

10907 Pattern Recognition

Lecturers

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Problem set 6

Math

Exercise 1 (Patchwise MLPs as convolutional nets $\star\star$).

Let $x \in \mathbb{R}^{H \times W \times 1}$ be a single-channel image and let k be a patch size (so we work with square patches of size $k \times k$). We will define an image-to-image architecture which uses an MLP instead of a convnet. It will simply apply the same MLP to all overlapping patches. For each *valid* location (i, j) (so $0 \leq i \leq H - k$, $0 \leq j \leq W - k$), define the patch

$$P_{ij} := x[i : i + k - 1, j : j + k - 1] \in \mathbb{R}^{k \times k}, \quad p_{ij} := \text{vec}(P_{ij}) \in \mathbb{R}^{k^2}.$$

Consider an MLP applied with *shared weights* to every patch:

$$\begin{aligned} z_{ij}^{(1)} &= W_1 p_{ij} + b_1, \quad W_1 \in \mathbb{R}^{D \times k^2}, \quad b_1 \in \mathbb{R}^D, \\ a_{ij} &= \sigma(z_{ij}^{(1)}), \\ y_{ij} &= W_2 a_{ij} + b_2, \quad W_2 \in \mathbb{R}^{C_{\text{out}} \times D}, \quad b_2 \in \mathbb{R}^{C_{\text{out}}}. \end{aligned}$$

where σ is a pointwise activation function.

Show that

1. *The first layer of the MLP can be implemented as a convolution:* Concretely: each row of the matrix W_1 is a (vectorized) filter impulse response, and the D rows constitute the D channels at the output of this layer.
2. *The second layer can be implemented as a 1×1 convolution.*

Now consider an MLP with L layers.

$$z_{ij}^{(1)} = W_1 p_{ij} + b_1, \quad z_{ij}^{(\ell+1)} = W_{\ell+1} \sigma(z_{ij}^{(\ell)}) + b_{\ell+1}, \quad \ell = 1, \dots, L-1.$$

Argue that the network is equivalent to: one $k \times k$ convolution for W_1 , followed by $L-1$ many 1×1 convolutions for W_2, \dots, W_L , with σ between them. In other words, this MLP-based architecture is a convnet in disguise!

Solution:

1. For a valid location (i, j) , the output of the d^{th} feature is

$$z_{ij}^{(1)}[d] = W_1[d, :] p_{ij} + b_1[d],$$

where $W_1[d, :] \in \mathbb{R}^{k^2}$. Assume $\text{vec}(\cdot)$ uses appropriate ordering and reshape $W_1[d, :]$ into a $k \times k$ matrix \tilde{W}_d defined by

$$\tilde{W}_d[l, m] = W_1[d, lk + m], \quad l, m \in \{0, 1, \dots, k-1\}.$$

Then

$$\begin{aligned} W_1[d, :] p_{ij} &= \sum_{n=0}^{k^2-1} W_1[d, n] p_{ij}[n] \\ &= \sum_{l=0}^{k-1} \sum_{m=0}^{k-1} \tilde{W}_d[l, m] P_{ij}[l, m]. \end{aligned}$$

Using the definition of the patch $P_{ij}[l, m] = x[i + l, j + m]$, we obtain

$$z_{ij}^{(1)}[d] = \sum_{l=0}^{k-1} \sum_{m=0}^{k-1} \tilde{W}_d[l, m] x[i + l, j + m] + b_1[d].$$

Here, \tilde{W}_d is a valid $k \times k$ convolution kernel and bias $b_1[d]$. Repeating this for $d = 1, \dots, D$ yield D filters and therefore D output channels.

2. A 1×1 convolution operates only along the channel dimension. The activation

$$a_{ij} \in \mathbb{R}^D$$

represents the D feature channels at spatial location (i, j) . The second layer computes

$$y_{ij} = W_2 a_{ij} + b_2, \quad W_2 \in \mathbb{R}^{C_{\text{out}} \times D}.$$

Define a 1×1 convolution kernel $\tilde{W}_2 \in \mathbb{R}^{C_{\text{out}} \times D \times 1 \times 1}$ by

$$\tilde{W}_2[c, d, 0, 0] = W_2[c, d].$$

Then for each output channel c ,

$$y_{ij}[c] = \sum_{d=0}^{D-1} \tilde{W}_2[c, d, 0, 0] a_{ij}[d] + b_2[c],$$

which is exactly the definition of a 1×1 convolution.

