Lecture 8: Nonlinearity (Polynomial Regression, Spline & GAM)

BIOS635 02/04/2020

Moving beyond linearity

The truth is never linear!

Or almost never!

But often the linearity assumption is good enough.

When its not ...

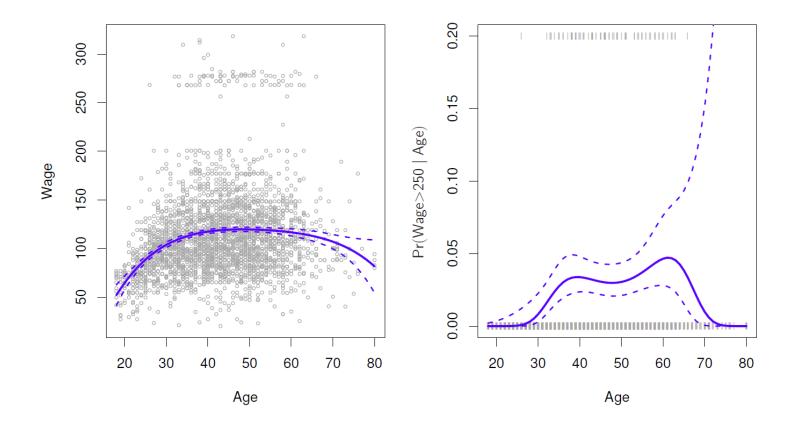
- polynomials,
- step functions,
- splines,
- local regression, and
- generalized additive models

offer a lot of flexibility, without losing the ease and interpretability of linear models.

Polynomial regression

$$y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \beta_3 x_i^3 + \ldots + \beta_d x_i^d + \epsilon_i$$

Degree-4 Polynomial



Polynomial regression: linear

- Create new variables $X_1 = X$, $X_2 = X^2$, etc and then treat as multiple linear regression.
- Not really interested in the coefficients; more interested in the fitted function values at any value x_0 :

$$\hat{f}(x_0) = \hat{\beta}_0 + \hat{\beta}_1 x_0 + \hat{\beta}_2 x_0^2 + \hat{\beta}_3 x_0^3 + \hat{\beta}_4 x_0^4.$$

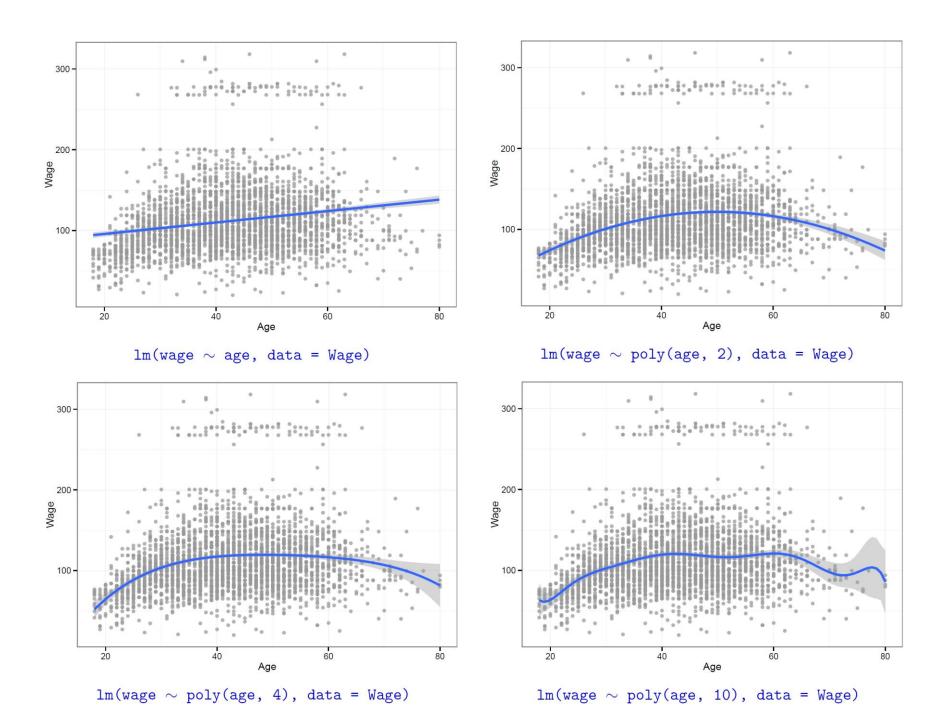
- Since $\hat{f}(x_0)$ is a linear function of the $\hat{\beta}_{\ell}$, can get a simple expression for *pointwise-variances* $\operatorname{Var}[\hat{f}(x_0)]$ at any value x_0 . In the figure we have computed the fit and pointwise standard errors on a grid of values for x_0 . We show $\hat{f}(x_0) \pm 2 \cdot \operatorname{se}[\hat{f}(x_0)]$.
- We either fix the degree d at some reasonably low value, else use cross-validation to choose d.

