

**1. Text8.1****(a)**H0:  $\rho = 0$ ; Ha:  $\rho > 0$ 

Durbin Watson statistic is  $d=0.6208403$ . By checking table A.6,  $d_{low}=1.29$  and  $d_{up}=1.45$ . Since  $d < d_{low}$ , so we reject H0 and conclude that the  $\rho > 0$  and there is first-order correlation.

```
> reg<-lm(H~P,data=text8.1)
> library(car)
> dwt(reg)
lag Autocorrelation D-W Statistic p-value
1      0.6511468      0.6208403      0
Alternative hypothesis: rho != 0
```

**(b)**

H0: there is no autocorrelation ; Ha: there is positive autocorrelation

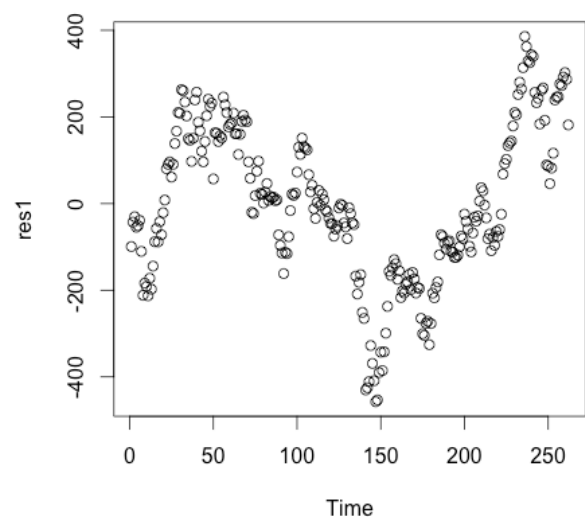
 $n1$  = # of positive residuals = 14;  $n2$  = # of negative residuals = 11 $E(R) = (2*n1*n2)/(n1+n2) + 1 = 13.32$ ; $VAR(R) = (2*n1*n2*(2*n1*n2-n1-n2))/((n1+n2-1)*(n1+n2)^2) = 5.8109$  $R$  = # of runs = 6 $Z$  statistics =  $(R-E(R))/\sqrt{var\_R} = -3.0366 \sim Z(0.05) = -1.645$ Since  $Z$  statistics  $< Z(0.05)$ , we reject H0 and conclude there is a positive autocorrelation.

```
> res = resid(reg)
> n1<-0
> n2<-0
> for (i in (1:25)) {if (res[i] > 0) {n1=n1+1} else {n2=n2+1}}
> E_R <- (2*n1*n2)/(n1+n2) + 1
> E_R
[1] 13.32
> var_R <- (2*n1*n2*(2*n1*n2-n1-n2))/((n1+n2-1)*(n1+n2)^2)
> var_R
[1] 5.810933
> R<-6
> Z<- (R-E_R)/sqrt(var_R)
> Z
[1] -3.036604
```

**2. Text8.4****(a)**

```
> reg1<-lm(DJIA~Time, data=text8.4)
> res1<-resid(reg1)
> plot(Time,res1)
```

From the plot of Residuals vs. Time, I can know the distribution of residues for time are not random and there is two clear cycles. So, the residuals for time are time dependent.



**(b)**

By considering lag =1, we use  $DJIA_{p-1}$  to predict  $DJIA_p$ . Since there is no such  $DJIA_0$  corresponding to  $DJIA_1$  and no  $DJIA_{p+1}$  corresponding to  $DJIA_p$ , So I create a new djia variable without  $DJIA_1$  and new djia\_lag variable without  $DJIA_p$ .

