PH240C Assignment 2

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Problem 1.

(a) Write the log likelihood and plot it for a sensible choice of values Likelihood function:

$$L_n(\lambda) = \prod_{i=1}^n \frac{\lambda^{x_i}}{x_i!} e^{-\lambda}$$
$$= \frac{\lambda^{\sum_{i=1}^n x_i} e^{-n\lambda}}{\prod_{i=1}^n x_i!}$$

Log-Likelihood function:

$$l_n(\lambda) = (\sum_{i=1}^n x_i) \log \lambda - n\lambda - \sum_{i=1}^n \log x_i!$$

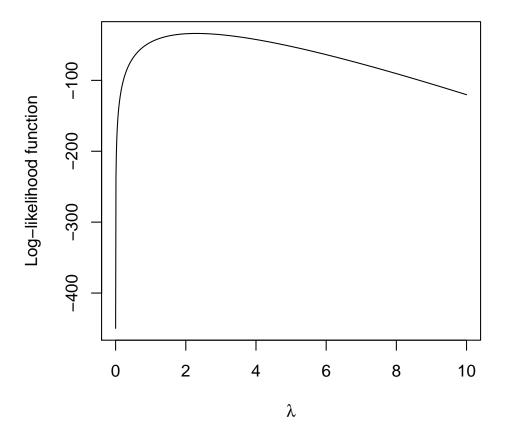
```
data1 = c(1,2,6,2,3,3,2,3,2,1,5,2,2,4,2,0,2,1,2,1)

lambda = seq(0.0001,10, length.out = 1000)

loglik = function(lambda,data){
    n = length(data)
    sumX = sum(data)
    sumLogFac = sum(log(factorial(data)))
    loglik = sumX*log(lambda) - n*lambda - sumLogFac
    return(loglik)
}

ll = loglik(lambda, data1)
plot(ll ~ lambda, main="Log-likelihood function for poisson distribution", xlab=expression(lambda),
    ylab = "Log-likelihood function", type ="l")
```

Log-likelihood function for poisson distribution



(b) Find the MLE analytically: Score function:

$$\frac{dl_n(\lambda)}{d\lambda} = \frac{\sum_{i=1}^n}{\lambda} - n$$

Set it equal to 0:

$$\frac{\sum_{i=1}^{n}}{\lambda} - n = 0$$

And, solve for λ :

$$\hat{\lambda} = \frac{\sum_{i=1}^{n} x_i}{n} = \bar{x}$$

Verify if we achieve maximum at the estimate using second derivative:

$$\frac{dl_n^2(\lambda)}{d^2\lambda} = -\frac{\sum_{i=1}^n x_i}{\lambda^2} < 0 \text{ since both the numerator and the denominator are positive.}$$

Therefore, the MLE of λ is \bar{x}

```
# analytically:
mean(data1)
## [1] 2.3
# numerically:
optimize(loglik, interval = c(0,4), data1, maximum =TRUE)
```

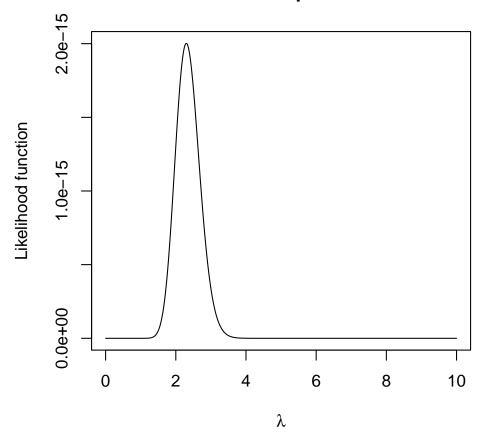
```
## $maximum
## [1] 2.3
##
## $objective
## [1] -33.84
```

From above, we can see that the results of finding MLE of λ both analytically and numerically are the same.

(c) $\theta = \log(\lambda)$, so $\lambda = e^{\theta}$. If we plug-in the e^{θ} to the likihood function and log-likehood function above, we will then have $L_n(\theta)$, and $l_n(\theta)$ correspondingly. The parameter space for θ : $\theta \in [-\infty, \infty]$

```
lik = function(data, lambda) {
    sumX = sum(data)
    n = length(data)
    facProd = prod(factorial(data))
    lik = (lambda^sumX * exp(-n*lambda)) / facProd
    return(lik)
}
l = lik(data1,lambda)
plot(1 ~ lambda, main="Likelihood function for poisson distribution", xlab=expression(lambda),
    ylab = "Likelihood function", type ="l")
```