Polynomial regression: logistic

• Logistic regression follows naturally. For example, in figure we model

$$\Pr(y_i > 250|x_i) = \frac{\exp(\beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots + \beta_d x_i^d)}{1 + \exp(\beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots + \beta_d x_i^d)}.$$

- To get confidence intervals, compute upper and lower bounds on *on the logit scale*, and then invert to get on probability scale.
- Can do separately on several variables—just stack the variables into one matrix, and separate out the pieces afterwards (see GAMs later).
- Caveat: polynomials have notorious tail behavior very bad for extrapolation.
- Can fit using $y \sim poly(x, degree = 3)$ in formula.



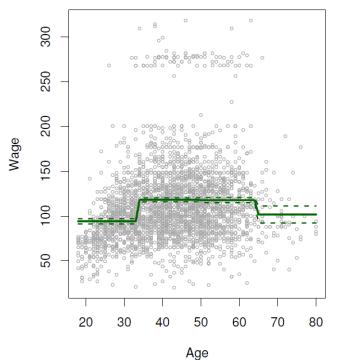
Step functions

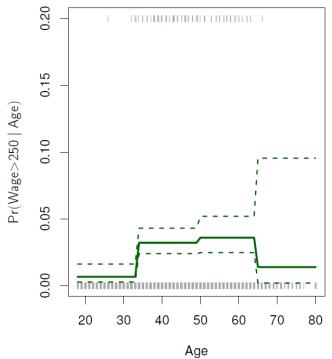
Another way of creating transformations of a variable — cut the variable into distinct regions.

$$C_1(X) = I(X < 35), \quad C_2(X) = I(35 \le X < 50), \dots, C_3(X) = I(X \ge 65)$$

Piecewise Constant

Age	C_1	C_2	C_3
18		0	0
24	I	0	0
45	0	I	0
67	0	0	1
54	0	I	0
:		:	:





Step functions, cont'd

- Easy to work with. Creates a series of dummy variables representing each group.
- Useful way of creating interactions that are easy to interpret. For example, interaction effect of Year and Age:

$$I(\texttt{Year} < 2005) \cdot \texttt{Age}, \quad I(\texttt{Year} \geq 2005) \cdot \texttt{Age}$$

would allow for different linear functions in each age category.

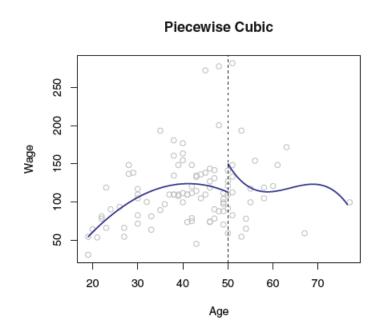
- In R: I(year < 2005) or cut(age, c(18, 25, 40, 65, 90)).
- Choice of cutpoints or *knots* can be problematic. For creating nonlinearities, smoother alternatives such as *splines* are available.

Piecewise polynomials

• Instead of a single polynomial in X over its whole domain, we can rather use different polynomials in regions defined by knots. E.g. (see figure)

$$y_i = \begin{cases} \beta_{01} + \beta_{11}x_i + \beta_{21}x_i^2 + \beta_{31}x_i^3 + \epsilon_i & \text{if } x_i < c; \\ \beta_{02} + \beta_{12}x_i + \beta_{22}x_i^2 + \beta_{32}x_i^3 + \epsilon_i & \text{if } x_i \ge c. \end{cases}$$

- Better to add constraints to the polynomials, e.g. continuity.
- Splines have the "maximum" amount of continuity.

