Recall:

Let $A \subset S$. The hitting time T_A of A is defined by $T_A = \min \{ n > 0 : X_n \in A \}$.

 T_A is the \underline{first} (positive) time the Markov chain hits A.

 $T_A = \infty$ if $X_n \notin A$ for all n > 0 (if the Markov chain never hits A).

Let $a \in S$. Let $T_a = T_{\{a\}}$.

$$\rho_{xy} = P_x(T_y < \infty).$$

(The probability that the Markov chain eventually visits y starting from x.)

State y is **recurrent** if $\rho_{vv} = 1$.

 $(\rho_{yy} = P_y(T_y < \infty) = 1 \implies \text{the Markov chain is guaranteed to return to } y \text{ starting from } y.)$

State *y* is **transient** if $\rho_{yy} < 1$.

($\rho_{yy} < 1 \implies$ the Markov chain is NOT guaranteed to return to y starting from y.)

Let y be a recurrent state.

 $m_y = E_y(T_y)$ — mean return time to y for a Markov chain starting at y.

$$E_{x}(1_{y}(X_{n})) = P_{x}(X_{n} = y) = P^{n}(x, y).$$

 $N_n(y) = \sum_{m=1}^n 1_y(X_m)$ - the number of times the Markov chain visits state y up to time n.

$$G_n(x,y) = E_x(N_n(y)) = \sum_{m=1}^n P^m(x,y)$$
 - expected number of visits to y starting at x up to time n .

Let y be a transient state.

$$\Rightarrow N(y) = \lim_{n \to \infty} N_n(y) < \infty \text{ with probability 1,}$$

$$G(x, y) = \lim_{n \to \infty} G_n(x, y) < \infty, \qquad x \in S.$$

$$\Rightarrow \lim_{n \to \infty} \frac{N_n(y)}{n} = 0 \text{ with probability 1.}$$

$$\lim_{n \to \infty} \frac{G_n(x, y)}{n} = 0.$$

Theorem Let y be a recurrent state.

$$\Rightarrow \lim_{n \to \infty} \frac{N_n(y)}{n} = \frac{1\left\{T_y < \infty\right\}}{m_y} \quad \text{with probability 1,}$$

$$\lim_{n \to \infty} \frac{G_n(x, y)}{n} = \frac{\rho_{xy}}{m_y}, \qquad x \in S.$$

$$\frac{N_n(y)}{n} = \text{proportion of the first } n \text{ units of time that the chain is in state } y.$$

$$\frac{G_n(x,y)}{n} = \text{expected value of this proportion for a chain starting at } x.$$

Corollary Let C be an irreducible closed set of recurrent states.

$$\Rightarrow \lim_{n\to\infty} \frac{G_n(x,y)}{n} = \frac{1}{m_y}, \qquad x,y\in C.$$

If $P(X_0 \in C) = 1$, then with probability 1

$$\lim_{n \to \infty} \frac{N_n(y)}{n} = \frac{1}{m_y}, \qquad y \in C.$$

If the chain returns to y every m_y steps (on average), then the proportion of time the chain is in y in the long run is about $\frac{1}{m_y}$.

y is **positive recurrent** if $m_{\nu} < \infty$.

y is **null recurrent** if $m_y = \infty$.

Theorem x is positive recurrent, x leads to y

 \Rightarrow y is positive recurrent.

An irreducible Markov chain (it is possible to get to any state from any state) is either a transient chain (all of its states are transient), or a null recurrent chain (all of its states are null recurrent), or a positive recurrent chain (all of its states are positive recurrent).

An irreducible Markov chain having a finite number of states is positive recurrent (all of its states are positive recurrent).

Silly example 1. Let 0 . Consider a Markov Chain with the states being all nonnegative integer numbers, having transition function

$$P(k, k+1) = 1-p,$$
 $P(k, 0) = p,$ $k \ge 0.$

Determine whether state 0 (and the entire chain) is transient, null recurrent, or positive recurrent.

$$P_0(T_0 \ge m) = P(0,1) \times P(1,2) \times P(2,3) \times ... \times P(m-3, m-2) \times P(m-2, m-1)$$

= $(1-p)^{m-1}$.

$$\Rightarrow P_0(T_0 = \infty) = \lim_{m \to \infty} P_0(T_0 \ge m) = 0.$$

$$\Rightarrow$$
 $\rho_{00} = P_0(T_0 < \infty) = 1.$ \Rightarrow 0 is recurrent.

$$E_0(T_0) = \sum_{m=1}^{\infty} P(T_0 \ge m) = \sum_{m=1}^{\infty} (1-p)^{m-1} = \frac{1}{1-(1-p)} = \frac{1}{p} < \infty.$$

 \Rightarrow 0 is positive recurrent.

The Markov chain is irreducible (it is possible to get to any state from any state).

⇒ The Markov chain is positive recurrent.

Spoiler: Markov chain is irreducible, positive recurrent.

