stat443 assignment2

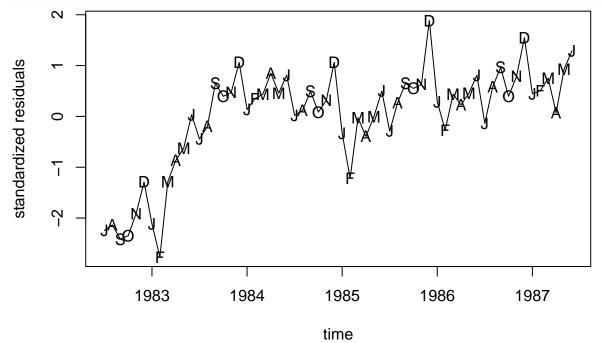
Yiming Shen 20891774

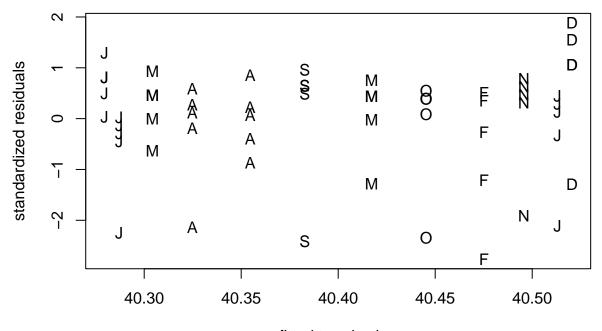
13/02/2023

Problem 1

```
library(TSA)
##
## Attaching package: 'TSA'
  The following objects are masked from 'package:stats':
##
##
       acf, arima
  The following object is masked from 'package:utils':
##
##
##
       tar
data("hours")
(a)
har. <- harmonic(hours)</pre>
hours.lm1 <- lm(hours ~ har.)
summary(hours.lm1)
##
## Call:
## lm(formula = hours ~ har.)
## Residuals:
       Min
                1Q Median
                                ЗQ
                                       Max
## -1.6752 -0.1935 0.1956 0.3825 1.1795
##
## Coefficients:
##
                   Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                   40.40000
                               0.08473 476.783
                                                 <2e-16 ***
## har.cos(2*pi*t) 0.11297
                               0.11983
                                        0.943
                                                  0.350
## har.sin(2*pi*t) -0.04533
                               0.11983 -0.378
                                                  0.707
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.6564 on 57 degrees of freedom
## Multiple R-squared: 0.01778, Adjusted R-squared: -0.01668
## F-statistic: 0.5159 on 2 and 57 DF, p-value: 0.5997
```

(b)

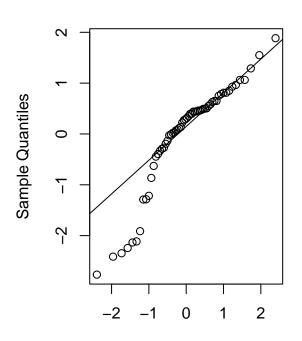


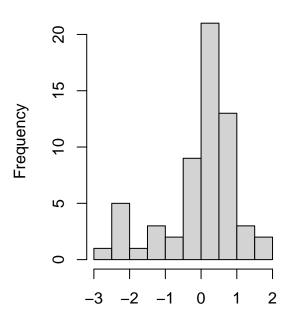


fitted trend value

Normal Q-Q Plot

Histogram for Residuals





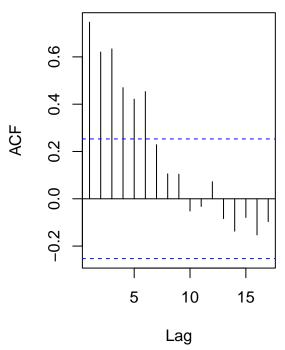
Theoretical Quantiles

standardized residuals

```
# comments: There are deviations on lower end of QQ-plot and the residuals are
# not normally distributed.
# The histogram also indicates a left-skewed distribution, which means that
# there is a heavy left tail.

# ACF plot
acf(rstudent(hours.lm1))
# comments: from the ACF plot, since many spikes go beyond the confidence
# band, so a strong correlation among residual values can be observed.
```