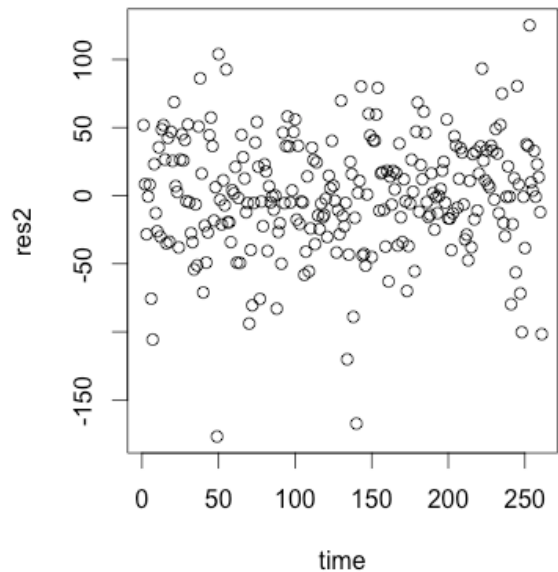
```
> djia<-text8.4$DJIA
> djia<-djia[-1]
> djia_lag<-text8.4$DJIA
> djia_lag<-djia_lag[-262]
> reg2<-lm(djia~djia_lag)
> summary(reg2)
```

```
Call:
lm(formula = djia ~ djia_lag)
```

```
Residuals:
    Min       1Q   Median       3Q      Max
-176.878  -22.397   -0.641   26.478  125.139
```

```
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 36.898384  42.989100   0.858   0.392
djia_lag     0.994459   0.007477 133.002 <2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 42.37 on 259 degrees of freedom
Multiple R-squared:  0.9856,    Adjusted R-squared:  0.9855
F-statistic: 1.769e+04 on 1 and 259 DF,  p-value: < 2.2e-16
```



I run the regression model and see  $DJIA_{p-1}$  is highly significant including in the model. Therefore, it is an adequate model. Then I check if there is any evidence of autocorrelation in residuals by plot residuals vs. time. The graph shows there is not an autocorrelation. And I did Durbin Watson test to check.

```
> dwt(reg2)
lag Autocorrelation D-W Statistic p-value
 1      0.1066715     1.758642   0.034
Alternative hypothesis: rho != 0
```

DW statistics = 1.758642 which is close to 2. So, there is no autocorrelation.

(c)

```
> log_djia <- log(djia)
> reg3<-lm(log_djia~djia_lag)
> summary(reg3)
```

Call:

```
lm(formula = log_djia ~ djia_lag)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.0304723	-0.0040038	0.0002664	0.0047480	0.0187148

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	7.678e+00	7.793e-03	985.2	<2e-16 ***
djia_lag	1.701e-04	1.355e-06	125.5	<2e-16 ***

---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

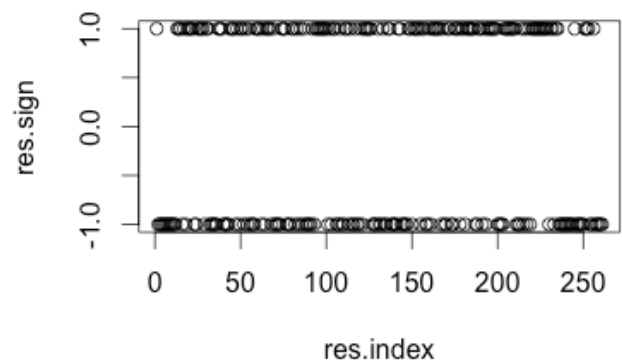
Residual standard error: 0.007681 on 259 degrees of freedom  
Multiple R-squared: 0.9838, Adjusted R-squared: 0.9838  
F-statistic: 1.576e+04 on 1 and 259 DF, p-value: < 2.2e-16

```
> library(car)
```

```
> dwt(reg3)
```

lag	Autocorrelation	D-W Statistic	p-value
1	0.1762719	1.619276	0.002

Alternative hypothesis: rho != 0



Used  $\log(djia)$  to run the regression and also see  $DJIA_{p-1}$  is highly significant including in the model. Then I check if there is any evidence of autocorrelation in residuals by plot above. The graph shows there is no autocorrelation. So, I did Durbin Watson test to check. DW statistics = 1.619276 which is close to 2. So, there is no autocorrelation.

### 3. Text8.5

(a)

The adequate model is  $\text{lm}(djia \sim djia\_lag)$  in Problem (b).

