



Computer  
Science

# **CSC380: Principles of Data Science**

## **Probability Primer 2**

Xinchen Yu

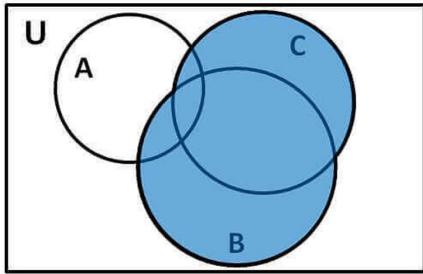
- What is probability?
- Axioms
- Event = set  $\Rightarrow$  use set theory!
- Set theory + axiom 3 is quite useful
- Draw diagrams
- Lots of jargons
  
- Make your own cheatsheet.

# Review

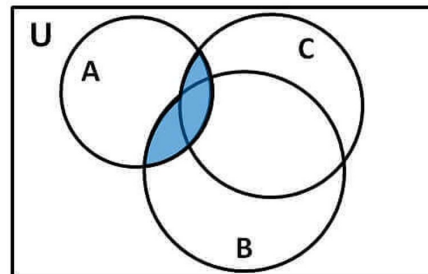
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- $A \cap (B \cup C) = (A \cap B) \cup (A \cap C)$

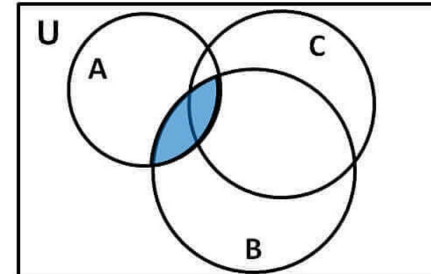
distributive law by Venn diagram



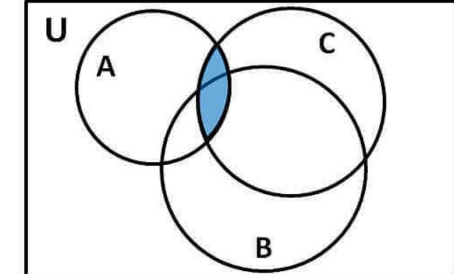
$(B \cup C)$



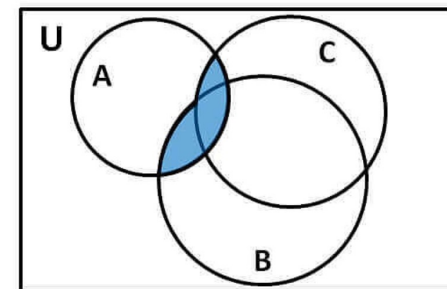
$A \cap (B \cup C)$



$(A \cap B)$

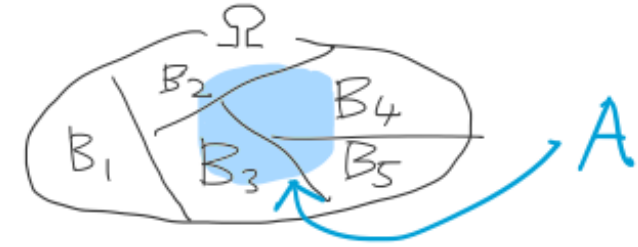


$(A \cap C)$



$(A \cap B) \cup (A \cap C)$

- $A \cap (B \cup C) = (A \cap B) \cup (A \cap C)$
- $A = A \cap \Omega = A \cap (\cup_i B_i) = \cup_i (A \cap B_i)$   
 $= A \cap (B_1 \cup B_2 \cup B_3 \dots \cup B_n)$   
 $= (A \cap B_1) \cup (A \cap B_2) \cup (A \cap B_3) \dots \cup (A \cap B_n)$



**Law of total probability:** Let  $A$  be an event. For any events  $B_1, B_2, \dots$  that partitions  $\Omega$ , we have

