

**Textbook**

3.14  $H_0: \beta(\text{gender}) = \beta(\text{experience}) = \beta(\text{months}) = 0$

$H_a$ : at least one of them is not zero

$r = 3$ . (test 3 parameters in  $H_0$ ).

$MSH_0 = SSH_0 / r = (SSE_0 - SSE) / r = (38460756 - 22657938) / 3 = 5267606$ .  $MSE = 257477$

And we have  $F = MSH_0 / MSE = 20.45855 \sim F(0.95, 3, 88) = 2.708186$ .

Since F-statistics is larger than  $F(0.95, 3, 88)$ , we reject the  $H_0$  and conclude at least one of gender, beta and months is indeed needed to add into the model.

5.6 (a)  $\text{cor}(Y, \text{Horsepower}) = 0.8611619$ . (Y: price)

(b) In the model 3 with interaction terms, all variables except Horsepower becomes insignificant (P-value larger than 0.05). So we should not use model 3 to predict Y. In model 2, Germany variable is highly insignificant since P-value is larger than 0.05. So, we need to exclude Germany first.

After exclude Germany from the category of Country (by substituting "Germany" with "Others") run the regression and get:

```
lm(formula = Y ~ Horsepower + NewCountry, data = Text5.6)
```

Residuals:

| Min      | 1Q      | Median  | 3Q     | Max     |
|----------|---------|---------|--------|---------|
| -11.1073 | -2.3465 | -0.1521 | 2.1123 | 12.4388 |

Coefficients:

|                 | Estimate  | Std. Error | t value | Pr(> t )    |
|-----------------|-----------|------------|---------|-------------|
| (Intercept)     | -4.244450 | 1.461595   | -2.904  | 0.00479 **  |
| Horsepower      | 0.174976  | 0.009769   | 17.911  | < 2e-16 *** |
| NewCountryUSA   | -3.195623 | 1.238057   | -2.581  | 0.01172 *   |
| NewCountryJapan | -3.856653 | 1.243542   | -3.101  | 0.00268 **  |

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.703 on 78 degrees of freedom

(8 observations deleted due to missingness)

Multiple R-squared: 0.8056, Adjusted R-squared: 0.7981

F-statistic: 107.8 on 3 and 78 DF, p-value: < 2.2e-16

The new model 2 is:  $\hat{Y} = -4.24445 + 0.174976\text{Horsepower} - 3.195623\text{USA} - 3.856653\text{Japan}$   
 Given  $\text{USA} = 1$ ,  $\text{Japan} = 0$ ,  $\text{Germany} = 0$  and  $\text{Horsepower} = 100$ , LS estimated price = 10.05753 thousands.

(c) Let Horsepower equal to a fixed number  $x$ ,  $x \geq 0$ .

The coefficient of Japan is smallest in the new model 2, So, we can get  $\hat{Y}(\text{Japan})$  is the smallest and Japan car is least expensive car by holding the horsepower fixed.

(d)  $H_0: \beta(\text{HP*USA}) = 0$ ;  $H_a: \beta(\text{HP*USA}) \neq 0$

From output of model 3, we can see p-value of HP\*USA is 0.2204, which is larger than 0.05, so we fail to reject  $H_0$  and conclude there is not an interaction between horsepower and USA.

$H_0: \beta(\text{HP*Japan}) = 0$ ;  $H_a: \beta(\text{HP*Japan}) \neq 0$

From output of model 3, we can see p-value of HP\*Japan is 0.0631, which is larger than 0.05, so we fail to reject  $H_0$  and conclude there is not an interaction between horsepower and Japan.

$H_0: \beta(\text{HP*Germany}) = 0$ ;  $H_a: \beta(\text{HP*Germany}) \neq 0$

From output of model 3, we can see p-value of HP\*Germany is 0.1560, which is larger than 0.05, so we fail to reject  $H_0$  and conclude there is not an interaction between horsepower and Germany.

Overall, we can conclude there is an interaction between Country and Horsepower.

(e) Because (d) shows there is not interaction between Country and Horsepower, so all interaction variables can be excluded. Since Germany is not significant in Model 2, so the new Model 2 is good to be used here.

$H_0: \beta(\text{USA}) = 0$ ;  $H_a: \beta(\text{USA}) \neq 0$

We can see p-value of USA is 0.01172, which is smaller than 0.05, so USA is significant in the model.

