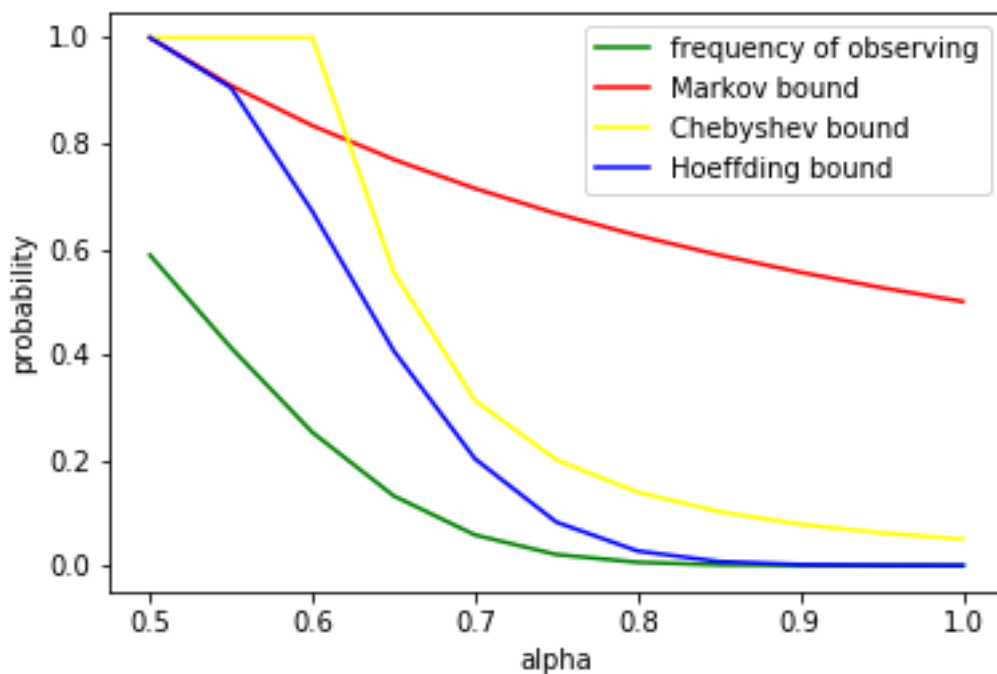


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1 Illustration of Hoeffding's Inequality



2. We can see that, the *Hoeffding's bound* converges to *frequency of observing* much faster than the *Chebyshev bound*. Especially, when $\alpha \geq 0.85$ the *Hoeffding's bound* almost equal to *frequency of observing*.

3. From *python*, we can get the probability with Hoeffding's bound

$$\begin{aligned} P\left(\frac{1}{20} \sum_{i=1}^{20} X_i \geq 0.95\right) &= 0.0003035391380788668 \\ P\left(\frac{1}{20} \sum_{i=1}^{20} X_i \geq 1\right) &= 4.5399929762484854e - 05 \end{aligned}$$

From Assignment 1, we can get the exact probability

$$\begin{aligned} P\left(\frac{1}{20} \sum_{i=1}^{20} X_i \geq 0.95\right) &= 2.002716064453125e - 05 \\ P\left(\frac{1}{20} \sum_{i=1}^{20} X_i \geq 1\right) &= 9.5367431640625e - 07 \end{aligned}$$

The Hoeffding's bound in $\alpha = 0.95$ is really close to the exact probability. But with the α increasing, the Hoeffding's bound become much closer and the convergence is much faster, at least at the rate of e^{-20} .