$$P(A) = \sum_i P(A \cap B_i)$$

# Review

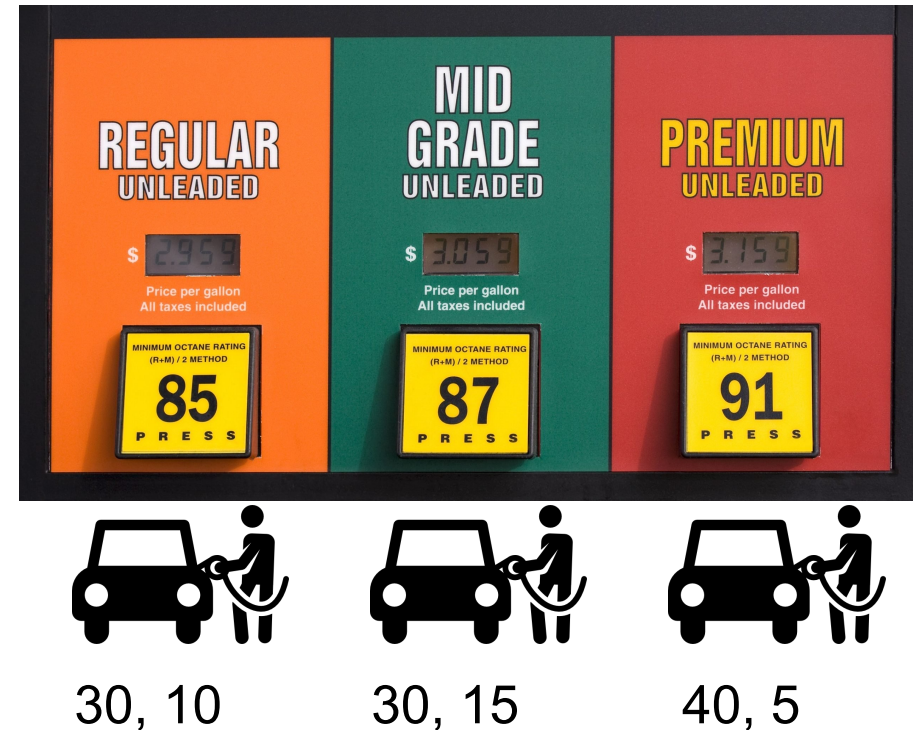
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$$P(A) = \sum_i P(A \cap B_i)$$

A: the customer (100)

B: fill gas

- $B_1$ : unleaded (30)
- $B_2$ : mid grade (30)
- $B_3$ : premium (40)



$P(A = \text{student})$

$= P(A = \text{student}, B = B_1) + P(A = \text{student}, B = B_2) + P(A = \text{student}, B = B_3)$

$= P(A = \text{student} | B = B_1)P(B = B_1) + P(A = \text{student} | B = B_2)P(B = B_2) + P(A = \text{student} | B = B_3)P(B = B_3)$

- Numpy package
- Conditional probability
- Independence

# Numpy Library

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*Package containing many useful numerical functions...*



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## ANACONDA DISTRIBUTION

The world's most popular open-source Python distribution platform

`conda install numpy`

If you use pip:

`pip install numpy`



*...we are interested in `numpy.random` at the moment*

# numpy.random

## numpy.random.randint

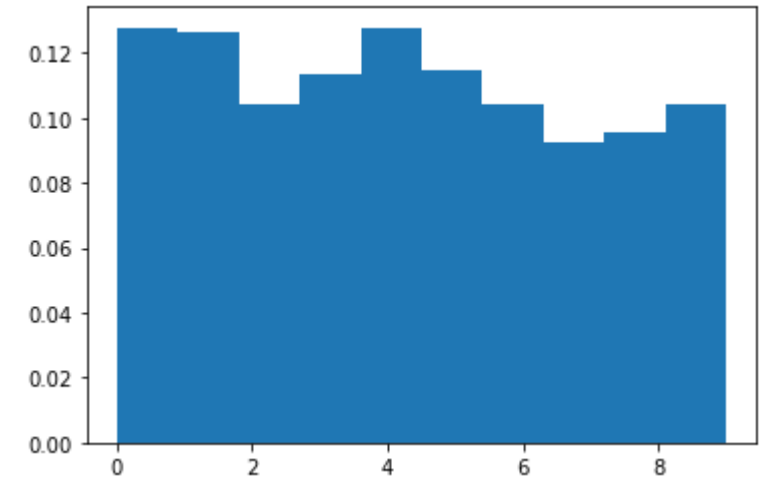
`numpy.random.randint(low, high=None, size=None, dtype='i')`

Return random integers from *low* (inclusive) to *high* (exclusive).

Return random integers from the “discrete uniform” distribution of the specified dtype in the “half-open” interval [*low*, *high*). If *high* is None (the default), then results are from [0, *low*).

