

Hen percbchatu сlyutatuu BEWЧЧИТУ

[85.] $f_x(x) = \begin{cases} c(x^2 + 2x), & x \in [0,1] \\ 0, & x \notin [0,1] \end{cases}$

a) тод ашаттанаңа c

$$\int_{-\infty}^{\infty} f_x(x) dx = 1 = \int_0^1 c(x^2 + 2x) dx = c \left[\frac{x^3}{3} + 2 \frac{x^2}{2} \right]_0^1 = c \left[\frac{1}{3} + 2 \cdot \frac{1}{2} \right] = c \left(\frac{1}{3} + 1 \right) = c \left(\frac{4}{3} \right)$$

~~$\Rightarrow c = \frac{3}{4}$~~

2) $E(X) \text{ и } D(X)$

$$E(X) = \int_{-\infty}^{\infty} x \cdot f_x(x) dx = \int_0^1 x \cdot c \cdot (x^2 + 2x) dx = c \left[\frac{x^4}{4} + 2 \frac{x^3}{3} \right]_0^1 = c \left(\frac{1}{4} + \frac{2}{3} \right) = c \cdot \frac{11}{12} = \frac{3}{4} \cdot \frac{11}{12} = \frac{11}{16}$$

~~$D(X) = E(X^2) - (E(X))^2$~~

$$E(X^2) = \int_0^1 x^2 \cdot f_x(x) dx = \int_0^1 x^2 \cdot c \cdot (x^2 + 2x) dx = c \left[\frac{x^5}{5} + 2 \frac{x^4}{4} \right]_0^1 = c \left(\frac{1}{5} + \frac{2}{4} \right) = c \left(\frac{1}{5} + \frac{1}{2} \right) = c \left(\frac{2+5}{10} \right) = \frac{3}{4} \cdot \frac{7}{10} = \frac{21}{40}$$

~~$D(X) = \frac{21}{40} - \left(\frac{11}{16} \right)^2$~~

3) $P(X < \frac{1}{4})$ w.e. $P(X < \frac{11}{16})$

$$P(X < \frac{11}{16}) = \int_0^{\frac{11}{16}} f_X(x) dx =$$

$$P(X < \frac{11}{16}) = F_X(\frac{11}{16})$$

$$F_X(y) = ?$$

$$F_X(y) = \begin{cases} 0, & y < 0 \\ c\left(\frac{y^3}{3} + y^2\right), & y \in [0, 1] \\ 1, & y > 1 \end{cases}$$

$$P(X < y) = \int_0^y c(x^2 + 2x) dx =$$
$$= c \left[\frac{x^3}{3} + 2 \frac{x^2}{2} \right]_0^y$$

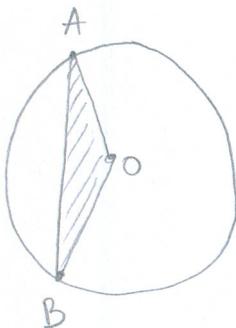
$$\rightarrow c \left[\frac{x^3}{3} + 2 \frac{x^2}{2} \right]_0^{\frac{11}{16}} \approx 0,455$$

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4) $E[X^2 + 3X] = ?$

$$E[X^2 + 3X] = EX^2 + 3EX = \frac{21}{16} + 3 \cdot \frac{11}{16} = \frac{21}{16} + \frac{33}{16}$$

86.]



$E[S_{AOB}] = ?$ → arc length to my area to ΔAOB

$$f \sim U(0, \pi)$$

$$S_{AOB} = \frac{\sin \theta}{2}$$

$$E[S_{AOB}] = E \frac{\sin \theta}{2} = \int_0^{\pi} \frac{\sin x}{2} \cdot \frac{1}{\pi} dx =$$
$$= \frac{1}{2\pi} \left[-\cos x \right]_0^{\pi} = \frac{1}{\pi}$$

- 87) $X \sim U(0,7)$** - X - десенкостно разпределение 6. години на гаже апарацii
 В случаи то дефектът ще биде застенч на искано година или преди това
 Y - времето до смята на апарацii

$$P(Y < 4), EY \text{ и } DY = ?$$

↳ Още въвеждаме също дефектът апарацii Година след смята

$$Y|w) = F_Y(x|w), x|w) < 5 \\ 5, x|w) \geq 5, \text{ where } Y = X \cdot 1_{\{X < 5\}} + 5 \cdot 1_{\{X \geq 5\}}$$

$$\Rightarrow P(Y < 4) = P(X < 4) = F_X(4) = \cancel{\frac{4}{7}} * F_Y(x) = P(W(0,7) < x) = \\ = \int_0^x \frac{1}{7} dy = \frac{x}{7}$$

$$\text{↳ } EY = E(X \cdot 1_{\{X < 5\}} + 5 \cdot 1_{\{X \geq 5\}}) =$$

$$EY = E(X \cdot 1_{\{X < 5\}}) + E(5 \cdot 1_{\{X \geq 5\}}) = \\ = \int_0^5 x \cdot f_X(x) dx + 5 \cdot \int_5^7 f_X(x) dx = \\ = \left[\frac{x^2}{2} \cdot \frac{1}{7} \right]_0^5 + 5 \cdot \left[\frac{2}{7} \right]_5^7 = \frac{25}{14} + 5 \cdot \frac{2}{7} = \frac{25}{14} + \frac{10}{7} = \underline{\underline{\frac{45}{14}}}$$

$E1_A = P(A)$

$$DY = EY^2 - (EY)^2 \\ EY^2 = \int_0^5 x^2 \cdot \frac{1}{7} dx + 5^2 \int_5^7 \frac{2}{7} dx = \left[\frac{x^3}{3} \cdot \frac{1}{7} \right]_0^5 + 25 \left[\frac{2}{7} \right]_5^7 = \\ = \frac{125}{21} + \frac{50}{7} = \frac{125}{21} + \frac{150}{21} = \underline{\underline{\frac{275}{21}}}$$

$$DY = \frac{275}{21} - \left(\frac{45}{14} \right)^2$$

↳ $n=1000$ апарацii

$$X = "gj" ce променлива преди първата година" = P(Y < 5) \\ X \sim Ber(1, \frac{5}{7})$$

→ средната възможност за променливата $\text{Bin}(1000, \frac{5}{7})$,

$$\frac{\frac{2}{\sqrt{\pi}} \cdot \left(e^{-\frac{x^2}{2}} - e^{-\frac{y^2}{2}} \right)}{\frac{2}{\sqrt{\pi}} \cdot \left(1 - e^{-\frac{x^2}{2}} \right)} = \frac{\frac{2}{\sqrt{\pi}} \cdot \left(e^{-\frac{x^2}{2}} - e^{-\frac{y^2}{2}} \right)}{\frac{2}{\sqrt{\pi}} \cdot \left(e^{-\frac{x^2}{2}} \right)} = P(I|A)$$

$$\frac{2}{\sqrt{\pi}} \cdot \left(e^{-\frac{x^2}{2}} - e^{-\frac{y^2}{2}} \right) = \frac{2}{\sqrt{\pi}} \cdot \left(1 - e^{-\frac{x^2+y^2}{2}} \right) = \left(1 - e^{-\frac{x^2+y^2}{2}} \right) = P(I|A)P(I)$$

$$\begin{aligned} & \frac{2}{\sqrt{\pi}} \cdot \left(e^{-\frac{x^2}{2}} - e^{-\frac{y^2}{2}} \right) = \frac{2}{\sqrt{\pi}} \cdot \left(e^{-\frac{x^2}{2}} + e^{-\frac{y^2}{2}} - 2e^{-\frac{x^2+y^2}{2}} \right) = \\ & = \frac{2}{\sqrt{\pi}} \cdot \left(e^{-\frac{x^2}{2}} + e^{-\frac{y^2}{2}} - e^{-\frac{x^2+y^2}{2}} \right) = \\ & = P(A|I) + P(A|II) = P(I|A) \end{aligned}$$

$$P(I|A) = \frac{P(A|I)P(I)}{P(A|I)P(I) + P(A|II)P(I)}$$

$$P(I|A) = \frac{P(I|A)}{P(I|A) + P(I|II)}$$

I - náplňa na sklo
II - náplňa na vodu

$$P(I|A) = ?$$

neplatíme u skupky no ceny zdrojů tečky + voda je výhodnější

$$x_2 = s_{min} = p \quad x_2 = \frac{1}{5}$$

$$x_1 = 8m_i n \Rightarrow x_1 = \frac{1}{x} = 8 \Rightarrow x = 8 \cdot \frac{1}{x} = \frac{1}{8}$$

$$x_1 \exp(x_1) \quad x_2 \exp(x_2)$$

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X = "Время на пребывания на участке"

$$\mathbb{E}X = 30 \text{ min} \rightarrow \lambda = \frac{1}{30} \text{ min}^{-1} \quad \cancel{\mathbb{E}X = 30 \text{ min}} = \frac{1}{2} \quad \lambda = 2$$

небольшое время на пребывание
Большое время в 11:30

то среднее время пребывания на участке =?

