

Hierarchical Bayesian Personalized Recommendation: A Case Study and Beyond

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ABSTRACT

Items in modern recommender systems are often organized in hierarchical structures. These hierarchical structures and the data within them provide valuable information for building personalized recommendation systems. In this paper, we propose a general hierarchical Bayesian learning framework, i.e., *HBayes*, to learn both the structures and associated latent factors. Furthermore, we develop a variational inference algorithm that is able to learn model parameters with fast empirical convergence rate. The proposed *HBayes* is evaluated on two real-world datasets from different domains. The results demonstrate the benefits of our approach on item recommendation tasks, and show that it can outperform the state-of-the-art models in terms of precision, recall, F1 measurement and normalized discounted cumulative gain.

KEYWORDS

Hierarchical Bayesian; Recommendation; Personalization

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1 INTRODUCTION

Real-world organizations in business domains operate in a multi-item, multi-level environment. Items and their corresponding information collected by these organizations often reflect a hierarchical structure. For examples, products in retail stores are usually stored in hierarchical inventories. News on web pages is created and placed hierarchically in most websites. These hierarchical structures and the data within them provide a large amount of information when building effective recommendation systems. Especially in e-commerce domain, all products are displayed in a site-wide hierarchical catalog and how to build an accurate recommendation engine on top of it becomes one of the keys to majority companies' business success.

However, how to utilize the rich information behind hierarchical structures to make personalized and accurate product recommendations still remains challenging due to the unique characteristics of hierarchical structures and the modeling trade-offs arising from them. Briefly, most well-established recommendation algorithms cannot naturally take hierarchical structures as additional inputs

and flattening hierarchical structures usually doesn't work well. It will not only blow up the entire feature space but introduce noise when training the recommendation models. On the other hand, completely ignoring hierarchies will lead to recommendation inaccuracies. The most common way to alleviate this problem is to feed every piece of data from the hierarchy into a complex deep neural network and hope the network itself can figure out a way to use the hierarchical knowledge. However, such approaches usually behave more like black boxes. They are difficult to debug and cannot provide any interpretation of their intermediate results and outcomes.

In this work, we propose and develop a hierarchical Bayesian, a.k.a., *HBayes*, modeling framework that is able to flexibly capture various relations between items in hierarchical structures from different recommendation scenarios. By introducing latent variables, all hierarchical structures are encoded as conditionally independences in *HBayes* graphical models. Moreover, we develop a variational inference algorithm for learning parameters of *HBayes*.

To illustrate the modeling power of the proposed *HBayes* approach, we introduce *HBayes* by using a real-world apparel recommendation problem as an example. As an illustration, we form apparel styles, product brands and apparel items into a three-level hierarchy. We add additional latent variables as the apparel style membership variables to capture the diverse and hidden style properties of each brand. Furthermore, we include user-specific features into *HBayes* and extend *HBayes* into the supervised learning setting where user feedback events such as clicks and conversions are incorporated. Please note that our *HBayes* framework is not only limited to apparel recommendation. In the end, we show its flexibility and effectiveness on a music recommendation problem as well.

Overall this paper makes the following contributions:

- It presents a generalized hierarchical Bayesian learning framework to learn models from rich data with hierarchies.
- It provides a variational inference algorithm that can learn the model parameters with few iterations.
- We evaluate our *HBayes* and its benefits comprehensively in tasks of apparel recommendation on a real-world data set.
- We test our framework in different recommendation scenarios to show its generalization and applicability.

The remainder of the paper is organized as follows: Section 2 provides a review of existing recommendation algorithms and their extensions in hierarchical learning settings. Section 3 introduces the notations and our generalized *HBayes* learning framework and its

variational inference algorithm. In Section 4, we conduct experiments in a real-world e-commerce data set to show the effectiveness of our proposed recommendation algorithm in different aspects. In addition, we test our model on a music recommendation data set to illustrate the generalization and extended ability of HBayes. We summarize our work and outline potential future extensions in Section 5.

2 RELATED WORK

2.1 Recommender Systems

With the explosive growth of digital information over the Internet, the recommender system is considered to be one of the most effective approaches to overcome such information overload. Traditional recommendation algorithms such as item-based approaches learn interactions between users and items and recommend items to users who share similar historical behaviors: collaborative filtering [27, 29] and matrix factorization [25] are both effective approaches under this category. Content-based approaches including [18, 19, 34] take use of the auxiliary information of both users and items for recommending items to users that are close in the content space. Furthermore, session based recommender systems are developed by analyzing user session information and user visiting patterns [13, 30]. These particular recommendation techniques take *time* as an additional input dimension, aiming at explicitly modeling user interests over time. Koren [17] develops a collaborative filtering type of approach with predictions from static average values combining with a dynamic changing factor. Yin et.al. [33] proposes a user-tag-specific temporal interest model to track user interests over time by maximizing the time weighted data likelihood.

2.2 Bayesian Approaches for Recommendation

Recently, there are works using Bayesian inferencing for recommendation tasks. Rendle et.al. [24] combines the Bayesian inference and matrix factorization together for learning users implicit feedbacks (click & purchase) that is able to directly optimize the recommendation ranking results. Ben-Elazar et.al. [1] and Zhang et.al. [35] take user preference consistency into account and develop a variational Bayesian personalized ranking model for better music recommendation. However, these approaches do not leverage the item structural information when building their Bayesian models. Given the fact that the hierarchical structural information widely exists in real-world recommendation scenarios such as e-commerce, social network, music, etc. failing to utilize such information makes these Bayesian approaches inefficient and inaccurate.

2.3 Hierarchical Recommender Systems

Hierarchical information for recommender systems, as mentioned, is a natural yet powerful entity structure that encodes human knowledge by means of tree-based dependency constraints between entities defined by a group of latent parameters. Recommender system hierarchies, in particular, could be either explicit or implicit; either approximate or exact. There are recommendation approaches take use of such information in order to promote items to users who have explicitly visited hierarchically related items or have shown preferences to items that belong to the same sub-categories. In social networks, Shepitsen et.al. [28] rely on the hierarchies

generated by user-taggings, a.k.a., *folksonomies*, to build a better personalized recommender system. In e-commerce, Wang et.al. [32] introduce a hierarchical matrix factorization approach that exploits the intrinsic structural information to alleviate cold-start and data sparsity problems. Despite the fact that these hierarchical recommender systems have received some success, there are still challenges such as: 1. how to infer the hierarchical structure efficiently and accurately if it is not explicit; 2. how to better understand the hierarchical topologies discovered by recommendation approaches; 3. how to utilize the inferred hierarchical information for precise data explanation all remain unsolved.