Sample a discrete uniform random variable,

```
import matplotlib.pyplot as plt
X = np.random.randint(0, 10, 1000)
count, bins, ignored = plt.hist(X, 10, density=True)
plt.show()
```



- **Caution** Interval is [*low*, *high*) and upper bound is **exclusive**
- *Size* argument accepts tuples for sampling ndarrays (multidimensional arrays)



# numpy.random

*Allows sampling from many common distributions*

Set (global) random seed as,

```
import numpy as np  
  
seed = 12345  
np.random.seed(seed)
```

- 😊 easier to debug (otherwise, you may have 'stochastic' bug)
- ☹️ can be risky

E.g., buy into the result based on a particular seed, publish a report.  
... turns out, you get a widely different result if you use a different seed!

Recommendation: change the seed every now and then

# Conditional Probability

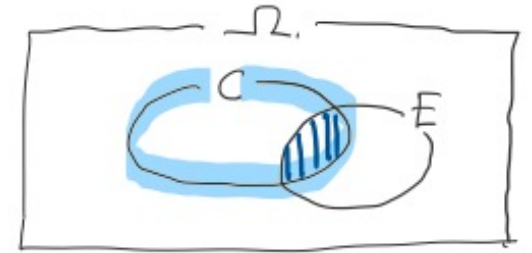
- Two fair dice example:

- Suppose I roll two dice secretly and tell you that one of the dice is 2. C
- In this situation, find the probability of two dice summing to 6. E

```
import numpy as np
for n in [10,100,1000,10_000,100_000, 1_000_000]:
    res_dice1 = np.random.randint(6,size=n) + 1
    res_dice2 = np.random.randint(6,size=n) + 1
    res = [(res_dice1[i], res_dice2[i]) for i in range(len(res_dice1))]

    conditioned = list(filter(lambda x: x[0] == 2 or x[1] == 2, res))
    n_eff = len(conditioned)

    cnt = len(list(filter(lambda x: x[0] + x[1] == 6, conditioned)))
    print("n=%9d, n_eff=%9d, result: %.4f " % (n, n_eff, cnt/n_eff))
```



compare:  
without conditioning,  
it was 0.138.

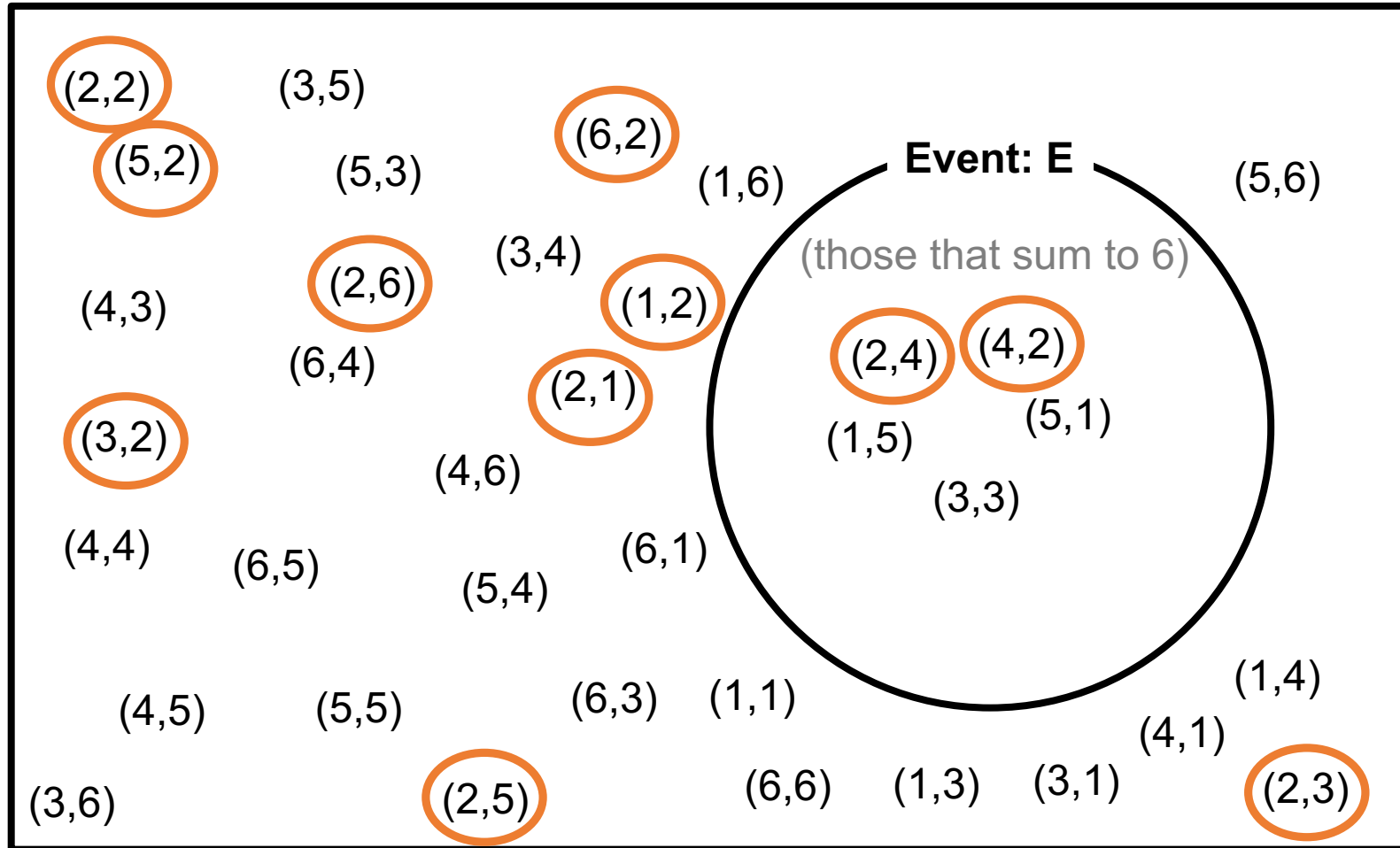
|    |         |          |        |           |        |
|----|---------|----------|--------|-----------|--------|
| n= | 10      | , n_eff= | 4      | , result: | 0.0000 |
| n= | 100     | , n_eff= | 32     | , result: | 0.2500 |
| n= | 1000    | , n_eff= | 300    | , result: | 0.1733 |
| n= | 10000   | , n_eff= | 3002   | , result: | 0.1742 |
| n= | 100000  | , n_eff= | 30590  | , result: | 0.1823 |
| n= | 1000000 | , n_eff= | 305616 | , result: | 0.1818 |

