

$\beta$  는 어떤가? 진짜가?

우선의 말.

$$x^T y = x^T x \beta$$

Least Square Estimate

$$\nabla_{\beta} (\beta^T A \beta) = \nabla_{\beta} \sum_{i,j} \beta_i A_{ij} \beta_j$$

$$(loss) = \|y - X\beta\|^2$$

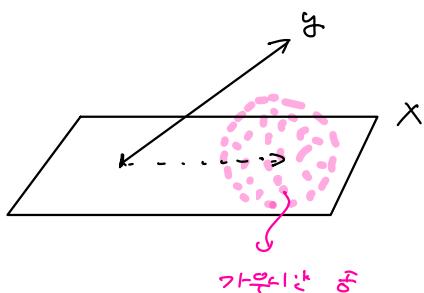
$$= (A + A^T) \beta$$

$$\begin{aligned} &= (y - X\beta)^T (y - X\beta) \\ &= (y^T - \beta^T x^T) (y - X\beta) \\ &\Rightarrow y^T y - y^T x \beta - \beta^T x^T y + \beta^T x^T x \beta \end{aligned}$$

$$\nabla_{\beta} (loss) = -2 x^T y + 2 x^T x \beta = \odot \text{ (극값)}$$

hence  $x^T y = x^T x \beta$  우선의 말

Maximum Likelihood Estimate



$$\text{let } Y = X\beta + \varepsilon \sim N(X\beta, \sigma^2 I)$$

$$\text{then } p(Y) \underset{\text{Likelihood}}{=} \frac{1}{(2\pi)^{d/2} \cdot \sigma^d} \exp\left(-\frac{1}{2\sigma^2} \|y - X\beta\|^2\right)$$

squared L2 Norm

$p(\theta)$  : prior

$p(\text{data}|\theta)$  : likelihood

$p(\text{data})$  : evidence.

$$p(\theta|\text{data}) : \text{posterior} = \frac{p(\text{data}|\theta) \cdot p(\theta)}{p(\text{data})}$$

Bernoulli R.V "동일 단위기"

$$p(x_i|\theta) = \begin{cases} \theta & x_i = 1 \\ 1-\theta & x_i = 0 \end{cases}$$

$$= \theta^{x_i} (1-\theta)^{1-x_i}$$

likelihood  $\propto$   $\theta^{\sum x_i} (1-\theta)^{n-\sum x_i}$

$$p(x|\theta) = \theta^{\sum x_i} (1-\theta)^{n-\sum x_i}$$

say  $\nabla_{\theta} \log p(x|\theta) = 0$  (MLE)

$$\nabla_{\theta} \log p(x|\theta) = \frac{\sum x_i}{\theta} - \frac{(n-\sum x_i)}{1-\theta}$$

then ...  $\theta = \frac{\sum x_i}{n}$

$$\nabla_{\theta} \log p(x|\theta) = \frac{\sum x_i}{\theta} - \frac{(n-\sum x_i)}{1-\theta}$$

$$\text{Beta}(\alpha, \beta) = x^{\alpha-1} (1-x)^{\beta-1} \cdot \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha) \cdot \Gamma(\beta)}$$

Note.

$$\Gamma(n) = (n-1)!$$

where  $0 \leq x \leq 1$

posterior의 관점

실험 결과  $\frac{H}{7} \frac{T}{3}$

$$\nabla_{\theta} \log p(\theta|x) = 0 \quad (\text{MAP})$$

$$\text{Beta}(1,1)$$

$$p(\theta|x) \propto \frac{1}{\theta} \cdot \theta^7 (1-\theta)^3 = \theta^7 (1-\theta)^3$$

$$\frac{p(\theta)}{p(\theta)} \cdot \frac{p(x|\theta)}{p(x|\theta)}$$

Beta(8,4)

