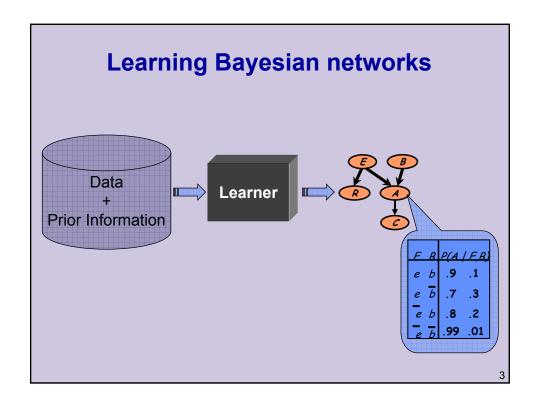
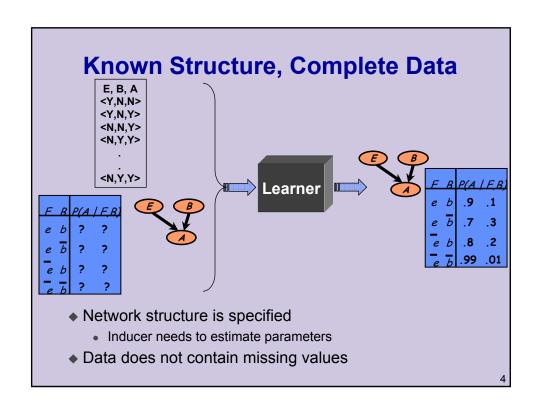
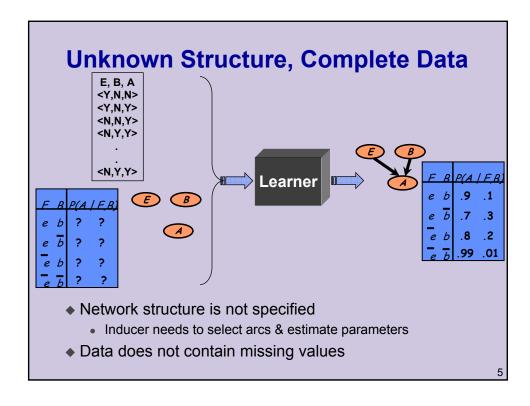
# **Bayesian Networks Learning**Parameter Estimation

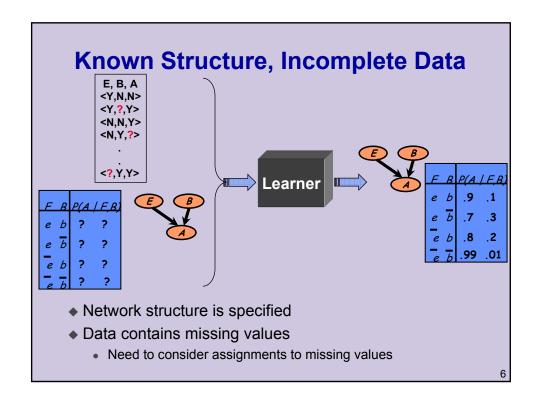
#### **Learning Bayesian networks**

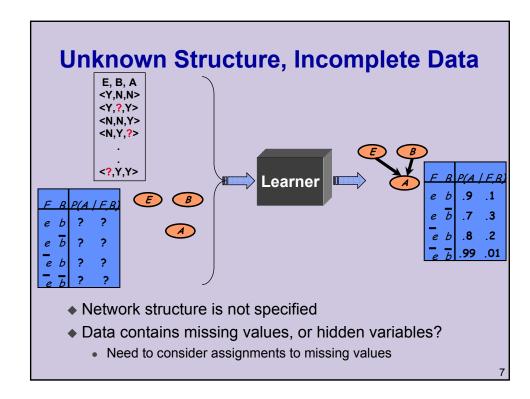
- ◆ Assume the domain X={X1, ..., Xn} is governed by some underlying distribution P\*(X)
- ♦  $P^*$  is induced by some Bayesian network  $B^*=(G,\Theta)$
- ◆ Given a data set D={x[1], ..., x[M]} of M samples from P\*
- Samples are i.i.d. independent and identically distributed
- ◆ The task is to recover the Bayesian network model