3. Now consider an L -layer MLP applied with shared weights to all patches:

$$z_{ij}^{(1)} = W_1 p_{ij} + b_1, \quad z_{ij}^{(\ell+1)} = W_{\ell+1} \sigma(z_{ij}^{(\ell)}) + b_{\ell+1}, \quad \ell = 1, \dots, L-1.$$

From part (1), the first layer W_1 depends on the $k \times k$ neighbourhood of x around (i, j) and is therefore equivalent to a valid $k \times k$ convolution producing feature maps $z^{(1)}$.

For every subsequent layer $\ell \geq 2$, the mapping

$$z_{ij}^{(\ell)} = W_{\ell} \sigma(z_{ij}^{(\ell-1)}) + b_{\ell}$$

applies the same affine transformation independently at each spatial location (i, j) and does not mix spatial information. From (2), each such layer is equivalent to a 1×1 convolution followed by the nonlinearity σ .

Hence, the entire network is equivalent to a convolutional architecture consisting of:

- one $k \times k$ convolution implementing W_1 and b_1 ,
- followed by $L - 1$ successive 1×1 convolutions implementing W_2, \dots, W_L ,
- with the activation function σ applied between layers.

Therefore, an MLP applied to all overlapping patches with shared weights is a convolutional network in disguise.

Exercise 2 (Closed-form forward noising process \star).

Recall the forward diffusion process in $1D$,

$$x_t = \sqrt{\alpha_t} x_{t-1} + \sqrt{1 - \alpha_t} \epsilon_t, \quad \epsilon_t \sim \mathcal{N}(0, 1)$$

with $0 < \alpha_t < 1$ (NB: close to 1), and all ϵ_t independent of each other and of x_0 .

1. Show that, conditional on x_{t-1} , x_t is normally distributed. Concretely,

$$x_t | x_{t-1} \sim \mathcal{N}(\sqrt{\alpha_t} x_{t-1}, 1 - \alpha_t).$$

2. Let $\bar{\alpha}_t = \prod_{s=1}^t \alpha_s$. Prove by induction that

$$x_t | x_0 \sim \mathcal{N}(\sqrt{\bar{\alpha}_t} x_0, 1 - \bar{\alpha}_t)$$

Solution:

1. We are given the forward diffusion update

$$x_t = \sqrt{\alpha_t} x_{t-1} + \sqrt{1 - \alpha_t} \epsilon_t, \quad \epsilon_t \sim \mathcal{N}(0, 1),$$

where ϵ_t is independent of x_{t-1} . Thus,

$$\sqrt{1 - \alpha_t} \epsilon_t \sim \mathcal{N}(0, 1 - \alpha_t).$$

Conditioning on x_{t-1} amounts to treating it as deterministic, which shifts the mean of the Gaussian. Therefore,

$$x_t | x_{t-1} = \sqrt{\alpha_t} x_{t-1} + \sqrt{1 - \alpha_t} \epsilon_t \sim \mathcal{N}\left(\sqrt{\alpha_t} x_{t-1}, 1 - \alpha_t\right).$$

2. Define $\bar{\alpha}_t = \prod_{s=1}^t \alpha_s$. We prove by induction on t that

$$x_t | x_0 \sim \mathcal{N}\left(\sqrt{\bar{\alpha}_t} x_0, 1 - \bar{\alpha}_t\right).$$

Base case ($t = 1$). Using the update with $t = 1$,

$$x_1 = \sqrt{\alpha_1} x_0 + \sqrt{1 - \alpha_1} \epsilon_1.$$

Conditioning on x_0 , the first term is deterministic and the second is Gaussian with variance $1 - \alpha_1$. Hence,

$$x_1 | x_0 \sim \mathcal{N}\left(\sqrt{\alpha_1} x_0, 1 - \alpha_1\right).$$

Since $\bar{\alpha}_1 = \alpha_1$, this matches the claim.

