
Section 6: Principal Component and Maximum Covariance Analyses

Maximum Covariance Analysis (MCA)

Purpose

- **Purpose**
 - Find correlated patterns in TWO datasets (i.e. are frequently met simultaneously). E.g. patterns of temperature that go along with patterns of pressure.
- **Applications**
 - Study the coupling between parameters to understand physical mechanisms of climate variations. E.g. how does atmospheric circulation influence the distribution of temperature?
 - Statistical downscaling. (Translate GCM derived climate change scenarios to the local/regional scale.)
 - Reconstruction / Forecasting

Paired Multivariate Datasets

- **Two paired multivariate datasets**

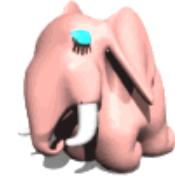
$$\mathbf{x}(i) = (x_1(i), \dots, x_n(i)) \quad \mathbf{y}(i) = (y_1(i), \dots, y_m(i))$$

- $i=1, \dots, N$, joint observations of \mathbf{x} , \mathbf{y}
- Two data clouds in two phase spaces
- Centered (means subtracted!), (just for more simple notation!)
- E.g. \mathbf{x} : anomalies of SLP Europe/Atlantic (n gridpoints),
 \mathbf{y} : Temperature anomalies in CH (m stations)

- **Data Matrices**

$$\mathbf{X} = \begin{bmatrix} x_1(1) & \cdots & x_n(1) \\ \vdots & \ddots & \vdots \\ x_1(N) & \cdots & x_n(N) \end{bmatrix} \quad \mathbf{Y} = \dots$$

Cross-Covariance



Call me
cross-covarofant

- **Cross-covariance matrix**

$$\mathbf{S}_{xy} := \begin{bmatrix} \text{cov}(x_1, y_1) & \text{cov}(x_1, y_j) & \text{cov}(x_1, y_m) \\ \text{cov}(x_i, y_1) & \text{cov}(x_i, y_j) & \vdots \\ \text{cov}(x_n, y_1) & \dots & \text{cov}(x_n, y_m) \end{bmatrix} = \frac{1}{(N-1)} \mathbf{X}^T \cdot \mathbf{Y}$$

- Univariate cross-covariances between all pairs of components
- $n \times m$ matrix
- In general not square and not symmetric
- diagonal elements do not have special meaning (i.e. different from var. matrix)

Cross-Correlation

- **Cross-Correlation matrix**

$$\mathbf{C}_{xy} := \begin{bmatrix} \text{cor}(x_1, y_1) & \text{cor}(x_1, y_j) & \text{cor}(x_1, y_m) \\ \text{cor}(x_i, y_1) & \text{cor}(x_i, y_j) & \vdots \\ \text{cor}(x_n, y_1) & \dots & \text{cor}(x_n, y_m) \end{bmatrix}$$

- Analogous to \mathbf{S}_{xy} but divided by standard deviations
- All matrix elements in $\{-1, +1\}$

$$\mathbf{C}_{xy} = \frac{1}{N-1} \mathbf{D}_x^{-\frac{1}{2}} \cdot \mathbf{X}^T \cdot \mathbf{Y} \cdot \mathbf{D}_y^{-\frac{1}{2}} = \mathbf{D}_x^{-\frac{1}{2}} \cdot \mathbf{S}_{xy} \cdot \mathbf{D}_y^{-\frac{1}{2}}$$

$$\mathbf{D}_x^{-\frac{1}{2}} = \left[\begin{array}{c} 1/\sigma_{x_k} \end{array} \right] \quad \mathbf{D}_y^{-\frac{1}{2}} = \left[\begin{array}{c} 1/\sigma_{y_k} \end{array} \right]$$

MCA Mathematical Procedure

- **Singular Value Decomposition (SVD)**

>> Appendix A

$$\mathbf{S}_{xy} = \text{cov}(\mathbf{X}, \mathbf{Y}) \quad \text{the cross-covariance matrix } (n \times m)$$

- There are r real numbers $\{\omega_1, \omega_2, \dots, \omega_r\}$, $\omega_k > 0$, *singular values*, (sorted in decreasing order $\omega_1 \geq \omega_2 \geq \omega_3 \dots \geq \omega_r$), ...
- ... and r vectors $\{\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_r\}$, n -dimensional, unit-length, orthogonal, and r vectors $\{\mathbf{v}_1, \mathbf{v}_2, \dots, \mathbf{v}_r\}$, m -dimensional, unit-length, orthogonal, called *left* (\mathbf{u}_k) and *right* (\mathbf{v}_k) *singular vectors*, ...
- ... such that:

$$\mathbf{S}_{xy} = \mathbf{U}^T \cdot \Omega \cdot \mathbf{V} \quad \Omega = [\omega_k], \quad \text{diagonal, } r \times r$$

$$\mathbf{U} = [\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_r], \quad \mathbf{V} = [\mathbf{v}_1, \mathbf{v}_2, \dots, \mathbf{v}_r], \quad \text{vectors in columns}$$

R: svd

MCA Interpretation

- **Singular Vectors, Coefficients**

- Consider singular vectors anchored in center of data clouds, spanning subspaces of the phasespaces of each dataset (new variables). They are not necessarily a basis system, because $r \leq n, r \leq m$
- Consider projections of data $\mathbf{x}(i)$ onto singular vectors as new data coordinates:

$$a_j(i) = \mathbf{x}(i)^T \cdot \mathbf{u}_j, \quad b_j(i) = \mathbf{y}(i)^T \cdot \mathbf{v}_j, \quad j = 1, \dots, r$$



projection of data vector on singular vector

- New coordinates are linear combinations of original variables.
- $a_j(i)$: left coefficients, (also left SVD scores)
 $b_j(i)$: right coefficients, (also right SVD scores)

MCA Interpretation

- **Singular values, cross-covariance**

- Ω is the cross-covariance matrix of the new coordinates $\{a_k\}$, $\{b_k\}$:

$$\text{cov}(a_i(.), b_j(.)) = 0 \quad \text{for } i \neq j, \quad \text{cov}(a_i(.), b_i(.)) = \omega_i \quad \gg \text{Appendix B}$$

- Coordinates corresponding to different indices of singular vectors are mutually uncorrelated.
 - The first pair of singular vectors $\{\mathbf{u}_1, \mathbf{v}_1\}$ are the phase-space directions, for which the projections have the largest possible cross-covariance. *First coupled mode*.
 - Subsequent vector pairs $\{\mathbf{u}_k, \mathbf{v}_k\}$ maximize cross-covariance subject to orthogonality to previous pairs. k^{th} *coupled mode*.

MCA Interpretation

- **Within space variance**
 - Singular vectors do not maximize variance in individual spaces.
 - Singular vectors are not necessarily aligned along directions of large data spread or cloud symmetry.
 - Left and right coefficients are in general not uncorrelated between themselves:

$$\text{var}(a_i, a_j) \neq 0 \quad \text{var}(b_i, b_j) \neq 0 \quad \text{for } i \neq j$$

