# Engineering Mathematics 2 Seoul National University

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### Exercise 8.1

Since *X* and *Y* are uniform random variables, we have the following density functions for  $x, y \in [0, 1]$ .

$$f_X(x) = 1, \ f_Y(y) = 1$$

The functions differ by whether X + Y is larger than 1 or not. If  $0 \le X + Y \le 1$ , we obtain

$$f_{X+Y}(z) = \int_0^z f_X(x) f_Y(z - x) dx = \int_0^z 1 \cdot 1 dx = z$$
$$F_{X+Y}(z) = \int_{-\infty}^z f_{X+Y}(z) dz = \int_0^z z dz = \frac{z^2}{2}.$$

If  $1 < X + Y \le 2$ , we obtain

$$f_{X+Y}(z) = \int_{z-1}^{1} f_X(x) f_Y(z-x) dx = \int_{z-1}^{1} 1 \cdot 1 dx = 2 - z$$
$$F_{X+Y}(z) = \int_{-\infty}^{z} f_{X+Y}(z) dz = \int_{0}^{1} z dz + \int_{1}^{z} (2-z) dz = 2z - \frac{z^2}{2} - 1.$$

### Exercise 8.4

Let X and Y the minute we each arrive at. Both are uniform random variables over [0, 60], and we should find the probability  $\Pr(|X - Y| \le 15)$ . We have the following density functions for  $x, y \in [0, 60]$ .

$$f_X(x) = \frac{1}{60}, \ f_Y(y) = \frac{1}{60}$$

The density and distribution functions for X-Y differ by whether X-Y is larger than 0 or not. If  $-60 \le X-Y \le 0$ , we obtain

$$f_{X-Y}(z) = \int_0^{60+z} f_X(x) f_Y(x-z) dx = \frac{60+z}{3600}$$
$$F_{X-Y}(z) = \int_{-\infty}^z f_{X-Y}(z) dz = \int_{-60}^z \frac{60+z}{3600} dz = \frac{z}{60} + \frac{z^2}{7200} + \frac{1}{2}.$$

If  $0 \le X - Y \le 60$ , we obtain

$$f_{X-Y}(z) = \int_{z}^{60} f_X(x) f_Y(x-z) dx = \frac{60-z}{3600}$$

$$F_{X-Y}(z) = \int_{-\infty}^{z} f_{X-Y}(z) dz = \int_{-60}^{0} \frac{60+z}{3600} dz + \int_{0}^{z} \frac{60-z}{3600} dz = \frac{z}{60} - \frac{z^2}{7200} + \frac{1}{2}.$$

Therefore

$$\Pr(|X - Y| \le 15) = \Pr(X - Y \le 15) - \Pr(X - Y \le -15) = F_{X+Y}(15) - F_{X+Y}(-15) = \frac{7}{16}.$$

### Exercise 8.9

Let X a uniform random variable in (0,1) and Y an exponentially distributed random variable with parameter  $\lambda$ . Then the connection between the two can be written as

$$Y = -\frac{\log X}{\lambda}.$$

This can be shown by retrieving the distribution function of *Y* , which is

$$F_Y(y) = \Pr(Y \le y) = \Pr\left(-\frac{\log X}{\lambda} \le y\right) = \Pr(X \ge e^{-\lambda y}) = 1 - F_X(e^{-\lambda y})$$
$$= \begin{cases} 0 & (y < 0) \\ 1 - e^{-\lambda y} & (y \ge 0) \end{cases}.$$

### Exercise 8.19

Let S the set of buses we want to take as soon as they arrive. Also let  $W_i$  the wait time of the ith line and W the wait time until the first bus's arrival, then  $W_i \sim \exp\left(1/\mu_i\right)$  and  $W = \min_{i \in S} W_i \sim \exp\left(\sum_{i \in S} \left(1/\mu_i\right)\right)$ . Therefore

$$\mathbf{E}[W] = \frac{1}{\sum_{i \in S} (1/\mu_i)}.$$

The probability of  $W_i$  being the minimum is  $\frac{1/\mu_i}{\sum_{i \in S} (1/\mu_i)}$ , so each elements of S will be chosen as the first bus with this probability. Now let T the travel time of the chosen bus, then

$$\mathbf{E}[T] = \sum_{i \in S} t_i \cdot \frac{1/\mu_i}{\sum_{j \in S} (1/\mu_j)} = \frac{\sum_{i \in S} (t_i/\mu_i)}{\sum_{i \in S} (1/\mu_i)}$$

Therefore the total expected time to cross town when we choose a bus from S is as the following.

$$\mathbf{E}[W] + \mathbf{E}[T] = \frac{1}{\sum_{i \in S} (1/\mu_i)} + \frac{\sum_{i \in S} (t_i/\mu_i)}{\sum_{i \in S} (1/\mu_i)} = \frac{1 + \sum_{i \in S} (t_i/\mu_i)}{\sum_{i \in S} (1/\mu_i)}$$

Our concern is to choose S that minimizes the above probability. If we name the bus lines  $1, 2, \dots, n$  in order of increasing  $t_i$ , we can restrict the range of possible tries of S to one of  $\{\{1\}, \{1, 2\}, \dots, \{1, 2, \dots, n\}\}$ . If S contains a specific bus line, lines with shorter travel time than that should also be in S.