

CSC380: Principles of Data Science

Probability 2

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Summary: calculating probabilities

If we know that all outcomes are equally likely, we can use

We will use combinatorics to do counting

$$P(E) = \frac{|E|}{|S|}$$
 Number of elements in event set Number of possible outcomes (e.g. 36)

- If |E| is hard to calculate directly, we can try
 - the rules of probability
 - the Law of Total Probability, using an appropriate partition of sample space S

Rules of probability

To recap and summarize:

Rules of Probability

- 1. Non-negativity: All probabilities are between 0 and 1 (inclusive)
- **2.** Unity of the sample space: P(S) = 1
- 3. Complement Rule: $P(E^C) = 1 P(E)$
- 4. Probability of Unions:
 - (a) In general, $P(E \cup F) = P(E) + P(F) P(E \cap F)$
 - (b) If E and F are disjoint, then $P(E \cup F) = P(E) + P(F)$

Overview

- Conditional probability
- Probabilistic reasoning
 - contingency table
 - probability trees
- Bayes rule
- Independence of events
- Probability and combinatorics

Conditional Probability

Example: Seat Belts

		C		
		Buck.	Unbuck.	Marginal
Parent	Buck.	0.48	0.12	0.60
	Unbuck.	0.10	0.30	0.40
	Marginal	0.58	0.42	1.00

Table: Probability Estimates for Seat Belt Status

Suppose we pick a family from US at random:

- What is the probability of the event "Child is Buckled"?
- What should our new estimate be if we know that "Parent is Buckled"?

Example: blood types

		Antigen B		
		Absent	Present	Marginal
Antigen A	Absent	0.44	0.10	0.54
	Present	0.42	0.04	0.46
	Marginal	0.86	0.14	1.00

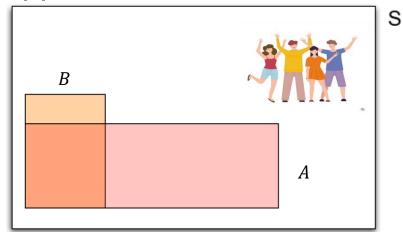
Table: Probability Estimates for U.S. Blood Types

- A: "presence of antigen A", and B: "presence of antigen B"
- Suppose someone of an unknown blood type gets a test that reveals the presence of antigen A.
 - What is the chance that event A happens to them now?
 - What is the chance that event B happens to them now?

Relative area

• A: antigen A present B: antigen B present

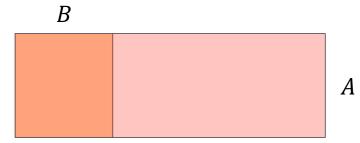
Given that A happens, what is the chance of B happening?



- Another way to think about this:
 - Restricted to people with antigen A present, what is the fraction of those people with antigen B?

Relative area

Let's zoom into people with antigen A present.



- It's just as if the sample space had shrunk to include only A
- Now, probabilities correspond to proportions of A
- What does the orange square represent in the original sample space?
 - $A \cap B$
- How would we find the probability of B given A?

Conditioning changes the sample space

 Before we knew anything, anything in sample space S could occur.

After we know A happened, we are only choosing from within A.
e.g., A: even numbers, B: get a 2

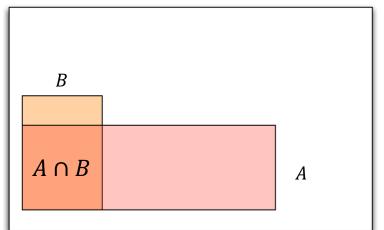
• The set A becomes our new sample space

Instead of asking "In what proportion of S is B true?", we now ask
 "In what proportion of A is B true?" e.g., 1/6 vs 1/3

Conditional Probability

• To find the conditional probability of B given A, consider the ways B can occur in the context of A (i.e., $A \cap B$), out of all the ways A can occur:

$$P(B \mid A) = \frac{P(A \cap B)}{P(A)}$$



Every Probability is a Conditional Probability

We can consider the original probabilities to be conditioned on the event
 S: at first what we know is that "something in S" occurs. E.g.

$$P(B) = P(B|S)$$

$$P(B \cap C) = P(B \cap C|S)$$

- P(B|S) in words: what proportion of S does B happen?
- If we then learn that A occurs, A becomes our restricted sample space.
 P(B|A) in words: what proportion of A does B happen?

Joint Probability and Conditional Probability

• We can rearrange $P(B \mid A) = \frac{P(A \cap B)}{P(A)}$ and derive:

The "Chain Rule" of Probability

For any events, A and B, the joint probability $P(A \cap B)$ can be computed as

$$P(A \cap B) = P(B|A) \times P(A)$$

Or, since $P(A \cap B) = P(B \cap A)$

$$P(A \cap B) = P(A \mid B) \times P(B)$$

Terminology

When we have two events A and B...

