

Bioinformatics and Statistical Genetics

Elective specialization for MDS/MIRI/MAI students

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**UNIVERSITAT POLITÈCNICA
DE CATALUNYA
BARCELONATECH**

Syllabus

Bioinformatics and **Statistical Genetics**

1. Introduction to statistical genetics
2. Hardy-Weinberg equilibrium
3. Linkage disequilibrium and haplotype estimation
4. Population substructure
5. Genetic association analysis
6. Relatedness analysis (allele sharing)

5 December 2023

Content

Genetic Association Studies

1. Introduction
2. Allele based tests
3. Genotype based tests
4. Quantitative traits
5. Multiple polymorphisms
6. Computer exercise

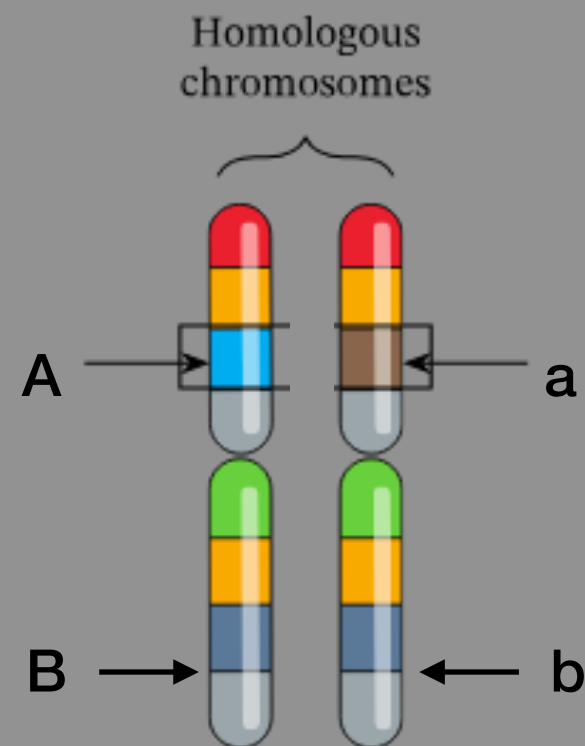
Genetic association studies

Explain Like I'm 5 - RECALL

The **Hardy-Weinberg Equilibrium** is fundamental in population genetics: **Linkage disequilibrium**

Haplotype estimation

Population substructure



Finally....phenotype-marker
association studies!

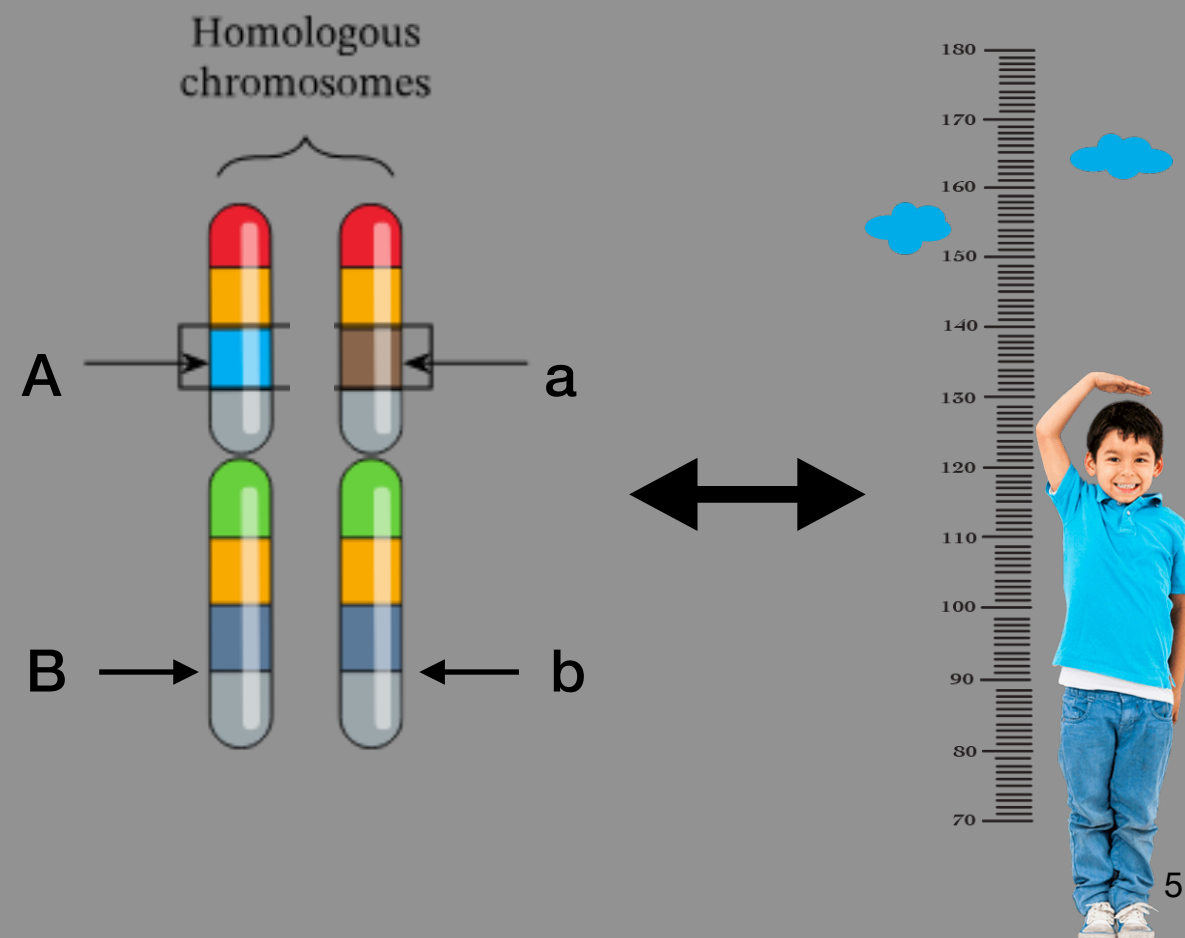
Genetic association studies

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Finally....phenotype-marker
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Genetic association studies

Why? (Still ELI5)

To what extent certain traits are genetic?

What are the genetic mechanisms of that trait?

To find gene-trait association that allows for:

- **Disease diagnosis:** determines if a particular person has a disease.
- **Prognosis:** determines if a particular person will develop the disease in the future.
- **Conversion:** determines if a particular person converts from one stage to another stage of a disease.
- **Therapy:** determines if a therapy is successful.
- **Aging:** determines general processes of aging but not the disease, determines the biological age of a person.
- **Individualised therapy:** determines the likelihood of benefiting from a specific therapy.
- ...

Genetic association studies

Introduction

RECALL:

A **trait (phenotype)** is a specific characteristic of an individual. Traits can be determined by genes, environmental factors or by a combination of both. Traits can be qualitative (such as eye color) or quantitative (such as height or blood pressure).

Goal:

Investigate associations between markers and a trait (disease).

NOTE:

Genetic association studies test for a **correlation** between a trait (usually disease status) and genetic variation to identify candidate genes or genome regions that **contribute** to that specific trait/disease.

Designs:

- Unrelated subjects (population-based)
- Related subjects from pedigrees (family-based)

Family-based association studies are often aimed at finding rare variants underlying rare conditions or rare sub-phenotypes of a common condition.

Genetic association studies

Introduction

Goal:

Investigate associations between markers and a trait (disease).

Early studies investigated rare conditions that show clear Mendelian segregation through families, and very successfully located those genetic variations because they carry 100% of the risk

Examples:

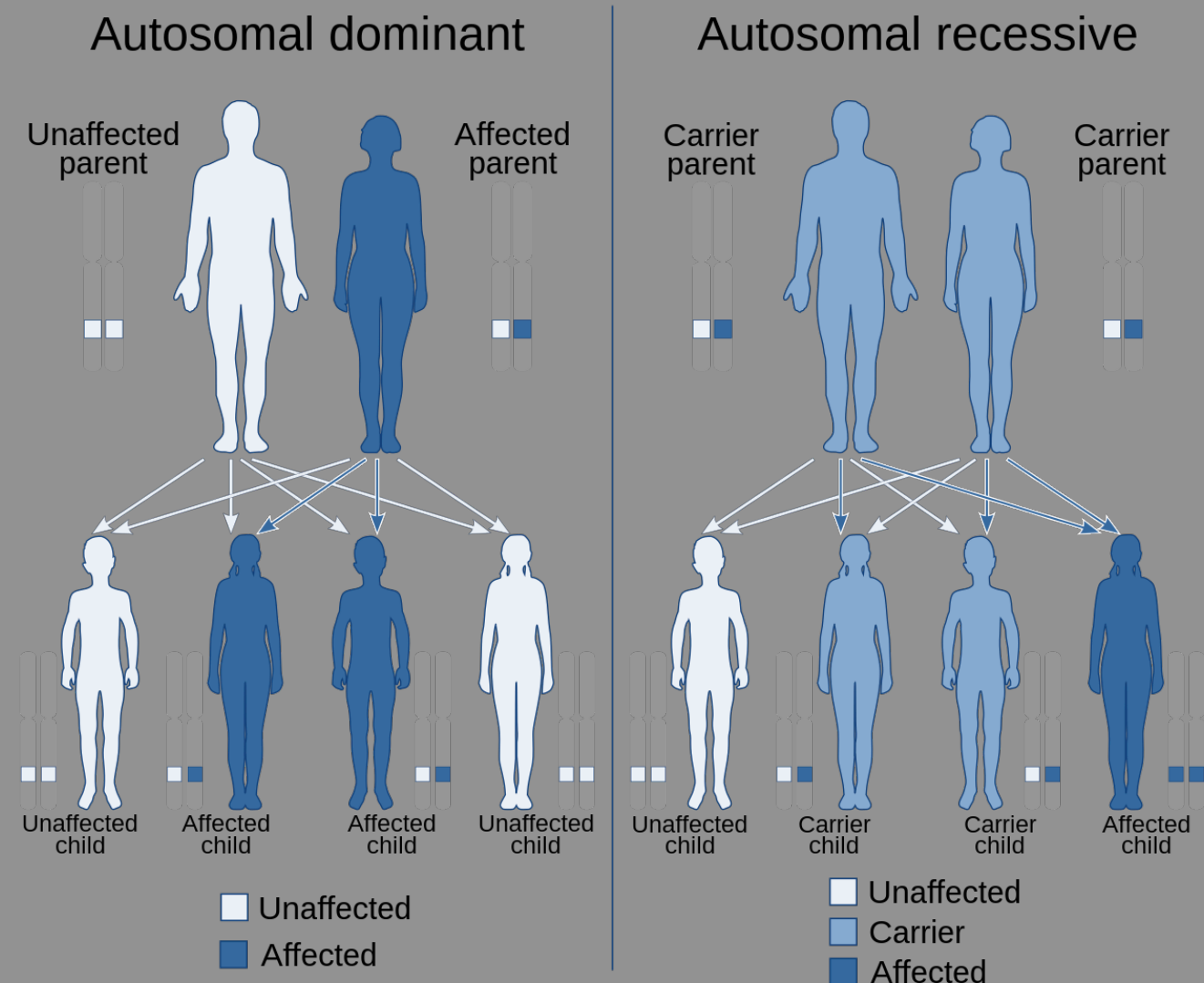
Huntington's disease is a neurodegenerative disease where patients show variety of movement disorders, with involuntary and uncoordinated body movements.

Common autosomal dominant inheritance disorders

- Huntington's disease
- Marfan syndrome (cardiovascular)
- Achondroplasia
- ...about 200 disorders

Common autosomal recessive inheritance disorders

- Spinal muscular atrophy
- Cystic fibrosis
-about 450 disorders
-



Genetic association studies

Introduction

Goal:

Investigate associations between markers and a trait (disease).

Complex diseases are caused by a combination of genetic, environmental, and lifestyle factors,and determining the heritability (genetic contribution to the trait) its difficult as there is no clear mendelian inheritance patterns.