|    |         |          |        |           |        |
|----|---------|----------|--------|-----------|--------|
| n= | 10      | , n_eff= | 3      | , result: | 0.3333 |
| n= | 100     | , n_eff= | 32     | , result: | 0.0625 |
| n= | 1000    | , n_eff= | 343    | , result: | 0.2245 |
| n= | 10000   | , n_eff= | 3062   | , result: | 0.1897 |
| n= | 100000  | , n_eff= | 30651  | , result: | 0.1811 |
| n= | 1000000 | , n_eff= | 305580 | , result: | 0.1808 |

# Random Events and Probability

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*What is the probability of having two numbers sum to 6 given one of dice is 2?*



Each outcome is equally likely.  
by the **independence**  
(will learn this concept later)

=>  $1/36$

# sum to 6:

=>  $5$

# one of dice is 2:

=>  $11$

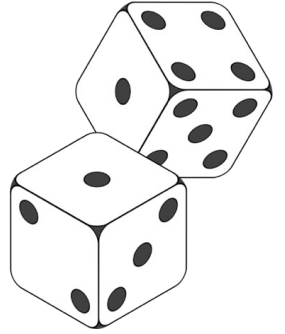
# sum to 6 and one of dice  
is 2:

=>  $2$

answer:

$2/11 = 0.181818....$

## Two fair dice example

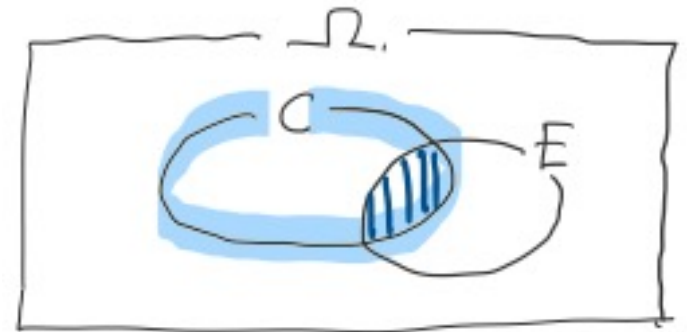


- Find the probability of one of the dice is 2 (event  $C$ ) and two dice summing to 6 (  $E$  )

$$P(E \cap C)$$

- I secretly tell you one of the dice is 2, find the probability of two dice summing to 6.

