

(1) The joint density of X and Y is given by

$$f(x, y) = \begin{cases} xe^{-(x+y)}, & x > 0, y > 0 \\ 0, & \text{otherwise.} \end{cases}$$

Are X and Y independent? If, instead, $f(x, y)$ were given by

$$f(x, y) = \begin{cases} 2, & 0 < x < y, 0 < y < 1; \\ 0, & \text{otherwise,} \end{cases}$$

would X and Y be independent?

marginal PDF

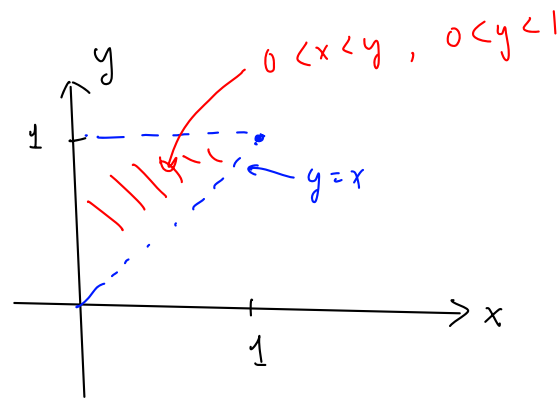
$$X, Y \text{ are independent} \Leftrightarrow f_{X,Y}(x, y) = f_X(x) f_Y(y)$$

$$f_X(x) = \int_0^{\infty} f(x, y) dy = \int_0^{\infty} x e^{-(x+y)} dy = x e^{-x} \int_0^{\infty} e^{-y} dy = x e^{-x} [-e^{-y}]_0^{\infty} = x e^{-x} \quad (x > 0)$$

$$f_Y(y) = \int_0^{\infty} f(x, y) dx = \int_0^{\infty} x e^{-(x+y)} dx = e^{-y} \int_0^{\infty} \underbrace{x e^{-x}}_1 dx = e^{-y} \left\{ [-x e^{-x}]_0^{\infty} + \int_0^{\infty} e^{-x} dx \right\} = e^{-y} [0 + 1] = e^{-y} \quad (y > 0)$$

Since $f(x, y) = f_X(x) f_Y(y)$ for $x > 0, y > 0 \Rightarrow X, Y$ are indep.

$$f(x, y) = \begin{cases} 2 & 0 < x < y, 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$$



$$f_X(x) = \int_x^1 f(x, y) dy = \int_x^1 2 dy = 2(1-x) \quad 0 < x < 1 \quad \text{check} \quad \int_0^1 2(1-x) dx = 1$$

$$f_Y(y) = \int_0^y f(x, y) dx = \int_0^y 2 dx = 2y \quad 0 < y < 1 \quad \text{check} \quad \int_0^1 2y dy = 1$$

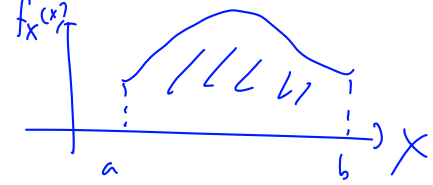
Since $f(x, y) \neq f_X(x) f_Y(y) \Rightarrow X, Y$ are not independent
on $0 < x < y, 0 < y < 1$

(2) Let

$$f(x, y) = 24xy, \quad 0 \leq x \leq 1, 0 \leq y \leq 1-x$$

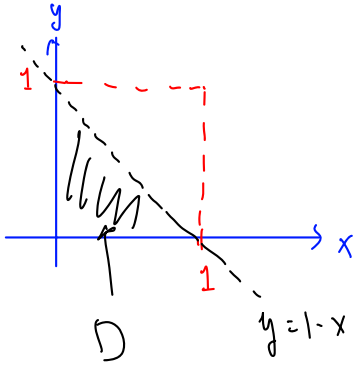
and let it be 0 otherwise.

- (a) Show that $f(x, y)$ is a joint density function.
- (b) Find $\mathbb{E}(X)$ and $\mathbb{E}(Y)$.
- (c) Show that X and Y are not independent.



