Examples for 04/05/2022

Brownian motion:

- 1) B(0) = 0.
- 2) If $0 \le s_1 \le t_1 \le s_2 \le t_2$, $B(t_1) B(s_1)$ and $B(t_2) B(s_2)$ are independent.
- 3) B(t)-B(s) has a Normal distribution with mean 0 and variance $\sigma^2(t-s)$, $0 \le s \le t$.

(B(t)|B(s)=b) has a Normal distribution with mean with mean b

and variance
$$\sigma^2(t-s)$$
, $0 \le s \le t$.

(B(s)|B(t)=b) has a Normal distribution with mean with mean $\frac{s}{t}b$

and variance
$$\frac{s(t-s)}{t} \sigma^2$$
, $0 \le s \le t$.

$$r_{\mathrm{B}}(s,t) = \mathrm{Cov}(\mathrm{B}(s),\mathrm{B}(t)) = \sigma^2 \min(s,t), \qquad s,t \ge 0.$$

- 1. Let B(t) be a Brownian motion with parameter $\sigma = 5$.
- a) Find the probability P(B(10) < 18).

$$P(B(10) < 18) = P\left(Z < \frac{(18-0)-0}{\sqrt{5^2(10-0)}}\right) \approx P(Z < 1.14) = 0.8729.$$

b) Find the probability $P(B(10) < 18 \mid B(6) = 8)$.

$$P(B(10) < 18 | B(6) = 8) = P\left(Z < \frac{18 - 8}{\sqrt{5^2 (10 - 6)}}\right) = P(Z < 1.00) = 0.8413.$$

c) Find the probability $P(B(6) < 8 \mid B(10) = 18)$.

(B(s)|B(t)=b) has a Normal distribution with mean with mean $\frac{s}{t}b$ and variance $\frac{s(t-s)}{t}\sigma^2$, $0 \le s \le t$.

$$P(B(6) < 8 \mid B(10) = 18) = P\left(Z < \frac{8 - \frac{6}{10} \cdot 18}{\sqrt{\frac{6 \cdot (10 - 6)}{10} \cdot 5^2}}\right) = P\left(Z < \frac{8 - 10.8}{\sqrt{60}}\right)$$

$$\approx P(Z < -0.36) = \mathbf{0.3594}.$$

d) Find the probability P(B(10)-B(6) < 18).

$$P(B(10)-B(6)<18) = P\left(Z<\frac{18-0}{\sqrt{5^2(10-6)}}\right) = P(Z<1.80) = 0.9641.$$

e) Find the probability P(B(10) + B(6) < 18).

$$E(B(10) + B(6)) = 0.$$

$$Var(B(10) + B(6)) = Var(B(10)) + 2 Cov(B(10), B(6)) + Var(B(6))$$
$$= 10 \sigma^2 + 2 \times 6 \sigma^2 + 6 \sigma^2 = 28 \sigma^2 = 700.$$

$$P(B(10) + B(6) < 18) = P(Z < \frac{18 - 0}{\sqrt{700}}) \approx P(Z < 0.68) = 0.7517.$$

Let B(t) be a Brownian motion with parameter σ .

Let T_a denote the first time the Brownian motion process hits a > 0.

$$P(B(t) \ge a) = P(B(t) \ge a \mid T_a \le t) \cdot P(T_a \le t) + P(B(t) \ge a \mid T_a > t) \cdot P(T_a > t).$$

By symmetry,
$$P(B(t) \ge a \mid T_a \le t) = \frac{1}{2}$$
.