3 THE HBAYES FRAMEWORK

In this work, we develop our generalized hierarchical Bayesian modeling framework that is able to capture the hierarchical structural relations and latent relations in the real-world recommendation scenarios. In our framework, we incorporate both user specific features and user events into model learning so that more personalized recommendation results can be achieved. Furthermore, we present a variational inference algorithm for the HBayes framework to provide fast parameter estimation.

In the following, we will describe how HBayes works by using apparel recommendation as an illustrative example. Using a real-world example helps explain the latent variables, hierarchy structural relations, and the conditional independence assumptions implied in HBayes. Please note that even we explain HBayes in apparel recommendation, the framework itself can be generalized into other recommendation scenarios with little modification.

3.1 Generative Process

In the real-world scenario, each item or product has to come with a brand and a brand may have more than one items in the hierarchical structures. Therefore, we denote each event t as a 4-tuple (Item, Brand, User, IsClick), i.e. (X_t, b_t, u_t, y_t) . X_t represents the item features associated with event t and y_t is the binary label that indicates whether user u_t has clicked X_t or not. b_t is the brand of item X_t .

Furthermore, we expand the hierarchy by a hidden factor, i.e., “style”. Products from each brand b_t tend to exhibit different styles or tastes, which are unknown but exist. In this paper, brands are represented as random mixtures over latent styles, where each style is characterized by a distribution over all the items. Let S, B, U and N be the total number of styles, brands, users and events.

The generative process of HBayes can be described as follows:

- Step 1. Draw a multivariate Gaussian prior for each user k , i.e., $U_k \sim \mathcal{N}(0, \delta_u^{-1} \mathbf{I})$ where $k \in \{1, \dots, U\}$.
- Step 2. Draw a multivariate Gaussian prior for each style j , i.e., $S_j \sim \mathcal{N}(\mathbf{w}, \delta_s^{-1} \mathbf{I})$ where $j \in \{1, \dots, S\}$.
- Step 3. Draw a style proportion distribution θ for each brand i , $\theta \sim \text{Dir}(\gamma)$ where $i \in \{1, \dots, B\}$.
- Step 4. For each brand i :
 - Step 4.1 Draw style assignment z_i for brand i where the selected style p is sampled from $\text{Mult}(\theta)$. z_i is a $S \times 1$ one hot encoding vector that $z_{i,p} = 1$ and $z_{i,j} = 0$ for $j = 1, \dots, p-1, p+1, \dots, S$.
 - Step 4.2 Draw $B_i \sim \mathcal{N}(S_p, \delta_b^{-1} \mathbf{I})$.

Step 5. For each event t , draw y_t from Bernoulli distribution where the probability p is defined as $p(y_t | \mathbf{x}_t, \mathbf{B}_{b_t}, \mathbf{U}_{u_t})$.

where δ_s , δ_u and δ_b are the scalar precision parameters and \mathbf{w} is the prior mean of \mathbf{S}_j . $\text{Dir}(\cdot)$ and $\text{Mult}(\cdot)$ represent Dirichlet distribution and multinomial distribution, respectively.

In this work, in order to have the higher modeling flexibility and capacity, we also treat each distribution's parameter as a random variable and define hyper-priors over them. More specifically, We draw the prior mean \mathbf{w} from $\mathcal{N}(\mathbf{0}, \delta_w^{-1} \mathbf{I})$. For δ_w , δ_s , δ_u and δ_b , we define Gamma priors over them i.e., $p(\delta_*) = \mathcal{G}(\alpha, \beta)$, where $\delta_* \in \{\delta_w, \delta_s, \delta_u, \delta_b\}$.

The family of probability distributions corresponding to this generative process is depicted as a graphical model in Figure 1.

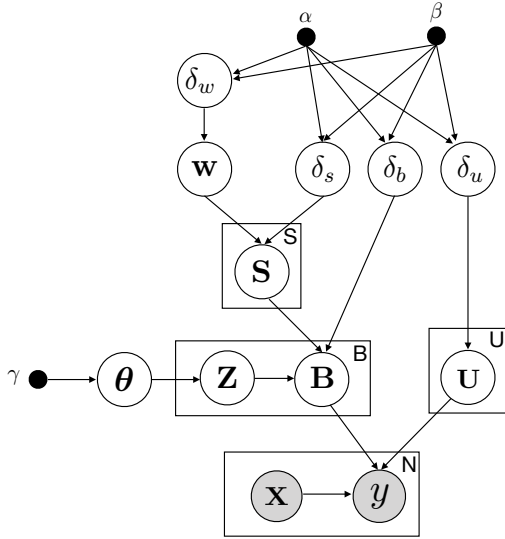


Figure 1: A graphical model representation of HBayes.

3.2 Probability Priors & Models

In this work, we model the probability of a single event t given $(\mathbf{X}_t, b_t, u_t, y_t)$ as

$$p(y_t | \mathbf{X}_t, \mathbf{B}_{b_t}, \mathbf{U}_{u_t}) = \left[\sigma(h_t) \right]^{y_t} \cdot \left[1 - \sigma(h_t) \right]^{1-y_t} \quad (1)$$

where $\sigma(\cdot)$ is a logistic function, i.e. $\sigma(x) = (1 + e^{-x})^{-1}$. $h_t = \mathbf{X}_t^T (\mathbf{B}_{b_t} + \mathbf{U}_{u_t})$. \mathbf{X}^T is the vector transpose of \mathbf{X} . \mathbf{U}_{u_t} represents user specific information encoded in HBayes for user u_t and \mathbf{B}_{b_t} denotes the brand b_t 's specific information.

As mentioned in Step 3 of the generative process of HBayes, each brand i 's style proportion distribution θ follows a Dirichlet distribution: $p(\theta) \sim \text{Dir}(\gamma)$, which is defined as follows:

$$\text{Dir}(\theta | \gamma) = \frac{\Gamma(\sum_{j=1}^S \gamma_j)}{\prod_{j=1}^S \Gamma(\gamma_j)} \prod_{j=1}^S \theta_j^{\gamma_j-1}$$

where γ is the S -dimensional Dirichlet hyper-parameter. We initialize γ_j by $\frac{1}{S}$.

Furthermore, in Step 4 of the generative process of HBayes, a brand is modeled as a random mixture over latent styles. Hence, we model the brand parameters by a mixture of multivariate Gaussian distribution defined as follows:

$$p(\mathbf{B}_i | z_i, S, \delta_b) = \prod_j p(\mathbf{B}_i | \mathbf{S}_j, \delta_b)^{\mathbb{I}(z_{i,j}=1)} = \prod_j \mathcal{N}(\mathbf{B}_i; \mathbf{S}_j, \delta_b^{-1} \mathbf{I})^{\mathbb{I}(z_{i,j}=1)}$$

where $z_{i,j}$ is the j th element of \mathbf{z}_i and $\mathbb{I}(\xi)$ is an indicator function that $\mathbb{I}(\xi) = 1$ if the statement ξ is true; $\mathbb{I}(\xi) = 0$ otherwise.

