

## EE2 Mathematics – Probability & Statistics

### Solution 4

1. If  $X$  is the number of particles emitted in a half-second interval, then  $X \sim \text{Poisson}(1.6)$ , so

$$\begin{aligned} P(X \leq 2) &= P(X = 0) + P(X = 1) + P(X = 2) \\ &= \frac{e^{-1.6} 1.6^0}{0!} + \frac{e^{-1.6} 1.6^1}{1!} + \frac{e^{-1.6} 1.6^2}{2!} \\ &\approx 0.202 + 0.323 + 0.258 = 0.783 \end{aligned}$$

2. If  $T \sim \text{Exp}(2)$ , we know that  $F_T(t) = 1 - e^{-2t}$  for  $t > 0$ .

(a)  $P(T \leq 1) = F_T(1) = 1 - e^{-2} \approx 0.865$

(b)  $P(T > 3) = 1 - F_T(3) = e^{-6} \approx 0.0025$

(c)  $P(T > 3 | T > 2) = \frac{P(T > 3 \cap T > 2)}{P(T > 2)} = \frac{1 - F_T(3)}{1 - F_T(2)} = \frac{e^{-6}}{e^{-4}} \approx 0.135$ .

This is the same as  $P(T > 1)$ , because the exponential is memoryless.

3. Consider the random variable  $X$  with probability density function

$$f_X(x) = k \cos x, \quad -\frac{\pi}{2} \leq x \leq \frac{\pi}{2}$$

and zero otherwise.

(a)  $\int_{-\infty}^{\infty} f_X(x) dx = \int_{-\pi/2}^{\pi/2} k \cos x dx = k [\sin x]_{-\pi/2}^{\pi/2} = 2k$ , so  $k = 1/2$  in order to guarantee  $\int_{-\infty}^{\infty} f_X(x) dx = 1$ .

(b)  $E(X) = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-\pi/2}^{\pi/2} x \frac{\cos x}{2} dx = 0$ , by symmetry.

$$\begin{aligned} \text{Var}(X) &= E(X^2) = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \frac{1}{2} \int_{-\pi/2}^{\pi/2} x^2 \cos x dx \\ &= \frac{1}{2} \left[ [x^2 \sin x]_{-\pi/2}^{\pi/2} - 2 \int_{-\pi/2}^{\pi/2} x \sin x dx \right] \\ &\quad \text{(using } u = x^2 \text{ and } dv = \cos x dx) \\ &= \frac{1}{2} \left[ [x^2 \sin x + 2x \cos x]_{-\pi/2}^{\pi/2} - 2 \int_{-\pi/2}^{\pi/2} \cos x dx \right] \\ &\quad \text{(using } u = x \text{ and } dv = \sin x dx) \\ &= \frac{1}{2} [x^2 \sin x + 2x \cos x - 2 \sin x]_{-\pi/2}^{\pi/2} = \frac{\pi^2}{4} - 2 \approx 0.467. \end{aligned}$$

(c) Start by finding the CDF. For  $t \in [-\pi/2, \pi/2]$  we have

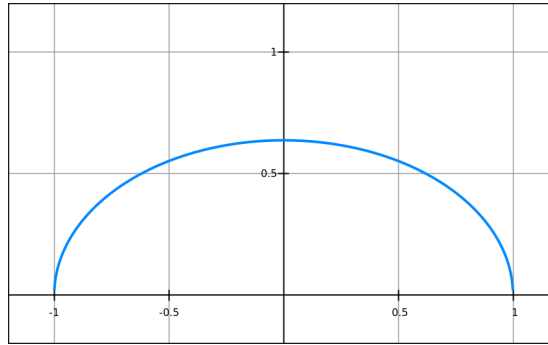
$$F_X(x) = \int_{-\infty}^x f_X(u) du = \int_{-\pi/2}^x \frac{\cos u}{2} du = \left[ \frac{\sin u}{2} \right]_{-\pi/2}^x = \frac{\sin x + 1}{2}.$$

The full CDF is, thus,

$$F_X(x) = \begin{cases} 0 & x < -\pi/2 \\ (\sin x + 1)/2 & -\pi/2 \leq x \leq \pi/2 \\ 1 & x > \pi/2 \end{cases}$$

We then have  $F_X(q_u) = 0.75 \Leftrightarrow q_u = \arcsin(1/2) = \pi/6 \approx 0.523$ . By symmetry,  $F_X(q_l) = 0.25 \Leftrightarrow q_l = \arcsin(-1/2) = -\pi/6 \approx -0.523$ , and  $\text{IQR} = \pi/3 \approx 1.046$ .