By continuity,
$$P(B(t) \ge a \mid T_a > t) = 0.$$

Thus
$$F_{T_a}(t) = P(T_a \le t) = 2P(B(t) \ge a) = 2P(Z \ge \frac{a}{\sigma \sqrt{t}})$$

$$= \int_{\frac{a}{\sigma \sqrt{t}}}^{\infty} \frac{2}{\sqrt{2\pi}} e^{-z^2/2} dz, \qquad t > 0.$$

$$f_{T_a}(t) = F_{T_a}'(t) = \frac{2}{\sqrt{2\pi}} e^{-a^2/(2\sigma^2 t)} \cdot \left(\frac{a}{2\sigma\sqrt{t^3}}\right)$$
$$= \frac{a}{\sqrt{2\pi}\sigma t^{3/2}} e^{-a^2/(2\sigma^2 t)}, \qquad t > 0.$$

If W has a Normal distribution with mean 0 and variance $\frac{\sigma^2}{a^2}$, then T_a has the same distribution as $\frac{1}{W^2}$.

$$P(T_a > t) = 1 - F_{T_a}(t) = P(-\frac{a}{\sigma \sqrt{t}} < Z < \frac{a}{\sigma \sqrt{t}}) \rightarrow 0 \text{ as } t \rightarrow \infty.$$

$$P(T_a = \infty) = \lim_{t \to \infty} P(T_a > t) = 0. \qquad \Rightarrow \qquad P(T_a < \infty) = 1.$$

Fact: Let X be a nonnegative continuous random variable with p.d.f. f(x) and c.d.f. F(x). Then

$$E(X) = \int_{0}^{\infty} (1 - F(x)) dx.$$

Proof:

$$E(X) = \int_{0}^{\infty} x f(x) dx = \int_{0}^{\infty} \left(\int_{0}^{x} dy \right) f(x) dx = \int_{0}^{\infty} \left(\int_{0}^{x} f(x) dy \right) dx$$

$$\int_{0}^{\infty} \left(\int_{0}^{x} f(x) dy \right) dx = \int_{0}^{\infty} \left(\int_{y}^{\infty} f(x) dx \right) dy$$

$$\sum_{0}^{\infty} \left(\int_{0}^{\infty} f(x) dy \right) dx = \int_{0}^{\infty} \left(\int_{y}^{\infty} f(x) dx \right) dy$$

$$\Rightarrow E(X) = \int_{0}^{\infty} \left(\int_{y}^{\infty} f(x) dx \right) dy = \int_{0}^{\infty} P(X > y) dy = \int_{0}^{\infty} (1 - F(y)) dy.$$

$$\begin{split} \mathrm{E}(\mathrm{T}_a) &= \int\limits_0^\infty \left(1 - \mathrm{F}_{\mathrm{T}_a}(t)\right) dt = \int\limits_0^\infty \int\limits_0^{\frac{a}{\sigma\sqrt{t}}} \frac{2}{\sqrt{2\pi}} \, e^{-z^2/2} dz \, dt \\ &= \int\limits_0^\infty \int\limits_0^{\frac{a^2}{\sigma^2 z^2}} \frac{2}{\sqrt{2\pi}} \, e^{-z^2/2} dt \, dz = \int\limits_0^\infty \frac{2}{\sqrt{2\pi}} \, \frac{a^2}{\sigma^2 z^2} \, e^{-z^2/2} dz = \infty. \end{split}$$

This is not surprising, even though $P(T_a < \infty) = 1$, since a Brownian motion is the limit of a (null recurrent) symmetric random walk.

Let
$$M(t) = \max_{0 \le s \le t} B(s), t > 0.$$

By continuity, $P(M(t) \ge a) = P(\max_{0 \le s \le t} B(s) \ge a) = P(T_a \le t), a > 0, t \ge 0.$

$$F_{M(t)}(a) = 1 - P(\max_{0 \le s \le t} B(s) \ge a) = 1 - P(T_a \le t) = 1 - \int_{a}^{\infty} \frac{2}{\sqrt{2\pi}} e^{-z^2/2} dz,$$

$$a > 0.$$

$$f_{M(t)}(a) = \frac{2}{\sqrt{2\pi}} e^{-a^2/(2\sigma^2 t)} \cdot \left(\frac{1}{\sigma\sqrt{t}}\right), \qquad a > 0, \ t \ge 0.$$

 $\max_{0 \le s \le t} B(t)$ has the same probability distribution as |B(t)|.