Inductive step. Assume for some $t - 1 \geq 1$ that

$$x_{t-1} | x_0 \sim \mathcal{N}\left(\sqrt{\bar{\alpha}_{t-1}} x_0, 1 - \bar{\alpha}_{t-1}\right).$$

Using the recursion,

$$x_t = \sqrt{\alpha_t} x_{t-1} + \sqrt{1 - \alpha_t} \epsilon_t,$$

and conditioning on x_0 , we have:

- By the inductive hypothesis, $x_{t-1} | x_0$ is Gaussian, so

$$\sqrt{\alpha_t} x_{t-1} | x_0 \sim \mathcal{N}\left(\sqrt{\alpha_t \bar{\alpha}_{t-1}} x_0, \alpha_t(1 - \bar{\alpha}_{t-1})\right).$$

- Also, ϵ_t is independent of x_{t-1} and x_0 , hence

$$\sqrt{1 - \alpha_t} \epsilon_t \mid x_0 \sim \mathcal{N}(0, 1 - \alpha_t),$$

and it is independent of $\sqrt{\alpha_t} x_{t-1} \mid x_0$.

The sum of independent Gaussians is Gaussian, with mean equal to the sum of means and variance equal to the sum of variances. Therefore,

$$\begin{aligned} x_t \mid x_0 &\sim \mathcal{N}\left(\sqrt{\alpha_t \bar{\alpha}_{t-1}} x_0, \alpha_t(1 - \bar{\alpha}_{t-1}) + (1 - \alpha_t)\right) \\ &= \mathcal{N}\left(\sqrt{\alpha_t \bar{\alpha}_{t-1}} x_0, \alpha_t - \alpha_t \bar{\alpha}_{t-1} + 1 - \alpha_t\right) \\ &= \mathcal{N}\left(\sqrt{\bar{\alpha}_t} x_0, 1 - \bar{\alpha}_t\right), \end{aligned}$$

since $\bar{\alpha}_t = \alpha_t \bar{\alpha}_{t-1}$. This completes the induction.

Exercise 3 (What is normal and what is not **).

We again work with the 1D diffusion process from the last exercise, and we want to show that the transition probabilities for the reverse process are Gaussian *IF* we know and condition on x_0 (which in practice we don't—we want to sample it).

- Using the result of the previous exercise for $t - 1$ and t , write both x_{t-1} and x_t as affine functions of x_0 and independent standard normals. That is to say, show that there are some numbers a, b, c, d, e such that

$$x_{t-1} = a x_0 + b z_1, \quad x_t = c x_0 + d z_1 + e z_2,$$

with independent $z_1, z_2 \sim \mathcal{N}(0, 1)$.

- Use this representation to compute the following quantities conditional on x_0 :

$$\mathbb{E}[x_{t-1} \mid x_0], \quad \mathbb{E}[x_t \mid x_0], \quad \text{Var}(x_{t-1} \mid x_0), \quad \text{Var}(x_t \mid x_0), \quad \text{Cov}(x_{t-1}, x_t \mid x_0).$$

- Recall that if $\begin{pmatrix} u \\ v \end{pmatrix}$ is jointly Gaussian, then

$$U \mid V = v \sim \mathcal{N}\left(\mathbb{E}[U] + \frac{\text{Cov}(U, V)}{\text{Var}(V)}(v - \mathbb{E}[V]), \text{Var}(U) - \frac{\text{Cov}(U, V)^2}{\text{Var}(V)}\right).$$

Use this with $U = x_{t-1}$, $V = x_t$ to show that

$$x_{t-1} \mid (x_t, x_0)$$

is Gaussian with the mean

$$\tilde{\mu}_t(x_t, x_0) = \frac{1}{\sqrt{\alpha_t}} \left(x_t - \frac{1 - \alpha_t}{1 - \bar{\alpha}_t} (x_t - \sqrt{\bar{\alpha}_t} x_0) \right).$$

Solution:

- From the previous exercise we know that, for $\bar{\alpha}_t = \prod_{s=1}^t \alpha_s$,

$$x_t \mid x_0 \sim \mathcal{N}\left(\sqrt{\bar{\alpha}_t} x_0, 1 - \bar{\alpha}_t\right), \quad x_{t-1} \mid x_0 \sim \mathcal{N}\left(\sqrt{\bar{\alpha}_{t-1}} x_0, 1 - \bar{\alpha}_{t-1}\right).$$