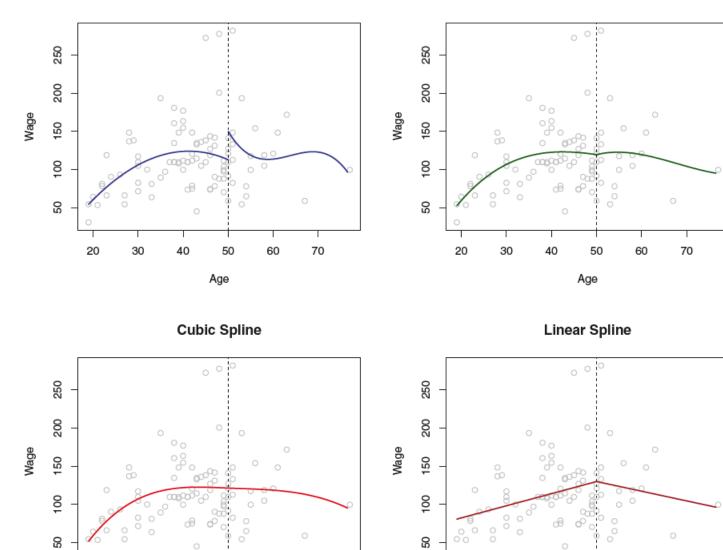




Age

Continuous Piecewise Cubic

Age



Linear splines

A linear spline with knots at ξ_k , k = 1, ..., K is a piecewise linear polynomial continuous at each knot.

We can represent this model as

$$y_i = \beta_0 + \beta_1 b_1(x_i) + \beta_2 b_2(x_i) + \dots + \beta_{K+3} b_{K+3}(x_i) + \epsilon_i,$$

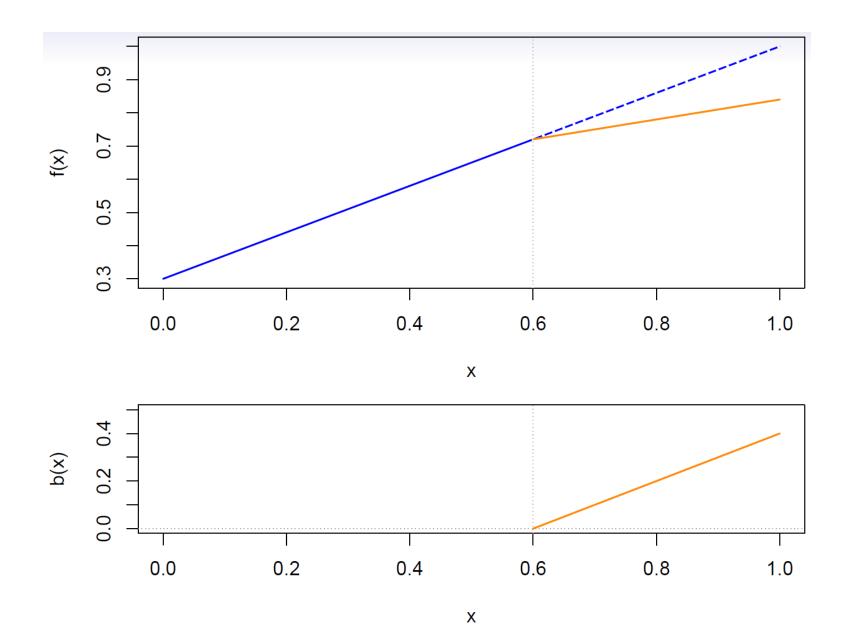
where the b_k are basis functions.

$$b_1(x_i) = x_i$$

 $b_{k+1}(x_i) = (x_i - \xi_k)_+, \quad k = 1, \dots, K$

Here the $()_{+}$ means positive part; i.e.

$$(x_i - \xi_k)_+ = \begin{cases} x_i - \xi_k & \text{if } x_i > \xi_k \\ 0 & \text{otherwise} \end{cases}$$



Cubic splines

A cubic spline with knots at ξ_k , k = 1, ..., K is a piecewise cubic polynomial with continuous derivatives up to order 2 at each knot.