#### **Parameter Estimation**

- ◆ Network structure is specified
- Data set D consists of fully observed instances of the network variables

$$D=\{x[1], ..., x[M]\}$$

Estimate network parameters

#### **Parameter Estimation**

- Use a set D={x[1], ..., x[M]} of training samples drawn independently from a parametric model P(x : θ) to estimate the unknown parameter vector θ
- ◆ Parameter estimation: a classic problem in statistics
  - Maximum-Likelihood (ML) estimation
  - Bayesian estimation

9

#### **Maximum-Likelihood Estimation**

- ◆ IID data samples D={x[1], ..., x[M]}
- ◆ Likelihood function

$$L(\theta:D) = P(D \mid \theta) = \prod_{k=1}^{M} P(x[k]:\theta)$$

the likelihood of  $\theta$  w.r.t. the set of samples

- ML estimate of  $\theta$  is, by definition, the value  $\hat{\theta}$  that maximizes  $L(\theta:D)$
- it is usually easier to work with the the log-likelihood function

$$l(\theta:D) = \log P(D \mid \theta) = \sum_{k=1}^{M} \log P(x[k]:\theta)$$

# **Example: Discrete Case**

◆ Single binary variable

$$P(X=1) = \theta$$
,  $P(X=0) = 1 - \theta$ 

$$P(x \mid \theta) = \theta^{x} (1 - \theta)^{1 - x}$$

Bernoulli distribution

$$L(\theta:D) = \prod_{k=1}^{M} P(x[k]:\theta) = \theta^{N_1} (1-\theta)^{N_0}$$

◆ Sufficient statistics:

N<sub>1</sub>: number of 1's in D, N<sub>0</sub>: number of 0's in D

1.

- ♦ Log-likelihood  $I(\theta:D) = N_1 \ln \theta + N_0 \ln (1 - \theta)$
- ML estimation

$$\hat{\theta}_{ML} = \frac{N_1}{N_1 + N_0} = \frac{N_1}{M}$$

• Multi-valued discrete random variables {1, ..., K}  $\theta_i = P(X = i)$ 

$$P(x \mid \boldsymbol{\theta}) = \prod_{i=1}^{K} \theta_{i}^{\delta_{xi}}$$

$$L(\mathbf{\theta}:D) = \prod_{j=1}^{M} P(x[j]:\theta) = \prod_{i=1}^{K} \theta_{i}^{N_{i}}$$

• Sufficient statistics N<sub>i</sub>: the # of times i appears in D

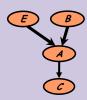
$$\hat{\theta}_{iML} = \frac{N_i}{\sum_{i} N_j} = \frac{N_i}{M}$$

13

## **MLE for Bayesian Networks**

◆ Training data has the form:

$$D = \begin{bmatrix} E[1] & B[1] & A[1] & C[1] \\ \vdots & \vdots & \vdots \\ E[M] & B[M] & A[M] & C[M] \end{bmatrix}$$



#### **Likelihood Function**

◆ By definition of network, we get



$$\mathcal{L}(\Theta:D) = \prod_{m} P(E[m], B[m], A[m], C[m]:\Theta)$$

$$= \prod_{m} \begin{pmatrix} P(E[m]:\Theta) \\ P(B[m]:\Theta) \\ P(A[m] \mid B[m], E[m]:\Theta) \\ P(C[m] \mid A[m]:\Theta) \end{pmatrix}$$

