

# BST 222 - Homework 2

Xiaowei Zeng

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## Problem 1

The SSA standard mortality table for 2019 is found at [Mortality Table](#). Compute the net present value of a series of two payments of \$1000 at the end of year 1 and at the end of year 2, each payment contingent on at least one of two people being alive at the time of the scheduled payment. Individual 1 is male aged 80 and individual 2 is female aged 75. You can assume that the chance of surviving a year at a given age is as given in the table. Use a discount rate of 5%.

Table 1: Mortality Table (Extracted)

Individual	Gender	Age	Death probability
1	Male	80	0.056237
		81	0.062360
2	Female	75	0.024080
		76	0.026831

**Solution:** Pension calculation:

$$V = \sum_{t=1}^{\infty} \frac{M}{(1+i)^t} \prod_{s=0}^{t-1} P(A+S)$$

where  $M$  is the annual amount paid at the end of the year,  $i$  is the discount rate,  $A$  is the initial age, and  $P(x)$  is the probability of surviving until the end of year at age  $x$ .

a. At the end of year 1:  $M = 1000, i = 0.05, t = 1$ .

$$\begin{aligned} P(\text{at least one survives}) &= 1 - P(\text{both die}) \\ &= 1 - P(1 \text{ dies}) \times P(2 \text{ dies}) \\ &= 1 - 0.056237 \times 0.024080 \\ &= 0.9986458 \end{aligned}$$

$$V_1 = \frac{1000}{1 + 0.05} * 0.9986458 = 951.0912$$

$\therefore$  the net present value is \$951.0912 at the end of year 1.

b. At the end of year 2:  $M = 1000, i = 0.05, t = 2$ .

**Case 1:** Both survive 2 years.

$$\begin{aligned} P(\text{Case 1}) &= P(1 \text{ survives in year 1}) \times P(1 \text{ survives in year 2}) \\ &\quad \times P(2 \text{ survives in year 1}) \times P(2 \text{ survives in year 2}) \\ &= (1 - 0.056237) \times (1 - 0.062360) \times (1 - 0.024080) \times (1 - 0.026831) \\ &= 0.84043 \end{aligned}$$

**Case 2:** Individual 1 survives 2 years, individual 2 dies in year 1.

$$\begin{aligned} P(\text{Case 2}) &= P(1 \text{ survives in year 1}) \times P(1 \text{ survives in year 2}) \\ &\quad \times P(2 \text{ dies in year 1}) \\ &= (1 - 0.056237) \times (1 - 0.062360) \times 0.024080 \\ &= 0.02130863 \end{aligned}$$

**Case 3:** Individual 1 survives 2 years, individual 2 dies in year 2.

$$\begin{aligned} P(\text{Case 3}) &= P(1 \text{ survives in year 1}) \times P(1 \text{ survives in year 2}) \\ &\quad \times P(2 \text{ survives in year 1}) \times P(2 \text{ dies in year 2}) \\ &= (1 - 0.056237) \times (1 - 0.062360) \times (1 - 0.024080) \times 0.026831 \\ &= 0.02317129 \end{aligned}$$

**Case 4:** Individual 2 survives 2 years, individual 1 dies in year 1.

$$\begin{aligned} P(\text{Case 4}) &= P(2 \text{ survives in year 1}) \times P(2 \text{ survives in year 2}) \\ &\quad \times P(1 \text{ dies in year 1}) \\ &= (1 - 0.024080) \times (1 - 0.026831) \times 0.056237 \\ &= 0.05341025 \end{aligned}$$

**Case 5:** Individual 2 survives 2 years, individual 1 dies in year 2.

$$\begin{aligned} P(\text{Case 5}) &= P(2 \text{ survives in year 1}) \times P(2 \text{ survives in year 2}) \\ &\quad \times P(1 \text{ survives in year 1}) \times P(1 \text{ dies in year 2}) \\ &= (1 - 0.024080) \times (1 - 0.026831) \times (1 - 0.056237) \times 0.062360 \\ &= 0.05589482 \end{aligned}$$

Therefore, we have

$$\begin{aligned}
 P(\text{at least one survives}) &= \sum_{i=1}^5 P(\text{Case } i) \\
 &= 0.84043 + 0.02130863 + 0.02317129 \\
 &\quad + 0.05341025 + 0.05589482 \\
 &= 0.994215
 \end{aligned}$$

$$V_2 = \frac{1000}{(1 + 0.05)^2} * 0.994215 = 901.7823$$

$$V_1 + V_2 = 951.0912 + 901.7823 = 1852.8735$$

$\therefore$  the net present value is \$1852.8735 at the end of year 2.

