Practice Week 6 Question 3

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Here, we will go through Practice Week 6 Question 3(a) and 3(b) in details and I will explain a few different methods to approach 3(b).

For 0 , we say X follows a negative binomial distribution if for <math>n = r, r + 1, ..., if

$$P(X = n) = \binom{n-1}{r-1} p^r (1-p)^{n-r}.$$

In short, we can write $X \sim \text{NegBin}(r, p)$. The generalized binomial theorem states that for any real numbers x, y and any real number r, we have

$$(x+y)^r = \sum_{k=0}^{\infty} \binom{r}{k} x^k y^{r-k}.$$

This in fact generalises to $r \in \mathbb{C}$. If we consider r as a non-negative integer and use the convention that $\binom{r}{k} = 0$, if r < k, then we have the simpler version of the binomial theorem, i.e.,

$$(x+y)^r = \sum_{k=0}^r \binom{r}{k} x^k y^{r-k}.$$

For both, we define $\binom{r}{k} = \frac{r \cdot (r-1) \cdot \ldots \cdot (r-k+1)}{k!}$.

For Question 3(a), we have

= 1.

$$\begin{split} \sum_{n=r}^{\infty} P(X=n) &= \sum_{n=r}^{\infty} \binom{n-1}{r-1} p^r (1-p)^{n-r} \\ &= \sum_{n=r}^{\infty} \binom{n-1}{n-r} p^r (1-p)^{n-r} \\ &= \sum_{j=0}^{\infty} \binom{n-1}{n-r} p^r (1-p)^{n-r} \\ &= \sum_{j=0}^{\infty} \binom{n-1}{n-r} p^r (1-p)^{n-r} \\ &= p^r \sum_{j=0}^{\infty} \binom{j+r-1}{j} (1-p)^j \\ &= p^r \sum_{j=0}^{\infty} \frac{(r+j-1) \cdot (r+j-2) \cdot \ldots \cdot (r+1) \cdot r}{j!} (1-p)^j \\ &= p^r \sum_{j=0}^{\infty} (-1)^j \frac{(-r-j+1) \cdot (-r-j+2) \cdot \ldots \cdot (-r-1) \cdot (-r)}{j!} (1-p)^j \\ &= p^r \sum_{j=0}^{\infty} (-1)^j \binom{-r}{j} (1-p)^j \\ &= p^r \sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j \cdot 1^{-r-j} \\ &= p^r (p-1+1)^{-r} \end{split}$$
 by generalised binomial theorem

For Question 3(b), I will provide two different methods to approach it.

Method 1. We need to compute $\mathbb{E}(X)$ and $Var(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2$. For any $k \in \mathbb{Z}$, $k \geq 1$, we can compute $\mathbb{E}(X^k)$ as follows:

$$\begin{split} \mathbb{E}(X^k) &= \sum_x x^k P(X=x) \\ &= \sum_{n=r}^\infty n^k \binom{n-1}{r-1} p^r (1-p)^{n-r} \\ &= r \sum_{n=r}^\infty n^{k-1} \binom{n}{r} p^r (1-p)^{n-r} \\ &= \frac{r}{p} \sum_{n=r}^\infty n^{k-1} \binom{n}{r} p^{r+1} (1-p)^{n-r} \\ &= \frac{r}{p} \sum_{m=r+1}^\infty (m-1)^{k-1} \binom{m-1}{r} p^{r+1} (1-p)^{m-r-1} \\ &= \frac{r}{p} \cdot \mathbb{E}\left[(Y-1)^{k-1} \right] \end{split} \qquad \text{based on pmf of the distribution where } Y \sim \text{NegBin}(r+1,p). \end{split}$$

Hence, we have

$$\mathbb{E}(X) = \frac{r}{p} \cdot \mathbb{E}(Y - 1)^0 = \frac{r}{p}.$$

We can also compute

$$\mathbb{E}(X^2) = \frac{r}{p}\mathbb{E}(Y-1) = \frac{r}{p}(\mathbb{E}(Y)-1) = \frac{r}{p}\left(\frac{r+1}{p}-1\right).$$

Hence,

$$Var(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 = \frac{r(1-p)}{p^2}.$$

This method is a more combinatorial approach and allows us to calculate $\mathbb{E}(X^k)$ for all k in a recursive manner. $\mathbb{E}(X^k)$ is the k-th moment of the negative binomial distribution (r, p). The moments of a distribution essentially define the distribution itself.