- *Cumulative Explained Variance Fraction* of first l modes

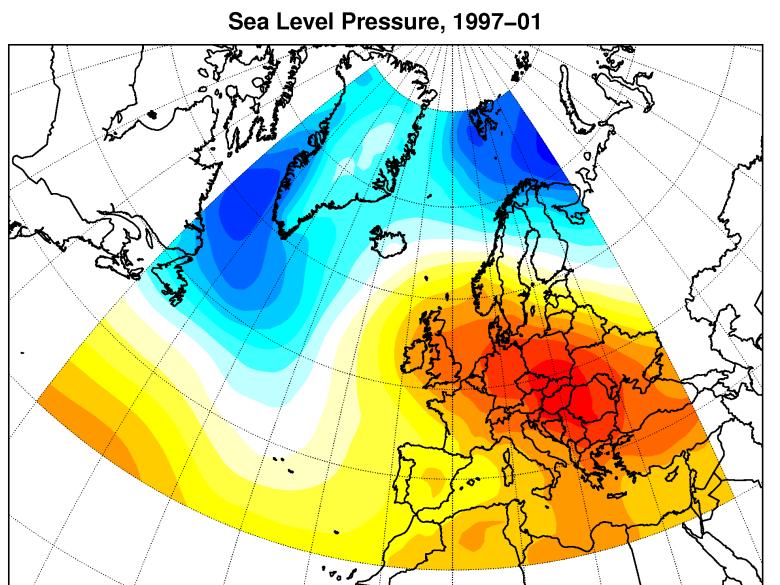
$$CEVF_x^l = \sum_k^l \text{var}(a_k) / \text{tr}(\mathbf{S}_{xx}) \quad CEVF_y^l = \sum_k^l \text{var}(b_k) / \text{tr}(\mathbf{S}_{yy})$$

MCA Interpretation

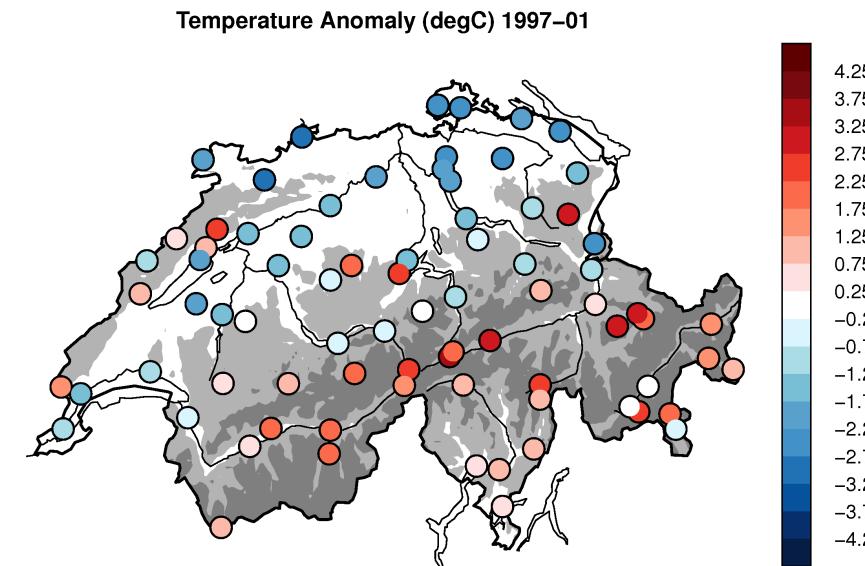
- Singular vectors describe patterns of anomalies in each dataset that tend to occur simultaneously (are linearly correlated). *Modes of co-variability, coupled modes, canonical pairs.*
- Coefficients represent amplitude (emphasis) of the respective patterns in each sample.
- The first few coefficient pairs have large cross-covariance and often show high correlations. *Dominant modes.*

Example: Swiss T <>> SLP

How does sea level pressure influence the distribution of winter-time temperature anomalies in Switzerland?



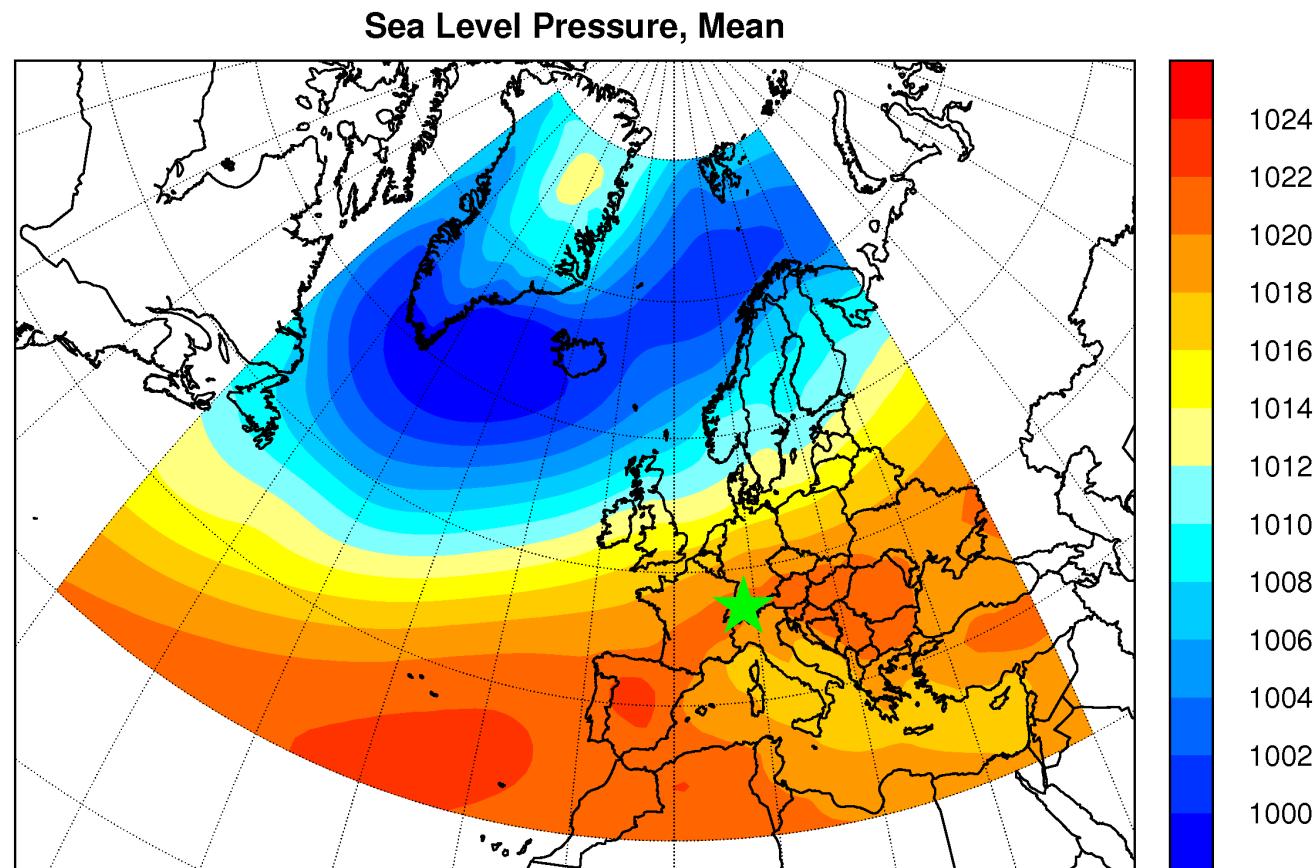
SLP: DJF, 1981 – 2010
(60°W – 40°E , 30°N – 80°N)
101 x 51 grid points, 1 degrees
ECMWF, ERA-Interim