Again we can represent this model with truncated power basis functions

$$y_{i} = \beta_{0} + \beta_{1}b_{1}(x_{i}) + \beta_{2}b_{2}(x_{i}) + \dots + \beta_{K+3}b_{K+3}(x_{i}) + \epsilon_{i},$$

$$b_{1}(x_{i}) = x_{i}$$

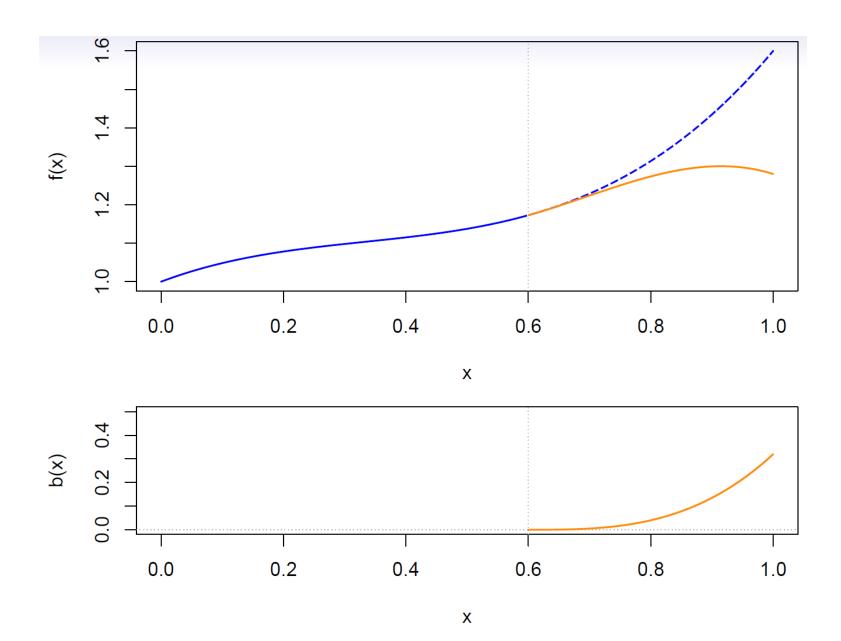
$$b_{2}(x_{i}) = x_{i}^{2}$$

$$b_{3}(x_{i}) = x_{i}^{3}$$

$$b_{k+3}(x_{i}) = (x_{i} - \xi_{k})_{+}^{3}, \quad k = 1, \dots, K$$

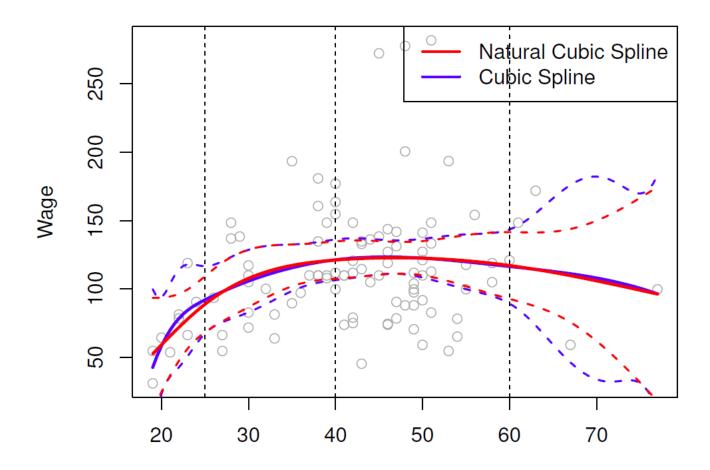
where

$$(x_i - \xi_k)_+^3 = \begin{cases} (x_i - \xi_k)^3 & \text{if } x_i > \xi_k \\ 0 & \text{otherwise} \end{cases}$$



Natural cubic splines

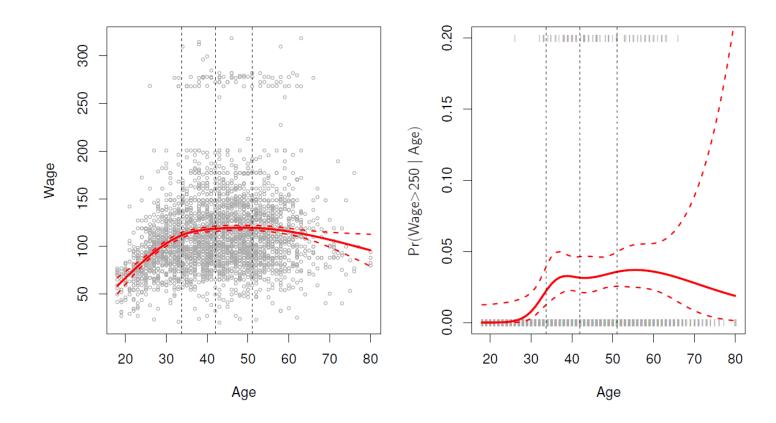
Second derivatives of the spline polynomials are set equal to zero at the end points of interpolation. This forces the spline to be a straight line outside the interval.