Using $z_1 \sim \mathcal{N}(0, 1)$, we can write:

$$x_{t-1} = \sqrt{\bar{\alpha}_{t-1}} x_0 + \sqrt{1 - \bar{\alpha}_{t-1}} z_1.$$



Using the forward update

$$x_t = \sqrt{\alpha_t} x_{t-1} + \sqrt{1 - \alpha_t} \epsilon_t, \quad \epsilon_t \sim \mathcal{N}(0, 1),$$

with ϵ_t independent of x_{t-1} and x_0 . Substitute the expression for x_{t-1} :

$$\begin{aligned} x_t &= \sqrt{\alpha_t} \left(\sqrt{\bar{\alpha}_{t-1}} x_0 + \sqrt{1 - \bar{\alpha}_{t-1}} z_1 \right) + \sqrt{1 - \alpha_t} \epsilon_t \\ &= \sqrt{\alpha_t \bar{\alpha}_{t-1}} x_0 + \sqrt{\alpha_t (1 - \bar{\alpha}_{t-1})} z_1 + \sqrt{1 - \alpha_t} \epsilon_t. \end{aligned}$$

Let $z_2 := \epsilon_t$, so $z_2 \sim \mathcal{N}(0, 1)$ and z_2 is independent of z_1 . Since $\bar{\alpha}_t = \alpha_t \bar{\alpha}_{t-1}$, we have:

$$x_{t-1} = a x_0 + b z_1, \quad x_t = c x_0 + d z_1 + e z_2,$$

with

$$a = \sqrt{\bar{\alpha}_{t-1}}, \quad b = \sqrt{1 - \bar{\alpha}_{t-1}}, \quad c = \sqrt{\bar{\alpha}_t}, \quad d = \sqrt{\alpha_t (1 - \bar{\alpha}_{t-1})}, \quad e = \sqrt{1 - \alpha_t}.$$

2. Using the representation with independent $z_1, z_2 \sim \mathcal{N}(0, 1)$ and conditioning on x_0 :

$$x_{t-1} = a x_0 + b z_1, \quad x_t = c x_0 + d z_1 + e z_2.$$

Since $\mathbb{E}[z_1] = \mathbb{E}[z_2] = 0$ and $\text{Var}(z_1) = \text{Var}(z_2) = 1$, we get

$$\mathbb{E}[x_{t-1} | x_0] = a x_0 = \sqrt{\bar{\alpha}_{t-1}} x_0, \quad \mathbb{E}[x_t | x_0] = c x_0 = \sqrt{\bar{\alpha}_t} x_0.$$

Also,

$$\text{Var}(x_{t-1} | x_0) = \text{Var}(b z_1) = b^2 = 1 - \bar{\alpha}_{t-1},$$

and, using independence of z_1 and z_2 ,

$$\text{Var}(x_t | x_0) = \text{Var}(d z_1 + e z_2) = d^2 + e^2 = \alpha_t (1 - \bar{\alpha}_{t-1}) + (1 - \alpha_t) = 1 - \bar{\alpha}_t.$$

The conditional covariance is

$$\begin{aligned} \text{Cov}(x_{t-1}, x_t | x_0) &= \text{Cov}(a x_0 + b z_1, c x_0 + d z_1 + e z_2 | x_0) \\ &= \text{Cov}(b z_1, d z_1 + e z_2) \\ &= bd \text{Var}(z_1) + be \text{Cov}(z_1, z_2) \\ &= bd \\ &= \sqrt{1 - \bar{\alpha}_{t-1}} \sqrt{\alpha_t (1 - \bar{\alpha}_{t-1})} = \sqrt{\bar{\alpha}_t} (1 - \bar{\alpha}_{t-1}). \end{aligned}$$

3. Conditioning on x_0 , $(x_{t-1}, x_t) | x_0$ is jointly Gaussian because it is an affine function of the jointly Gaussian vector (z_1, z_2) . Assume $U = x_{t-1}$ and $V = x_t$.