- In almost any complex trait that has been studied, **many loci contribute to standing genetic variation**...so that mutations in many genes contribute to genetic variation in the population.
- On average, the proportion of variance explained at the individual variants is small.
- Each variant is only one of the many genetic and environmental causal factors, each of which are neither necessary nor sufficient to individually cause the disease. Thus, they **predispose** to—rather than directly result in—its development.

Examples:

Complex diseases are also referred as multi-factorial traits, non-communicable diseases or chronic diseases

- Diabetes
- Stroke
- Asthma
- Obesity
- Hypothyroidism
- Cancer
- Schizophrenia
- Depression
- Epilepsy
- ...

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Genetic association studies

Introduction - population-based studies

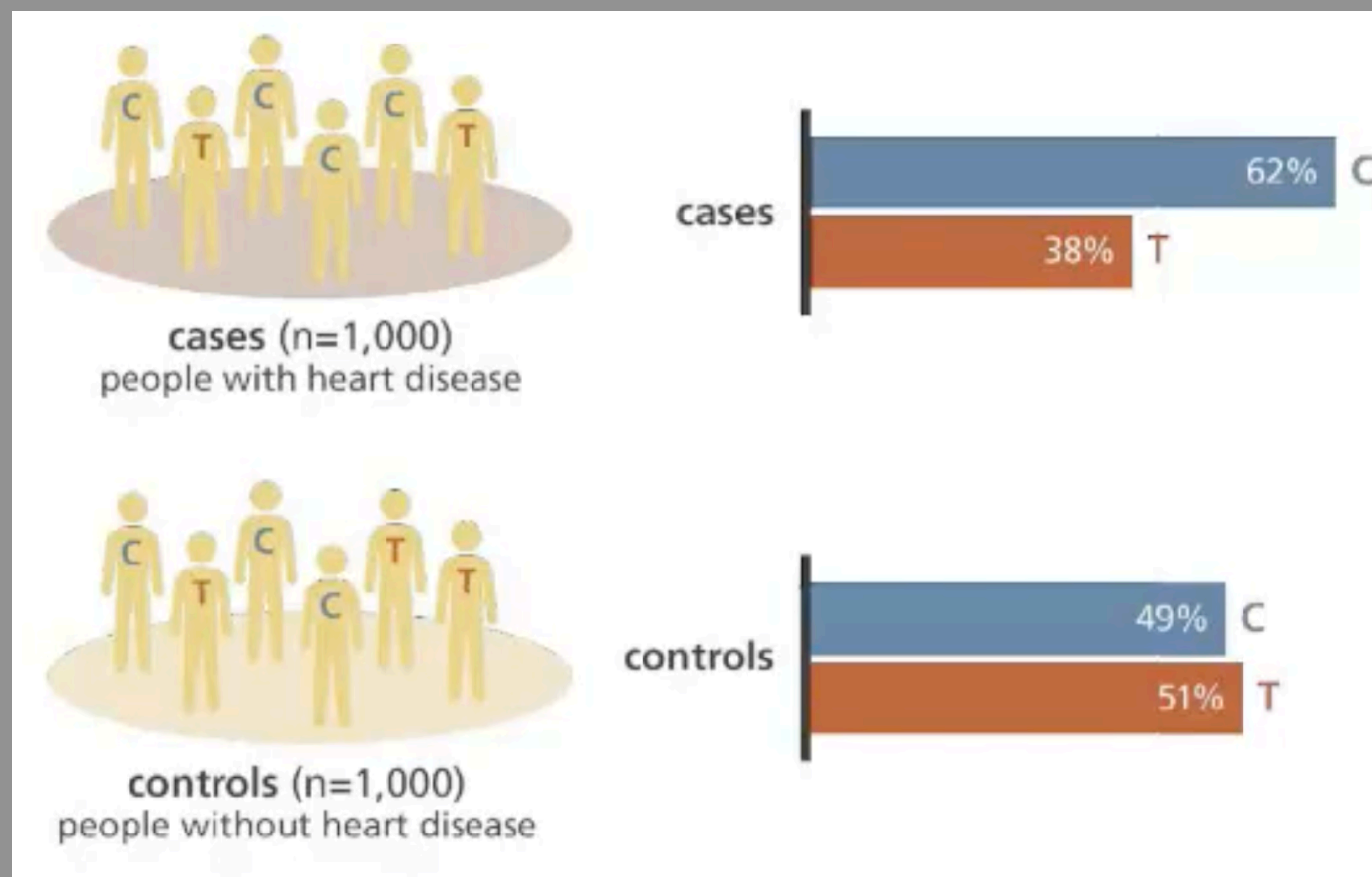
- We will focus on population-based association studies, executed on unrelated subjects. Two types:
 - **Allele-based tests and genotype-based tests:** are hypothesis driven, where a candidate locus is being tested.
 - **Multiple polymorphisms:** involve genome wide association studies conducted without prior hypothesis.

	id	rs34684677	rs1839115	rs4727804	rs4727805	rs200888633	rs12534908
1	NA18939	T/G	C/T	G/A	T/G	T/G	G/A
2	NA18940	G/G	T/T	A/A	G/G	T/G	A/A
3	NA18941	G/G	T/T	A/A	G/G	T/G	A/A
4	NA18942	G/G	T/T	A/A	G/G	T/T	A/A
5	NA18943	G/G	T/T	A/A	G/G	T/T	A/A
6	NA18944	T/T	C/C	G/G	T/G	G/G	G/G
7	NA18945	G/G	T/T	A/A	G/G	G/G	A/A
8	NA18946	T/G	C/T	G/A	G/G	G/G	G/A
9	NA18947	T/G	C/T	G/A	G/G	T/G	G/A
10	NA18948	G/G	T/T	A/A	G/G	G/G	A/A
11	NA18949	T/G	C/T	G/A	T/G	T/G	G/A
12	NA18950	G/G	T/T	A/A	G/G	T/G	A/A
13	NA18951	G/G	T/T	A/A	G/G	T/G	A/A
14	NA18952	T/G	C/C	G/G	T/G	T/G	G/G

Genetic association studies

Introduction - population-based studies

- We will focus on population-based association studies, executed on unrelated subjects. Two types:
 - **Allele-based tests and genotype-based tests:** are hypothesis driven, where a candidate locus is being tested with a **case-control studies**.



The success rate of candidate gene case-control studies has been very poor. In 2002:

- 603 published disease-genetic variant associations found that only 6 appeared to be independently replicated

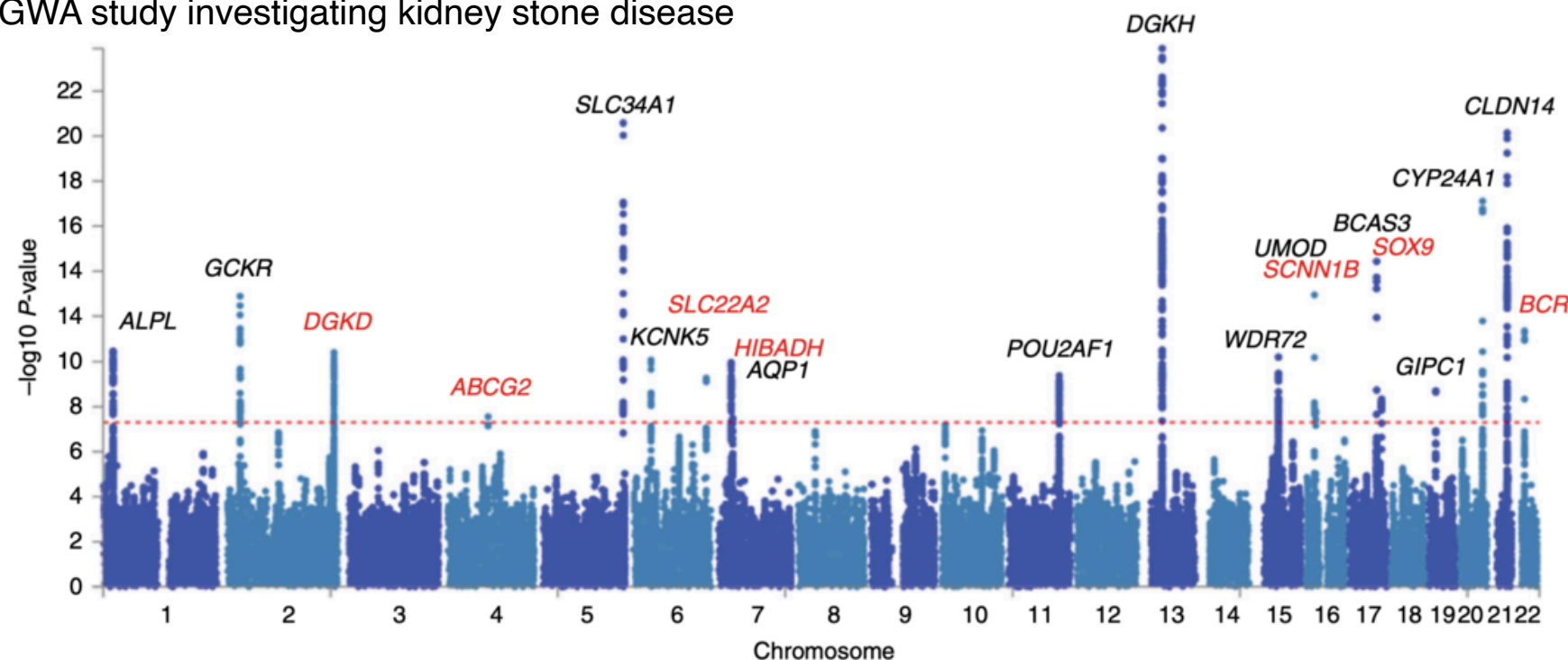
Complex traits are not caused by common genetic polymorphisms but by multiple rare ones

Genetic association studies

Introduction - population-based studies

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GWA study investigating kidney stone disease



Typically, a single test statistic (for case-control studies, a chi-squared (χ^2) comparison of absolute genotype counts) is calculated for each variant passing quality control.

Increasing number of studies are being extended from case-control studies to population-based cohorts.

Manhattan plot depicting several strongly associated risk loci. Each dot represents a SNP, with the X-axis showing genomic location and Y-axis showing association level. The peaks indicate genetic variants that are found more often in individuals with kidney stones.

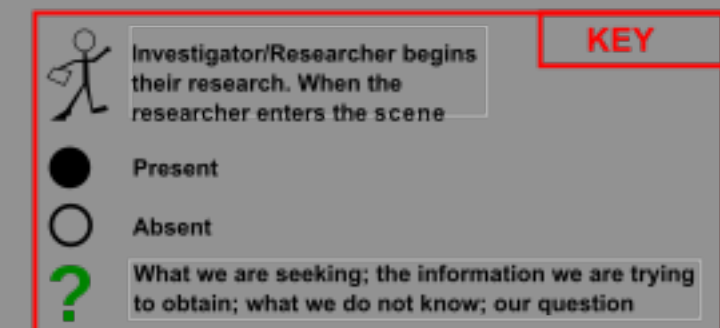
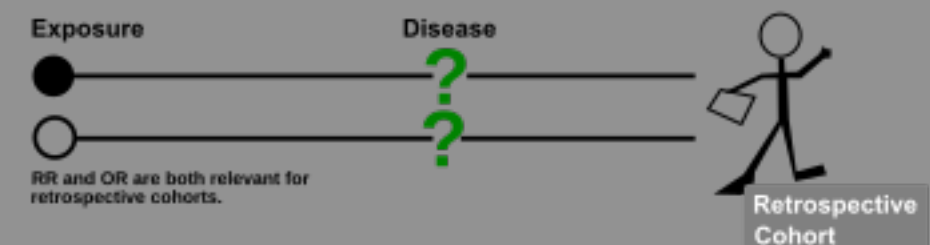
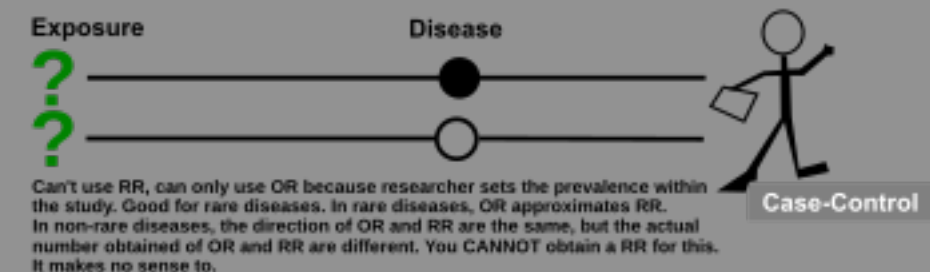
When we are looking for regions of the genome or SNP that is causal for a gene, we often find that a whole bunch of SNPs are associated with the disease. It's not that they all cause disease, it is just that a whole bunch are correlated with the causal SNP (passenger mutations). Thus it is our job to identify the causal needle in the haystack.