2 The effect of scale (range) and normalisation of random variables in Hoeffding's inequality

Theorem 2.3

$$P\left(\sum_{i=1}^n X_i - E\left[\sum_{i=1}^n X_i\right] \geq \epsilon\right) \leq e^{\frac{-2\epsilon^2}{\sum_{i=1}^n (b_i - a_i)^2}}$$

Because $X_i \in [0, 1]$, so

$$P\left(\sum_{i=1}^n X_i - E\left[\sum_{i=1}^n X_i\right] \geq \epsilon\right) \leq e^{\frac{-2\epsilon^2}{\sum_{i=1}^n (1-0)^2}}$$

Let $\epsilon = nk$, left side

$$\begin{aligned} P\left(\sum_{i=1}^n X_i - E\left[\sum_{i=1}^n X_i\right] \geq nk\right) &= P\left(\frac{1}{n} \sum_{i=1}^n X_i - \frac{1}{n} E\left[\sum_{i=1}^n X_i\right] \geq k\right) \\ &= P\left(\frac{1}{n} \sum_{i=1}^n X_i - \frac{1}{n} \sum_{i=1}^n E[X_i] \geq k\right) \\ &= P\left(\frac{1}{n} \sum_{i=1}^n X_i - \frac{1}{n} \sum_{i=1}^n \mu \geq k\right) \\ &= P\left(\frac{1}{n} \sum_{i=1}^n X_i - \frac{1}{n} \sum_{i=1}^n \mu \geq k\right) \\ &= P\left(\frac{1}{n} \sum_{i=1}^n X_i - \frac{1}{n} n\mu \geq k\right) \\ &= P\left(\frac{1}{n} \sum_{i=1}^n X_i - \mu \geq k\right) \end{aligned}$$

right side

$$e^{\frac{-2n^2k^2}{\sum_{i=1}^n 1^2}} = e^{\frac{-2n^2k^2}{n}} = e^{-2nk^2}$$

so Corollary 2.5 be proved $P\left(\frac{1}{n} \sum_{i=1}^n X_i - \mu \geq k\right) \leq e^{-2nk^2}$

3 Distribution of Student's Grades

1. Markov's inequality:

$$P(X \geq \epsilon) \leq \frac{E[X]}{\epsilon}$$

$$\begin{aligned} P(\hat{Z} \leq z) &= P(-\hat{Z} \geq -z) \\ &= P(100 - \hat{Z} \geq 100 - z) \\ &= P(\hat{Q} \geq 100 - z) \leq \frac{E[\hat{Q}]}{100 - z} \end{aligned}$$

$$\begin{aligned} E[\hat{Q}] &= E[100 - \hat{Z}] = E[100] - E[\hat{Z}] = 100 - 50 = 50 \\ \frac{50}{100 - z} &= 0.05 \\ z &= -900 \end{aligned}$$

2. Chebyshev's inequality:

$$P(|X - E[X]| \geq \epsilon) \leq \frac{Var[X]}{\epsilon^2}$$

$$\begin{aligned} P(\hat{Z} \leq z) &= P(\hat{Q} \geq 100 - z) \\ &= P(\hat{Q} - E[\hat{Q}] \geq 50 - z) \\ &\leq \frac{Var[\hat{Q}]}{(50 - z)^2} \end{aligned}$$

$$\begin{aligned} Var[\hat{Q}] &= E[\hat{Q}^2] - E[\hat{Q}]^2 = 2500 \\ \frac{2500}{(50 - z)^2} &= 0.05 \\ z &= -173.6 \end{aligned}$$

3. Hoeffding's inequality:

$$P(\sum_{i=1}^n X_i - E[\sum_{i=1}^n X_i] \leq -\epsilon) \leq e^{\frac{-2\epsilon^2}{\sum_{i=1}^n (b_i - a_i)^2}}$$

$z = 3.7$

4. So, we can find that, only Hoeffding's inequality can get a non-vacuous value is 3.7.

4 The Airline Question

1. We can see it as a binomial distribution
Markov's bound:

| | X_1 | X_2 |
|-----|-------|-------|
| X | 1 | 0 |
| P | 0.95 | 0.05 |

$$\begin{aligned}
 P\left(\sum_{i=1}^{100} X_i \geq 100\right) &\leq \frac{E[\sum_{i=1}^{100} X_i]}{100} \\
 &= \frac{100 * 0.95}{100} \\
 &= 0.95
 \end{aligned}$$