Therefore, the log joint likelihood of the dataset \mathcal{D} , latent variable \mathbf{Z} and the parameter Θ by given hyper-parameters $\mathcal{H} = \{\gamma, \alpha, \beta\}$ could be written as follows:

$$\begin{aligned} & \log(p(\mathcal{D}, \mathbf{Z}, \Theta | \mathcal{H})) \\ &= \sum_{t=1}^N \log p(y_t | \mathbf{X}_t, \mathbf{B}_{b_t}, \mathbf{U}_{u_t}) + \sum_{i=1}^B \log p(\mathbf{B}_i | z_i, S, \delta_b) \\ &+ \sum_{i=1}^B \log p(z_i | \theta) + \sum_{j=1}^S \log p(\mathbf{S}_j | \mathbf{w}, \delta_s) + \log p(\theta | \gamma) \\ &+ \sum_k^U \log p(\mathbf{U}_k | \delta_u) + \log p(\mathbf{w} | \delta_w) + \log p(\delta_w | \alpha, \beta) \\ &+ \log p(\delta_u | \alpha, \beta) + \log p(\delta_b | \alpha, \beta) + \log p(\delta_s | \alpha, \beta) \end{aligned} \quad (2)$$

We use Θ to denote all model parameters:

$$\Theta = \{\{\mathbf{U}_k\}, \{\mathbf{B}_i\}, \{\mathbf{S}_j\}, \mathbf{w}, \theta, \delta_u, \delta_b, \delta_s, \delta_w\},$$

where $k \in \{1, \dots, U\}$, $i \in \{1, \dots, B\}$, $j \in \{1, \dots, S\}$.

3.3 Optimization

Since both \mathbf{Z} and Θ defined by the HBayes are unobserved, we cannot learn our HBayes directly. Instead, we infer the expectations of these latent variables and compute the expected log likelihood of the log joint probability with respect to the latent variables distribution, i.e., Q function defined in eq.(3). In the following, we omit the explicit conditioning on \mathcal{H} for notational brevity.

$$Q = \int_{\Theta} \sum_{\mathbf{Z}} p(\mathbf{Z}, \Theta | \mathcal{D}) \log(p(\mathcal{D}, \mathbf{Z}, \Theta)) d\Theta \quad (3)$$

From the Bayes rule, we can see that the posteriors distribution of \mathbf{Z} and Θ can be represented by $p(\mathbf{Z}, \Theta | \mathcal{D}) = \frac{p(\mathcal{D}, \mathbf{Z}, \Theta)}{p(\mathcal{D})}$. However, this above distribution is intractable to compute in general [9]. To tackle this problem, a wide variety of approximate inference algorithms are developed, such as Laplace approximation [26], variational approximation [2], and Markov Chain Monte Carlo (MCMC) approach [3], etc.

In this work, we choose to solve this problem by using variational Bayes approximation [2]. More specifically, we approximate the original posterior distribution $p(\mathbf{Z}, \Theta | \mathcal{D})$ with a tractable distribution $q(\mathbf{Z}, \Theta)$ such that instead of maximizing the Q function defined in eq.(3), we maximize the variational free energy defined as

$$Q'(q) = \int_{\Theta} \sum_Z q(Z, \Theta) \log \frac{p(\mathcal{D}, Z, \Theta)}{q(Z, \Theta)} d\Theta \quad (4)$$

which is also equal to minimize the KL divergence of $p(Z, \Theta|\mathcal{D})$ and $q(Z, \Theta)$.

Here we choose to apply *Mean Field* approximation technique to approximate $p(Z, \Theta|\mathcal{D})$, where we assume independence among all different variables (Z and Θ) and define $q(Z, \Theta)$ as follows:

$$q(Z, \Theta) = q(Z) \cdot \prod_{k=1}^K q(\mathbf{U}_k) \cdot \prod_{j=1}^S q(\mathbf{S}_j) \cdot \prod_{i=1}^B q(\mathbf{B}_i) \cdot q(\mathbf{w}) \cdot q(\boldsymbol{\theta}) \cdot q(\delta_u) \cdot q(\delta_b) \cdot q(\delta_s) \cdot q(\delta_w) \quad (5)$$

where q denotes different distribution functions for notation brevity. Details of choices of different distributions will be discussed in Section 3.3.2.

3.3.1 Sigmoid Approximation. The Gaussian priors from our log joint probability (see eq.(2)) are not conjugate to the data likelihood due to the fact that our events are modeled by a sigmoid function (see eq.(1)). In order to conduct tractable inference on $Q'(q)$, we apply a variational lower bound approximation on eq.(1) that has the “squared exponential” form. Therefore, they are conjugate to the Gaussian priors.

$$\sigma(h_t) \geq \sigma(\xi_t) \exp \left\{ \frac{1}{2} (h_t - \xi_t) - \lambda_t (h_t^2 - \xi_t^2) \right\}$$

where $\lambda_t = \frac{1}{2\xi_t} [\sigma(\xi_t) - \frac{1}{2}]$ and ξ_t is a variational parameter. This lower bound is derived using the convex inequality. The similar problem was discussed in [14, 15].

Therefore, each event likelihood can be expressed as follows:

$$\begin{aligned} [\sigma(h_t)]^{y_t} \cdot [1 - \sigma(h_t)]^{1-y_t} &= \exp(y_t h_t) \sigma(-h_t) \\ &\geq \sigma(\xi_t) \exp(y_t h_t - \frac{1}{2} (h_t + \xi_t) - \lambda_t (h_t^2 - \xi_t^2)) \end{aligned} \quad (6)$$

By using the sigmoid approximation in eq.(6), our variational free energy $Q'(q)$ (eq.(4)) can be bounded as:

$$Q'(q) \geq Q'_\xi(q) = \int_{\Theta} \sum_Z q(Z, \Theta) \log \frac{p_\xi(\mathcal{D}, Z, \Theta)}{q(Z, \Theta)} d\Theta \quad (7)$$

In the following, we will maximize the lower bound of the variational free energy $Q'_\xi(q)$ for parameter estimation.

