

## Problem 2.1

Let  $X$  be a random variable with the following cumulative distribution function (CDF).

$$F(x) = \begin{cases} 1 - e^{-(x/\alpha)^\beta} & \text{for } x > 0 \\ 0 & \text{otherwise} \end{cases}$$

where  $\alpha$  and  $\beta > 0$  are constants. The distribution of  $X$  is called a **Weibull distribution**.

- a) Find the PDF of the random variable  $X$ .  
 b) Let  $Y = (X/\alpha)^\beta$ . Identify the distribution of  $Y$ .

a) For  $x > 0$ ,  $f(x) = \frac{d}{dx} (1 - e^{-(\frac{x}{\alpha})^\beta})$  for  $x \leq 0$ ,  $f(x) = 0$ .

$$= -e^{-(\frac{x}{\alpha})^\beta} \cdot \frac{d}{dx} \left( -\frac{x^\beta}{\alpha^\beta} \right)$$

$$= \frac{\beta e^{-(\frac{x}{\alpha})^\beta}}{\alpha^\beta x^{1-\beta}}$$

b)  $F_Y(y) = P(Y \leq y) = P\left(\frac{X^\beta}{\alpha^\beta} \leq y\right) = P\left(X \leq \alpha y^{\frac{1}{\beta}}\right) = 1 - e^{-\left(\frac{\alpha y^{\frac{1}{\beta}}}{\alpha}\right)^\beta}$

$$= 1 - e^{-y}$$

$$f_Y(y) = e^{-y} \quad \text{for } y > 0.$$

$$= 0 \quad \text{otherwise} \quad \Rightarrow Y \sim \text{Exp}(1)$$

## Problem 2.2

Let  $X \sim \text{Exp}(\lambda)$ . Determine the distributions of  $Y_1 = \lambda X$  and  $Y_2 = 2\lambda X$ .

$$M_{Y_1}(t) = M_X(\lambda t) = \frac{\lambda}{\lambda - \lambda t} = \frac{1}{1-t} \Rightarrow \text{Gamma}(1, 1), \quad t < 1.$$

$$M_{Y_2}(t) = M_X(2\lambda t) = \frac{\lambda}{\lambda - 2\lambda t} = \frac{1}{1-2t} \Rightarrow \text{Gamma}(1, 2), \quad t < \frac{1}{2}.$$

$$\text{Gamma}\left(1, \frac{1}{\lambda}\right) \equiv \text{Exp}(\lambda)$$

### Problem 2.3

Let  $X_1, \dots, X_5$  be i.i.d.  $\sim N(0, 1)$  and let  $\bar{X} = \frac{1}{5} \sum_{i=1}^5 X_i$  be the sample mean. Let  $Y \sim N(0, 1)$  be another random variable which is independent from  $X_1, \dots, X_5$ . Find the distributions of the following random variables and justify your answers.

a) Random variable  $W = \sum_{i=1}^5 X_i^2$

b) Random variable  $U = \sum_{i=1}^5 (X_i - \bar{X})^2$

c) Random variable  $2Y/\sqrt{U}$

d) Random variable  $2(5\bar{X}^2 + Y^2)/U$

$$T = \frac{C}{\sqrt{\frac{D}{K}}} \quad \begin{matrix} C \sim N(0, 1) \\ D \sim \chi_k^2 \end{matrix}$$

(a)  $W \sim \chi_5^2$

$$\bar{X} \sim N(0, \frac{1}{5})$$

(b)  $S^2 = \frac{1}{n-1} \sum (X_i - \bar{X})^2 \quad \mu = 0, \sigma^2 = 1$

$$U = 4S^2 \sim \chi_4^2$$

(c)  $\frac{2Y}{\sqrt{U}} \sim \frac{N(0, 1)}{\sqrt{\frac{\chi_4^2}{4}}} = t_4$

$$\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \sim N(0, 1)$$

(d)  $\begin{matrix} \bar{X} \sim N(0, \frac{1}{5}) \Rightarrow \sqrt{5}\bar{X} \sim N(0, 1) \Rightarrow 5\bar{X}^2 \sim \text{Gamma}(\frac{1}{2}, 2) \\ Y \sim N(0, 1) \Rightarrow Y^2 \sim \text{Gamma}(\frac{1}{2}, 2) \end{matrix}$

