

# CEF, OLS, and Causal Identification

Causal Inference

Spring 2025

# Overview: From Statistics to Causality

- **Previously covered:**

- ▶ The Potential Outcomes Framework ( $Y(0)$ ,  $Y(1)$ )
- ▶ Identification and Estimation

- **Today's Goal:**

- ▶ Linking (Causal) Identification and (Statistical) Estimation
- ① **Causal Estimand:** The Average Treatment Effect ( $E[Y(1) - Y(0)]$ ).
- ② **Identification Assumptions:** Ignorability and Positivity.
- ③ **Identified Statistical Object:** The CEF ( $E[Y|D, X]$ ).
- ④ **Estimation Strategy:** Using OLS to approximate the CEF.

- **The Causal Estimand** is the *theoretical* target in the *population*.
- It is what we want to know, independent of any specific data or estimator.
- **Average Treatment Effect (ATE):**

$$\tau_{ATE} = E[Y_i(1) - Y_i(0)]$$

- **Fundamental Problem of Causal Inference:** We never observe both  $Y_i(1)$  and  $Y_i(0)$  for the same unit.
- **Selection Bias:** Observed differences (SDM) do not equal the ATE unless  $E[Y(0)|D = 1] = E[Y(0)|D = 0]$ .

# Identification Assumptions

- Identification is the process of expressing the **Causal Estimand** in terms of **Observed Data**.

## ① Conditional Ignorability (CIA):

$$\{Y(0), Y(1)\} \perp D \mid X$$

*This implies:*  $E[Y(d)|X] = E[Y|D = d, X]$

## ② Positivity (Overlap):

$$0 < P(D = 1 \mid X = x) < 1 \quad \forall x$$

*This ensures:* We have both treated and control units across all values of  $X$ . If  $P(D = 1|X) = 1$ , we have no counterfactual.

- Once identified, our target is no longer a “potential” outcome; it is a **Statistical Estimand**: the **Conditional Expectation Function (CEF)**.
- Under CIA, the ATE is identified as:

$$\tau_{ATE} = E_X[\underbrace{E[Y|D = 1, X] - E[Y|D = 0, X]}_{\delta(X) \text{ (The Stratum (conditional on } X \text{) Effect)}}]$$

- **Law of Iterated Expectations (LIE)** to recover population ATE

## OLS as an Estimation Strategy

- $Y = \beta D + \gamma X + \epsilon$ , where  $X$  is a set of strata dummies
- By the **Frisch-Waugh-Lovell Theorem**(QPM1, partialling out approach):
  - ▶  $D - E[D|X]$ : “remaining” variation
  - ▶ The OLS estimator for  $\beta$  is:

$$\begin{aligned}\beta_{OLS} &= \frac{\text{Cov}(\text{residuals}_D, Y)}{\text{Var}(\text{residuals}_D)} \\ &= \frac{E[\text{residuals}_D Y] - E[\text{residuals}_D]E[Y]}{\text{Var}(\text{residuals}_D)} \\ &= \frac{E[(D - E[D|X])Y]}{E[(D - E[D|X])^2]}\end{aligned}$$

- The denominator  $E[(D - E[D|X])^2]$  is the expected conditional variance,  $E_X[\text{Var}(D|X)]$ .
- Since  $D$  is binary,  $\text{Var}(D|X) = P(D = 1|X)(1 - P(D = 1|X))$ .

## OLS as an Estimation Strategy

- We can rewrite the numerator using the Law of Iterated Expectations:

$$E[(D - E[D|X])Y] = E_X[E[(D - E[D|X])Y | X]]$$

- Focus on the inner expectation  $E[(D - E[D|X])Y | X]$ :

$$E[DY | X] - E[D | X]E[Y | X]$$

- By the definition of  $E[DY|X]$ :

$$E[DY|X] = P(D = 1|X)E[1 \cdot Y|D = 1, X] + P(D = 0 | X)E[0 \cdot Y | D = 0, X]$$

