

~~(x,y)~~

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C is closed if $x \in C$ and $p_{xy} > 0 \Rightarrow y \in C$. (Note if C is closed and $x \in C$, then $\mathbb{P}_x(X_n \in C) = 1 \forall n$) D is irreducible if $x, y \in D$ implies $p_{xy} > 0$ ($p_{yx} > 0$).

Theorem: Let C be a finite closed set. Then, C contains a recurrent set. If C is irreducible, then all states in C are recurrent.

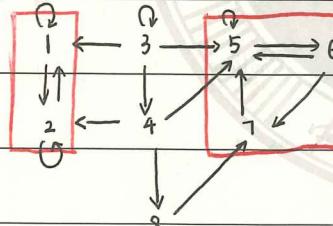
Pf. In view of the previous theorem, we only need to show the first claim.

If otherwise, $\forall y \in C$, $p_{yy} < 1$ and $\mathbb{E}_x N(y) = \frac{p_{xy}}{1-p_{yy}}$. But,

$$\sum_{y \in C} \mathbb{E}_x N(y) = \sum_{y \in C} \sum_{n=1}^{\infty} p^n(x, y) = \sum_{n=1}^{\infty} \sum_{y \in C} p^n(x, y) = \sum_{n=1}^{\infty} 1$$

C is closed \square

Example:



	1	2	3	4	5	6	7	8
1	0.4	0.6	0	0	0	0	0	0
2	0.2	0.8	0	0	0	0	0	0
3	0.1	0	0.2	0.3	0.4	0	0	0
4	0	0.6	0	0	0.2	0	0	0.2
5	0	0	0	0	0.3	0.7	0	0
6	0	0	0	0	0.2	0	0.8	0
7	0	0	0	0	0	0	1	0
8	0	0	0	0	0	0	0.1	0.9

(i) $p_{31} > 0$ and $p_{13} = 0$, so 3 must be transient. Similarly, $p_{42} > 0$ and $p_{24} = 0$, so 4 must be transient; $p_{78} > 0$ and $p_{87} = 0$, so 8 is transient.

(ii) $\{1, 2\}$ and $\{5, 6, 7\}$ are irreducible closed sets, ^{so} they are recurrent.

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This reasoning can be used to classify transient & recurrent states when S is finite.

(i) There is a y with $p_{xy} > 0$ and $p_{yy} = 0$, then x must be ~~closed~~ transient.

(ii) $p_{xy} > 0$ implies $p_{yx} > 0$. Then, let $C_x = \{y : p_{xy} > 0\}$. We claim that C_x is an irreducible closed set, so all states in C_x is recurrent.

[If $y, z \in C_x$, then $p_{yz} \geq p_{yx} p_{xz} > 0$. Hence, C_x is irreducible. If $p_{yw} > 0$ for $y \in C_x$, then $p_{xw} \geq p_{xy} p_{yw} > 0$, so $w \in C_x$. Hence, C_x is closed.]

Theorem (Decomposition theorem) Let $R = \{x : p_{xx} = 1\}$ be the set of recurrent states of a Markov chain. Then, $R = \bigcup_i R_i$, where each R_i is closed and irreducible.

Rmk. For the study of ^a recurrent state x , we can consider a single irreducible closed set containing x .

Pf.: $\forall x \in R$, let $C_x = \{y : p_{xy} > 0\}$. We have shown that C_x is a irreducible closed set.

If $C_y \cap C_x \neq \emptyset$, then we have $C_x = C_y$. ($\text{If } \exists z \in C_x \cap C_y, \text{ then } p_{xyz} \geq p_{xz} p_{zy} > 0 \Rightarrow z \in C_x \Rightarrow C_y \subseteq C_x$. Similarly, $C_x \subseteq C_y$.)

Hence, we have either $C_x \cap C_y = \emptyset$ or $C_x = C_y$. This decomposes R into disjoint unions of equivalent classes. \square

Example: (Branching process) Suppose the probability of no children is positive

(i.e., $p_0 > 0$). Then, all states are transient: $p_{k0} > 0$ but $p_{kk} = 0 \quad \forall k \geq 1$. The state 0 is recurrent since $p_{(0,0)} = 1$. 0 is an absorbing state.

