$$\frac{d}{do} \left(e^{t} (1-e)^{h} \right) = t e^{t-1} (1-e)^{h-1} - e^{t} h (1-e)^{h-1}$$

$$= e^{t-1} (1-e)^{h-1} \left[t (1-e) - h e \right]$$

$$\begin{aligned} &\frac{d^{2}}{d\theta^{2}}\left(\Theta^{\frac{k-1}{2}}\cdot\left(1-\theta\right)^{\frac{k-1}{2}}\cdot\left[t\cdot\left(1-\theta\right)-\lambda\cdot\theta\right]\right) \\ &=\left[\left(t\cdot1\right)\cdot\underline{\theta}^{\frac{k-2}{2}}\cdot\left(1-\underline{\theta}\right)^{\frac{k-1}{2}}-\underline{\theta}^{\frac{k+1}{2}}\cdot\left(\lambda-1\right)\cdot\left(1-\underline{\theta}\right)^{\frac{k-2}{2}}\cdot\left[\frac{t\cdot\left(1-\underline{\theta}\right)}{t\cdot\left(1-\theta\right)}-\lambda\cdot\underline{\theta}\right]-\left[\underline{\theta}^{\frac{k+1}{2}}\cdot\left(1-\underline{\theta}\right)^{\frac{k-1}{2}}\cdot\left[t+\lambda\right]\right] \\ &=\frac{t\cdot\left(t-1\right)\cdot\underline{\theta}^{\frac{k-2}{2}}\cdot\left(1-\underline{\theta}\right)^{\frac{k}{2}}-\frac{t\cdot\left(1-\underline{\theta}\right)^{\frac{k-1}{2}}\cdot\left(1-\underline{\theta}\right)^{\frac{k-1}{2}}-\frac{t\cdot\left(1-\theta\right)^{\frac{k-1}{2}}\cdot\left(1-\theta\right)^{$$

$$\frac{d}{d\sigma} \left(\log \left(\Theta^{\dagger} (1-\theta)^{h} \right) \right) = \frac{d}{d\sigma} \left(f \log \left(\sigma \right) + h \log \left(1-\sigma \right) \right)$$

$$= \frac{f}{\sigma} - \frac{h}{\sigma \sigma}$$

$$d/d\sigma \left(\frac{1}{6} - \frac{1}{4} - \sigma\right) = -\frac{1}{6}\sigma^2 + \frac{1}{4}(1-\sigma)^2 = \frac{1}{6}\sigma^2$$

for
$$f$$
 to have a maximum at an arbitrary point x_k we need $f(x_k) = 0$

$$f''(x_k) < 0$$

as for
$$\log(f(x)) = g(x)$$
, we have $\int g'(x) = 0 \in 0$ $\int f'(x) = 0$

$$\int g''(x) = \int f'(x) - \int f'(x)$$

so any local musiumm in f is also a local

noxinum of 8

- 2) concluding, we have that the computation of the by-likelihood is trivial compared to the standard likelihood, and
 the logarithm preserves the monohory of the original fraction templaing that are can compute the lay-likelihood when bying
 to some likelihood optimization problems.
- (8) find posterior distribution $\frac{\partial}{\partial x} p(\sigma/D) = p(\sigma) \frac{(D|\sigma)}{\rho(D)} \frac{\rho(\sigma)}{\rho(D)}$ prior is Sefer(a, b) = $\rho(\sigma)$
 - · so we know that p(D|O) is modeled as a binomial distribution with para. O
 - · since the andon uniable disjocites from a Benearli experiment are can assume that the variable is i.i.d.
 - =) likelihood = P(DIO) = om (1-0) N-m leading constant his so impact in maximizing likellood

$$P(0|0) = \left[e^{m} (1-0)^{n-m} \cdot \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} e^{a-1} (1-0)^{b-1} \right] / P(0)$$

$$\frac{d}{dt} = \frac{dt}{dt} = \frac{dt$$

-) we need to show that
$$\frac{m+a}{N+a-5} = \lambda \left(\frac{a}{a+5}\right) + (1-\lambda) O_{MLF}$$

The = ang max
$$\psi(D|\Phi)$$
 = ang max $\Phi^{m}(1-\theta)^{N-m}$ = ang max $\log(\Phi^{m}(1-\theta)^{N-m})$

= ang max $\log(\Phi)$ + $(N-m)\log(1-\Phi)$ = 0 to to

$$\frac{m}{\Phi} = \frac{(N-m)}{1-\Phi} = 0 = \frac{1-\Phi}{m} = \frac{N-m}{m} = 0 = \frac{1-M-m-m}{M} = 0 = \frac{1-M-m-m}{N-2m}$$

$$\frac{m}{\sigma} = \frac{(N-m)}{1-\sigma} = 0 = \frac{1-\sigma}{\sigma} = \frac{N-m}{\sigma} = \frac{1}{\sigma} = \frac{N-m-m}{\sigma} = \frac{1}{\sigma} = \frac{N-m-m}{\sigma} = \frac{1}{\sigma} = \frac{N-m-m}{\sigma} = \frac{1}{\sigma} = \frac{N-m-m}{\sigma} = \frac{1}{\sigma} = \frac{1}$$

$$\frac{m+a}{N+a-b} = \frac{\lambda \left(\frac{a}{a+b}\right) + \left(\frac{1-\lambda}{a-b}\right) \left(\frac{m}{N-lm}\right) = \frac{\lambda \left(\frac{c}{a+b}\right) + \frac{m}{N-lm} - \frac{\lambda m}{N+lm}}{N-lm} = \frac{\lambda \left(\frac{a}{a+b}\right) + \frac{m}{N-lm}}{\frac{m+a}{N+a-b} - \frac{m}{N-lm}}$$

$$\frac{m+a}{N+a-b} = \frac{m}{N-lm}$$

$$\frac{a+b}{N-lm} = \frac{m}{N-lm}$$

$$\frac{a+b}{N-lm} = \frac{m}{N-lm}$$

a value in
$$[0,1]$$

$$f(\lambda|x) \propto f(x|\lambda) p(\lambda) = \frac{\lambda^{2} - \lambda^{2}}{x!} \cdot \frac{\lambda^{2}}{\Gamma(a)} \lambda^{a-1} e^{-\lambda \lambda}$$

$$\langle \lambda \rangle^{\times} = \lambda^{-\lambda} \cdot \lambda^{\alpha-1} e^{-5\lambda} = \lambda^{\times} e^{-1} \cdot \lambda^{(b+1)}$$

$$|a_{1}| = |a_{1}| + |a_{2}| + |a_{3}| + |a_{4}| + |a_{$$

$$= 2) \times \lambda_{MFF} = any \max_{\lambda} (p(\lambda|x)) = \frac{x+\alpha-1}{b+1}$$