Chebyshev's bound:

$$\begin{aligned}
 P\left(\sum_{i=1}^{100} X_i \geq 100\right) &= P\left(\sum_{i=1}^{100} X_i - E\left[\sum_{i=1}^{100} X_i\right] \geq 100 - E\left[\sum_{i=1}^{100} X_i\right]\right) \\
 &\leq \frac{Var[\sum_{i=1}^{100} X_i]}{(100 - E[\sum_{i=1}^{100} X_i])^2} \\
 &= \frac{100 * 0.95 * 0.05}{(100 - 95)^2} \\
 &= 0.19
 \end{aligned}$$

Hoeffding's bound:

$$\begin{aligned}
 P\left(\sum_{i=1}^{100} X_i \geq 100\right) &= P\left(\sum_{i=1}^{100} X_i - E\left[\sum_{i=1}^{100} X_i\right] \geq 100 - E\left[\sum_{i=1}^{100} X_i\right]\right) \\
 &\leq 2e^{\frac{-2\epsilon^2}{\sum_{i=1}^{100} (1-0)^2}} \\
 &= 0.61
 \end{aligned}$$

So, we choose the Chebyshev's bound.

2. From 1. we use Chebyshev's inequality

$$\begin{aligned}
P\left(\sum_{i=1}^{10000} X_i = 9500\right) &\leq P\left(\sum_{i=1}^{10000} X_i \geq 9500\right) \\
&= P\left(\sum_{i=1}^{10000} X_i - E\left[\sum_{i=1}^{10000} X_i\right] \geq 9500 - E\left[\sum_{i=1}^{10000} X_i\right]\right) \\
&\leq \frac{\text{Var}\left[\sum_{i=1}^{10000} X_i\right]}{\left(9500 - E\left[\sum_{i=1}^{10000} X_i\right]\right)^2} \\
&= \frac{np(1-p)}{(9500 - np)^2}
\end{aligned}$$

Then, we calculate partial Derivative of p , to find the extremum.

5 Logistic Regression

5.1 Cross-entropy error measure

(a) The likelihood for i.i.d S :

$$\prod_i^N P(y_i|x_i)$$

We assume $p = [y = +1]; q = h(x_n)$, the likelihood of the training set:

$$\prod_i q^{(N-n)p} (1-q)^{n(1-p)}$$

so, the negative log-likelihood, divided by N is

$$\begin{aligned}
-\frac{1}{N} \ln\left(\prod_i^N P(y_i|x_i)\right) &= -\frac{1}{N} \ln\left(\prod_i q^{(N-n)p} (1-q)^{n(1-p)}\right) \\
&= \sum_i^N p \ln \frac{1}{q} + (1-p) \ln \frac{1}{1-q} \\
E_{in}(w) &= \sum_i^N [y_n = +1] \ln \frac{1}{h(x_n)} + [y_n = -1] \ln \frac{1}{1-h(x_n)}
\end{aligned}$$

The $h(x_n)$ from maximum likelihood minimizes the negative log-likelihood, and also minimizes the sample error.

(b) the error function:

$$\begin{aligned}
\frac{1}{N} \sum_{n=1}^N \ln(1 + e^{-y_n w^T x_n}) &= -\frac{1}{N} \sum_{n=1}^N \ln(\theta(-y_n w^T x_n)) \\
&= -\frac{1}{N} \sum_{i=1}^N \ln(P(y_i | x_i)) \\
&= -\frac{1}{N} \ln\left(\prod_i P(y_i | x_i)\right)
\end{aligned}$$

from(a)

$$\frac{1}{N} \sum_{n=1}^N \ln(1 + e^{-y_n w^T x_n}) = \sum_i^N [y_n = +1] \ln \frac{1}{h(x_n)} + [y_n = -1] \ln \frac{1}{1-h(x_n)}$$