3.3.2 Parameter Estimation. We develop a Variational Bayes (VB) algorithm for HBayes parameter estimation, where in the E-step, we compute the expectation of the hidden variables Z and in the M-step, we try to find Θ that maximizes lower bound of the variational free energy $Q'_\xi(q)$ (eq.(7)). In the VB algorithm, we use coordinate ascent variational inference (CAVI) [2] to optimize $Q'_\xi(q)$. CAVI iteratively optimizes each factor of the mean field variational distribution, while holding the others fixed.

update expectation of Z : We assume each brand's style membership latent variable is independent and therefore, $q(Z) = \prod_{i=1}^B q(\mathbf{z}_i)$. For each \mathbf{z}_i , we parameterize $q(\mathbf{z}_i)$ as a multinomial distribution, i.e.,

$$q(\mathbf{z}_i) = \prod_{j=1}^S \mu_{i,j}^{\mathbb{I}(z_{i,j}=1)}$$

and the update rule for $\mu_{i,j}$ is

$$\mu_{i,j} = \frac{\rho_{i,j}}{\sum_{p=1}^S \rho_{i,p}} \quad (8)$$

$$\ln(\rho_{i,j}) = \mathbb{E}[\ln(\theta_j)] + \frac{d}{2} \mathbb{E}[\ln(\delta_b)] - \frac{d}{2} \ln(2\pi) - \frac{1}{2} \mathbb{E}[\delta_b(\mathbf{B}_i - \mathbf{S}_j)^T (\mathbf{B}_i - \mathbf{S}_j)] \quad (9)$$

where the expectation $\mathbb{E}[\cdot]$ is with respect to the (currently fixed) variational density over Θ i.e., $\mathbb{E}[\cdot] = \mathbb{E}_{-Z}[\cdot]$ in this part. Furthermore, in the following, we note that:

$$\mathbb{E}[z_{i,j}] = \mu_{i,j} \quad (10)$$

Parametrization and update rule of $q(\theta)$: For the style proportion distribution θ , we parameterize $q(\theta)$ as a Dirichlet distribution, i.e., $q(\theta) = \text{Dir}(\theta; \gamma)$, and the update rule for γ are

$$\gamma_j = \gamma_j + \sum_{i=1}^B \mu_{i,j}, j = 1, \dots, S \quad (11)$$

Parametrization and update rule of $q(\mathbf{U}_k)$: For each user k , $k = 1, \dots, U$, we parameterize $q(\mathbf{U}_k)$ as a multivariate normal distribution, i.e., $q(\mathbf{U}_k) = \mathcal{N}(\mathbf{U}_k; \boldsymbol{\mu}_k^u, \Sigma_k^u)$, and the update rule for $\boldsymbol{\mu}_k^u, \Sigma_k^u$ are

$$\Sigma_k^u = [\delta_u \mathbf{I} + \sum_{t=1}^N \mathbb{I}(u_t = k) 2\lambda_t \mathbf{X}_t \mathbf{X}_t^T]^{-1} \quad (12)$$

$$\boldsymbol{\mu}_k^u = \Sigma_k^u \left[\sum_{t=1}^N \mathbb{I}(u_t = k) (y_t - \frac{1}{2} - 2\lambda_t \mathbf{X}_t^T \mathbb{E}[\mathbf{B}_{b_t}]) \mathbf{X}_t \right] \quad (13)$$

Parametrization and update rule of $q(\mathbf{B}_i)$: For each brand i , $i = 1, \dots, B$, we parameterize $q(\mathbf{B}_i)$ as a multivariate normal distribution, i.e., $q(\mathbf{B}_i) = \mathcal{N}(\mathbf{B}_i; \boldsymbol{\mu}_i^b, \Sigma_i^b)$, and the update rule for $\boldsymbol{\mu}_i^b, \Sigma_i^b$ are

$$\Sigma_i^b = \left[\delta_b \sum_{j=1}^S \mu_{i,j} \mathbf{I} + \sum_{t=1}^N \mathbb{I}(b_t = i) 2\lambda_t \mathbf{X}_t \mathbf{X}_t^T \right]^{-1} \quad (14)$$

$$\boldsymbol{\mu}_i^b = \Sigma_i^b \left[\delta_b \sum_{j=1}^S \mu_{i,j} \mathbb{E}[\mathbf{S}_j] + \sum_{t=1}^N \mathbb{I}(b_t = i) (y_t - \frac{1}{2} - 2\lambda_t \mathbf{X}_t^T \mathbb{E}[\mathbf{U}_{u_t}]) \mathbf{X}_t \right] \quad (15)$$

Parametrization and update rule of $q(\mathbf{S}_j)$: For each style j , $j = 1, \dots, S$, we parameterize $q(\mathbf{S}_j)$ as a multivariate normal distribution, i.e., $q(\mathbf{S}_j) = \mathcal{N}(\mathbf{S}_j; \boldsymbol{\mu}_j^s, \Sigma_j^s)$, and the update rule for $\boldsymbol{\mu}_j^s, \Sigma_j^s$ are

$$\Sigma_j^s = [\delta_s + \delta_b \sum_{i=1}^B \mu_{i,j}]^{-1} \mathbf{I} \quad (16)$$

$$\boldsymbol{\mu}_j^s = \Sigma_j^s \left[\delta_s \mathbb{E}[\mathbf{w}] + \delta_b \sum_{i=1}^B \mu_{i,j} \mathbb{E}[\mathbf{B}_i] \right] \quad (17)$$

Parametrization and update rule of $q(\mathbf{w})$: For the mean variable of style prior \mathbf{w} , we parameterize $q(\mathbf{w})$ as a multivariate normal distribution, i.e., $q(\mathbf{w}) = \mathcal{N}(\mathbf{w}; \boldsymbol{\mu}^w, \Sigma^w)$, and the update rule for $\boldsymbol{\mu}^w$, Σ^w are

$$\Sigma^w = [\delta_w + \delta_s \cdot S]^{-1} \mathbf{I} \quad (18)$$

$$\boldsymbol{\mu}^w = \Sigma^w \left[\delta_s \sum_{j=1}^S \mathbb{E}[\mathbf{S}_j] \right] \quad (19)$$

Parametrization and update rule of $q(\delta_u)$, $q(\delta_b)$, $q(\delta_s)$ and $q(\delta_w)$: For all the precision parameters' distributions, we parameterize them as a Gamma distribution, i.e., $p(\delta_*) = \mathcal{G}(\delta_*; \alpha_*, \beta_*)$, where $\delta_* \in \{\delta_w, \delta_s, \delta_u, \delta_b\}$ and the update rule for α and β are

$$\alpha_u = \alpha + \frac{dU}{2}, \quad \beta = \beta + \frac{1}{2} \sum_{k=1}^U \mathbb{E}[\mathbf{U}_k^T \mathbf{U}_k] \quad (20)$$

$$\alpha_b = \alpha + \frac{dB}{2}, \quad \beta = \beta + \frac{1}{2} \sum_{i=1}^B \sum_{j=1}^S \mu_{i,j} \mathbb{E}[(\mathbf{B}_i - \mathbf{S}_j)^T (\mathbf{B}_i - \mathbf{S}_j)] \quad (21)$$

$$\alpha_s = \alpha + \frac{dS}{2}, \quad \beta = \beta + \frac{1}{2} \sum_{j=1}^S \mathbb{E}[(\mathbf{S}_j - \mathbf{w})^T (\mathbf{S}_j - \mathbf{w})] \quad (22)$$

$$\alpha_w = \alpha + \frac{d}{2}, \quad \beta = \beta + \frac{1}{2} \mathbb{E}[\mathbf{w}^T \mathbf{w}] \quad (23)$$