We have

$$\mathbb{E}[U | x_0] = \sqrt{\bar{\alpha}_{t-1}} x_0, \quad \mathbb{E}[V | x_0] = \sqrt{\bar{\alpha}_t} x_0,$$

$$\text{Var}(U | x_0) = 1 - \bar{\alpha}_{t-1}, \quad \text{Var}(V | x_0) = 1 - \bar{\alpha}_t,$$

$$\text{Cov}(U, V | x_0) = \sqrt{\bar{\alpha}_t} (1 - \bar{\alpha}_{t-1}).$$

Hence,

$$\begin{aligned} \mathbb{E}[x_{t-1} | x_t, x_0] &= \mathbb{E}[U | x_0] + \frac{\text{Cov}(U, V | x_0)}{\text{Var}(V | x_0)} (x_t - \mathbb{E}[V | x_0]) \\ &= \sqrt{\bar{\alpha}_{t-1}} x_0 + \frac{\sqrt{\bar{\alpha}_t} (1 - \bar{\alpha}_{t-1})}{1 - \bar{\alpha}_t} (x_t - \sqrt{\bar{\alpha}_t} x_0). \end{aligned}$$

Now use $\bar{\alpha}_t = \alpha_t \bar{\alpha}_{t-1}$ so that $\sqrt{\bar{\alpha}_{t-1}} = \frac{\sqrt{\bar{\alpha}_t}}{\sqrt{\alpha_t}}$, and also

$$1 - \bar{\alpha}_t = (1 - \alpha_t) + \alpha_t(1 - \bar{\alpha}_{t-1}) \implies \alpha_t(1 - \bar{\alpha}_{t-1}) = (1 - \bar{\alpha}_t) - (1 - \alpha_t).$$

Therefore,

$$\begin{aligned} \mathbb{E}[x_{t-1} | x_t, x_0] &= \frac{1}{\sqrt{\alpha_t}} \sqrt{\bar{\alpha}_t} x_0 + \frac{1}{\sqrt{\alpha_t}} \frac{\alpha_t(1 - \bar{\alpha}_{t-1})}{1 - \bar{\alpha}_t} (x_t - \sqrt{\bar{\alpha}_t} x_0) \\ &= \frac{1}{\sqrt{\alpha_t}} \left[\sqrt{\bar{\alpha}_t} x_0 + \frac{(1 - \bar{\alpha}_t) - (1 - \alpha_t)}{1 - \bar{\alpha}_t} (x_t - \sqrt{\bar{\alpha}_t} x_0) \right] \\ &= \frac{1}{\sqrt{\alpha_t}} \left[\sqrt{\bar{\alpha}_t} x_0 + \left(1 - \frac{1 - \alpha_t}{1 - \bar{\alpha}_t}\right) (x_t - \sqrt{\bar{\alpha}_t} x_0) \right] \\ &= \frac{1}{\sqrt{\alpha_t}} \left[\sqrt{\bar{\alpha}_t} x_0 + \left(x_t - \sqrt{\bar{\alpha}_t} x_0\right) - \frac{1 - \alpha_t}{1 - \bar{\alpha}_t} (x_t - \sqrt{\bar{\alpha}_t} x_0) \right] \\ &= \frac{1}{\sqrt{\alpha_t}} \left[x_t - \frac{1 - \alpha_t}{1 - \bar{\alpha}_t} (x_t - \sqrt{\bar{\alpha}_t} x_0) \right]. \end{aligned}$$

Thus $x_{t-1} | (x_t, x_0)$ is Gaussian with mean

$$\tilde{\mu}_t(x_t, x_0) = \frac{1}{\sqrt{\alpha_t}} \left(x_t - \frac{1 - \alpha_t}{1 - \bar{\alpha}_t} (x_t - \sqrt{\bar{\alpha}_t} x_0) \right),$$

Exercise 4 (Predicting x_{t-1} vs predicting noise ε).