Method 2. This will be a calculus-based approach where we need to analyse the convergence of infinite series under differentiation.

$$\begin{split} \mathbb{E}(X) &= \sum_{n=r}^{\infty} n \binom{n-1}{r-1} p^r (1-p)^{n-r} \\ &= \sum_{n=r}^{\infty} n \binom{n-1}{n-r} p^r (1-p)^{n-r} & \text{by symmetry} \\ &= \sum_{j=0}^{\infty} (j+r) \binom{r+j-1}{j} p^r (1-p)^j & \text{let } j = n-r \\ &= rp^r \sum_{j=0}^{\infty} \binom{j+r-1}{j} (1-p)^j + p^r \sum_{j=0}^{\infty} \binom{r+j-1}{j} j (1-p)^j \\ &= r + p^r \sum_{j=0}^{\infty} \binom{r+j-1}{j} j (1-p)^j & \text{from (a): } p^r \sum_{j=0}^{\infty} \binom{j+r-1}{j} (1-p)^j = 1 \\ &= r + p^r \sum_{j=0}^{\infty} \binom{-r}{j} j (p-1)^j & \text{from (a): } p^r \sum_{j=0}^{\infty} \binom{j+r-1}{j} (1-p)^j = 1 \\ &= r + p^r (p-1) \sum_{j=0}^{\infty} \binom{-r}{j} j (p-1)^{j-1} \\ &= r + p^r (p-1) \sum_{j=0}^{\infty} \binom{-r}{j} \frac{d}{dp} (p-1)^j & \text{interchanging summation and differentiation} \\ &= r + p^r (p-1) \frac{d}{dp} (p^{-r}) & \text{from (a)} \\ &= r - \frac{rp^r (p-1)}{p^{r+1}} = \frac{r}{p}. \end{split}$$

We can compute $\mathbb{E}(X^2)$ in a similar manner.

$$\begin{split} \mathbb{E}(X^2) &= \sum_{n=r}^{\infty} n^2 \binom{n-1}{r-1} p^r (1-p)^{n-r} \\ &= \sum_{j=0}^{\infty} (j+r)^2 \binom{r+j-1}{j} p^r (1-p)^j \\ &= r^2 \underbrace{p^r \sum_{j=0}^{\infty} \binom{r+j-1}{j} (1-p)^j + 2r \underbrace{p^r \sum_{j=0}^{\infty} j \binom{r+j-1}{j} (1-p)^j + p^r \sum_{j=0}^{\infty} j^2 \binom{r+j-1}{j} (1-p)^j}_{=1, \text{ from (a)}} + p^r \underbrace{\sum_{j=0}^{\infty} j \binom{r+j-1}{j} (1-p)^j}_{=\frac{r}{p}-r, \text{ from } \mathbb{E}(X)} \\ &= r^2 + 2r \left(\frac{r}{p} - r\right) + p^r \underbrace{\sum_{j=0}^{\infty} j (j-1) \binom{r+j-1}{j} (1-p)^j + p^r \underbrace{\sum_{j=0}^{\infty} j \binom{r+j-1}{j} (1-p)^j}_{=\frac{r}{p}-r}}_{=r^2 + (2r+1) \binom{r}{p} - r} + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} j (j-1) (p-1)^{j-2}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} \frac{d^2}{dp^2} (p-1)^j}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} \frac{d^2}{dp^2} (p-1)^j}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} (p-1)^j}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j} \binom{-r}{j}}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j}}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j}}_{j=0}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (1-p)^2 \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (2r+1) \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r + p^r (2r+1) \underbrace{\sum_{j=0}^{\infty} \binom{-r}{j}}_{j=0} \\ &= r^2 + (2r+1) \binom{r}{p} - r +$$