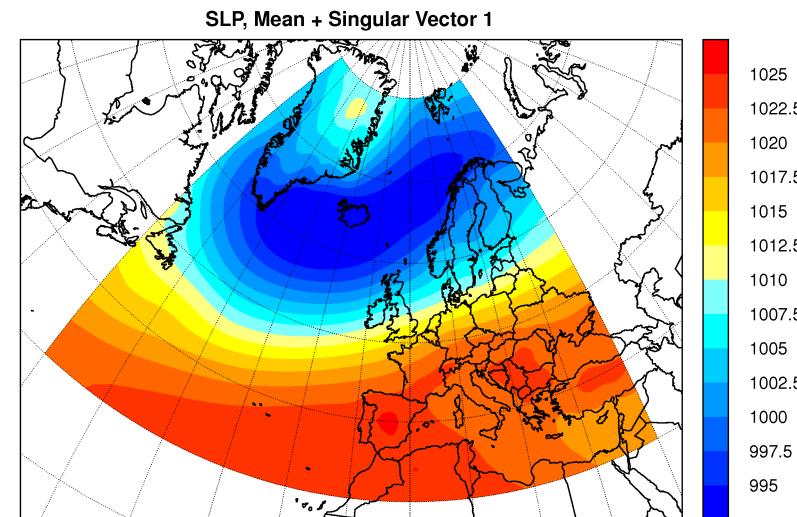
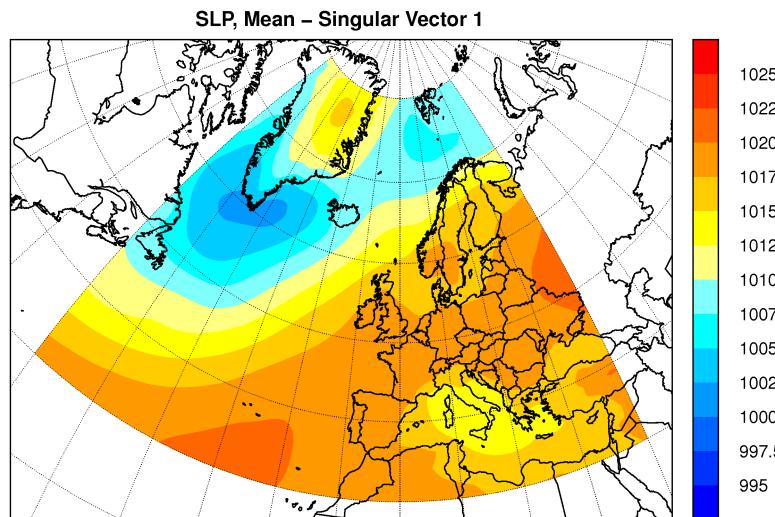
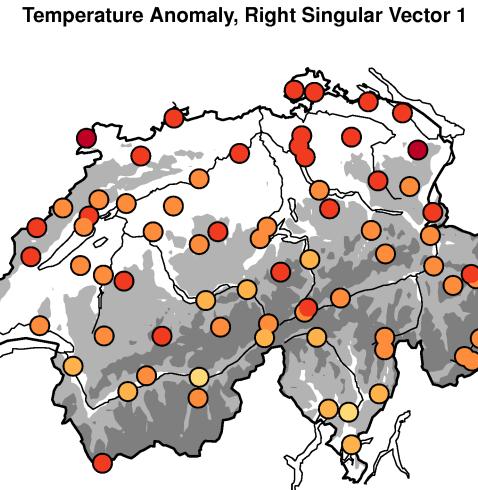
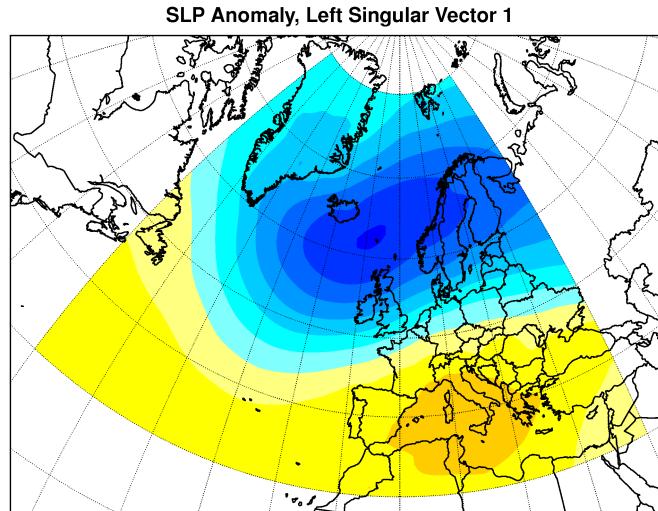
T: DJF, 1981 – 2010
79 stations, MeteoSwiss
Anomaly from Mean

Begert et al. 2005, Dee et al. 2011

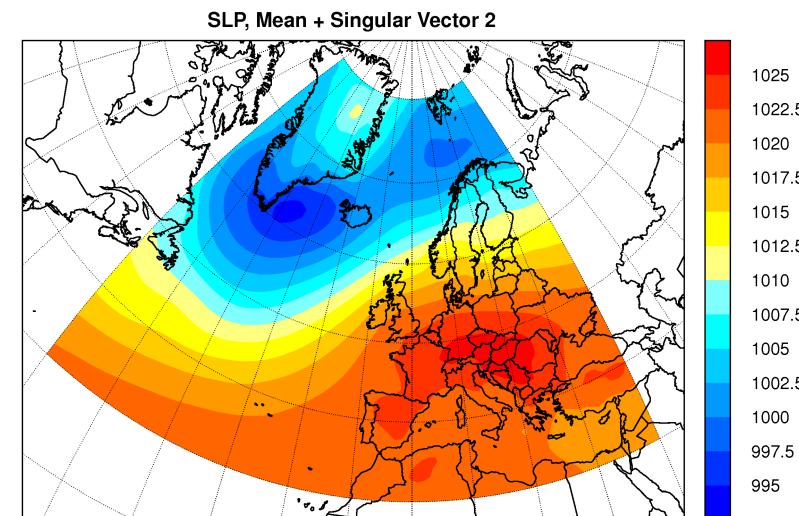
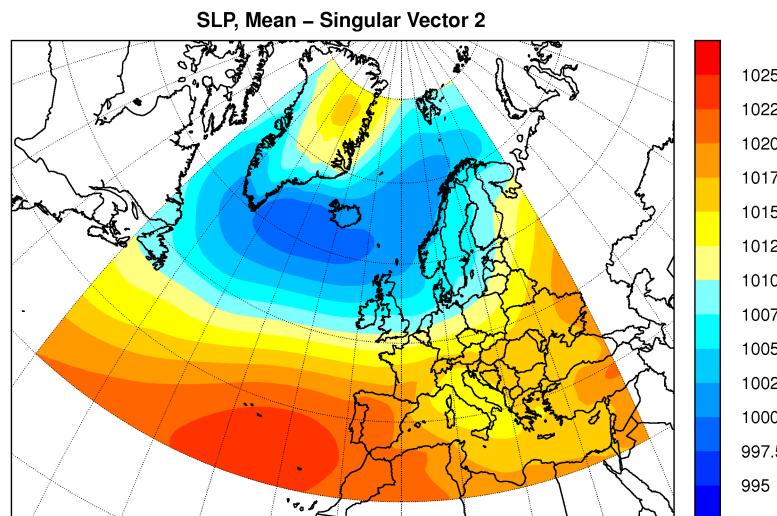
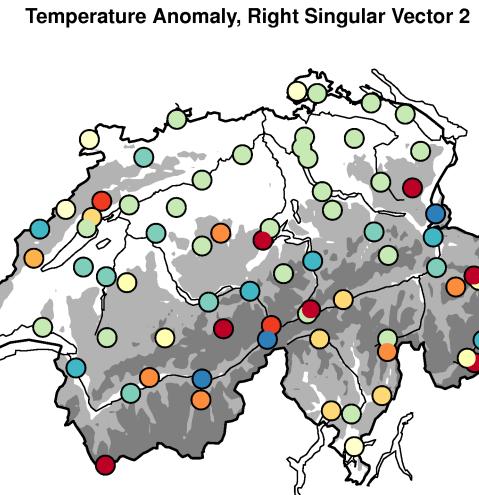
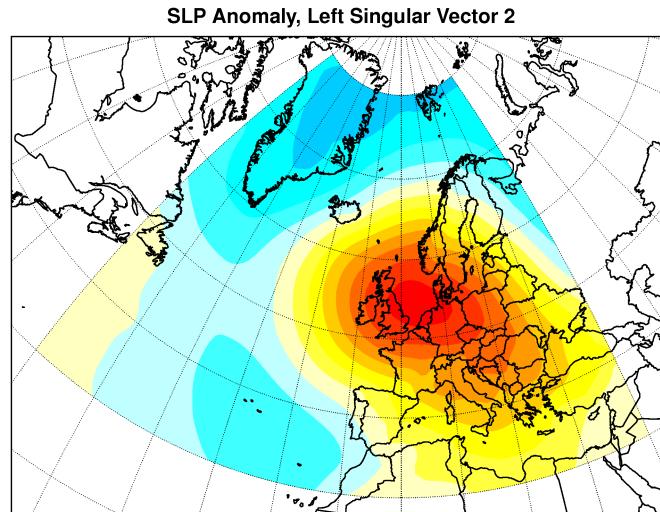
Example: Mean winter SLP