$H_0: \beta(\text{Japan}) = 0$ ;  $H_a: \beta(\text{Japan}) \neq 0$

We can see p-value of Japan is 0.00268, which is smaller than 0.05, so Japan is significant in the model.

USA and Japan are significant and indeed needed included in the model. Therefore, we can conclude Country is an important predictor of the price of a car.

(f) Since Germany is not significant in Model 2 and there is no need include Germany in the model, we can reduce Germany and let it joins with Other category together.

(g)

$H_0: \beta_{\text{USA}} = \beta_{\text{Japan}}$  (also can be  $\beta_{\text{USA}} - \beta_{\text{Japan}} = 0$ );  $H_a: \beta_{\text{USA}} \neq \beta_{\text{Japan}}$

T-statistics:  $T = [(\beta_{\text{USA}} - \beta_{\text{Japan}}) - 0] / \text{SE}(\beta_{\text{USA}} - \beta_{\text{Japan}}) \sim t(n-2, 0.025) = t(88, 0.025)$

And  $\text{SE}(\beta_{\text{USA}} - \beta_{\text{Japan}}) = \sqrt{\text{Var}(\beta_{\text{USA}}) + \text{Var}(\beta_{\text{Japan}}) - 2 \cdot \text{Cov}(\beta_{\text{USA}}, \beta_{\text{Japan}})}$

5.9 After change the baseline to be "0" for D and set a new variable  $\mathbf{Z} = \mathbf{G} \cdot \mathbf{I}$ , run the regression:

|               |
|---------------|
| Coefficients: |
|---------------|

|             | Estimate   | Std. Error | t value | Pr(> t )     |
|-------------|------------|------------|---------|--------------|
| (Intercept) | 0.5111627  | 0.0321992  | 15.875  | 2.40e-10 *** |
| I           | -0.0201077 | 0.0168979  | -1.190  | 0.2539       |
| D           | 0.0546159  | 0.0205705  | 2.655   | 0.0188 *     |
| W           | 0.0133905  | 0.0422639  | 0.317   | 0.7560       |
| Z           | 0.0096901  | 0.0017712  | 5.471   | 8.24e-05 *** |
| P           | -0.0007224 | 0.0040046  | -0.180  | 0.8594       |
| N           | -0.0051822 | 0.0038083  | -1.361  | 0.1951       |

$$(a) V^{\wedge} = 0.5111627 - 0.0201077 I + 0.0546159 D + 0.0133905 W + 0.0096901 G * I - 0.0007224 * P - 0.0051822 * N$$

The coefficient of D is 0.0546159. If a Democratic incumbent is running for election,  $D=1$ , the Democratic share of the two-party presidential vote will increase 0.0633485 units. If a Republic incumbent is running for election,  $D=-1$ , the Democratic share of the two-party presidential vote will decrease 0.0633485 units. If Other incumbent is running for election,  $D=0$ , the Democratic share of the two-party presidential vote will not change.

$$V^{\wedge} = 0.5657786 + -0.0201077 I + 0.0133905 W + 0.0096901 G * I - 0.0007224 * P - 0.0051822 * N \text{ when } D=1$$

$$V^{\wedge} = 0.4565468 + -0.0201077 I + 0.0133905 W + 0.0096901 G * I - 0.0007224 * P - 0.0051822 * N \text{ when } D=-1$$

$$V^{\wedge} = 0.5111627 + -0.0201077 I + 0.0133905 W + 0.0096901 G * I - 0.0007224 * P - 0.0051822 * N \text{ when } D=0$$

(b) We don't need to keep I in the above model since the p-value of I is 0.2539, which is larger than 0.05. I is not significant and should be excluded from the model.

(c) Yes.  $G * I$ 's corresponding P-value is 8.24e-05, which is highly significant.  $G * I$  should be included in the model.

(d) After run many different models, the best model I found is below:

| Coefficients: |           |            |         |              |
|---------------|-----------|------------|---------|--------------|
|               | Estimate  | Std. Error | t value | Pr(> t )     |
| (Intercept)   | 0.481837  | 0.008522   | 56.543  | < 2e-16 ***  |
| D             | 0.050000  | 0.014198   | 3.522   | 0.002619 **  |
| $G * I$       | 0.007626  | 0.001603   | 4.757   | 0.000183 *** |
| $I * P$       | -0.003705 | 0.001994   | -1.858  | 0.080534 .   |

$$V^{\wedge} = 0.481837 + 0.050000 D + 0.007626 G * I - 0.003705 I * P$$

I tried all possible interaction and found out I\*P is a significant interaction added into the model. In the best model, all variables are significant under 90% CI test.