Series rstudent(hours.lm1)



```
# test for normality
shapiro.test(rstudent(hours.lm1))
```

```
##
    Shapiro-Wilk normality test
##
##
## data: rstudent(hours.lm1)
## W = 0.88612, p-value = 4.259e-05
# comments: since p-value = 4.259e-05 < 0.05,
# so we have strong evidence against our null hypothesis: data come from a
# normal distribution.
# test for constant variance
seg <- as.vector(season(hours))</pre>
fligner.test(rstudent(hours.lm1), seg)
##
##
   Fligner-Killeen test of homogeneity of variances
## data: rstudent(hours.lm1) and seg
## Fligner-Killeen:med chi-squared = 2.7252, df = 11, p-value = 0.9939
# comments: since p-value = 0.9939 > 0.05,
# so we have no evidence against our null hypothesis: data have
# constant variance.
# test for randomness
library("randtests")
library("lawstat")
```

```
##
## Attaching package: 'lawstat'
## The following object is masked from 'package:randtests':
##
       runs.test
difference.sign.test(rstudent(hours.lm1))
## Difference Sign Test
##
## data: rstudent(hours.lm1)
## statistic = 3.77, n = 60, p-value = 0.0001632
## alternative hypothesis: nonrandomness
# comments: since p-value = 0.0001632 < 0.05,
\# so we have strong evidence against our null hypothesis: the data is random.
runs.test(rstudent(hours.lm1))
##
## Runs Test - Two sided
##
## data: rstudent(hours.lm1)
## Standardized Runs Statistic = -2.8646, p-value = 0.004176
# comments: same as difference sign test, since p-value = 0.004176,
# so we have strong evidence against our null hypothesis: the data is random.
```

(c)

The cosine trend model is not appropriate for the trend estimation. Because from residual vs. time, we observed a quadratic trend. From QQ-plot and histogram, we found that the residuals are not normally distributed, which violated the normality assumption. From ACF plot, a strong correlation among residual values can be observed, which violated the uncorrelation assumption.

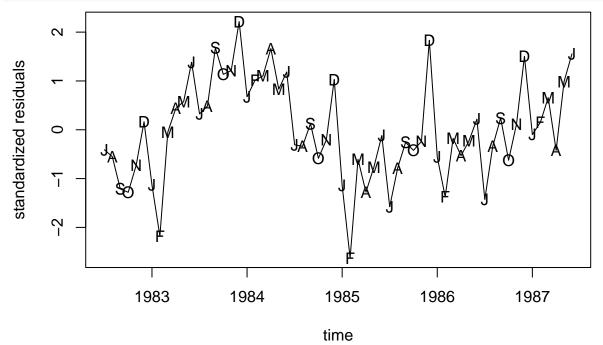
(d)

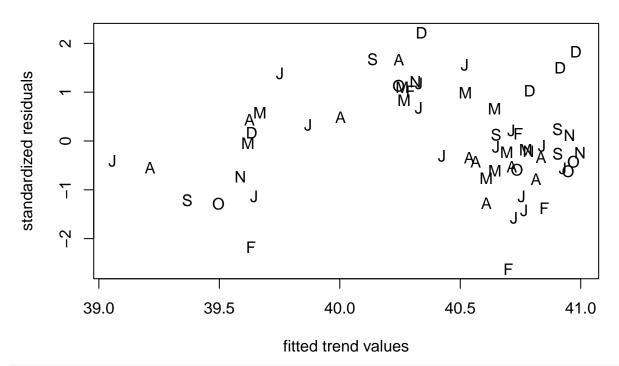
```
# fit a combination of cosine trend model and quadratic model
hours.lm2 <- lm(hours~har.+time(hours)+I(time(hours)^2))
summary(hours.lm2)

##
## Call:
## lm(formula = hours ~ har. + time(hours) + I(time(hours)^2))
###</pre>
```

```
## Residuals:
##
       Min
                 1Q
                    Median
                                   3Q
## -0.99955 -0.23753 -0.05985 0.28763 0.86045
## Coefficients:
##
                     Estimate Std. Error t value Pr(>|t|)
                   -5.061e+05 1.134e+05 -4.462 4.06e-05 ***
## (Intercept)
                   7.478e-02 7.586e-02
                                         0.986
## har.cos(2*pi*t)
                                                  0.3286
## har.sin(2*pi*t) -1.382e-01 7.665e-02
                                        -1.804
                                                  0.0768 .
## time(hours)
                    5.097e+02 1.143e+02
                                         4.460 4.09e-05 ***
## I(time(hours)^2) -1.283e-01 2.879e-02 -4.457 4.13e-05 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4149 on 55 degrees of freedom
## Multiple R-squared: 0.6213, Adjusted R-squared: 0.5938
## F-statistic: 22.56 on 4 and 55 DF, p-value: 4.573e-11
```