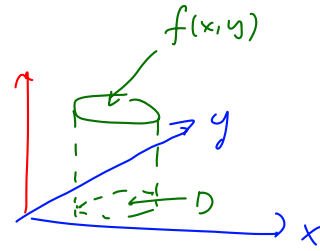
$$\int_a^b f_X(x) dx = 1$$

(a) Show $\iint_D (24xy) dA = 1$



$$\begin{aligned} \int_0^1 \int_0^{1-x} (24xy) dy dx \\ &= \int_0^1 24x \left[\frac{1}{2} y^2 \right]_{y=0}^{y=1-x} dx \\ &= \int_0^1 24x \left[\frac{1}{2} (1-x)^2 \right] dx \end{aligned}$$

$$= 12 \int_0^1 x(1-x)^2 dx = \dots = 1$$



$$\iint_D f(x, y) dA = 1$$

(b) $\mathbb{E}(X) = \int_0^1 x f_X(x) dx$

$$= \int_0^1 x \cdot 12x(1-x)^2 dx$$

$$= \frac{2}{5}$$

$$f_X(x) = \int_0^{1-x} 24xy dy = 24x \left[\frac{1}{2} y^2 \right]_0^{1-x} = 12x(1-x)^2$$

$$f_Y(y) = \int_0^{1-y} 24xy dx = 12y(1-y)^2$$

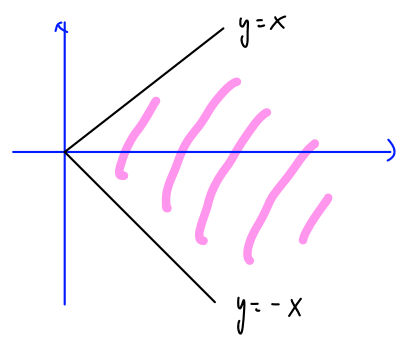
$$\mathbb{E}(Y) = \int_0^1 y f_Y(y) dy = \int_0^1 y \cdot 12y(1-y)^2 dy = \frac{2}{5}$$

(c) $f_{X,Y}(x,y) \neq f_X(x) f_Y(y) \Rightarrow$ not independent

(3) The joint density of X and Y is

$$f(x, y) = c(x^2 - y^2)e^{-x}, \quad 0 \leq x < \infty, \quad -x \leq y \leq x.$$

Find the conditional CDF of Y , given $X = x$.



Conditional PDF

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)} = \frac{c(x^2 - y^2)e^{-x}}{\int_{-x}^x c(x^2 - y^2)e^{-x} dy}$$

$$= \frac{x^2 - y^2}{\left[x^2 y - \frac{1}{3} y^3 \right]_{y=-x}^{y=x}} = \frac{x^2 - y^2}{\frac{4}{3} x^3} \quad -x \leq y \leq x$$

$$f_{Y|X}(y|x) = \begin{cases} \frac{x^2 - y^2}{\frac{4}{3} x^3} & -x \leq y \leq x \\ 0 & \text{otherwise.} \end{cases}$$

Conditional CDF of $Y | X=x$

$$F_{Y|X}(y|x) = \int_{-x}^y \frac{x^2 - t^2}{\frac{4}{3} x^3} dt = \frac{3}{4x^3} \left[x^2 t - \frac{1}{3} t^3 \right]_{t=-x}^{t=y}$$

$$= \frac{3}{4x^3} \left[x^2 y - \frac{1}{3} y^3 + x^3 - \frac{1}{3} x^3 \right]$$

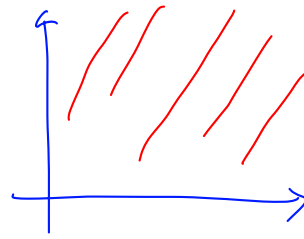
$$F_{Y|X}(y|x) = \begin{cases} 0 & y < -x \\ \left[\frac{3y}{4x} - \frac{y^3}{4x^3} + \frac{1}{2} \right] & -x \leq y \leq x \\ 1 & y > x \end{cases}$$

(4) The joint density function of X and Y is given by

$$f(x, y) = x e^{-x(y+1)}, \quad x > 0, y > 0$$

(a) Find the conditional density of X , given $Y = y$, and that of Y , given $X = x$.