$$\frac{P(E \cap C)}{P(C)}$$



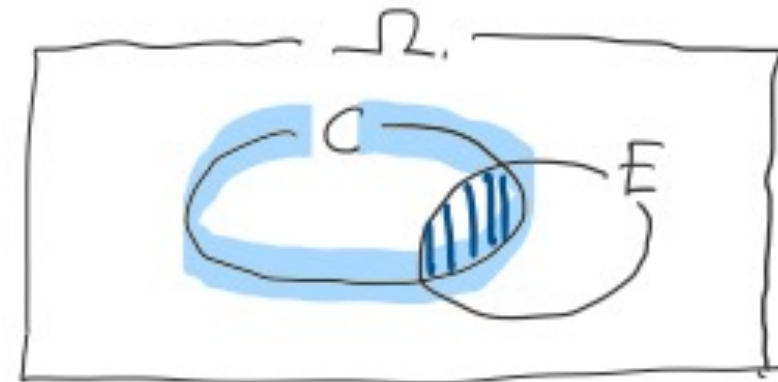
- Two fair dice example:

- Suppose I roll two dice and secretly tell you that **one of the dice is 2.**  $C$
- **In this situation**, find the probability of **two dice summing to 6.**  $E$

- Turns out, such a probability can be computed by  $\frac{P(E \cap C)}{P(C)}$
- It's like “zooming in” to the condition.
- This happens a lot in practice, so let's give it a notation:

$$P(E|C) := \frac{P(E \cap C)}{P(C)}$$

Say: probability of “ $E$  given  $C$ ”, “ $E$  conditioned on  $C$ ”



“it's the ratio”

Q: Conditional probability  $P(A|B)$  could be undefined. When?

- A: The denominator can be 0 already. In this case, numerator is also 0!

Note  $P(A|B) \neq P(B|A)$  in general!

$$P(A|B) := \frac{P(A \cap B)}{P(B)}$$

E.g., throw a fair die.  $X$  := outcome.  $A = \{X=4\}$ ,  $B = \{X \text{ is even}\}$

**Question:**  $P(A | B) = P(B | A)$  ?

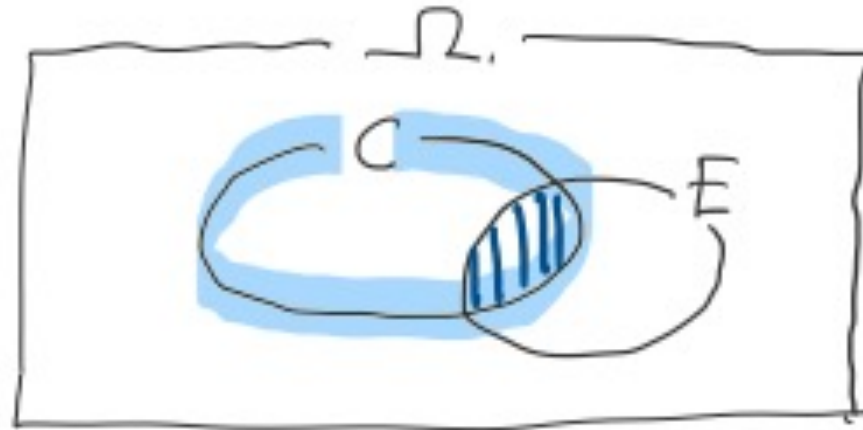
- $P(A) = 1/6$
- $P(B) = 1/2$
- $P(A \cap B) = 1/6$
- Therefore,  $P(A|B) = 1/3$ ,  $P(B|A) = 1$

## Chain rule

- $P(A \cap B) = P(A|B)P(B)$  ← just a rearrangement of definition:  $P(A|B) := \frac{P(A \cap B)}{P(B)}$
- $P(A \cap B \cap C) = P(A|B \cap C)P(B \cap C) = P(A|B \cap C)P(B|C)P(C)$
- $P(E_1 \cap E_2 \cap \dots \cap E_n) = P(E_1) \prod_{i=2}^n P(E_i | \cap_{j=1}^{i-1} E_j)$  valid for any ordering!



- $P(E \cap C) = P(E|C)P(C) = P(C|E)P(E)$



“it’s the ratio”

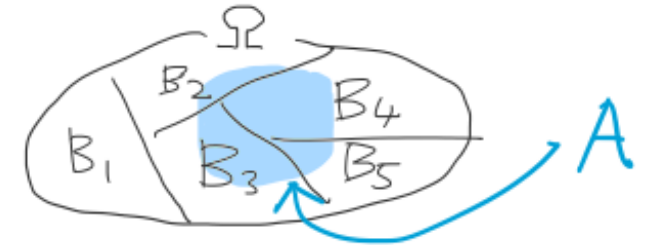
# Conditional Probability

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Recall: let  $A$  be an event. For events  $B_1, B_2, \dots$  that partitions  $\Omega$ , we have

$$\begin{aligned} P(E \cap C) \\ &= P(E|C)P(C) \\ &= P(C|E)P(E) \end{aligned}$$

$$P(A) = \sum_i P(A \cap B_i)$$



$$A = A \cap \Omega = A \cap (\cup_i B_i) = \cup_i (A \cap B_i)$$

Check axiom 3 & distributive law!

