Advanced Quantitative Research Methodology, Lecture Notes: Theories of Inference¹

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3. A more reasonable, limited goal. Let $M = \{M^*, \theta\}$, where M^* is assumed & θ is to be estimated:

$$P(\theta|y, M^*) \equiv P(\theta|y)$$

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- 8. The two differ on the rest

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6. $L(\theta|y)$ is a function: for y fixed at the observed values, it gives the "likelihood" of any value of θ .

7. Likelihood: a relative measure of uncertainty, changing with the data

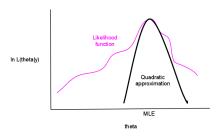
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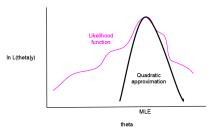
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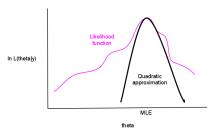


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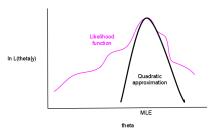
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- Uncertainty of point estimate: curvature at the maximum

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- \bullet P(θ), the prior density the way Bayes differs from likelihood

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- 4. A philosophical assumption that nonsample information should matter (as it always does) and be formalized and included in all inferences.

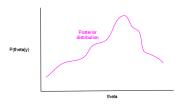
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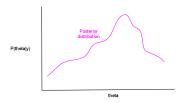
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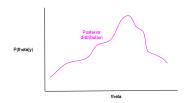
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- 4. If variables A and B are both unknown, then the distribution of A alone is $P(A) = \int P(A, B) dB = \int P(A|B) P(B) dB$.

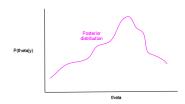




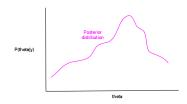
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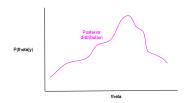
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- If $P(\theta) = 1$, i.e., is uniform in the relevant region, then $L(\theta|y) = P(\theta|y)$.

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$$(TS)_{\beta}|(\beta=0)\equiv \frac{b-\beta}{\hat{\sigma}_b}\equiv \frac{b}{\hat{\sigma}_b}\sim N(0,1).$$

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$$\begin{cases} \beta > 0 \text{ (I was right)} & \text{if } (TS) > (CV) \\ \beta = 0 \text{ (I was wrong)} & \text{if } (TS) \leq (CV) \end{cases}$$

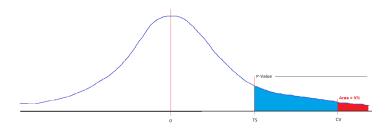
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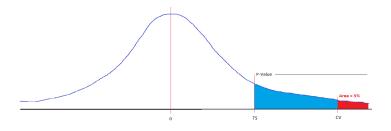
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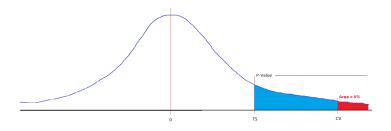
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And then first collect your data. You may not revise your hypothesis or your theory.

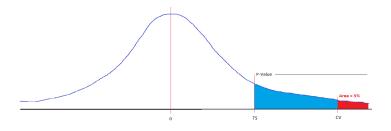




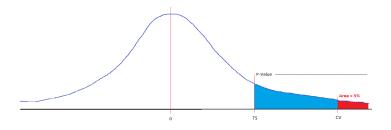
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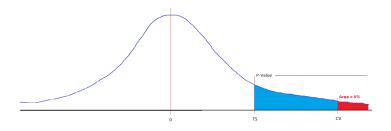
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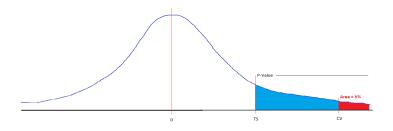
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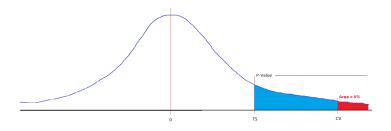
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- We can use likelihood to compute hypothesis tests and p-values.

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- 4. Methods for applied researchers: either useful or irrelevant

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- The key: No assumptions can be trusted; all theories of inference condition on assumptions and so data analysts always struggle trying to understand and get around them

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• What can you do with this probability density?

