

MATH 154 Homework 1 Solutions: Spring 2024

1. (a) There are two die; choose one to roll a 6 with probability $\frac{1}{6}$. The other needs to roll a non-6 with probability $\frac{5}{6}$. Thus, the probability of exactly one six is $2 \cdot \frac{1}{6} \cdot \frac{5}{6} = \frac{5}{18}$
- (b) The probability of a single roll being odd is $\frac{1}{2}$. Due to independence, the probability of two rolls both being odd is $\frac{1}{4}$.
- (c) With two die, we can sum to 4 by rolling 1,3; 2,2; 3,1. Since each of the 36 combinations are equally likely, the probability of rolling a cumulative 4 is $\frac{1}{12}$
- (d) We can get a sum divisible by 3 by summing two numbers that are 0 mod 3, or one that is 1 mod 3 and one that is 2 mod 3. Adding probabilities and using the product rule, we get

$$\frac{1}{3} \cdot \frac{1}{3} + 2 \cdot \frac{1}{3} \cdot \frac{1}{3} = \frac{1}{3}$$

2. (a) If heads appears for the first time on the n th throw, tails were flipped on the first $n - 1$ throws. The probability is thus

$$(1 - p)^{n-1}p$$

- (b) If n is odd, the probability is 0. If n is even, the probability is

$$\binom{n}{\frac{n}{2}} p^{\frac{n}{2}} (1 - p)^{\frac{n}{2}}$$

since we choose $\frac{n}{2}$ flips to be heads and assign the appropriate probabilities.

- (c) Using similar logic as above, the probability of exactly two heads is

$$\binom{n}{2} p^2 (1 - p)^{n-2}$$

- (d) Using complimentary counting, probability of at least two heads is 1 minus the probability of zero or one, giving

$$1 - \binom{n}{0} (1 - p)^n - \binom{n}{1} p (1 - p)^{n-1}$$

3. The principle of inclusion and exclusion gives that

$$\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B).$$

However, we also have that

$$\mathbb{P}(A \cup B) = \mathbb{P}(\text{exactly one of } A, B) + \mathbb{P}(A \cap B).$$

Thus,

$$\mathbb{P}(\text{exactly one of } A, B) + \mathbb{P}(A \cap B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$$

$$\mathbb{P}(\text{exactly one of } A, B) = \mathbb{P}(A) + \mathbb{P}(B) - 2\mathbb{P}(A \cap B).$$

4. (a) By the definition of conditional probability,

$$\mathbb{P}(A^c|B^c \cap C^c)\mathbb{P}(B^c|C^c)\mathbb{P}(C^c) = \frac{\mathbb{P}(A^c \cap B^c \cap C^c)}{\mathbb{P}(B^c \cap C^c)} \cdot \frac{\mathbb{P}(B^c \cap C^c)}{\mathbb{P}(C^c)} \cdot \mathbb{P}(C^c) = \mathbb{P}(A^c \cap B^c \cap C^c).$$

By DeMorgan's Law,

$$\mathbb{P}(A^c \cap B^c \cap C^c) = \mathbb{P}((A^c \cup B^c \cup C^c)^c)$$

Thus, we have

$$\begin{aligned}\mathbb{P}(A \cup B \cup C) &= 1 - \mathbb{P}((A^c \cup B^c \cup C^c)^c) \\ &= 1 - \mathbb{P}(A^c|B^c \cap C^c)\mathbb{P}(B^c|C^c)\mathbb{P}(C^c).\end{aligned}$$

- (b) No. Imagine you have two coins, one fair, one that always returns heads. Choose a coin at random. Let A be the event that the first toss is H, and B be the event that the second is H. Let C be the event that the fair coin is chosen. Note that $\mathbb{P}(A|C) = \mathbb{P}(B|C) = \frac{1}{2}$, with $\mathbb{P}(A \cap B|C) = \frac{1}{4} = \mathbb{P}(A|C)\mathbb{P}(B|C)$. However,

$$\mathbb{P}(A) = \mathbb{P}(A|C)\mathbb{P}(C) + \mathbb{P}(A|C^c)\mathbb{P}(C^c) = \frac{1}{2} \cdot \frac{1}{2} + 1 \cdot \frac{1}{2} = \frac{3}{4}.$$

$\mathbb{P}(B) = \frac{3}{4}$ similarly. However,

$$\mathbb{P}(A \cap B) = \mathbb{P}(A \cap B|C)\mathbb{P}(C) + \mathbb{P}(A \cap B|C^c)\mathbb{P}(C^c) = \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2} + 1 \cdot 1 \cdot \frac{1}{2} = \frac{5}{8}$$

and conditional independence does not imply independence.

