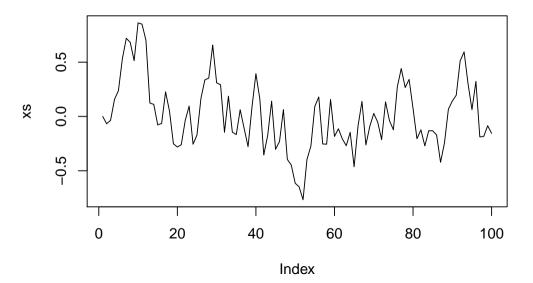
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and we are interested in estimating ρ . Here is some simulated example data:

```
> rho <- 0.75
> k <- 100
> xs <- numeric(k)
> xs[1] <- 0
> for (i in 2:k) {
+     xs[i] <- rho * xs[i - 1] + (rbeta(1, 2, 2) - 0.5)
+ }</pre>
```



Estimating ρ

```
One way we can estimate \rho is to perform OLS on X(t) given X(t-1):
> (est <- lm(xs[2:k] ~ xs[1:(k - 1)] - 1)) # no intercept term
Call:
lm(formula = xs[2:k] ~ xs[1:(k-1)] - 1)
Coefficients:
xs[1:(k-1)]
         0.71
```

What is the variation? Can we get bootstrap confidence intervals?

Parametric bootstrap

Here we need **bootstrap samples with dependence**:

$$X(0)^*, X(1)^*, \dots, X(t)^*$$
 such that: $X(s)^* = \rho X(s-1)^* + e(s)^*$

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We have already seen parametric bootstrapping in which we use a model to generate the bootstrap samples.

Idea: starting from the observed X(0), draw from the **estimated distribution of** e to create a **bootstrap series**:

$$X(1)^* = \hat{
ho}X(0) + e(1)^*, \dots, X(t)^* = \hat{
ho}X(t-1)^* + e(t)^*$$

```
> rhohat <- coef(est)</pre>
> et_hat <- xs[2:k] - predict(est) # estimate residuals
> est_rho_stat <- function(x, index) {</pre>
+ estar <- x[index]
   new series <- numeric(k)</pre>
   new_series[1] <- xs[1] # starts at the same point
+ for (i in 2:k) {
     new series[i] <- rhohat * new series[i - 1] + estar[i]
   7
   coef(lm(new_series[2:k] \sim new_series[1:(k-1)] - 1))
+ }
```

Cls for ρ

```
> boot_rho <- boot(et_hat, est_rho_stat, R = 1000)</pre>
> boot.ci(boot_rho, type = "perc")
BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
Based on 1000 bootstrap replicates
CALL:
boot.ci(boot.out = boot_rho, type = "perc")
Intervals:
Level Percentile
95% (0.5263, 0.8143)
Calculations and Intervals on Original Scale
```

Two Sample Problems

In the previous example, all the e(t) were identical, but the X(t) were not independent.

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Two sample problems consider the situation of having two independent samples

$$X_1,\ldots,X_n \stackrel{\text{iid}}{\sim} F \quad Y_1,\ldots,Y_m \stackrel{\text{iid}}{\sim} G$$

and estimating quantities such as the difference of means:

$$\Delta = \mathsf{E}(X) - \mathsf{E}(Y)$$

(notice: if F = G, $\Delta = 0$, though not necessarily the converse.)

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In fact, we might be interested in knowing if the two populations have the same distribution (or same mean, etc). In particular, we'll focus on $\Delta = E\left(\text{San Antonio}\right) - E\left(\text{Phoenix}\right).$

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In fact, we might be interested in knowing if the two populations have the same distribution (or same mean, etc). In particular, we'll focus on $\Delta = E$ (San Antonio) - E (Phoenix).

We'll bootstrap in the two groups separately to get estimates of means and then combine them.

Estimating the difference of means

The boot function has a strata (groups) argument we can use.

```
> mean_diff <- function(x, index) {</pre>
      xstar <- x[index, ] # boot will handle stratification for us
      mean(xstar$READ PCTZ[xstar$PH.AZ], na.rm = TRUE) -
          mean(xstar$READ PCTZ[!xstar$PH.AZ]. na.rm = TRUE)
+ }
> gam.boot <- boot(gamoran.</pre>
                    statistic = mean diff.
+
                    strata = gamoran$PH.AZ,
                   R = 1000
```

CI

```
> (gbci <- boot.ci(gam.boot, type = "basic"))</pre>
BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
Based on 1000 bootstrap replicates
CALL:
boot.ci(boot.out = gam.boot, type = "basic")
Intervals:
Level Basic
95% (0.1398, 0.3782)
Calculations and Intervals on Original Scale
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We saw that the 95% confidence interval for \triangle did not include zero:

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Additionally, we can reject any hypothesis that says that San Antonio has a higher average reading score. (All at $\alpha=0.05$ level).

Let's be clear that this is not a **causal conclusion**, merely that on average San Antonio students tend to score lower.

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- For dependent data, generate samples following dependence.
- For structured data, incorporate the structure.