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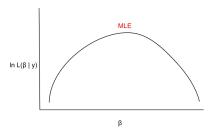
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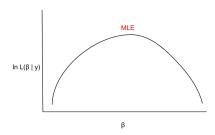
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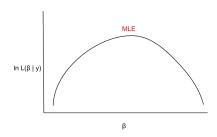
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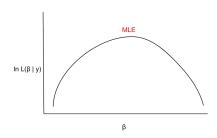
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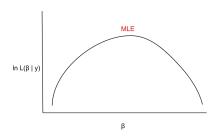
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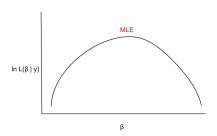
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① Consistency (from the Law of Large Numbers). As $n \to \infty$, the sampling distribution of the MLE collapses to a spike over the parameter value

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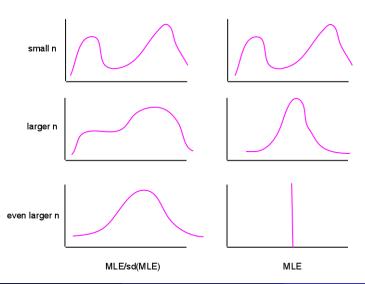
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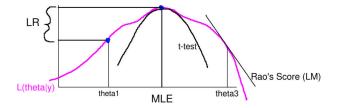
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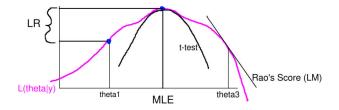
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 - Do the LLN and CLT (the 2 most important theorems in statistics) contradict each other?
- Asymptotic efficiency. The MLE contains as much information as can be packed into a point estimator.

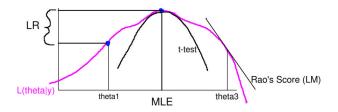
Sampling distributions of the MLE: CLT vs LLN



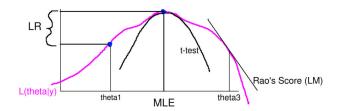




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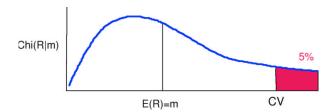
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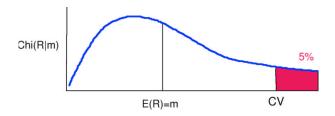
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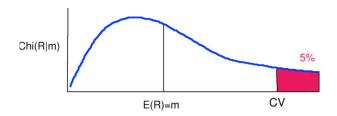
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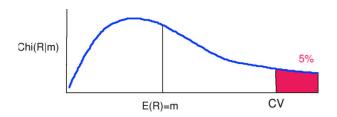
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Meaning of the likelihood ratio

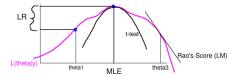


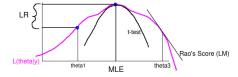
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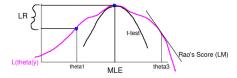


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- So only if r >> m will the test parameters be clearly different from zero.
- Disadvantage: Too many likelihood ratio tests may be required to test all points of interest

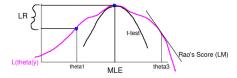




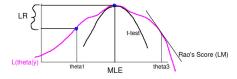
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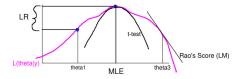
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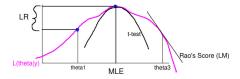


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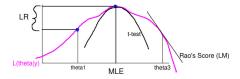
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8. We invert the curvature to provide a statistical interpretation:

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 - We'll discuss later how how to improve the approximation.

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- Compute and save $\hat{V}(\hat{\theta})$, which is $k+2 \times k+2$

• Mathematical Form:

$$\ln L(\beta, \sigma^2 | y) = \sum_{i=1}^n \sum_{t=1}^T -\frac{1}{2} \left[\ln \sigma^2 + \frac{(y_{it} - X_{it}\beta)^2}{\sigma^2} \right]$$

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X <- as.matrix(cbind(1, X))
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• Calling it:

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ll.normal(c(2,1,2,1,33,4,3.2),x,y)
ll.normal(c(2,1,2,1,33,4,3.7),x,y)
ll.normal(c(2,1,2,1,33,4,3.5),x,y)
```

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- Repeat Steps 1–3 M=1,000 times, and plot a histogram of the results.

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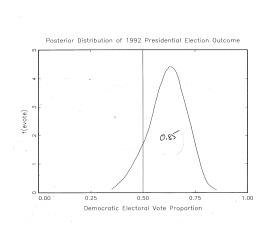
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Actual Results (calculated before the election) for 1992

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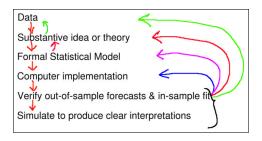
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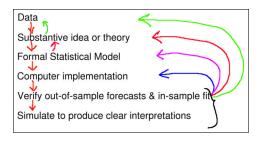
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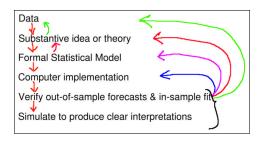
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• For what applications would this model be informative?

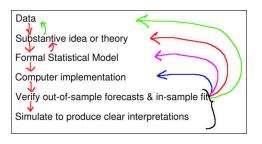




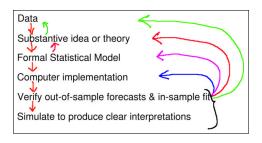
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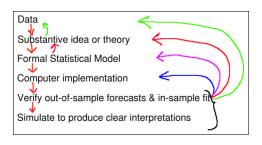
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- 5. Don't miss any parts.