$Y =$ "Время, которое требуется на участок пребывания в клинической практике"



$$\mathbb{E}Y = ?$$

$$Y = x_2 \cdot 1_{\{x_1 \leq \frac{1}{2}\}} + \left(x_2 + \left(x_1 - \frac{1}{2} \right) \right) \cdot 1_{\{x_1 > \frac{1}{2}\}}$$

что небольшое время на участке пребывания в клинической практике

в случае если зона пребывания в клинической практике на участке $x_2 + \text{разница между временем пребывания и 11:30}$

$$\mathbb{E}Y = x_2 + \left(x_1 - \frac{1}{2} \right) \cdot 1_{\{x_1 > \frac{1}{2}\}}$$

$$\mathbb{E}Y = \mathbb{E}x_2 + \mathbb{E}\left(x_1 - \frac{1}{2}\right) \cdot 1_{\{x_1 > \frac{1}{2}\}} =$$

$$= \frac{1}{2} + \int_{-\infty}^{\infty} \left(x - \frac{1}{2} \right) \cdot 2e^{-2x} dx =$$

$$= \frac{1}{2} + \int_{\frac{1}{2}}^{\infty} \left(x - \frac{1}{2} \right) de^{-2x} = \frac{1}{2} - \left[\left(x - \frac{1}{2} \right) e^{-2x} \right]_{\frac{1}{2}}^{\infty} + \int_{\frac{1}{2}}^{\infty} e^{-2x} dx =$$

$$= \frac{1}{2} - \frac{1}{2} e^{-2x} \Big|_{\frac{1}{2}}^{\infty} = \frac{1}{2} \left(1 + e^{-1} \right)$$

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* Четиригурарат е да се направи с вс средно 0

$$P(x_i) = \frac{p}{n}, \text{ където } n = 2000 \text{ лв/ч}$$

P) # както се създава га е вероятност $1475 \leq 1535$)=?

→ нормално разпределение

$$x_i = \begin{cases} 1, & \text{еси на i-то} \\ 0, & \text{иначе} \end{cases} \quad x_i \sim \text{Ber}(p)$$

$$P(1475 \leq x_1 + x_2 + \dots + x_{2000} \leq 1535) = ?$$

• прилагаме ЦГТ:

$$D \text{ на вероятност} = p(1-p)$$

$$\frac{3}{n} \cdot \left(1 - \frac{3}{n}\right) = \frac{3}{n} \cdot \frac{1}{n} = \frac{3}{16}$$

$$\delta^2 = \frac{3}{16} \Rightarrow \delta = \sqrt{\frac{3}{16}}$$

$$P(1475 \leq x_1 + x_2 + \dots + x_n \leq 1535) =$$

$$= P\left(\frac{1475 - \mu_n}{\sigma_n} \leq \frac{x_1 + \dots + x_n - \mu_n}{\sigma_n} \leq \frac{1535 - \mu_n}{\sigma_n}\right) =$$

$$= P\left(\frac{1475 - \frac{3}{n} \cdot 2000}{\sqrt{\frac{3}{n} \cdot 2000}} \leq \frac{x_1 + \dots + x_n - \mu_n}{\sigma_n} \leq \frac{1535 - \frac{3}{n} \cdot 2000}{\sqrt{\frac{3}{n} \cdot 2000}}\right) =$$

$$= P\left(\frac{-25}{\sqrt{19,3649}} \leq \frac{x_1 + \dots + x_n - \mu_n}{\sigma_n} \leq \frac{35}{\sqrt{19,3649}}\right) =$$

$n > 30, \delta^2 < \infty$

$$\approx P(-1,29 \leq N(0,1) \leq 1,83) = \int_{-1,29}^{1,83} \frac{1}{2\pi} e^{-\frac{x^2}{2}} dx$$

$$\int e^{-\frac{x^2}{2}} dx \rightarrow \text{нормална вероятност}$$

$$\Rightarrow P(N(0,1) \leq x) = F_{N(0,1)}(x) = \Phi(x)$$

$$P(-1,29 \leq N(0,1) \leq 1,83) = \Phi(1,83) - \Phi(-1,29) =$$

$$= 0,9664 - 0,0985 = \underline{\underline{0,8679}}$$

$f_{x,y}(x,y) = \begin{cases} cx^2 + 1 & , x,y \geq 0, x+2y \leq 1 \\ 0 & \text{иначе} \end{cases}$

a) с, наименьшее значение на x и ограничено на y

$$x, y \geq 0$$

$$x \leq 1-2y \Leftrightarrow y \leq \frac{1-x}{2}$$

$$\begin{aligned} I &= \int_0^{\frac{1-x}{2}} \int_0^{1-x} cx^2 + 1 \, dy \, dx = \int_0^1 cx^2 + 1 [y]_0^{\frac{1-x}{2}} \, dx = \\ &= \int_0^1 (cx^2 + 1) \Big|_{\frac{1-x}{2}} \, dx = \frac{1}{2} \int_0^1 (cx^2 + 1)(1-x) \, dx = \\ &= \frac{1}{2} \int_0^1 cx^2 - cx^3 + 1 - x \, dx = \frac{1}{2} \left[\frac{cx^3}{3} - \frac{cx^4}{4} + x - \frac{x^2}{2} \right]_0^1 = \\ &= \frac{1}{2} \left[\frac{c}{3} - \frac{c}{4} + 1 - \frac{1}{2} \right] = \frac{1}{2} \left[\frac{4c}{12} - \frac{3c}{12} + \frac{2-1}{2} \right] = \\ &= \frac{1}{2} \left[\frac{c}{12} + \frac{1}{2} \right] = \frac{c}{24} + \frac{1}{4} = \frac{c}{24} + \frac{6}{24} = \frac{6c}{24} \end{aligned}$$

$$\Rightarrow 6c = 24$$

$$c = 24/6$$

$$\cancel{c = 18}$$

DANH THOCT TA HA X

$$f_X(x) = \int_0^{\frac{1-x}{2}} cx^2 + 1 dy = (cx^2 + 1) \left(\frac{1-x}{2} \right)$$

→ OzarGatte TO HA Y

↓

$$\begin{aligned} EY &= \int_0^1 \int_0^{\frac{1-x}{2}} y(cx^2 + 1) dy dx = \int_0^1 (cx^2 + 1) \left(\frac{1-x}{2} \right)^2 \cdot \frac{1}{2} dx = \\ &= \int_0^1 (cx^2 + 1) \left| \frac{1-2x+x^2}{8} \right| \cdot \frac{1}{2} dx = \frac{1}{8} \int_0^1 (cx^2 + 1) (1-2x+x^2) dx = \\ &= \frac{1}{8} \int_0^1 cx^2 - 2cx^3 + cx^4 + 1 - 2x + x^2 dx = \\ &= \frac{1}{8} \left[\frac{cx^3}{3} - \frac{2cx^4}{4} + \frac{cx^5}{5} + x - \frac{2x^2}{2} + \frac{x^3}{3} \right]_0^1 = \\ &= \frac{1}{8} \left[\frac{18}{3} - \frac{36}{4} + \frac{18}{5} + x - \cancel{\frac{x}{2}} + \frac{1}{3} \right] = \\ &= \frac{1}{8} \left[\frac{19}{3} - \frac{36}{4} + \frac{18}{5} \right] = \frac{1}{8} \left[\frac{380 - 540 + 216}{60} \right] = \\ &= \frac{1}{8} \left[\frac{56}{60} \right] = \cancel{\frac{7}{60}} \end{aligned}$$

$Cx^2 + 1$

8) $\mathbb{E}[Y | X = 1/2]$

$$f_{X,Y}(y | 1/2) = \frac{f_{X,Y}(1/2, y)}{f_X(1/2)} = \frac{c(1/2)^2 + 1}{(c(1/2)^2 + 1) \cdot 1} =$$

$$= \frac{1}{\frac{1}{2}} = \frac{1}{\frac{1}{h}} = h$$

$$\Rightarrow \text{sa } y \in (0, \frac{1}{h}) \Rightarrow \mathbb{E}[Y | X = \frac{1}{2}] = \mathbb{E}[U(0, \frac{1}{h})] =$$

$$\Rightarrow \mathbb{E}[U(a, b)] = \frac{b-a}{2} \Rightarrow \frac{\frac{1}{h} - 0}{2} = \frac{\frac{1}{h}}{2} = \frac{1}{8}$$

b) Probabilitätsrechnung für $Z = X + 2Y$

$$\begin{vmatrix} Z = X + 2Y \\ W = X \end{vmatrix} \Rightarrow \begin{vmatrix} X = w \\ Y = \frac{Z-w}{2} \end{vmatrix}$$

$$|J| = \begin{vmatrix} 1 & 0 \\ -\frac{1}{2} & \frac{1}{2} \end{vmatrix} = \frac{1}{2}$$

$$\Rightarrow f_{Z,W}(z,w) = f_{X,Y}(w, \frac{z-w}{2}) \cdot \frac{1}{2} = (cw^2 + 1) \cdot \frac{1}{2}$$

$$\text{sa } w, \frac{z-w}{2} \in (0,1) \text{ w.e. } z \in (0,1) \cup w \in (0, z)$$

$$\Rightarrow f_Z(t) = \int_0^t [(cw^2 + 1) \cdot \frac{1}{2}] dw = \frac{1}{2} \int_0^t cw^2 + 1 dw =$$