15

#### **Likelihood Function**

◆ Rewriting terms, we get



$$\mathcal{L}(\Theta: \mathcal{D}) = \prod_{m} P(E[m], \mathcal{B}[m], \mathcal{A}[m], \mathcal{C}[m]: \Theta)$$

$$\prod_{m} P(E[m]: \Theta)$$

$$= \prod_{m} P(\mathcal{B}[m]: \Theta)$$

$$= \prod_{m} P(\mathcal{A}[m] \mid \mathcal{B}[m], E[m]: \Theta)$$

$$\prod_{m} P(\mathcal{C}[m] \mid \mathcal{A}[m]: \Theta)$$

#### **Likelihood Function**

◆ Rewriting terms, we get



$$L(\Theta:D) = \prod_{m} P(E[m], B[m], A[m], C[m]: \Theta)$$

$$\prod_{m} P(E[m]: \Theta_{E})$$

$$= \prod_{m} P(B[m]: \Theta_{B})$$

$$= \prod_{m} P(A[m] | B[m], E[m]: \Theta_{A|BE})$$

$$\prod_{m} P(C[m] | A[m]: \Theta_{C|A})$$

17

# **General Bayesian Networks**

Generalizing for any Bayesian network:

$$L(\Theta:D) = \prod_{m} P(x_1[m], ..., x_n[m]: \Theta)$$

$$= \prod_{i} \prod_{m} P(x_i[m] | Pa_i[m]: \Theta_i)$$

$$= \prod_{i} L_i(\Theta_i:D) \longrightarrow \text{local likelihood function}$$

Global Decomposition of the likelihood function

⇒ Independent estimation problems

#### **MLE**

 Assuming discrete variables (CPTs) leads to further decomposition → local decomposition of the likelihood function

$$L_{i}(\Theta_{i}:D) = \prod_{m} P(x_{i}[m] | Pa_{i}[m]: \Theta_{i})$$

$$= \prod_{pa_{i}} \prod_{m,Pa_{i}[m]=pa_{i}} P(x_{i}[m] | pa_{i}: \Theta_{X_{i}|pa_{i}})$$

$$= \prod_{pa_{i}} \prod_{x_{i}} \theta_{x_{i}|pa_{i}}^{N(x_{i},pa_{i})}$$

$$\hat{\theta}_{x_{i}|pa_{i}} = \frac{N(x_{i}, pa_{i})}{N(pa_{i})}$$

19

#### **Bayesian Inference**

- Represent uncertainty about parameters using a probability distribution over parameters
- Learning using Bayes rule

$$P(\theta \mid x[1], ... x[M]) = \frac{P(x[1], ... x[M] \mid \theta) P(\theta)}{P(x[1], ... x[M])}$$
Posterior

Probability of data

## **Example: Discrete Variable**

Single binary variable

$$P(X=1) = \theta$$
,  $P(X=0) = 1 - \theta$ 

$$L(\theta:D) = \theta^{N_1} (1-\theta)^{N_0}$$

• What prior  $p(\theta)$  to use?

21

#### **Beta distribution**

Beta 
$$(\theta \mid \alpha_1, \alpha_0) = \frac{\Gamma(\alpha_1 + \alpha_0)}{\Gamma(\alpha_1)\Gamma(\alpha_0)} \theta^{\alpha_1 - 1} (1 - \theta)^{\alpha_0 - 1}$$

$$0 \le \theta \le 1$$

$$E(\theta) = \frac{\alpha_1}{\alpha_1 + \alpha_0}$$

- The parameters α<sub>1</sub> and α<sub>2</sub> are positive reals, often called hyperparameters
- Gamma Function

$$\Gamma(x+1) = x\Gamma(x)$$
  
 $\Gamma(1) = 1, \Gamma(x) = (x-1)!$  for interger  $x$ 

Assume the prior is a Beta distribution

$$p(\theta) = Beta (\theta | \alpha_1, \alpha_0) = c \theta^{\alpha_1 - 1} (1 - \theta)^{\alpha_0 - 1}$$

The posterior density  $p(\theta \mid D)$ 

$$p(\theta \mid \mathsf{D}) = c' \cdot p(\mathsf{D} \mid \theta) p(\theta)$$
$$= Beta(\theta \mid N_1 + \alpha_1, N_0 + \alpha_0)$$