Example: (Birth - Death chain) State space $S = \{0, 1, 2, \dots\}$. Let $p(i, i+1) = p_i^{>0}$

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$$T_a = \inf \{n \geq 1 : X_n = a\}, T_b = \inf \{n \geq 1 : X_n = b\}$$

$$p(i, i+1) = q_i^{>0} \quad \text{and} \quad p(i, i) = r_i. \quad p_i + q_i + r_i = 1. \quad \text{Let } q_0 = 0. \quad N = \inf \{n : X_n = 0\}.$$

We now define a function φ so that $\varphi(X_{N+n})$ is a martingale. First, we have

$$\text{For any } 0 \leq x \leq b, \text{ let } \varphi(x) = \mathbb{P}_x(T_a < T_b). \quad \text{Then, using the Markov property, we get: } \varphi(x) = p_x \mathbb{P}_{x+1}(T_a < T_b) + q_x \mathbb{P}_{x-1}(T_a < T_b) + r_x \varphi(x+1) + q_x \varphi(x-1).$$

$$\text{Moreover, } \varphi(a-1) = 1, \quad \varphi(b+1) = 0. \quad \Rightarrow \varphi(b+1) = p_a \varphi(a+1) + q_a \varphi(a-1) + r_a \varphi(a).$$

$$\Rightarrow q_a (\varphi(a) - \varphi(a-1)) = p_a (\varphi(a+1) - \varphi(a)) \Rightarrow \varphi(a+1) - \varphi(a) = \frac{q_a}{p_a} (\varphi(a) - \varphi(a-1)).$$

$$\text{For } x \in [a, b], \quad \varphi(x) = q_a + p_x \mathbb{P}_x(T_a < T_b) + r_x \varphi(x+1) \quad \text{for } a \leq x \leq b-1.$$

$$\Rightarrow \varphi(b) - [(p_a + q_a) \varphi(a) - p_a] / p_a.$$

We now define a function φ such that $\varphi(X_{n+1})$ is a martingale,

~~Let $\varphi(a) = 0$ and $\varphi(b) = 1$~~ For the martingale property to hold, we need

$$\varphi(x) = \mathbb{E}_x \varphi(X_{n+1} | X_n = x) = p_x \varphi(x+1) + r_x \varphi(x) + q_x \varphi(x-1).$$

$$\Rightarrow p_x (\varphi(x+1) - \varphi(x)) = \frac{q_x}{p_x} (\varphi(x) - \varphi(x-1))$$

$$\Rightarrow \varphi(k+1) - \varphi(k) = \prod_{j=1}^k \frac{q_j}{p_j} \quad \text{for } k \geq 1. \Rightarrow \varphi(n) = \sum_{k=0}^n \prod_{j=1}^k \frac{q_j}{p_j} \quad \text{for } n \geq 1.$$

Let $T_a = \inf \{n \geq 1 : X_n = a\}$. For any $a < x < b$, we let $T = T_a \wedge T_b$.

Then, $\varphi(X_{n+1})$ is a martingale. Thus, we have

$$\varphi(x) = \mathbb{E}_x \varphi(X_{n+1}) \xrightarrow{n \rightarrow \infty} \varphi(x) = \mathbb{E}_x \varphi(X_T) = (\varphi(a) \mathbb{P}_x(T_a < T_b)$$

$$\Rightarrow \varphi(x) = \varphi(a) \mathbb{P}_x(T_a < T_b) + \varphi(b) [1 - \mathbb{P}_x(T_a < T_b)] + (\varphi(b) \mathbb{P}_x(T_b < T_a))$$

$$\Rightarrow \mathbb{P}_x(T_a < T_b) = \frac{\varphi(b) - \varphi(a)}{\varphi(b) - \varphi(a)}. \quad \text{In particular, let } a=0 \text{ and } b=M, \text{ we get}$$

$$\mathbb{P}_x(T_a < T_M) = \frac{\varphi(M) - \varphi(a)}{\varphi(M) - \varphi(0)} = \frac{\varphi(x)}{\varphi(M)}$$

Letting $M \rightarrow \infty$, we have $\mathbb{P}_x(T_a < \infty) \rightarrow \mathbb{P}_x(T_a < \infty) = \lim_{M \rightarrow \infty} \frac{\varphi(x)}{\varphi(M)}$

Thus, we get that 0 is recurrent if and only if $\varphi(M) \rightarrow \infty$ as $M \rightarrow \infty$, i.e.,

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$$P_{10} = \sum_{m=0}^{\infty} \prod_{j=1}^m \frac{q_j}{p_j} = \infty. \quad \text{If } P_{10} < \infty, \text{ then } P_{10}(T_0 = \infty) = \frac{P_{10}}{P_{100}}.$$

If 0 is recurrent/transient, then all states are recurrent/transient since the chain is irreducible.