Genetic association studies

Introduction - some notes on study designs

- **Case-control studies:** observational study in which two groups that only differ in the trait/outcome of interest are compared.
- **Cohort studies:** a type of longitudinal (and observational) study that recruit participants of a specific population and observe them over time.
- Cross-sectional studies: observational (and descriptive) study that involve the analysis of a data set in a particular point in time. Can't inform about causal relationship with a trait nor the risk of a disease.
- Other study designs may involve the analysis of an intervention (experimental studies):
 - Randomized Controlled Trials
 - Uncontrolled Trials
- The choice of study design or data resource depends on the required sample size, the experimental question, the availability of existing data or...the ease by which new data can be obtained.

Observational Study Designs: Case Control vs Cohort



OR: odds ratio
RR: relative risk

Genetic association studies

Introduction - population-based studies

Tests of association at the level of **alleles**

- We are sampling alleles
- Alleles assumed to be independent
- Rely on the Hardy-Weinberg equilibrium assumption
- Statistics on the alleles by trait cross table
 - Chi-square test
 - Fisher exact test
 - Odds ratio

Tests of association at the level of the **genotypes**

- We are sampling individuals
- Hardy-Weinberg equilibrium assumption is not needed
- Co-dominant, dominant and recessive Chi-square tests
- Cochran-Armitage trend test
- Logistic regression
- Multiple regression and mixed effects models

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Genetic association studies

The data

The trait (Y_i) (e.g. disease) we wish to understand is binary (dichotomous) so that:

- $Y_i = 1$ individual i has the trait
- $Y_i = 0$, individual i does not have the trait.

The marker is a bi-allelic polymorphism (e.g. AA, Aa and aa)

The data table:

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n

r_0 , r_1 and r_2 refer to the number of observed individuals with genotype aa, aA and AA respectively, in the case group

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Genetic association studies


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Risk of developing the disease for each allele : $R_{aa} = \frac{r_0}{n_0}$, $R_{aA} = \frac{r_1}{n_1}$ and $R_{AA} = \frac{r_2}{n_2}$

Genetic association studies

Allele based test

- Let p be the allele frequency of the A allele and $P_{cases}(A)$ the frequency of allele A in the case group and $P_{control}(A)$ the frequency of allele A in the control group.

- Hypothesis:

$$H_0 : P_{cases}(A) = P_{control}(A)$$

$$H_0 : P_{cases}(A) \neq P_{control}(A)$$

- The test assumes Hardy-Weinberg equilibrium.
- The test is a χ^2 test for independence in a 2×2 table of alleles, which can be obtained from the genotype count table.

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n

	a	A	Total	\hat{p}
Cases	$r_a = 2r_0 + r_1$	$r_A = 2r_2 + r_1$	$2r$	$r_A/(2r)$
Controls	$s_a = 2s_0 + s_1$	$s_A = 2s_2 + s_1$	$2s$	$s_A/(2s)$
Total	$n_a = 2n_0 + n_1$	$n_A = 2n_2 + n_1$	$2n$	$n_A/(2n)$

- Statistical tests:
 - Chi square test for independence
 - Fisher's exact test
- Odds ratio for the effect size

Genetic association studies

Allele based test - Pearson's χ^2 test

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Total	$n_a = 2n_0 + n_1$	$n_A = 2n_2 + n_1$	$2n$	$n_A/(2n)$

- Chi square test for independence

$$X^2 = \sum_{i,j} \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

Expected count:

$$e_{ij} = \text{total row } i \times \text{total column } j / \text{total of table}$$

- If H_0 true, then $X^2 \sim \chi_1^2$

Genetic association studies

Allele based test - Pearson's χ^2 test

Example

- A polymorphism in the Dopamine receptor is supposed to be involved in Schizophrenia. In a case-control study, the following data were obtained:

	11	12	22	Total		1	2	Total
Cases	7	69	57	133	Cases	83	183	266
Controls	20	56	33	109	Controls	96	122	218
Total	27	125	90	242	Total	179	305	484

$$r_1 = 2 \cdot r_{11} + r_{12} = 2 \cdot 7 + 69 = 83$$

$$r_2 = 2 \cdot r_{22} + r_{12} = 2 \cdot 57 + 69 = 183$$

$$\dots$$

- The test is a χ^2 test for independence in a 2×2 table of alleles.

$$X^2 = \sum_{i,j} \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

$$X^2 = \frac{(83 - 98.38)^2}{98.38} + \dots + \frac{(122 - 137.38)^2}{137.38} = 8.4671$$

$$p\text{-value} = P(\chi^2 \leq 8.4761) = 0.0036$$

$$e_{11} = 266 \cdot 179/484 = 98.38$$

$$e_{12} = 218 \cdot 179/484 = 80.62$$

$$\dots$$

	1	2	Total
Cases	98.38	167.62	266
Controls	80.62	137.38	218
Total	179	305	484

Expected allele count:

$$e_{ij} = \text{total row } i \times \text{total column } j / \text{total of table}$$

We REJECT the null hypothesis that $H_0 : P_{cases}(A) = P_{control}(A)$

Genetic association studies

Allele based test - Pearson's χ^2 test

Example in R

- A polymorphism in the Dopamine receptor is supposed to be involved in Schizophrenia. In a case-control study, the following data were obtained:

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- The test is a χ^2 test for independence in a 2×2 table of alleles.

```
> X <- matrix(c(7,69,57,20,56,33),byrow=TRUE,ncol=3)
> colnames(X) <- c("11","12","22")
> rownames(X) <- c("Cases","Controls")

> Y <- cbind(2*X[,1]+X[,2],2*X[,3]+X[,2])
> colnames(Y) <- c("1","2")

> chisq.test(Y,correct=FALSE)
Pearson's Chi-squared test
data: Y
X-squared = 8.4671, df = 1, p-value = 0.003616
```

Genetic association studies

Allele based test - Fisher exact test

- Recall: $H_0 : P_{cases}(A) = P_{control}(A)$
- Calculation of Fisher's exact test involves direct calculation of the probability P from the number of samples observed n and its counts.
- Often used for tables with low counts with a 2x2 matrix, or when $e_{ij} < 5$.
- Like the χ^2 test, events must be independent

Example in R (same data)

```
>Y
      1  2
cases  83 183
controls 96 122

> fisher.test(Y)
Fisher's Exact Test for Count Data
data: Y

p-value = 0.00448

alternative hypothesis: true odds ratio is not equal to 1
95 percent confidence interval:
0.3903016 0.8512261
sample estimates: odds ratio 0.5770451
```

Genetic association studies

Allele based test - Odds ratio

- Odds of an event p :

$$Odds(p) = \frac{p}{1-p} \quad \text{for example:} \quad Odds(disease) = \frac{P(disease)}{P(nodisease)}$$

- The odds ratio (OR) compares the odds of an event in two groups and provides a relative measure of effect in case-control studies. Quantifies the strength of an association between two events. Odds can range from 0 to infinity.
- The OR compares the odds of the disease for the two alleles:

	A	B
Cases	n_{11}	n_{12}
Controls	n_{21}	n_{22}

$$OR = \frac{n_{11}/n_{21}}{n_{12}/n_{22}} = \frac{n_{11} \cdot n_{22}}{n_{12} \cdot n_{21}}$$

Odds of disease with allele A
 Odds of disease with allele B

- OR = 1 independence
- OR > 1 allele A associated with higher odds of an outcome
- OR < 1 allele A associated with lower odds of an outcome

- The distribution of the log odds ratio $L(OR)$ is approximately normal $L(OR) \sim N(\ln(OR), V(\ln(OR)))$
- Variance of OR is known, which allows for the calculation of confidence intervals for the OR

$$V(\ln(OR)) = \frac{1}{n_{11}} + \frac{1}{n_{12}} + \frac{1}{n_{21}} + \frac{1}{n_{22}}$$

Genetic association studies

Allele based test - Odds ratio

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- Odds ratio

$$OR = \frac{n_{11}/n_{21}}{n_{12}/n_{22}} = \frac{n_{11} \cdot n_{22}}{n_{12} \cdot n_{21}} = \frac{83 \cdot 122}{96 \cdot 183} = 0.576$$

Odds that a person with allele 1 will suffer the disease to the odds that a person will still get a disease with allele 2.

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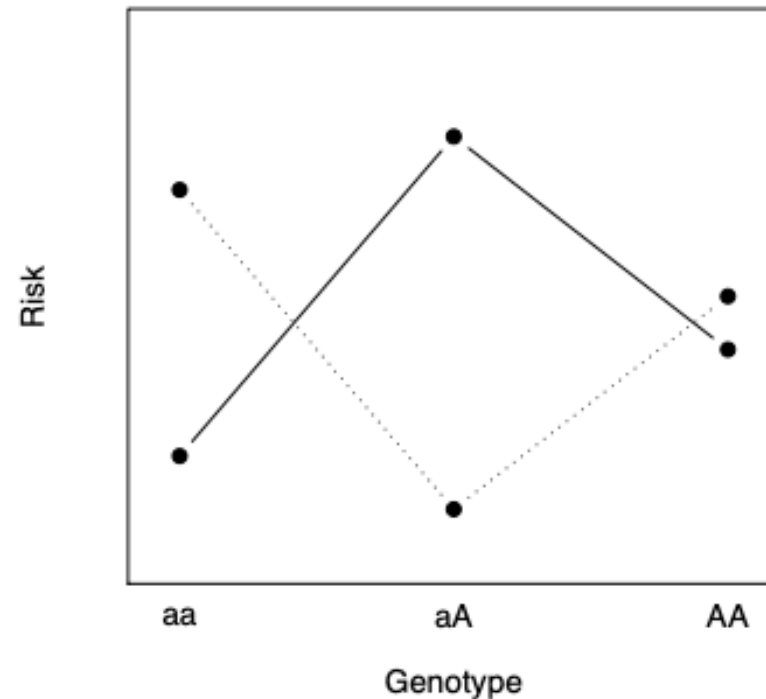
We can test for association using different genetic models:

