Differential Expression Analysis Methods Comparison

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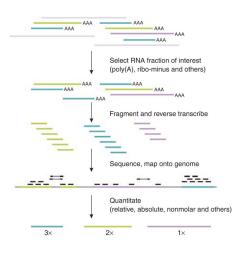
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RNA fragmentation, sequencing, and alignment



(Pepke, Wold, and Mortazavi (2009) http://www.nature.com/nmeth/journal/v6/n11s/fig_tab/nmeth.1371_F5.html)

RNAseq data

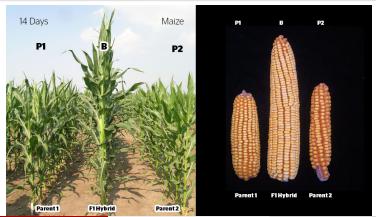
| Genes | B73 | B73 | B73 | B73 | Mo17 | Mo17 | Mo17 | Mo17 |
|------------------|------|------|------|------|------|------|------|------|
| | Rep1 | Rep2 | Rep3 | Rep4 | Rep1 | Rep2 | Rep3 | Rep4 |
| AC148152.3_FG001 | 3 | 4 | 6 | 0 | 8 | 17 | 18 | 20 |
| AC148152.3_FG008 | 3 | 3 | 4 | 1 | 31 | 40 | 45 | 49 |
| AC152495.1_FG002 | 33 | 46 | 18 | 13 | 4 | 0 | 2 | 6 |
| AC152495.1_FG017 | 41 | 44 | 16 | 13 | 2 | 2 | 2 | 0 |
| AC184130.4_FG012 | 24 | 47 | 18 | 21 | 110 | 144 | 121 | 96 |
| AC184133.3_FG001 | 0 | 1 | 1 | 0 | 14 | 13 | 4 | 9 |
| AC148152.3_FG005 | 2323 | 1533 | 1932 | 1945 | 2070 | 1582 | 2196 | 1882 |
| AC148167.6_FG001 | 672 | 598 | 728 | 713 | 743 | 655 | 821 | 824 |
| AC149475.2_FG002 | 459 | 438 | 451 | 483 | 467 | 448 | 634 | 532 |
| AC149475.2_FG003 | 1184 | 976 | 1131 | 1206 | 891 | 743 | 1288 | 1107 |
| AC149475.2_FG005 | 551 | 535 | 360 | 353 | 550 | 524 | 492 | 440 |
| AC149475.2_FG007 | 245 | 214 | 169 | 159 | 297 | 262 | 210 | 302 |

• DE genes: expression in Genotype Variety B73 is different from that in another Genotype Variety Mo17

Differential Expression Genes

Definition

A gene is regarded as differentially expressed (DE) when the expected count reads of this gene corresponding to one genotype variety differs from that of another genotype variety.



Differential expression analysis

Definition

For a given gene, we use statistical testing to decide whether an observed difference in read counts is significant, i.e., whether it is greater than what would be expected just due to natural random variation.

- Normalization
 Estimated normalization factors should ensure that a gene with the same expression level in two samples is not detected as DE.
- Assumed distribution Negative binomial
- Parameter estimation Mean, Dispersion
- Test for DE Exact test, Wald test, t-test

Negative Binomial Model in Generalized Linear Model Framework (Part 1)

Let

- g (g = 1, ..., G) identify the gene,
- i (i = 1, 2) identify the genotype variety,
- j (j = 1, 2, 3, 4)
- Y_{gij} be the RNAseq counts of gene g, genotype variety i, replicate j

We assume

$$Y_{\mathrm{g}ij} \overset{\mathrm{ind}}{\sim} \mathrm{NB}\left(\mu_{\mathrm{g}ij}, \phi_{\mathrm{g}}\right)$$

where

- ullet μ_{gij} are means of read counts of gene g genotype i replicate j,
- \bullet ϕ_{g} allow for gene-specific overdispersion

Negative Binomial Model in Generalized Linear Model Framework (Part 2)

In the generalized linear model (GLM) setting, the mean response, μ_{gij} is linked to a linear predictor with natural log link:

$$log(\mu_{gij}) = x_i^T \beta_g + log(N_{ij})$$

where

- x_i is row of the design matrix containing the covariates indicating this sample belongs to variety i,
- $\beta_g = (\beta_{g1}, \beta_{g2})$ is a vector of regression parameters
- N_{ij} is the normalized library size of replicate j in variety i

Hierarchical model for RNA-seq counts

We assume

$$Y_{gij} \stackrel{ind}{\sim} \mathsf{NB}\left(\mu_{gij}, \phi_{g}\right)$$

where

- $\mu_{gij} = \exp(x_i^T \beta_g + \log(N_{ij}))$
- $\lambda_{gi} = x_i^T \beta_g, \gamma_{ij} = \log(N_{ij})$, then γ_{ij} are normalization factors
- $\phi_g = \exp(\psi_g)$ allow for gene-specific overdispersion

We reparameterized the mean dispersion structure into the genespecific average β_{g1} and half-variety difference β_{g2}

$$\beta_{g1} = \frac{\lambda_{g1} + \lambda_{g2}}{2}, \beta_{g2} = \frac{\lambda_{g1} - \lambda_{g2}}{2}$$

we also assume

$$\beta_{\mathbf{g}1} \overset{\mathit{ind}}{\sim} \mathsf{N}\left(\eta_{\beta_{1}}, \sigma_{\beta_{1}}^{2}\right), \beta_{\mathbf{g}2} \overset{\mathit{ind}}{\sim} \mathsf{N}\left(\eta_{\beta_{2}}, \sigma_{\beta_{2}}^{2}\right), \psi_{\mathbf{g}} \overset{\mathit{ind}}{\sim} \mathsf{N}\left(\eta_{\psi}, \sigma_{\psi}^{2}\right)$$

 $\beta_{g1}, \beta_{g2}, \psi_{g}$ are independent to each other.