4.  $\int_{-\infty}^{\infty} f_X(x) dx = K \int_{-1}^1 \sqrt{1-x^2} dx = K\pi/2$ , because the integrand is the unit semicircle. Thus,  $K = 2/\pi$  in order to guarantee  $\int_{-\infty}^{\infty} f_X(x) dx = 1$ .



Alternatively, make the change of variable  $x = \sin u$  with  $-\pi/2 \leq u \leq \pi/2$  such that

$$\begin{aligned} K \int_{-1}^1 \sqrt{1-x^2} dx &= K \int_{-\pi/2}^{\pi/2} \cos^2 u du \\ &= \frac{K}{2} \int_{-\pi/2}^{\pi/2} (1 + \cos 2u) du \\ &= \frac{K}{2} \left[ u + \frac{1}{2} \sin 2u \right]_{-\pi/2}^{\pi/2} = K \frac{\pi}{2}. \end{aligned}$$

Thus,  $K = 2/\pi$ .

5.  $f(y)$  is always positive, so we just need to check that it integrates to 1,

$$\int_{-\infty}^{\infty} f(y) dy = \int_0^{\infty} \frac{y f_X(y)}{\mu} dy = \frac{1}{\mu} \underbrace{\int_0^{\infty} y f_X(y) dy}_{=E(X)=\mu} = 1.$$

Now let  $Y$  be a random variable with this density function, and notice that

$$E(Y^k) = \int_{-\infty}^{\infty} y^k f(y) dy = \int_0^{\infty} \frac{y^{k+1} f_X(y)}{\mu} dy = \frac{E(X^{k+1})}{\mu},$$

which implies  $E(X^{k+1}) = E(X) E(Y^k)$ . We can now prove the inequality directly:

$$\begin{aligned} E(X^3) E(X) &= E(X) E(Y^2) E(X) E(Y^0) = E(X)^2 E(Y^2) \\ &\geq E(X)^2 E(Y)^2 \quad (\text{recall that } \text{Var}(Y) = E(Y^2) - E(Y)^2 \geq 0) \\ &= \{E(X^2)\}^2 \quad (\text{from above, recall that } E(X^2) = E(X) E(Y)) \end{aligned}$$

6. (a)  $F_X(u) = \int_{-\infty}^u f_X(x) dx = \frac{1}{\pi} \int_{-\infty}^u \frac{1}{(1+x^2)} dx$ . Make the substitution  $x = \tan \theta$  (with  $dx = \sec^2 \theta d\theta$ ) such that

$$\begin{aligned} \int_{-\infty}^u \frac{1}{(1+x^2)} dx &= \int_{\theta_l}^{\theta_u} \underbrace{\frac{1}{(1+\tan^2 \theta)}}_{=\sec^2 \theta} \sec^2 \theta d\theta \\ &= \int_{\theta_l}^{\theta_u} d\theta = \theta_u - \theta_l = \arctan u + \frac{\pi}{2} \end{aligned}$$

where  $\theta_u = \arctan u$  and  $\theta_l = -\pi/2$ . We finally get  $F_X(u) = \frac{1}{\pi} \arctan u + \frac{1}{2}$ , for all  $u$ .

(b)  $F_X(u) = \int_{-\infty}^u f_X(x) dx = \int_{-\infty}^u \frac{e^{-x}}{(1+e^{-x})^2} dx = \left[ \frac{1}{1+e^{-x}} \right]_{-\infty}^u = \frac{1}{1+e^{-u}},$   
for all  $u$ .

(c)  $F_X(u) = \int_{-\infty}^u f_X(x) dx = \int_0^u \frac{(a-1)}{(1+x)^a} dx = \left[ \frac{-1}{(1+x)^{a-1}} \right]_0^u = 1 - \frac{1}{(1+u)^{a-1}},$   
for  $u > 0$ .

(d)  $F_X(u) = \int_{-\infty}^u f_X(x) dx = \int_0^u c\tau x^{\tau-1} e^{-cx^\tau} dx = \left[ -e^{-cx^\tau} \right]_0^u = 1 - e^{-cu^\tau},$  for all  $u > 0$ .