Update rule of ξ : For the variational parameters $\xi_t, t = 1, \dots, N$, in order to maximize $Q'_\xi(q)$ such that the bound on $Q'(q)$ is tight [2], the update rule is:

$$\xi_t = \sqrt{\mathbb{E}[(\mathbf{X}_t^T (\mathbf{B}_{b_t} + \mathbf{U}_{u_t}))^2]} \quad (24)$$

3.3.3 Summary. The parameter estimation method for the HBayes is summarized by Algorithm 1.

Algorithm 1 Parameter Estimation in HBayes

- 1: **INPUT:**
 - 2: Hyper-parameters \mathcal{H} : $\mathcal{H} = \{\alpha, \beta, \gamma\}$
 - 3: Data samples \mathcal{D} : $(\mathbf{X}_t, b_t, u_t, y_t), t = 1, \dots, N$
 - 4: **procedure** LEARNING HBayes
 - 5: **repeat**
 - 6: E-step: compute expectation of \mathbf{Z} by eq.(10).
 - 7: M-step: estimate $\{\mathbf{U}_k\}, \{\mathbf{B}_i\}, \{\mathbf{S}_j\}, \mathbf{w}, \boldsymbol{\theta}, \delta_u, \delta_b, \delta_s, \delta_w$ by eq.(11) - eq.(24).
 - 8: **until** Convergence
 - 9: **return** $\boldsymbol{\theta}$
-

3.4 Prediction

In the recommender system, the task is to generate top K products list for each user. So, given the user u^* , it's natural to expose top M products based on the probability of the positive outcome. For the m^{th} item, the probability is calculated as:

$$\begin{aligned} \hat{y}_m &= p(y_m = 1 | \mathbf{X}_m, \mathcal{D}, \mathcal{H}) \approx \int \sigma(h_m) q(\mathbf{Z}, \boldsymbol{\theta}) d\boldsymbol{\theta} \\ &= \int \sigma(h_m) \mathcal{N}(h_m | \mu_m, \sigma_m^2) dh_m \approx \sigma\left(\frac{\mu_m}{\sqrt{1 + \pi \sigma_m^2 / 8}}\right) \end{aligned}$$

where h_m is a random variable with Gaussian distribution:

$$\begin{aligned} h_m &= \mathbf{X}_m^T (\mathbf{B}_{b_m} + \mathbf{U}_{u^*}) \sim \mathcal{N}(h_m; \mu_m, \sigma_m^2) \\ \mu_m &= \mathbb{E}[\mathbf{X}_m^T (\mathbf{B}_{b_m} + \mathbf{U}_{u^*})] \\ \sigma_m^2 &= \mathbb{E}[(\mathbf{X}_m^T (\mathbf{B}_{b_m} + \mathbf{U}_{u^*}) - \mu_m)^2] \end{aligned}$$

4 EXPERIMENT

In this section, we conduct several experiments on two data sets: (1) our case study: the real-world e-commerce apparel data set; (2) the publicly available music data set. For both data sets, we compare HBayes against several other *state-of-the-art* recommendation approaches which are briefly mentioned as follows:

HSR [31] is an item-based recommendation approach that employs a special non-negative matrix factorization for exploring the implicit hierarchical structure of users and items so the user preference towards certain products is better understood.

HPF [12] generates a hierarchical *Poisson* factorization model for better modeling users' rating towards certain items based upon each one's latent preferences. Unlike proposed HBayes, HPF does not leverage the entity content feature for constructing the hierarchical structure.

SVD++ [16, 21] combines the collaborative filtering and latent factor approaches, so to provide the more accurate neighboring based recommendation results.

CoClustering [11] is a collaborative filtering approach based on weighted co-clustering improvements that simultaneously cluster users and items.

Factorization Machine (FM) [22, 23] combines support vector machines (SVM) with factorization models. It takes the advantage of SVM meanwhile overcomes the feature sparsity issues. In this paper, we adopt the LibFM implementation mentioned in [23] specifically as another baseline.

LambdaMART [5] is the boosted tree version of LambdaRank [10], which is based on RankNet [4]. LambdaMART proves to be a very successful approach for ranking as well as recommendation.

4.1 Evaluation Metrics

Throughout the experiments, we compare HBayes against other baselines on the testing held-out dataset under the 5-folds cross-validation settings. For each fold, after fitting the model on the training set, we rank on the test set by each model, and generate the top M samples with maximal ranking scores for recommendation. Regarding metrics, we adopt precisions, recalls as well as F1-scores for evaluating the item retrieval quality and normalized discounted information gain (NDCG) for evaluating the recommendation ranking quality. More specifically, each metric mentioned is defined as:

$$\begin{aligned}
\text{Precision@K} &= \frac{\# \text{ of products clicked in top } K}{K} \\
\text{Recall@K} &= \frac{\# \text{ of products clicked in top } K}{\# \text{ of products the user clicked}} \\
\text{F1-score@K} &= 2 \cdot \frac{\text{Precision@K} \cdot \text{Recall@K}}{\text{Precision@K} + \text{Recall@K}}
\end{aligned}$$

discounted cumulative gain (DCG) measures the ranking quality based on the item positions from the resulting list, which is defined as:

$$\text{DCG@K} = \sum_{i=1}^K \frac{r_i}{\log_2(i+1)} = r_1 + \sum_{i=2}^K \frac{r_i}{\log_2(i+1)}$$

re-rank all potential products to produce the ideal maximal possible DCG through out first K items, and its corresponding DCG score is called the ideal discounted cumulative gain (IDCG). By using both DCG and IDCG, the normalized discounted cumulative gain (NDCG) is defined as:

$$\text{NDCG@K} = \frac{\text{DCG@K}}{\text{IDCG@K}}$$

4.2 Recommendation on Apparel Data

The first apparel data set is collected from a large e-commerce company. In this dataset, each sample represents a particular apparel product which is recorded by various features including: categories, titles, and other properties, etc. Meanwhile, the user click information is also recorded and translated into data labels. Throughout the experiment, positive labels indicate that certain recommended products are clicked by the user, whereas negative samples indicate that the recommended products are skipped by the user which usually implies that the user is 'lack of interest' towards that certain item. By data cleaning and preprocessing: (1) merging duplicated histories; (2) removing users of too few records, the post-processed data set ends up with 895 users, 81223 products, 5535 brands with 380595 uniquely observed user-item pairs. In average, each user has 425 products records, ranging from 105 to 2048, and 61.2% of the users have fewer than 425 product clicking records. For each item, we encode the popularity and category features into a 20 dimensional feature vector; title and product property into a 50 dimensional feature vector. Combining with other features, the total dimension of each sample ends up with 140.