Then, we have

$$\mathbb{E}(X^2) = r^2 + (2r+1)\left(\frac{r}{p} - r\right) + p^r(1-p)^2 \frac{d^2}{dp^2}(p^{-r})$$
$$= r^2 + (2r+1)\left(\frac{r}{p} - r\right) + \frac{(1-p)^2 r(r+1)}{p^2}$$

Thus.

$$Var(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 = \frac{r^2p^2 + pr(2r+1)(1-p) + r(r+1)(1-p)^2}{p^2} - \frac{r^2}{p^2} = \frac{r(1-p)}{p^2}.$$

Extra information on calculus: In both the computations of $\mathbb{E}(X)$ and $\mathbb{E}(X^2)$, we interchange differentiation and summation. This always works for a finite series, since differentiation is linear. However, it is not always the case for infinite series. Given a power series centered at x = a,

$$f(x) = \sum_{n=0}^{\infty} c_n (x-a)^n$$
, for $|x-a| < R$.

R is the radius of convergence. f is differentiable on the open interval of convergence (a - R, a + R) and we can interchange summation and differentiation to do term-by-term differentiation, i.e.,

$$f'(x) = \frac{d}{dx} \sum_{n=0}^{\infty} c_n (x-a)^n = \sum_{n=0}^{\infty} c_n \frac{d}{dx} (x-a)^n = \sum_{n=1}^{\infty} n c_n (x-a)^{n-1}.$$

f' also converges on (a-R,a+R). In our example in $\mathbb{E}(X)$, we have to check the interval of convergence of $\sum_{j=0}^{\infty} {r\choose j} (p-1)^j$.

We can use the ratio test.

$$L = \lim_{n \to \infty} \left| \frac{\binom{-r}{n+1}(p-1)^{n+1}}{\binom{-r}{n}(p-1)^n} \right|$$

$$= |p-1| \lim_{n \to \infty} \left| \frac{(-r)(-r-1)\dots(-r-n)}{(-r)(-r-1)\dots(-r-n+1)} \cdot \frac{n!}{(n+1)!} \right|$$

$$= |p-1| \lim_{n \to \infty} \frac{n+r}{n+1} = |p-1|.$$

By ratio test, the power series converges if and only if |p-1| < 1. From our assumption of the negative binomial distribution that 0 , this falls within the interval of convergence, hence the power series is always differentiable for <math>0 and we are able to interchange differentiation and summation.

Method 3. Suppose you already know $\mathbb{E}(X) = r/p$, this gives an easier way to compute $\mathbb{E}(X^2)$. The same method has been applied to find the variance of geometric distribution.

$$\begin{split} \mathbb{E}(X(X+1)) &= \sum_{n=r}^{\infty} n(n+1) \binom{n-1}{r-1} p^r (1-p)^{n-r} \\ &= \sum_{n=r}^{\infty} n(n+1) \frac{(n-1)!}{(r-1)!(n-r)!} p^r (1-p)^{n-r} \\ &= \sum_{n=r}^{\infty} \frac{(n+1)!}{(r-1)!(n-r)!} p^r (1-p)^{n-r} \\ &= \sum_{n=r}^{\infty} \frac{(n+1)!}{(r+1)!(n-r)!} \cdot r(r+1) \cdot p^r (1-p)^{n-r} \\ &= r(r+1) \sum_{n=r}^{\infty} \binom{n+1}{r+1} p^r (1-p)^{n-r} \\ &= r(r+1) \sum_{n=r}^{\infty} \binom{n+1}{r+1} p^{r+2} (1-p)^{n-r} \\ &= \frac{r(r+1)}{p^2} \sum_{n=r+2}^{\infty} \binom{n-1}{r+1} p^{r+2} (1-p)^{n-r-2} \\ &= \frac{r(r+1)}{p^2} \sum_{n=r+2}^{\infty} P(Y=n) \qquad \qquad \text{where } Y \sim \text{NegBin}(r+2,p) \\ &= \frac{r(r+1)}{p^2} \end{split}$$

. Hence,

$$Var(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 = (\mathbb{E}(X(X+1)) - \mathbb{E}(X)) - \mathbb{E}(X)^2 = \frac{r(1-p)}{p^2}.$$