$$\approx$$

Uniform Distribution of  $\theta$

$$\theta_{\text{MAP}} = \theta_{\text{MLE}}$$

$$= 0.7$$

$$\text{Beta}(2,2)$$

$$p(\theta|x) \propto \frac{1}{\theta} \cdot (1-\theta)^1 \cdot \theta^7 (1-\theta)^3 = \theta^8 (1-\theta)^4$$

$$\frac{p(\theta)}{p(\theta)} \cdot \frac{p(x|\theta)}{p(x|\theta)}$$

Beta(9,5)

$$\approx \theta_{\text{MAP}} = 0.666\ldots$$

$$\bullet H(X) = \mathbb{E}[-\log p(x)] \geq 0$$

$$\bullet h(x) = -\log p(x)$$

$$\bullet H(X, Y) = H(X) + H(Y|X)$$

• if  $x, y$  independent

$$h(x, y) = h(x) + h(y)$$

$$\begin{array}{ccc} H \uparrow & & H \downarrow \\ (\frac{1}{2}, \frac{1}{2}) & \xrightarrow{\text{Eq}} & \frac{1}{2} \text{ vs } \frac{1}{2} \\ \frac{1}{2} \text{ vs } \frac{1}{2} & \xrightarrow{\text{Eq}} & \frac{1}{2} \text{ vs } \frac{1}{2} \end{array}$$

### KL Divergence

$$D_{KL}(P||\tilde{P})$$

$$= \sum_x p(x) \log \frac{p(x)}{\tilde{p}(x)}$$

$$= \int p(x) \log \frac{p(x)}{\tilde{p}(x)} dx \geq 0$$

def

$$\begin{aligned} & \mathbb{E}_{p(x)} \left[ -\log \frac{\tilde{p}(x)}{p(x)} \right] \\ & \geq -\log \mathbb{E}_{p(x)} \left[ \frac{\tilde{p}(x)}{p(x)} \right] \\ & = -\log \sum_x p(x) \cdot \frac{\tilde{p}(x)}{p(x)} \geq 0 \end{aligned}$$

Const

Const = 1

(x) symmetric

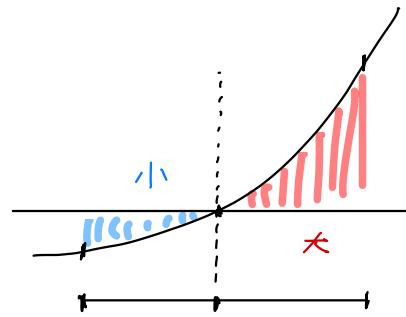
(x) triangular inequality

### Jensen's Inequality

f convex

$$\text{then } \mathbb{E}[f(x)] \geq f(\mathbb{E}[x])$$

$$\Downarrow f(x) = \text{Const}$$



### Mutual Information

$$I(X; Y) = D_{KL}(P(x,y) || p(x)p(y))$$

$$= \sum p(x,y) \log \frac{p(x,y)}{p(x)p(y)}$$

$$= H(X) + H(Y) - H(X, Y)$$

$$= H(X) - H(X|Y)$$

### Cross Entropy

$$\bullet H_p(\tilde{p}) = \mathbb{E}_{\tilde{p}}[-\log \tilde{p}(x)]$$

$$= -\sum_x p(x) \log \tilde{p}(x)$$

$$\bullet D_{KL}(P||\tilde{p}) = \mathbb{E}_P \left[ \log \frac{p(x)}{\tilde{p}(x)} \right]$$

$$= H_p(\tilde{p}) - H(p) \geq 0$$

fixed

## overfitting

단일 변수에  
의존

$$y = \begin{bmatrix} 1 & x^1 & \cdots & x^m \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_m \end{bmatrix}$$

