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$$\int_{X} (x) = \frac{1}{\sqrt{2\eta}} \exp\left(-\frac{x^{2}}{2}\right) - \infty < x < \infty$$

$$CDF of Y : F_{Y}(y) = f(Y \le y) = f(e^{X} \le y)$$

$$= f(X \le \ln y)$$

$$= \int_{-\infty}^{\ln y} \frac{1}{\sqrt{2\eta}} \exp\left(-\frac{x^{2}}{2}\right) dX$$

$$= \int_{-\infty}^{\ln y} f(x) dx$$

$$= f(g(y))g'(y)$$
PDF of Y : $f_{Y}(y) = \frac{1}{dy} F_{Y}(y) = \frac{1}{\sqrt{2\eta}} \exp\left(-\frac{(\ln y)^{2}}{2}\right) \cdot \frac{1}{y}$
O $\langle y < \infty \rangle$

$$E(Y) = \int_{0}^{\infty} y f_{Y}(y) dy = \int_{0}^{\infty} \frac{1}{\sqrt{2\eta}} exp\left(-\frac{(\ln y)^{2}}{2}\right) dy$$

Let
$$\frac{\ln y}{\sqrt{z}} = u \implies \frac{1}{y} dy = \int z du$$

As $y \rightarrow w$, $u \rightarrow w$; $y \rightarrow 0$, $u \rightarrow -w$, $y = e^{\int w} \Rightarrow dy = \int z e^{\int w} du$

$$F(Y) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp(-u^2) \cdot \sqrt{2} e^{\sqrt{2}\pi} du$$

$$= \frac{1}{\sqrt{11}} \int_{-\infty}^{\infty} \exp(-u^2 + \sqrt{2}u) du$$

$$= \frac{1}{\sqrt{11}} \int_{-\infty}^{\infty} \exp(-(u - \frac{1}{\sqrt{12}})^2 + \frac{1}{2}) du$$

$$= \frac{1}{\sqrt{11}} \int_{-\infty}^{\infty} e^{\frac{1}{2}} e^{-t^2} dt$$

$$= \frac{1}{\sqrt{11}} \int_{-\infty}^{\infty} e^{\frac{1}{2}} e^{-t^2} dt$$

$$= e^{\frac{1}{2}}$$
Since
$$\int_{-\infty}^{\infty} e^{-t^2} dt = \sqrt{11} \left(\frac{Gaussian}{Integral} \right)$$

Problem 3. The median of a continuous random variable having cumulative distribution function F is that value m such that $F(m) = \frac{1}{2}$. That is, a random variable is just as likely to be larger than its median as it is to be smaller. Find the median of X if X has the following distribution.

- (a) Uniform distribution over (a, b);
- (b) Standard normal distribution, N(0,1);
- (c) Exponential distribution with parameter λ ;
- (d) Lognormal distribution with parameters 0 and 1.

(b)
$$\int_{X}^{1} (r) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) - \infty < x < \omega$$

$$\frac{1}{2} = F(m) = P\left(X \le m\right) = \int_{-\infty}^{m} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) dx$$

$$= \int_{-\infty}^{0} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) dx + \int_{0}^{m} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) dx$$

$$= \frac{1}{2}$$

$$\therefore \int_{0}^{m} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) dx = 0 \quad \text{The integrand } \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) \ge 0 \quad \forall x.$$
For the integral to be zero, we need $m = 0$ (Area under graph = 0)
$$\int_{-\infty}^{0} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) dx = P(X \le 0) = \frac{1}{2} \quad \text{(since N(0, 1) is symmetric about } x = 0)$$

$$\begin{aligned} f_{\chi}(x) &= \pi e^{-\lambda x} & x \ge 0 \\ \frac{1}{2} &= F(m) = \int_{0}^{m} \pi e^{-\lambda x} dx = \left[-e^{-\lambda x} \right]_{0}^{m} = \left[-e^{-\lambda m} \right]$$