7. (a)  $f_X(x) = e^{-kx} x^{r-1} k^r / (r-1)!, x > 0, r = 1, 2, 3, \dots, k > 0$ .

$$E(X) = \int_0^{\infty} x e^{-kx} x^{r-1} \frac{k^r}{(r-1)!} dx = \frac{k^r}{(r-1)!} \int_0^{\infty} e^{-kx} x^r dx.$$

Making use of integration by parts with  $u = x^r$  and  $dv = e^{-kx} dx$ , we get

$$\begin{aligned} E(X) &= \frac{k^r}{(r-1)!} \left[ \frac{-x^r e^{-kx}}{k} \right]_0^{\infty} + \frac{k^r}{(r-1)!} \frac{r}{k} \underbrace{\int_0^{\infty} e^{-kx} x^{r-1} dx}_{=\frac{(r-1)!}{k^r} (f_X(x) \text{ is a valid pdf})} \\ &= \frac{k^r}{(r-1)!} [0 - 0] + \frac{k^r}{(r-1)!} \frac{r}{k} \frac{(r-1)!}{k^r} = \frac{r}{k}. \end{aligned}$$

$$E(X^2) = \int_0^\infty x^2 e^{-kx} x^{r-1} \frac{k^r}{(r-1)!} dx = \frac{k^r}{(r-1)!} \int_0^\infty e^{-kx} x^{r+1} dx.$$

Here again, we can use integration by parts

$$\begin{aligned} \int e^{-kx} x^{r+1} dx &= \frac{-x^{r+1} e^{-kx}}{k} + \frac{r+1}{k} \int e^{-kx} x^r dx \\ &= \frac{-x^{r+1} e^{-kx}}{k} + \frac{r+1}{k} \left[ \frac{-x^r e^{-kx}}{k} + \frac{r}{k} \int e^{-kx} x^{r-1} dx \right] \end{aligned}$$

such that

$$\begin{aligned} \int_0^\infty e^{-kx} x^{r+1} dx &= \left[ \frac{-x^{r+1} e^{-kx}}{k} - \frac{r+1}{k} \frac{x^r e^{-kx}}{k} \right]_0^\infty + \frac{r+1}{k} \frac{r}{k} \underbrace{\int_0^\infty e^{-kx} x^{r-1} dx}_{= \frac{(r-1)!}{k^r}} \\ &= [0 - 0] + \frac{r+1}{k} \frac{r}{k} \frac{(r-1)!}{k^r}. \end{aligned}$$

We finally get

$$E(X^2) = \frac{k^r}{(r-1)!} \frac{r+1}{k} \frac{r}{k} \frac{(r-1)!}{k^r} = \frac{r(r+1)}{k^2}$$

and

$$\text{Var}(X) = E(X^2) - E(X)^2 = \frac{r(r+1)}{k^2} - \frac{r^2}{k^2} = \frac{r}{k^2}$$

(b)  $f_X(x) = \binom{x+a-1}{x} p^x (1-p)^a$ ,  $x \in \mathbb{N}_0$ ,  $a \in \mathbb{N}^*$ ,  $p \in (0, 1)$ .

$$\begin{aligned}
E(X) &= \sum_{x=0}^{\infty} x \binom{x+a-1}{x} p^x (1-p)^a \\
&= \sum_{x=1}^{\infty} x \frac{(x+a-1)!}{x!(a-1)!} p^x (1-p)^a \quad (\text{ignore the term for } x=0) \\
&= \sum_{x=1}^{\infty} \frac{(x+a-1)!}{(x-1)!(a-1)!} p^x (1-p)^a \quad (\text{cancel out } x) \\
&= \sum_{y=0}^{\infty} \frac{(y+a)!}{y!(a-1)!} p^{y+1} (1-p)^a \quad (\text{set } y = x-1) \\
&= \sum_{y=0}^{\infty} \frac{(y+b-1)!}{y!(b-2)!} p^{y+1} (1-p)^{b-1} \quad (\text{set } b = a+1) \\
&= \frac{(b-1)p}{1-p} \underbrace{\sum_{y=0}^{\infty} \frac{(y+b-1)!}{y!(b-1)!} p^y (1-p)^b}_{=1} \\
&\quad (f_Y(y) \text{ is a valid neg. binomial pmf}) \\
&= \frac{ap}{1-p}
\end{aligned}$$