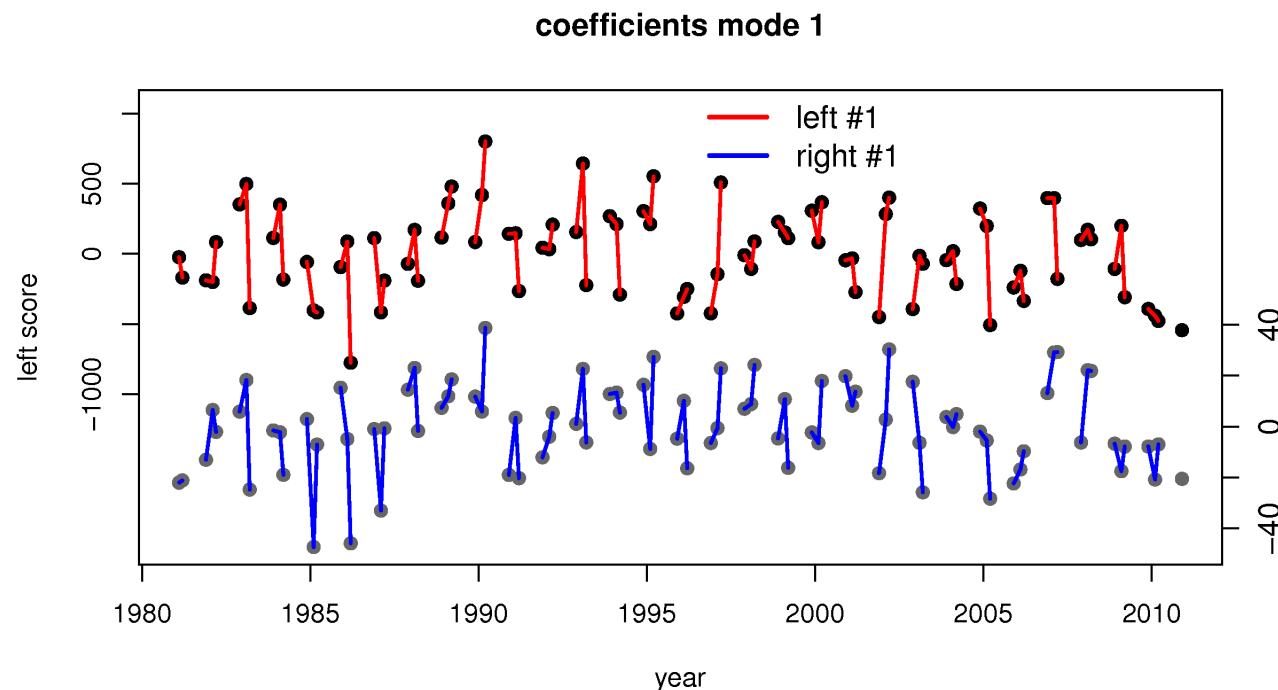
Example: Singular Vector Pair 1



Example: Singular Vector Pair 2

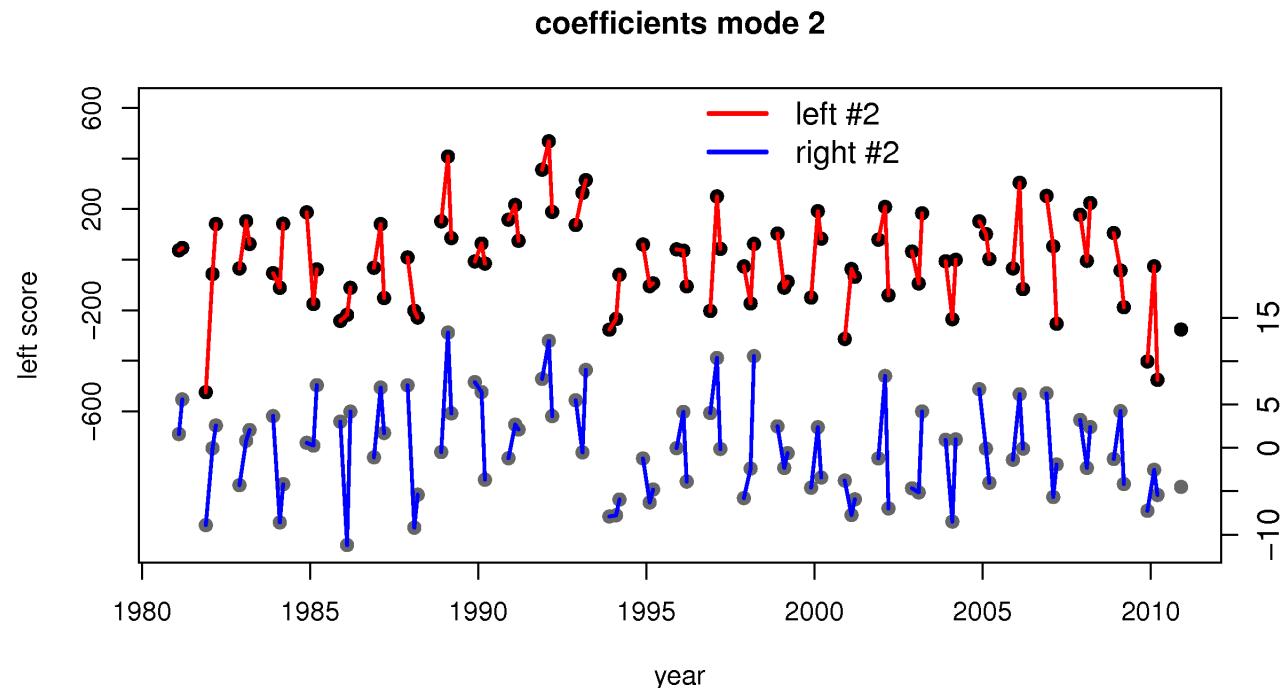


Example: Coefficients



Mode	Correlation
1	0.63
2	0.68
3	0.56
4	0.43
5	0.39
6	0.37
7	0.42
8	0.31

Example: Coefficients



Mode	Correlation
1	0.63
2	0.68
3	0.56
4	0.43
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8	0.31

Measuring Cross-Covariance

- **Total squared cross-covariance:**

$$\|S_{xy}\|_F := \sum_i^n \sum_j^m s_{ij}^2 = \sum_i^r \omega_i^2 \quad \text{Frobenius Norm}$$

- Note all matrix elements (different from PCA!)

- **Squared cross-covariance fraction:**

- of singular vector pair k :

$$SCF_k = \omega_k^2 / \sum_i^r \omega_i^2$$

- Note squared quantities (different from PCA!)