$$= \frac{1}{2} \left[\frac{cw^3}{3} + w \right]_0^t = \frac{1}{2} \left[\frac{ct^3}{3} + t \right] = \frac{3t^3 + \frac{t}{2}}{6}$$

sa $t \in (0,1)$

$$2) f_{x,y}(x,y) = \begin{cases} cx^3y & , x,y \geq 0, x+y \leq 1 \\ 0 & , \text{ otherwise} \end{cases}$$

$$\begin{array}{l} x \leq 1-y \\ y \leq 1-x \end{array}$$

a) C, nábožná funkcia na X a oznámenie na Y , $CDF(X,Y)$

$$A = x$$

$$I = \int_0^1 \int_0^{1-x} cx^3y \, dy \, dx = \int_0^1 cx^3 \left[\frac{(1-x)^2}{2} \right] \, dx =$$

$$= \frac{1}{2} \int_0^1 cx^3 (1-x)^2 \, dx = \frac{1}{2} \int_0^1 (x^3 (1-2x+x^2)) \, dx =$$

$$= \frac{1}{2} c \int_0^1 x^3 - 2x^4 + x^5 \, dx = \frac{1}{2} c \left[\frac{x^4}{4} - \frac{2x^5}{5} + \frac{x^6}{6} \right] =$$

$$= \frac{1}{2} c \left[\frac{1}{4} - \frac{2}{5} + \frac{1}{6} \right] = \frac{1}{2} c \left[\frac{30 - 48 + 20}{120} \right] =$$

$$= \frac{1}{2} c \cdot \frac{2}{120} = \frac{c}{120} \rightarrow \underline{\underline{c=120}}$$

- nábožná funkcia na X

$$f_X(x) = \int_0^{1-x} cx^3y \, dy = cx^3 \left[\frac{(1-x)^2}{2} \right], \quad x \in [0,1]$$

- oznámenie na Y

$$f_Y(y) = \int_0^1 cx^3y \, dx = c \left[\frac{(1-y)^4}{4} y \right] =$$

$$f_Y(y) = \int_0^1 y \cdot cx^3y \, dy \, dx = c \int_0^1 y \left[\frac{x^3}{3} \right] \Big|_{1-y}^1 = \frac{1}{3} c \left(y^3 |_{1-y}^1 - y^3 x^3 \Big|_{1-y}^1 \right) =$$

$$= y^3 - y^6 - 3y^4 + 3y^5$$

$$\mathbb{E}XY = \int_0^1 \int_0^x xy^3 x dy dx = \int_0^1 3x^2 \int_0^x y dy dx =$$

$$= \int_0^1 3x^2 \cdot \frac{x^2}{2} dx = \int_0^1 \frac{3}{2} x^4 dx = \frac{3}{2} \left[\frac{x^5}{5} \right]_0^1 = \frac{3}{2} \cdot \frac{1}{5} = \underline{\underline{\frac{3}{10}}}$$

$$\Rightarrow \text{Cov}(X,Y) = \mathbb{E}XY - \mathbb{E}X\mathbb{E}Y =$$

$$= \frac{3}{10} - \frac{3}{4} \cdot \frac{3}{8} = \frac{3}{10} - \frac{9}{32} = \frac{96 - 90}{320} = \frac{6}{320} = \underline{\underline{\frac{3}{160}}}$$

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1) $f_{X,Y}(x,y) = \begin{cases} cx+1 & , x,y \geq 0, x+y \leq 1 \\ 0 & , \text{otherwise} \end{cases} , \begin{matrix} x \leq 1-y \\ y \leq 1-x \end{matrix}$

a) c u $\text{Cov}(X,Y) = ?$

$$I = \iint_D (cx+1) dy dx = \int_0^1 cx+1 \int_0^{1-x} dy dx = \int_0^1 (cx+1)(1-x) dx =$$

$$= \int_0^1 cx - cx^2 + 1 - x dx = \left[\frac{cx^2}{2} - \frac{cx^3}{3} + x - \frac{x^2}{2} \right]_0^1 =$$

$$= \left[\frac{c}{2} - \frac{c}{3} + 1 - \frac{1}{2} \right] = \left[\frac{3c-2c}{6} + \frac{1}{2} \right] = \frac{c}{6} + \frac{1}{2} = \frac{c+3}{6}$$

$$\Rightarrow \frac{c+3}{6} = c+3 = 6 \rightarrow c = 6-3 \rightarrow \underline{\underline{c=3}}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$E(XY) = \int_0^{1-x} cx+1 dy = (cx+1)(1-x)$$

$$\underline{\underline{E(X)}} = \int_0^1 x(cx+1)(1-x) dx = \int_0^1 (cx^2+x)(1-x) dx =$$

$$= \int_0^1 cx^2 - cx^3 + x - x^2 dx = \left[\frac{cx^3}{3} - \frac{cx^4}{4} + \frac{x^2}{2} - \frac{x^3}{3} \right]_0^1 =$$

$$= \left[\frac{c}{3} - \frac{c}{4} + \frac{1}{2} - \frac{1}{3} \right] = \frac{3}{3} - \frac{3}{4} + \frac{1}{2} - \frac{1}{3} = \frac{12-9+6-4}{12} = \frac{5}{12}$$

$$E(Y|y) = \int_0^{1-y} cx+1 dx = c \left[\frac{x^2}{2} + x \right]_0^{1-y} =$$

$$= c \left[\frac{(1-y)^2}{2} + (1-y) \right] = \frac{c}{2} (1-y)^2 + (1-y)$$

$$\underline{\underline{E(Y)}} = \frac{c}{2} \int_0^1 y(1-y)^2 dy + \int_0^1 y(1-y) dy - \frac{c}{2} \int_0^1 y^2 - 2y^2 + y^3 dy + \int_0^1 y - y^2 dy =$$

$$= \frac{c}{2} \left[\frac{y^2}{2} - \frac{2y^3}{3} + \frac{y^4}{4} \right]_0^1 + \left[\frac{y^2}{2} - \frac{y^3}{3} \right]_0^1 =$$

$$= \frac{c}{2} \left[\frac{1}{2} - \frac{2}{3} + \frac{1}{4} \right] + \left[\frac{1}{2} - \frac{1}{3} \right] = \frac{c}{2} \left[\frac{6-8+3}{12} \right] + \frac{1}{6} =$$

$$= \frac{c}{2} \cdot \frac{1}{12} + \frac{1}{6} = \frac{3}{2} \cdot \frac{1}{12} + \frac{1}{6} = \frac{1}{8} + \frac{1}{6} = \frac{6+8}{48} = \frac{14}{48} = \frac{7}{24}$$

$$\text{Ex 4} = \int_0^1 \int_0^{1-x} xy(cx+1) dy dx = \int_0^1 x(cx+1) \int_0^{1-x} y dy dx =$$

$$= \int_0^1 x(cx+1) \cdot \frac{(1-x)^2}{2} dx = \frac{1}{2} \int_0^1 (cx^2+x)(1-x)^2 dx =$$

$$= \frac{1}{2} c \int_0^1 x^2 (1-x) dx + \frac{1}{2} \int_0^1 x(1-x)^2 dx =$$

$$= \frac{1}{2} c \int_0^1 x^2 - 2x^3 + x^4 dx + \frac{1}{2} \int_0^1 x - 2x^2 + x^3 dx =$$

$$= \frac{1}{2} c \left[\frac{1}{3} - \frac{2}{5} + \frac{1}{5} \right] + \frac{1}{2} \left[\frac{1}{2} - \frac{2}{3} + \frac{1}{4} \right] =$$

$$= \frac{3}{2} \left[\frac{10-15+6}{30} \right] + \frac{1}{2} \left[\frac{6-8+3}{12} \right] = \frac{3}{2} \cdot \frac{1}{30} + \frac{1}{2} \cdot \left(-\frac{1}{12} \right) =$$

$$= \frac{1}{20} + \frac{1}{24} = \frac{6+5}{120} = \cancel{\frac{11}{120}}$$

$$\Rightarrow \text{Cov}(X, Y) = \frac{11}{120} - \frac{5}{12} \cdot \frac{7}{24} = \frac{11}{120} - \frac{35}{288} = \frac{132-175}{1440} =$$

$$= -\cancel{\frac{43}{1440}}$$

$$8) \quad \mathbb{E}(X \mid Y=1/2)$$

$$f_{x,y}(x|y=\frac{1}{2}) = \frac{f_{x,y}(x, \frac{1}{2})}{f_Y(\frac{1}{2})} = \frac{cx+1}{\frac{5}{2}(1-y)^2 + 1-y}$$

$$\Rightarrow \frac{3}{2}(1-\frac{1}{2})^2 + 1 - \frac{1}{2} = \frac{3}{2} \cdot \frac{1}{4} + \frac{1}{2} = \frac{3}{8} + \frac{4}{8} = \underline{\underline{\frac{7}{8}}}$$

$$\Rightarrow f_{x,y}(x|\frac{1}{2}) = \frac{cx+1}{\frac{7}{8}}, \quad x \in [0, \frac{1}{2}]$$

$$\mathbb{E}[X \mid Y=\frac{1}{2}] = \int_0^{1/2} x \cdot \frac{cx+1}{\frac{7}{8}} dx =$$

$$= \frac{8}{7} \int_0^{1/2} cx^2 + x dx = \frac{8}{7} \left[c \cdot \frac{x^3}{3} + \frac{x^2}{2} \right] = \frac{8}{7} \left[c \cdot \frac{1}{24} + \frac{1}{8} \right] = \underline{\underline{c}}$$

$$= \frac{8}{7} \left[\frac{3}{24} + \frac{3}{24} \right] = \frac{8}{7} \cdot \frac{6}{24} = \frac{6}{21} = \underline{\underline{\frac{2}{7}}}$$