Consider a single step of the 1D forward process

$$x_t = \sqrt{\alpha_t} x_{t-1} + \sqrt{\beta_t} \varepsilon_t, \quad \beta_t = 1 - \alpha_t, \quad \varepsilon_t \sim \mathcal{N}(0, 1).$$

We compare two possible training objectives:

1. directly regress x_{t-1} from x_t ,
2. predict the noise ε_t from x_t and estimate x_{t-1} as $(x_t - \text{estimated noise})$

Assume we parameterize the reverse mean via a noise-predictor network $\varepsilon_\theta(x_t, t)$ as

$$\mu_\theta(x_t, t) = \frac{1}{\sqrt{\alpha_t}} \left(x_t - \sqrt{\beta_t} \varepsilon_\theta(x_t, t) \right).$$

1. Express the true x_{t-1} in the same form:

$$x_{t-1} = \frac{1}{\sqrt{\alpha_t}} \left(x_t - \sqrt{\beta_t} \varepsilon_t \right).$$

2. Show that

$$x_{t-1} - \mu_\theta(x_t, t) = \frac{\sqrt{\beta_t}}{\sqrt{\alpha_t}} (\varepsilon_\theta(x_t, t) - \varepsilon_t).$$

3. Deduce that the two MSE losses

$$L_x(\theta) = \mathbb{E}[(x_{t-1} - \mu_\theta(x_t, t))^2], \quad L_\varepsilon(\theta) = \mathbb{E}[(\varepsilon_t - \varepsilon_\theta(x_t, t))^2]$$

are proportional:

$$L_x(\theta) = \frac{\beta_t}{\alpha_t} L_\varepsilon(\theta),$$

where the factor β_t/α_t does not depend on θ . Conclude that minimizing MSE on x_{t-1} is equivalent to minimizing MSE on ε_t in the sense that they will give the same θ and ultimately the same estimate of x_{t-1} .

Solution:

1. we solve for x_{t-1} by rearranging the above equation:

$$\sqrt{\alpha_t} x_{t-1} = x_t - \sqrt{\beta_t} \varepsilon_t.$$

Dividing by $\sqrt{\alpha_t}$ gives

$$x_{t-1} = \frac{1}{\sqrt{\alpha_t}} (x_t - \sqrt{\beta_t} \varepsilon_t).$$

2. By definition, the parameterized reverse mean is

$$\mu_\theta(x_t, t) = \frac{1}{\sqrt{\alpha_t}} (x_t - \sqrt{\beta_t} \varepsilon_\theta(x_t, t)).$$

Subtracting $\mu_\theta(x_t, t)$ from the true x_{t-1} (from part (1)):

$$\begin{aligned} x_{t-1} - \mu_\theta(x_t, t) &= \frac{1}{\sqrt{\alpha_t}} (x_t - \sqrt{\beta_t} \varepsilon_t) - \frac{1}{\sqrt{\alpha_t}} (x_t - \sqrt{\beta_t} \varepsilon_\theta(x_t, t)) \\ &= \frac{1}{\sqrt{\alpha_t}} (-\sqrt{\beta_t} \varepsilon_t + \sqrt{\beta_t} \varepsilon_\theta(x_t, t)) \\ &= \frac{\sqrt{\beta_t}}{\sqrt{\alpha_t}} (\varepsilon_\theta(x_t, t) - \varepsilon_t). \end{aligned}$$

3. The two MSE losses are:

$$L_x(\theta) = \mathbb{E}[(x_{t-1} - \mu_\theta(x_t, t))^2], \quad L_\varepsilon(\theta) = \mathbb{E}[(\varepsilon_t - \varepsilon_\theta(x_t, t))^2].$$

From part (2):

$$x_{t-1} - \mu_\theta(x_t, t) = \frac{\sqrt{\beta_t}}{\sqrt{\alpha_t}} (\varepsilon_\theta(x_t, t) - \varepsilon_t),$$

and squaring both sides gives

$$(x_{t-1} - \mu_\theta(x_t, t))^2 = \frac{\beta_t}{\alpha_t} (\varepsilon_\theta(x_t, t) - \varepsilon_t)^2.$$

Taking expectation on both sides,

$$\begin{aligned} L_x(\theta) &= \mathbb{E}[(x_{t-1} - \mu_\theta(x_t, t))^2] \\ &= \mathbb{E}\left[\frac{\beta_t}{\alpha_t} (\varepsilon_\theta(x_t, t) - \varepsilon_t)^2\right] \\ &= \frac{\beta_t}{\alpha_t} \mathbb{E}[(\varepsilon_\theta(x_t, t) - \varepsilon_t)^2] \\ &= \frac{\beta_t}{\alpha_t} L_\varepsilon(\theta). \end{aligned}$$

The factor β_t/α_t does not depend on θ . Therefore, minimizing $L_x(\theta)$ w.r.t θ is equivalent to minimizing $L_\varepsilon(\theta)$ w.r.t θ .