- The property that the posterior distribution follows the same parametric form as the prior distribution is called conjugacy
- ◆ Beta prior is a conjugate family for the binomial distribution

$$P(X=1 \mid D) = \int P(X=1 \mid \theta) p(\theta \mid D) d\theta$$
$$= \int \theta p(\theta \mid D) d\theta = \frac{N_1 + \alpha_1}{N_1 + N_0 + \alpha_1 + \alpha_0} \equiv \hat{\theta}_{BE}$$

- It can be proved that:
   If the prior is well-behaved i.e.
  - If the prior is well-behaved i.e. does not assign 0 density to any feasible parameter value, then both MLE and Bayesian estimate converge to the same value in the limit
- ◆ Both almost surely converge to the underlying distribution P(X)
- But the ML and Bayesian approaches behave differently when the number of samples is small

◆ Multi-valued discrete random variables {1, ..., k}

$$\theta_{i} = P(X = i)$$

$$P(D \mid \theta) = \prod_{i=1}^{k} \theta_{i}^{N_{i}}$$

Sufficient statistics N<sub>i</sub>: the # of times i appears in D

• Assume the prior  $p(\theta)$  is a Dirichlet distribution  $Dir(\theta|\alpha)$ 

25

Dirichlet distribution with hyperparameters  $\alpha_i$ 's

Dir(
$$\boldsymbol{\theta} \mid \boldsymbol{\alpha}$$
) =  $\frac{\Gamma(\alpha)}{\prod_{i=1}^{k} \Gamma(\alpha_i)} \prod_{i=1}^{k} \theta_i^{\alpha_i - 1}$   
 $0 \le \theta_i \le 1$ ,  $\sum_{i=1}^{k} \theta_i = 1$ ,  $\alpha = \sum_{i=1}^{k} \alpha_i$   
 $E(\theta_i) = \frac{\alpha_i}{\alpha}$ 

- ◆ Multi-valued discrete random variables {1, ..., k}
- Assume the prior p(θ) is a Dirichlet distribution Dir(θ|α) with hyperparameters α<sub>i</sub>'s
- Then the posterior density p(θ |D) is also a Dirichlet distribution with hyperparameters α<sub>1</sub>+N<sub>1</sub>,...,α<sub>k</sub>+N<sub>k</sub>

$$p(\mathbf{\theta} \mid D) = c \cdot P(\mathbf{D} \mid \theta) p(\mathbf{\theta})$$
$$= Dir(\mathbf{\theta} \mid \mathbf{N} + \boldsymbol{\alpha})$$

 Dirichlet prior is a conjugate family for the multinomial distribution

27

◆ Bayesian estimates

$$P(X = i \mid D) = \int P(X = i \mid \mathbf{\theta}) p(\mathbf{\theta} \mid D) d\mathbf{\theta}$$
$$= \int \theta_i Dir(\mathbf{\theta} \mid \mathbf{N} + \boldsymbol{\alpha}) d\mathbf{\theta} = \frac{N_i + \alpha_i}{M + \alpha} \equiv \hat{\theta}_{iBE}$$

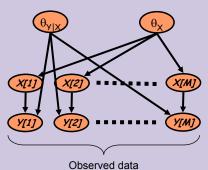
- $\bullet$  The hyperparameters  $\alpha_{\it i}$  can be thought of as "<code>imaginary</code>" counts from our prior experience
- α: imaginary equivalent sample size
- Let  $p_i$  be prior belief about  $\theta_i$ :  $\alpha_i = \alpha p_i$
- The larger the equivalent sample size, the more confident we are in our prior
- Laplace estimates:  $\alpha$ =k,  $\alpha$ <sub>i</sub>= 1

# **Summary of Bayesian estimation**

- ◆ Treat the unknown parameters as random variables
- ◆ Assume a prior distribution for the unknown parameters
- Update the distribution of the parameters based on data
- Finally compute p(x|D)

20

# **Bayesian Estimation in BNs**



- ◆ Meta-network for P(Θ,D)
- ◆ Priors for each parameter group are independent
- Data instances are independent given the unknown parameters