```
> reg4<-lm(djia~djia_lag, data = data_130)
> summary(reg4)
```

Call:

```
lm(formula = djia ~ djia_lag, data = data_130)
```

Residuals:

Min	1Q	Median	3Q	Max
-170.440	-23.477	0.806	25.819	104.042

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	239.96956	108.59402	2.21	0.0289 *
djia_lag	0.95732	0.01964	48.74	<2e-16 ***

---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 41.23 on 127 degrees of freedom  
(1 observation deleted due to missingness)

Multiple R-squared: 0.9492, Adjusted R-squared: 0.9488  
F-statistic: 2375 on 1 and 127 DF, p-value: < 2.2e-16

```
> res4<-resid(reg4)
```

```
> anova(reg4) #get residual mean square
```

Analysis of Variance Table

Response: djia

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
djia_lag	1	4037704	4037704	2375.2	< 2.2e-16 ***
Residuals	127	215891	1700		

The residual mean square is **1700**.

(b)

After predict first 15 days of July, we get predicts and also the predict error (the difference between predict and actual DJIA).

```
> data_15<-subset(text8.4,Time<=145)
> data_15<-subset(data_15,Time>=131)
> for (i in 1:15 ) {data_15$predict[i] = 239.96956 + 0.95732*data_15$djia_lag[i]}
> #c
> for (i in 1:15) {data_15$pre_error[i]<-data_15$djia[i]-data_15$predict[i]}
> SSE =0
> for (i in 1:15) {SSE<-SSE+(data_15$pre_error[i])^2}
> SSE
[1] 63856.89
> ave_SSE<-SSE/15
> ave_SSE
[1] 4257.126
```

	row.names	Date	djia	Time	djia_lag	predict	pre_error
1	131	7/1/96	5729.98	131	5654.63	5653.260	76.720048
2	132	7/2/96	5720.38	132	5729.98	5725.394	-5.014014
3	133	7/3/96	5703.02	133	5720.38	5716.204	-13.183742
4	134	7/4/96	5703.02	134	5703.02	5699.585	3.435334
5	135	7/5/96	5588.14	135	5703.02	5699.585	-111.444666
6	136	7/8/96	5550.83	136	5588.14	5589.608	-38.777745
7	137	7/9/96	5581.86	137	5550.83	5553.890	27.969864
8	138	7/10/96	5603.65	138	5581.86	5583.596	20.054225
9	139	7/11/96	5520.50	139	5603.65	5604.456	-83.955778
10	140	7/12/96	5510.56	140	5520.50	5524.855	-14.294620
11	141	7/15/96	5349.51	141	5510.56	5515.339	-165.828859
12	142	7/16/96	5358.76	142	5349.51	5361.162	-2.402473
13	143	7/17/96	5376.88	143	5358.76	5370.018	6.862317
14	144	7/18/96	5464.18	144	5376.88	5387.364	76.815678
15	145	7/19/96	5426.82	145	5464.18	5470.938	-44.118358

(c)

From the code in (b), we get the SSE = **63856.89**. Average of squared predict error =  $SSE/(n) = 4257.126$ . Residual mean square = 1700 is much less than Average of squared predict error.

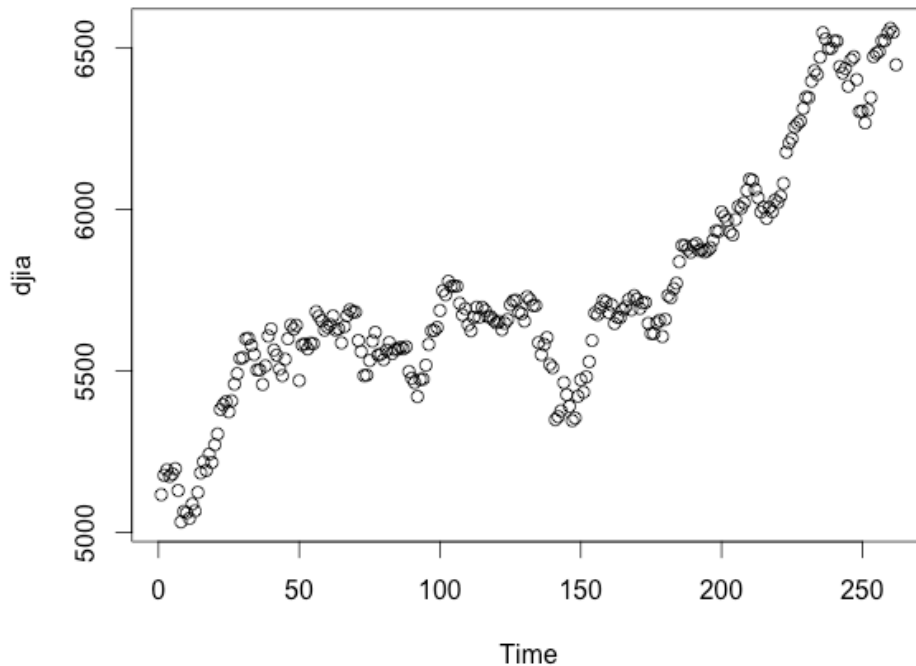
(d)