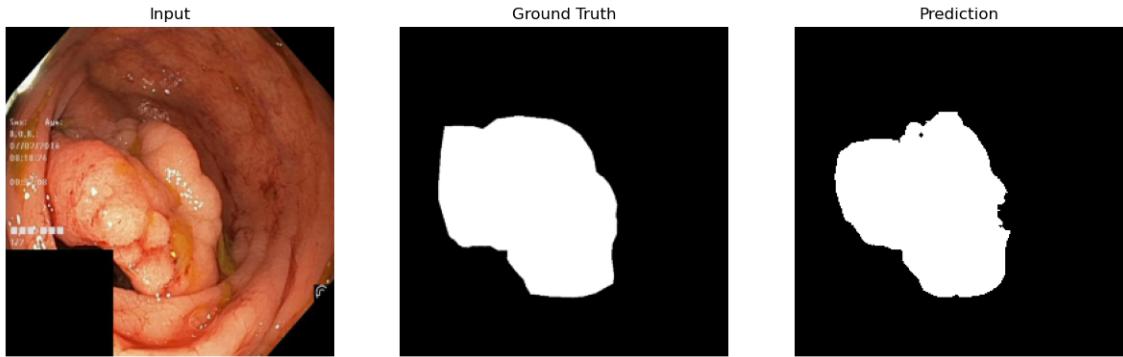


Figure 1: Segmentation and prediction for a sample image in the dataset.

Coding

For this problem set, we do not provide Gradescope autograding, because the expected outputs are plots and other visualizations, which are difficult to assess automatically. Instead, we will include reference plots in the model solution.

Exercise 5 (Segmentation — **).

Pixel-wise image segmentation is an important computer vision task that involves dividing an image into different segments representing various objects or regions of interest. Such a method can aid in medical diagnostic scenarios where experts can sift through large amounts of data quickly or assist in the automatic diagnosis of several conditions. In this exercise, we will focus on segmenting polyps used in gastrointestinal disease detection. We will use a popular dataset, Kvasir SEG (<https://datasets.simula.no/kvasir-seg/>) [1] to train and test segmentation networks. The data is already contained in the `python_scripts.zip` file, so you do not need to manually download it. The segmentation network $f_\theta(\cdot)$ simply takes an image x as input and predicts the segmentation map y . To train such types of networks, we need to ensure that the predicted output $f_\theta(x)$ is as close to the true output y as possible. Closeness can be defined in several ways, and one such method is Binary Cross Entropy (*BCE*) loss, commonly used in binary classification problems. Binary cross-entropy loss is defined as

$$\text{BCE}(y, p) = -\frac{1}{N^2} \sum_{i,j} \left(y[i, j] \log(p[i, j]) + (1 - y[i, j]) \log(1 - p[i, j]) \right).$$

where $p = \sigma(f_\theta(x))$ and σ denotes the sigmoid function. In the notebook, apply a sigmoid to the network output before computing BCE, Dice loss, and IoU. Additionally, several metrics have been developed that are effective for segmentation problems, with one of the recent popular ones being the Dice loss [2], which is defined as

$$\text{DiceLoss}(y, p) = 1 - \frac{2 \sum_{i,j} y[i, j] p[i, j]}{\sum_{i,j} (y[i, j] + p[i, j]) + \epsilon}, \quad \epsilon > 0$$

To train the network in the assignment the two losses are combined as follows:

$$\min_{\theta} \mathbb{E}_{y, x \sim p_{x,y}} \left(\text{BCE}(y, \sigma(f_\theta(x))) + \lambda_{\text{DiceLoss}} \text{DiceLoss}(y, \sigma(f_\theta(x))) \right)$$

where $\lambda_{\text{DiceLoss}} \geq 0$ is a user-defined parameter. A sample image and its segmentation are given in Figure 1. The quality of segmentation is typically evaluated using the Intersection over Union (IoU) metric. We have provided the function to compute this metric.