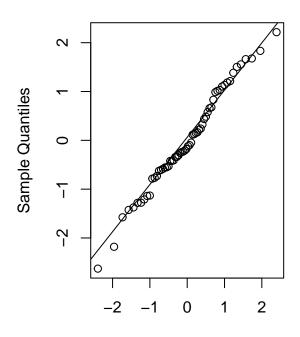
(e)

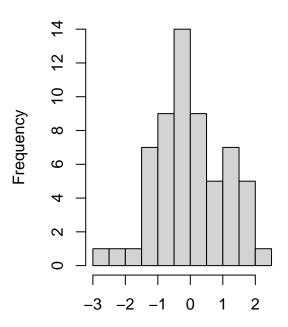




Normal Q-Q Plot

istogram for Residuals (Quadratic N





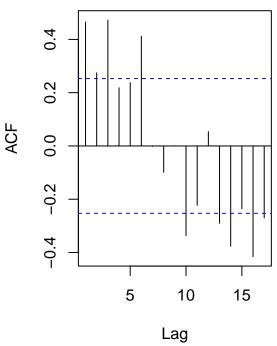
Theoretical Quantiles

standardized residuals

```
# comments: The residuals follow normal distribution strictly.

# ACF plot
acf(rstudent(hours.lm2))
# comments: from the ACF plot, since many spikes go beyond the confidence
# band, so a strong correlation among residual values can be observed.
```

Series rstudent(hours.lm2)



```
# test for normality
shapiro.test(rstudent(hours.lm2))
```

```
##
##
    Shapiro-Wilk normality test
##
## data: rstudent(hours.lm2)
## W = 0.98671, p-value = 0.7581
# comments: since p-value = 0.7581 > 0.05,
# so we have no evidence against our null hypothesis: data come from a
# normal distribution.
# test for constant variance
seg <- as.vector(season(hours))</pre>
fligner.test(rstudent(hours.lm2), seg)
##
##
   Fligner-Killeen test of homogeneity of variances
## data: rstudent(hours.lm2) and seg
## Fligner-Killeen:med chi-squared = 6.6642, df = 11, p-value = 0.8256
# comments: since p-value = 0.8256 > 0.05,
# so we have no evidence against our null hypothesis: data have
# constant variance.
# test for randomness
library("randtests")
library("lawstat")
```

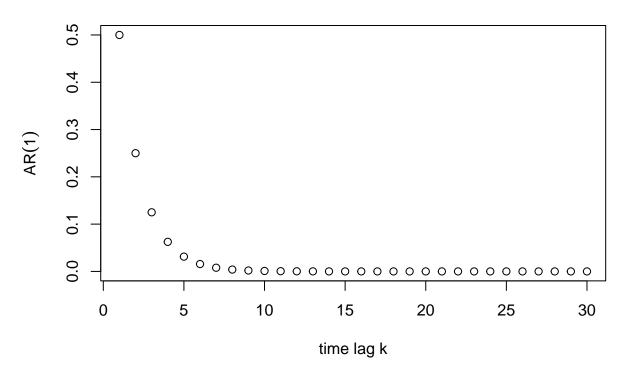
```
difference.sign.test(rstudent(hours.lm2))
##
## Difference Sign Test
##
## data: rstudent(hours.lm2)
## statistic = 2.4394, n = 60, p-value = 0.01471
## alternative hypothesis: nonrandomness
# comments: since p-value = 0.01471 < 0.05,
# so we have evidence against our null hypothesis: the data is random.
runs.test(rstudent(hours.lm2))
##
## Runs Test - Two sided
##
## data: rstudent(hours.lm2)
## Standardized Runs Statistic = -2.8646, p-value = 0.004176
# comments: same as difference sign test, since p-value = 0.004176 < 0.05
# so we have strong evidence against our null hypothesis: the data is random.
```