2. $N(\mu, 10^2)$

~~45%~~

15% ойн хөтөрөгчийн сүрьеэртэй нийтийн ойн 250gr гэж таан

a) нариамжийн $\mu = ?$

$$P(N(\mu, 10^2) < 250) = 15\%$$

$$P\left(N(0,1) < \frac{250-\mu}{10^2}\right) = 15\%$$

Ойн хадны ялангаа:

15%, 0,15

$$\frac{250-\mu}{10^2} \approx -1,04$$

$$\rightarrow \frac{250-\mu}{10^2} = -1,04$$

$$250-\mu = -1,04 \cdot 10^2$$

$$250-\mu = -10,4$$

$$-\mu = -10,4 - 250 \quad (-1)$$

$$\mu = 10,4 + 250$$

$$\mu = 260,4$$

b) про-дуктадын хөтөрөгчийн хувь ойн 280gr

$$P(N(\mu, 10^2) > 280) = P(N(0,1) > \frac{280-\mu}{10}) \approx$$

$$1 - \Phi(-1,96) = \Phi(-1,96) = 0,0250 \approx \underline{\underline{2,5\%}}$$

6) Намерете δ , ако средната е 97% от кореспондентният интервал 250 и 270 гр грах.
 $\hookrightarrow N(250, \delta^2)$

$$P(250 \leq N(250, \delta^2) \leq 270) =$$

$$= 1 - 2 P(N(250, \delta^2) < 250) =$$

$$= 1 - 2 P(N(0,1) < \frac{250-250}{\delta}) = \text{no standard} \quad 95\%$$

$$\hookrightarrow Този δ ще е $\Phi^{-1}\left(\frac{20}{\delta}\right) = 0,015$$$

$$\hookrightarrow -\frac{20}{\delta} \approx -2,17 \Rightarrow \delta = 9,2166$$

- 3] Кошмарът е същността на загадки.
 Сърдитите га се решат. Ето как същността на загадки е 80%, 70%, 60% и 40%.
 Ако решим загадка за 10 минути в 0 чинаре
 очакваме ето формула $2 + \text{минут}/10$
 Ако решим за 100 думи, приемаме:
- Сърдитите средният резултат на загадка 1 юле бъде
 ноц 8 минути.

Ако X_i е резултатът на загадка 1 $\sim Ber\left(\frac{8}{10}\right)$ $\rightarrow 80\%$.

$$EX_i = 10 \cdot \frac{8}{10} = 8 \quad \text{и} \quad DX = 100 \cdot \frac{8}{100} \cdot \frac{2}{10} = 16$$

Ако $S_n = X_1 + \dots + X_{100}$ същността $\frac{S_n - n \cdot \mu}{\sqrt{Dn}} = N(0,1)$

$$P\left(\frac{S_{100}}{\sqrt{100}} < 8\right) = P(S_{100} < 800) = P\left(\frac{S_{100} - 100 \cdot 8}{\sqrt{100}} < \frac{8}{\sqrt{100}}\right) \downarrow$$

$$\hookrightarrow \Phi(0,2) = 0,5793 \approx 58\% \quad \downarrow \frac{8}{\sqrt{100}} = \frac{8}{10} = 0,8$$

8) Ozak Galdaia cpegtta oyetka

Heca $Y_i = \# \text{moran ha ayygetti i } \sim X_i^{(1)} + X_i^{(2)} + X_i^{(3)} + X_i^{(4)}$
 xbagens $X_i^{(1)} \sim 10 \cdot \text{Ber}\left(\frac{8}{10}\right)$, $X_i^{(2)} \sim 10 \cdot \text{Ber}\left(\frac{7}{10}\right)$, $X_i^{(3)} \sim 10 \cdot \text{Ber}\left(\frac{6}{10}\right)$,
 $\text{u } X_i^{(4)} \sim 10 \cdot \text{Ber}\left(\frac{4}{10}\right)$ ra hez. an. Gen.

$$\mathbb{E} Y_i = 10 \left(\frac{8}{10} + \frac{7}{10} + \frac{6}{10} + \frac{4}{10} \right) = \underline{\underline{25}}$$

$$\mathbb{D} Y_i = 100 \left(\frac{8}{10} \cdot \frac{2}{10} + \frac{7}{10} \cdot \frac{3}{10} + \frac{6}{10} \cdot \frac{4}{10} + \frac{4}{10} \cdot \frac{6}{10} \right) = 100 \cdot \frac{16+21+24+24}{100} = \\ = 85$$

↳ Oyenkawia ha i e $Z_i := 2 + \frac{Y_i}{5}$

$$\mathbb{E} Z_i = \mathbb{E} \left[2 + \frac{Y_i}{5} \right] = 6,5$$

$$\mathbb{D} Z_i = \mathbb{D} \left[2 + \frac{Y_i}{5} \right] = \frac{\mathbb{D} Y_i}{100} = \frac{85}{100} = \underline{\underline{\frac{17}{20}}}$$

6) Cepogutnoscia cpegtta oyetka go e noite 11. no-Burora oñ
 ozak Galdaia =?

$$P \left| \frac{Z_1 + \dots + Z_{100}}{100} - 6,5 > \frac{1}{100} \cdot 6,5 \right. =$$

$$P \left| \frac{Z_1 + \dots + Z_{100} - 6,5 \cdot 100}{\sqrt{100 \cdot \frac{17}{20}}} > \frac{\frac{6,5}{100}}{\sqrt{100 \cdot \frac{17}{20}}} \right. \rightarrow \frac{0,065}{\sqrt{1062}} \sim 0,0213$$

J

C

$$\boxed{1} \quad E[X_i] = 5011$$

$$DX_i = 4000$$

$$n=1 \quad P(X_1 + \dots + X_n > 10^6)$$

$$S_n = \sum_{i=1}^n X_i \quad P\left(\frac{S_n - E[S_n]}{\sqrt{D(S_n)}} > \frac{10^6 - 5011 \cdot n}{\sqrt{4000 \cdot n}}\right)$$

$$\stackrel{\text{LFT}}{\approx} \stackrel{\text{N}(0,1)}{\approx} P\left(N(0,1) > \frac{10^6 - 5011n}{\sqrt{4000n}}\right) = 99\%$$

$$\Rightarrow \frac{10^6 - 5011}{\sqrt{4000n}} < -2,33 \quad \rightarrow \underline{n \geq 200}$$

$$\Rightarrow E[S_n] = n \cdot 5011$$

$$DS_n = n \cdot \sigma^2 = n \cdot 4000$$

↳ ukane ga tähemikus mõistet n , se

$$P(S_n \geq 10^6) \geq 99\% (0,99)$$

$$\Rightarrow P\left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \geq \frac{10^6 - n\mu}{\sigma\sqrt{n}}\right) \geq 0,99$$

$$\downarrow \quad P\left(N(0,1) \geq \frac{10^6 - n\mu}{\sigma\sqrt{n}}\right) \geq 0,99$$

$$\frac{10^6 - 5011n}{\sqrt{4000n}} \leq -2,33$$

$$\frac{10^6 - 5011n}{4000n} \leq -2,33$$

$$10^6 - 5011n \leq -2,33 \cdot 4000n$$

$$10^6 - 5011n \leq -2,33 \cdot 63,25 \sqrt{n}$$

$$10^6 - 5011n \leq -147,36 \sqrt{n}$$

$$10^6 - 5011n + 147,36 \sqrt{n} = 0$$

$$10^6 - 5011x^2 + 147,36x = 0$$

• решаем квадратное уравнение

$$-5011x^2 + 147,36x + 10^6 = 0$$

$$a = -5011, b = 147,36, c = 10^6$$

$$x = \frac{-b \pm \sqrt{b^2 - 4ac}}{2a}$$

$$\sqrt{21714,97 + 20044 \cdot 10^6}$$

$$\sqrt{20044021714,97}$$

$$x = \frac{-147,36 \pm \sqrt{20044021714,97}}{-10022} \approx 141576,9109$$

$$x = \frac{-147,36 \pm 141576}{-10022}$$

$$x = \frac{-147,36 - 141576}{-10022} = \frac{-141723,36}{-10022} \approx 14,16$$

