

# Towards Optimal Variance Reduction in Online Experiments

Ying Jin

Department of Statistics, Stanford University



Joint work with Shan Ba at LinkedIn Applied Research  
*Conference on Digital Experimentation, November 5, 2021*

# Randomized experiments in tech companies

- ▶ A/B testing, randomized experiments, online controlled experiments...
  - ▶ Units are randomly assigned to treated / control groups
  - ▶ Measure outcomes after a period
  - ▶ Evaluate and compare the outcomes in two groups
  - ▶ If the effect is significantly positive, then adopt the new feature

# Randomized experiments in tech companies

- ▶ A/B testing, randomized experiments, online controlled experiments...
  - ▶ Units are randomly assigned to treated / control groups
  - ▶ Measure outcomes after a period
  - ▶ Evaluate and compare the outcomes in two groups
  - ▶ If the effect is significantly positive, then adopt the new feature
- ▶ Powerful hypothesis testing is important
  - ▶ shorter experimental horizon, smaller sample, avoid potentially negative impacts

# Randomized experiments in tech companies

- ▶ A/B testing, randomized experiments, online controlled experiments...
  - ▶ Units are randomly assigned to treated / control groups
  - ▶ Measure outcomes after a period
  - ▶ Evaluate and compare the outcomes in two groups
  - ▶ If the effect is significantly positive, then adopt the new feature
- ▶ Powerful hypothesis testing is important
  - ▶ shorter experimental horizon, smaller sample, avoid potentially negative impacts
- ▶ Desire estimator of treatment effects with **smaller variance**
  - ▶ What treatment effect?
  - ▶ How to reduce variance?
  - ▶ What is the best effort?

# Overview of the work

- ▶ A rigorous statistical framework for variance reduction of count and ratio metrics
- ▶ Methodology of unbiased variance reduction with flexible machine learning tools and large numbers of covariates
- ▶ Optimality in the sense of semiparametric efficiency of all procedures
- ▶ Performance on (simulated and) real data

# Potential outcome framework

- ▶ i.i.d. units  $i = 1, \dots, n$  from  $\mathbb{P}$ .
- ▶ Potential outcomes  $Y_i(1), Y_i(0)$ .
- ▶ Treatment  $T_i \stackrel{\text{i.i.d.}}{\sim} \text{Bernoulli}(p)$ .
- ▶ SUTVA: observe  $Y_i(1)$  for  $T_i = 1$  and  $Y_i(0)$  for  $T_i = 0$ .
- ▶  $n_t = \sum_i T_i$  size of treated group,  $n_c = n - n_t$  size of control group.

# Randomization and metrics

- ▶ Case 1: Count metric  $\tau = \mathbb{E}[Y(1)] - \mathbb{E}[Y(0)]$
- ▶ Default estimator: difference-in-mean

$$\hat{\tau}_{\text{DIM}} = \frac{1}{n_t} \sum_{i \text{ treated}} Y_i - \frac{1}{n_c} \sum_{i \text{ control}} Y_i.$$

# Randomization and metrics

► Case 2: ratio metric

$$\frac{\sum_{i \text{ treated}} Y_i}{\sum_{i \text{ treated}} Z_i} = \frac{\frac{1}{n_t} \sum_{i \text{ treated}} Y_i}{\frac{1}{n_t} \sum_{i \text{ treated}} Z_i}, \quad \frac{\sum_{i \text{ control}} Y_i}{\sum_{i \text{ control}} Z_i} = \frac{\frac{1}{n_c} \sum_{i \text{ control}} Y_i}{\frac{1}{n_c} \sum_{i \text{ control}} Z_i},$$

where  $i$  stands for a cluster,  $Y_i$  is the aggregated outcome,  $Z_i$  is the size of cluster.



# Randomization and metrics

- ▶ Case 2: ratio metric

$$\frac{\sum_{i \text{ treated}} Y_i}{\sum_{i \text{ treated}} Z_i} = \frac{\frac{1}{n_t} \sum_{i \text{ treated}} Y_i}{\frac{1}{n_t} \sum_{i \text{ treated}} Z_i}, \quad \frac{\sum_{i \text{ control}} Y_i}{\sum_{i \text{ control}} Z_i} = \frac{\frac{1}{n_c} \sum_{i \text{ control}} Y_i}{\frac{1}{n_c} \sum_{i \text{ control}} Z_i},$$

where  $i$  stands for a cluster,  $Y_i$  is the aggregated outcome,  $Z_i$  is the size of cluster.

- ▶ **Stable denominator assumption:**  $Z_i = Z_i(1) = Z_i(0)$ , not influenced by the treatment.