4.2.1 Feature Analysis. The apparel data are composed of four types of features: (1) product popularity; (2) product category; (3) product title; (4) product properties. We briefly explain each feature's physical meaning and how we process the data as follows:

Product Popularity (POP): product popularity is a measure of the prevalence of certain items in the dataset. In general, customers have preferences for a particular product during a period of time. This phenomenon is pretty common for apparel products (e.g. apparels' popularity in certain styles may be affected by the public, especially by those important public figures). For a particular product i , the popularity is defined as: $\text{POP}_i := \frac{n_{x_i}}{N_x}$, where n_{x_i} are the number of orders or the contribution of gross merchandise volume (GMV) for product i , and $N_x = \sum_{\forall x_i} n_{x_i}$ is the summation of n_{x_i}

across all products in the dataset.

Product Category (CID): In e-commerce, items are clustered into different groups based on the alliance/ similarity of the item functionalities and utilities. In e-commerce websites, such an explicit hierarchy usually exists. We encode each item's category into a high dimensional vector via one-hot encoding and adjust the feature weights by the popularity of such category.

Product Title (TITLE): product titles are created by vendors and they are typically in the forms of natural languages indicating the item functionality and utility. Examples could be like 'INMAN short sleeve round neck triple color block stripe T-shirt 2017'. We preprocess the product titles by generating the sentence embedding based on [7]. The main idea is to average the wording weights in the title sentence based on the inverse document frequency (IDF) value of each individual word involved.

Product Property Features (PROP): other product metadata features are also provided in the apparel dataset. For instance, the color feature of items takes values such like: 'black', 'white', 'red', etc, and the sizing feature takes values such like: 'S', 'M', 'L', 'XL', etc. Similar to category features, product properties are first encoded into binary vectors $\mathbf{x}_i \in \{0, 1\}^{|N|}$, where N denotes the set of all possible product values. Then the property binary vectors are hashed into fixed length (50) vectors.

On one hand, by utilizing more features HBayes in general reaches better performance in terms of precision-recall metrics. We report PR-AUC in the table (1) to prove this argument; on the other hand, more features need more training time. Figure (8) reports the changes in likelihood against training time cost (in minutes) for each feature using cases¹. It is shown that by only taking POP features, model spends less than 5 minutes to converge while taking POP+CID+TITLE+PROP features, the model needs more than 4 hours in training.

Features	PR AUC
POP	0.0406
POP+CID	0.0414
POP+CID+TITLE	0.0489
POP+CID+TITLE+PROP	0.0491

Table 1: Model performance under different feature combinations in terms of PR AUC

4.2.2 Performance Comparison. We first report the model performance regarding precision, recall, as well as F1-score for HBayes against other baselines in Figure (2,3,4). As shown, when each method recommends fewer number of products ($K = 5$), HBayes does not show the superiority regarding with recalls, with the increment of recommended items, the recall for HBayes becomes much better against others, which implies HBayes is really efficient in terms of finding items that people tend to take interest in. In the

¹The experiment is conducted via the same Linux Qual core 2.8 GHz Intel Core i7 MacBook with 16 Gigabytes of memory

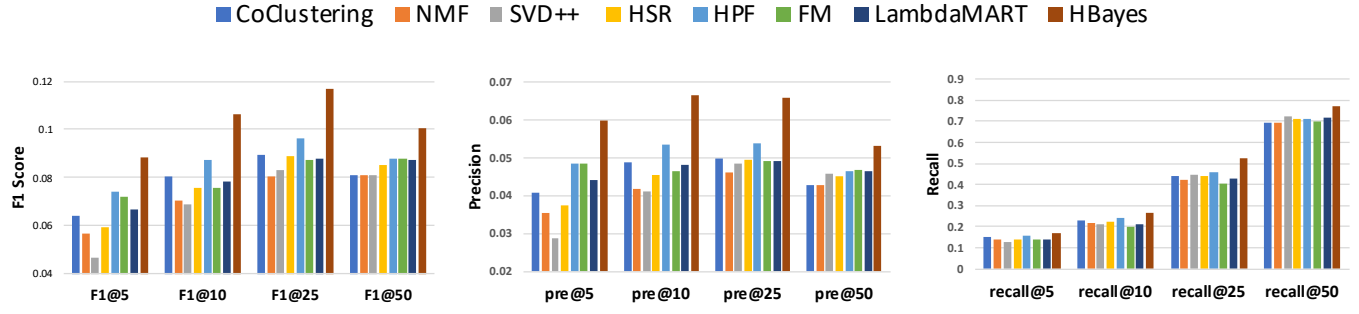


Figure 2: F1@K on Apparel data.

Figure 3: Precision@K on Apparel data.

Figure 4: Recall@K on Apparel data.

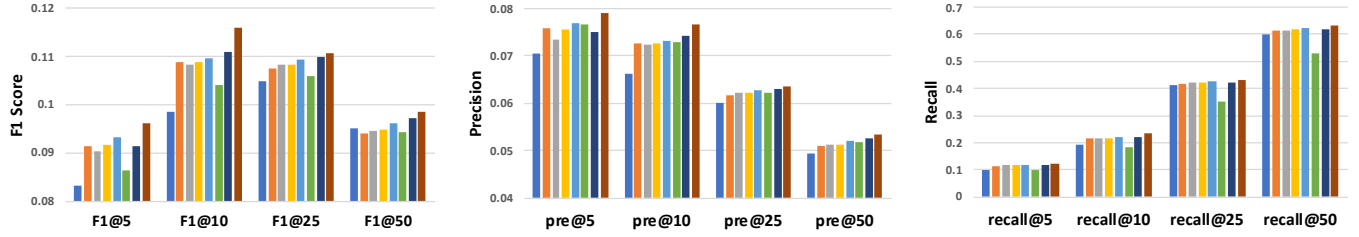


Figure 5: F1@K on Music data.

Figure 6: Precision@K on Music data.

Figure 7: Recall@K on Music data.