**Law of total probability:** If  $A \in \mathcal{F}$  and  $\{B_i \in \mathcal{F}\}_i$  partitions  $\Omega$ , then

$$P(A) = \sum_i P(A, B_i) = \sum_i P(B_i)P(A|B_i)$$

Shortcut:  
 $P(A, B) := P(A \cap B)$

$$= \sum_i P(A)P(B_i|A)$$

(by definition)

- $P(A) = \sum_i P(A, B_i) = \sum_i P(B_i)P(A|B_i)$

$P(A = \text{student})$

$$= P(A = \text{student}|B = B_1)P(B = B_1) + P(A = \text{student}|B = B_2)P(B = B_2) + P(A = \text{student}|B = B_3)P(B = B_3)$$

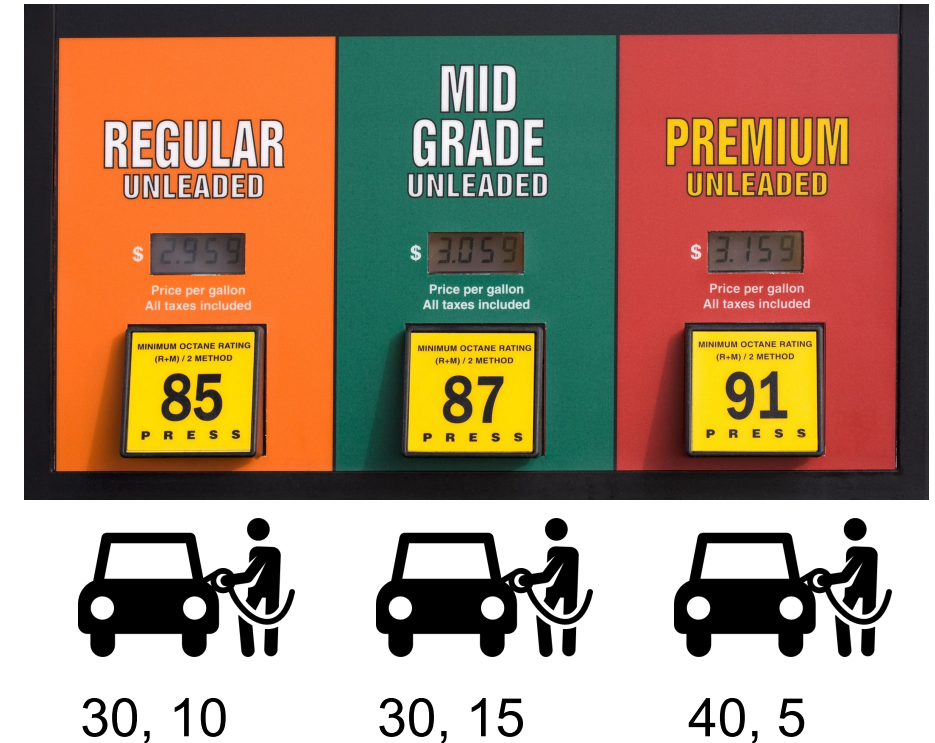
$P(A = \text{student})$

$$= 10/30 \times 30/100 + 15/30 \times 30/100 + 5/40 \times 40/100$$

- $\sum_i P(B_i|A) = 1$

$$P(B_1|A = \text{student}) + P(B_2|A = \text{student}) + P(B_3|A = \text{student})$$

$$= \frac{10}{10+15+5} + \frac{15}{10+15+5} + \frac{5}{10+15+5} = 1$$



The Public Health Department gives us the following information:

- A test for the disease yields a positive result (+) 90% of the time when the disease is present (Y)  
 $P(+ | Y) = 0.9$
- A test for the disease yields a positive result 1% of the time when the disease is not present (N)  
 $P(+ | N) = 0.01$
- One person in 1,000 has the disease.  
 $P(Y) = 0.001$

Q: What is the probability that a person with positive test has the disease?  $P(Y | +)$ ?

Pick a person **uniformly at random** from the population. Apply the test. When test=+, what is the probability of this person having the disease (Y) ?

