

Linear Regression HW1 MATH PART

① Prove: $\sum \hat{y}_i \hat{e}_i = 0$

Proof: We know $\sum \hat{e}_i = 0$. In SLR, we minimize

$$Q = \sum_{i=1}^n (Y_i - (\beta_0 + \beta_1 x_i))^2$$

$$\text{Set } \frac{\partial Q}{\partial \beta_0} = \sum_{i=1}^n 2(Y_i - (\beta_0 + \beta_1 x_i))(-1) = 0$$

$$\frac{\partial Q}{\partial \beta_1} = -2 \sum_{i=1}^n (Y_i - (\beta_0 + \beta_1 x_i)) x_i = 0$$

let $\hat{y}_i = \beta_0 + \beta_1 x_i$, $\hat{e}_i = y_i - \hat{y}_i$
and we get

$$\sum_{i=1}^n (Y_i - \hat{y}_i) = 0 \Rightarrow \sum_{i=1}^n \hat{e}_i = 0$$

$$\sum_{i=1}^n x_i (Y_i - \hat{y}_i) = 0 \Rightarrow \sum_{i=1}^n x_i \hat{e}_i = 0$$

Since $\beta_1 x_i + \beta_0 = \hat{y}_i$,

$$\frac{\hat{y}_i - \beta_0}{\beta_1} = x_i$$

$$\text{so } \sum_{i=1}^n x_i \hat{e}_i = 0$$

$$\Rightarrow \sum_{i=1}^n \frac{(\hat{y}_i - \beta_0)}{\beta_1} \hat{e}_i = 0$$

$$\Rightarrow \frac{1}{\beta_1} \sum_{i=1}^n (\hat{y}_i \hat{e}_i - \beta_0 \hat{e}_i) = 0$$

$$\Rightarrow \frac{1}{\beta_1} \left(\sum_{i=1}^n \hat{y}_i \hat{e}_i - \beta_0 \sum_{i=1}^n \hat{e}_i \right) = 0 \quad \text{and} \quad \sum_{i=1}^n \hat{e}_i = 0$$

$$\text{so } \frac{1}{\beta_1} \sum_{i=1}^n \hat{y}_i \hat{e}_i = 0$$

$$\Rightarrow \sum_{i=1}^n \hat{y}_i \hat{e}_i = 0$$

$$\textcircled{2} \sum (y_i - \hat{y}_i)(\hat{y}_i - \bar{y})$$

$$= \sum \hat{e}_i (\hat{y}_i - \bar{y})$$

$$= \sum \hat{e}_i \hat{y}_i - \sum \hat{e}_i \bar{y} \quad \text{as proved in } \textcircled{1}, \sum_{i=1}^n \hat{e}_i \hat{y}_i = 0$$

$$\text{and } \sum_{i=1}^n \hat{e}_i = 0$$

$$= 0$$

B.)

$$\hat{\beta}_{y|x} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} = \frac{\text{cov}(x, y)}{\text{var}(x)} = r_{xy} \frac{s_y}{s_x}.$$

$$\text{sd}(x) = \text{sd}(y) \Rightarrow \text{var}(x) = \text{var}(y)$$

$$\Rightarrow \hat{\beta}_{y|x} = \frac{\text{cov}(x, y)}{\text{var}(y)} = \hat{\beta}_{x|y}$$

$$s_y = s_x \Rightarrow \frac{s_y}{s_x} = 1 \Rightarrow \hat{\beta}_{y|x} = \hat{\beta}_{x|y} = r_{xy}.$$

$\hat{\alpha} = \bar{y} - \hat{\beta} \bar{x}$. \bar{y} & \bar{x} are constants.
So $\hat{\alpha}_{y|x} = \hat{\alpha}_{x|y}$ as $\hat{\beta}_{y|x} = \hat{\beta}_{x|y}$

C) let $\alpha_{x|y} = \alpha_{y|x} = \alpha$, $\beta_{x|y} = \beta_{y|x} = \beta$.

then we have $y = \beta x + \alpha$ and $x = \beta y + \alpha$

$$\Rightarrow \frac{1}{\beta} x - \frac{\alpha}{\beta} = y \text{ \& \; } \beta x + \alpha = y$$

Since $\frac{1}{\beta} = \beta$ and $-\frac{\alpha}{\beta} = \alpha$ can't be realized they're different.

3. This is a biased sample. The children who had difficulty in reading cannot represent the whole population, because proper randomization is not achieved. Also, they should set up a control group and an experimental group, to make sure that the improvement comes from the treatment. Otherwise, I could propose that

the improvement is due to other reasons such as the natural mental development of kids in a year.