Fitting splines in R

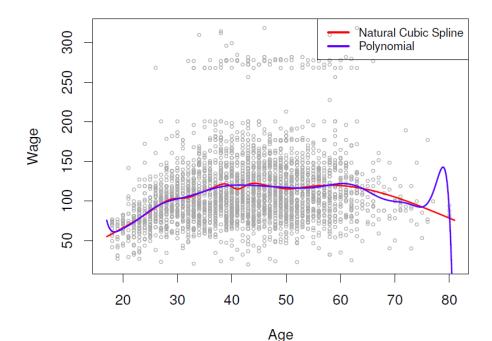
Fitting splines in R is easy: bs(x, ...) for any degree splines, and ns(x, ...) for natural cubic splines, in package splines.

Natural Cubic Spline



Knot placement & degree of freedom

- One strategy is to decide K, the number of knots, and then place them at appropriate quantiles of the observed X.
- A cubic spline with K knots has K+4 parameters or degrees of freedom.
- A natural spline with K knots has K degrees of freedom.



Comparison of a degree-14 polynomial and a natural cubic spline, each with 15df.

```
ns(age, df=14)
poly(age, deg=14)
```

Smoothing splines

• The smoothing spline estimator is the solution \hat{g} to the problem

minimize
$$\underbrace{\sum_{i=1}^{n}(y_i-g(x_i))^2}_{\text{RSS}} + \underbrace{\lambda \int g''(t)^2 dt}_{\text{Roughness penalty}}$$

- This is a penalized regression problem
- We're saying we want a function that:
 - Fits the data well; and
 - isn't too wiggly
- Large $\lambda \implies \hat{g}$ will have low variability (& higher bias)
- Small $\lambda \implies \hat{g}$ will have high variability (& lower bias)

How is this at all related to splines?

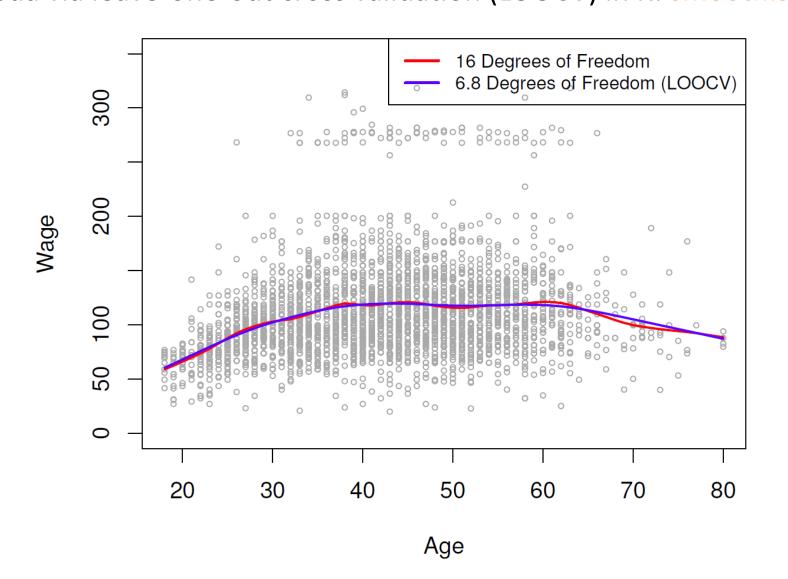
Smoothing splines, cont'd

minimize
$$\underbrace{\sum_{i=1}^{n}(y_i-g(x_i))^2}_{\text{RSS}} + \underbrace{\lambda \int g''(t)^2 dt}_{\text{Roughness penalty}} \tag{*}$$

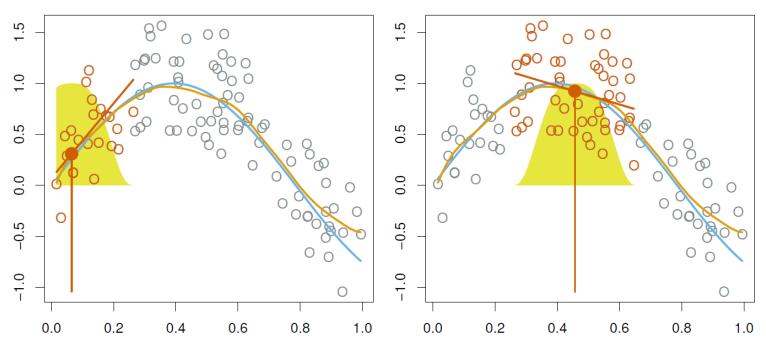
It turns out...