• проверим, что $n = x^2$

$$n = (14,16)^2 \approx 199,9396 \approx 200$$

$$n \geq 200$$

$$\Pr(\max(x_1, \dots, x_n) \leq 10^4) = \left(1 - \Pr(x_i > 10^4)\right)^n$$

$$1 - \Pr(x_i > 10^4) \approx \frac{63.25}{10^4 - 500} = 0.9999999999999999$$

$\Pr(x_i > 10^4)$ is small \rightarrow

$$\Pr(\max(x_1, \dots, x_n) \leq 10^4) \approx 1$$

$$\Pr(\max(x_1, \dots, x_n) > 10^4) = 1 - \Pr(\max(x_1, \dots, x_n) \leq 10^4)$$

$$\frac{388}{6889} \approx \frac{63.25}{10^4 - 500} = \frac{63.25}{9999500} = \frac{1}{1588}$$

$$\Pr(\max(x_1, \dots, x_n) > 10^4) \approx 1/1588$$

large max value \rightarrow small probability

$$\Pr(\max(x_1, \dots, x_n) > 10^4) \approx 1/1588$$

2. $X \sim \text{Exp}(2)$ u $Y \sim U(0,3)$

$$f_X(x) = \begin{cases} 2e^{-2x}, & \text{ako } x > 0 \\ 0, & \text{u ina } x \end{cases}$$

$$f_{Y|X}(y|x) = \begin{cases} \frac{1}{3}, & \text{ako } 0 < y < 3 \\ 0, & \text{u ina } y \end{cases}$$

$$\text{Cor}(X, Y) = ?$$

$$\text{P}(X < Y) = ?$$

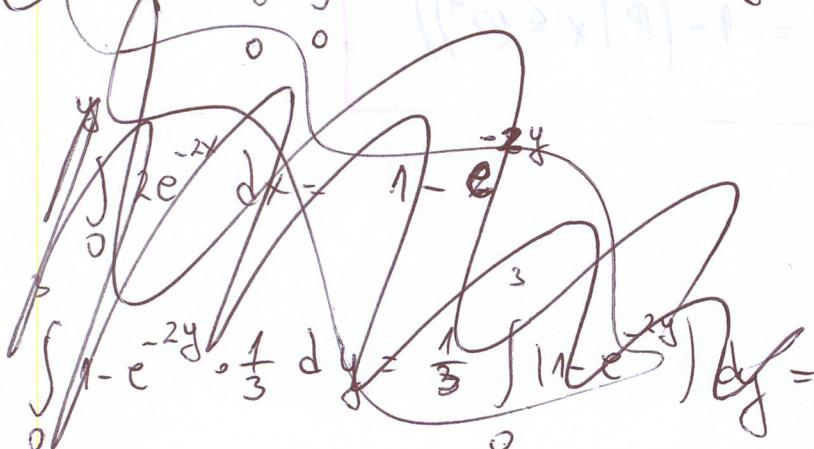
$$F_{XY}(x,y) = ?$$

a) Tačku rano $X \sim U(0,2)$ teža bučina. $\text{Cor}(X, Y) = 0$

$$E[X] = \frac{1}{2}, \quad DX = \frac{1}{2^2} = \frac{1}{4}$$

$$E[Y] = \frac{a+b}{2} = \frac{3}{2}, \quad DY = \frac{(b-a)^2}{12} = \frac{3^2}{12} = \frac{9}{12} = \frac{3}{4}$$

~~$P(X < Y) = \int \int f_{X,Y}(x,y) dx dy$~~



$$P(X < Y) = ?$$

$$P(X < Y) = \iint f_{XY}(x,y) dx dy$$

↳ ибо X и Y — независимы, то $P(X < Y)$ можно выразить в виде $P(X < Y | Y = y) P(Y = y)$.
Наше значение $f_{XY}(x,y) = f_X(x) f_Y(y)$

$$f_{XY}(x,y) = f_X(x) f_Y(y)$$

$$P(X < Y) = \iint f_X(x) f_Y(y) dx dy = \left| \int_0^3 \int_0^y 2e^{-2x} dy dx \right| \cdot \frac{1}{3} dy$$

$$\Rightarrow \int_0^y 2e^{-2x} dx = \left[-e^{-2x} \right]_0^y = 1 - e^{-2y} \rightarrow \text{Corresponding variable}$$

↳ Таблица

$$\int_0^3 (1 - e^{-2y}) \cdot \frac{1}{3} dy = \frac{1}{3} \int_0^3 1 - e^{-2y} dy$$

$$\Rightarrow \int_0^3 1 dy = 1 \cdot (3 - 0) = 3$$

$$\int_0^3 e^{-2y} dy = \left[\frac{1}{2} e^{-2y} \right]_0^3 = -\frac{1}{2} e^{-2y} + \frac{1}{2} = -\frac{1}{2} e^{-6} + \frac{1}{2}$$

$$\Rightarrow P(X < Y) = \frac{1}{3} \left| 3 - \left(-\frac{1}{2} e^{-6} + \frac{1}{2} \right) \right| = \frac{1}{3} \left| 3 + \frac{1}{2} e^{-6} - \frac{1}{2} \right| =$$

$$= \frac{1}{3} \left| \frac{5}{2} + \frac{1}{2} e^{-6} \right| = \frac{1}{3} \left| \frac{5}{2} + \frac{e^{-6}}{2} \right| = \frac{5}{6} + \frac{e^{-6}}{6} = \frac{1}{6} \left(5 + e^{-6} \right)$$

$$z = \frac{x}{y}$$

#TODO

$$\begin{aligned} z &= \frac{x}{y} \\ w &= \frac{y}{x} \end{aligned} \Rightarrow \quad \left| \begin{array}{l} x = w \\ y = \frac{w}{z} \end{array} \right.$$

$$J = \begin{vmatrix} \frac{\partial x}{\partial w} & \frac{\partial x}{\partial z} \\ \frac{\partial y}{\partial w} & \frac{\partial y}{\partial z} \end{vmatrix} = \begin{vmatrix} 1 & 0 \\ \frac{1}{z} & -\frac{w}{z^2} \end{vmatrix} = \begin{vmatrix} \frac{w}{z^2} \end{vmatrix}$$

$$f_{z,w}(z,w) = f_{x,y}(w, \frac{w}{z}) \circ J$$

$$f_{x,y}(x,y) = f_x(x) f_y(y) = 2e^{-2x} \cdot \frac{1}{3} = \frac{2}{3} e^{-2x}$$

$$f_{z,w}(z,w) = \frac{2}{3} e^{-2w} \cdot \frac{w}{z^2} = \frac{2}{3} \cdot \frac{we^{-2w}}{z^2}$$

$$f_z(z) = \int_0^z f_{z,w}(z,w) dw = \int_0^z \frac{2}{3} \cdot \frac{we^{-2w}}{z^2} dw = \frac{2}{3z^2} \int_0^z we^{-2w} dw$$

$$= \left(\frac{1}{2} - \frac{1}{2} e^{-2z} \right) \Big|_0^z = \left(\frac{1}{2} + \frac{1}{2} e^{-2z} \right) - \frac{1}{2} e^{-2z} = \frac{1}{2} + \frac{1}{2} e^{-2z} - \frac{1}{2} e^{-2z} = \frac{1}{2}$$

3. $X \perp Y \sim \text{Exp}(\frac{1}{\alpha}), Y \sim \text{Exp}(\frac{1}{\beta})$

a) $\bar{X} = (\bar{x}_1 + \dots + \bar{x}_n)/n$

#TODO

$$P(|\alpha - \bar{X}| > c) \leq 1\%$$

Ow UMT so ronem n:

$$\frac{\bar{X} - \alpha}{\frac{1}{\alpha} \sqrt{n}} \approx Z \sim N(0,1) \text{ . Basis des abgesu G:}$$

$$P(|Z| > \frac{c}{\sqrt{n}}) \leq 1\%$$

$$\Leftrightarrow P(|Z| > \frac{c}{\sqrt{n}}) \leq 0,5\%$$

$$\Rightarrow \frac{c}{\sqrt{n}} \approx 2,57 \Rightarrow c \approx 2,57 \cdot \sqrt{n}$$

$$\left(\frac{2,57}{\sqrt{n}}\right) \cdot \left[1 + \frac{1}{\beta}\right] = \left[\frac{1}{\beta} + \frac{2,57}{\sqrt{n}}\right] \cdot \frac{1}{\beta}$$