Likelihood function for poisson distribution

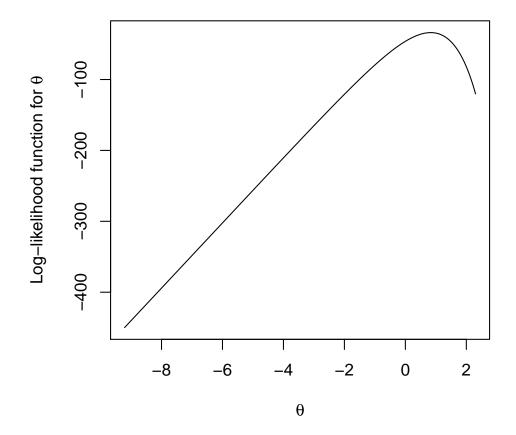


(d) Log-Likelihood function w.r.t θ :

$$l_n(\theta) = (\sum_{i=1}^n x_i)\theta - ne^{\theta} - \sum_{i=1}^n \log x_i!$$

```
# MLE of theta numerically:
loglik_t = function(theta, data){
    sumX = sum(data)
    n = length(data)
    sumLogFac = sum(log(factorial(data)))
    loglik_t = sumX*theta - n*exp(theta) - sumLogFac
    return(loglik_t)
}
theta = log(lambda)
ll_t = loglik_t(theta, data1)
plot(ll_t~theta, xlab=expression(theta),
    ylab = expression(paste("Log-likelihood function for ",theta)), type ="l")
title(main="Log-likelihood function for poisson distribution-different parametrization", cex.main=0.7)
```

Log-likelihood function for poisson distribution-different parametrization



```
optimize(loglik_t, interval = c(-1,2), data1, maximum =TRUE)

## $maximum
## [1] 0.8329
##
## $objective
## [1] -33.84

log(2.3)
## [1] 0.8329
```

From the above, we can see that numerically, $\theta_n = \log(\lambda_n)$

Problem 2.

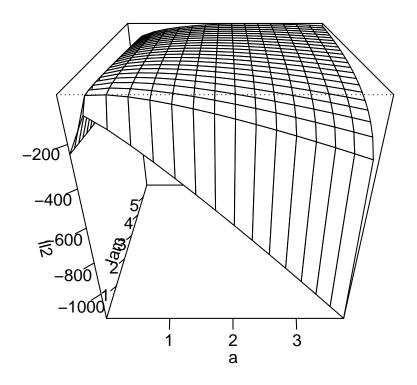
(a) Write the log likelihood and plot it for a sensible choice of values Likelihood function:

$$L_n(\alpha, \lambda) = \prod_{i=1}^n \frac{1}{\Gamma(\alpha)} \lambda^{\alpha} y_i^{\alpha - 1} e^{-\lambda y_i}$$

Log-likelihood function:

$$l_n(\alpha, \lambda) = -n \log \Gamma(\alpha) + n\alpha \log(\lambda) + (\alpha - 1) \sum_{i=1}^{n} \log y_i - \lambda \sum_{i=1}^{n} y_i$$

```
data2 = c(0.15812, 0.30070, 0.48016, 0.49813, 0.20042, 0.26716, 0.80124,
          0.10914, 0.57169, 0.83686, 1.57027, 0.10458, 0.58490, 1.14454,
          0.61595, 0.28155, 0.13236, 0.36252, 0.08614, 0.27907, 0.46010,
          0.03824, 0.76581, 0.30369, 0.42404, 0.57530, 0.26987, 0.22416,
          0.07673, 1.09659)
loglik2 = function(a,lam, data){
  n = length(data)
  # gammar of alpha
  ga = factorial(a-1)
  sumY = sum(data)
  sumLog = sum(log(data))
  loglik2 = -n*log(ga) + n*a*log(lam) + (a-1)*sumLog - lam*sumY
  return(loglik2)
a = seq(0.0001, 4, by = 0.25)
lam = seq(0.0001, 6, by = 0.25)
pars = c(a=a,lam=lam)
112 = outer(a,lam,loglik2,data2)
persp(a,lam,ll2, ticktype = "detail", phi = 30)
```



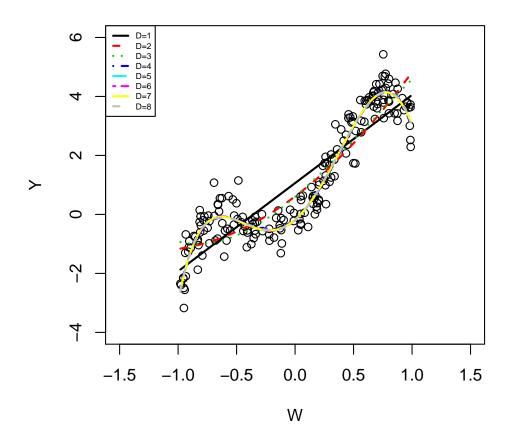
(b)