(b) Find the density function of $Z = XY$.



$$(a) \quad f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{x e^{-x(y+1)}}{\int_0^{\infty} x e^{-x(y+1)} dx}$$

$$= \frac{x e^{-x(y+1)}}{\left[-\frac{x}{y+1} e^{-x(y+1)} \right]_0^{\infty} + \frac{1}{y+1} \int_0^{\infty} e^{-x(y+1)} dx}$$

$$= \frac{x e^{-x(y+1)}}{\frac{1}{y+1} \cdot \left[-\frac{1}{y+1} e^{-x(y+1)} \right]_0^{\infty}} = (y+1)^2 x e^{-x(y+1)} \quad \underline{\underline{x > 0}}$$

$$f_{Y|X} = \frac{f(x, y)}{f_X(x)} = \frac{x e^{-x(y+1)}}{\int_0^{\infty} x e^{-x(y+1)} dy} = \frac{x e^{-x(y+1)}}{x e^{-x} \int_0^{\infty} e^{-xy} dy} = \frac{e^{-xy}}{\left[-\frac{1}{x} e^{-xy} \right]_{y=0}^{y=\infty}} = x e^{-xy} \quad \underline{\underline{y > 0}}$$

(b) PDF of $Z = XY$ $f(x, y) = x e^{-x(y+1)} \quad x > 0, y > 0.$

CDF of $Z = XY$: $F_Z(z) = P(XY \leq z) = P\left(Y \leq \frac{z}{x}\right)$

$$= \int_0^{\infty} \int_0^{\frac{z}{x}} f(x, y) dy dx$$

$$= \int_0^{\infty} \int_0^{\frac{z}{x}} x e^{-x(y+1)} dy dx$$

$$= \int_0^{\infty} x e^{-x} \left[-\frac{1}{x} e^{-xy} \right]_{y=0}^{y=\frac{z}{x}} dx$$

$$\frac{d}{dx} \int_a^x f(t) dt = f(x)$$

$$= \int_0^{\infty} -e^{-x} (e^{-x} - 1) dx$$

$$= 1 - e^{-z} \quad z > 0$$

$$\text{PDF of } Z = f_Z(z) = \frac{d}{dz} F_Z(z) = \begin{cases} e^{-z} & z > 0 \\ 0 & \text{otherwise} \end{cases}$$

Test 2

Discrete RV :

- Binomial $P(X=k) = \binom{n}{k} p^k (1-p)^{n-k} \quad X=0,1,2,\dots,n$
- Poisson $P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!} \quad X=0,1,2,\dots$
- Geometric $P(X=k) = (1-p)^{k-1} p \quad X=1,2,\dots$

Continuous RV

- Unif (a,b) $f_X(x) = \begin{cases} \frac{1}{b-a} & x \in (a,b) \\ 0 & \text{else} \end{cases}$
- Normal $f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \quad -\infty < x < \infty$

Approximation

$n \geq 100$, $np \leq 10$: Poisson to approx Binomial.

$$X \sim \text{Bin}(n, p) \xrightarrow{\text{Approx}} X \sim \text{Po}(np)$$

$np(1-p) \geq 10$: $X \sim \text{Bin}(n, p) \xrightarrow{\text{Approx}} X \sim N(\overset{\mu}{np}, \overset{\sigma^2}{np(1-p)})$

Binomial

Normal

Continuity correction.

$$P(X=k) \approx P(k-0.5 \leq X \leq k+0.5)$$

$$P(X > k) \approx P(X \geq k+0.5)$$

$$P(X \geq k) \approx P(X \geq k-0.5)$$

$$P(X < k) \approx P(X \leq k - 0.5)$$

$$P(X \leq k) \approx P(X \leq k + 0.5)$$

Geometric Series

$$\sum_{k=0}^{\infty} r^k = \frac{1}{1-r} \quad |r| < 1$$

$$\sum_{k=0}^n r^k = \frac{1-r^{n+1}}{1-r}$$

Binomial Series

$$\sum_{k=0}^n \binom{n}{k} a^k b^{n-k} = (a+b)^n$$

Taylor expansion for e^x

$$e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!} \quad x \in \mathbb{R}$$

X_1, X_2, \dots, X_n having ^{same} CDF F

$$Y = \min\{X_1, \dots, X_n\}$$

$$\begin{aligned} \text{CDF of } Y &= F_Y(y) = P(Y \leq y) = P(\min\{X_1, \dots, X_n\} \leq y) \\ &= 1 - P(\min\{ \} \geq y) \\ &= 1 - P(X_1 \geq y) P(X_2 \geq y) \dots P(X_n \geq y) \\ &= 1 - (1 - F(y))^n \end{aligned}$$

$$Z = \max\{X_1, \dots, X_n\}$$

$$F_Z(z) = (F(z))^n$$