4

$$f_{x,y}(x,y) = cx^2 + 1, \quad x,y \geq 0$$

$$x+2y \leq 1$$

u 0 иначе

или $(0 \leq x-y \leq 1)$ a) С, областью на x и ограничениями y

$$\begin{aligned} I &= \int_0^1 \int_0^{1-x/2} cx^2 + 1 \, dy \, dx = \int_0^1 (cx^2 + 1) \Big|_0^{1-x/2} \, dx = \frac{1}{2} \int_0^1 (cx^2 + 1)(1-x) \, dx = \\ &= \frac{1}{2} \int_0^1 cx^2 - cx^3 + 1 - x \, dx = \frac{1}{2} \left[\frac{cx^3}{3} - \frac{cx^4}{4} + 1 - \frac{x^2}{2} \right]_0^1 = \\ &= \frac{1}{2} \left[\frac{c}{3} - \frac{c}{4} + 1 - \frac{1}{2} \right] = \frac{1}{2} \left[\frac{4c-3c}{12} + \frac{2-1}{2} \right] = \\ &= \frac{1}{2} \left[\frac{c}{12} + \frac{1}{2} \right] = \frac{c}{24} + \frac{1}{4} = \frac{c+6}{24} \end{aligned}$$

$$\Rightarrow c+6 = 24$$

$$c = 24 - 6$$

$$\underline{\underline{c = 18}}$$

С, областью на x

$$f_{x|y}(x) = \int_0^{1-x} cx^2 + 1 \, dy = [cx^2 + 1] \Big|_0^{1-x} =$$

$$\begin{aligned}
 Q4 &= \int_0^1 \int_0^{\frac{1-x}{2}} y(cx^2+1) dy dx = \\
 &= \int_0^1 \int_0^{\frac{1-x}{2}} \frac{y^2}{2} (cx^2+1) dy dx = \int_0^1 (cx^2+1) \cdot \frac{(1-x)^2}{2} dx = \\
 &= \int_0^1 (cx^2+1) \cdot \frac{1-(x-1)^2}{2} dx = \int_0^1 (cx^2+1) \frac{(1-x)^2}{8} dx = \\
 &= \frac{1}{8} \int_0^1 (cx^2+1) (1-2x+x^2) dx = \frac{1}{8} \int_0^1 (cx^2-2cx^3+cx^4+1-2x+x^2) dx = \\
 &= \frac{1}{8} \left[\frac{cx^3}{3} - \frac{2cx^4}{4} + \frac{cx^5}{5} + 1 - \frac{2x^2}{2} + \frac{x^3}{3} \right]_0^1 = \\
 &= \frac{1}{8} \left[\frac{c}{3} - \frac{c}{2} + \frac{c}{5} + 1 - 1 + \frac{1}{3} \right] = \\
 &= \frac{1}{8} \left[\frac{10c-15c+6c}{30} + \frac{1}{3} \right] = \frac{1}{8} \left[\frac{c}{30} + \frac{1}{3} \right] = \cancel{\frac{1}{8} \left[\frac{c}{30} + \frac{1}{3} \right]} \\
 &= \frac{1}{8} \left[\frac{18}{30} + \frac{1}{3} \right] = \frac{18}{240} + \frac{1}{24} = \frac{18+10}{240} = \frac{28}{240} = \frac{7}{60}
 \end{aligned}$$

$$8) E[Y | X=1/2] = ?$$

AAA

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)} = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

$$f_X(x) = [cx^2 + 1] \cdot \left| \frac{1-x}{2} \right| = (18x^2 + 1) \left| \frac{1-x}{2} \right| = g(x^2+1)(1-x)$$

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)} = \frac{18(x^2+1)}{g(x^2+1)(1-x)} = \frac{2}{1-x}$$

to same value $x = \frac{1}{2}$

$$\Rightarrow f_{Y|X}(y|x) = \frac{2}{1-\frac{1}{2}} = \frac{2}{\frac{2-1}{2}} = \frac{2}{\frac{1}{2}} = \underline{\underline{4}}$$

$$\frac{1-x}{2} = \frac{1-\frac{1}{2}}{2} = \frac{\frac{1}{2}}{2} = \frac{1}{4}$$

$$\Rightarrow E[Y | X = \frac{1}{2}] = \int_0^4 y f_{Y|X}(y|\frac{1}{2}) dy = \int_0^4 y \cdot 4 dy =$$

$$= 4 \int_0^4 y dy = 4 \cdot \left[\frac{y^2}{2} \right]_0^4 = 4 \cdot \frac{(4)^2}{2} = 4 \cdot \frac{16}{2} = 4 \cdot \frac{1}{16 \cdot 2} = \underline{\underline{\frac{1}{8}}}$$

HYPOTHETISCH

h.1. OCHO BTRU NOTGUTU A

2024 / KH

4.1. $f(x| -x) = f(x|x)$, $x \in \mathbb{R}$

$F_x(0) = ?$

$$\begin{aligned} F_x(0) &= \int_{-\infty}^0 f(x|x) dx = \int_{-\infty}^{\infty} f(x|x) dx - \int_{\infty}^{\infty} f(x|x) dx = \\ &= \int_{-\infty}^{\infty} f(x|x) dx - \int_{-\infty}^0 f(x|x) dx = 1 - F_x(0) \Rightarrow \end{aligned}$$

$$F_x(0) = 1 - F_x(0)$$

$$2F_x(0) = 1$$

$F_x(0) = 1/2$

$$\int_{-\infty}^{\infty} f(x|x) dx = - \int_{-\infty}^0 f(x|-u) du = \int_{-\infty}^0 f(x|u) du = f_x(0)$$

$\begin{matrix} u := -x \\ du := -dx \end{matrix}$

$f_x(-x) = ?$

$$f_x(-x) = \int_{-\infty}^0 f_x(t) dt \underset{\begin{matrix} u := -t \\ du := -dt \end{matrix}}{=} \int_{\infty}^0 f_x(u) du = \underline{1 - F_x(x)}$$

$P(|X| < x) = ?$

$$\begin{aligned} P(|X| \leq x) &= P(-x < X \leq x) = F_x(x) - F_x(-x) = F_x(x) - (1 - F_x(x)) = \\ &\quad \underline{1 - F_x(x)} \end{aligned}$$

$$= F_x(x) - 1 + F_x(x) = \underline{2F_x(x) - 1}$$

4.3.

$$x \sim \text{Exp}(\lambda)$$

$$x \sim f_x(x) = \lambda e^{-\lambda x} \cdot 1_{(0, \infty)}(x)$$

a) $y = -x$

b) $y = 2x - 1$

c) $y = \sqrt{x}$

a) $g(x) = -x$

$$g^{-1}(y) = h(g) = -y$$

$$f(-x|y) = f_x(-y) \cdot 1_{(-1)} = \lambda e^{\lambda x} \cdot 1_{(0, \infty)}(-y) = \lambda e^{\lambda x} \cdot 1_{(-\infty, 0)}(y)$$

b) $g(x) = 2x - 1$
 \downarrow
 $y \quad g^{-1}(y)$

$$y = 2g^{-1}(y) - 1$$

$$2g^{-1}(y) = y + 1$$

$$g^{-1}(y) = \frac{y+1}{2}$$

$$\frac{d}{dy} g^{-1}(y) = \frac{1}{2}$$

$$f_{2x-1|y} = f_x(g^{-1}(y)) |(g^{-1}(y))'| = f_x\left(\frac{y+1}{2}\right) \cdot \left|\frac{1}{2}\right| =$$

$$= \frac{\lambda}{2} e^{-\lambda \left|\frac{y+1}{2}\right|} \cdot 1_{(0, \infty)}\left(\frac{y+1}{2}\right) = \frac{\lambda}{2} e^{-\frac{\lambda}{2}(y+1)^2} \cdot 1_{(-1, \infty)}(y)$$

$$0 < \frac{y+1}{2} < \infty$$

$$2 \cdot 0 < y+1 < \infty$$

$$0 < y+1 < \infty$$

$$-1 < y < \infty$$

$$6) y = rx$$

$$g(x) = rx, \text{ za } x > 0$$

$$g^{-1}(y) = y^2, \text{ za } y > 0$$

$$\begin{aligned} \varphi_{rx}(y) &= \lambda e^{-\lambda y^2} \cdot 1_{(0, \infty)}(y^2) = \\ &= \lambda e^{-\lambda y^2} \cdot |2y| \cdot 1_{(0, \infty)}(y) = \\ &= \lambda e^{-\lambda y^2} \cdot 2y \cdot 1_{(0, \infty)}(y) \end{aligned}$$

$$a) g(x) = -x, \quad g^{-1}(y) = -y$$

$$\cancel{\frac{d}{dy} g^{-1}(y) = \frac{d}{dy}(-y) = -1} \quad \begin{array}{l} \text{проверка} \\ \text{обратная функция} \end{array}$$

$$\varphi_y(y) = \varphi_x(g^{-1}(y) \cdot |g^{-1}(y)|') = \varphi_x(-y)(-1)$$

$$\varphi_y(y) = \lambda e^{-\lambda(-y)} \cdot 1 = \lambda e^{\lambda y}, \text{ za } y \leq 0$$

8)

$$1) Y = 2x - 1$$

1) Напишите уравнение обратимое к ϕ -изоморфизму

~~Изоморфизм~~

$$Y = 2x + 1 \Rightarrow Y - 1 = 2x$$

$$x = \frac{y+1}{2}$$

$$\Rightarrow g^{-1}(y) = \boxed{\frac{y+1}{2}}$$

2) Напишите производную обратимого ϕ -изоморфизма

$$(g^{-1}(y))' = \frac{1}{2} \cdot (y+1)', \quad (y+1)' \text{ существует } y \neq -1, \text{ так как}$$

~~если~~ ~~тогда~~

$$\underline{(g^{-1}(y))'} = \frac{1}{2}$$

3) Запишите формулу для нахождения производной

$$f'(y) = \lambda \times |g^{-1}(y)| \cdot |(g^{-1}(y))'| =$$

$$= \lambda \times \frac{y+1}{2} \times \left| \frac{1}{2} \right|$$

$g^{-1}(y)$

$$\Rightarrow f'(y) = \lambda e^{-\lambda \left(\frac{y+1}{2} \right)} \cdot \frac{1}{2} = \frac{1}{2} e^{\frac{-\lambda(y+1)}{2}}$$