## Problem 2

The data set `tongue` from `KMsurv` is described in KM section 1.11.

Description:

The 'tongue' data frame has 80 rows and 3 columns.

Format:

This data frame contains the following columns:

1. `type`: Tumor DNA profile (1=Aneuploid Tumor, 2=Diploid Tumor)
2. `time`: Time to death or on-study time, weeks
3. `delta`: Death indicator (0=alive, 1=dead)

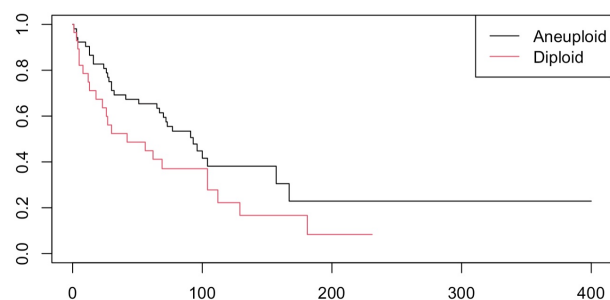
- (a) Construct and plot the Kaplan-Meier survival function estimates for the two types of tumors, first without the confidence limits and then with the confidence limits.
- (b) Use `survdif` to test the hypothesis that the two true survival curves are the same. Interpret the results.
- (c) Compare the two Kaplan-Meier curves to the Nelson-Aalen estimates graphically. Interpret the results.
- (d) To investigate the proportionality of the two hazard curves, compute and plot the Nelson-Aalen cumulative hazards, the ratio of the cumulative hazards, and the smoothed hazards using `muhaz`. Does it look as if the hazards are proportional?

**Solution:** Load the packages and data in R.

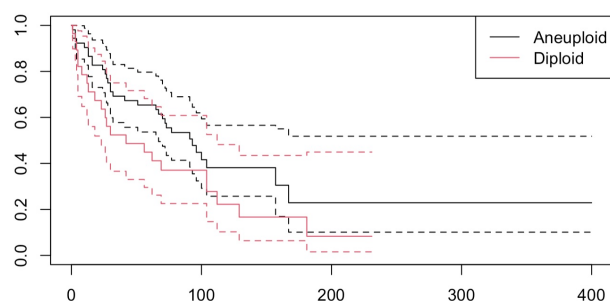
```
1 library(KMsurv)
2 library(survival)
3 library(muhaz)
4 data('tongue')
5 dfsurv = Surv(tongue$time, tongue$delta) # create a survival object
```

(a) Plot the Kaplan-Meier curve for the two types of tumor.

```
1 # Without CI.
2 plot(survfit(dfsurv ~ type, data = tongue), col = 1:2)
3 legend('topright', c("Aneuploid", 'Diploid Tumor'), col = 1:2)
```



```
1 # With CI.
2 plot(survfit(dfsurv ~ type, data = tongue), col = 1:2, conf.int = T)
3 legend('topright', c("Aneuploid", 'Diploid Tumor'), col = 1:2)
```



(b) Suppose that the significance level  $\alpha$  is 0.05.

```
1 survdiff(dfsurv ~ type, data = tongue)
```

Call:

```
survdiff(formula = dfsurv ~ type, data = tongue)
```

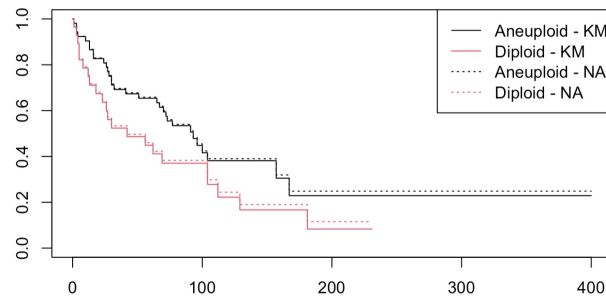
	N	Observed	Expected	$(O-E)^2/E$	$(O-E)^2/V$
type=1	52	31	36.6	0.843	2.79
type=2	28	22	16.4	1.873	2.79

Chisq= 2.8 on 1 degrees of freedom, p= 0.09

The p-value of the test is  $0.09 > 0.05$ , so we **cannot reject** the null hypothesis that the two true survival curves are the same.