(f)

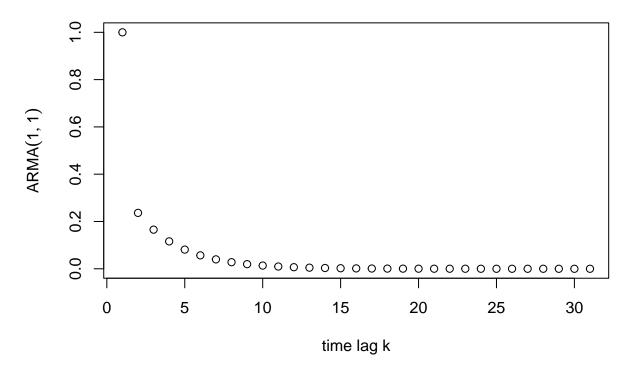
The combined model is relative appropriate for the trend estimation. Because from residual vs. time, we observed no obvious pattern anymore. From QQ-plot and histogram, we found that the residuals are normally distributed. From residual vs. estimated value, there is no obvious change in the variability at different estimated levels. But for ACF plot, a correlation still exists among residual values. I think that the combined model significantly improved the quality of fit.

Problem 2

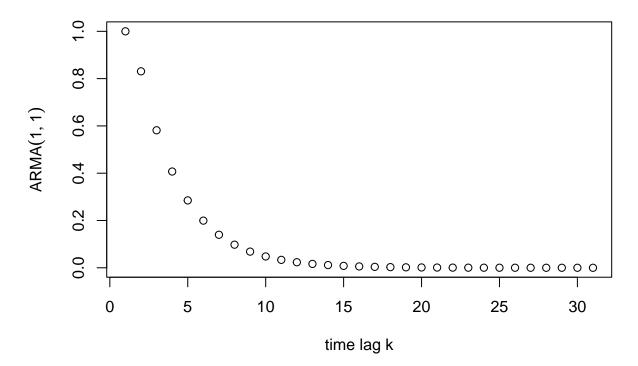
ACF of AR(1) Process with phi=0.5



ACF of ARMA(1,1) Process with phi=0.7 theta=0.5

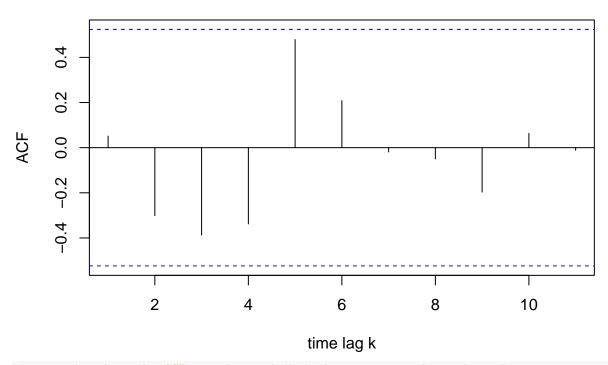


ACF of ARMA(1,1) Process with phi=0.7 theta=-0.5



Problem 4

Correlogram for time series Y

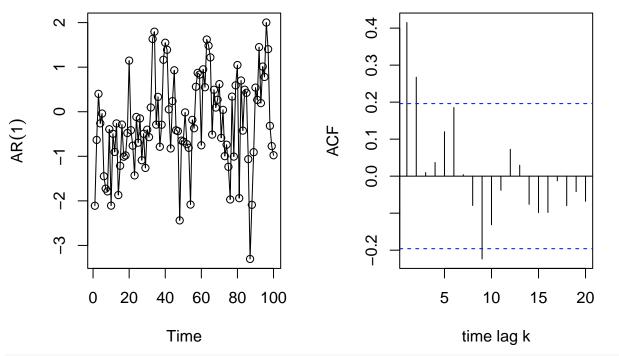


comments: from the ACF, we observed that there is no spike go beyond # confidence band, so no strong correlation is observed.

Problem 5

Generating from $Y_t=0.4 Y_{t-1}+ Z_t \text{ proc}$

ACF of $Y_{t}=0.4 Y_{t-1}+ Z_{t}$ process



comments on ACF: from the ACF, we observe that the magnitude of the ACF # decays exponentially as the time lag k increases.