불필요한  
큰 차원

$\approx$  overfitting

## Regularization

### Regularization

$$(\text{loss}) = \|Y - X\beta\|^2 + \lambda \|\beta\|^2$$

$$\nabla_{\beta} (\text{loss}) = \nabla_{\beta} (Y^T Y - Y^T X\beta - \beta^T X^T Y + \beta^T X^T X\beta) + \nabla_{\beta} \lambda \beta^T \beta$$

$$= 2 \left[ \underbrace{-X^T(Y - X\beta)}_{\text{제한 조건}} + \underbrace{\lambda \beta}_{\text{정규화 조건}} \right] = 0$$

## Kernel

$$K(x_1, x_2) = \underbrace{\phi(x_1)^T}_{\text{low dimension}} \underbrace{\phi(x_2)}_{\text{high dimension}}$$

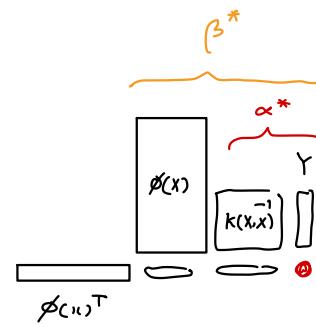
$$(\text{loss}) = \|Y - \phi(X)^T \beta\|^2 + \lambda \|\beta\|^2$$

$$= Y^T Y - Y^T \phi(X)^T \beta - \beta^T \phi(X) Y + \beta^T \phi(X) \phi(X)^T \beta + \lambda \beta^T \beta$$

$$\nabla_{\beta} (\text{loss}) = 2 \left[ -\phi(X) (Y - \phi(X)^T \beta) + \lambda \beta \right] = 0$$

$$\begin{aligned} \beta^* &= (\phi(X) \phi(X)^T + \lambda I)^{-1} \phi(X) Y \\ &= \phi(X) (\underbrace{\phi(X)^T \phi(X)}_{\alpha^*} + \lambda I)^{-1} Y \end{aligned}$$

$$\begin{aligned} f^*(x) &= \phi(x)^T \beta^* \\ &= K(x, X) \alpha^* \end{aligned}$$



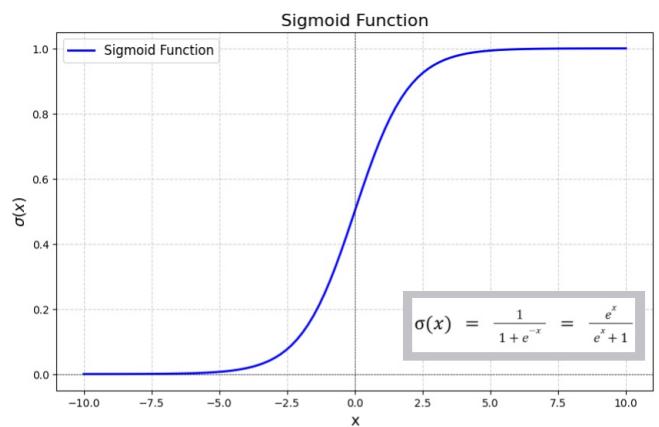
## Sigmoid

임의의 확률 확률 분포를 계산하는 법 .

↑  
if probability =  $\frac{e^z}{1+e^z}$   $[0, 1]$

then odds =  $e^z$   $[0, \infty)$

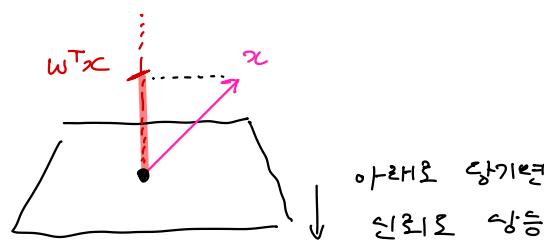
$\log(\text{odds})$  =  $z$   $(-\infty, \infty)$