Following similar derivations, we compute

$$\begin{aligned}
E(X^2) &= \sum_{x=0}^{\infty} x^2 \binom{x+a-1}{x} p^x (1-p)^a \\
&= \sum_{x=1}^{\infty} x^2 \frac{(x+a-1)!}{x!(a-1)!} p^x (1-p)^a \quad (\text{ignore the term for } x=0) \\
&= \sum_{x=1}^{\infty} x \frac{(x+a-1)!}{(x-1)!(a-1)!} p^x (1-p)^a \quad (\text{cancel out } x) \\
&= \sum_{y=0}^{\infty} (y+1) \frac{(y+a)!}{y!(a-1)!} p^{y+1} (1-p)^a \quad (\text{set } y = x-1) \\
&= \sum_{y=0}^{\infty} (y+1) \frac{(y+b-1)!}{y!(b-2)!} p^{y+1} (1-p)^{b-1} \quad (\text{set } b = a+1) \\
&= \frac{(b-1)p}{1-p} \left[ \underbrace{\sum_{y=0}^{\infty} y \frac{(y+b-1)!}{y!(b-1)!} p^y (1-p)^b}_{=E(Y)=\frac{bp}{1-p}} + \underbrace{\sum_{y=0}^{\infty} \frac{(y+b-1)!}{y!(b-1)!} p^y (1-p)^b}_{=1} \right] \\
&\quad (f_Y(y) \text{ is a valid neg. binomial pmf}) \\
&= \frac{ap}{1-p} \left[ \frac{(a+1)p}{1-p} + 1 \right] = \frac{ap(ap+1)}{(1-p)^2}
\end{aligned}$$

such that

$$\text{Var}(X) = E(X^2) - E(X)^2 = \frac{ap(ap+1)}{(1-p)^2} - \left( \frac{ap}{1-p} \right)^2 = \frac{ap}{(1-p)^2}$$

8. (a) i. Exponential

$$\begin{aligned}
f_X(x) &= \begin{cases} \theta e^{-\theta x} & x > 0 \\ 0 & \text{otherwise} \end{cases} \\
F_X(x) &= \begin{cases} 1 - e^{-\theta x} & x > 0 \\ 0 & \text{otherwise} \end{cases} \\
\bar{F}_X(x) &= 1 - F_X(x) = \begin{cases} e^{-\theta x} & x > 0 \\ 1 & \text{otherwise} \end{cases} \\
\lambda(x) &= \frac{f_X(x)}{\bar{F}_X(x)} = \begin{cases} \theta & x > 0 \\ 0 & \text{otherwise} \end{cases}
\end{aligned}$$

The hazard rate is constant for all  $x > 0$ .

ii. Weibull

$$\begin{aligned}
f_X(x) &= c\tau x^{\tau-1} e^{-cx^\tau} & (0 < x < \infty) \\
F_X(x) &= 1 - e^{-cx^\tau} & (0 < x < \infty) \\
\bar{F}_X(x) &= 1 - F_X(x) = e^{-cx^\tau} & (0 < x < \infty) \\
\lambda(x) &= \frac{f_X(x)}{\bar{F}_X(x)} = c\tau x^{\tau-1} & (0 < x < \infty)
\end{aligned}$$

Increasing hazard

iii. Pareto

$$\begin{aligned}
f_X(x) &= \frac{(a-1)}{(1+x)^a} & (0 < x < \infty) \\
F_X(x) &= 1 - \frac{1}{(1+x)^{(a-1)}} & (0 < x < \infty) \\
\bar{F}_X(x) &= 1 - F_X(x) = \frac{1}{(1+x)^{(a-1)}} & (0 < x < \infty) \\
\lambda(x) &= \frac{f_X(x)}{\bar{F}_X(x)} = \frac{(a-1)}{(1+x)} & (0 < x < \infty)
\end{aligned}$$

Decreasing hazard

(b) Using the hint

$$\begin{aligned}
\frac{d}{dx} \log \frac{\bar{F}(x+y)}{\bar{F}(x)} &= \frac{d}{dx} [\log \bar{F}(x+y) - \log \bar{F}(x)] \\
&= \frac{-f_X(x+y)}{\bar{F}(x+y)} + \frac{f_X(x)}{\bar{F}(x)} = \lambda(x) - \lambda(x+y)
\end{aligned}$$

Hence the hazard rate does not decrease (i.e.  $\lambda(x) - \lambda(x+y) \leq 0$ ) if  $\frac{d}{dx} \log \frac{\bar{F}(x+y)}{\bar{F}(x)} \leq 0$ , i.e.  $\frac{\bar{F}(x+y)}{\bar{F}(x)}$  does not increase as  $x$  increases.