Let  $D = 5\bar{X}^2 + Y^2$  ( $\because \bar{X}, Y$  independent).

$$M_D(t) = (1-2t)^{-\frac{1}{2}} (1-2t)^{-\frac{1}{2}} = (1-2t)^{-1}$$

$$D \sim \text{Gamma}(1, 2) \sim \chi_2^2$$

$$\frac{2(5\bar{X}^2 + Y^2)}{U} \sim \frac{2\chi_2^2}{\frac{\chi_4^2}{4}} = \frac{\chi_2^2}{\frac{\chi_4^2}{4}} \sim F(2, 4)$$

Problem 2.4

Let  $X_1, \dots, X_{16}$  be i.i.d.  $\sim N(0, 1)$  and let  $\bar{X}$  be the sample mean.

- Find the constant  $c$  such that  $P(|\bar{X}| < c) = 0.5$ .
- Find the mean and the variance of the sample variance  $S^2$ .

$$(a) \quad \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \sim N(0, 1) \Rightarrow 4\bar{X} \sim N(0, 1)$$

$$P(|\bar{X}| < c) = P(-4c < 4\bar{X} < 4c)$$

$$= 2\Phi(4c) - 1 = 0.5$$

$$\Phi(4c) = 0.75 \Rightarrow 4c = 0.674 \Rightarrow c = 0.169.$$

$$(b) \quad \frac{15S^2}{\sigma^2} \sim \chi_6^2 \Rightarrow 15S^2 \sim \text{Gamma}\left(\frac{15}{2}, 2\right)$$

$$E(15S^2) = 15 \Rightarrow E(S^2) = 1.$$

$$\text{Var}(15S^2) = \frac{15}{2}(2)(2) = 30 \Rightarrow \text{Var}(S^2) = \frac{2}{15}.$$

### Problem 2.5

Let  $X_1, \dots, X_n \sim \text{Poisson}(1)$  be an i.i.d. random sample. For each  $n \in \{10, 100, 1000, 10000\}$ , compute the following.

- The exact probability  $\Pr(\sum_{i=1}^n X_i \leq n)$  for each  $n$ .
- The approximation of  $\Pr(\sum_{i=1}^n X_i \leq n)$  obtained from the Central Limit Theorem (CLT).

Hint: note that  $\sqrt{n}(\bar{X} - \mu)/\sigma$  converges in distribution to  $N(0, 1)$  for  $n \rightarrow \infty$  where  $\mu = E[X_1]$  and  $\sigma = \text{Var}[X_1]$ .

$$(a) \quad X_1 + \dots + X_n \sim p_0(n).$$

$$\Pr\left(\sum_{i=1}^{10} X_i \leq 10\right) = P(Y \leq 10) = 0.583$$

$$\Pr\left(\sum_{i=1}^{100} X_i \leq 100\right) = 0.527$$

$$\Pr\left(\sum_{i=1}^{1000} X_i \leq 1000\right) = 0.50841$$

$$\Pr\left(\sum_{i=1}^{10000} X_i \leq 10000\right) = 0.50266.$$

$$(b) \quad \mu = E(X_1) = 1 \quad \sigma = \sqrt{\text{Var}(X_1)} = 1$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \sim N(1, 1)$$

$$P\left(\sum_{i=1}^n X_i \leq n\right) = P(n\bar{X} \leq n)$$

$$= P(\bar{X} \leq 1)$$

$$= P(Z \leq 0) = 0.5 \quad \text{for large } n \geq 30.$$

## MH3500 Statistics

### Tutorial 2

AY2022/23 Semester 2

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