# Randomization and metrics

- ▶ Case 2: ratio metric

$$\frac{\sum_{i \text{ treated}} Y_i}{\sum_{i \text{ treated}} Z_i} = \frac{\frac{1}{n_t} \sum_{i \text{ treated}} Y_i}{\frac{1}{n_t} \sum_{i \text{ treated}} Z_i}, \quad \frac{\sum_{i \text{ control}} Y_i}{\sum_{i \text{ control}} Z_i} = \frac{\frac{1}{n_c} \sum_{i \text{ control}} Y_i}{\frac{1}{n_c} \sum_{i \text{ control}} Z_i},$$

where  $i$  stands for a cluster,  $Y_i$  is the aggregated outcome,  $Z_i$  is the size of cluster.

- ▶ **Stable denominator assumption:**  $Z_i = Z_i(1) = Z_i(0)$ , not influenced by the treatment.
- ▶ If SDA holds, the population quantity is  $\delta' = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i]}$ .

# Randomization and metrics

- ▶ Case 2: ratio metric

$$\frac{\sum_{i \text{ treated}} Y_i}{\sum_{i \text{ treated}} Z_i} = \frac{\frac{1}{n_t} \sum_{i \text{ treated}} Y_i}{\frac{1}{n_t} \sum_{i \text{ treated}} Z_i}, \quad \frac{\sum_{i \text{ control}} Y_i}{\sum_{i \text{ control}} Z_i} = \frac{\frac{1}{n_c} \sum_{i \text{ control}} Y_i}{\frac{1}{n_c} \sum_{i \text{ control}} Z_i},$$

where  $i$  stands for a cluster,  $Y_i$  is the aggregated outcome,  $Z_i$  is the size of cluster.

- ▶ **Stable denominator assumption:**  $Z_i = Z_i(1) = Z_i(0)$ , not influenced by the treatment.
- ▶ If SDA holds, the population quantity is  $\delta' = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i]}$ .
- ▶ If SDA does not hold, the population quantity is  $\delta = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i(1)]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i(0)]}$ .

# Methods in the literature

- ▶ For count metrics...
- ▶ Earlier: linear adjustment [Lin, 2013], CUPED [Deng et al., 2013]..
  - ▶ linearity is restrictive
  - ▶ cannot handle large numbers of covariates
- ▶ Recently: machine learning, large numbers of covariates [Guo et al., 2021]
  - ▶ optimality of procedures?

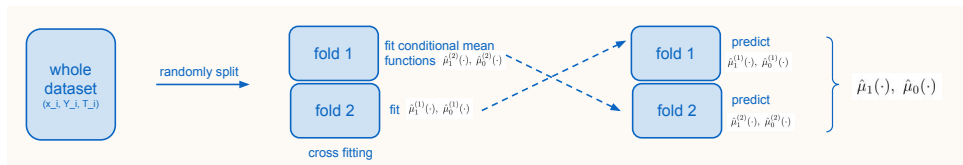
# Methods in the literature

- ▶ For count metrics...
- ▶ Earlier: linear adjustment [Lin, 2013], CUPED [Deng et al., 2013]..
  - ▶ linearity is restrictive
  - ▶ cannot handle large numbers of covariates
- ▶ Recently: machine learning, large numbers of covariates [Guo et al., 2021]
  - ▶ optimality of procedures?
- ▶ For ratio metrics...
  - ▶ less studied, lack of rigorous statistical framework and guarantee
  - ▶ CUPED extension cannot use general covariates
  - ▶ optimality of procedures?

# High-level idea: fit-then-debias

- ▶ Why is diff-in-mean estimator not efficient?
  - ▶ Decreased sample size (half of  $n$  units are treated / control)
  - ▶ Unpaired comparison (variance is the sum of treated and control)
  - ▶ Ideal estimator:  $\frac{1}{n} \sum_{i=1}^n [Y_i(1) - Y_i(0)]$  (for count metrics)
- ▶ General idea:
  - ▶ with covariates  $X_i$ , use ML estimators to predict the missing outcome and plug in to perform pairwise comparison  $\Rightarrow \frac{1}{n} \sum_{i=1}^n [\hat{\mu}_1(X_i) - \hat{\mu}_0(X_i)]$ ?
  - ▶ **Bias** can be larger than variance!
- ▶ De-biasing techniques
  - ▶ use **cross-fitting** to fit the estimators
  - ▶ add a **de-biasing term** to correct for the bias of  $\hat{\mu}_w$