Truncation

- **Retain only first l coupled modes**
- **Projection onto first l modes**
 - yields an approximation of the original data

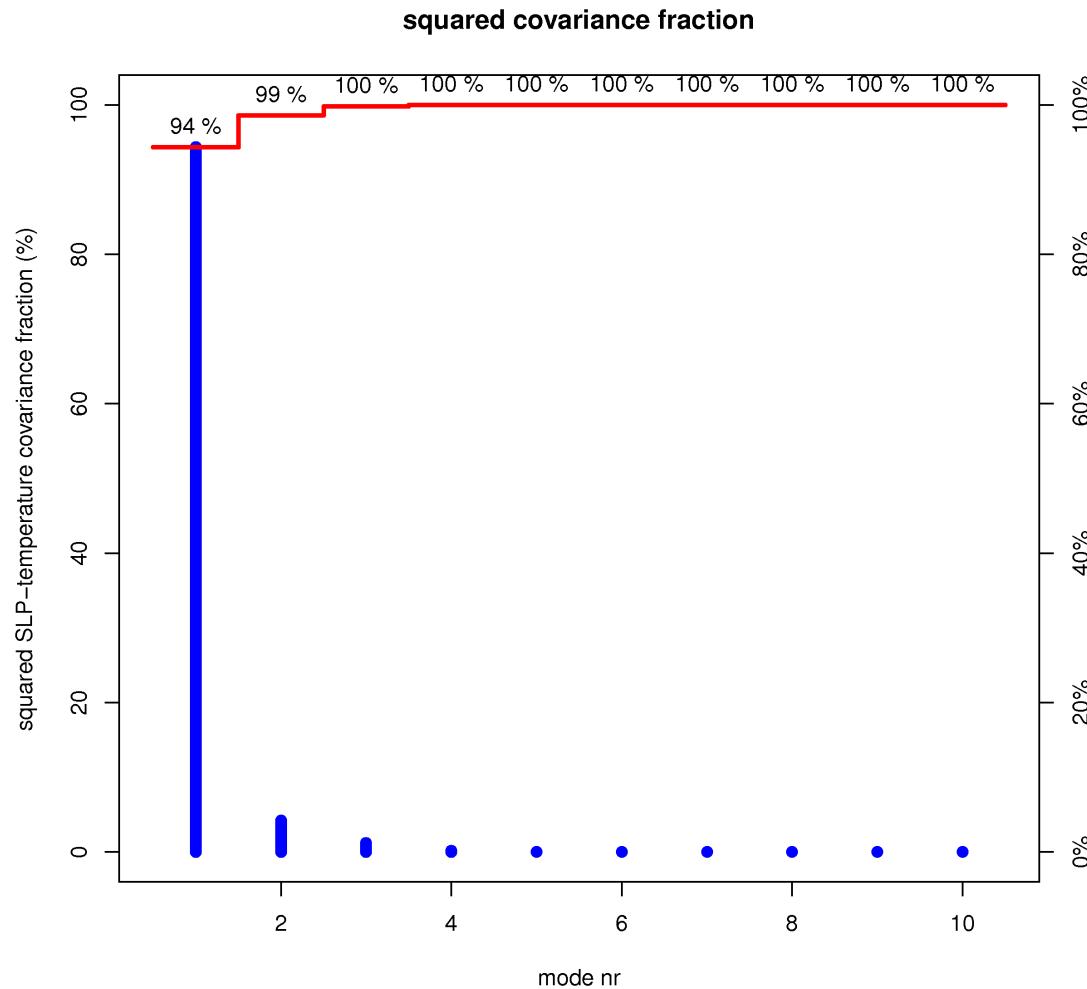
$$\tilde{\mathbf{x}}(i) = \sum_k^l a_k(i) \cdot \mathbf{u}_k \quad \tilde{\mathbf{y}}(i) = \sum_k^l b_k(i) \cdot \mathbf{v}_k \quad l \leq r$$

- **Cumulative squared cross-covariance fraction:**

Residual cross-covariance,
i.e. not reproduced by approximation

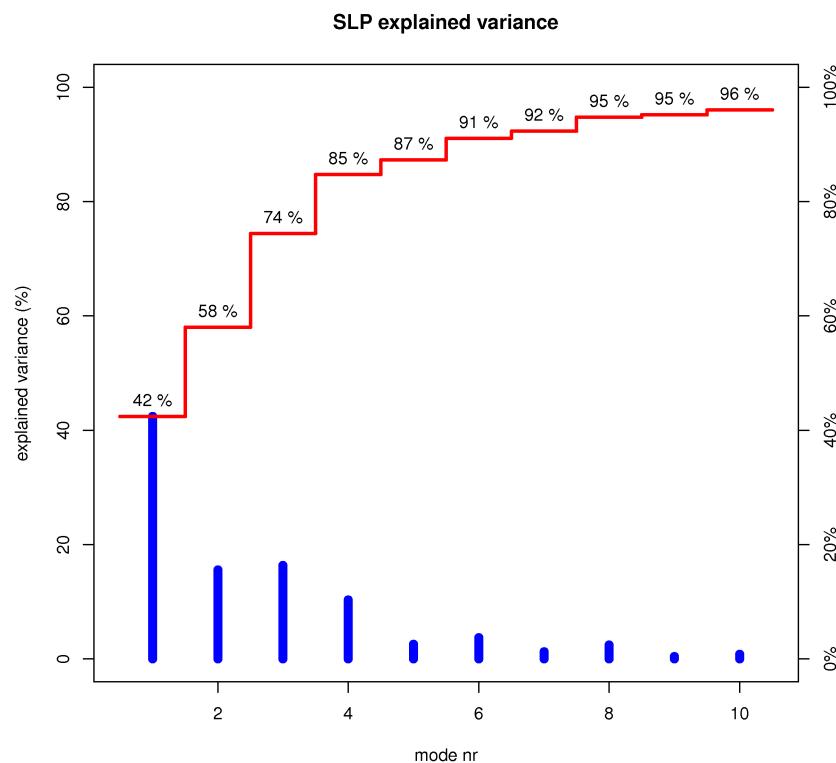
$$CSCF_l = 1 - \underbrace{\frac{\|\mathbf{S}_{xy} - \tilde{\mathbf{S}}_{xy}\|_F}{\|\mathbf{S}_{xy}\|_F}}_{\text{Residual cross-covariance, i.e. not reproduced by approximation}} = \sum_k^l \omega_k^2 / \sum_k^r \omega_k^2 = \sum_k^l SCF_k$$

Example: Squared Covariance Fraction

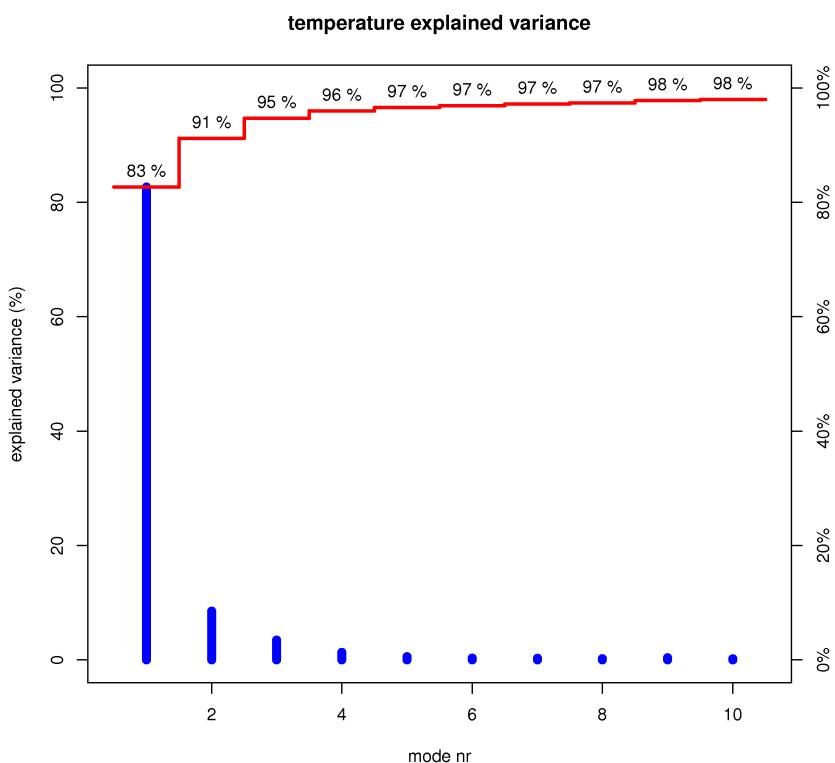


Example: Explained Variance

Sea Level Pressure



Temperature



Reconstruction / Prediction

- **Purpose**
 - Exploit cross-covariance to reconstruct/predict a right data vector from a left data vector (or vice versa).
- **Linear model between left and right coefficients**

 right coefficients  left coefficients (predictors)