(d)
$$f_{Y}(y) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(\ln y)^{2}}{2}\right) \cdot \frac{1}{y}$$
 o $\langle y \rangle \ll 1$

$$\frac{1}{2} = F(m) = \int_{0}^{m} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(\ln y)^{2}}{2}\right) \cdot \frac{1}{y} dy$$

$$= \int_{-\infty}^{\ln m} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u^{2}}{2}\right) du$$

$$= \int_{-\infty}^{0} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u^{2}}{2}\right) du + \int_{0}^{\ln m} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u^{2}}{2}\right) du$$

$$= \frac{1}{2} \exp\left(-\frac{u^{2}}{2}\right) du + \int_{0}^{\ln m} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u^{2}}{2}\right) du$$

$$= \frac{1}{2} \exp\left(-\frac{u^{2}}{2}\right) du = 0 \implies S_{\text{AMME}} \text{ as } (b), \text{ we need } \ln m = 0 \implies m = 1$$

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Problem 4. A standard Cauchy random variable has density function

$$f(x) = \frac{1}{\pi(1+x^2)}, -\infty < x < \infty.$$

Show that if X is a standard Cauchy random variable, then 1/X is also a standard Cauchy random variable.

Let
$$Y = \frac{1}{X}$$

$$F_{Y}(y) = P(Y \le y) = P(\frac{1}{X} \le y) = P(X \ge \frac{1}{y}) = \int_{\frac{1}{y}}^{\frac{1}{y}} \frac{1}{\pi(1+X^{2})} dx = -\int_{\infty}^{\frac{1}{y}} \frac{1}{\pi(1+X^{2})} dx$$

$$f_{Y}(y) = \frac{1}{\alpha y} F_{Y}(y) = -\frac{1}{\pi(1+(\frac{1}{y})^{2})} \cdot (\frac{1}{y})' = \frac{1}{\pi(1+(\frac{1}{y})^{2})} \cdot \frac{1}{y^{2}} = \frac{1}{\pi(1+y^{2})}$$

Problem 6. A point is chosen at random on a line segment of length L. Find the probability that the ratio of the shorter to the longer segment is less than $\frac{1}{4}$.

Let X a Unif (O, L) denoting the position of the point on line segment.

Let Y be the ratio of the shorter to the longer segment.

$$Y = \min \left\{ \frac{X}{L-X}, \frac{L-X}{X} \right\}$$

$$P(Y \leftarrow \frac{1}{4}) = 1 - P(\min \left\{ \frac{x}{L - x}, \frac{1 - x}{x} \right\} > \frac{1}{4})$$

$$= 1 - P(\frac{x}{L - x}) > \frac{1}{4} \quad \text{and} \quad \frac{1 - x}{x} > \frac{1}{4})$$

$$= 1 - P(4x > 1 - x) \quad \text{and} \quad 4L - 4x > x)$$

$$= 1 - P(\frac{1}{5} + x) = 1 - \int_{\frac{1}{5}}^{\frac{x}{5}} \frac{1}{5} dx$$

$$= 1 - \frac{1}{5} \left(\frac{4L}{5} - \frac{L}{5} \right) = \frac{2}{5}$$

Problem 8. Let X be a random variable with the following density function

$$f_X(x) = \begin{cases} \frac{1}{3}, & -1 < x < 2, \\ 0, & \text{otherwise.} \end{cases}$$

Find the CDF and PDF of $Y = X^2$.

For
$$0 \le y \le 1$$
,
 $f_{X}(y) = P(Y \le y) = P(X^{2} \le y) = P(-\sqrt{y} \le X \le \sqrt{y}) = \int_{-\sqrt{y}}^{\sqrt{y}} \frac{1}{3} dx = \frac{2\sqrt{y}}{3}$

For
$$1 \le y \le 4$$
,
 $f_{Y}(y) : P(Y \le y) = P(Y \le 1) + P(Y \le y) = \frac{2\sqrt{1}}{3} + P(X^2 \le y)$

$$= \frac{2}{3} + P(1 \le x \le \lceil y \mid) = \frac{2}{3} + \int_{1}^{\sqrt{y}} \frac{1}{3} dx = \frac{1}{3} (\sqrt{y} - 1) + \frac{2}{3} = \frac{1}{3} (\sqrt{y} - 1)$$

$$\begin{cases}
\frac{1}{3\sqrt{y}} & 0 \leq y \leq 1 \\
\frac{1}{6\sqrt{y}} & 1 \leq y \leq 4
\end{cases}$$
otherwise