• Conditional probability: P(A|B), $P(A^c|B)$, P(B|A) etc.

• Joint probability: P(A,B) or $P(A^c,B)$ or ...

• Marginal probability: P(A) or $P(A^c)$

Example revisited: blood types

		Antigen B		
		Absent	Present	Marginal
Antigen A	Absent	0.44	0.10	0.54
	Present	0.42	0.04	0.46
	Marginal	0.86	0.14	1.00

Table: Probability Estimates for U.S. Blood Types

- Suppose someone of an unknown blood type gets a test that reveals the presence of antigen A.
 - What is $P(A \mid A)$?

$$P(A \mid A) = \frac{P(A \cap A)}{P(A)} = 1$$

What is P(B | A)?

$$P(B \mid A) = \frac{P(A \cap B)}{P(A)} = \frac{0.04}{0.46} = 0.087$$

Example revisited: Seat Belts

A: parent is buckled

C: child is buckled

		C		
		Buck.	Unbuck.	Marginal
Parent	Buck.	0.48	0.12	0.60
	Unbuck.	0.10	0.30	0.40
Marginal		0.58	0.42	1.00

Table: Probability Estimates for Seat Belt Status

Suppose we pick a family from US at random:

- What is the probability of the event "Child is Buckled"? P(C)
- What should our new estimate be if we know that ("given that") Parent is Buckled? $P(C \mid A)$

Example revisited: Seat Belts

A: parent is buckled

C: child is buckled

		C		
		Buck.	Unbuck.	Marginal
Parent	Buck.	0.48	0.12	0.60
	Unbuck.	0.10	0.30	0.40
Marginal		0.58	0.42	1.00

Table: Probability Estimates for Seat Belt Status

Suppose we pick a family from the US at random:

•
$$P(C) = 0.58$$

$$P(C \mid A) = \frac{P(C \cap A)}{P(A)} = \frac{0.48}{0.60} = 0.8$$

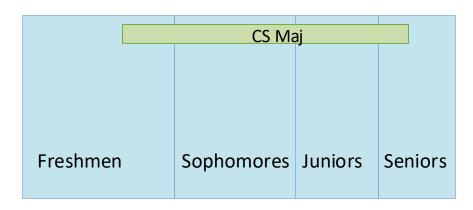
Larger than P(C)

 Suppose we see a buckled parent, it is much more likely that we see their child buckled

Law of Total Probability, revisited

Law of Total Probability Suppose $B_1, ..., B_n$ form a partition of the sample space S. Then,

$$P(A) = P(A, B_1) + \dots + P(A, B_n)$$



Law of Total Probability, revisited

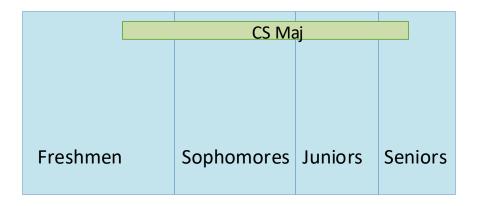
Expanding each $P(A, B_i) = P(A \mid B_i)P(B_i)$, we have:

$$P(A) = \sum_{i=1}^{n} P(A \mid B_i) P(B_i)$$

A: student in CS major

 B_i : student in class year i

 $P(A \mid B_i)$ The fraction of CS major in class year i



Law of Total Probability, revisited

Example Suppose UA has an equal number of students in the 4 class years, and the fraction of CS major in these 4 class years are 10%, 10%, 20%, 80% respectively. What is fraction of CS majors?

•
$$P(B_1) = P(B_2) = P(B_3) = P(B_4) = 0.25$$

•
$$P(C \mid B_1) = 0.1, ..., P(C \mid B_4) = 0.8$$

• Calculate P(C) by:

$$P(C) = \sum_{i=1}^{4} P(C \mid B_i) P(B_i) = 30\%$$

Probabilistic reasoning

Probabilistic reasoning

- We have some prior belief of an event A happening
 - P(A), prior probability
 - e.g. me infected by COVID
- We see some new evidence B
 - e.g. I test COVID positive



- How does seeing B affect our belief about A?
 - $P(A \mid B)$, posterior probability



Another example: detector

A store owner discovers that some of her employees have taken cash. She decides to use a detector to discover who they are.

- Suppose that 10% of employees stole.
- The detector buzzes 80% of the time that someone stole, and 20% of the time that someone not stole
- Is the detector reliable? In other words, if the detector buzzes, what's the probability that the person did stole?

H: employee not stole

B: lie detector buzzes

Another example: detector

Suppose that 10% of employees stole.

H: employee did not stole P(H) = 0.9

 The detector buzzes 80% of the time that someone stoles, and 20% of the time that someone not stole.

