

# Generalized Method of Moments (GMM)

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## 1 1. What is GMM?

### Idea

The **Generalized Method of Moments (GMM)** is a flexible estimation framework that uses the **moment conditions** implied by economic theory or statistical models.

If a model implies:

$$E[m(W_i, \theta_0)] = 0,$$

where  $m(\cdot)$  is a  $q \times 1$  vector of functions of the data  $W_i$  and parameters  $\theta$ , and  $q \geq k = \dim(\theta)$ , then GMM estimates  $\theta_0$  by making the sample analog as close to zero as possible.

$$\hat{\theta}_{GMM} = \arg \min_{\theta} g_n(\theta)' W_n g_n(\theta),$$

where

$$g_n(\theta) = \frac{1}{n} \sum_{i=1}^n m(W_i, \theta),$$

and  $W_n$  is a positive definite **weighting matrix**.

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### Intuition

- The model says: "At the truth, certain sample moments should equal zero." - GMM chooses  $\hat{\theta}$  to make the sample moments as close to zero as possible. - Flexible: does not require specifying a full likelihood (unlike MLE).

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## 2 2. OLS as GMM

### Model

$$y_i = x_i' \beta + u_i, \quad E[u_i | x_i] = 0.$$

## Moment Condition

Conditional mean assumption implies:

$$E[x_i u_i] = E[x_i(y_i - x_i' \beta)] = 0.$$

So define  $m_i(\beta) = x_i(y_i - x_i' \beta)$ .

Sample moment:

$$g_n(\beta) = \frac{1}{n} X'(Y - X\beta).$$

## GMM Criterion

$$Q_n(\beta) = g_n(\beta)' W g_n(\beta).$$

Taking  $W = (E[x_i x_i'])^{-1}$  (or sample analog  $(X'X/n)^{-1}$ ):

$$Q_n(\beta) = (Y - X\beta)' X (X'X)^{-1} X' (Y - X\beta).$$

## Minimization

Differentiate and set to zero:

$$\frac{\partial Q_n}{\partial \beta} = -2X'(Y - X\beta) = 0 \quad \Rightarrow \quad \hat{\beta}_{OLS} = (X'X)^{-1} X' Y.$$

**Therefore:** OLS is GMM with moment  $E[x_i(y_i - x_i' \beta)] = 0$  and weight  $W = (X'X/n)^{-1}$ .

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## 3 3. GLS as GMM

### Model

$$y = X\beta + u, \quad E[u|X] = 0, \quad Var(u|X) = \Omega.$$

Here  $\Omega$  may be non-scalar (heteroskedastic or correlated).

### Moment Condition

Same as OLS:

$$E[X'u] = 0.$$

But now moments have covariance:

$$E[X'uu'X] = X'\Omega X.$$

## Efficient GMM Weight

Optimal weight  $W = (X'\Omega X)^{-1}$  corresponds to weighting each moment by its variance.

Criterion:

$$Q(\beta) = (Y - X\beta)' \Omega^{-1} (Y - X\beta).$$

Minimization gives:

$$\hat{\beta}_{GLS} = (X'\Omega^{-1}X)^{-1}X'\Omega^{-1}Y.$$

**Therefore:** GLS is GMM with moment  $E[X'u] = 0$  and optimal weighting matrix  $W = \Omega^{-1}$ .

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## 4 4. IV / 2SLS as GMM

### Model

$$Y = X\beta + u, \quad E[Zu] = 0, \quad E[ZX] \text{ full rank.}$$

Here  $Z$  are instruments ( $n \times q, q \geq k$ ).

### Moment Condition

$$E[Z'(Y - X\beta)] = 0.$$

Sample analog:

$$g_n(\beta) = \frac{1}{n} Z'(Y - X\beta).$$

### GMM Criterion

$$Q_n(\beta) = g_n(\beta)' W g_n(\beta) = (Y - X\beta)' Z W Z' (Y - X\beta).$$

First-order condition:

$$X' Z W Z' (Y - X\beta) = 0.$$

### Solution

$$\hat{\beta}_{GMM} = (X' Z W Z' X)^{-1} X' Z W Z' Y.$$

#### Special Cases:

- If  $W = (Z'Z)^{-1}$ , then

$$\hat{\beta}_{GMM} = (X' P_Z X)^{-1} X' P_Z Y = \hat{\beta}_{2SLS},$$

where  $P_Z = Z(Z'Z)^{-1}Z'$ .

- If  $W = \Omega^{-1}$ , where  $\Omega = E[Z'uu'Z]$ , then we get the **efficient GMM / optimal IV estimator**.
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## Relationship Among OLS, GLS, IV

Estimator	Moment Condition	Weighting Matrix $W$	Estimator
OLS	$E[X'(Y - X\beta)] = 0$	$(X'X/n)^{-1}$	$(X'X)^{-1}X'Y$
GLS	$E[X'(Y - X\beta)] = 0$	$\Omega^{-1}$	$(X'\Omega^{-1}X)^{-1}X'\Omega^{-1}Y$
IV (2SLS)	$E[Z'(Y - X\beta)] = 0$	$(Z'Z/n)^{-1}$	$(X'P_Z X)^{-1}X'P_Z Y$

All are special cases of:

$$\hat{\beta}_{GMM} = (X'Z W Z' X)^{-1} X' Z W Z' Y,$$

where  $(Z, W)$  define the model's moment restrictions and weighting scheme.