# **Bayesian Estimation in BNs**

◆ Global parameter independence assumption

$$P(\Theta) = \prod P(\Theta_{X_i|Pa_i})$$

Global Decomposition of the likelihood function

$$L(\Theta:D) = P(D \mid \Theta) = \prod L_i(\Theta_{X_i \mid Pa_i}:D)$$

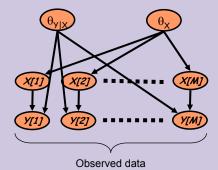
Posterior parameter independence

$$P(\Theta \mid D) = \prod P(\Theta_{X_i \mid Pa_i} \mid D)$$

We can solve the prediction problem for each CPD independently

31

## **Bayesian Estimation in BNs**

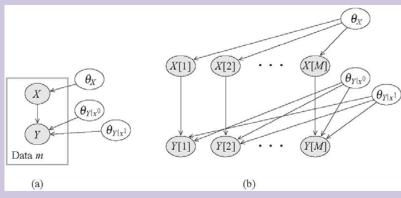


◆ This can be "read" from the network by d-separation
 Complete data ⇒

#### posteriors on parameters are independent

Can compute posterior over parameters separately!

# **Bayesian Estimation in BNs - CPTs**



◆ Local parameter independence assumption

$$P(\Theta_{X_i|Pa_i}) = \prod_{pa_i} P(\Theta_{X_i|pa_i})$$

33

# **Bayesian Nets & Bayesian Prediction**

Posterior parameter independence

$$P(\Theta \mid D) = \prod_{i} \prod_{pa_{i}} P(\Theta_{X_{i} \mid pa_{i}} \mid D)$$

Assume Dirichlet prior

$$P(\Theta_{X_i|pa_i}) = Dir(\Theta_{X_i|pa_i} \mid \alpha_{x_i|pa_i}, \ldots)$$

◆ Then

$$P(\Theta_{X_i|pa_i} \mid D) = Dir(\Theta_{X_i|pa_i} \mid \alpha_{x_i|pa_i} + N(x_i, pa_i), \ldots)$$

#### **Bayesian Nets & Bayesian Prediction**

◆ Bayesian estimation

$$\widetilde{\theta}_{x_i|pa_i} = P(X_i = x_i | Pa_i = pa_i, D)$$

$$= \frac{\alpha_{x_i|pa_i} + N(x_i, pa_i)}{\alpha_{pa_i} + N(pa_i)}$$

35

#### **Assessing Priors for Bayesian Nets**

The BDe prior

 Introduce an equivalent sample size α and a prior distribution P'(X), and set

$$\alpha_{x_i|pa_i} = \alpha P'(x_i, pa_i)$$

• We can represent P' as a BN  $(G_0, \Theta_0)$ , and set

$$\alpha_{x_i|pa_i} = \alpha P(x_i, pa_i^G \mid G_0, \Theta_0)$$

Use BN inference to compute this

◆ E.g., empty BN with uniform distribution

$$\alpha_{x_i|pa_i} = \frac{\alpha}{|X_i||Pa_i|}$$

# **Learning Parameters: Summary**

- ◆ Estimation relies on sufficient statistics
  - For multinomials: counts  $N(x_i, pa_i)$
  - Parameter estimation

$$\widehat{\boldsymbol{\theta}}_{x_i|pa_i} = \frac{\mathcal{N}(x_i, pa_i)}{\mathcal{N}(pa_i)} \qquad \qquad \widetilde{\boldsymbol{\theta}}_{x_i|pa_i} = \frac{\alpha_{x_i|pa_i} + N(x_i, pa_i)}{\alpha_{pa_i} + N(pa_i)}$$

$$\qquad \qquad \text{MLE} \qquad \qquad \text{Bayesian (Dirichlet)}$$

- ◆ Both are asymptotically equivalent and consistent
- ◆ Both can be implemented in an on-line manner by accumulating sufficient statistics