Note that  $\sigma' = \sigma(1-\sigma)$

## Logistic Regression (MLE)

$$\sigma' = \sigma(1-\sigma)$$



$$\begin{aligned}\nabla_w \sigma_i &= \sigma_i(1-\sigma_i) \nabla_w (w^T x_i) \\ &= \sigma_i(1-\sigma_i) x_i \\ &\propto x\end{aligned}$$

$$p(y_i | x_i, w) = (\sigma(w^T x_i))^{y_i} (1 - \sigma(w^T x_i))^{1-y_i}$$

$$\mathcal{L} = \prod_{i=1}^n p_i$$

1 data

$$\log p_i = (y_i \log \sigma_i) + ((1-y_i) \log (1-\sigma_i))$$

$$\nabla_w (-\log p_i) = (y_i (\sigma_i - 1) x_i) + ((1-y_i) \sigma_i x_i)$$

$$= \begin{pmatrix} y_i \sigma_i x_i \\ -y_i x_i \end{pmatrix} + \begin{pmatrix} \sigma_i x_i \\ -y_i \sigma_i x_i \end{pmatrix}$$

$$= (\sigma_i - y_i) x_i$$

여기서  $\sigma_i$

최적화변수

우리

$$\text{hence } \nabla_w (-\log \mathcal{L}) = \sum_i (\sigma(w^T x_i) - y_i) x_i$$

$$\begin{aligned}\nabla_w \nabla_w (-\log p_i) &= [\nabla_w (\sigma_i - y_i)] x_i^T \\ &= \sigma_i' x_i x_i^T\end{aligned}$$

for any  $v$

$$v^T (x^T \sigma_i x) v$$

$$= (x v)^T \sigma_i (x v)$$

$$= \sum_i (\sigma_i') ((x v)_i)^2 \geq 0$$

$$\nabla_w \nabla_w (-\log \mathcal{L}) = \begin{bmatrix} 1 & \dots \end{bmatrix} \begin{bmatrix} \sigma_1' & \dots \\ \vdots & \ddots & \sigma_n' \end{bmatrix} \begin{bmatrix} -x_1 \\ \vdots \\ -x_n \end{bmatrix}$$

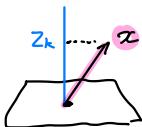
positive  
semi-definite  
 $\langle v, Hv \rangle \geq 0$

$-\log \mathcal{L}$  convex  
有 global min

$$\text{because } \sigma' = \sigma(1-\sigma) \geq 0$$

## Multiclass classification

Let  $z_{ik} = w_k^T x$



$$\nabla_{w_m} (w_k^T x) = \begin{cases} x & \text{if } m=k \\ 0 & \text{if } m \neq k \end{cases}$$

Then  $p_{ik} = \frac{e^{z_{ik}}}{\sum_j e^{z_j}}$

$$= \delta_{mk} \cdot x$$

$$\nabla_{w_m} (p_{ik}) = \frac{\sum_j e^{z_j} (e^{z_{ik}} \cdot (\nabla z_{ik})) - e^{z_{ik}} (\sum_j e^{z_j} \nabla z_j)}{(\sum_j e^{z_j})^2}$$

$$= p_{ik} (\delta_{mk} - p_{im}) x$$

$$\nabla_{w_m} \left( -\log \frac{\text{likelihood}}{\text{of } x} \right) = \sum_{k=1}^c y_{ik} \nabla (-\log p_{ik})$$

$$= \sum_{k=1}^c y_{ik} \cdot \frac{1}{p_{ik}} \cdot p_{ik} (\delta_{mk} - p_{im}) x$$

$$= (p_{im} - y_{im}) x$$

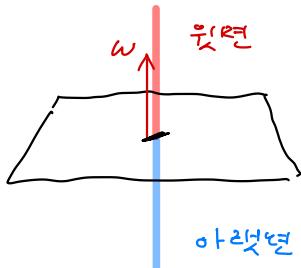
$$\nabla_{w_m} (-\log \mathcal{L}) = \sum_{i=1}^N (p_{im} - y_{im}) x_i$$