```
> data_132<-subset(text8.4,Time>130)
> for (i in 1:132 ) {data_132$predict[i] = 239.96956 + 0.95732*data_132$djia_lag[i]}
> for (i in 1:132) {data_132$pre_error[i]<-data_132$djia[i]-data_132$predict[i]}
> SSE =0
> for (i in 1:132) {SSE<-SSE+(data_132$pre_error[i])^2}
> SSE
[1] 319809.7
> ave_SSE<-SSE/132
> ave_SSE
[1] 2422.801
```

After predict second half of the year, we get predicts and also the predict error (the difference between predict and actual DJIA).

We get the SSE = **319809.7**. Average of squared predict error =  $SSE/(n) = 2422.801$ .

Residual mean square =1700 is still much less than Average of squared predict error.  
(e)



The scatter plot of Time vs. DJIA shows there is a significant drop of DJIA with first 15 days of July and the increasing trend is very smooth for DJIA in second half of the year. This is the reason the Average of squared predict error of DJIA within first 15 days of July **is much larger** than Average of squared predict error of DJIA within second half of the year.

```
Call:
lm(formula = Y ~ X1 + X2 + X3 + X4 + X5 + X6 + X7 + X8 + X9,
    data = text11.5)
```

```
Residuals:
    Min       1Q   Median       3Q      Max
-3.7729 -1.9801 -0.0868  1.6615  4.2618
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	15.31044	5.96093	2.568	0.0223 *
X1	1.95413	1.03833	1.882	0.0808 .
X2	6.84552	4.33529	1.579	0.1367
X3	0.13761	0.49436	0.278	0.7848
X4	2.78143	4.39482	0.633	0.5370
X5	2.05076	1.38457	1.481	0.1607
X6	-0.55590	2.39791	-0.232	0.8200
X7	-1.24516	3.42293	-0.364	0.7215
X8	-0.03800	0.06726	-0.565	0.5810
X9	1.70446	1.95317	0.873	0.3976

```
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 2.973 on 14 degrees of freedom
Multiple R-squared:  0.8512,    Adjusted R-squared:  0.7555
F-statistic: 10.88 on 9 and 14 DF, p-value: 0.000015
```

#### 4. Text11.5

(a)

No. Since all predictive variables, except X1, are not significant adding the model. Also, adjust  $R^2 = 75.55\%$  is not good enough.

**(b)**

```
Call:
lm(formula = Y ~ X1 + X6 + X8, data = text11.5)

Residuals:
    Min       1Q   Median       3Q      Max
-3.7486 -2.4082 -0.3594  2.1378  6.5353

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 14.796013   4.971105   2.976 0.007462 **
X1           3.489464   0.729368   4.784 0.000113 ***
X6          -0.415515   1.182262  -0.351 0.728921
X8           0.004923   0.063597   0.077 0.939062
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.123 on 20 degrees of freedom
Multiple R-squared:  0.7655,    Adjusted R-squared:  0.7303
F-statistic: 21.76 on 3 and 20 DF,  p-value: 1.653e-06
```

No. Since X6 (Number of rooms) and X8 (Age of the home) are highly insignificant in the model since their P-values are very large; and adjusted  $R^2$  is low. So the model does not adequately describe the sale price.