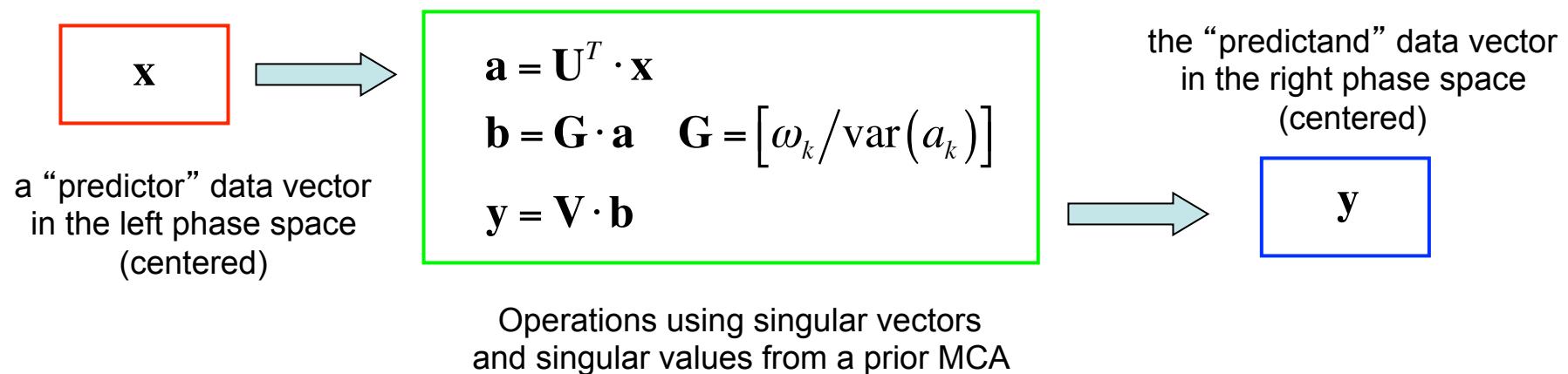
$$b_k(i) = \beta_{k0} + \sum_l^r \beta_{kl} \cdot a_l(i) + \varepsilon_k(i) \quad \leftarrow \text{noise}$$

  regression coefficients

- Simplifies to:

$$b_k(i) = \beta_{kk} \cdot a_k(i), \quad \beta_{kk} = \frac{\omega_k}{\text{var}(a_k)} \quad \begin{array}{l} a_k, b_l \text{ centered} \Rightarrow \beta_{k0} = 0 \\ \text{var}(a_k, b_l) = 0 \Rightarrow \beta_{kl} = 0, k \neq l \end{array}$$

Reconstruction / Prediction



- **MCA reconstruction equation:**

$$y = V \cdot G \cdot U^T \cdot x$$

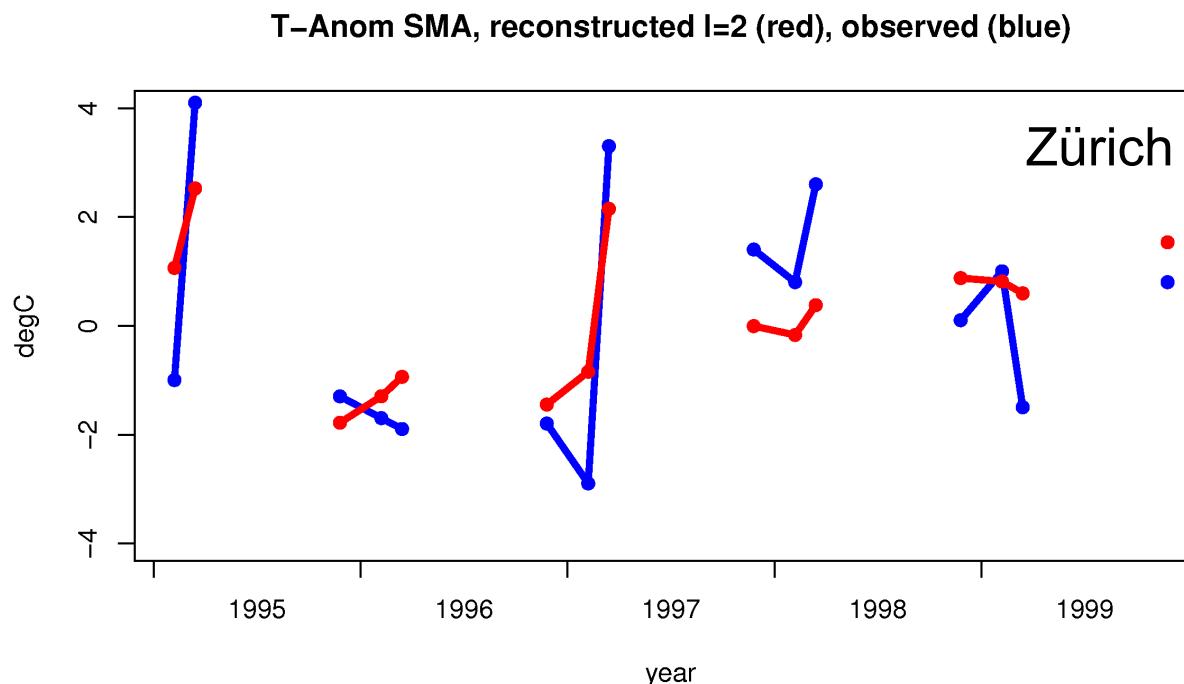
- Possibly only using a few leading coupled modes. I.e. all matrices truncated to the number of desired modes.

Reconstruction / Prediction

- **Accuracy of reconstruction depends on**
 - The degree of covariance between the two fields.
 - The cumulative squared cross-covariance fraction explained by the modes that are retained (truncation). The more modes the better the reconstruction.
 - The variance explained by the right singular vectors in the right phase space. I.e. the component of the right space that is related to the left space.