- Codominant model
- Dominant model
- Recessive model
- Additive model

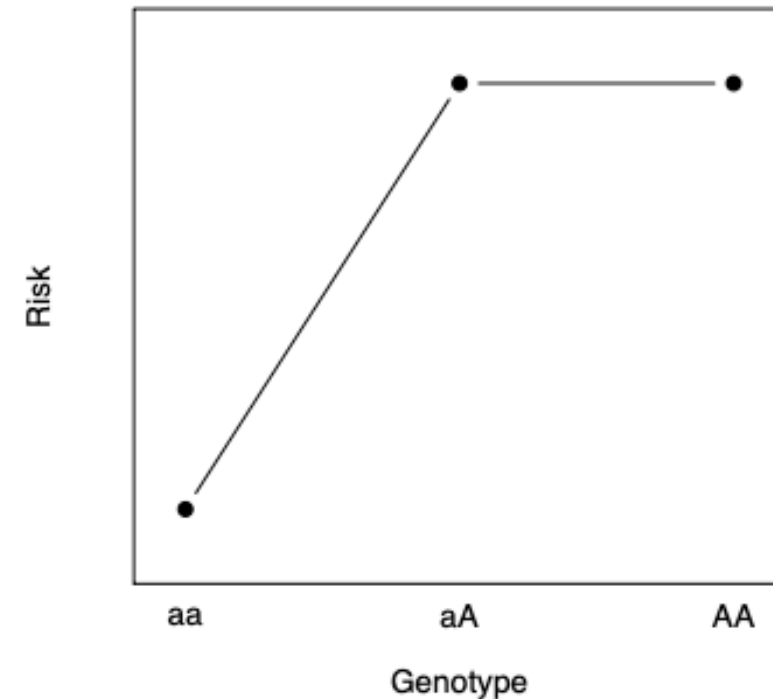
Genetic association studies

Genotype based tests

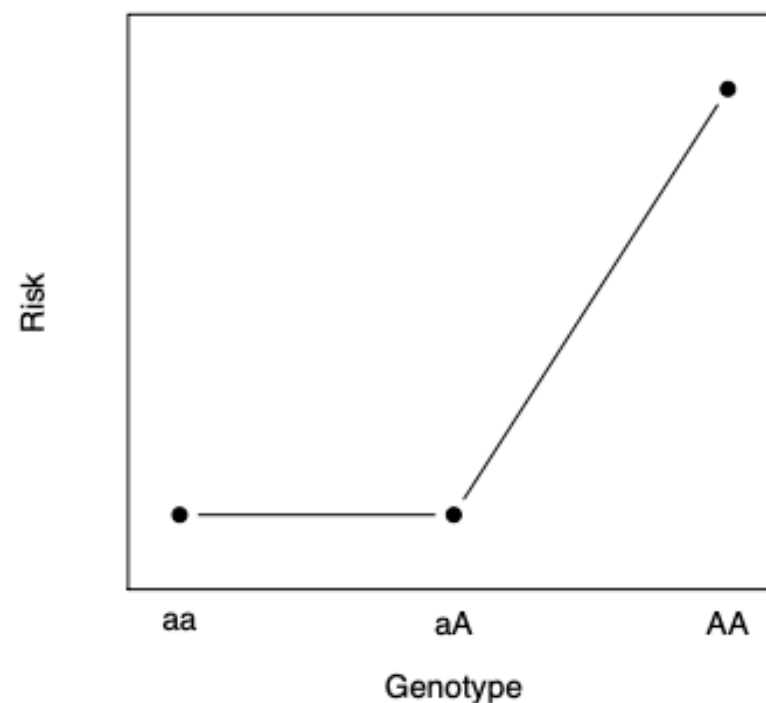
Co-dominant model



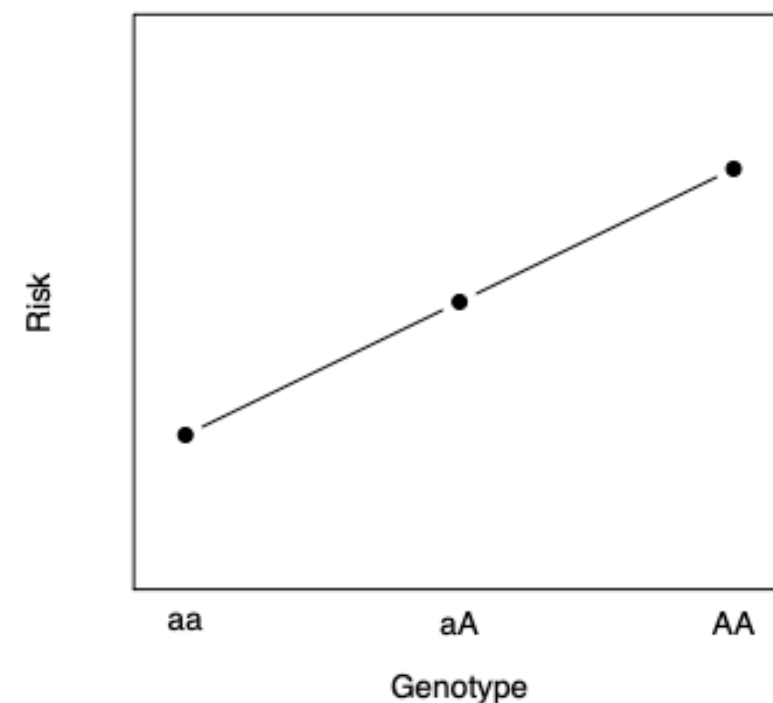
Dominant model



Recessive model



Additive model



Genetic association studies

Genotype based tests - Codominant test

- We test the null hypothesis of no effect of the marker on the trait. Formally:

- $$H_0 : P(Y = 1 | AA) = P(Y = 1 | aA) = P(Y = 1 | aa)$$

- $$H_1 : \text{at least one pair different}$$

- Test statistic:

$$X^2 = \sum_{i,j} \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

- Under H_0 , the test statistic X^2 follows a χ^2_2 distribution (df=2)
- The test makes no assumptions about the relationship between genotype and trait.
- Under H_1 , each genotype can have a different disease rate.
- The test can reject the null if the data support heterozygote advantage (overdominance).

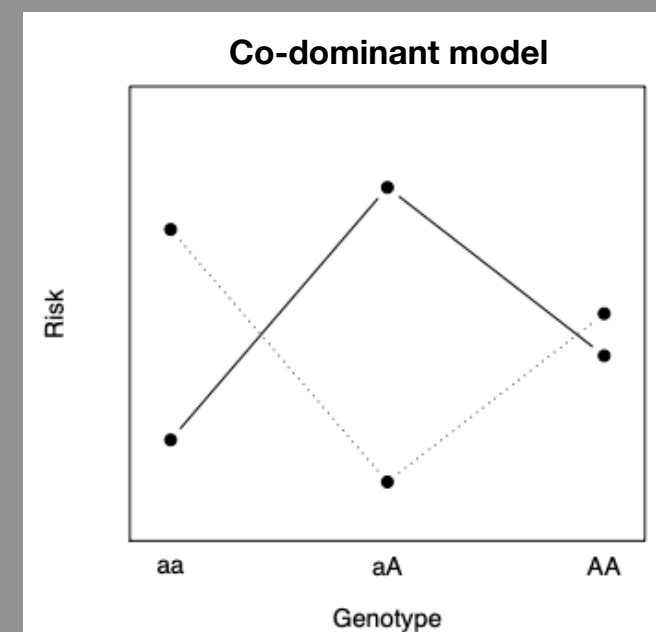
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Genetic association studies

Genotype based tests - Codominant

Example

- TNF genotype (G/A polymorphism) is on a study on acne patients and controls

- Estimate $\hat{p}(Y = 1) = \frac{r}{n} = \frac{113}{227} = 0.497$

- The expected counts for cases:

- $exp(Y = 1 | aa) = \hat{p}(Y = 1) \cdot n_0 = \frac{r \cdot n_0}{n} = 113 \cdot 165/227 = 82.136$

- $exp(Y = 1 | aA) = \hat{p}(Y = 1) \cdot n_1 = \frac{r \cdot n_1}{n} = 113 \cdot 58/227 = 28.872$

- $exp(Y = 1 | AA) = \hat{p}(Y = 1) \cdot n_2 = \frac{r \cdot n_2}{n} = 113 \cdot 4/227 = 1.991$

- The expected counts for controls:

- $exp(Y = 0 | aa) = (1 - \hat{p}(Y = 1)) \cdot n_0 = \frac{s \cdot n_0}{n} = 114 \cdot 165/227 = 82.863$

- $exp(Y = 0 | aA) = (1 - \hat{p}(Y = 1)) \cdot n_1 = \frac{s \cdot n_1}{n} = 114 \cdot 58/227 = 29.127$

- $exp(Y = 0 | AA) = (1 - \hat{p}(Y = 1)) \cdot n_2 = \frac{s \cdot n_2}{n} = 114 \cdot 4/227 = 2.008$

- The chi-square statistics:

- $X^2 = \sum_{i,j} \frac{(observed - expected)^2}{expected} = \frac{(66 - 82.136)^2}{82.136} + \dots + \frac{(0 - 2.008)^2}{2.008} = 24.113$

- P-value = $P(\chi^2_2 \geq 24.113) = 5.806e - 06$

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n

	GG	GA	AA	Total
Cases	66	43	4	113
Controls	99	15	0	114
Total	165	58	4	227

Genetic association studies

Genotype based tests - Codominant test

Example - R code

- TNF genotype (G/A polymorphism) is on a study on acne patients and controls

	GG	GA	AA	Total
Cases	66	43	4	113
Controls	99	15	0	114
Total	165	58	4	227

```
> X <- matrix(c(66,43,4,99,15,0),byrow=TRUE,ncol=3)
> colnames(X) <- c("GG","GA","AA")
> rownames(X) <- c("Acne","Control")
> X
```

```
      GG GA AA
Acne   66 43  4
Control 99 15  0
```

```
> results <- chisq.test(X)
Warning message:
In chisq.test(X) : Chi-squared approximation may be incorrect
```

```
> print(results)
```

Pearson's Chi-squared test

X-squared = 24.1133, df = 2, **p-value = 5.806e-06**

```
> results$expected
      GG      GA      AA
Acne  82.13656 28.87225 1.991189
Control 82.86344 29.12775 2.008811
```

```
> fisher.test(X)
Fisher's Exact Test for Count Data
data: X
p-value = 1.97e-06
alternative hypothesis: two.sided
```

We reject the null hypothesis that the probability of disease with all the genotypes is the same

Genetic association studies

Genotype based tests - Dominant test

- Columns in the original data table are combined to reflect the dominant model.

	aa	aA or AA	Total
Cases	r_0	$r_1 + r_2$	r
Controls	s_0	$s_1 + s_2$	s
Total	n_0	$n_1 + n_2$	n

- Hypothesis:
 - $H_0 : P(Y = 1 | aa) = P(Y = 1 | (aA + AA))$, so that the disease does not depend on presence of A
 - $H_1 : \text{disease depends on the presence of A}$
- Test statistic:

$$X^2 = \sum_{\text{genotypes}} \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

- Under H_0 , the test statistic X^2 follows a χ_1^2 distribution (df=1)

The trait (Y_i) (e.g. disease) we wish to understand is binary (dichotomous) so that:

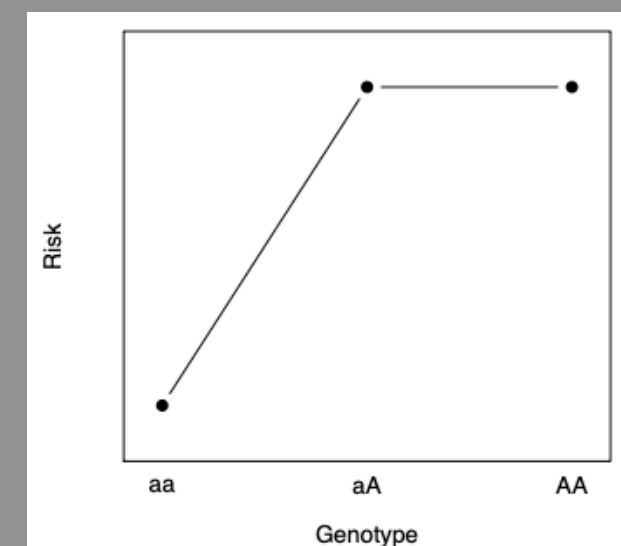
- $Y_i = 1$ individual i has the trait
- $Y_i = 0$, individual i does not have the trait.