Example: Let B(t) be a Brownian motion with parameter $\sigma^2 = 1$. Then the process $X(t) = |B(t)|, t \ge 0$, is called Brownian motion reflected at the origin.

a) Find the density function of X(t), $t \ge 0$.

$$F_{X(t)}(x) = P(X(t) \le x) = P(-x \le B(t) \le x) = F_{B(t)}(x) - F_{B(t)}(-x).$$
 $x > 0.$

$$f_{X(t)}(x) = F'_{X(t)}(x) = f_{B(t)}(x) + f_{B(t)}(-x)$$

$$= \frac{1}{\sqrt{2\pi t}} e^{-x^2/2t} + \frac{1}{\sqrt{2\pi t}} e^{-(-x)^2/2t}$$

$$= \frac{2}{\sqrt{2\pi t}} e^{-x^2/2t}, \qquad x > 0.$$

b) Find E[X(t)] and Var[X(t)].

Hint: Compute E[X(t)] directly after $u = x^2$ and $E[(X(t))^2]$ by comparing the integral with the integral representing the variance of a random variable that is N(0, t).

$$E[X(t)] = \int_{0}^{\infty} x \frac{2}{\sqrt{2\pi t}} e^{-x^{2}/2t} dx = \dots$$

$$u = \frac{x^{2}}{2t} \qquad du = \frac{1}{t} x dx$$

$$\dots = \int_{0}^{\infty} \frac{2\sqrt{t}}{\sqrt{2\pi}} e^{-u} du = \frac{2\sqrt{t}}{\sqrt{2\pi}} = \frac{\sqrt{2t}}{\sqrt{\pi}}.$$

$$E[(X(t))^{2}] = \int_{0}^{\infty} x^{2} \frac{2}{\sqrt{2\pi t}} e^{-x^{2}/2t} dx = \int_{-\infty}^{\infty} x^{2} \frac{1}{\sqrt{2\pi t}} e^{-x^{2}/2t} dx$$
$$= E(W^{2}) = t, \qquad \text{where } W \sim N(0, t).$$

OR

$$E[(X(t))^{2}] = \int_{0}^{\infty} x^{2} \frac{2}{\sqrt{2\pi t}} e^{-x^{2}/2t} dx = \dots$$

$$u = \frac{x^{2}}{2t} \qquad du = \frac{1}{t} x dx$$

$$\dots = \int_{0}^{\infty} \frac{2t}{\sqrt{\pi}} \sqrt{u} e^{-u} du = \frac{2t}{\sqrt{\pi}} \int_{0}^{\infty} u^{1/2} e^{-u} du = \frac{2t}{\sqrt{\pi}} \Gamma\left(\frac{3}{2}\right)$$

$$= \frac{2t}{\sqrt{\pi}} \cdot \frac{1}{2} \Gamma\left(\frac{1}{2}\right) = t.$$

$$Var[X(t)] = E[(X(t))^2] - [E(X(t))]^2 = t - \frac{2t}{\pi} = t \left(1 - \frac{2}{\pi}\right).$$

- 2. Let B(t) be a Brownian motion.
- a) $P(T_2 < T_{-1})$.

$$P(T_2 < T_{-1}) = P(up \ A = 2 \ before \ down \ B = 1) = \frac{B}{A + B} = \frac{1}{2 + 1} = \frac{1}{3}.$$

b) $P(T_2 < T_{-1} < T_3).$

$$P(T_2 < T_{-1} < T_3) = P(\text{first, up 2 before down 1,}$$

then (from 2) down 3 (to -1) before up 1 (to 3))
$$= \frac{1}{2+1} \times \frac{1}{1+3} = \frac{1}{12}.$$