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## 5 5. Intuition and Efficiency

- Each estimator arises from choosing **different moments and weights**.
- OLS uses its own regressors as instruments ( $Z = X$ ), with homoskedastic errors.
- GLS uses known error covariance  $\Omega$  as a weight.
- IV/2SLS uses external instruments  $Z$  when regressors are endogenous.

**Efficient GMM:** When the weighting matrix equals the inverse of the true moment covariance,  $W = \Omega^{-1}$ , GMM achieves the Cramér–Rao lower bound among all estimators that use the same moments.

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## 6 6. Visual Summary

GMM unifies all linear estimators:  $\hat{\beta}_{GMM} = (X'Z W Z' X)^{-1} X' Z W Z' Y$

$$Z = X, W = (X'X)^{-1} \Rightarrow \text{OLS}$$

$$Z = X, W = \Omega^{-1} \Rightarrow \text{GLS}$$

$$Z \neq X, W = (Z'Z)^{-1} \Rightarrow \text{IV / 2SLS}$$

$$Z \neq X, W = \Omega^{-1} \Rightarrow \text{Efficient GMM (Optimal IV)}$$

## 7 3. GMM and Extremum Estimators

### Definition

An **extremum estimator** solves:

$$\hat{\theta} = \arg \max_{\theta} Q_n(\theta)$$

(or  $\arg \min$ ). Examples: MLE, NLS, GMM.

GMM is a particular extremum estimator where:

$$Q_n(\theta) = -g_n(\theta)'W_n g_n(\theta),$$

so it fits perfectly within the extremum framework.

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### **Consistency and Asymptotic Normality**

Under regularity conditions:

$$\sqrt{n}(\hat{\theta} - \theta_0) \xrightarrow{d} N\left(0, (G'WG)^{-1}G'W\Omega WG(G'WG)^{-1}\right),$$

where:

$$G = E\left[\frac{\partial m(W_i, \theta_0)}{\partial \theta'}\right], \quad \Omega = E[m(W_i, \theta_0)m(W_i, \theta_0)'].$$


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### **Optimal GMM**

Efficiency in GMM depends on choice of  $W$ . The asymptotic variance is minimized when:

$$W = \Omega^{-1}.$$

Then:

$$\text{Avar}(\hat{\theta}_{opt}) = (G'\Omega^{-1}G)^{-1}.$$


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## **8 4. Detailed Econometrics: Steps, Testing, and Implementation**

### **Step 1: Model moments**

Specify  $q$  moment conditions:

$$E[m(W_i, \theta_0)] = 0.$$


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### **Step 2: Sample moments**

Compute sample analog:

$$g_n(\theta) = \frac{1}{n} \sum_i m(W_i, \theta).$$


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### **Step 3: Choose initial weighting matrix**

- Often start with  $W_n = I_q$  (identity) or  $(Z'Z/n)^{-1}$ . - Compute preliminary GMM estimate  $\tilde{\theta}$ .

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#### Step 4: Estimate the optimal weighting matrix

Estimate

$$\hat{\Omega} = \frac{1}{n} \sum_i m(W_i, \tilde{\theta}) m(W_i, \tilde{\theta})'.$$

Set

$$\hat{W} = \hat{\Omega}^{-1}.$$

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#### Step 5: Two-step efficient GMM

Re-estimate:

$$\hat{\theta}_{GMM} = \arg \min_{\theta} g_n(\theta)' \hat{W} g_n(\theta).$$

Iterating again yields the same asymptotic efficiency.

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#### Step 6: Over-identification Test (Sargan–Hansen $J$ test)

If  $q > k$ , the model is over-identified.

$$J = n g_n(\hat{\theta}_{GMM})' \hat{W} g_n(\hat{\theta}_{GMM}) \sim \chi^2_{q-k}.$$

Interpretation: - Small  $J$ : moments consistent  $\Rightarrow$  instruments valid. - Large  $J$ : reject over-identifying restrictions  $\Rightarrow$  possible invalid instruments.

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#### Step 7: Robust Standard Errors

Variance–covariance (sandwich) estimator:

$$\widehat{\text{Var}}(\hat{\theta}_{GMM}) = (G' \hat{W} G)^{-1} G' \hat{W} \hat{\Omega} \hat{W} G (G' \hat{W} G)^{-1}.$$

In practice, software packages compute this automatically.

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## 9 5. GMM in Practice

- **When to use:** models defined by moment restrictions (e.g., IV, dynamic panels, consumption Euler equations).
- **Strengths:** general, robust, efficient under weak distributional assumptions.
- **Weaknesses:** finite-sample bias, sensitivity to weighting matrix, weak identification.

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## 10 6. Summary Table

Estimator	Moment condition	Weight W	Efficiency criterion
OLS	$E[X'u] = 0$	$(X'X)^{-1}$	Homoskedasticity
GLS	$E[X'u] = 0$	$\Omega^{-1}$	Heteroskedasticity / known $\Omega$
IV / 2SLS	$E[Z'u] = 0$	$(Z'Z)^{-1}$	Linear IV
Efficient GMM	$E[m(W_i, \theta)] = 0$	$\Omega^{-1}$	Optimal weighting

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## 11 7. Intuitive Summary

**GMM unifies classical estimators:** OLS, GLS, IV are all special cases. It replaces “derive a formula” with “impose a moment condition.”

**GMM is flexible:** Works for nonlinear, dynamic, and structural models. In large samples, efficient GMM = MLE when both are correctly specified.

**Philosophy:** “Estimate parameters so that the theoretical moments match the data moments as closely as possible.”