Example: Reconstruction

- **Reconstruction for 1995-1999 (winter months) using**
 - MCA calibrated for 1981-1994 & 2000-2010, i.e. without test period
 - SLP (left field) as predictor
 - 2 leading coupled modes

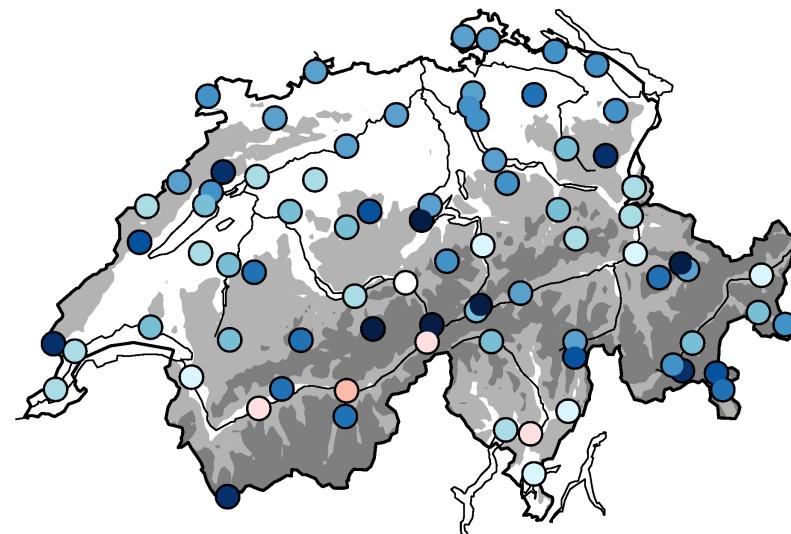


Example: Reconstruction

February 1996

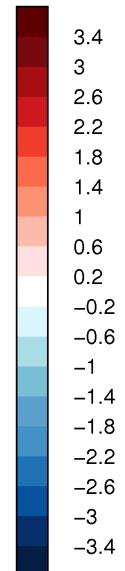
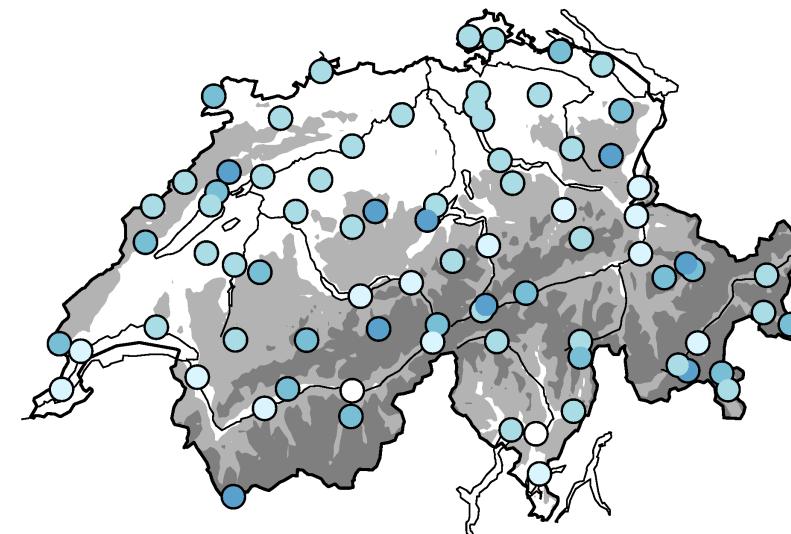
Observed

Temperature Anomaly (degC) 1996–02 observed



Reconstructed

Temperature Anomaly (degC) 1996–02 reconstructed l=2

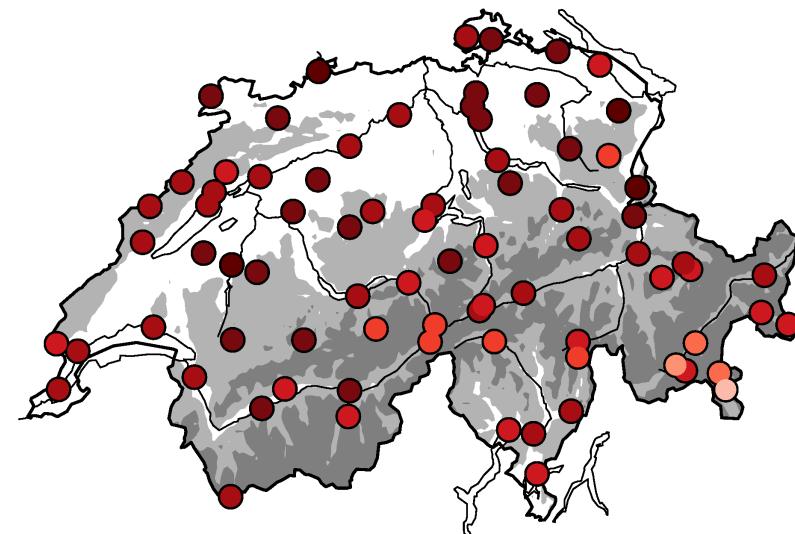


Example: Reconstruction

February 1997

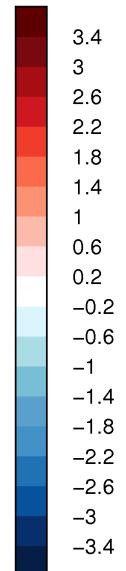
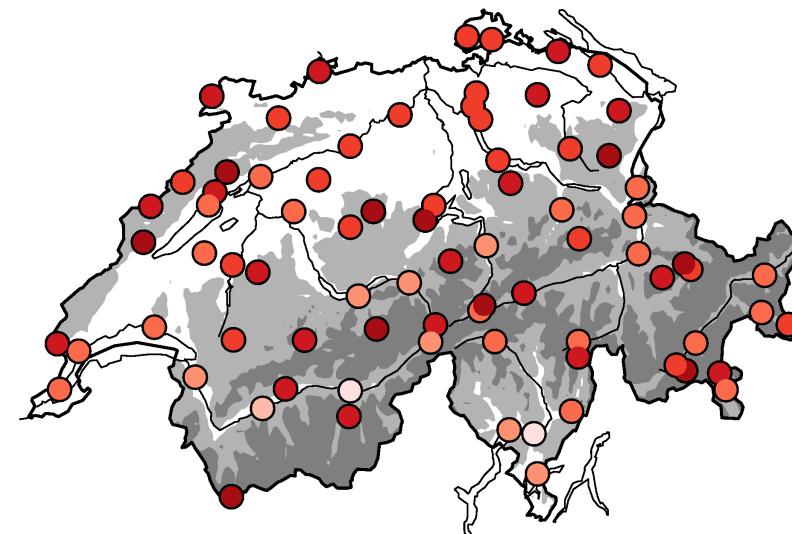
Observed

Temperature Anomaly (degC) 1997–02 observed



Reconstructed

Temperature Anomaly (degC) 1997–02 reconstructed l=2

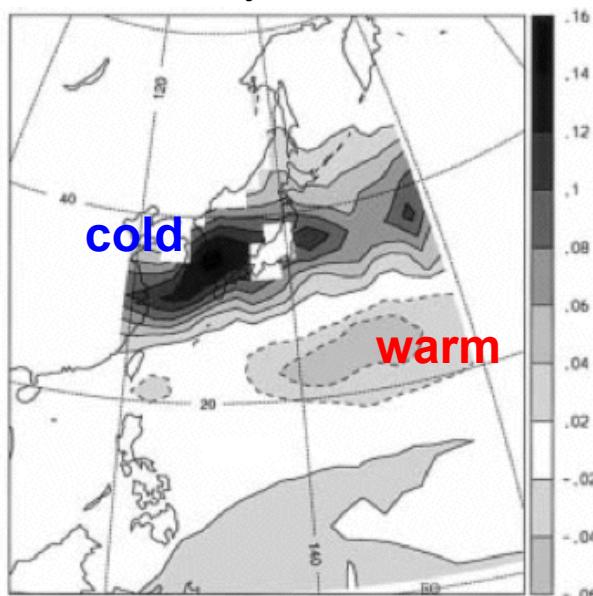


Example

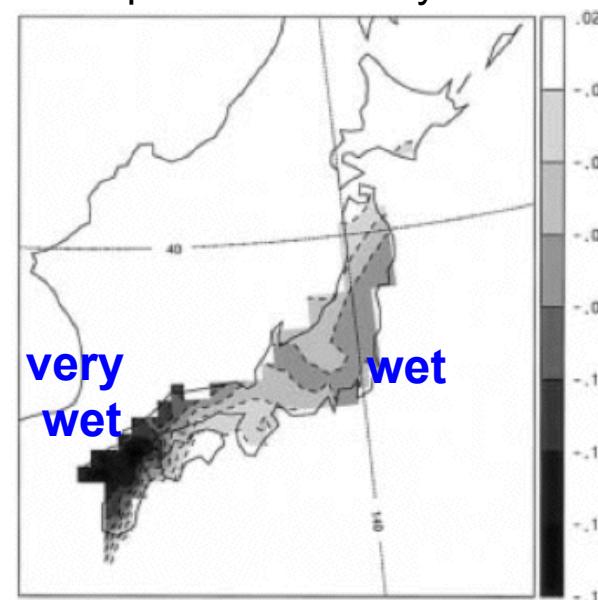
- **Baiu (JJA) Precipitation in Japan coupling to SST**

1st MCA mode

SST anomaly



Precipitation anomaly



What is cause,
what is effect?