# Procedure for count metrics



- Step 1: sample splitting into  $\mathcal{D}^{(k)}$ ,  $k = 1, \dots, K$ .
- Step 2: use  $\mathcal{D} \setminus \mathcal{D}^{(k)}$  to estimate  $\hat{\mu}_1^{(k)}$  and  $\hat{\mu}_0^{(k)}$ .
- Step 3: Plug into  $\mathcal{D}^{(k)}$  to obtain  $\hat{\mu}_1(X_i) = \hat{\mu}_1^{(k)}(X_i)$ ,  $\hat{\mu}_0(X_i) = \hat{\mu}_0^{(k)}(X_i)$  for all  $i \in \mathcal{D}^{(k)}$ .
- Step 4: Estimator

$$\hat{\theta} = \frac{1}{n} \sum_{i=1}^n (\hat{\mu}_1(X_i) - \hat{\mu}_0(X_i)) + \frac{1}{n_t} \sum_{i \text{ treated}} (Y_i(1) - \hat{\mu}_1(X_i)) - \frac{1}{n_c} \sum_{i \text{ control}} (Y_i(0) - \hat{\mu}_0(X_i)),$$

# Procedure for count metrics

- ▶ Estimator

$$\hat{\theta} = \frac{1}{n} \sum_{i=1}^n (\hat{\mu}_1(X_i) - \hat{\mu}_0(X_i)) + \frac{1}{n_t} \sum_{i \text{ treated}} (Y_i(1) - \hat{\mu}_1(X_i)) - \frac{1}{n_c} \sum_{i \text{ control}} (Y_i(0) - \hat{\mu}_0(X_i)),$$

- ▶ Valid inference when  $\|\hat{\mu}_w^{(k)} - \mu_w^*\|_2 \xrightarrow{P} 0$  for deterministic functions  $\mu_w^*$ ,  $w \in \{0, 1\}$ .
- ▶ Semiparametrically efficient when  $\mu_w^*(x) = \mathbb{E}[Y(w) | X = x]$ ,  $w \in \{0, 1\}$ .



## Count metrics: implications on optimality

- ▶ When estimators converge in  $L_2$  to true conditional mean functions,

$$\text{Var}(\hat{\theta}) \approx \underbrace{\frac{1}{n} \text{Var}(\mu_1(X_i) - \mu_0(X_i))}_{\text{(i) predictable part}} + \underbrace{\frac{1}{n_t} \text{Var}(Y_i(1) - \mu_1(X_i)) + \frac{1}{n_c} \text{Var}(Y_i(0) - \mu_0(X_i))}_{\text{(ii) irreducible variance}}$$

- ▶ (i) is the best efforts in predicting  $Y(1) - Y(0)$  given  $X$ .
- ▶ (ii) is the intrinsic uncertainty that cannot be eliminated.

# Count metrics: implications on optimality

- ▶ When estimators converge in  $L_2$  to true conditional mean functions,

$$\text{Var}(\hat{\theta}) \approx \underbrace{\frac{1}{n} \text{Var}(\mu_1(X_i) - \mu_0(X_i))}_{\text{(i) predictable part}} + \underbrace{\frac{1}{n_t} \text{Var}(Y_i(1) - \mu_1(X_i)) + \frac{1}{n_c} \text{Var}(Y_i(0) - \mu_0(X_i))}_{\text{(ii) irreducible variance}}$$

- ▶ (i) is the best efforts in predicting  $Y(1) - Y(0)$  given  $X$ .
  - ▶ (ii) is the intrinsic uncertainty that cannot be eliminated.
- ▶ Take-away message on optimality:
  - ▶ Should target for **conditional mean** functions
    - ⇒ This method is better than CUPED when there is **nonlinearity**

# Count metrics: implications on optimality

- ▶ When estimators converge in  $L_2$  to true conditional mean functions,

$$\text{Var}(\hat{\theta}) \approx \underbrace{\frac{1}{n} \text{Var}(\mu_1(X_i) - \mu_0(X_i))}_{\text{(i) predictable part}} + \underbrace{\frac{1}{n_t} \text{Var}(Y_i(1) - \mu_1(X_i)) + \frac{1}{n_c} \text{Var}(Y_i(0) - \mu_0(X_i))}_{\text{(ii) irreducible variance}}$$

- ▶ (i) is the best efforts in predicting  $Y(1) - Y(0)$  given  $X$ .
- ▶ (ii) is the intrinsic uncertainty that cannot be eliminated.
- ▶ Take-away message on optimality:
  - ▶ Should target for **conditional mean** functions
    - ⇒ This method is better than CUPED when there is **nonlinearity**
  - ▶ Estimate for **two groups**  $\mathbb{E}[Y(1) | X = x]$  and  $\mathbb{E}[Y(0) | X = x]$  separately
    - ⇒ This method is better than one single estimator when there is **treatment heterogeneity**