 \Rightarrow There is stationary distribution:

$$\pi(0) = p \,\pi(0) + p \,\pi(1) + p \,\pi(2) + p \,\pi(3) + p \,\pi(4) + p \,\pi(5) + p \,\pi(6) + \ldots,$$

$$\pi(1) = (1-p)\pi(0),$$

$$\pi(2) = (1-p)\pi(1) = (1-p)^2\pi(0),$$

$$\pi(3) = (1-p)\pi(2) = (1-p)^3\pi(0),$$

$$\pi(4) = (1-p)\pi(3) = (1-p)^4\pi(0),$$

$$\pi(5) = (1-p)\pi(4) = (1-p)^5\pi(0),$$

$$\pi(6) = (1-p)\pi(5) = (1-p)^6\pi(0),$$

...

$$\pi(0) + \pi(1) + \pi(2) + \pi(3) + \pi(4) + \pi(5) + \pi(6) + \dots = 1.$$

$$\Rightarrow \pi(0) = p, \ \pi(1) = (1-p)p, \ \pi(2) = (1-p)^2 p, \ \pi(3) = (1-p)^3 p, \dots$$
That is, $\pi(x) = (1-p)^x p, \ x \ge 0.$

Note that
$$m_0 = \frac{1}{p}$$
 and $\pi(0) = p = \frac{1}{m_0}$.

Silly example **2.** Consider a Markov Chain with the states being all nonnegative integer numbers, having transition function

$$P(k, k+1) = \frac{k+1}{k+2}, \qquad P(k, 0) = \frac{1}{k+2}, \qquad k \ge 0.$$

Determine whether state 0 (and the entire chain) is transient, null recurrent, or positive recurrent.

$$P_{0}(T_{0} \ge m) = P(0,1) \times P(1,2) \times P(2,3) \times ... \times P(m-3, m-2) \times P(m-2, m-1)$$

$$= \frac{1}{2} \cdot \frac{2}{3} \cdot \frac{3}{4} \cdot ... \cdot \frac{m-2}{m-1} \cdot \frac{m-1}{m} = \frac{1}{m}.$$

$$\Rightarrow P_0(T_0 = \infty) = \lim_{m \to \infty} P_0(T_0 \ge m) = 0.$$

$$\Rightarrow$$
 $\rho_{00} = P_0(T_0 < \infty) = 1.$ \Rightarrow 0 is recurrent.

$$E_0(T_0) = \sum_{m=1}^{\infty} P(T_0 \ge m) = \sum_{m=1}^{\infty} \frac{1}{m} = \infty.$$

 \Rightarrow 0 is null recurrent.

The Markov chain is irreducible (it is possible to get to any state from any state).

⇒ The Markov chain is null recurrent.

Spoiler: Markov chain is irreducible, null recurrent.

 \Rightarrow There is no stationary distribution:

Silly example **2.5.** Consider a Markov Chain with the states being all nonnegative integer numbers, having transition function

$$P(k, k+1) = \frac{1}{k+2}, \qquad P(k, 0) = \frac{k+1}{k+2}, \qquad k \ge 0.$$

Determine whether state 0 (and the entire chain) is transient, null recurrent, or positive recurrent.

$$P_{0}(T_{0} \ge m) = P(0,1) \times P(1,2) \times P(2,3) \times ... \times P(m-3, m-2) \times P(m-2, m-1)$$
$$= \frac{1}{2} \cdot \frac{1}{3} \cdot \frac{1}{4} \cdot ... \cdot \frac{1}{m-1} \cdot \frac{1}{m} = \frac{1}{m!}.$$

$$\Rightarrow P_0(T_0 = \infty) = \lim_{m \to \infty} P_0(T_0 \ge m) = 0.$$

$$\Rightarrow$$
 $\rho_{00} = P_0(T_0 < \infty) = 1.$ \Rightarrow 0 is recurrent.

$$E_0(T_0) = \sum_{m=1}^{\infty} P(T_0 \ge m) = \sum_{m=1}^{\infty} \frac{1}{m!} = e - 1 < \infty.$$

 \Rightarrow 0 is positive recurrent.

The Markov chain is irreducible (it is possible to get to any state from any state).

⇒ The Markov chain is positive recurrent.

Spoiler: Markov chain is irreducible, positive recurrent.

 \Rightarrow There is stationary distribution:

	0	1	2	3	4	5	6	7	
0	1/2	1/2	0	0	0	0	0	0	•••
1	2/3	0	1/3	0	0	0	0	0	
2	3/4	0	0	1/4	0	0	0	0	
3	4/5	0	0	0	1/5	0	0	0	•••
4	5/6	0	0	0	0	1/6	0	0	
5	6/7	0	0	0	0	0	1/7	0	•••
6	7/8	0	0	0	0	0	0	1/8	•••
•••			•••	•••	•••	•••	•••		•••

$$\pi(0) = \frac{1}{2} \pi(0) + \frac{2}{3} \pi(1) + \frac{3}{4} \pi(2) + \frac{4}{5} \pi(3) + \frac{5}{6} \pi(4) + \frac{6}{7} \pi(5) + \dots,$$

$$\pi(1) = \frac{1}{2} \pi(0) = \frac{1}{2!} \pi(0),$$

$$\pi(2) = \frac{1}{3} \pi(1) = \frac{1}{3!} \pi(0),$$

$$\pi(3) = \frac{1}{4} \pi(2) = \frac{1}{4!} \pi(0),$$

$$\pi(4) = \frac{1}{5}\pi(3) = \frac{1}{5!}\pi(0),$$

$$\pi(5) = \frac{1}{6} \pi(4) = \frac{1}{6!} \pi(0),$$

...