$\text{conf}(x_i) \quad \sum_i z_{ii}$

$$\text{SVM} \quad \text{정의} \quad w^T x + b = 0$$

정의 ①

$$\forall i \quad y_i (w^T x_i + b) \geq 0$$

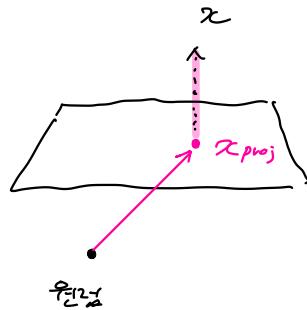


$$y = 1 \quad w^T x + b \geq 0$$

$$y = -1 \quad w^T x + b \leq 0$$

정의 ②

$$\min_w |w^T x + b| = 1$$



$$x = x_{\text{proj}} + d \frac{w}{\|w\|}$$

$$d = \frac{|w^T x + b|}{\|w\|}$$

$$= \frac{1}{\|w\|} \quad (\text{가장 가까운 점})$$

$$\frac{1}{2} \|w\|^2$$

$$y_i (w^T x_i + b) \geq 1$$

$$\text{minimize} \quad \frac{1}{2} \|w\|^2 + \max_{\alpha_i \geq 0} \alpha_i (1 - y_i (w^T x_i + b)) = \begin{cases} 0 & y_i (w^T x_i + b) > 1 \\ \infty & \text{else} \end{cases}$$

$$\min_{w, b} \max_{\alpha_i \geq 0} \left[ \frac{1}{2} \|w\|^2 + \sum_i \alpha_i (1 - y_i (w^T x_i + b)) \right] \geq \max_{\alpha_i \geq 0} \min_{w, b} \left[ \dots \right]$$

$$= \frac{1}{2} w^T w + \sum_i \alpha_i - w^T \left( \sum_i \alpha_i y_i x_i \right) - b \sum_i \alpha_i$$

maximize

$$\sum_i \alpha_i - \frac{1}{2} w^T w$$

with

$$\left\{ \begin{array}{l} \nabla_w = w - \sum_i \alpha_i y_i x_i = 0 \\ \frac{\partial}{\partial b} = \sum_i \alpha_i y_i = 0 \end{array} \right.$$

say  $i \in \mathcal{I}$  whenever  $\alpha_i > 0$

$$\text{i.e. } y_i(w^T x_i + b) = 1$$

then  $w = \sum_i \alpha_i y_i x_i \quad (i \in \mathcal{I})$

$$b = y_i - w^T x_i \quad (i \in \mathcal{I})$$

$$w^T w = \sum_i \alpha_i y_i (w^T x_i) \quad (i \in \mathcal{I})$$

$$= \sum_i \alpha_i (1 - y_i b) \quad \text{because } y_i(w^T x_i + b) = 1$$

$$= \sum_i \alpha_i - b \sum_i \alpha_i y_i$$

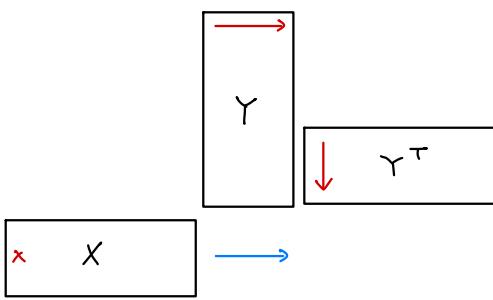
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$$\frac{|w^T w + b|}{\|w\|} = \frac{1}{\sqrt{\sum \alpha_i}}$$