Empirical Bayes Method

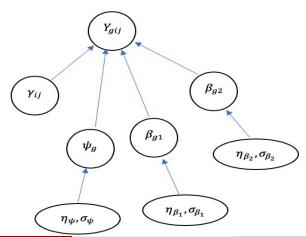
Let

- $\theta = (\theta_1, ..., \theta_G)$ (g = 1, ..., G) where $\theta_g = (\beta_{g1}, \beta_{g2}, \psi_g)$,
- ullet γ_{ij} is the normalized facor for replicate j in variety i,
- $\pi=(\eta,\sigma)$, where $\eta=(\eta_{eta_1},\eta_{eta_2},\eta_{\psi}), \sigma=(\sigma_{eta_1},\sigma_{eta_2},\sigma_{\psi})$

Then,

- ullet $\hat{\gamma}$ was obtained from trimmed mean of M values (TMM)
- ullet $\hat{\psi}_{ extsf{g}}$ was got through the adjusted profile likelihood (APL)
- $\hat{\beta}_{g1}, \hat{\beta}_{g1}$ was retrieved by fitting the generalized linear model with log link function
- $\hat{\pi}=(\hat{\eta},\hat{\sigma})$, where $\hat{\eta}=\sum_{g=1}^G\hat{\beta}/G,\hat{\sigma}^2=\sum_{g=1}^G(\hat{\beta}-\hat{\eta})^2/(G-1)$

Empirical Bayes Method (cont)



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Null Hypothesis for DE Analysis

$$H_0:\beta_{g2}=0$$

which is equivalent to $\lambda_{g1}=\lambda_{g2}$ Statistics used to do the DE analysis is based on the posterior probabilities of β_{g2} as

$$P(DE_{g}|y,\hat{\pi},\hat{\gamma}) = \min(P(\beta_{g2} < 0|y,\hat{\pi},\hat{\gamma}), P(\beta_{g2} > 0|y,\hat{\pi},\hat{\gamma}))$$

where
$$P(\beta_{g2} < 0 | y, \hat{\pi}, \hat{\gamma}) = \frac{1}{M} \sum_{m=1}^{M} I(\beta_{g2}^{(m)} < 0)$$
, and

$$P(\beta_{g2} > 0|y, \hat{\pi}, \hat{\gamma}) = \frac{1}{M} \sum_{m=1}^{M} I(\beta_{g2}^{(m)} > 0)$$

Alternative Methods

Normalization

edgeR used gene-wise trimmed median of means (TMM), while DESeq, DESeq2, sSeq, EBSeq used sample-wise size factor.

Dispersion estimation

edgeR used Cox-Reid approximate conditional inderence (CRACI) moderate towards the mean while DESeq, DESeq2 used CRACI with focus on maximum individual dispersion estimate; sSeq estimated dispersion by pooling all the samples using the method of moments(MM), and then shrinking the gene-wise estimates through minimizing the mean-square error; EBSeq also estimated the gene-specific varainces via MM.

Test for DE

edgeR, sSeq used exact test for 2 factors; DESeq, DESeq2 used Wald test for 2 factors;

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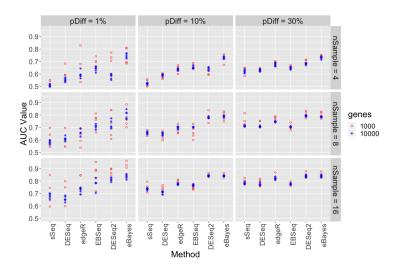
Simulation studies

Parameter estimation use edgeR: $\hat{\beta}_{g1}, \hat{\beta}_{g2}, \hat{\phi}_{g}$, and the normalized library sizes N_{ij}

Simulation scenario set up: nGenes, nSamples, pDiff

Simulation model: $Y_{gij} \stackrel{ind}{\sim} \text{NB}\left(\mu_{gij}, \phi_g\right)$ with $\mu_{gij} = \exp(x_i^T \beta_g + \log(N_{ij}))$ where N_{ij} is the normalized library size. For non-DE genes, we set $\mu_{g1} = \mu_{g2}$.

AUC Plot



Summary of the Results

Effect of nGenes: not obvious

Effect of pDiff: smaller pDiff \rightarrow larger differences between eBayes and other methods

Effect of nSample: smaller nSample -> larger differences between eBayes and other mehods

Discussion

For the future research, we could:

- (1) add more methods: baySeq, ShrinkSeq, NOISeq, SAMseq;
- (2) include more varieties;
- (3) consider the flow cell effects;
- (4) improve the eBayes by refining the hierarchical model