5.10 (a)

```
lm(formula = V ~ factor(I) + D1 + D2 + factor(W) + Z + P + N)
```

Residuals:

| Min       | 1Q        | Median    | 3Q       | Max      |
|-----------|-----------|-----------|----------|----------|
| -0.044201 | -0.022728 | -0.002548 | 0.011671 | 0.084681 |

Coefficients:

|             | Estimate   | Std. Error | t value | Pr(> t )     |
|-------------|------------|------------|---------|--------------|
| (Intercept) | 0.5260743  | 0.0369736  | 14.228  | 2.64e-09 *** |
| factor(I)1  | -0.0411965 | 0.0349716  | -1.178  | 0.259912     |
| D1          | 0.0633485  | 0.0312177  | 2.029   | 0.063423 .   |
| D2          | -0.0469714 | 0.0291912  | -1.609  | 0.131600     |
| factor(W)1  | 0.0123948  | 0.0436938  | 0.284   | 0.781127     |
| G*I         | 0.0094222  | 0.0019580  | 4.812   | 0.000339 *** |
| P           | -0.0006963 | 0.0041333  | -0.168  | 0.868808     |
| N           | -0.0051083 | 0.0039349  | -1.298  | 0.216773     |

$V^{\wedge} = 0.5260743 - 0.0411965I + 0.0633485D1 - 0.0469714D2 + 0.0123948W + 0.0094222G*I - 0.0006963P - 0.0051083N$ .

A Democratic incumbent is running for election, the Democratic share of the two-party presidential vote will estimate increase 0.0633485 units.

A Republic incumbent is running for election, the Democratic share of the two-party presidential vote will estimate decrease 0.0469714 units.

(b) when  $\alpha_1 = -\alpha_2$ ,  $V = \beta_0 + \beta_1 I + \alpha_1 (D1 - D2) + \beta_3 W + \beta_4 (G*I) + \beta_5 P + \beta_6 N + \text{error}$ . 5.11 is a special case of 5.12 since  $D2=0$  when  $D1=1$ , which is category "1" above.  $D1=0$  when  $D2=1$ , which is category "-1" above.  $D1=0, D2=0$  is category "0" above. So we can consider  $(D1-D2)$  being a variable with 3 categories, which is variable D from 5.9.

(c) when  $\alpha_1 = -\alpha_2$ , we get reduce model under  $H_0$ :  $V = \beta_0 + \beta_1 I + \alpha_1 (D1 - D2) + \beta_3 W + \beta_4 (G*I) + \beta_5 P + \beta_6 N + \text{error}$ , which is the model from 5.9.  
Full model from 5.10.

$H_0$ : Reduced model = full model ;  $H_a$  : reduced model  $\neq$  full model

Run the ANOVA

Model 1:  $V \sim I + D + W + G:I + P + N$

Model 2:  $V \sim I + D1 + D2 + W + G:I + P + N$

|   | Res.Df | RSS      | Df | Sum of Sq  | F      | Pr(>F) |
|---|--------|----------|----|------------|--------|--------|
| 1 | 14     | 0.023686 |    |            |        |        |
| 2 | 13     | 0.023423 | 1  | 0.00026227 | 0.1456 | 0.709  |

P-value = 0.709 larger than 0.05, so we fail to reject  $H_0$  and conclude reduced model = full model, so  $\alpha_1 = -\alpha_2$ .

### Prof's book

3.4 (a) Yes. The fitted equation is checked.

```
lm(formula = log.Salary. ~ YrsEm + PriorYr + Education + Super +
  factor(Female) + factor(Advertising) + factor(Engineering) +
  factor(Sales))
```

Residuals:

| Min       | 1Q        | Median    | 3Q       | Max      |
|-----------|-----------|-----------|----------|----------|
| -0.089659 | -0.024036 | -0.004498 | 0.028587 | 0.089410 |

