

2.2 Composite quantile regression

$$\begin{aligned} \arg \min_{b_{\theta_k}, \beta} \sum_{i=1}^n \sum_{k=1}^K \omega_k \rho_{\theta_k} \{y_i - b_{\theta_k} - \mathbf{x}_i^T \beta\}, \\ \Updownarrow \\ L(y|\alpha, \beta) = \prod_{i=1}^n \sum_{k=1}^K \omega_k \theta_k (1 - \theta_k) \exp\{-\rho_{\theta_k}(y_i - b_{\theta_k} - \mathbf{x}_i^T \beta)\} \end{aligned} \quad (4)$$

To solve (4), Huang and Chen [2015] introduce a cluster matrix \mathbf{C} , whose i, k th element C_{ik} is equal to 1 if the i th subject belongs to the k th cluster, otherwise $C_{ik} = 0$. It is treated as a missing value, and they consider following complete likelihood

$$\begin{aligned} \prod_{i=1}^n \prod_{k=1}^K \left[\omega_k \theta_k (1 - \theta_k) \exp\{-\rho_{\theta_k}(y_i - b_{\theta_k} - \mathbf{x}_i^T \beta)\} \right]^{C_{ik}} \\ = \prod_{i=1}^n \prod_{k=1}^K \left[\theta_k (1 - \theta_k) \exp\{-\rho_{\theta_k}(y_i - b_{\theta_k} - \mathbf{x}_i^T \beta)\} \right]^{C_{ik}} \times \prod_{i=1}^n \prod_{k=1}^K \omega_k^{C_{ik}} \\ = P(y|C) \times P(C), \end{aligned} \quad (5)$$

where $P(C)$ is marginal distribution which is multinomial distribution.

Then, the posterior distribution of β is

$f(\beta|y, \mathbf{X}, \tau, \lambda, \mathbf{b}, \omega, \mathbf{C}) \propto$ Complete likelihood \times Priors

$$\begin{aligned} & \propto \prod_{i=1}^n \prod_{k=1}^K \left[\omega_k \theta_k (1 - \theta_k) \exp\{-\rho_{\theta_k}(y_i - b_{\theta_k} - \mathbf{x}_i^T \beta)\} \right]^{C_{ik}} \times \pi(\beta|\tau, \lambda) \pi(\omega) \\ & \propto \exp\left\{-\sum_{i=1}^n \sum_{k=1}^K \omega_k \rho_{\theta_k} \{y_i - b_{\theta_k} - \mathbf{x}_i^T \beta\} - \tau \lambda \sum_{j=1}^p |\beta_j|\right\} : \text{(Not sure, but no need to derive)} \end{aligned} \quad (6)$$

: We are draw samples from the posterior distribution using MCMC.

First, consider hierarchical model as follows:

$$\begin{aligned}
y_i &= b_{\theta_k} + \mathbf{x}_i^T \boldsymbol{\beta} + \xi_{1k} \tilde{\nu}_i + \tau^{-1/2} \xi_{2k} \sqrt{\tilde{\nu}_i} z_i, \text{ if } y_i \in \text{cluster } k. \\
\tilde{\nu}|\tau &\sim \prod_{i=1}^n \tau \exp(-\tau \tilde{\nu}_i) \\
\mathbf{z} &\sim \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} z_i^2\right) \\
\boldsymbol{\beta}, \mathbf{s}|\eta^2 &\sim \prod_{j=1}^p \frac{1}{\sqrt{2\pi s_j}} \exp\left(-\frac{\beta_j^2}{2s_j}\right) \prod_{j=1}^p \frac{\eta^2}{2} \exp\left(-\frac{\eta^2}{2} s_j\right) \\
(\tau, \eta^2) &\sim \text{Gamma}(b, d) \\
\boldsymbol{\omega} &\sim \text{Dirichlet}(\alpha_1, \dots, \alpha_K)
\end{aligned}$$

Then, the complete likelihood based on ALD form is

$$f(\mathbf{y}|\mathbf{X}, \tau, \boldsymbol{\beta}, \boldsymbol{\nu}, \mathbf{s}, \eta^2, \mathbf{b}, \boldsymbol{\omega}, \mathbf{C}) = \prod_{i=1}^n \prod_{k=1}^K \left(\frac{1}{\sqrt{2\pi\tau^{-1}\xi_{2k}^2\tilde{\nu}_i}} \right)^{C_{ik}} \exp\left\{ -\frac{1}{2} \sum_{i=1}^n \sum_{k=1}^K \frac{C_{ik}(y_i - b_{\theta_k} - \mathbf{x}_i^T \boldsymbol{\beta} - \xi_{1k} \tilde{\nu}_i)^2}{\tau^{-1}\xi_{2k}^2\tilde{\nu}_i} \right\}, \quad (7)$$

and compute conditional distributions of the parameters as follows.

- $f(\tilde{\nu}_i|\mathbf{X}, \mathbf{y}, \tilde{\nu}_{-i}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}) \propto f(\mathbf{y}|\mathbf{X}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}) \times \pi(\tilde{\nu}_i|\tau) :$
Inverse Gaussian distribution $\left(\lambda = \left(\sum_{k=1}^K \frac{C_{ik}\xi_{1k}^2}{\xi_{2k}^2} + 2 \right) \tau, \mu = \sqrt{\sum_{k=1}^K \frac{C_{ik}(\xi_{1k}^2 + 2\xi_{2k}^2)}{(y_i - b_{\theta_k} - \mathbf{x}_i^T \boldsymbol{\beta})^2}} \right)$
- $f(s_j|\mathbf{X}, \mathbf{y}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}_{-j}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}) \propto \pi(\beta_j|s_j) \times \pi(s_j|\eta^2) :$ Inverse Gaussian distribution $(\lambda = \eta^2, \mu = \sqrt{\eta^2/\beta_j^2})$
- $f(\beta_j|\mathbf{X}, \mathbf{y}, \tilde{\nu}, \boldsymbol{\beta}_{-j}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}) \propto f(\mathbf{y}|\mathbf{X}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}) \times \pi(\beta_j|s_j) :$
 $N\left(\frac{\sum_{i=1}^n x_{ij} \tilde{y}_i / \tilde{\sigma}_i^2}{\sum_{i=1}^n x_{ij}^2 / \tilde{\sigma}_i^2 + 1/s_j}, \frac{1}{\sum_{i=1}^n x_{ij}^2 / \tilde{\sigma}_i^2 + 1/s_j} \right)$, where
- $f(\tau|\mathbf{X}, \mathbf{y}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \eta^2, \boldsymbol{\omega}, \mathbf{C}) \propto f(\mathbf{y}|\mathbf{X}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}) \times \pi(\tilde{\nu}|\tau) \times \pi(\tau) :$ Gamma distribution
- $f(\eta^2|\mathbf{X}, \mathbf{y}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \boldsymbol{\omega}, \mathbf{C}) \propto \pi(\mathbf{s}|\eta^2) \times \pi(\eta^2) :$ Gamma distribution
- $f(\boldsymbol{\omega}|\mathbf{X}, \mathbf{y}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \mathbf{C}) :$ Dirichlet distribution
- $f(\mathbf{C}_i|\mathbf{X}, \mathbf{y}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}_{-i}) :$ Multinomial distribution
- $f(b_{\theta_k}|\mathbf{X}, \mathbf{y}, \tilde{\nu}, \boldsymbol{\beta}, \mathbf{b}, \mathbf{s}, \tau, \eta^2, \boldsymbol{\omega}, \mathbf{C}) :$ Gaussian distribution

They are all known distributions [\[Huang and Chen, 2015\]](#).