say  $W = XY$

then  $\frac{\partial L}{\partial X} = \frac{\partial L}{\partial W} \cdot Y^T$

$$\frac{\partial L}{\partial Y} = X^T \frac{\partial L}{\partial W}$$



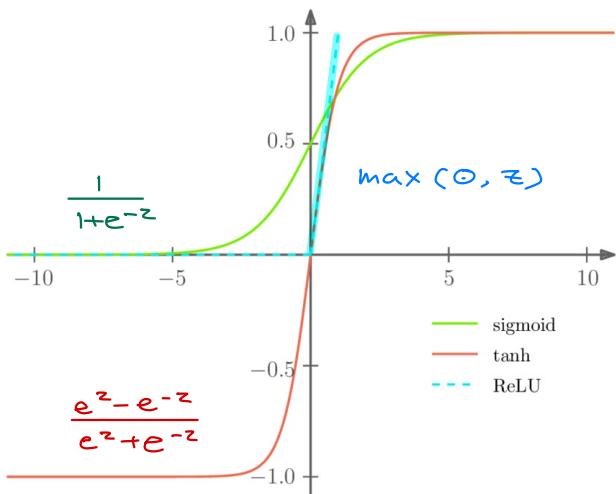
minimize

Gradient Descent  $\theta_{k+1} = \theta_k - \alpha \nabla L(\theta_k)$

Full Batch GD  $\tilde{\nabla} f(x) = \frac{1}{N} \sum_{i=1}^N \nabla f_i(x)$

Stochastic GD  $\tilde{\nabla} f(x) = \nabla f_i(x)$

Mini Batch GD



$$\delta' = \delta(1-\delta)$$

$$(\tanh)' = 1 - (\tanh)^2$$

$$(\text{ReLU})' = \begin{cases} 1 & z > 0 \\ 0 & \text{else} \end{cases}$$

Note.  $\delta \leftarrow \tanh \frac{z}{2}$

$(\tanh \frac{z}{2})$  linear

Let  $z = \sum_{i=1}^{\text{Dim}} w_i x_i$  s.t.  $\begin{cases} w: \text{i.i.d, } E[w_i] = 0 \quad \text{Var}[w_i] = \sigma^2 \\ x: \text{i.i.d independent of } w \\ E[x_i] \neq 0 \text{ (in general)} \end{cases}$

Let  $S = E[z^2]$  denote the 2nd moment of the input.

We know that  $E[z] = 0$

Note. independent

then  $\text{Var}(z) = \sum_{i=1}^{\text{Dim}} \text{Var}(w_i x_i)$

$$= \sum_{i=1}^{\text{Dim}} E[w_i^2 x_i^2] - E[w_i x_i]^2 = \text{Dim } \sigma^2 S$$

$\downarrow \quad \downarrow$

$$\frac{E[w_i^2]}{\sigma^2} \frac{E[x_i^2]}{S} \quad \frac{E[w_i]^2}{0} \frac{E[x_i]^2}{0}$$

$$\begin{aligned} & \iint xy p(x,y) dx dy \\ &= \iint xy p(x) p(y) dx dy \end{aligned}$$

in Xavier Init we want  $E[z^2] = E[x^2]$

$$\text{Dim } \sigma^2 S = S \quad \sigma = \frac{1}{\sqrt{\text{Dim}}}$$

now, let  $h = \text{ReLU}(z)$

then  $E[h^2] = \int_0^\infty z^2 p(z) dz$

$$= \frac{1}{2} \int_{-\infty}^\infty z^2 p(z) dz = \frac{1}{2} E[z^2]$$

$$= \frac{1}{2} \text{Dim } \sigma^2 S$$

$$\begin{cases} \text{Xavier} & \sigma_{\text{ref}}^2 = \frac{1}{\text{Dim}} \\ \text{He} & \sigma_{\text{ref}}^2 = \frac{2}{\text{Dim}} \end{cases}$$

then  $w \sim N(0, \sigma_{\text{ref}}^2)$

or  $w \sim \text{Uni}(-\sqrt{3}\sigma_{\text{ref}}, \sqrt{3}\sigma_{\text{ref}})$

in He Init we want  $E[h^2] = E[x^2]$

$$\frac{1}{2} \text{Dim } \sigma^2 S = S \quad \sigma = \sqrt{\frac{2}{\text{Dim}}}$$

as variance of  $\text{Uni}(-b, b)$  becomes

$$\begin{aligned} & \int_{-b}^b x^2 \cdot \frac{1}{2b} dx \\ &= \left[ \frac{x^3}{6b} \right]_{-b}^b = \frac{b^3}{3} \end{aligned}$$

learning rate decay

① at a few fixed points?