- (c) We can use the option `type = 'fleming-harrington'` in `survfit` to get the Nelson-Aalen estimates. Note that Fleming Harrington method will reduce to Nelson-Aalen method when the data is unweighted.

```
1 plot(survfit(dfsurv ~ type, data = tongue), col = 1:2)
2 lines(survfit(dfsurv ~ type, data = tongue, type = 'fleming-harrington'),
3       col = 1:2, lty = 3)
4 legend('topright', c("Aneuploid - KM", 'Diploid - KM',
5                       "Aneuploid - NA", 'Diploid - NA'),
6       col = c(1:2, 1:2), lty = c(1, 1, 3, 3), lwd = 1)
```



From the graph above, we can learn that, the Nelson-Aalen estimates for the survival function are a bit larger than the Kaplan-Meier estimates for each type of tumor. This is because

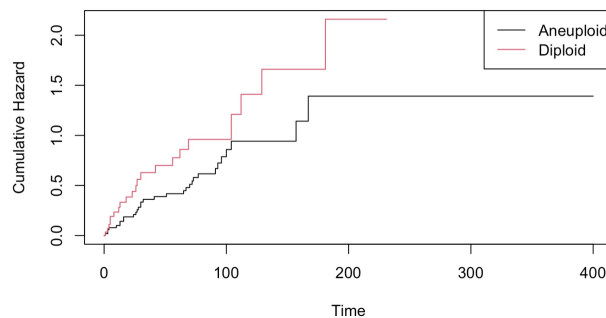
$$\hat{S}_{KM}(t) = \prod_{t_i < t} \left[ 1 - \frac{d_i}{Y_i} \right], \hat{S}_{NA}(t) = \prod_{t_i < t} \exp \left( -\frac{d_i}{Y_i} \right).$$

$$\because \exp(-x) - (1 - x) > 0, \forall x \in (0, 1),$$

$$\therefore \hat{S}_{NA}(t) > \hat{S}_{KM}(t), \forall x \in (0, 1).$$

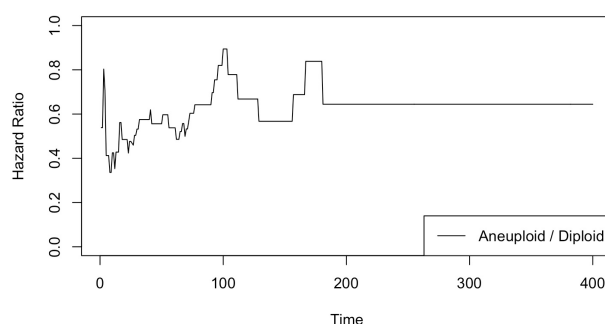
- (d) Nelson-Aalen cumulative hazards:

```
1 plot(survfit(dfsurv ~ type, data = tongue, type = 'fleming-harrington'),
2       col = 1:2, fun = 'cumhaz', xlab = 'Time', ylab = 'Cumulative Hazard')
3 legend('topright', c("Aneuploid", 'Diploid'), col = 1:2, lwd = 1)
```



Ratio of the cumulative hazards:

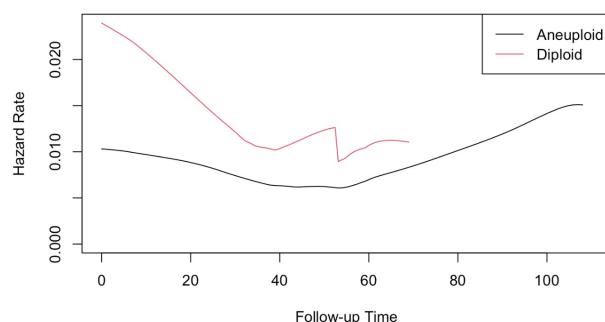
```
1 naft = survfit(dfsurv ~ type, data = tongue, type = 'fleming-harrington')
2 timevec = 1:400
3 sf1 = stepfun(naft[1]$time, c(1, naft[1]$surv))
4 sf2 = stepfun(naft[2]$time, c(1, naft[2]$surv))
5 cumhaz1 = - log(sf1(timevec))
6 cumhaz2 = - log(sf2(timevec))
7 plot(timevec, cumhaz1 / cumhaz2, type = 'l', ylim = c(0, 1),
8       xlab = 'Time', ylab = 'Hazard Ratio')
9 legend('bottomright', 'Aneuploid / Diploid', lwd = 1)
```



The graph shows that the ratio of Aneuploid / Diploid fluctuates slightly between 0.4 and 0.8 and is always below 1, so we can assume that the hazards are proportional.

Smoothed hazards:

```
1 plot(muhaz(tongue$time, tongue$delta, tongue$type == 2), col = 2,
2       xlim = c(0, 110))
3 lines(muhaz(tongue$time, tongue$delta, tongue$type == 1), col = 1)
4 legend('topright', c("Aneuploid", "Diploid"), col = 1:2, lwd = 1)
```



According to the graph, the two hazard curves have similar trends and are roughly paralleled, so we can assume that the hazards are proportional.