# Conditional Probability

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What we know:

$$P(+ | Y) = 0.9$$

$$P(+ | N) = 0.01$$

$$P(Y) = 0.001$$

$\Rightarrow$

$$P(- | Y) = 0.1$$

$$P(- | N) = 0.99$$

$$P(N) = 0.999$$

Question:  $P(Y | +)$

$$= \frac{P(Y, +)}{P(+)}$$

$$P(+)=P(+,Y)+P(+,N)$$

$$P(+,Y)=P(+|Y)P(Y)$$

$$P(+,N)=P(+|N)P(N)$$

Law of total probability

$$P(A)=\sum_iP(A,B_i)=\sum_iP(B_i)P(A|B_i)$$

The answer is 0.0826...

When we have two events A and B...

- Conditional probability:  $P(A|B)$ ,  $P(A^c|B)$ ,  $P(B|A)$  etc.
- Joint probability:  $P(A, B)$  or  $P(A^c, B)$  or ...
- Marginal probability:  $P(A)$  or  $P(A^c)$

# Conditional Probability

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Tip: Make a table of **joint probabilities**

$$P(+ | Y) = 0.9$$

$$P(+ | N) = 0.01$$

$$P(Y) = 0.001$$

Each cell is  $P(\text{column event} \cap \text{row event}) = P(T=t \cap D=d) = P(T=t | D=d) P(D=d)$

|           | Test = +                     | Test = -                     |       |
|-----------|------------------------------|------------------------------|-------|
| Disease=Y | $0.9 \cdot 0.001 = 0.0009$   | $0.1 \cdot 0.001 = 0.0001$   | 0.001 |
| Disease=N | $0.01 \cdot 0.999 = 0.00999$ | $0.99 \cdot 0.999 = 0.98901$ | 0.999 |
|           | 0.01089                      | 0.98911                      |       |

Workflow:

- make a table, then fill in the cells.
- write down the target  $P(A|B)$  all in terms of joint probabilities and marginal probabilities.

$P(\text{test} = +)$

We can directly calculate:

$$P(Y|+) = \frac{P(Y,+)}{P(+)} = \frac{P(+|Y)P(Y)}{P(+)}$$

## Bayes rule

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

proof: definition and definition!

⇒ particularly useful in practice: infer  $P(A|B)$  given  $P(B|A)$ !

$P(A)$ : **prior** probability

e.g.,  $A$ =‘dice sum to 6’,  $B$ =‘one of the die is 2’

$P(A|B)$ : **posterior** probability

e.g.,  $A$ =‘disease=Y’,  $B$ =‘test=+’



# Independence

# Independence

- Informally, given two events A and B, they are **independent** if the probability of A is not affected by whether B is true or false (and vice versa)
  - E.g., A = “die1=1” and B=“die2=1” are independent.  
⇒ we know that the probability of die1 being 1 would not be changed just because die2=1.
- Mathematically, this can be written as  $P(A|B) = P(A)$  or  $P(B|A) = P(B)$ .
- E.g., A = “die1=6” and B=“two dice sum to 6” are **not independent**. quiz candidate
  - ∴ intuitively, when B is true, A can never happen! So,  $P(A|B)=0$  but  $P(A) = 1/6$ .
- E.g., A = “die1=1” and B=“two dice sum to 6” are **not independent**. quiz candidate
  - ∴  $P(A) = 1/6 = 0.166\dots$  . However,  $P(A|B) = 1/5 = 0.2$

$$P(A|B) = \frac{P(A, B)}{P(B)}$$

# Independence

[Def] Two events  $A$  and  $B$  are **independent** if

$$P(A, B) = P(A)P(B)$$

$A \perp B$  means  $A$  and  $B$  are independent

“joint probability is product of two marginal probabilities”

=> note: symmetric!

Also, a set of events  $\{A_i\}_{i=1}^n$  ( $n$  can be  $\infty$ ) are **mutually independent** if

for every  $J \subseteq \{1, \dots, n\}$ , we have  $P(\cap_{i \in J} A_i) = \prod_{i \in J} P(A_i)$

- Ex) recall two fair dice

- We took it for granted that  $P((1,1))$  is  $1/36$ .
- But why is it true, really?
- To be rigorous,

$$P(\text{die1} = 1, \text{die2} = 1) = P(\text{die1} = 1)P(\text{die2} = 1) = \frac{1}{6} \cdot \frac{1}{6}$$

due to independence.

$$\text{or, ...} = P(\text{die1}=1 \mid \text{die2}=1) * P(\text{die2}=1) = P(\text{die1}=1) * P(\text{die2}=1)$$