In this exercise, you will test two types of networks: U-Net [3], a popular network that has been successfully used for several biomedical applications, and a simple Vision Transformer (ViT)

[4], which employs a transformer-like architecture for computer vision tasks. We have provided you with a simple ViT architecture in the file `./models/vit.py` and the U-Net architecture in the file `./models/unet.py`. We have also provided you with a notebook `segment.ipynb` where the necessary libraries and data are loaded. Your task in this assignment is:

1. Split the data into train, validation, and test sets, using 20% for validation and 20% for testing. The validation set is used to tune the hyperparameters of the network, and the test set is used to report the results.
2. Choose an appropriate batch size and initialize the data loader for each set. Hint: use `TensorDataset`.
3. Define the U-Net model (`UNet(3, 1).to(device)`), the loss, and the optimizer. We have provided you with the following parameters; you can choose to vary them:
 - (a) `EPOCHS`: number of epochs trained ,
 - (b) `LR`: learning rate,
 - (c) `LAMBDA_DICE`: the weight $\lambda_{DiceLoss}$ in the loss function .

Use `torch.nn.BCELoss()` for BCE loss and `smp.losses.DiceLoss('binary')` for dice loss. You can use the Adam optimizer or experiment with other optimizers. Note that the loss function is defined for a single image; however, you can use it for multiple images at once.

4. Complete the training loop of the code. Evaluate the network's performance using the validation set in terms of the IoU metric using the provided function. You can choose how often you want to evaluate your network on the validation set. Plot the loss curves.
5. Evaluate the model on the test set and report the IoU metric. Plot a few examples from your test set and the corresponding predictions. Note that the plots should show results from several cases, including examples where it worked well, cases where it missed the segmentation completely, and cases where the network partially recovered the segmentations.
6. Report on the hyperparameters used and observations on the test set in the cell `Observations and Results`.
7. Report on the possible reasons for using binary cross-entropy loss for training the model.
8. Repeat the above steps using the vision transformer. We have provided you with a simple vision transformer model in the class `SegmentViT`. Complete the steps again and note down your observations and results in the cell `ViT Observations and Results` and compare the results with those of U-Net.

References

- [1] Debesh Jha, Pia H. Smetsrud, Michael A. Riegler, Pål Halvorsen, Thomas de Lange, Dag Johansen, and Håvard D. Johansen, *Kvasir-seg: A segmented polyp dataset*, in *MultiMedia Modeling: 26th International Conference, MMM 2020, Daejeon, South Korea, January 5–8, 2020, Proceedings, Part II 26*, pages=451–462, year=2020, organization=Springer.
- [2] C. H. Sudre, W. Li, T. Vercauteren, S. Ourselin, and M. Jorge Cardoso, *Generalised dice overlap as a deep learning loss function for highly unbalanced segmentations*, in *Deep Learning in Medical Image Analysis and Multimodal Learning for Clinical Decision Support: Third International Workshop, DLMIA 2017, and 7th International Workshop, ML-CDS 2017, Held in Conjunction with MICCAI 2017, Québec City, QC, Canada, September 14, Proceedings 3*, pp. 240–248, 2017, Springer.

- [3] Olaf Ronneberger, Philipp Fischer, and Thomas Brox, *U-net: Convolutional networks for biomedical image segmentation*, in *Medical Image Computing and Computer-Assisted Intervention–MICCAI 2015: 18th International Conference, Munich, Germany, October 5–9, 2015, Proceedings, Part III* 18, pages=234–241, year=2015, organization=Springer
- [4] Alexey Dosovitskiy, Lucas Beyer, Alexander Kolesnikov, Dirk Weissenborn, Xiaohua Zhai, Thomas Unterthiner, Mostafa Dehghani, Matthias Minderer, Georg Heigold, Sylvain Gelly, and others, *An image is worth 16x16 words: Transformers for image recognition at scale*, *arXiv preprint arXiv:2010.11929*, year=2020.