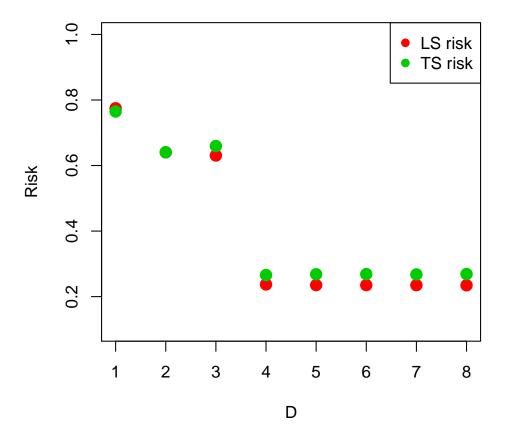
```
negloglik2 = function(pars, data){
  a = pars[1]
  lam = pars[2]
  n = length(data)
  # gammar of alpha
  ga = factorial(a-1)
  sumY = sum(data)
  sumLog = sum(log(data))
  loglik2 = -(-n*log(ga) + n*a*log(lam) + (a-1)*sumLog - lam*sumY)
  return(loglik2)
optim(c(0.001,0.001), negloglik2, data=data2)
## $par
## [1] 1.677 3.694
##
## $value
## [1] 4.235
##
## $counts
## function gradient
##
       123
##
## $convergence
## [1] 0
##
```

```
## $message
## NULL
```

Problem 3. Answers to part(a) and par(b) are combined below:

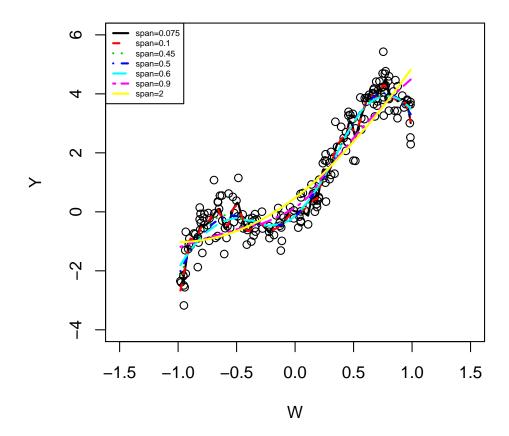
```
# reading in the data
ls = read.table("learning.txt")
ts = read.table("test.txt")
### Polynomial using lm ###
############################
D = 1:8
# create an empty list to store model for later
fitLm = vector("list", length(D))
# create an empty matrix to store fitted Y for training set
YhatLm.ts = matrix(NA, nrow(ts), length(D))
# create an empty matrix to store fitted Y for learning set
YhatLm.ls = matrix(NA, nrow(ls), length(D))
# create vectors of length(D) to store LS risk and TS risk for each model
LSRiskLm = rep(NA, length(D))
TSRiskLm = rep(NA, length(D))
w.ls = ls$W
w.ts = ts$W
W.ls = NULL
W.ts = rep(1, nrow(ts))
for(d in 1:length(D))
  W.ls = cbind(W.ls, w.ls^D[d])
 W.ts = cbind(W.ts, w.ts^D[d] )
  fitLm[[d]] = lm(ls$Y ~ W.ls)
  YhatLm.ts[,d] = W.ts %*% as.matrix(fitLm[[d]]$coef)
  YhatLm.ls[,d] = fitLm[[d]]$fitted
 LSRiskLm[d] = mean(fitLm[[d]]$residuals^2)
  TSRiskLm[d] = mean((YhatLm.ts[,d] - ts$Y)^2, na.rm=TRUE)
# lm fits
plot(ls$W, ls$Y, xlim=c(-1.5, 1.5), ylim=c(-4, 6),
     xlab="W", ylab="Y", main="Mystery model: lm fits, Y \sim 1 + W + ... + W^{\rm D}",
     cex.main = 0.5)
matplot(ls$W, YhatLm.ls, type="1", lty=1:length(D), lwd=2,
        col=1:length(D), add=TRUE)
legend("topleft", paste("D=",D,sep=""), lty=1:length(D),
       lwd=2, col=1:length(D), cex = 0.5)
```