1st mode explains 70% of total squared covariance fraction.
SST and Precip. Coefficients are correlated by $r=0.47$ ($r^2=0.22$).

Fukutome et al. 2003

Canonical Correlation Analysis (CCA)

- **Similar to MCA but searching for maximum correlation**
- **Procedure**
 - Conduct SVD with cross-correlation matrix
- **Interpretation (like for MCA)**
 - Canonical pairs describe coupled modes, emphasis on correlation
- **CCA or MCA?**
 - MCA focuses on covariance: Modes tend to be large where the variance is large. Danger that physical modes are confounded by large variance.
 - CCA focuses on correlation: Coupling is identified also if associated variance is low. Danger that physical modes are confounded by small (insignificant) variations (sampling problems).

see Wilks 2005, Chap 12 for details

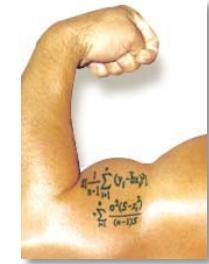
Summary

- **PCA & MCA are interesting techniques for characterising the (co-)variability in spatial datasets.**
- **Results require careful interpretation, corroboration in sensitivity experiments.**
- **Well established instruments in climate science:**
 - Technique of data reduction
 - Parsimonious reconstruction models (e.g. in historical climatology, paleoclimatology).
 - Evaluation of physical mechanisms in climate models.
 - Statistical (empirical) forecasting, climate change downscaling.
 - Source of hypothesis building for later modelling exercises.

Section 6: Principal Component and Maximum Covariance Analyses

Appendix MCA

Appendix A



- **Singular Value Decomposition (SVD)**
 - If \mathbf{Q} is a real-valued $n \times m$ matrix (e.g. a cross-covariance matrix) with rank $r \leq \min(n, m)$
 - Then there exist real-valued matrices \mathbf{U} , \mathbf{V} , Ω such that:

$$\mathbf{Q} = \mathbf{U} \cdot \Omega \cdot \mathbf{V}^T$$

- Ω : a diagonal $r \times r$ matrix with diagonal elements:

$$\omega_1 \geq \omega_2 \geq \dots \geq \omega_r > 0 \quad \text{the singular values}$$

- \mathbf{U} : a $n \times r$ matrix $\mathbf{U} = [\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_r]$, with orthonormal vectors: $\mathbf{U}^T \cdot \mathbf{U} = \mathbf{I}$

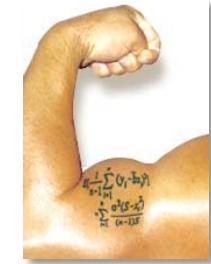
$$\mathbf{u}_k \quad k=1 \dots r \quad \text{the } \textit{left singular vectors}$$

- \mathbf{V} : a $m \times r$ matrix $\mathbf{V} = [\mathbf{v}_1, \mathbf{v}_2, \dots, \mathbf{v}_r]$, with orthonormal vectors: $\mathbf{V}^T \cdot \mathbf{V} = \mathbf{I}$

$$\mathbf{v}_k \quad k=1 \dots r \quad \text{the } \textit{right singular vectors}$$

See Wilks 2005, Sec. 9.3.5

Appendix A



- The singular vectors $\{\mathbf{u}_k\}$ and $\{\mathbf{v}_k\}$ ($k=1, \dots, r$) constitute orthonormal coordinate systems for r -dim subspaces of the n -dim (left) and m -dim (right) phase-spaces.
- I.e. the coordinate systems are not necessarily a basis system for the full phasespaces.
- The sum of squared singular values

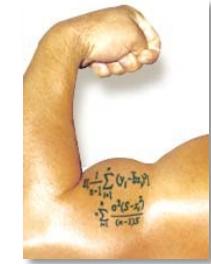
$$\|Q\|_F^2 := \sum_i^n \sum_j^m q_{ij}^2 = \sum_i^r \omega_i^2$$

the *Frobenius Norm*

= sum of squared matrix elements
= sum of squared singular values

see e.g. Bretherton et al. 1992

Appendix A



- **Calculation of SVD**

- Singular vectors $\{\mathbf{u}_k\}$ are eigenvectors of $\mathbf{Q}\mathbf{Q}^T$ (symmetric, $n \times n$) and $\{\mathbf{v}_k\}$ of $\mathbf{Q}^T\mathbf{Q}$ (symmetric, $m \times m$)

$$(\mathbf{Q} \cdot \mathbf{Q}^T) \cdot \mathbf{U} = \mathbf{U} \Omega \mathbf{V}^T \cdot \mathbf{V} \Omega^T \mathbf{U}^T \cdot \mathbf{U} = \mathbf{U} \cdot \Omega^2, \quad \text{dito for } \mathbf{V}$$

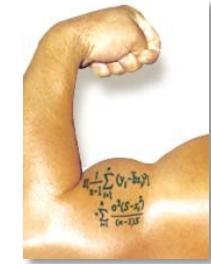
- Singular left vectors can be obtained directly from singular right vectors (and vice versa):

$$\mathbf{Q}\mathbf{V} = \mathbf{U}\Omega, \quad \mathbf{Q}^T\mathbf{U} = \mathbf{V}\Omega^T$$

$$\mathbf{Q}\mathbf{v}_k = \omega_k \mathbf{u}_k, \quad \mathbf{Q}^T \mathbf{u}_k = \omega_k \mathbf{v}_k, \quad k = 1, \dots, r$$

- In practice there is dedicated software to calculate SVD

Appendix B



- **Transformation of Cross-Covariance Matrix**

- If \mathbf{X} , \mathbf{Y} are centered data matrices, the cross-covariance matrix is:

$$\mathbf{S}_{xy} = \frac{1}{N-1} \mathbf{X}^T \cdot \mathbf{Y}$$

- Let \mathbf{U} , \mathbf{V} be transformation matrices (with projection vectors in columns), then the data matrices in transformed coordinates are:

$$\mathbf{A} = \mathbf{X} \cdot \mathbf{U}, \quad \mathbf{B} = \mathbf{Y} \cdot \mathbf{V}$$

- The cross-covariance matrix of the transformed variables is:

$$\mathbf{S}_{ab} = \frac{1}{N-1} \mathbf{A}^T \cdot \mathbf{B} = \frac{1}{N-1} \mathbf{U}^T \mathbf{X}^T \cdot \mathbf{Y} \mathbf{V} = \mathbf{U}^T \mathbf{S}_{xy} \mathbf{V}$$

- I.e. when \mathbf{U} , \mathbf{V} are chosen the singular vector systems of \mathbf{S}_{xy} then:

$$\begin{aligned} \mathbf{S}_{xy} &= \mathbf{U} \Omega \mathbf{V}^T \\ \mathbf{U}^T \cdot \mathbf{U} &= \mathbf{1}, \mathbf{V}^T \cdot \mathbf{V} = \mathbf{1} \end{aligned} \Rightarrow \mathbf{S}_{ab} = \mathbf{U}^T \mathbf{S}_{xy} \mathbf{V} = \mathbf{U}^T \mathbf{U} \Omega \mathbf{V}^T \mathbf{V} = \Omega$$