- Since  $D$  is a binary,  $E[D | X] = P(D = 1 | X)$
- Substitute the definition of  $E[Y|X]$  for a binary  $D$ :

$$E[Y|X] = P(D = 1|X)E[Y|D = 1, X] + P(D = 0|X)E[Y|D = 0, X]$$

## OLS as an Estimation Strategy

$$\begin{aligned} & E[DY \mid X] - E[D \mid X]E[Y \mid X] \\ &= P(D = 1 \mid X)E[Y \mid D = 1, X] - P(D = 1 \mid X)E[Y \mid X] \\ &= P(D = 1 \mid X)(E[Y \mid D = 1, X] - E[Y \mid X]) \\ &= P(D = 1 \mid X) \left( E[Y \mid D = 1, X] \right. \\ &\quad \left. - P(D = 1 \mid X)E[Y \mid D = 1, X] - P(D = 0 \mid X)E[Y \mid D = 0, X] \right) \\ &= P(D = 1 \mid X) \left( \right. \\ &\quad \left. P(D = 1 \mid X)E[Y \mid D = 1, X] + P(D = 0 \mid X)E[Y \mid D = 1, X] \right. \\ &\quad \left. - P(D = 1 \mid X)E[Y \mid D = 1, X] - P(D = 0 \mid X)E[Y \mid D = 0, X] \right. \\ &\quad \left. \right) \\ &= P(D = 1 \mid X)P(D = 0 \mid X) \left( E[Y \mid D = 1, X] - E[Y \mid D = 0, X] \right) \end{aligned}$$



## OLS as an Estimation Strategy

- Back into the full formula:

$$\beta_{OLS} = \frac{E_X[\overbrace{P(D=1|X)(1-P(D=1|X))}^{\text{Var}(D|X)} \cdot \overbrace{(E[Y|D=1,X] - E[Y|D=0,X])}^{\delta(X)}]}{E_X[P(D=1|X)(1-P(D=1|X))]}$$

- If  $X$  is discrete, this is exactly a weighted average of  $\delta(X)$ :

$$\beta_{OLS} = \frac{\sum_x w(x) \delta(x)}{\sum_x w(x)} \quad \text{where } w(x) = P(X=x) \text{Var}(D|X=x)$$

- $\delta(x)$  is the treatment effect within a specific group  $X=x$ .
- $w(x)$  is the weight given to that group by OLS.

# OLS (Estimation) vs. ATE (Estimand)

## 1 ATE (Estimand):

- ▶ Weights each  $X = x$  strictly by its population size,  $P(X = x)$ .
- ▶  $\tau_{ATE} = \sum_x P(X = x) \delta(X = x)$ .

## 2 OLS (Estimation Strategy):

- ▶ Weights each  $X = x$  by  $P(X = x) \text{Var}(D \mid X = x)$
- ▶ Linear approximation

# R Application: Wage Gender Gap and Working Hours

- **Simple Difference:** `r mean(log_earn[female==1]) - mean(log_earn[female==0])`
  - ▶ Statistical Estimand:  $E[Y|D = 1] - E[Y|D = 0]$ .
  - ▶ Only causal if Selection Bias is zero.
- **Binned CEF:** `r group_by(hours) %>% summarise(mean(log_earn))`
  - ▶ Identified Statistical Object:  $E[Y|X]$ .
  - ▶ Does not assume linearity; shows non-linearities OLS might miss.
- **Multivariate Regression:** `r lm(log_earn ~ female + hours)`
  - ▶ Estimation Strategy:  $\beta_{OLS}$  is the variance-weighted average of the gender gap across different categories of hours worked.

# Conclusion

- **Distinguish between:**
  - ▶ **Causal Estimand:** The ATE (The theoretical target).
  - ▶ **Identified Object:** The CEF (The target translated to data).
  - ▶ **Estimator:** OLS (The strategy to calculate the identified object).
- **The OLS Weighting Result:**  $OLS \neq ATE$  if treatment effects  $\delta(x)$  are **heterogeneous** across strata.
- OLS is an efficient estimator for a common effect, but a variance-weighted average for heterogeneous effects.