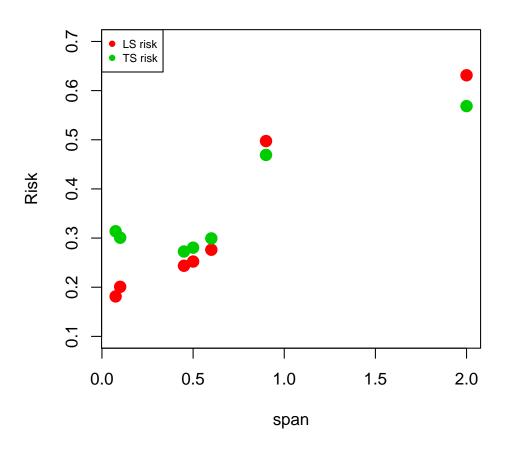




```
## Comment: From the plot, we can see that both LS empirical risk and TS empirical risk are
## decreasing as the D becomes bigger. I think D=4 is the best fit. It has low TS empirical risk,
## and also avoid posibilities of overfitting
##############
### loess ####
#############
span \leftarrow c(0.075, 0.1, 0.45, 0.5, 0.6, 0.9, 2)
# Create an empty matrix to store fitted Y for LS
YhatLoess <- matrix(NA, nrow(ls), length(span))</pre>
LSRiskLoess <- TSRiskLoess <- rep(NA,length(span))
for(j in 1:length(span))
  fit <- loess(Y ~ W, span=span[j], data=ls)</pre>
  YhatLoess[,j] <- fit$fitted
 LSRiskLoess[j] <- mean(fit$residuals^2)
  pred <- predict(fit,data.frame(W=ts$W))</pre>
  TSRiskLoess[j] <- mean((pred-ts$Y)^2, na.rm=TRUE)</pre>
# loess fits
plot(ls$W, ls$Y, xlim=c(-1.5, 1.5), ylim=c(-4, 6), xlab="W", ylab="Y",
     main=paste("Mystery model: loess fits, span=", deparse(span), sep=""),
     cex.main = 0.5)
matplot(ls$W, YhatLoess, type="l", lty=1:length(span),
        lwd=2, col=1:length(span), add=TRUE)
legend("topleft", paste("span=",span,sep=""), lty=1:length(span),
```

Mystery model: loess fits, span=c(0.075, 0.1, 0.45, 0.5, 0.6, 0.9, 2)





Comment: From this plot, we can see that LS empirical risk is monotonically increasing as we
increase the span. We can identify a minimum TS empirical risk clearly from the plot when span=0.45.
Therefore, I think span=0.45 is the best fit.

lapply(fitLm, function(z) round(z\$coefficients,2))

<pre>lapply(fitLm, function(z) round(z\$coefficients,2))</pre>									
## ## ##	[[1]] (Intercept) 1.06	W.ls 2.99							
	(Intercept) 0.61	W.ls1 3.00	W.ls2 1.22						
##	[[3]] (Intercept) 0.60	W.ls1 3.39	W.ls2 1.23	W.ls3 -0.64					
## ##	[[4]] (Intercept)	W.ls1	W.1s2	W.ls3	W.ls4				
## ##	-	3.25	8.20	-0.43	-7.98				
	(Intercept)	W.ls1 3.54	W.ls2 8.30	W.ls3 -1.72	W.ls4 -8.09	W.ls5 1.14			
##	[[6]] (Intercept)	W.ls1	W.ls2	W.1s3	W.ls4	W.ls5			

```
3.54 8.43 -1.72 -8.46 1.12
##
       -0.18
##
        W.ls6
##
        0.26
##
## [[7]]
                           W.ls2
## (Intercept)
                 W.ls1
                                      W.ls3
                                                 W.ls4
                                                             W.ls5
   -0.18
                            8.39
                                       -0.77
                                                  -8.36
                                                            -0.95
##
                  3.44
##
       W.ls6
                  W.ls7
##
       0.18
                  1.28
##
## [[8]]
                           W.ls2
## (Intercept)
                 W.ls1
                                      W.ls3
                                                 W.ls4
                                                           W.ls5
     -0.21
##
                  3.42
                             9.30
                                       -0.59
                                                 -13.32
                                                            -1.36
      W.ls6
##
                  W.ls7
                             W.ls8
##
       8.83
                  1.56
                             -4.67
# lm fits: LS and TS risk
round(rbind(D, LSRiskLm, TSRiskLm), 4)
           [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
## D
          1.0000 2.0000 3.0000 4.0000 5.0000 6.0000 7.0000 8.0000
## LSRiskLm 0.7747 0.6403 0.6307 0.2371 0.2351 0.2351 0.2350 0.2345
## TSRiskLm 0.7648 0.6404 0.6595 0.2658 0.2681 0.2686 0.2677 0.2689
#loess fits: LS and TS risk
round(rbind(span, LSRiskLoess, TSRiskLoess), 4)
##
              [,1] [,2] [,3] [,4] [,5] [,6] [,7]
         0.0750 0.1000 0.4500 0.5000 0.6000 0.9000 2.0000
## LSRiskLoess 0.1813 0.2009 0.2437 0.2522 0.2763 0.4973 0.6312
## TSRiskLoess 0.3139 0.3006 0.2727 0.2805 0.2994 0.4691 0.5685
```