

(3) Let  $X_1, \dots, X_n$  be independent Poisson random variables with mean 1, and  $X = \sum_{i=1}^n X_i$ .

(a) Use the Markov's inequality to obtain a bound on

$$\mathbb{P}r \left\{ \sum_{i=1}^n X_i \geq \frac{3}{2}n \right\}.$$

(b) Use Chebyshev's inequality to obtain a bound on the same probability in part (a).

(a)  $X_i \sim \text{Po}(1)$

$X_1 \sim \text{Po}(\lambda_1)$   
 $X_2 \sim \text{Po}(\lambda_2)$  indep  $\Rightarrow X_1 + X_2 \sim \text{Po}(\lambda_1 + \lambda_2)$

Let  $X = \sum X_i \sim \text{Po}(n)$

$$P\left(\sum X_i \geq \frac{3}{2}n\right) \leq \frac{E(\sum X_i)}{\frac{3}{2}n} = \frac{n}{\frac{3}{2}n} = \frac{2}{3}$$

Markov's Inequality

$X$  nonnegative RV.

$k > 0$ .

$$P(X > k) \leq \frac{E(X)}{k}$$

Chebyshev's

$X$  is RV

$\text{Var}(X) < \infty$

$k > 0$

$$P(|X - E(X)| > k)$$

$$\leq \frac{\text{Var}(X)}{k^2}$$

(b)  $P\left(\sum X_i \geq \frac{3}{2}n\right) = P\left(X - n \geq \frac{n}{2}\right)$

$$\leq P\left(|X - n| \geq \frac{n}{2}\right)$$

$$\leq \frac{\text{Var}(X)}{\left(\frac{n}{2}\right)^2} = \frac{n}{\left(\frac{n}{2}\right)^2} = \frac{4}{n}$$

if  $A \subseteq B$

$\Downarrow$

$P(A) \leq P(B)$

(4) Flip a fair coin 100 times. By using Central Limit Theorem, estimate the probability of more than 55 heads.

CLT: If  $X_1, \dots, X_n$  i.i.d. RV ← identically & independently distributed.

$$\text{and } \mu = E(X_1), \quad \sigma^2 = \text{Var}(X_1)$$

Then for  $n > 30$  (recommended)

$$\sum_{i=1}^n X_i \sim N(n\mu, n\sigma^2) \quad \text{approximation}$$

Let  $X_i \sim \text{Bin}(1, \frac{1}{2})$  be the RV, where  $X_i = 1$  if  $i^{\text{th}}$  flip is a head.  
 $X_i = 0$  otherwise.

$$\text{Find } P(\sum X_i > 55)$$

$$\mu = E(X_1) = \frac{1}{2}$$

$$\sigma^2 = \text{Var}(X_1) = \frac{1}{4}$$

$$\sum X_i \sim N\left(\frac{1}{2}n, \frac{1}{4}n\right) \quad n=100$$

$$\approx P(\sum X_i \geq 55.5)$$

$$= P\left(Z \geq \frac{55.5 - 50}{\sqrt{25}}\right) = \underline{\hspace{2cm}}$$

(5) If  $X$  is a gamma random variable with parameters  $(n, 1)$ , approximately how large need  $n$  be so that

$$\Pr \left\{ \left| \frac{X}{n} - 1 \right| > 0.01 \right\} \leq 0.01?$$

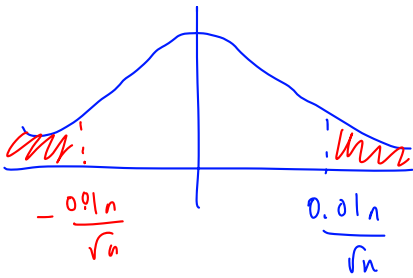
Fact :  $X_1 \sim \text{Gamma}(\alpha_1, \beta)$   
 $X_2 \sim \text{Gamma}(\alpha_2, \beta)$  ind  $\Leftrightarrow X_1 + X_2 \sim \text{Gamma}(\alpha_1 + \alpha_2, \beta)$

Let  $X_i \sim \text{Gamma}(1, 1)$  for  $i=1, \dots, n$  be independent RV.