**(c)**

No. After tried several model, I found a model works better than just adding Taxes as a predictive variable. Left side is the better model; all variables are significant including in the model with adjusted  $R^2$  80.23%. Right side is the model with only Taxes as predictive variable, which has a lower adjusted  $R^2$  75.3%. So, the new model works better.

```
> reg2<-lm(Y~X1+X2+X5+X2:X7, data = text11.5)
> summary(reg2)
```

```
Call:
lm(formula = Y ~ X1 + X2 + X5 + X2:X7, data = text11.5)

Residuals:
    Min       1Q   Median       3Q      Max
-4.2039 -2.4101 -0.0615  2.0829  3.8468

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  6.0241    3.4729   1.735 0.099007 .
X1           2.4923    0.5125   4.863 0.000108 ***
X2          14.9721    5.4718   2.736 0.013121 *
X5           2.0956    1.1930   1.757 0.095089 .
X2:X7        -2.0201    1.0706  -1.887 0.074562 .
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.674 on 19 degrees of freedom
Multiple R-squared:  0.8367,    Adjusted R-squared:  0.8023
F-statistic: 24.33 on 4 and 19 DF,  p-value: 2.993e-07
```

```
Call:
lm(formula = Y ~ X1)

Residuals:
    Min       1Q   Median       3Q      Max
-3.8445 -2.3340 -0.3841  1.9689  6.3005

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 13.3553    2.5955   5.146 3.71e-05 ***
X1           3.3215    0.3939   8.433 2.44e-08 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.988 on 22 degrees of freedom
Multiple R-squared:  0.7637,    Adjusted R-squared:  0.753
F-statistic: 71.11 on 1 and 22 DF,  p-value: 2.435e-08
```

**5. Text11.6****(a)**

No. As the result show of the regression model including all predictive variable, only X5 is significant in the model. Adjust  $R^2$  is not good. So, I would not include all the variables to predict the gasoline consumption of the cars.

```
Call:
lm(formula = Y ~ X1 + X2 + X3 + X4 + X5 + X6 + X7 + X8 + X9 +
    X10 + factor(X11), data = text9.3)
```

```
Residuals:
    Min       1Q   Median       3Q      Max
-5.3498 -1.6236 -0.6002  1.5155  5.2815
```

```
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  17.773204   30.508775    0.583   0.5674
X1           -0.077946    0.058607   -1.330   0.2001
X2           -0.073399    0.088924   -0.825   0.4199
X3            0.121115    0.091353    1.326   0.2015
X4            1.329034    3.099535    0.429   0.6732
X5            5.975989    3.158647    1.892   0.0747 .
X6            0.304178    1.289094    0.236   0.8161
X7           -3.198576    3.105435   -1.030   0.3167
X8            0.185362    0.129252    1.434   0.1687
X9           -0.399146    0.323812   -1.233   0.2336
X10          -0.005193    0.005893   -0.881   0.3898
factor(X11)1  0.598655    3.020681    0.198   0.8451
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 3.226 on 18 degrees of freedom
Multiple R-squared:  0.8353,    Adjusted R-squared:  0.7346
F-statistic: 8.297 on 11 and 18 DF,  p-value: 5.287e-05
```

**(b)**

Among these regression models, I would choose regression of Y on X8, X5 and X10. All the three variables are significant including in the model and the model has highest adjusted  $R^2$  78.08% among these given models.

```
Call:
lm(formula = Y ~ X8 + X5 + X10)
```

```
Residuals:
    Min       1Q   Median       3Q      Max
-4.5930 -1.9674 -0.6438  2.0314  5.8823
```

```
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  4.494972   11.764747    0.382   0.7055
X8            0.218119    0.087764    2.485   0.0197 *
X5            2.607338    1.263791    2.063   0.0492 *
X10          -0.009482    0.001993   -4.759 6.35e-05 ***
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 2.932 on 26 degrees of freedom
Multiple R-squared:  0.8035,    Adjusted R-squared:  0.7808
F-statistic: 35.43 on 3 and 26 DF,  p-value: 2.47e-09
```



I found a better model, which shows below. Two interaction terms added in the model and all variable s are significant in the model. The adjust  $R^2$  88.15% is higher.