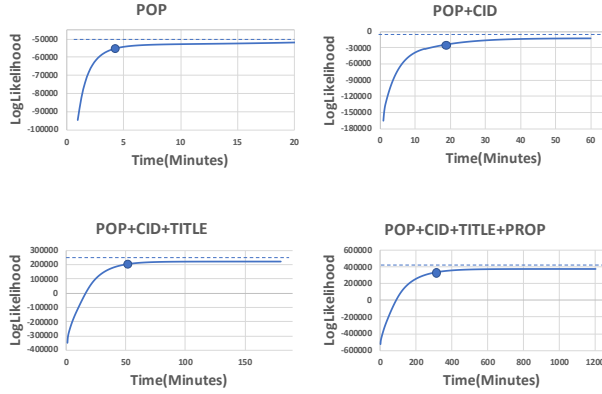


Figure 8: Training time under different feature combinations on e-commerce recommendations

Method	NDCG@5	NDCG@10	NDCG@25	NDCG@50
CoClustering	0.1288	0.1637	0.2365	0.3050
NMF	0.1249	0.0156	0.2272	0.3020
SVD++	0.1138	0.1487	0.2287	0.3073
HSR	0.1266	0.1603	0.2354	0.3107
HPF	0.1412	0.1757	0.2503	0.3229
FM	0.1363	0.1592	0.2291	0.3117
LambdaMART	0.1287	0.1585	0.2304	0.3123
HBayes	0.1557	0.1974	0.2871	0.3590

Table 2: NDCG on apparel recommendations

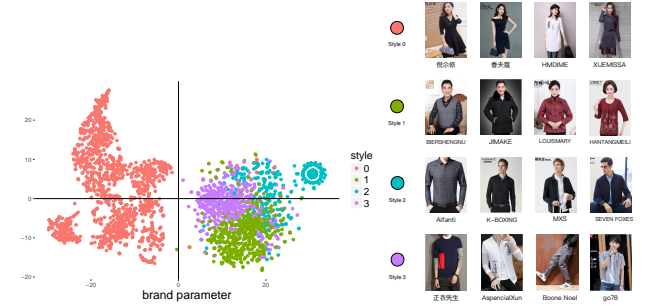


Figure 9: tSNE of four latent apparel style clusters

sense of precision, HBayes is consistently better than other baseline methods which implies HBayes is much more accurate in terms of item classification under different K. Given the performance of precisions and recalls, HBayes is much better regarding F1-score with different K for apparel recommendation.

Regarding the ranking quality, we use NDCG to report each method's performance in Table (2). HBayes is superior against other baseline methods through out different K recommended. Specially, HBayes beats the second best HPF at K = 5 by 10.3%, at K = 10 by 12.4%, at K = 25 by 14.7% and at K = 50 by 11.2%.

4.2.3 Model Learning Analysis. Like mentioned in Sec.3, HBayes learns the latent style clusters, and group different brands of products based on their different hidden style representations. Figure (9) shows the tSNE [20] representations of different apparel clusters learned by HBayes and we randomly pick 4 samples out of each cluster and display each product image at the right subfigure. As shown, cluster one which takes the majority proportion of apparel

items seems about stylish female youth garment. This intuitively makes sense because the majority of apparel customers are female youth for e-commerce websites; as a result, most apparels are focusing on the young female audience as well. The second cluster seems about senior customers who are elder in age. Interestingly, the third cluster and the fourth cluster that are closely tied up are both about male customer youth. However, the third cluster seems focusing more on office business garment while the fourth cluster seems more about Korean-pop street styles. This indicates us that the HBayes indeed learns the meaningful intrinsic garment styles from apparel items by leveraging customer behavior data.

4.3 Recommendation on Last.fm Music Data

The second data set is collected from Last.fm dataset [6] and Free Music Archive (FMA) [8]. Last.fm is a publicly available dataset which contains the whole listening habits (till May, 5th 2009) for 1000 users. FMA is an open and easily accessible dataset providing 917 GiB and 343 days of Creative Commons-licensed audio from 106574 tracks, 16341 artists and 14854 albums, arranged in a hierarchical taxonomy of 161 genres. It also provides full-length and high-quality audios with precomputed features. In our experiment, tracks in Last.fm dataset were further intersected with FMA dataset for better feature generation. The resulting dataset contains 500 users, 16328 tracks and 36 genres.

4.3.1 Performance Comparison. We conduct similar experiments as we do for apparel dataset and report precisions, recalls as well as F1-scores in Figure (5,6,7). Although HPF and HBayes share similar performance regarding recalls along with different K, HBayes is dominant for precisions at different K, especially when K is small (5, 10), which indicates HBayes is very efficient and precise for helping users pick up the songs they prefer even when the recommended item lists are short. Combining the two, HBayes is superior in terms of F1-scores for different K recommended.

For ranking qualities, we report the NDCG performance in Table (3). Similar as e-commerce apparel data, HBayes is the best approach and HPF is the second best one throughout different K items recommended. Specifically, HBayes beats HPF for 2.7% at K = 5, 4.3% at K = 10, 2.1% at K = 25, and 2.5% at K = 50 separately.

Method	NDCG@5	NDCG@10	NDCG@25	NDCG@50
CoClustering	0.2415	0.2314	0.2289	0.2349
NMF	0.2556	0.2494	0.2368	0.2431
SVD++	0.2493	0.2478	0.2381	0.2439
HSR	0.2544	0.2495	0.2384	0.2448
HPF	0.2584	0.2513	0.2405	0.2474
FM	0.2527	0.2453	0.2284	0.2333
LambdaMART	0.2372	0.2337	0.2272	0.2218
HBayes	0.2655	0.2620	0.2455	0.2537

Table 3: NDCG on Last.fm recommendations

5 CONCLUSION

In this paper, we propose a novel generalized learning framework that learns both the entity hierarchical structure and its latent factors as well as build a personalized Bayesian recommender system, *HBayes*. By utilizing variational Bayesian inference approach,

HBayes is able to converge efficiently with few iterations. In empirical studies, we walk through a practical case study of e-commerce apparel data and another publicly available music recommendation data. Experiment results show that HBayes beats other *state-of-the-art* recommendation approaches in the sense of precisions, recalls, F1-score, as well as NDCG metrics, due to the fact that HBayes is able to extract the data characteristics as well as capture user hidden preferences.