The marker is a bi-allelic polymorphism (e.g. AA, Aa and aa)

The original data table:

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n

Dominant model



Genetic association studies

Genotype based tests - Dominant test

Example - R code

- TNF genotype (G/A polymorphism) is on a study on acne patients and controls

	GG	GA	AA	Total
Cases	66	43	4	113
Controls	99	15	0	114
Total	165	58	4	227

```
> Y <- cbind(X[,1],X[,2]+X[,3])
> colnames(Y) <- c("GG","GA or AA")
> rownames(Y) <- c("Acne","Control")
> Y
```

```
      GG GA or AA
Acne   66      47
Control 99      15
```

```
> results <- chisq.test(Y)
> print(results)
```

Pearson's Chi-squared test with Yates' continuity correction

```
data: Y
X-squared = 21.7021, df = 1, p-value = 3.184e-06
```

```
> results <- chisq.test(Y,correct=FALSE)
> print(results)
```

Pearson's Chi-squared test

```
data: Y
X-squared = 23.1122, df = 1, p-value = 1.528e-06
```

Genetic association studies

Genotype based tests - Recessive test

- Columns in the original data table are combined to reflect the dominant model.

	aa or aA	AA	Total
Cases	$r_0 + r_1$	r_2	r
Controls	$s_0 + s_1$	s_2	s
Total	$n_0 + n_1$	n_2	n

- Hypothesis:
 - $H_0 : P(Y = 1 | AA) = P(Y = 1 | (aA + aa))$, so that the probability of disease does not depend on being homozygote AA
 - H_1 : disease depends on being homozygote AA
- Test statistic:

$$X^2 = \sum_{\text{genotypes}} \frac{(\text{observed} - \text{expected})^2}{\text{expected}}$$

- Under H_0 , the test statistic X^2 follows a χ_1^2 distribution (df=1)

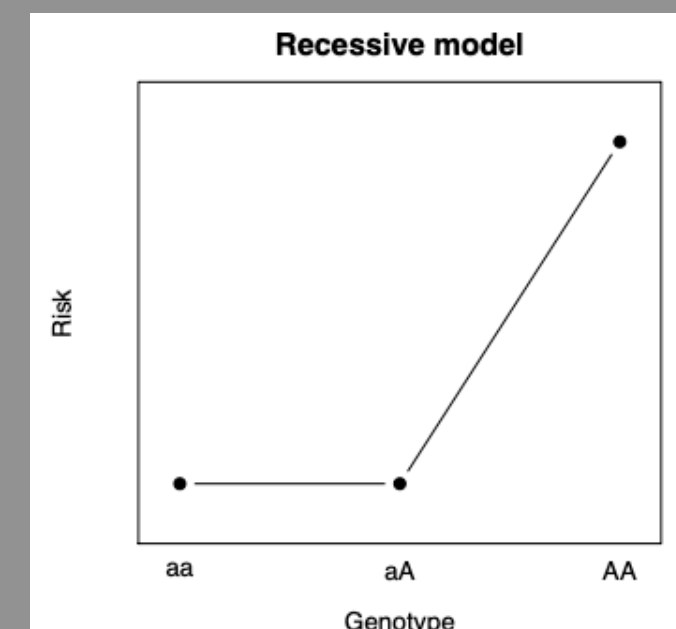
The trait (Y_i) (e.g. disease) we wish to understand is binary (dichotomous) so that:

- $Y_i = 1$ individual i has the trait
- $Y_i = 0$, individual i does not have the trait.

The marker is a bi-allelic polymorphism (e.g. AA, Aa and aa)

The original data table:

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n



Genetic association studies

Genotype based tests - Recessive test

Example - R code

- TNF genotype (G/A polymorphism) is on a study on acne patients and controls

	GG	GA	AA	Total
Cases	66	43	4	113
Controls	99	15	0	114
Total	165	58	4	227

```
> Y <- cbind(X[,1]+X[,2],X[,3])
> colnames(Y) <- c("GG", "GA or AA")
> rownames(Y) <- c("Acne", "Control")
> Y
      GG or GA  AA
Acne      109   4
Control   114   0
> results <- chisq.test(Y)
Warning message:
In chisq.test(Y) : Chi-squared approximation may be incorrect

> print(results)

Pearson's Chi-squared test with Yates' continuity correction
data: Y
X-squared = 2.3174, df = 1, p-value = 0.1279
```

```
> fisher.test(Y)
Fisher's Exact Test for Count Data
data: X
p-value = 0.05977
alternative hypothesis: true odds
ratio is not equal to 1
95 percent confidence interval:
 0.000000 1.485382
sample estimates:
odds ratio
```

We FAILED to reject the null hypothesis that the probability of disease does not depend on being homozygote AA

Genetic association study

Genotype based tests

TNF genotype (G/A polymorphism) is on a study on acne patients and controls

	GG	GA	AA	Total
Cases	66	43	4	113
Controls	99	15	0	114
Total	165	58	4	227

Significant association between G/A polymorphism and acne risk under dominant model

Allele model test (G vs A):

$$H_0 : P_{cases}(A) = P_{control}(A)$$

$$H_0 : P_{cases}(A) \neq P_{control}(A)$$

p-value = 1.072e-06 (Fisher Exact Test)
odds ratio (G/A) = 0.2423696

We reject the null hypothesis that the probability of A is the same in both cases and controls groups
Disease is 0.2 more frequent with allele A

Codominant model test (GG vs GA vs AA):

$$H_0 : P(Y = 1 | AA) = P(Y = 1 | aA) = P(Y = 1 | aa)$$

$$H_1 : \text{at least one pair different}$$

p-value = 1.97e-06 (Fisher Exact Test)

We reject the null hypothesis that the probability of disease with all the genotypes is the same

Recessive model test (GG + GA vs AA):

$$H_0 : \text{disease does not depend on being homozygote AA}$$

$$H_0 : P(Y = 1 | AA) = P(Y = 1 | (GG + GA))$$

$$H_1 : \text{disease depends on being homozygote AA}$$

p-value = 0.05977 (Fisher Exact Test)

We FAILED to reject the null hypothesis that the probability of disease does not depend on being homozygote AA

Dominant model test (GG vs GA + AA):

$$H_0 : \text{disease does not depend on presence of A}$$

$$H_0 : P(Y = 1 | GG) = P(Y = 1 | (GA + AA))$$

$$H_1 : \text{disease depends on the presence of A}$$

p-value = 1.528e-06 (χ^2 without cc)

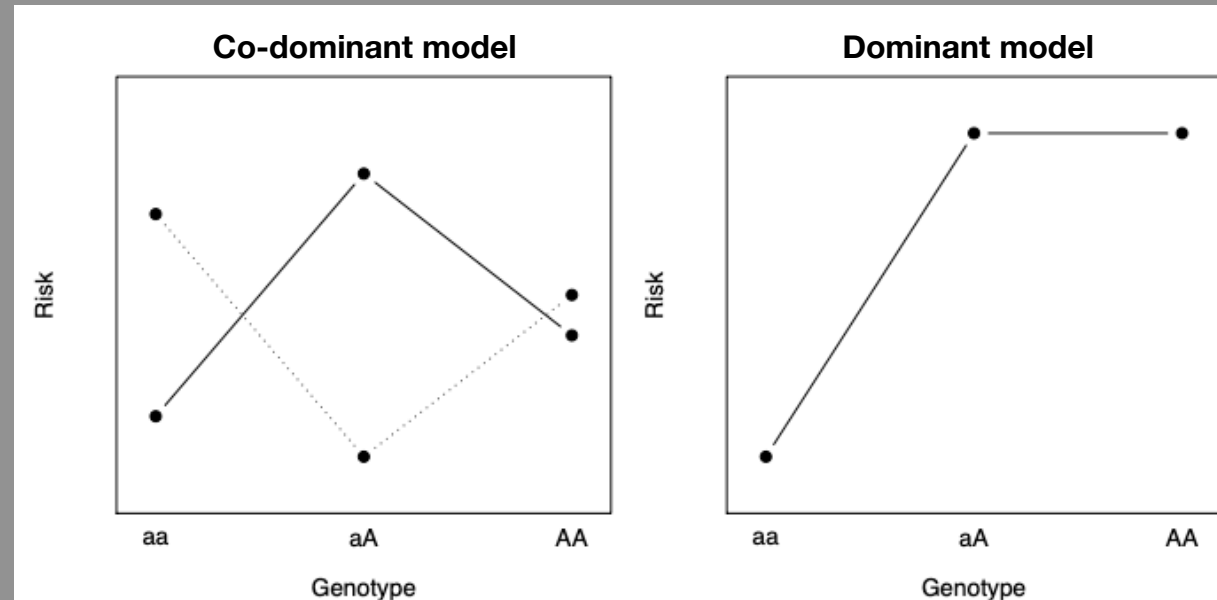
We reject the null hypothesis that the probability of disease does not depend on presence of A
Probability of disease is the same in GG vs GA/AA

Genetic association studies

Genotype based tests

Some codominant inheritance traits

- Alleles A and B in blood type
- HLA protein (cell surface antigen)
- ...

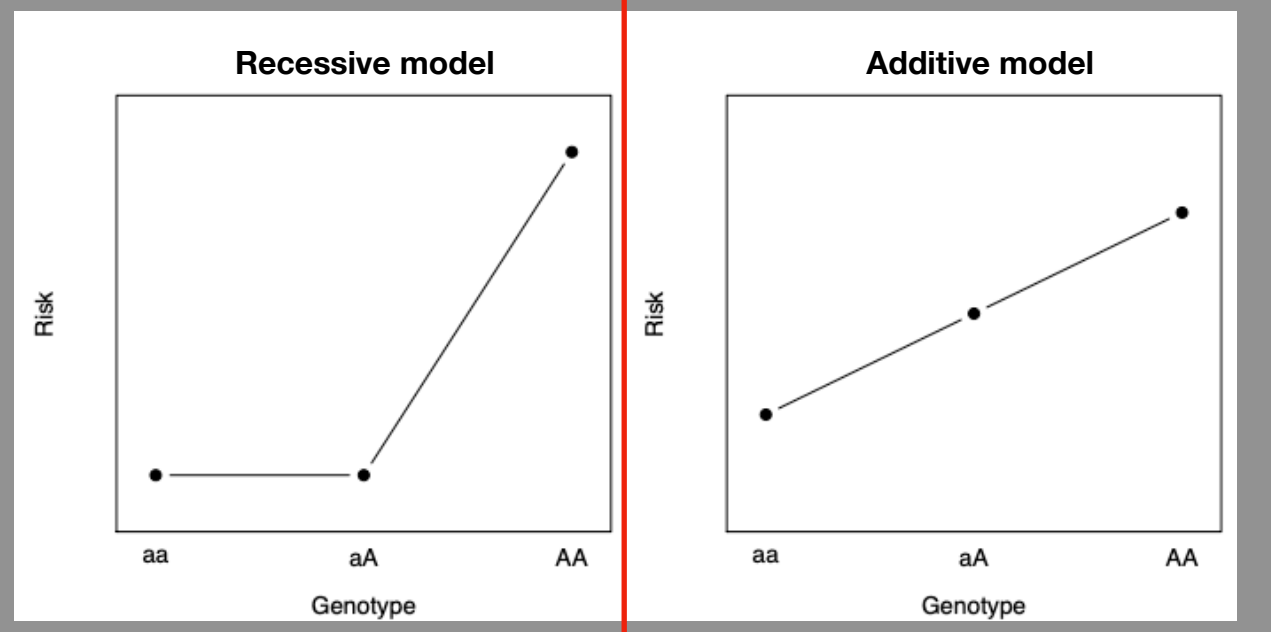


Common autosomal dominant inheritance disorders

- Huntington's disease
- Marfan syndrome (cardiovascular)
- Acondroplasia
- ...about 200 disorders
-