Call:

```
lm(formula = Y ~ X8 + X5 + X10 + X1:X3 + X1:X7)
```

Residuals:

	Min	1Q	Median	3Q	Max
Residuals	-5.4577	-0.9720	0.1114	1.3833	3.3444

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	2.124e+01	1.007e+01	2.110	0.045487 *
X8	2.104e-01	7.673e-02	2.742	0.011357 *
X5	2.091e+00	9.465e-01	2.209	0.036942 *
X10	-8.925e-03	2.762e-03	-3.232	0.003554 **
X1:X3	1.982e-04	4.123e-05	4.808	6.76e-05 ***
X1:X7	-3.337e-02	7.741e-03	-4.311	0.000239 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

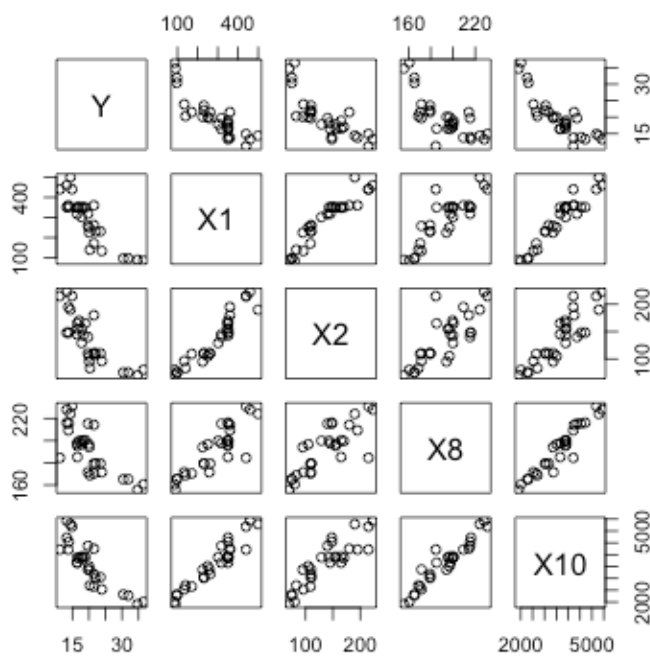
Residual standard error: 2.156 on 24 degrees of freedom

Multiple R-squared: 0.9019, Adjusted R-squared: 0.8815

F-statistic: 44.15 on 5 and 24 DF, p-value: 2.486e-11

(c)

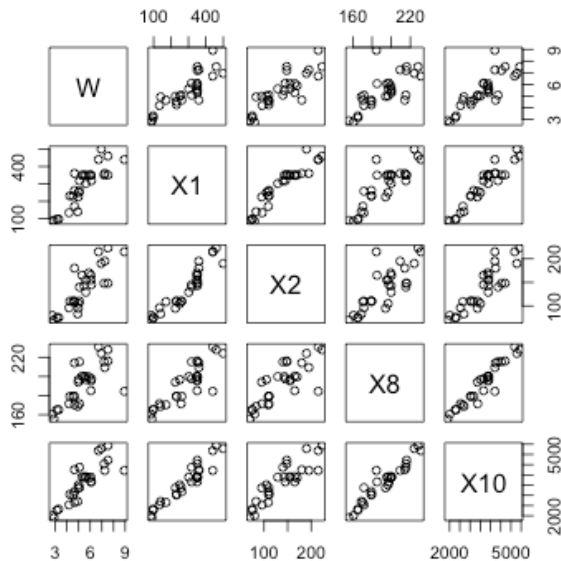
Yes. From the plot of Y vs. X1, X2, X8, X10 (first row of scatter matrix), we can see there is a exponential trend in each plot. So, Y vs X1, Y vs. X2, Y vs. X8 and Y vs. X10 have a exponential relationship, not a linear relationship.





(d)

Yes. From the plot of W vs. X1, X2, X8, X10 (first row of scatter matrix), we can see there is a linear trend in each plot. So, the relationship between W and the 11 predictor variables is more linear than that between Y and the 11 predictor variables.