$$\pi(0) + \pi(1) + \pi(2) + \pi(3) + \pi(4) + \pi(5) + \pi(6) + \dots = 1.$$

$$\Rightarrow 1 = \pi(0) + \frac{1}{2!} \pi(0) + \frac{1}{3!} \pi(0) + \frac{1}{4!} \pi(0) + \frac{1}{5!} \pi(0) + \frac{1}{6!} \pi(0) + \dots$$
$$= \pi(0) \cdot \sum_{k=1}^{\infty} \frac{1}{k!} = \pi(0) (e-1).$$

$$\Rightarrow \pi(0) = \frac{1}{e-1}. \qquad \pi(x) = \frac{1}{(x+1)!} \cdot \pi(0) = \frac{1}{(x+1)!} \cdot \frac{1}{e-1}, \quad x \ge 1.$$
That is, $\pi(x) = \frac{1}{(x+1)!} \cdot \frac{1}{e-1}, \quad x \ge 0.$

Note that
$$m_0 = e - 1$$
 and $\pi(0) = \frac{1}{e - 1} = \frac{1}{m_0}$.

Silly example **3.** Consider a Markov Chain with the states being all nonnegative integer numbers, having transition function

$$P(k, k+1) = \frac{\left(1 + \frac{1}{k+1}\right)^{k+1}}{\left(1 + \frac{1}{k+2}\right)^{k+2}}, \qquad P(k, 0) = 1 - \frac{\left(1 + \frac{1}{k+1}\right)^{k+1}}{\left(1 + \frac{1}{k+2}\right)^{k+2}}, \qquad k \ge 0.$$

Determine whether state 0 (and the entire chain) is transient, null recurrent, or positive recurrent.

$$P_0(T_0 \ge m) = P(0,1) \times P(1,2) \times P(2,3) \times ... \times P(m-3, m-2) \times P(m-2, m-1)$$

$$= \frac{2}{\frac{9}{4}} \cdot \frac{\frac{64}{27}}{\frac{64}{27}} \cdot \frac{\frac{64}{27}}{\frac{625}{256}} \cdot \dots \cdot \frac{\left(1 + \frac{1}{m-2}\right)^{m-2}}{\left(1 + \frac{1}{m-1}\right)^{m-1}} \cdot \frac{\left(1 + \frac{1}{m-1}\right)^{m-1}}{\left(1 + \frac{1}{m}\right)^{m}} = \frac{2}{\left(1 + \frac{1}{m}\right)^{m}}.$$

$$\Rightarrow P_0(T_0 = \infty) = \lim_{m \to \infty} P_0(T_0 \ge m) = \frac{2}{e} > 0.$$

$$\Rightarrow \rho_{00} = P_0(T_0 < \infty) = 1 - \frac{2}{e} < 1.$$

 \Rightarrow 0 is transient.

The Markov chain is irreducible (it is possible to get to any state from any state).

⇒ The Markov chain is transient.

Spoiler: Markov chain is irreducible, transient.

 \Rightarrow There is no stationary distribution:

Let π be a stationary distribution.

If x is transient or null recurrent, then $\pi(x) = 0$.

Theorem

An irreducible positive recurrent Markov chain has a unique stationary distribution π , given by

$$\pi(x) = \frac{1}{m_x}, \qquad x \in S.$$

Example:

Winter weather in Central Illinois.

Recall (Examples for 02/08/2022 (2)):

$$\pi(N) = \frac{2}{9}, \qquad \pi(R) = \frac{4}{9}, \qquad \pi(S) = \frac{3}{9}.$$

$$\pi(R) = \frac{4}{9},$$

$$\pi(S) = \frac{3}{9}$$

Recall (Examples for 01/20/2022 (2)):

$$m_{ij} = E_i(T_j)$$

$$m_{\rm NN} = \frac{9}{2}$$

$$m_{\rm NN} = \frac{9}{2}, \qquad m_{\rm RR} = \frac{9}{4}, \qquad m_{\rm SS} = 3.$$

$$m_{\rm SS} = 3$$

$$m_y = E_y(T_y) = m_{yy}$$

$$m_{\rm N}=\frac{9}{2},$$

$$m_{\rm R}=\frac{9}{4}$$

$$m_{\rm N} = \frac{9}{2}, \qquad m_{\rm R} = \frac{9}{4}, \qquad m_{\rm S} = 3 = \frac{9}{3}.$$

$$\pi(x) = \frac{1}{m_x},$$

$$x \in S$$
.