② cosine :  $\frac{1}{2} \alpha_0 \left( 1 + \cos \left( \frac{t\pi}{T} \right) \right)$

③ linear :  $\alpha_0 \left( 1 - \frac{t}{T} \right)$

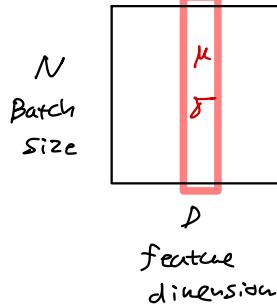
④ linear sqrt :  $\alpha_0 / \sqrt{t}$

linear warmup

: linearly increase

learning steps at  $t \approx 0$

Batch Normalization



$$\tilde{x}_{ij} = \frac{x_{ij} - \mu_j}{\sqrt{\delta_j^2 + \epsilon}}$$

$$\tilde{y}_{ij} = \gamma_j \tilde{x}_{ij} + \beta_j$$

(training)

$$\mu_{\text{run}} = m \cdot \mu_{\text{run}} + (1-m) \mu_{\text{batch}}$$

$$(\sigma^2)_{\text{run}} = m (\sigma^2)_{\text{run}} + (1-m) (\sigma^2)_{\text{batch}}$$

BN  $\Rightarrow$  Internal

Convergence Shift

shift in the  
mean/var of  
hidden activation

$\Rightarrow$  Smooth

loss landscape

: stable  
gradient

GD minimizes  $f(x)$

$$x_{t+1} = x_t + \gamma v$$

to minimize  $f(x_{t+1}) \approx f(x_t) + \langle \nabla f(x_t), \gamma v \rangle$

GD idea

Set  $v = - \frac{\nabla f(x_t)}{\| \nabla f(x_t) \|}$

Lemma 3.1

$f: \mathbb{R}^d \rightarrow \mathbb{R}$  continuously

differentiable function.

Let  $f$   $\beta$ -smooth :  $|\nabla f(x) - \nabla f(y)| \leq \beta \|x - y\|$   $(x, y)$

then  $f(y) \leq f(x) + \underbrace{\langle \nabla f(x), y - x \rangle}_{\text{梯度}} + \underbrace{\frac{\beta}{2} \|y - x\|^2}_{\text{凸性}}$

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GD :  $x_{t+1} = x_t - \eta \nabla f(x_t)$

$$f(x_{t+1}) \leq f(x_t) + \langle \nabla f(x_t), -\eta \nabla f(x_t) \rangle + \frac{\beta}{2} \|\eta \nabla f(x_t)\|^2$$

$$= \dots = f(x_t) - \left(\eta - \frac{\beta}{2} \eta^2\right) \|\nabla f(x_t)\|^2$$

small step

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SGD :  $x_{t+1} = x_t - \eta \tilde{\nabla} f(x_t)$

$$\text{then } f(x_{t+1}) \leq f(x_t) - \eta \langle \nabla f(x_t), \tilde{\nabla} f(x_t) \rangle + \frac{\beta}{2} \cdot \eta^2 \|\tilde{\nabla} f(x_t)\|^2$$

$$E_t[f(x_{t+1})] \leq f(x_t) - \eta \|\nabla f(x_t)\|^2 + \underbrace{\frac{\beta}{2} \eta^2 E_t[\|\tilde{\nabla} f(x_t)\|^2]}_{=: G}$$

$$E[\|\nabla f(x_t)\|^2] \leq \frac{1}{\eta} (E[f(x_t)] - E[f(x_{t+1})]) + \frac{\beta}{2} \eta^2 G$$