B: lie detector buzzes

$$P(B \mid H^{C}) = 0.8$$

 $P(B \mid H) = 0.2$

If the detector buzzes, what's the probability that the person stole?

$$P(H^C \mid B)$$

		Detecto		
		Pass	Buzz (B)	Marginal
Employee	Not stole (H)			
	Stole			
	Marginal			

$$P(H) = 0.9$$

 $P(B \mid H^{C}) = 0.8$
 $P(B \mid H) = 0.2$

$$P(H,B) = P(H) \cdot P(B \mid H) = 0.9 \times 0.2 = 0.18$$

		Detecto		
		Pass (B ^C)	Buzz (B)	Marginal
Employee	Not stole (H)		0.18	0.9
	Stole (H ^C)			0.1
	Marginal			

$$P(H) = 0.9$$

$$P(B \mid H^C) = 0.8$$

$$P(B \mid H) = 0.2$$

$$P(H) = P(H,B) + P(H,B^c) = 0.9$$

		Detector result		
		Pass (B ^C)	Buzz (B)	Marginal
Employee	Not stole (H)	0.72	0.18	0.9
	Stole (H ^C)			0.1
	Marginal			

$$P(H) = 0.9$$

$$P(B \mid H^C) = 0.8$$

$$P(B \mid H) = 0.2$$

		Detecto		
		Pass (B ^C)	Buzz (B)	Marginal
Employee	Not stole (H)	0.72	0.18	0.9
	Stole (H ^C)	0.02	0.08	0.1
	Marginal	0.74	0.26	1

$$P(H) = 0.9$$
$$P(B \mid H^C) = 0.8$$

$$P(B \mid H) = 0.2$$

		Detecto		
		Pass (B ^C)	Buzz (B)	Marginal
Employee	Not stole (H)	0.72	0.18	0.9
	Stole (H ^C)	0.02	0.08	0.1
	Marginal	0.74	0.26	1

• We have the full probability table. Can we calculate $P(H^C \mid B)$? Yes!

$$P(H^C \mid B) = \frac{P(H^C,B)}{P(B)}$$
 $\frac{0.08}{0.26} = 0.307$

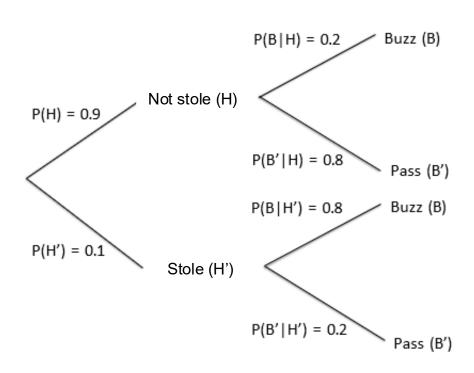
It seems like the detector is not very reliable...

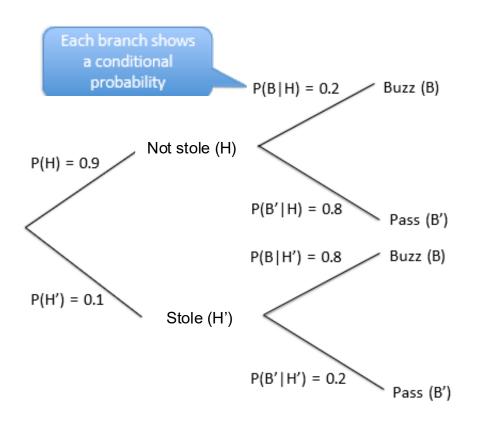
Recap

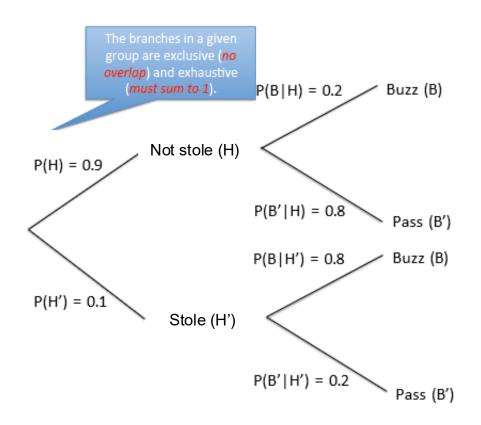
- Conditional probability: $P(B \mid A) = \frac{P(A \cap B)}{P(A)}$
- Law of total probability: $P(A) = \sum_{i=1}^{n} P(A \mid B_i) P(B_i)$
- If we know P(H), $P(B|H^C)$, P(B|H):
 - $P(H) \rightarrow P(H^C)$ Complement rule
 - $P(H), P(B|H) \rightarrow P(B,H)$ Conditional probability
 - $P(H^C), P(B|H^C) \rightarrow P(B,H^C)$ Conditional probability
 - $P(B) \rightarrow P(B, H) + P(B, H^C)$ Law of total probability
 - $P(B), P(B, H) \rightarrow P(H|B)$ Conditional probability
 - $P(B), P(B, H^C) \rightarrow P(H^C|B)$ Conditional probability
- We can get P(B), P(H|B), $P(H^C|B)$