$$E(X_i) = 1, \text{Var}(X_i) = 1$$

$\Rightarrow \sum X_i \sim \text{Gamma}(n, 1)$  can be approx by  $N(n \cdot 1, n \cdot 1)$  ↗  $\mu$  ↘  $\sigma^2$

$$\begin{aligned} 0.01 &\geq P\left(\left|\frac{X}{n} - 1\right| > 0.01\right) = P(|X - n| > 0.01n) \\ &= P(|\sum X_i - n| > 0.01n) \end{aligned}$$



$$\begin{aligned} &\approx P\left(|Z| > \frac{0.01n}{\sqrt{n}}\right) \\ &= 2 \left[ 1 - \Phi\left(\frac{0.01n}{\sqrt{n}}\right) \right] \end{aligned}$$

$$0.01 \geq 2 \left[ 1 - \Phi(0.01\sqrt{n}) \right]$$

$$0.005 \geq 1 - \Phi(0.01\sqrt{n})$$

$$\Phi(0.01\sqrt{n}) \geq 0.995$$

$$0.01\sqrt{n} \geq \Phi^{-1}(0.995)$$

$$n \geq \frac{1}{0.01} \left[ \Phi^{-1}(0.995) \right]^2 = \underline{\hspace{2cm}}$$

- (6) Civil engineers believe that  $W$ , the amount of weight (in units of 1000 pounds) that a certain span of a bridge can withstand without structural damage resulting, is normally distributed with mean 400 and standard deviation 40. Suppose that the weight (again, in units of 1000 pounds) of a car is a random variable with mean 3 and standard deviation 0.3. Approximately how many cars would have to be on the bridge span for the probability of structural damage to exceed 0.1?

$$W \sim N(400, 40^2)$$

Let  $X_i$  be the RV for the weight of car  $i$ .

$$E(X_i) = 3, \text{Var}(X_i) = 0.09$$

Find  $n$ , such that  $P\left(\sum_{i=1}^n X_i > W\right) > 0.1$

By CLT,  $X = \sum X_i \sim N(3n, 0.09n)$

$$0.1 < P(X - W > 0)$$

$$\approx P\left(Z > \frac{0 - (400 - 3n)}{\sqrt{40^2 + 0.09n}}\right)$$

$$\Phi\left(\frac{400 - 3n}{\sqrt{40^2 + 0.09n}}\right) > 0.1$$

$$\frac{400 - 3n}{\sqrt{40^2 + 0.09n}} > \Phi^{-1}(0.1) \Rightarrow \underline{\text{solve for } n.}$$

$$X_1 \sim N(\mu_1, \sigma_1^2)$$

$$X_2 \sim N(\mu_2, \sigma_2^2)$$

$$X_1 \pm X_2$$

$$\sim N(\mu_1 \pm \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$X - W \sim N(400 - 3n, 40^2 + 0.09n)$$

## Summary

- ① Probability : Bayes Rule  
Law of Total Probability
- ② Discrete RV : Bernoulli, Binomial, Geometric, Poisson.
- ③ Continuous RV : Normal, Uniform, Exponential, Gamma.

### $E(X)$ , $Var(X)$

④ Joint RV :  $f_X(x) = \int f(x,y) dy$   
 $f_Y(y) = \int f(x,y) dx$   
 $\iint_D f(x,y) dy dx = 1$

$f(x,y) = f_X \cdot f_Y$   
 $\Downarrow$   
independence.

$$Cov(X,Y) = E(XY) - E(X)E(Y).$$

$$\text{If } X, Y \text{ indep} \Rightarrow E(XY) = E(X)E(Y) \Rightarrow Cov(X,Y) = 0$$

Converse is FALSE.

$$\rho(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}} \quad -1 < \rho < 1$$

$$\text{If } X_i \text{ indep} \Rightarrow Var(\sum X_i) = \sum Var(X_i) \quad (\text{False in general})$$

$$E(\sum X_i) = \sum E(X_i) \quad (\text{always true})$$

⑤ Law of Total Expectation .

$$E(X) = \sum_i E(X|A_i) P(A_i)$$

$$E(X) = E_Y(\underbrace{E(X|Y)}_{\text{in terms of } Y})$$

in terms of  $Y$ .

Conditional Variance .

$$\text{Var}(Y|X=x) = E(Y^2|X=x) - [E(Y|X=x)]^2$$

Law of Total Variance

$$\text{Var}(Y) = \text{Var}(E(Y|X)) + E[\text{Var}(Y|X)]$$

⑥ Markov Ineq, Chebyshev Ineq, CLT

Normal Approximation, Poisson Approx to Binomial