## Ratio metrics: optimal procedures

- ▶ We develop two procedures for  $\delta$  and  $\delta'$  under different conditions of denominators.
- ▶ For the target  $\delta = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i(1)]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i(0)]}$  without SDA
  - ▶ fit-then-debias for  $\mathbb{E}[Y(w) | X = x]$ ,  $\mathbb{E}[Z(w) | X = x]$  respectively,  $w \in \{0, 1\}$ .
- ▶ For the target  $\delta' = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i]}$  under SDA,
  - ▶ pool all  $Z_i$  to estimate  $\mathbb{E}[Z]$
  - ▶ fit-then-debias for  $\mathbb{E}[Y(w) | X = x, Z = z]$ ,  $w \in \{0, 1\}$ .

## Ratio metrics: optimal procedures

- ▶ We develop two procedures for  $\delta$  and  $\delta'$  under different conditions of denominators.
- ▶ For the target  $\delta = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i(1)]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i(0)]}$  without SDA
  - ▶ fit-then-debias for  $\mathbb{E}[Y(w) | X = x]$ ,  $\mathbb{E}[Z(w) | X = x]$  respectively,  $w \in \{0, 1\}$ .
- ▶ For the target  $\delta' = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i]}$  under SDA,
  - ▶ pool all  $Z_i$  to estimate  $\mathbb{E}[Z]$
  - ▶ fit-then-debias for  $\mathbb{E}[Y(w) | X = x, Z = z]$ ,  $w \in \{0, 1\}$ .
- ▶ **Valid inference** when estimators converge to deterministic functions ( $L_2$  distance in probability)
- ▶ **Optimality** when they converge to true conditional mean functions

# Ratio metrics: optimal procedures

- Ratio metric: target quantity  $\delta = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i(1)]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i(0)]}$

$$\hat{\delta} = \frac{\sum_{i=1}^n A_i}{\sum_{i=1}^n B_i} - \frac{\sum_{i=1}^n C_i}{\sum_{i=1}^n D_i},$$

where  $A_i = \hat{\mu}_1^Y(X_i) + \frac{T_i}{\hat{p}}(Y_i - \hat{\mu}_1^Y(X_i)), \quad B_i = \hat{\mu}_1^Z(X_i) + \frac{T_i}{\hat{p}}(Z_i - \hat{\mu}_1^Z(X_i)),$

$$C_i = \underbrace{\hat{\mu}_0^Y(X_i)}_{\text{regressor + plug-in}} + \underbrace{\frac{1-T_i}{1-\hat{p}}(Y_i - \hat{\mu}_0^Y(X_i))}_{\text{de-bias}}, \quad D_i = \hat{\mu}_0^Z(X_i) + \frac{1-T_i}{1-\hat{p}}(Z_i - \hat{\mu}_0^Z(X_i)).$$

- Ratio metric: target quantity  $\delta' = \frac{\mathbb{E}[Y_i(1)]}{\mathbb{E}[Z_i]} - \frac{\mathbb{E}[Y_i(0)]}{\mathbb{E}[Z_i]}$

$$\hat{\delta}' = \frac{\sum_{i=1}^n \Gamma_i}{\sum_{i=1}^n Z_i}, \quad \text{where } \Gamma_i = \underbrace{\hat{\mu}_1(X_i, Z_i) - \hat{\mu}_0(X_i, Z_i)}_{\text{regressor + plug-in}} + \underbrace{\frac{T_i}{\hat{p}}(Y_i - \hat{\mu}_1(X_i, Z_i)) - \frac{1-T_i}{1-\hat{p}}(Y_i - \hat{\mu}_0(X_i, Z_i))}_{\text{de-bias}},$$

# Real data performance

- ▶ Count metric: LinkedIn Feed experiment, revenue metric
  - ▶  $n = 400,000$  subsample of users
  - ▶ Our estimator with random forest from `scikit-learn` python library
  - ▶ Incorporate user covariates
  - ▶ Reduces **22.22%** of variance compared to diff-in-mean, while CUPED reduces 15.91%

## Real data performance

- ▶ Ratio metric: enterprise experiment in LinkedIn Learning, 'learning engagement' metric
  - ▶  $n = 10,299$  enterprise accounts
  - ▶ Our estimator with XGBoost python library
  - ▶ Incorporate enterprise covariates
  - ▶ For  $\delta$ , reduces 12.35% of variance compared to diff-in-mean, while CUPED reduces 1.76%
  - ▶ For  $\delta'$ , reduces 83.6% of variance compared to diff-in-mean, while CUPED reduces 76.62%



Thanks!

Our paper: <https://arxiv.org/abs/2110.13406>