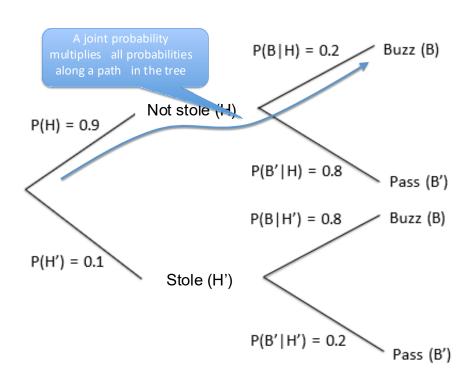
Today's plan

- Another tool: probability trees
- Bayes rule
- Bayes rule and law of total probability
- Independence

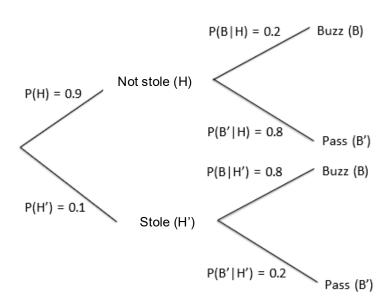








- What is P(Buzz, Stole)?
 - 0.08
- P(Buzz)?
 - Hint: which branches end up with buzzing?
 - · 0.26 (0.08+0.18)
- $P(Stole \mid Buzz)$?
 - Hint: which of the prev. branches contains the stole event?
 - 0.08 / 0.26



In-class activity: COVID test

The Public Health Department gives us the following information:

- A test for the disease yields a positive result (+) 90% of the time when the disease is present (Y)
 P(+ | Y) = 0.9, "sensitivity" of the test
- A test for the disease yields a positive result 1% of the time when the disease is not present (N)
- One person in 1,000 has the disease. P(Y) = 0.1%

<u>Draw a probability tree and use it to answer:</u> what is the probability that a person with positive test has the disease? $P(Y \mid +)$?

In-class activity: COVID test

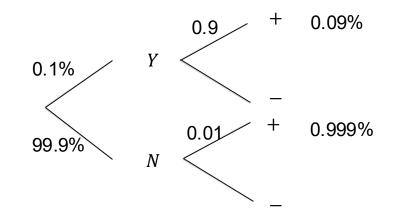
- Goal: calculate P(Y | +)
- Two branches are associated with positive test results +
 - What are the associated events?

•
$$P(+,Y) = P(+|Y|)P(Y) = 0.09\%$$

•
$$P(+,N) = P(-|N|)P(N) = 0.999\%$$

•
$$P(Y \mid +) = \frac{P(+,Y)}{P(+)} = \frac{0.09\%}{0.09\% + 0.999\%} \approx \frac{1}{12}$$

Conclusion: being tested positive does not mean much...



$$P(+ | Y) = 0.9$$

 $P(+ | N) = 0.01$
 $P(Y) = 0.001$

In-class activity: COVID test

Probabilistic reasoning tells as how does seeing a new evidence affect our prior belief about an event.

- Prior probability: one person in 1,000 has the disease: P(Y) = 0.1%
- New evidence: seen a person is tested positive
- Posterior probability: a person with positive test has the disease:

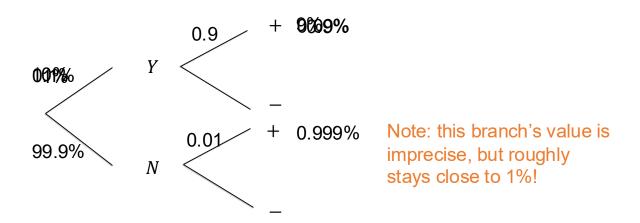
$$P(Y \mid +) = \frac{0.09\%}{0.09\% + 0.999\%} \approx \frac{1}{12}$$

COVID test: additional insights

- What would P(Y | +) look like, if instead:

1 in 100 people have COVID?
$$P(Y \mid +) = \frac{P(+,Y)}{P(+)} = \frac{0.09\%}{0.09\% + 0.999\%} \approx \frac{1}{12}$$

• 1 in 10?