5.2 Logistic regression loss gradient

For labels in $\{-1,1\}$

in-sample error:

$$E_{in}(w) = \frac{1}{N} \sum_{n=1}^N \ln(1 + e^{-y_n w^T x_n})$$

Partial derivative of w

$$\begin{aligned}
\nabla E_{in}(w) &= \frac{1}{N} \sum_{n=1}^N \frac{1}{1 + e^{-y_n w^T x_n}} (e^{-y_n w^T x_n} + 1)' \\
&= \frac{1}{N} \sum_{n=1}^N \frac{1}{1 + e^{-y_n w^T x_n}} e^{-y_n w^T x_n} (-y_n w^T x_n)' \\
&= \frac{1}{N} \sum_{n=1}^N \frac{1}{1 + \frac{1}{e^{y_n w^T x_n}}} e^{-y_n w^T x_n} (-y_n x_n) \\
&= \frac{1}{N} \sum_{n=1}^N \frac{e^{y_n w^T x_n}}{1 + e^{y_n w^T x_n}} e^{-y_n w^T x_n} (-y_n x_n) \\
&= \frac{1}{N} \sum_{n=1}^N \frac{1}{1 + e^{y_n w^T x_n}} (-y_n x_n) \\
&= -\frac{1}{N} \sum_{n=1}^N \frac{y_n x_n}{1 + e^{y_n w^T x_n}}
\end{aligned}$$

logistic function $\theta(s) = \frac{1}{1+e^{-s}}$, so

$$\begin{aligned}
\nabla E_{in}(w) &= -\frac{1}{N} \sum_{n=1}^N \frac{y_n x_n}{1 + e^{y_n w^T x_n}} \\
&= \frac{1}{N} \sum_{n=1}^N -y_n x_n \theta(-y_n w^T x_n)
\end{aligned}$$

its equals:

$$-\frac{1}{N} \sum_{n=1}^N [\frac{y_n+1}{2} - \theta(w^T x_n)] x_n$$

from this function, when the example is 'misclassified', the difference $\frac{y_n+1}{2} - \theta(w^T x_n)$ is larger than a correctly classified one.

For labels in $\{0,1\}$

from $\{-1,1\}$, we can get

$$-\frac{1}{N} \sum_{n=1}^N [\frac{y_n+1}{2} - \theta(w^T x_n)] x_n \frac{y_n+1}{2} \in [0,1]$$

for $\{0,1\}$, we can get $y \in [0,1]$ directly.

so

$$-\frac{1}{N} \sum_{n=1}^N [y_n - \theta(w^T x_n)] x_n \frac{y_n+1}{2} \in [0,1]$$

5.3 Logistic regression implementation

```
def gradientDescent(x,y):

    eta = [] %define eta
    t = [] %define the time of steps
    w = [] %define the parameter

    for i in range(0,t):

        loss = np.zeros(shape=(10,3))

        for j in range(0,10):

            hypothesis = np.dot(w,x1[j].T)
            loss[j] = np.dot(y[j],x1[j])/(1 + math.exp(np.dot(y[j],
                hypothesis)))

        gradient = loss.sum(axis=0)/(-10)

        w = w - eta * gradient

    return w
```

Firstly, we define $w(0)$. For the first step, we calculate $hx = w(0)^T x$, and then calculate the gradient $g_0 = -\frac{1}{N} \sum_{n=1}^N \frac{y_n x_n}{1+e^{y_n w(0)^T x_n}}$. Update the $w(0)$ with η and $gradient(0)$. After that, use $w(1)$ to repeat this process, until the $w(t)$.

I define the x with $[[1,0], [1,1], [1,2]...[1,8], [1,9]]$, y is a $1*10$ matrix and every elements is a random number 1 or 0. And assume $w = [0,0,0]$, $t = 10$, $\eta = 0.005$ we can get an final $w = [-0.21571744, -1.01579983, -0.21571744]$

5.4 Iris flower data

I use linear regression parameters as $w(0) = [0.53779839, -6.06028193, -0.54198853]$ and $\theta = 0.01$, $t = 10$.

so, we can get the parameters $w = [0.60310129, -6.05716725, -0.53041294]$

$$y = 0.60310129x_1 - 6.05716725x_2 - 0.53041294$$

By 0-1 loss:

$$\begin{aligned} Error_{train} &= \frac{61}{62} \\ Error_{test} &= \frac{13}{26} \end{aligned}$$