Common autosomal recessive inheritance disorders

- Spinal muscular atrophy
- Cystic fibrosis
-about 450 disorders
-



Genetic association studies

Genotype based tests - Additive test

- Basic idea: disease risk increases as a function of the number of alleles (0,1 or 2).
- There are two tests for the additive genetic model
 - The alleles test
 - Cochran-Armitage trend test
 - Used in categorical data analysis when the aim is to assess for the presence of an association between a variable with two categories and an ordinal variable with k categories.
 - Assumes ordering effect on the k categories of the second variable (i.e. treatment dosage or genotype)
 - Does not assume HWE hold in H_0
- In both cases, given that $P_{cases}(A)$ denotes the frequency of A alleles among cases and $P_{control}(A)$ denotes the frequency of A alleles among controls in the population, the hypothesis:

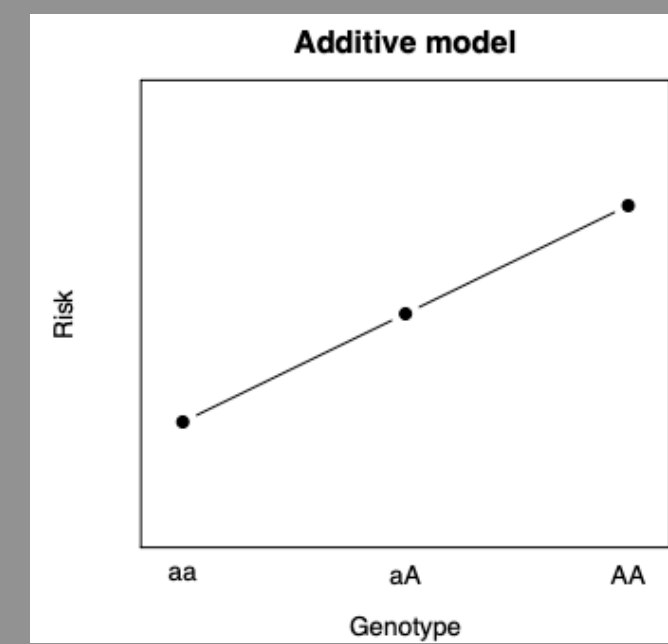
$$H_0 : P_{cases}(A) = P_{control}(A)$$

$$H_0 : P_{cases}(A) \neq P_{control}(A)$$

The marker is a bi-allelic polymorphism (e.g. AA, Aa and aa)

The original data table:

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n



Genetic association studies

Genotype based tests - Additive test

- Two options to resolve the additive model...

- **Option 1:** Assign weights by hand

- Test statistic $T = \sum_{i=1}^k t_i(r_i \cdot s - s_i \cdot r)$

- Variance $Var(T) = \frac{r \cdot s}{n} \left(\sum_{i=1}^k t_i r_i (n - n_i) - 2 \sum_{i=1}^{k-1} \sum_{j=i+1}^k t_i t_j n_i n_j \right)$

- If we're testing for linear trend on a bi-allelic marker $k = 3$ and $t = (0,1,2)$

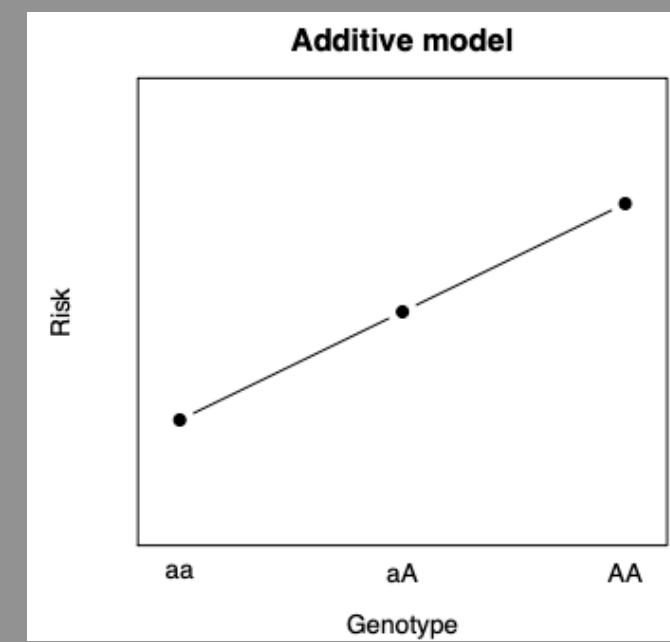
- Under H_0 , the test statistic T follows a χ_1^2 distribution (df=1), which corresponds to a standard normal distribution (zero mean and variance of 1) so that we can use the Z- score:

- The Z- score $Z = \frac{T}{\sqrt{var(T)}} \sim N(0,1)$

The marker is a bi-allelic polymorphism (e.g. AA, Aa and aa)

The original data table:

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n



Genetic association studies

Genotype based tests - Armitage trend

Example - Option 1

- TNF genotype (G/A polymorphism) is on a study on acne patients and controls
- As we're testing for linear trend on a bi-allelic marker $k = 3$ then $t = (0,1,2)$
- Test statistic and its variance:

$$T = \sum_{i=1}^k t_i(r_i \cdot s - s_i \cdot r) = 0 \cdot (66 \cdot 114 - 99 \cdot 113) + 1 \cdot (43 \cdot 114 - 15 \cdot 113) + 2 \cdot (4 \cdot 114 - 0 \cdot 113) = 4119$$

$$Var(T) = \frac{r \cdot s}{n} \left(\sum_{i=1}^k t_i r_i (n - n_i) - 2 \sum_{i=1}^{k-1} \sum_{j=i+1}^k t_i t_j n_i n_j \right) = \dots = 706081.3$$

$$\sum_{i=1}^{k-1} \sum_{j=i+1}^k t_i t_j n_i n_j = \sum_{j=i+1, i=1}^k t_i t_j n_i n_j + \sum_{j=i+1, i=2}^k t_i t_j n_i n_j = \sum_{j=2, i=1}^k t_i t_j n_i n_j + \sum_{j=3, i=2}^k t_i t_j n_i n_j = \sum_{j=2, i=1}^k t_i t_j n_i n_j + t_2 t_3 n_2 n_3 = t_1 t_2 n_1 n_2 + t_1 t_3 n_1 n_3 + t_2 t_3 n_2 n_3 = t_2 t_3 n_2 n_3$$

- The z-score $Z \sim N(0,1)$

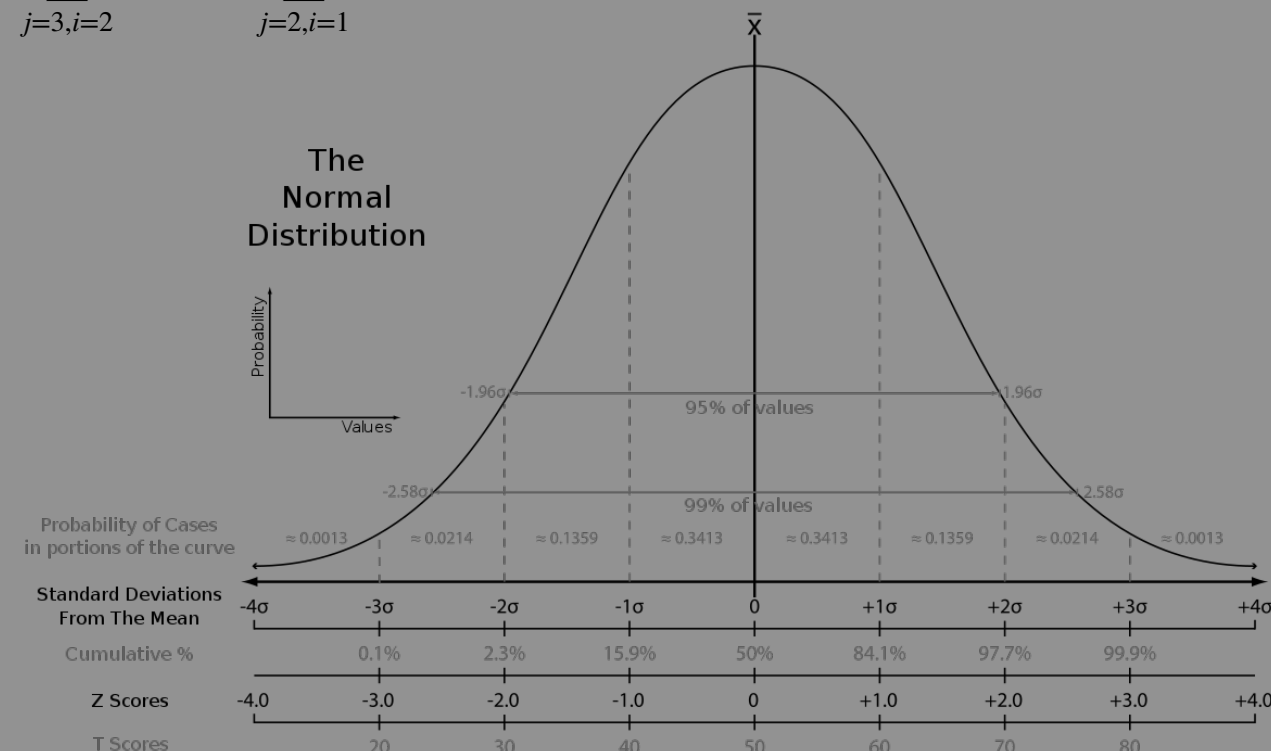
$$Z = \frac{T}{\sqrt{var(T)}} = \frac{4119}{\sqrt{706081.3}} = 4.9019$$

- P-value $p - value = P(Z \geq 4.9019) = 9.49e - 07$

We reject the null hypothesis that the probability of disease is the same with different number of A

	aa	aA	AA	Total
Cases	r_1	r_2	r_3	r
Controls	s_1	s_2	s_3	s
Total	n_1	n_2	n_3	n

	GG	GA	AA	Total
Cases	66	43	4	113
Controls	99	15	0	114
Total	165	58	4	227



Genetic association studies

Genotype based tests - Additive test

- **Option 2:** Consider a general linear regression model:

$$p(Y = 1 | X) = \beta_0 + \beta_1 X + \epsilon$$

- Y is the disease status ($Y_i = 1$ individual i has the trait, $Y_i = 0$, individual i does not have the trait)
- X is the number of A alleles (0=BB, 1=AB, 2=AA).
- Hypothesis:
 - $H_0 : \beta_1 = 0$
 - $H_1 : \beta_1 \neq 0$
- Test statistic:

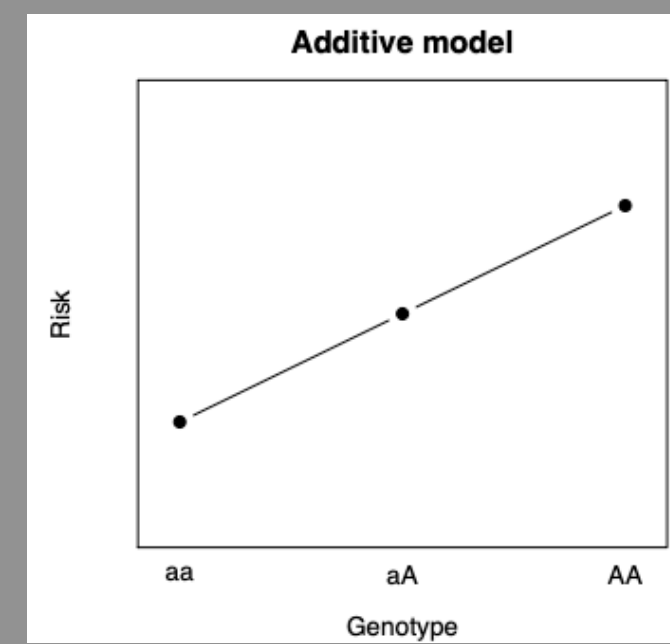
$$A = \frac{\hat{\beta}_1^2}{V(\hat{\beta}_1)} = n \cdot r_{xy}^2$$

- Under H_0 , the test statistic A follows a χ_1^2 distribution (df=1) from where we can obtain the p-value.