Insight: base rate P(Y) significantly affects $P(Y \mid +)$, hence the conclusions we draw

Conditional probability: additional note

- The rules of probability also applies to the rules of conditional probability
- Just replace P(E), P(F) with P(E|A), P(F|A)
 - But, need to condition on the same A in the same equation

Rules of Probability

- 1. Non-negativity: All probabilities are between 0 and 1 (inclusive)
- **2.** Unity of the sample space: P(S) = 1
- **3.** Complement Rule: $P(E^C) = 1 P(E)$
- 4. Probability of Unions:
 - (a) In general, $P(E \cup F) = P(E) + P(F) P(E \cap F)$
 - (b) If E and F are disjoint, then $P(E \cup F) = P(E) + P(F)$

Some examples

$$P(S|A) = 1$$

A: CS major

$$P(E|A) + P(E^C|A) = 1$$

• $P(E|A) + P(F|A) = P(E \cup F|A)$ for disjoint E and F

	CS Maj		
Freshmen	Sophomores	Juniors	Seniors

Bayes rule

Reversing conditional probabilities

- Is $P(A \mid B) = P(B \mid A)$ in general?
- Let's see..

$$P(A,B) = P(A \mid B) \cdot P(B) = P(B \mid A) \cdot P(A)$$

- Equal only when P(A) and P(B) are equal
- Let's take a look at a real-world example when they are unequal...

Reversing conditional probabilities

Q: Hearing a French accent means someone is French?

Event A: A person is from France.

Event B: A person speaks English with a French accent.

- In a diverse city, only 5% of people are from France
- Of those from France, 80% speak English with a French accent: P(B|A)
- Of those not from France, only 2% speak English with a French accent (due to schooling, mimicry, or neighboring countries)

What is P(A), P(B) and P(A|B)?

Reversing conditional probabilities

What is P(A), P(B) and P(B|A)?

- P(A) = 0.05
- $P(B) = P(A,B) + P(A^c,B) = P(B|A) \cdot P(A) + P(B|A^c) \cdot P(A^c) = 0.8 \cdot 0.05 + 0.02 \cdot (1 0.05) = 0.04 + 0.019 = 0.059$
- $P(A \mid B) = P(A, B)/P(B) = 0.04/0.059 \approx 0.678$

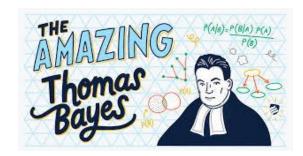
So $P(A) \neq P(B)$, also hearing a French accent doesn't guarantee someone is French: a ~68% chance

Bayes rule

Bayes rule For events A, B,

$$P(A \mid B) = \frac{P(A) \cdot P(B \mid A)}{P(B)}$$

- Very easy to derive from the chain rule, so remember that first.
- Named after Thomas Bayes (1701-1761), English philosopher & pastor



Bayes rule

Bayes rule For events A, B,

Prior probability Support of evidence

$$P(A \mid B) = \frac{P(A) \cdot P(B \mid A)}{P(B)}$$
Posterior probability
$$P(B) = \frac{P(A) \cdot P(B \mid A)}{P(B)}$$
Probability of evidence

Examples:

- A: I have COVID, B: my test shows positive
- A: employee stole B: the detector buzzes
- A: student is CS major B: student is a senior

Bayes rule: another example

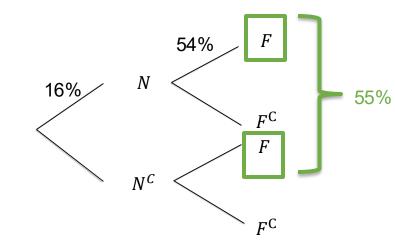
 In a class, 16% of the students are Nutrition Science majors, 55% students are female. Of the Nutrition Science majors, 54% are female.

What proportion of female students in the class are Nutrition Science

majors?

What is the probability tree of this?

We are looking for P(N | F)



Bayes rule: another example

 16% of the students are Nutrition Science majors, 55% are female. Of the Nutrition Science majors, 54% are female. What proportion of female students in the class are Nutrition Science majors?

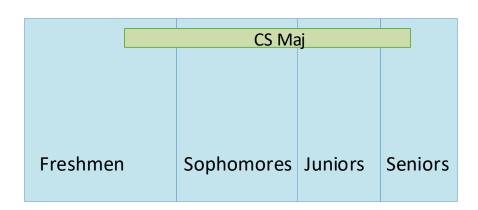
54%

- We can use $P(N \mid F) = \frac{P(N,F)}{P(F)}$ We know P(F) = 0.55
- Can we obtain P(N,F)?
 - We can use $P(N, F) = P(F \mid N) \cdot P(N) = 0.54 \times 0.16$
- Altogether, we have $P(N \mid F) = \frac{P(F \mid N) \cdot P(N)}{P(F)} = \frac{0.54 \times 0.16}{0.55}$

Bayes rule and Law of Total Probability

Bayes rule (equivalent form) For event A and $B_1, ..., B_n$ forming a partition of S,

$$P(B_i \mid A) = \frac{P(A \mid B_i) \cdot P(B_i)}{\sum_{j=1}^{n} P(A \mid B_j) \cdot P(B_j)} \qquad P(A)$$



. . . .

Bayes rule and Law of Total Probability

Example Suppose UA has an equal number of students in the 4 class years, and the fraction of CS major in these 4 class years are $10\% (P(C|B_1))$, 10%, 20%, 80% respectively.