(e)

After replacing Y by W in these regression model, X2 becomes significant in the model  $W \sim X2 + X10$ . But, X5 becomes insignificant in the model  $W \sim X5 + X8 + X10$ . So, the relationship between W and the 11 predictor variable is more linear, but **not** a linear relationship of W vs. each variables, like W vs. X5.

```
lm(formula = W ~ X2 + X10)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-1.82734	-0.39360	-0.05088	0.31576	1.95887

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	0.8397046	0.5651506	1.486	0.1489
X2	0.0147369	0.0067923	2.170	0.0390 *
X10	0.0007026	0.0003223	2.180	0.0381 *

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.7514 on 27 degrees of freedom

Multiple R-squared: 0.7576, Adjusted R-squared: 0.7396

F-statistic: 42.19 on 2 and 27 DF, p-value: 4.919e-09

Call:

```
lm(formula = Y ~ X2 + X10)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-5.389	-2.155	-1.266	3.044	7.575

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	40.203322	2.516464	15.976	2.77e-15 ***
X2	-0.026227	0.030244	-0.867	0.39349
X10	-0.004569	0.001435	-3.184	0.00364 **

---

```
Call:
lm(formula = W ~ X8 + X5 + X10)

Residuals:
    Min       1Q   Median       3Q      Max
-1.1439 -0.3867 -0.0915  0.4313  1.3642

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  8.8185512   2.6878166   3.281 0.002946 **
X8          -0.0743329   0.0200508  -3.707 0.000998 ***
X5           0.0778394   0.2887303   0.270 0.789602
X10          0.0029363   0.0004552   6.450 7.79e-07 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.6699 on 26 degrees of freedom
Multiple R-squared:  0.8144,    Adjusted R-squared:  0.793
F-statistic: 38.04 on 3 and 26 DF,  p-value: 1.178e-09
```

```
Call:
lm(formula = Y ~ X8 + X5 + X10)

Residuals:
    Min       1Q   Median       3Q      Max
-4.5930 -1.9674 -0.6438  2.0314  5.8823

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  4.494972   11.764747   0.382  0.7055
X8           0.218119   0.087764   2.485  0.0197 *
X5           2.607338   1.263791   2.063  0.0492 *
X10          -0.009482   0.001993  -4.759 6.35e-05 ***
```

So, among given regression by the question, I think  $\text{lm}(W \sim X8 + W10)$  is the model. Since all predict variables are significant and adjust  $R^2$  is 80.01%, which is the highest.

```
lm(formula = W ~ X8 + X10)
```

```
Residuals:
    Min       1Q   Median       3Q      Max
-1.1512 -0.3578 -0.1174  0.4314  1.3869

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  9.1675393   2.3147462   3.960 0.000491 ***
X8          -0.0744971   0.0196944  -3.783 0.000784 ***
X10          0.0029144   0.0004402   6.621 4.19e-07 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.6583 on 27 degrees of freedom
Multiple R-squared:  0.8139,    Adjusted R-squared:  0.8001
F-statistic: 59.05 on 2 and 27 DF,  p-value: 1.384e-10
```

After tried many different models, I think  $\text{lm}(W \sim X8 + W10)$  is the best among them. I didn't find a better model than  $\text{lm}(W \sim X8 + W10)$ .

(f)

Regressed Y on X13, we can see X3 is highly significant in the model and with very high adjusted  $R^2$  88.82%. So we definitely should include the X13 into the final model to predict the gasoline consumption of the cars.

```
Call:
lm(formula = Y ~ X13)

Residuals:
    Min       1Q   Median       3Q      Max
-3.7134 -1.2457 -0.0227  1.4211  4.3458

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  -11.37      2.10   -5.415 8.93e-06 ***
X13           566.00     37.21  15.213 4.59e-15 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.094 on 28 degrees of freedom
Multiple R-squared:  0.8921,    Adjusted R-squared:  0.8882
F-statistic: 231.4 on 1 and 28 DF,  p-value: 4.587e-15
```