Coefficients:

|                      | Estimate   | Std. Error | t value | Pr(> t )     |
|----------------------|------------|------------|---------|--------------|
| (Intercept)          | 4.4287934  | 0.0213399  | 207.535 | < 2e-16 ***  |
| YrsEm                | 0.0074788  | 0.0011931  | 6.269   | 2.72e-07 *** |
| PriorYr              | 0.0016839  | 0.0019568  | 0.861   | 0.395039     |
| Education            | 0.0170345  | 0.0033360  | 5.106   | 1.02e-05 *** |
| Super                | 0.0003901  | 0.0008056  | 0.484   | 0.631115     |
| factor(Female)1      | 0.0230683  | 0.0142917  | 1.614   | 0.115002     |
| factor(Advertising)1 | -0.0387774 | 0.0249146  | -1.556  | 0.128124     |
| factor(Engineering)1 | -0.0057292 | 0.0197703  | -0.290  | 0.773597     |
| factor(Sales)1       | -0.0937783 | 0.0225745  | -4.154  | 0.000185 *** |

Residual standard error: 0.04586 on 37 degrees of freedom

Multiple R-squared: 0.8634, Adjusted R-squared: 0.8338

F-statistic: 29.22 on 8 and 37 DF, p-value: 9.629e-14

(b)

Coefficients:

|                      | Estimate   | Std. Error | t value | Pr(> t )     |
|----------------------|------------|------------|---------|--------------|
| (Intercept)          | 4.3580834  | 0.0248414  | 175.436 | < 2e-16 ***  |
| YrsEm                | 0.0074788  | 0.0011931  | 6.269   | 2.72e-07 *** |
| PriorYr              | 0.0016839  | 0.0019568  | 0.861   | 0.395039     |
| Education            | 0.0170345  | 0.0033360  | 5.106   | 1.02e-05 *** |
| Super                | 0.0003901  | 0.0008056  | 0.484   | 0.631115     |
| factor(Male)1        | -0.0230683 | 0.0142917  | -1.614  | 0.115002     |
| factor(Advertising)1 | 0.0550009  | 0.0230111  | 2.390   | 0.022045 *   |
| factor(Engineering)1 | 0.0880491  | 0.0180562  | 4.876   | 2.07e-05 *** |
| factor(Marketing)1   | 0.0937783  | 0.0225745  | 4.154   | 0.000185 *** |

If we set Female and Sales to be baseline, the new regression is above. So the new coefficient for Male is -0.0230683, for Advertising is 0.0550009, for Engineering is 0.0550009, and for Marketing is 0.0550009.

(c) The coefficients for the other categories represent the differences from that group. Engg is nonsignificant with base categories Male&Marketing, so the influence on log(salary) is not much difference from Male and Marketing's influence on log(salary), so, here, Engg is not significant. However, Engg is significant with base categories Female&Sales, so the influence on log(salary) is much difference from Female and Sale's influence on log(salary), so, here, Engg is significant.

(d) after drop those three variables, run the regression:

Coefficients:

|             | Estimate  | Std. Error | t value | Pr(> t )     |
|-------------|-----------|------------|---------|--------------|
| (Intercept) | 4.439005  | 0.019804   | 224.142 | < 2e-16 ***  |
| YrsEm       | 0.007660  | 0.001208   | 6.341   | 1.57e-07 *** |
| Education   | 0.018371  | 0.003124   | 5.881   | 6.95e-07 *** |
| Advertising | -0.036488 | 0.025311   | -1.442  | 0.157208     |
| Engineering | -0.002507 | 0.020037   | -0.125  | 0.901046     |
| Sales       | -0.087593 | 0.022740   | -3.852  | 0.000414 *** |

If we do 95% CI test, Advertising and Engineering is not significant since its P-value is larger than 0.05.

$$3.5 \quad y = \begin{pmatrix} y1 \\ y2 \\ y3 \\ \vdots \\ yn \end{pmatrix}, X = \begin{pmatrix} x1 \\ x2 \\ x3 \\ \vdots \\ xn \end{pmatrix}, e = \begin{pmatrix} e1 \\ e2 \\ e3 \\ \vdots \\ en \end{pmatrix}, \beta = \beta. \text{ Given 3.7 formula: } \beta^{\wedge} = (X'X)^{-1}X'y$$

$$X'X = \sum_{i=1}^n x_i^2, (X'X)^{-1} = \frac{1}{\sum_{i=1}^n x_i^2}; X'y = \sum_{i=1}^n x_i y_i$$

$$\text{So } \beta^{\wedge} = (X'X)^{-1}X'y = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2}$$