- The solution to (*) is a natural cubic spline
- The solution has knots at every unique value of x
- The effective degrees of freedom of the solution is calculable
- $\lambda \longleftrightarrow df$

Specify df in R: smooth.spline(age, wage, df = 10)
Choose lambda via leave-one-out cross validation (LOOCV) in R: smooth.spline(age, wage)



Local regression



With a sliding weight function, we fit separate linear fits over the range of X by weighted least squares.

See text for more details, and loess() function in R.

Nonlinearity coding in R

Model	R command	
Degree-d polynomial regression	~ poly(x, degree = d)	
Step functions with knots c1, c2, c3	\sim cut(x, breaks = c(c1, c2, c3))	
Cubic spline	~ bs(x, df)	
Natural cubic spline	~ ns(x, df)	
Degree-d spline	~ bs(x, df, degree = d)	
Smoothing spline	~ s(x, df)	
Local linear regression	~ lo(x)	

Generalized additive models (GAMs)

Recall the Linear Regression Model

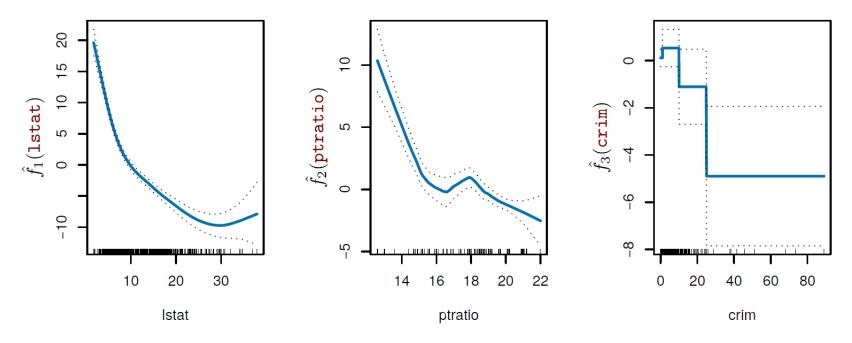
$$Y = \beta_0 + \sum_{j=1}^{p} \beta_j X_j + \epsilon$$

We can now extend this to the far more flexible Additive Model

$$Y = \beta_0 + \sum_{j=1}^{p} f_j(X_j) + \epsilon$$

- Each f_j can be any of the different methods we just talked about: Linear term $(\beta_j X_j)$, Polynomial, Step Function, Piecewise Polynomial, Degree-k spline, Natural cubic spline, Smoothing spline, Local linear regression fit, ...
- You can mix-and-match different kinds of terms
- The gam and mgcv packages enable Additive Models in R

GAMs: Boston housing data

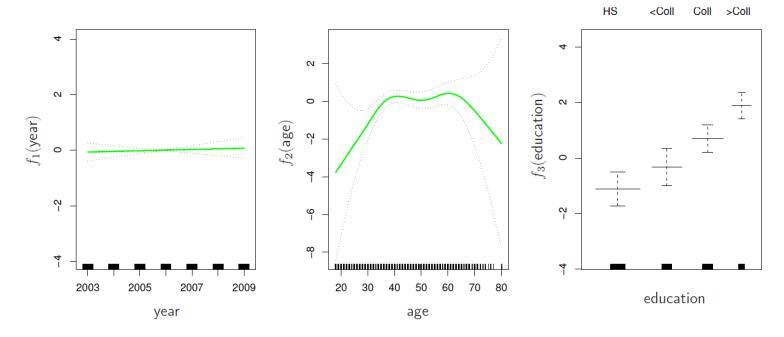


$$medv = f_1(lstat) + f_2(ptratio) + f_3(crim) + \epsilon$$

- $f_1(lstat)$ smoothing spline with 5 df
- $f_2(ptratio)$ local linear regression
- $f_3(\text{crim})$ step function with breaks at crim = 1, 10, 25

GAMs for classification

$$\log\left(\frac{p(X)}{1-p(X)}\right) = \beta_0 + f_1(X_1) + f_2(X_2) + \dots + f_p(X_p).$$



 $gam(I(wage > 250) \sim year + s(age, df = 5) + education, family = binomial)$