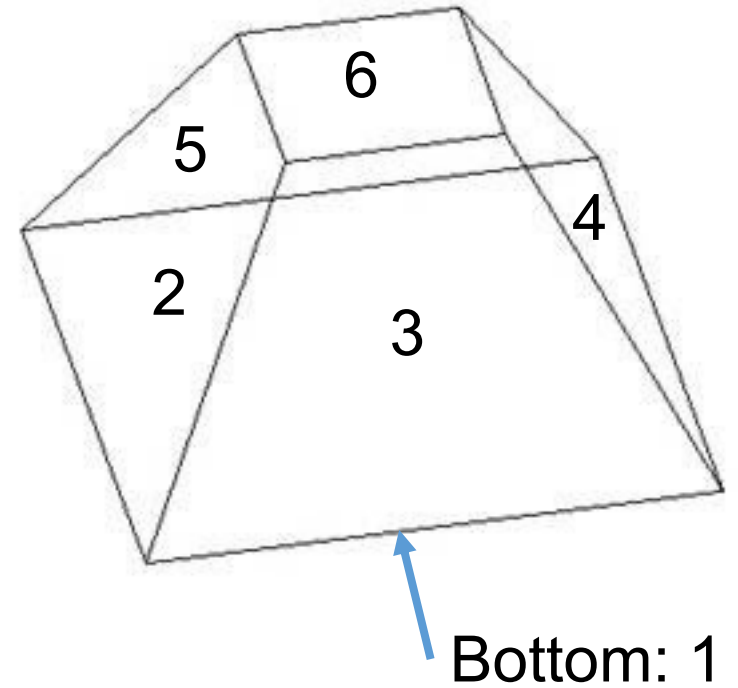
- E.g., two biased coin **C1** and **C2**. Suppose  $P(C1=H) = 0.3$  and  $P(C2=H) = 0.4$ . Compute the probability of  $P(C1=H, C2=T)$ .

$$0.3 \cdot 0.6 = 0.18$$

quiz candidate

# Independence

- Ex) Unfair die
  - Let  $A$  be the outcome of a single throw.
  - $P(A=1) \ll P(A=2) = \dots = P(A=5) \ll P(A=6)$   
say, 0.1                  0.15                  0.15.                  0.3
- Throw this die twice. What's the probability of observing (1,3)?
  - $P((1,3)) = 0.1 * 0.15 = 0.015$
- Similarly,  $P((6,3)) = 0.3 * 0.15 = 0.045$



# Example: Dependent Coin Flips

- First coin ( $X_1$ ): fair coin
- Second coin ( $X_2$ ):
  - if  $X_1=H$ , throw a **fair** coin.
  - If  $X_1=T$ , throw an **unfair** coin  $P(H) = 0.2$ ,  $P(T) = 0.8$

- Q: Are  $X_1=H$  and  $X_2=H$  independent or not?

$$P(X_1=H) = \underline{\hspace{2cm}} \quad 0.5$$

$$P(X_2=H) = \underline{\hspace{2cm}} \quad = P(X_2=H, X_1=H) + P(X_2=H, X_1=T) = 0.25 + 0.1 = 0.35$$

$$P(X_1=H, X_2=H) = \underline{\hspace{2cm}} \quad 0.25$$

$$P(X_1=H) * P(X_2=H) = 0.175$$

Quiz candidate

**Axiom 3:**

For any *finite* or *countably infinite* sequence of disjoint events  $E_1, E_2, E_3, \dots$ , 
$$P\left(\bigcup_{i \geq 1} E_i\right) = \sum_{i \geq 1} P(E_i)$$

**Inclusion-exclusion rule:**

$$P(E_1 \cup E_2) = P(E_1) + P(E_2) - P(E_1 \cap E_2)$$

**Law of total probability:** For events  $B_1, B_2, \dots$  that partitions  $\Omega$ ,

$$P(A) = \sum_i P(A \cap B_i)$$

**Conditional probability:**

$$P(E|C) := \frac{P(E \cap C)}{P(C)}$$

$(P(A|B) \neq P(B|A) \text{ in general})$

**Probability chain rule:**

$$P(A \cap B \cap C) = P(A|B \cap C)P(B|C)P(C)$$

**Bayes' rule:**

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

**Independence:**

(definition) A and B are independent if  $P(A, B) = P(A)P(B)$

(property) A and B are independent if and only if  $P(A|B) = P(A)$  (or  $P(B|A) = P(B)$ )