We have previously calculated that P(C) = 30%

If we see a CS major student, what is their most likely year class?

$$P(B_1 | C), ..., P(B_4 | C) \rightarrow \text{maximum}$$
?

Bayes rule and Law of Total Probability

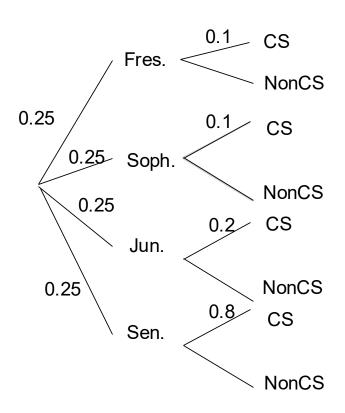
- Let's draw a probability tree...

After learning that the student is CS major:
$$P(B_1 \mid C) = \frac{0.25 \times 0.1}{P(C)}$$

$$P(B_4 \mid C) = \frac{0.25 \times 0.8}{P(C)}$$

- So most likely, this student is a senior
- Equivalent form: $P(B_i \mid C) \propto P(B_i)P(C \mid B_i)$

 - P(C) can be viewed as a *normalization factor*



Recap: conditional probability

• Conditional prob $P(B \mid A) = \frac{P(A, B)}{P(A)}$

The "Chain Rule" of Probability

For any events, A and B, the joint probability $P(A \cap B)$ can be computed as

$$P(A \cap B) = P(B|A) \times P(A)$$

Or, since $P(A \cap B) = P(B \cap A)$

$$P(A \cap B) = P(A \mid B) \times P(B)$$

Extension: chain rule for conditional probability

 If we deal with more than 3 events happening together, we can apply the chain rule of probability repeatedly:

Treat (B, C) as a single event $P(A, B, C) = P(A \mid B, C) P(B, C)$

$$= P(A \mid B, C) P(B \mid C) P(C)$$

Recap

- Useful tools:
 - Contingency table
 - Probability tree
- Bayes rule:

•
$$P(A \mid B) = \frac{P(A,B)}{P(B)} = \frac{P(A) \cdot P(B \mid A)}{P(B)}$$

Bayes rule + law of total probability:

•
$$P(B_i \mid A) = \frac{P(A|B_i) \cdot P(B_i)}{P(A)} = \frac{P(A|B_i) \cdot P(B_i)}{\sum_{j=1}^n P(A|B_j) \cdot P(B_j)}$$

Independence

Probabilistic Independence

- Event S: 10% of employees stole.
- Event R: There's a 5% chance of rain tomorrow.
- What's the probability an employee stole if it rains tomorrow?

Probably your intuition is that one conveys no information about the other. What does this mean about the relationship between P(R|S), and P(S)?

Probabilistic Independence

Independent Events

We say that event *A* is **independent** of event *B* if conditioning on *B* does not change the probability of *A*, that is if

$$P(A|B) = P(A)$$

- Is the independence symmetric?
- In other words, if P(A|B) = P(A), is P(B|A) = P(A)?

Probabilistic Independence

• If A is independent of B, then $P(A \mid B) = P(A)$. Is P(B|A) also equal to P(B)?

Using Bayes' rule, we have $P(B|A) = \frac{P(A \mid B)P(B)}{P(A)}$

So independence is indeed a symmetric notion

Independence: equivalent statement

 If A, B are independent, then their joint probability has a simple form:

$$P(A,B) = P(A \mid B)P(B)$$
$$= P(A) \cdot P(B)$$

This is an equivalent characterization of independence

Independence (version 2)

If A and B are independent events, then

$$P(A \cap B) = P(A)P(B)$$

Are these independent events?

A box has 2 yellow balls and 2 red balls.

- E: the 1st draw is yellow, F: the 2nd draw is yellow (with replacement)
- E: the 1st draw is yellow, F: the 2nd draw is yellow (without replacement)

Hint: calculate P(E), P(F) and P(E, F)

$$P(E,F) = P(E) \times P(F|E) = \frac{2}{4} \times \frac{1}{3} = \frac{1}{6}$$

$$P(E) = \frac{1}{2}$$

$$P(A) = \sum_{i=1}^{n} P(A \mid B_i) P(B_i) \qquad P(F) = \frac{2}{4} \cdot \frac{1}{3} + \frac{2}{4} \cdot \frac{2}{3} = \frac{1}{2}$$

Independence of several events

- We can generalize the notion of independence from two events to more than two.
 - E.g. A: employee stole; B: rain tomorrow, C: stock price up
- Events A_1, \dots, A_n are independent if for any subsets A_{i_1}, \dots, A_{i_j}

$$P\left(A_{i_1},\ldots,A_{i_j}\right) = P\left(A_{i_1}\right)\cdot\ldots\cdot P(A_{i_j})$$

Independence of several events

• If events A, B, C are independent, then

•
$$P(A, B, C) = P(A) \cdot P(B) \cdot P(C)$$

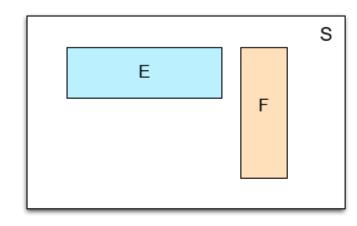
•
$$P(A,C) = P(A) \cdot P(C)$$

•
$$P(B,C) = P(B) \cdot P(C)$$

Rolling a die three times, the probability of sequence (1, 2, 3)?