и при этом $x \geq 0$, что означает $2x - 1 \geq -1$, $y \geq -1$

$$* \quad 0 < \frac{y+1}{2} < \infty$$

$$0 < y+1 < \infty$$

$$-1 < y < \infty$$

$$6) y = \sqrt{x}$$

$$g^{-1}(y) = y^2$$

$$y = \sqrt{x} / ^2$$

$$y^2 = x$$

$$(g^{-1}(y))' = (y^2)' = 2y$$

$$\Rightarrow f'(y) = f_x(g^{-1}(y) \cdot |(g^{-1}(y))'|) =$$

$$= f_x(y^2) \cdot |2y| =$$

$$= \lambda e^{-\lambda y^2} \cdot 2y = 2\lambda y^{-\lambda y^2}, \text{ so } A(10, \infty) (y)$$

$$0 < y^2 < \alpha / ^2$$

$$0^2 < y < \alpha ^2$$

$$0 < y < \alpha$$

L.5.]

$$f(x) = \begin{cases} c(x^2 + 2x), & x \in [0, 1] \\ 0, & x \notin [0, 1] \end{cases}$$

a) c?

b) $E[X]$, DX c) $P[X \leq t]$ d) $E[X^2 + 3X]$

$$a) I_1 = \int_{-\infty}^{\infty} f(x) dx = \int_0^1 c(x^2 + 2x) dx = c \left[\frac{x^3}{3} + \frac{2x^2}{2} \right]_0^1 =$$

$$= c \left(\frac{1}{3} + \frac{2}{2} \right) = c \left(\frac{2+6}{6} \right) = c \cdot \frac{8}{6} \rightarrow c \cdot \frac{4}{3}$$

$$1 = \frac{4}{3} c \rightarrow \frac{1}{c} = \frac{3}{4} \quad c = \frac{3}{4}$$

$$b) E[X] = \int_{-\infty}^{\infty} x \cdot f(x) dx = \int_0^1 x \cdot \frac{3}{4}(x^2 + 2x) dx =$$

$$\begin{aligned} &= \frac{3}{4} \int_0^1 x^3 + 2x^2 dx = \frac{3}{4} \left[\frac{x^4}{4} + \frac{2x^3}{3} \right]_0^1 = \frac{3}{2} \cdot \frac{5+4}{12} = \frac{3}{2} \cdot \frac{9}{12} = \frac{27}{24} = \frac{9}{8} \\ &= \frac{3}{2} \cdot \left[\frac{x^4}{4} \right]_0^1 + \frac{3}{2} \cdot \left[\frac{2x^3}{3} \right]_0^1 = \frac{3}{4} \cdot \frac{1}{4} + \frac{3}{2} \cdot \frac{1}{3} = \frac{3}{16} + \frac{1}{2} = \end{aligned}$$

$$\begin{aligned} &= \frac{3+8}{16} = \frac{11}{16} \end{aligned}$$

$$\begin{aligned} \mathbb{E}[X^2] &= \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_{-\infty}^1 x^2 \cdot \frac{3}{4}(x^2 + 2x) dx = \\ &= \frac{3}{4} \int_0^1 x^4 + 2x^3 dx = \frac{3}{4} \int_0^1 x^4 dx + \frac{3}{4} \cdot 2 \int_0^1 x^3 dx = \\ &= \frac{3}{4} \cdot \frac{1}{5} + \frac{3}{2} \cdot \frac{1}{4} = \frac{3}{20} + \frac{3}{8} = \frac{6+15}{40} = \underline{\underline{\frac{21}{40}}} \end{aligned}$$

$$DX = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = \frac{21}{40} - \underline{\underline{\left(\frac{11}{10}\right)^2}}$$

B) $\mathbb{P}|X \leq \mathbb{E}[X]| = \int_0^{\frac{11}{10}} \frac{3}{4}(x^2 + 2x) dx = \frac{3}{4} \int_0^{\frac{11}{10}} x^2 dx + 2 \cdot \frac{3}{4} \int_0^{\frac{11}{10}} x dx = \dots$

$F_X(x)$

r) $\mathbb{E}(x^2 + 3x) = \mathbb{E}x^2 + 3\mathbb{E}x = \dots \checkmark$

н.к. $x \in \text{H.C.G. c н.в. и.с.и. } f_x$

$y = x^2$, push pegeneet uens $H?$

$\mathbb{P}(X^2 \leq x) = \int_0^0 f_X(t) dt + \int_{-\sqrt{x}}^0 f_X(t) dt =$

$$x \in \text{H.C.B.} \subset \text{nabovouci } f(x)$$

$$y = x^2, \quad \sim x^2$$

$$\Leftrightarrow g^{-1}(y) = \sqrt{x^2}$$

$$y = x^2$$

$$x = \sqrt{y}$$

$$g^{-1}(y) = \pm \sqrt{y}$$

$$\Rightarrow f(g^{-1}(y)) =$$

$$(g^{-1}(y))^2 = \frac{1}{2\sqrt{y}}$$

$$= xb(xb + x)^2 = (xb + x)^2$$

$$= xb^2 + 2xb^2 + x^2b^2 = (3b^2 + x^2)b^2$$

~~$$= xb(4x^2 + 4x^2) = (8x^2)b^2$$~~

h-h

$x \in \text{H.c.6.} \subset \text{ набором } f_x$

$$y = x^2, \sim Y?$$

$$\rightarrow y = x^2 |'$$

$$x = \pm \sqrt{y}$$

→ unique

$$g_1^{-1}(y) = \sqrt{y}$$

$$g_2^{-1}(y) = -\sqrt{y}$$

$$\rightarrow (g_1^{-1}(y))' = \frac{1}{2\sqrt{y}}$$

$$(g_2^{-1}(y))' = -\frac{1}{2\sqrt{y}}$$

$$\begin{aligned}\text{• } f'(y) &= f_x(g_1^{-1}(y)) \left| (g_1^{-1}(y))' \right| + f_x(g_2^{-1}(y)) \left| (g_2^{-1}(y))' \right| = \\ &= f_x(\sqrt{y}) \cdot \left| \frac{1}{2\sqrt{y}} \right| + f_x(-\sqrt{y}) \cdot \left| -\frac{1}{2\sqrt{y}} \right| = \\ &= f_x(\sqrt{y}) \cdot \frac{1}{2\sqrt{y}} + f_x(-\sqrt{y}) \cdot \frac{1}{2\sqrt{y}} \\ &= \frac{1}{2\sqrt{y}} (f_x(\sqrt{y}) + f_x(-\sqrt{y})), \quad y \geq 0\end{aligned}$$

ЧАСТЬ ПЕРВАЯ
РАСПРЕДЕЛЕНИЕ

$$X \sim \text{Unif}(a, b)$$

$$f_X(x) = \frac{1}{b-a} \cdot I(a, b)(x)$$

$$Y = \frac{x-a}{b-a}, \quad Y \sim ?$$

$$\hookrightarrow g(x) = \frac{x-a}{b-a}$$

$$Y = \frac{x-a}{b-a}$$

$$y(b-a) = x-a$$

$$x = a + (b-a)y$$

$$(g^{-1}(y))' = b-a$$

$$f_Y(y) = f_X(g^{-1}(y)) |(g^{-1}(y))'| =$$

$$= \underbrace{f_X(a + (b-a)y)}_{\frac{1}{b-a}} \cdot (b-a) =$$

$$= \frac{1}{b-a} \cdot \frac{b-a}{b-a} = 1 \sim \text{Unif}(0, 1)$$

$$a < a + (b-a)y < b$$

$$0 < y < 1$$

4.15. $X \sim \text{Unif}(0,7)$ \rightarrow времето за бекът касът работи е
гадет апарат в гараж

Лема се на линейна функция или преобразуване

$Y = \text{"Времето до смята за апарат"}$

$$P(Y < 4) = P(\min(X, 5) < 4) = P(X < 4) \quad (I_{X < 4} + I_{X \geq 4})$$

Линейният закон $X \in [0, 7]$

Линейният закон $X \in [0, 7]$ е равномерно разпределен в интервала

$$P(X < 4) = \frac{4}{7}$$

$E[Y]$

Линейният закон $X \in [0, 7]$

Линейният закон $X \in [0, 7]$, т.е. $y = x$

Линейният закон $X \in [0, 7]$, т.е. $y = 5$

Линейният закон $X \in [0, 7]$, т.е. $y = x$

Линейният закон $X \in [0, 7]$, т.е. $y = 5$

$$f_Y(y) = \begin{cases} \frac{1}{7}, & 0 \leq y \leq 5 \\ 0, & \text{иначе} \end{cases}$$

$$\rightarrow E[Y] = \int_0^5 y f_Y(y) dy + \int_5^7 5 f_Y(y) dy$$

Смятане по нули

$$\int_0^5 y f_Y(y) dy = \int_0^5 y \cdot \frac{1}{7} dy = \frac{1}{7} \int_0^5 y dy = \frac{1}{7} \cdot \left[\frac{y^2}{2} \right]_0^5 = \frac{1}{7} \cdot \frac{25}{2} = \frac{25}{14}$$

$$\int_5^7 5 f_Y(y) dy = \int_5^7 5 \cdot \frac{1}{7} dy = \left(5 \cdot \frac{1}{7} \right) \int_5^7 1 dy = 5 \cdot \frac{1}{7} (7 - 5) = \frac{10}{7}$$

$$\Rightarrow E[Y] = \frac{25}{14} + \frac{10}{7} = \frac{25 + 20}{14} = \frac{45}{14}$$