(lower bound)

$$\sum_{t=0}^T E[\|\nabla f(x_t)\|^2] \leq \frac{1}{\eta} (f_0 - f^*) + \frac{\beta}{2} \eta^2 G T$$

$$\text{let } \eta = \frac{1}{\sqrt{T}}$$

$$\min_t (E[\|\nabla f(x_t)\|^2]) \leq \frac{1}{\eta T} (f_0 - f^*) + \frac{\beta}{2} \eta^2 G$$

$$\text{then } \min (E[\|\nabla f(x_t)\|^2]) = \mathcal{O}(\frac{1}{\sqrt{T}})$$

$$g(t) = f(x + t(y-x))$$

$$g'(t) = (y-x)^T \nabla f(x + t(y-x))$$

$$g''(t) = (y-x)^T \nabla^2 f(x + t(y-x)) (y-x) \leq \beta \|y-x\|^2$$


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$$\begin{aligned} \int_0^1 (1-s) g''(s) ds &= \left[ (1-s) g'(s) \right]_0^1 + \int_0^1 g'(s) ds \\ &= g(1) - g(0) - g'(0) \end{aligned}$$

$$\begin{aligned} \text{hence } f(y) &= f(x) + \langle \nabla f(x), y-x \rangle + \int_0^1 (1-s) (y-x)^T \nabla^2 f(x + s(y-x)) (y-x) ds \\ &\leq \int_0^1 (1-s) \beta \|y-x\|^2 ds \\ &= \frac{\beta}{2} \|y-x\|^2 \end{aligned}$$


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$$\text{let } \phi(t) = \nabla f(x + t(y-x))$$

$$\phi'(t) = \nabla^2 f(x + t(y-x)) (y-x)$$

$$\text{then } \nabla f(y) - \nabla f(x) = \int_0^1 \phi'(t) dt$$

$$= \int_0^1 \nabla^2 f(x + t(y-x)) dt \cdot (y-x)$$

$$\begin{aligned} (y-x)^T \int_0^1 \nabla^2 f(x + t(y-x)) dt \cdot (y-x) \\ = \langle \nabla f(y) - \nabla f(x), y-x \rangle \leq \| \nabla f(y) - \nabla f(x) \| \|y-x\| \\ \leq \beta \|y-x\|^2 \end{aligned}$$

GD

+Momentum

Nesterov Momentum

$$x_{t+1} = x_t - \alpha \nabla f(x_t)$$

$$x_{t+1} = x_t + v_{t+1}$$

$$x_{t+1} = x_t + v_{t+1}$$

$$v_{t+1} = \rho v_t - \alpha \nabla f(x_t)$$

$$v_{t+1} = \rho v_t - \alpha \nabla f(x_t + \rho v_t)$$



장속운동

$\|\nabla f(x)\|$  . 가파른 경로 .

“가파른 경로는 느리게  
단단한 경로는 빠르게”

AdaGrad

$$x = x - \alpha \frac{\nabla f(x_t)}{\sqrt{\sum (\nabla f(x_t))^2}}$$

$\rightarrow \infty$

RMSProp

$$x = x - \frac{\eta g_t}{\sqrt{E[g^2]_t}}$$

$$E[g^2]_t = \beta E[g^2]_{t-1} + (1-\beta) g_t^2$$

Adam

$$m_t = \beta_1 m_{t-1} + (1-\beta_1) g_t$$

$$m_0 = 0$$

$$v_t = \beta_2 v_{t-1} + (1-\beta_2) g_t^2$$

$$v_0 = 0$$

$$\hat{m}_t = \frac{m_t}{1-\beta_1^t}$$

$$\hat{v}_t = \frac{v_t}{1-\beta_2^t}$$

if  $g$  const

then

$$\begin{cases} m_k = (1-\beta_1^k) g \\ v_k = (1-\beta_2^k) g^2 \end{cases}$$