- (c) No. Suppose you roll a die. Let A be an event of rolling a 1 or 2; B be rolling 2,4,6; C be rolling 1,4. We have $\mathbb{P}(A) = \frac{1}{3}, \mathbb{P}(B) = \frac{1}{2}, \mathbb{P}(A \cup B) = \frac{1}{6} = \mathbb{P}(A)\mathbb{P}(B)$. However, we have $\mathbb{P}(A|C) = \frac{1}{2}, \mathbb{P}(B|C) = \frac{1}{2}, \mathbb{P}(A \cup B|C) = 0 \neq \mathbb{P}(A|C)\mathbb{P}(B|C)$. Thus, independence does not imply conditional independence.

Grading:

- (a) 4 points:

- 0 points - no correct steps
- 2 points - some correct steps
- 4 points - all correct steps

- (b) 3 points:

- 0 points - incorrect answer
- 1 point - correct answer, incorrect counterexample
- 3 points - correct answer, correct counterexample

- (c) 3 points:

- 0 points - incorrect answer
- 1 point - correct answer, incorrect counterexample
- 3 points - correct answer, correct counterexample

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5. By the definition of conditional probability,

$$\mathbb{P}(A_j|B) = \frac{\mathbb{P}(A_j \cap B)}{\mathbb{P}(B)}.$$

Using the definition of conditional probability in the numerator and the law of total probability in the denominator, we get

$$\mathbb{P}(A_j|B) = \frac{\mathbb{P}(B|A_j)\mathbb{P}(A_j)}{\sum_{i=1}^n \mathbb{P}(B|A_i)\mathbb{P}(A_i)}.$$

Grading:

- (a) +3 points - initial step of using definition of conditional probability
 - (b) +3 points - using definition of conditional probability to transform numerator
 - (c) +4 points - using LOTP to transform denominator
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6. Let $|A| = a$, $|B| = b$, $|A \cap B| = c$. If A and B are independent,

$$\frac{c}{p} = \mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B) = \frac{a}{p} \frac{b}{p} \Rightarrow cp = ab.$$

Since p is prime and $p|ab$, $p|a$ or $p|b$, implying that at least one of a or b is 0 or p , or, equivalently, at least one of A or B is \emptyset or Ω .

7. (a)

$$\begin{aligned} \mathbb{P}(N = 2|S = 4) &= \frac{\mathbb{P}(S = 4|N = 2)\mathbb{P}(N = 2)}{\mathbb{P}(S = 4)} \\ &= \frac{\frac{1}{12} \cdot \frac{1}{4}}{\sum_{i=1}^{\infty} \mathbb{P}(S = 4|N = i)\mathbb{P}(N = i)} \\ &= \frac{\frac{1}{48}}{\sum_{i=1}^4 \mathbb{P}(S = 4|N = i)\mathbb{P}(N = i)} \\ &= \frac{\frac{1}{48}}{\frac{1}{6} \cdot \frac{1}{2} + \frac{1}{12} \cdot \frac{1}{4} + \frac{3}{6^3} \cdot \frac{1}{8} + \frac{1}{6^4} \cdot \frac{1}{16}} \\ &= \frac{\frac{1}{12}}{\frac{2197}{20736}} \\ &= \frac{432}{2197}. \end{aligned}$$