последните
числа

трансформация, защото сме и

~~P~~

$$\text{P}(Y=y) = \text{P}(\text{f}_x^{-1}(y))$$

$\Rightarrow Y \in \text{Uniform}(0,1)$ f.e. u.a. da esponentielle Verteilung, woher

$$\text{P}(Y=y) = \text{P}(X \in \text{f}_x^{-1}(y)) = \text{P}(X \in \text{f}_x^{-1}(y)) = \text{P}(X \leq \text{f}_x^{-1}(y))$$

$$\text{P}(Y=y) = \text{P}(X \leq \text{f}_x^{-1}(y)) = \text{P}(Y \leq \text{f}_x^{-1}(y)) = \text{P}(Y \leq y) = \text{P}(Y \leq y)$$

Bei einem f.x. mit einer Verteilung, die uniforme Verteilung, woher

$$\text{P}(Y=y) = \text{P}(X \leq y) = \text{P}(Y \leq y)$$

$\Leftrightarrow Y = f_x(X)$, w.e. $f_y(y)$ eine dichte-funktion ist

temperaturen u. temperaturen der personen

z.B. dass ein f.x. mit einer exponentialen Verteilung ist

Also $f_x(x)$ ist uniform temperaturen u. temperaturen, wobei

4.12.

$$\overline{s} = \frac{t}{\ln(1.000, \frac{1}{2})} = \frac{t}{\ln(1.000, 5)} = 1000 \cdot \frac{t}{5} =$$

$$\text{Cpgeht} \sim \ln(1.000, \frac{1}{2})$$

Neigung s-wurzel logn+1

u. Also ca. neigung, 1000 anpassen, kann cpgeht zu einem weiteren

$$\overline{s} = \frac{t}{\ln(1.000, \frac{1}{2})} = \frac{t}{\ln(1.000, 5)} = (X\phi) - X\phi = X\phi$$

$$\overline{s} = \frac{t}{\ln(1.000, \frac{1}{2})} = \frac{t}{\ln(1.000, 5)} = \overline{s} + \frac{t}{\ln(1.000, 5)} = \overline{s} + \frac{t}{\ln(1.000, 5)}$$

$$\overline{s} = (s-t) \circ \frac{t}{\ln(1.000, 5)} = \int_{t}^{s} \frac{t}{\ln(1.000, 5)} dt = \int_{t}^{s} 25 \cdot \frac{t}{2} dt = \frac{25}{2} \int_{t}^{s} t^2 dt = \frac{25}{2} \cdot \frac{s^3 - t^3}{3} = \frac{25}{6} (s^3 - t^3)$$

$$\overline{s} = \frac{1}{125} \int_{s}^{s+t} \frac{t}{2} dt = \frac{1}{125} \cdot \frac{1}{3} \left[\frac{t^2}{2} \right]_{s}^{s+t} = \frac{1}{125} \cdot \frac{1}{3} \cdot \frac{(s+t)^2 - s^2}{2} = \frac{1}{125} \cdot \frac{1}{3} \cdot \frac{2st + 2t^2}{2} = \frac{1}{125} \cdot \frac{1}{3} \cdot (st + t^2) = I$$

$$\overline{s} = \text{P}_{\text{hp}}(Y \leq y) \int_{s}^{s+t} y^2 dy + \text{P}_{\text{hp}}(Y > y) \int_{s}^{s+t} y^2 dy = \int_{s}^{s+t} y^2 dy = \overline{s}$$

↳ Unit no-explicit

$$P(Y \leq y) = P(F_X(x) \leq y) = P(X \leq F_X^{-1}(y)) = F_X(F_X^{-1}(y)) = y,$$
$$y \in (0,1) \rightarrow Y \sim \text{Unif}(0,1)$$

4.22.

↑

↓

$$X_1 \sim \text{Exp}(\lambda_1)$$

$$X_2 \sim \text{Exp}(\lambda_2)$$

$$\text{if } X_1 = 8 \Rightarrow \lambda_1 = \frac{1}{8}$$
$$\text{if } X_2 = 5 \Rightarrow \lambda_2 = \frac{1}{5}$$

A = "minimum value < 4"

I - избыточный остаток

II - недостаточный остаток

$$P(I|A) = ?$$

$$P(I|A) = \frac{P(A|I) P(I)}{P(A|I) P(I) + P(A|II) P(II)} \rightarrow P(A)$$

$$P(I) = 1/2$$

$$P(II) = 1/2$$

$$P(A) = P(X_1 < 4) \cdot \frac{1}{2} + P(X_2 < 4) \cdot \frac{1}{2} =$$

$$= 1 - e^{-\lambda_1 \cdot 4} \cdot \frac{1}{2} + 1 - e^{-\lambda_2 \cdot 4} \cdot \frac{1}{2} =$$

$$= \frac{1}{2} \left(1 - e^{-4/\lambda_1} + 1 - e^{-4/\lambda_2} \right) = \frac{1}{2} \left(2 - e^{-\frac{4}{8}} - e^{-\frac{4}{5}} \right)$$

$$P(A|I) P(I) = P(X_1 < 4) \cdot \frac{1}{2} = \frac{1}{2} (1 - e^{-\lambda_1 \cdot 4}) = \frac{1}{2} (1 - e^{-\frac{4}{8}})$$

$$\Rightarrow P(I|A) = \frac{\frac{1}{2} (1 - e^{-\frac{4}{8}})}{\frac{1}{2} \left(2 - e^{-\frac{4}{8}} - e^{-\frac{4}{5}} \right)}$$

n-18.

$$X, Y \sim \text{Unif}(0,1)$$

$$X+Y \sim ?$$

(not fully worked out)

$$S = ab = 8 = AB$$

$$\frac{S}{2} = cd = 2 = CD$$

(ab) greatest

(AB) greatest

Find \rightarrow first choose mostsmallest \rightarrow square - I
middle \rightarrow square - II

$$S = (A + B)^2$$

$$(I)S - (II)(A+B) = (A)II$$

$$(A)I + (II)(II(A)B + (II)(II(A)B))$$

$$A + (I)B$$

$$A + (II)B$$

$$= \frac{1}{2} \cdot (A + 2AB) + \frac{1}{2} \cdot (A + 2AB) = (A)I$$

$$= \frac{1}{2} \cdot (A + 2AB) + \frac{1}{2} \cdot (A + 2AB) =$$

$$\left(\frac{1}{2} \cdot (A + 2AB) \right) \frac{1}{2} = \left(\frac{1}{2} \cdot (A + 2AB) + \frac{1}{2} \cdot (A + 2AB) \right) \frac{1}{2} =$$

$$\left(\frac{1}{2} \cdot (A + 2AB) \right) \frac{1}{2} = \left(\frac{1}{2} \cdot (A + 2AB) + \frac{1}{2} \cdot (A + 2AB) \right) \frac{1}{2} = (A)II +$$

$$\left(\frac{1}{2} \cdot (A + 2AB) \right) \frac{1}{2}$$

$$\left(\frac{1}{2} \cdot (A + 2AB) \right) \frac{1}{2} = (A)II +$$

+

$$X \sim N(\mu, \sigma^2)$$

$$Y = \frac{X-\mu}{\sigma}, \quad Y \sim ?$$

$$\mathbb{E}[Y] = \mathbb{E}\left[\frac{X-\mu}{\sigma}\right] = \frac{\mathbb{E}[X]-\mu}{\sigma} = \frac{\mu-\mu}{\sigma} = 0$$

$$D Y = D\left(\frac{X-\mu}{\sigma}\right) = \frac{1}{\sigma^2} D X = \frac{1}{\sigma^2} \cdot \sigma^2 = 1$$

$$\Rightarrow Y \sim N(0,1)$$

↳ cbc cнага на нрометнбике

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$Y = \frac{X-\mu}{\sigma}$$

$$g^{-1}(y) = \mu + \sigma y$$

$$Y = \frac{X-\mu}{\sigma}$$

$$\begin{aligned} y\sigma &= X - \mu \\ y &= \mu + \sigma y \end{aligned}$$

$$(g^{-1}(y))' = \sigma$$



$$f_Y(y) = f_X(g^{-1}(y)) \left| (g^{-1})'(y) \right| =$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{(\mu+\sigma y-\mu)^2}{2\sigma^2}} \cdot \sigma =$$

$$= \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{y^2}{2}} \Rightarrow Y \sim N(0,1)$$

Теорема на Мабер-Ланна

• позволява да се установи нормалното разпределение за априорниятата на сумата от много случаи чието беше имат по добър разпределение.

$$Z \sim N(0,1), \quad X_n \sim \text{Bin}(n,p) \quad p \in (0,1)$$

$$\Pr\left(\frac{X_n - np}{\sqrt{np(1-p)}} \leq x\right) \xrightarrow{n \rightarrow \infty} \Pr(Z \leq x)$$