The marker is a bi-allelic polymorphism (e.g. AA, Aa and aa)

The original data table:

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n



Genetic association studies

Genotype based tests - Armitage trend

Example - Option 2: linear model

- TNF genotype (G/A polymorphism) is on a study on acne patients and controls
- Test statistic: $A = n \cdot r_{xy}^2 = 227 \cdot (0.3257)^2 = 24.02$
- P-value $p\text{-value} = P(\chi^2 \geq 24.02) = 9.49e - 07$

	aa	aA	AA	Total
Cases	r_0	r_1	r_2	r
Controls	s_0	s_1	s_2	s
Total	n_0	n_1	n_2	n

	GG	GA	AA	Total
Cases	66	43	4	113
Controls	99	15	0	114
Total	165	58	4	227

```
Cases      <- c(66,43,4)
Controls   <- c(99,15,0)
```

```
X <- rbind(Cases,Controls)
rownames(X) <- c("Cases","Controls")
colnames(X) <- c("GG","GA","AA")
n <- sum(X)
```

```
cas <- rep(c(0,1,2),Cases)
con <- rep(c(0,1,2),Controls)
```

```
y <- c(rep(1,sum(Cases)), rep(0,sum(Controls)))
```

```
x <- c(cas,con)
```

```
r <- cor(x,y)
```

```
A <- n*(r^2)
```

```
pvalue <- pchisq(A,df=1,lower.tail=FALSE)
```

SHORTCUT: The coefficient of determination is equivalent to the correlation coefficient between x, y when a single intercept is included and the sample standard deviations (SD) are the same.

$$\hat{\beta} = \text{corr}(x, y) \cdot \frac{SD(y)}{SD(x)}$$

Content

Genetic Association Studies

Until now....

The marker is a bi-allelic polymorphism (e.g. AA, Aa and aa)

The phenotype is binary:

- $Y_i = 1$ individual i has the trait
- $Y_i = 0$, individual i does not have the trait.

1. Introduction
2. Allele based tests
3. Genotype based tests
4. Quantitative traits
5. Multiple polymorphisms
6. Computer exercise

Genetic association studies

Quantitative responses

The trait or disease of interest can be quantitative (e.g. height, lipid levels, blood pressure, etc...).

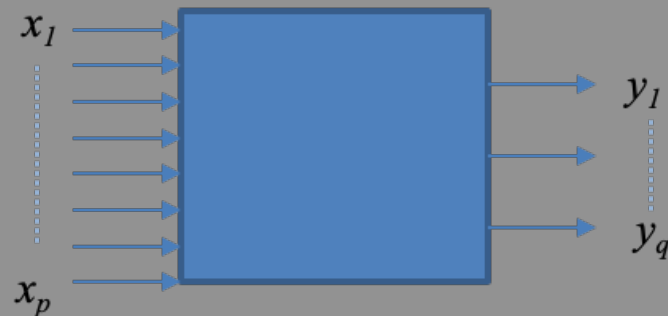
How to deal with quantitative outcomes?

- Multiple regression
- Mixed effects models (to account for correlation between individuals)

Genetic association studies

Genotype based tests - Regression models

- **Regression analysis:** estimate relationship between dependent variable y_1, \dots, y_q (outcome, output, response or label) and one or more independent variables x_1, \dots, x_p (predictors, features, covariates or explanatory variables)...where q can be $q = 1$, depending on the problem.



- **Example:**
 - Objective: prediction of the country median household income given a set of socio-economic independent variables (or predictors)
 - Variables: Geographic border to metro county, educational attainment measures, interstate highway density, population density, labor force participation rate, ...
 - Outcome: Country median household income
- **Linear regression:** each response y_1, \dots, y_q is modeled as a linear combination of the input x_1, \dots, x_p plus a random fluctuation ϵ so that:

$$y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i = \mathbf{x}_i^T \boldsymbol{\beta} + \epsilon_i$$

Genetic association studies

Genotype based tests - Regression models

- **Linear regression:** each response y_1, \dots, y_q is modeled as a linear combination of the input x_1, \dots, x_p plus a random fluctuation ϵ so that:

$$y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i = \mathbf{x}_i^T \boldsymbol{\beta} + \epsilon_i$$

Or in matrix notation, considering n individuals:

$$\begin{matrix} \vec{y}_k & & \vec{x}_1 & & \vec{x}_p & & \vec{\epsilon}_k \\ \left[\begin{array}{ccc} y_{11} & y_{1k} & y_{1q} \\ \vdots & \vdots & \vdots \\ y_{n1} & y_{nk} & y_{nq} \end{array} \right] & = & \left[\begin{array}{ccc} x_{11} & x_{12} & \dots & x_{1p} \\ \vdots & \vdots & \dots & \vdots \\ x_{n1} & x_{n2} & \dots & x_{np} \end{array} \right] & \left[\begin{array}{ccc} \beta_{11} & \beta_{1k} & \beta_{1q} \\ \beta_{21} & \beta_{2k} & \beta_{2q} \\ \vdots & \vdots & \vdots \\ \beta_{p1} & \beta_{pk} & \beta_{pq} \end{array} \right] & + & \left[\begin{array}{ccc} \epsilon_{11} & \epsilon_{1k} & \epsilon_{1q} \\ \vdots & \vdots & \vdots \\ \epsilon_{n1} & \epsilon_{nk} & \epsilon_{nq} \end{array} \right] \end{matrix}$$

$$\vec{y}_k = X \mathbf{b}_k + \mathbf{e}_k \quad k = 1, \dots, q$$

$$\underset{(n,q)}{Y} = \underset{(n,p)}{X} \underset{(p,q)}{B} + \underset{(n,q)}{E}$$

Y matrix of q response variables (centered)

X matrix of p explanatory variables (centered)

B is the matrix of $p \times q$ b_{jk} parameters

E random fluctuation matrix

Genetic association studies

Genotype based tests - Regression models

- **Linear mixed models:** extension of linear models that allow for both fixed and random effects. Used to analyze hierarchical data, longitudinal or correlated. Each response y_1, \dots, y_q is modeled as a linear combination of the fixed effects x_1, \dots, x_p , plus a linear combination of random effects z_1, \dots, z_p and a random fluctuation ϵ so that:

$$y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + u_0 + u_1 z_{i1} + \dots + u_p z_{ip} + \epsilon_i = \mathbf{x}_i^T \boldsymbol{\beta} + \mathbf{z}_i^T \mathbf{u} + \epsilon_i$$

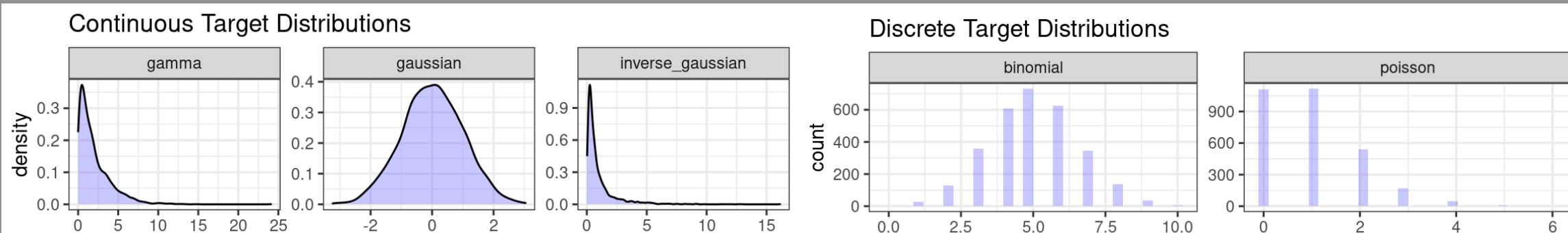
- **Logistic regression:** (subset of GLM) linear regression adjusted for binary outcomes by using the logistic or logit function as a link function:

$$\text{logit}(y_i) = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i = \mathbf{x}_i^T \boldsymbol{\beta} + \epsilon_i$$

- **Generalized linear model:** linear regression adjusted for outcome variables that are not continuous, allowing the response variable y_i to be related to the input via several link functions g - all from the exponential family. Choose the target distribution that is similar to the actual distribution of Y .

$$g(E(y)) = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i = \mathbf{x}_i^T \boldsymbol{\beta} + \epsilon_i$$

- Binomial (link = “logit”): outcome variable has 2 possible outcomes
- Gamma (link = “inverse”): outcome is continuous, non-negative and positive-skewed data, such as insurance claims and survival data.
- Poisson (link = ‘log’): outcome variable is a count/discrete variable
- ...



Content

Genetic Association Studies

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Genetic association studies

Multiple polymorphisms

RECALL:

Multiple polymorphisms studies: involve genome wide association studies (or multiple SNPs) conducted without prior hypothesis.

How to deal with multiple SNPs?

- Multiple regression models (for a small amount of SNPs)
- Regression with haplotypes
- Test all variants: Genome wide association studies (GWAS)

Genetic association studies

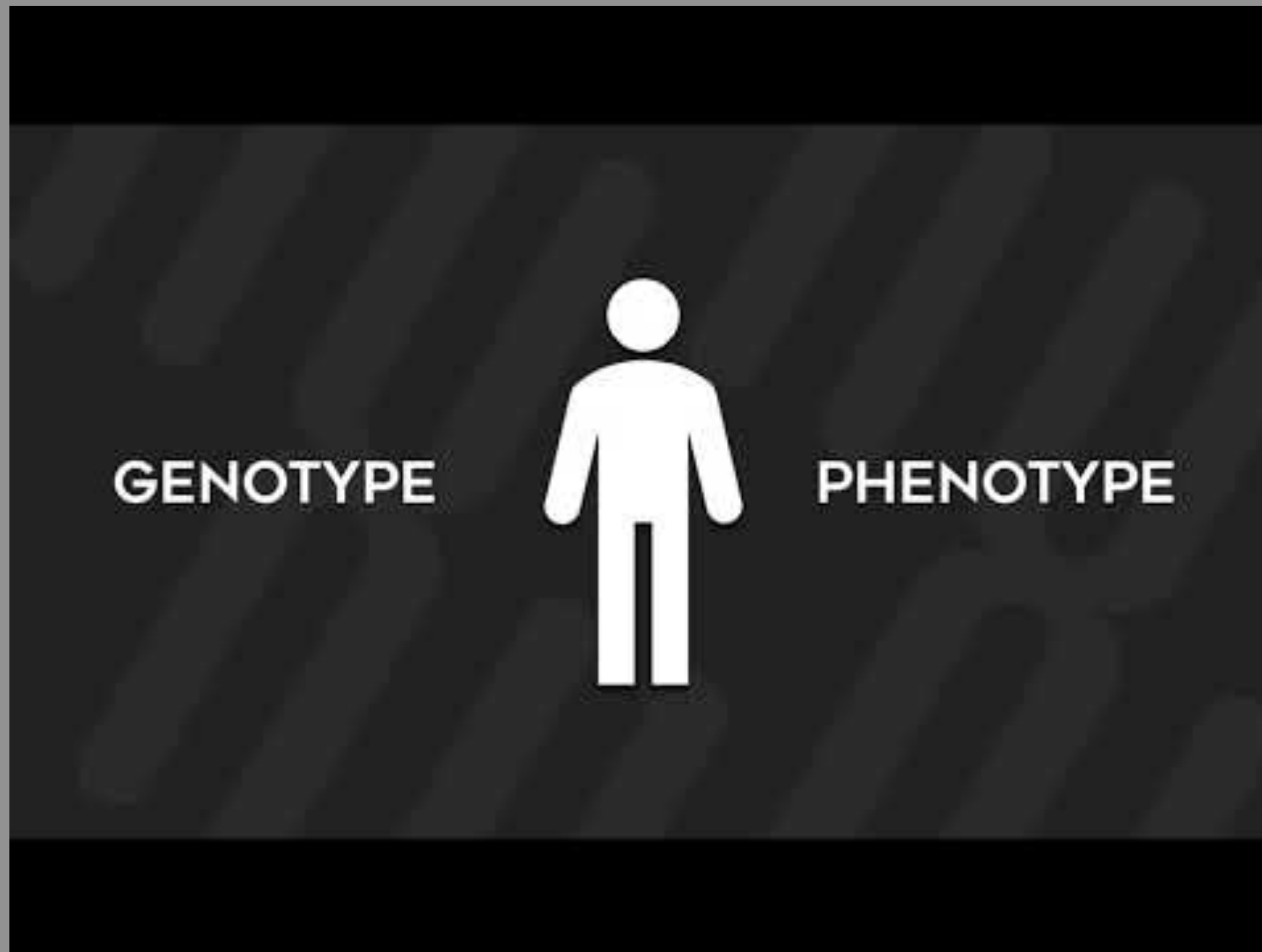
Multiple polymorphisms - multiple regression