(b)

$$\begin{aligned}
\mathbb{P}(S = 4|N \text{ even}) &= \frac{\mathbb{P}(S = 4|N = 2)\mathbb{P}(N = 2) + \mathbb{P}(S = 4|N = 4)\mathbb{P}(N = 4)}{\mathbb{P}(N \text{ even})} \\
&= \frac{\frac{1}{12} \cdot \frac{1}{4} + \frac{1}{6^4} \cdot \frac{1}{16}}{\frac{\frac{1}{4}}{1 - \frac{1}{4}}} \\
&= \frac{433}{6912}
\end{aligned}$$

(c) By Bayes Rule,

$$\begin{aligned}
\mathbb{P}(N = 2|S = 4, F = 1) &= \frac{\mathbb{P}(S = 4|N = 2, F = 1)\mathbb{P}(N = 2|F = 1)}{\mathbb{P}(S = 4|F = 1)} \\
&= \frac{\frac{1}{6} \cdot \mathbb{P}(N = 2)}{\sum_{i=1}^{\infty} \mathbb{P}(S = 4|F = 1, N = i)\mathbb{P}(N = i)} \\
&= \frac{\frac{1}{6} \cdot \frac{1}{4}}{\sum_{i=1}^4 \mathbb{P}(S = 4|F = 1, N = i)\mathbb{P}(N = i)} \\
&= \frac{\frac{1}{24}}{0 \cdot \frac{1}{2} + \frac{1}{6} \cdot \frac{1}{4} + \frac{2}{36} \cdot \frac{1}{8} + \frac{1}{6^3} \cdot \frac{1}{16}} \\
&= \frac{144}{169}.
\end{aligned}$$

8. Let X be the event that there is an open road from A to B , Y be the event that there is an open road from B to C , and Z be the event that there is no open route from A to C . We have

$$\mathbb{P}(X|Z) = \frac{\mathbb{P}(Z|X)\mathbb{P}(X)}{\mathbb{P}(Z)} = \frac{\mathbb{P}(Y^c)\mathbb{P}(X)}{\mathbb{P}(X^c) + \mathbb{P}(Y^c) - \mathbb{P}(X^c \cap Y^c)} = \frac{p^2 \cdot (1 - p^2)}{p^2 + p^2 - p^4} = \frac{1 - p^2}{2 - p^2}.$$

9. Since 0 is even, $p_0 = 1$. Consider $n - 1$ tosses for $n \geq 1$. If there were an odd number of H tossed in the first $n - 1$ tosses, the n th toss must be a head for the total count to be even after n tosses. In other words, we contribute a term

$$\mathbb{P}(n \text{ head})\mathbb{P}(\text{odd H after } n - 1) = p(1 - p_{n-1})$$

If there were an even number of H tossed in the first $n - 1$ tosses, the n th toss must be a tail for the total count to be even after n tosses. In other words, we contribute a term

$$\mathbb{P}(n \text{ tail})\mathbb{P}(\text{even H after } n - 1) = (1 - p)p_{n-1}.$$

Thus, we indeed have that

$$p_n = p(1 - p_{n-1}) + (1 - p)p_{n-1}.$$

Rearranging, we get

$$p_n = p + (1 - 2p)p_{n-1} \Rightarrow p_n - \frac{1}{2} = (1 - 2p)(p_{n-1} - \frac{1}{2}).$$

Define the sequence $a_n = p_n - \frac{1}{2}$. Since $p_0 = 1, a_0 = \frac{1}{2}$. The above gives the relation

$$a_n = (1 - 2p)a_{n-1},$$

which is simply a geometric series. Thus, a closed form for $a_n = \frac{1}{2}(1 - 2p)^n \Rightarrow p_n = \frac{1}{2} + \frac{1}{2}(1 - 2p)^n$. Plugging in this closed form expression also suffices for proof.

10. The crux of the passenger problem is that, regardless of how all $n - 1$ passengers behave beforehand, the last passenger will only have two potential seats to choose from; the first person's seat or their assigned seat. To see why, note that if the i th person's seat was open, the i th person *must* sit there, but the emptiness of that seat contradicts the fact that the i th person is already seated. Thus, the last passenger will always only have those two potential seats, regardless of n . Symmetry gives the answer $\frac{1}{2}$.
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