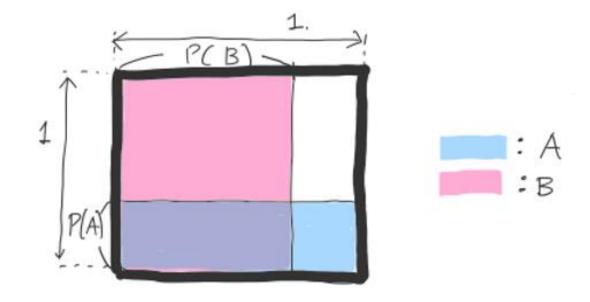
Independent vs. Disjoint Events

- Many people confuse independence with disjointness.
- They are very different!
- What does it mean for two events to be disjoint?
- If A and B are disjoint, then they cannot occur simultaneously; they are mutually exclusive; their intersection is the empty set.
- What does the Venn diagram look like?



Independent vs. Disjoint Events

- If A and B are independent, then P(B|A) = P(B).
- What does the Venn Diagram look like?



Independent vs. Disjoint Events

If A and B are disjoint, what is P(B|A)?

$$P(B \mid A) = \frac{P(A,B)}{P(A)} = 0!$$

- Disjointness is practically the opposite of independence: if A occurs, B doesn't occur
- Defining property of independent events:

$$P(A \cap B) = P(A)P(B)$$

Defining property of disjoint events:

$$P(A \cap B) = 0$$

In-class activity: the absent-minded diners

- Three friends decide to go out for a meal, but they forget where they're going to meet.
 - Fred decides to throw a coin. If it lands heads, he'll go to the Chinese restaurant; tails, and he'll go to the Italian restaurant.
 - George throws a coin, too: heads, it's the Italian restaurant; tails, it's the Mexican restaurant.
 - Ron decides he'll just go to the Italian restaurant.

- What's the probability that all three friends meet?
- What's the probability that one of them eats alone? 1 0.25 = 0.75

Summary

Conditional Probability Summary

- Representing conditional probabilities using contingency tables, Venn diagrams, and probability trees.
- The chain rule
- □ Bayes rule
- I The law of total probability
- Independent events
- Disjoint events

Probability and Combinatorics

Probability and Combinatorics

- Combinatorics (in CSc144) are useful in calculating probabilities
 - Permutations
 - Combinations
- Recall: when all outcomes are equally likely:

We will use combinatorics to do counting

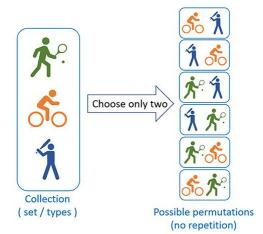
$$P(E) = \frac{|E|}{|S|}$$
 Number of outcomes in event set Number of possible outcomes (e.g. 36)

Permutation number

 If ordered selection of k items out of n is done without replacement, there are

$$n \times (n-1) \times \dots \times (n-k+1) = \frac{n!}{(n-k)!}$$

outcomes



Choose 2 from 3 sports for people A and B

Example 1: sampling without replacement

- Choose 2 from 20 people (including Alice and Bob) for a president (P) and a treasurer (T).
- Probability that Alice is P and Bob is T?

• Total possible ordered outcomes (P,T) = $20 \times 19 = 380$

•
$$P(E) = \frac{1}{380}$$

Example 2: Birthdays

- Probability that at least 2 in a group of 20 have same birthday?
- Sample space: $S = \{(n_1, ..., n_{20}): n_1, ..., n_{20} \in \{1, ..., 365\}\}$
- What is |S|?
 - 365^{20}

• E: set of outcomes where at least two have same birthday, e.g. (5,5,176, ..., 80)

Number of elements

Example 2: Birthdays

- Let's try to calculate |E|
- It turns out that it is easier to calculate $|E^{C}|$
- E^{C} : all 20 birthdays are different 185 13 359 .. 243 19

•
$$|E^C| = \frac{365!}{(365-20)!} = 365 \times 364 \times \dots \times 346^{\frac{\text{Choose 20 from 365 days for people 1, 2, ... 20}}{}$$

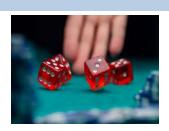
•
$$P(E) = 1 - P(E^C) = 1 - \frac{365!}{(365-20)! \cdot 365^{20}} \approx 0.411$$

This is quite high?! "Birthday paradox"

Repeated independent trials

Example The house or the player?