REFERENCES

- [1] Shay Ben-Elazar, Gal Lavee, Noam Koenigstein, Oren Barkan, Hilik Berezin, Ulrich Paquet, and Tal Zaccai. 2017. Groove Radio: A Bayesian Hierarchical Model for Personalized Playlist Generation. In *Proceedings of the Tenth ACM International Conference on Web Search and Data Mining (WSDM '17)*. ACM, New York, NY, USA, 445–453.
- [2] Christopher M Bishop. 2006. Pattern recognition and machine learning (information science and statistics) springer-verlag new york. Inc. Secaucus, NJ, USA (2006).
- [3] David M Blei, Andrew Y Ng, and Michael I Jordan. 2003. Latent dirichlet allocation. *Journal of machine Learning research* 3, Jan (2003), 993–1022.
- [4] Chris Burges, Tal Shaked, Erin Renshaw, Ari Lazier, Matt Deeds, Nicole Hamilton, and Greg Hullender. 2005. Learning to rank using gradient descent. In *Proceedings of the 22nd international conference on Machine learning*. ACM, 89–96.
- [5] Christopher JC Burges. 2010. From ranknet to lambdarank to lambdamart: An overview. *Learning* 11, 23-581 (2010), 81.
- [6] O. Celma. 2010. *Music Recommendation and Discovery in the Long Tail*. Springer.
- [7] Cedric De Boom, Steven Van Canneyt, Thomas Demeester, and Bart Dhoedt. 2016. Representation learning for very short texts using weighted word embedding aggregation. *Pattern Recognition Letters* 80 (2016), 150–156.
- [8] Michaël Defferrard, Kirell Benzi, Pierre Vandergheynst, and Xavier Bresson. 2017. FMA: A Dataset for Music Analysis. In *18th International Society for Music Information Retrieval Conference*.
- [9] James M Dickey. 1983. Multiple hypergeometric functions: Probabilistic interpretations and statistical uses. *J. Amer. Statist. Assoc.* 78, 383 (1983), 628–637.
- [10] Pinar Donmez, Krysta M Svore, and Christopher JC Burges. 2009. On the local optimality of LambdaRank. In *Proceedings of the 32nd international ACM SIGIR conference on Research and development in information retrieval*. ACM, 460–467.
- [11] Thomas George and Srujana Merugu. 2005. A scalable collaborative filtering framework based on co-clustering. In *Data Mining, Fifth IEEE international conference on*. IEEE, 4–pp.
- [12] Prem Gopalan, Jake M Hofman, and David M Blei. 2015. Scalable Recommendation with Hierarchical Poisson Factorization. In *UAI*. 326–335.
- [13] San Gultekin and John Paisley. 2014. A Collaborative Kalman Filter for Time-Evolving Dyadic Processes. In *Proceedings of the 2014 IEEE International Conference on Data Mining (ICDM '14)*. IEEE Computer Society, Washington, DC, USA, 140–149.
- [14] T Jaakkola and M Jordan. 1997. A variational approach to Bayesian logistic regression models and their extensions. In *Sixth International Workshop on Artificial Intelligence and Statistics*, Vol. 82. 4.
- [15] Michael I Jordan, Zoubin Ghahramani, Tommi S Jaakkola, and Lawrence K Saul. 1999. An introduction to variational methods for graphical models. *Machine learning* 37, 2 (1999), 183–233.
- [16] Yehuda Koren. 2008. Factorization meets the neighborhood: a multifaceted collaborative filtering model. In *Proceedings of the 14th ACM SIGKDD international conference on Knowledge discovery and data mining*. ACM, 426–434.
- [17] Yehuda Koren. 2010. Collaborative Filtering with Temporal Dynamics. *Commun. ACM* 53, 4 (April 2010), 89–97.
- [18] Yiqun Liu, Junwei Miao, Min Zhang, Shaoping Ma, and Liyun Ru. 2011. How do users describe their information need: Query recommendation based on snippet click model. *Expert Systems with Applications* 38, 11 (2011), 13847 – 13856.
- [19] P. Lops, M. de Gemmis, and G. Semeraro. 2011. *Content-based Recommender Systems: State of the Art and Trends*. 73.
- [20] Laurens van der Maaten and Geoffrey Hinton. 2008. Visualizing data using t-SNE. *Journal of machine learning research* 9, Nov (2008), 2579–2605.
- [21] Andriy Mnih and Ruslan R Salakhutdinov. 2008. Probabilistic matrix factorization. In *Advances in neural information processing systems*. 1257–1264.
- [22] Steffen Rendle. 2010. Factorization machines. In *Data Mining (ICDM), 2010 IEEE 10th International Conference on*. IEEE, 995–1000.
- [23] Steffen Rendle. 2012. Factorization machines with libfm. *ACM Transactions on Intelligent Systems and Technology (TIST)* 3, 3 (2012), 57.
- [24] Steffen Rendle, Christoph Freudenthaler, Zeno Gantner, and Lars Schmidt-Thieme. 2009. BPR: Bayesian personalized ranking from implicit feedback. In *Proceedings of the twenty-fifth conference on uncertainty in artificial intelligence*. AUAI Press, 452–461.

- [25] Steffen Rendle, Christoph Freudenthaler, and Lars Schmidt-Thieme. 2010. Factorizing Personalized Markov Chains for Next-basket Recommendation. In *Proceedings of the 19th International Conference on World Wide Web (WWW '10)*. ACM, New York, NY, USA, 811–820.
- [26] Håvard Rue, Sara Martino, and Nicolas Chopin. 2009. Approximate Bayesian inference for latent Gaussian models by using integrated nested Laplace approximations. *Journal of the royal statistical society: Series b (statistical methodology)* 71, 2 (2009), 319–392.
- [27] Badrul Sarwar, George Karypis, Joseph Konstan, and John Riedl. 2001. Item-based Collaborative Filtering Recommendation Algorithms. In *Proceedings of the 10th International Conference on World Wide Web (WWW '01)*. ACM, New York, NY, USA, 285–295.
- [28] Andriy Shepitsen, Jonathan Gemmell, Bamshad Mobasher, and Robin Burke. 2008. Personalized recommendation in social tagging systems using hierarchical clustering. In *Proceedings of the 2008 ACM conference on Recommender systems*. ACM, 259–266.
- [29] Xiaoyuan Su and Taghi M. Khoshgoftaar. 2009. A Survey of Collaborative Filtering Techniques. *Adv. in Artif. Intell.* 2009, Article 4 (Jan. 2009), 1 pages.
- [30] Jiliang Tang, Xia Hu, and Huan Liu. 2013. Social recommendation: a review. *Social Network Analysis and Mining* 3, 4 (1 2013), 1113–1133. DOI:<http://dx.doi.org/10.1007/s13278-013-0141-9>
- [31] Suhan Wang, Jiliang Tang, Yilin Wang, and Huan Liu. 2015. Exploring Implicit Hierarchical Structures for Recommender Systems.. In *IJCAI*. 1813–1819.
- [32] Suhan Wang, Jiliang Tang, Yilin Wang, and Huan Liu. 2018. Exploring Hierarchical Structures for Recommender Systems. *IEEE Transactions on Knowledge and Data Engineering* (2018).
- [33] Dawei Yin, Liangjie Hong, Zhenzhen Xue, and Brian D. Davison. 2011. Temporal Dynamics of User Interests in Tagging Systems. In *Proceedings of the Twenty-Fifth AAAI Conference on Artificial Intelligence (AAAI'11)*. AAAI Press, 1279–1285.
- [34] Quan Yuan, Gao Cong, Kaiqi Zhao, Zongyang Ma, and Aixin Sun. 2015. Who, Where, When, and What: A Nonparametric Bayesian Approach to Context-aware Recommendation and Search for Twitter Users. *ACM Trans. Inf. Syst.* 33, 1, Article 2 (Feb. 2015), 33 pages.
- [35] Yi Zhang and Jonathan Koren. 2007. Efficient bayesian hierarchical user modeling for recommendation system. In *Proceedings of the 30th annual international ACM SIGIR conference on Research and development in information retrieval*. ACM, 47–54.