- Helps understanding genetic mechanisms that contribute to a trait, providing an effect size for each SNP.
- Only works for small amount of SNPs, as we need more participants (data points) than SNPs (parameters).
- Marginal SNP effects can be estimated one SNP at a time but....multiple regression allows for multiple estimations.
- Additional covariate terms may be included
- Mult-SNP joint regression model of p SNPs:

$$y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i = \mathbf{x}_i^T \boldsymbol{\beta} + \epsilon_i$$

Genetic association studies

Multiple polymorphisms - Genome wide association studies (GWAS)

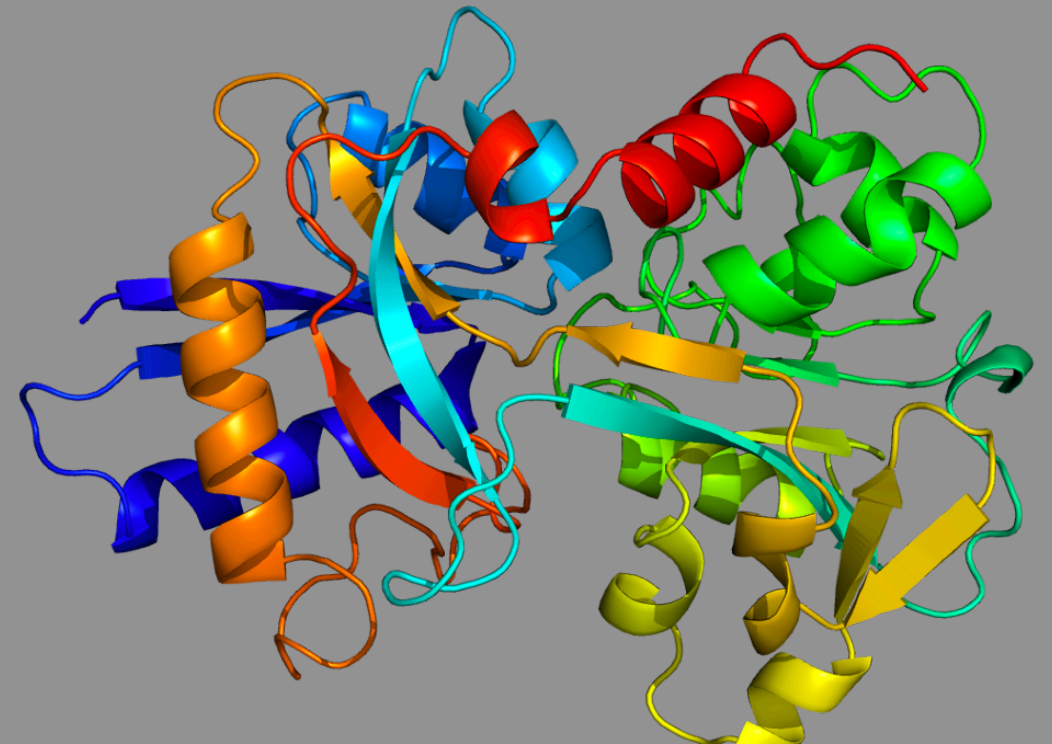


Genetic association studies

Multiple polymorphisms - Example study

GWAS for transferrin

- 2,362 individuals for which (adjusted) transferrin serum levels are available.
- 281,313 SNPs from all chromosomes.
- Filters: missing rate < 0.01 ;
- $MAF > 0.05$;
- $HWE\ p\text{-value} > 0.001$.
- We use an additive model for each SNP and fit this model with PLINK.



Transferrins are a family of proteins that mediate iron transport.

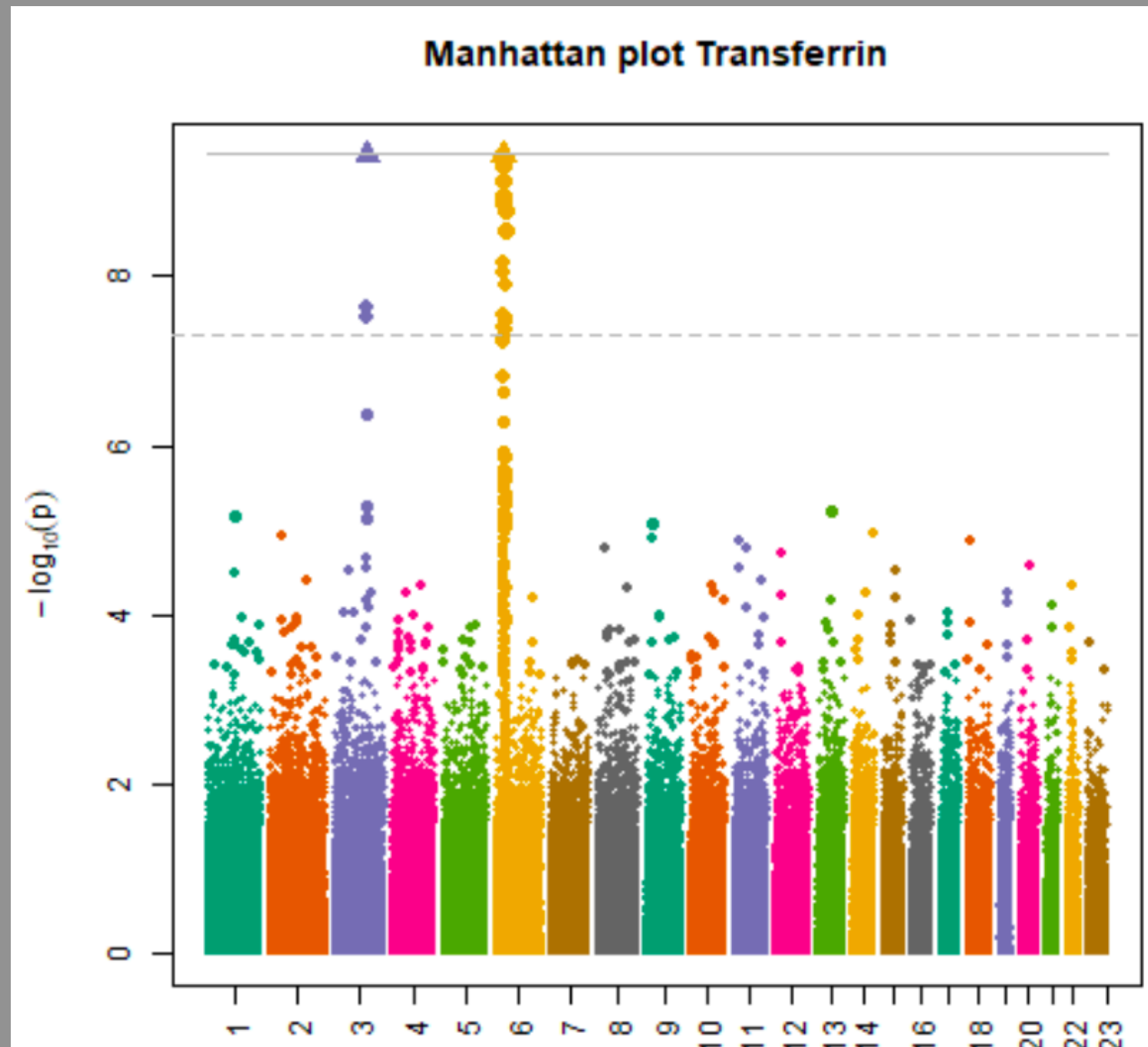
Gene coding for transferrin in humans is located in the chromosome band 3q21.

- The first point of reference is a number (or letter) which denotes the chromosome (e.g. 3q21 refers to chromosome 3)
- The second point of reference is a letter (p or q) to denote which arm the locus is positioned on (e.g. 3q21 is on the q arm)
- The third point of reference is a number corresponding to the G band location (e.g. 3q21 is at the longitudinal position 21)

Genetic association studies

Multiple polymorphisms - Example study

GWAS for transferrin



Genetic association studies

Multiple polymorphisms - Example study

GWAS for transferrin - top 25

	SNP	CHR	BP	B	SE	R2	T	P	-log10(P)
1	rs3811647	3	134966719	0.3832	0.02889	0.06936	13.260	8.965e-39	38.047450
2	rs6794945	3	135001153	0.3652	0.02940	0.06136	12.420	2.324e-34	33.633764
3	rs1800562	6	26201120	-0.5884	0.04988	0.05572	-11.800	2.968e-31	30.527536
4	rs13214703	6	28049366	-0.4378	0.04886	0.03292	-8.961	6.390e-19	18.194499
5	rs1358024	3	134966878	0.3290	0.03745	0.03168	8.785	2.941e-18	17.531505
6	rs2274089	6	25596562	-0.3791	0.04551	0.02856	-8.330	1.352e-16	15.869023
7	rs4525863	3	134918826	0.2399	0.03017	0.02609	7.951	2.845e-15	14.545918
8	rs1867503	3	134893338	0.2039	0.02864	0.02103	7.120	1.428e-12	11.845272
9	rs1867504	3	134893351	0.2039	0.02864	0.02103	7.120	1.428e-12	11.845272
10	rs9853615	3	135002671	-0.2083	0.02929	0.02098	-7.111	1.523e-12	11.817300
11	rs12216125	6	26105437	-0.1974	0.02891	0.01936	-6.826	1.107e-11	10.955852
12	rs9379818	6	26131185	-0.1931	0.02838	0.01925	-6.805	1.276e-11	10.894149
13	rs13194984	6	26608542	-0.2719	0.04060	0.01865	-6.698	2.638e-11	10.578725
14	rs932316	6	25749179	-0.2371	0.03557	0.01849	-6.664	3.309e-11	10.480303
15	rs17270561	6	25928418	-0.2183	0.03292	0.01829	-6.631	4.108e-11	10.386370
16	rs2013063	6	26102077	-0.1798	0.02823	0.01690	-6.369	2.285e-10	9.641114
17	rs1543680	6	26211156	-0.2036	0.03259	0.01627	-6.247	4.944e-10	9.305922
18	rs10484432	6	26116855	-0.2013	0.03256	0.01595	-6.183	7.390e-10	9.131356
19	rs2009610	6	26075047	-0.1959	0.03205	0.01558	-6.111	1.158e-09	8.936291
20	rs707889	6	26203910	-0.1969	0.03238	0.01543	-6.082	1.383e-09	8.859178
21	rs1029328	6	28555894	-0.2509	0.04150	0.01526	-6.047	1.709e-09	8.767258
22	rs11757000	6	28592848	-0.2307	0.03868	0.01486	-5.966	2.806e-09	8.551912
23	rs169219	6	26065371	0.1669	0.02870	0.01413	5.816	6.861e-09	8.163613
24	rs7748771	6	25463078	-0.2678	0.04645	0.01389	-5.765	9.249e-09	8.033905
25	rs3130253	6	29741991	-0.2769	0.04845	0.01365	-5.715	1.238e-08	7.907279

Statistical concerns:

- Effect of filters applied?
- Multiple testing problem?
- X-chromosome adequately dealt with?
- Family structure accounted for?
- Adjustment for covariates?
- Power?

Content

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Genetic association studies

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