- 3 dice are rolled:
- House wins if two dice are 6, otherwise player wins.
- What is the probability that the house wins?



Repeated independent trials (Bernoulli trials)

In general, we are interested in the question: suppose we repeatedly perform an experiment n independent times, each with success probability p, what is the probability that we succeed m times?

- The gambling example above: n = 3, $p = \frac{1}{6}$, m = 2
- Applications: sports analytics, gene mutations, etc.
- Named after Jacob Bernoulli (1655-1705)

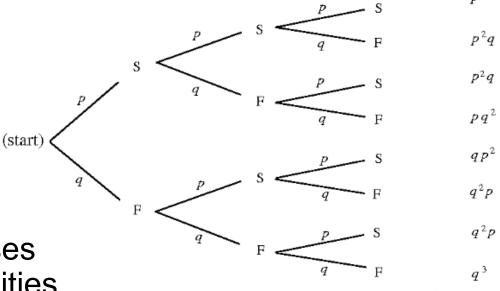


Repeated independent trials: analysis

- Let's draw a probability tree!
- n=3
- Let $q \coloneqq 1 p$

Observations:

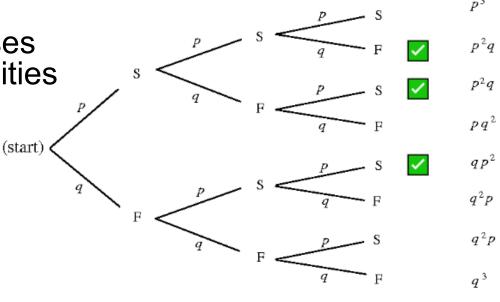
- $2^n = 8$ paths
- Paths with same #successes
 (m) have identical probabilities
 - They are equal to $p^m q^{n-m}$



Repeated independent trials: analysis

Observations:

- $2^n = 8$ paths
- Paths with same #successes
 (m) have identical probabilities
 - They are equal to $p^m q^{n-m}$
- How many paths have 2 successes?
 - . 3?
 - · 2?
 - 1?

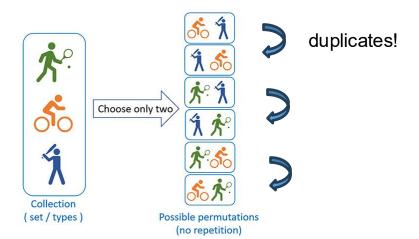


Combination number

 If unordered selection of k items out of n is done without replacement, there are

$$\frac{n!}{(n-k)! \ k!} =: \binom{n}{k}$$

outcomes



Repeated independent trials: analysis

• Out of all $2^3 = 8$ paths, the paths with 2 successes are: SSF, SFS, FSS # such paths is $\binom{3}{2}$

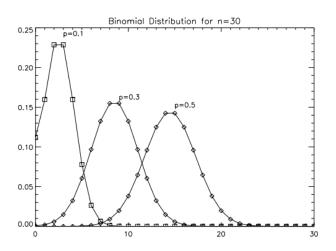
- In general, given n trials, #paths with m successes is $\binom{n}{m}$
 - select m different success positions out of n slots

• Thus, $P(m \text{ successes}) = \binom{n}{m} \cdot p^m q^{n-m}$

Repeated independent trials: conclusion

• In summary, in an experiment with n repeated independent trials with success probability p,

$$P(m \text{ successes}) = \binom{n}{m} \cdot p^m (1-p)^{n-m}$$



The (random) number of successes is said to follow a binomial distribution

.. Back to gambling

Example The house or the player?

- 4 dice are rolled:
- House wins if at least one die is a 6, otherwise player wins.
- What is the probability that the house wins?

- We have n = 4 repeated independent trials
- Here, "success" = "die is a 6"
- The asked probability is $P(\geq 1 \text{ success})$



.. Back to gambling

- We do n = 4 repeated independent trials
- Here, "success" = "die is a 6"
- The asked probability is $P(\geq 1 \text{ successes})$



•
$$P(\ge 1 \text{ successes}) = \sum_{i=1}^{4} P(i \text{ successes})$$

= $\sum_{i=1}^{4} {4 \choose i} \cdot \left(\frac{1}{6}\right)^i \left(\frac{5}{6}\right)^{4-i} = 0.518$

Take home message: the house always wins ©

.. Back to gambling

There is another easier way to think about this problem..

$$P(\ge 1 \text{ successes}) = 1 - P(0 \text{ successes})$$

Complementary rule

$$P(0 \text{ successes}) = P(\text{Fail}_1, \dots, \text{Fail}_4) = \left(\frac{5}{6}\right)^4$$

Independence of the 4 dice rolls