Countable space S

3. Stationary Measure: Given a transition probability p , a measure μ is said to be stationary if $\mu p = \mu$, i.e., $\sum_x \mu(x) p(x, y) = \mu(y)$. (*)

[$P_n(x, A) = P_{x0}(X_n \in A)$, $\mu p^n(A) = \int \mu(dx) p^n(x, A)$, regard as a "matrix multiplication".]

Given (*), we have $P_{\mu}(X_0 = y) = \sum_x \mu(x) p(x, y) = \mu(y)$. Using the Markov property and induction in n , we get $P_{\mu}(X_n = y) = \mu(y) \forall n \geq 1$.

If μ is a stationary distribution, we say μ is a stationary distribution. It represents a possible equilibrium for the chain.

There is a stronger notion of stationary measure: μ is said to be a reversible measure if it satisfies the detailed balance condition if $\mu(x) p(x, y) = \mu(y) p(y, x)$.

Summing over x , we get $\sum_x \mu(x) p(x, y) = \mu(y)$, i.e., μ is stationary.

"The amount of mass that moves from x to y in 1 jump is the same as the amount that moves from y to x ."

Example: Simple RW on graph $G = (V, E)$.



We claim that $\mu(x) = \deg(x)$ defines a reversible measure: for $x \sim y$, $\mu(x)p(x, y) = \deg(x) \frac{1}{\deg(y)} = 1$

Given a finite graph, we have $\sum_x \mu(x) = \sum_x \deg(x) = 2|E| < \infty$. $\boxed{= \mu(y)p(y, x) = \deg(y) \frac{1}{\deg(x)}}$

Then, we get a stationary distribution by dividing $\mu(x) = \frac{\deg(x)}{2|E|}$.

In particular, for the SRW on \mathbb{Z}^d , $\mu(x) = 1$ is a stationary measure.

$\boxed{\deg(x) = \sum_y c(x, y)}$ is a stationary measure.

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Asymmetric RW

Example: ~~(transient vs. recurrent)~~ $p_{ii, i+1} = p$, $p_{ii, i-1} = q = 1-p$, $p_{100} = r$, $p_{NN, N-1} = s$.

Let μ be a stationary measure, then it satisfies the equations

$$\text{(*) } \sum_x \mu(x) p(x, y) = p \mu(y-1) + q \mu(y+1) = \mu(y) \Rightarrow p \mu(y+1) - \mu(y) = \frac{p}{q} (\mu(y) - \mu(y-1)).$$

$$\begin{aligned} y=0: \quad & \mu(0) + p\mu(1) = \mu(0) \Rightarrow \mu(1) = \mu(0). \\ \text{(*) } y \geq 1: \quad & \mu(y) = \frac{\mu(y-1)}{1-p} + \frac{p}{q} (\mu(y) - \mu(y-1)) + \dots + \left(\frac{p}{q}\right)^{y-1} (\mu(1) - \mu(0)) \end{aligned}$$

$$= \mu(0) + \frac{1 - \left(\frac{p}{q}\right)^y}{1 - \frac{p}{q}} (\mu(1) - \mu(0)).$$

let $\mu(0) = 1$, $\mu(1) = \frac{p}{q}$, we get $\mu(y) = \left(\frac{p}{q}\right)^y$, $y \geq 0$. Similarly, it holds also for $y < 0$.

Example (Ehrenfest chain) $S = \{0, 1, \dots, r\}$, $p(k, k+1) = \frac{r-k}{r}$, $p(k, k-1) = \frac{k}{r}$.

We claim that $\mu(k) = 2^{-k} \binom{r}{k}$ is a stationary distribution. [μ corresponds to tossing r coins to determine which urn each ball is to be placed in, and the transitions of the chain correspond to picking a coin at random and turning it over.] We can check directly:

$$\mu(k+1) p(k+1, k) = 2^{-k-1} \binom{r}{k+1} \frac{r-k}{r} = 2^{-k} \binom{r-1}{k}$$

$$\mu(k) p(k, k-1) = 2^{-k} \binom{r}{k} \frac{r-k}{r} = 2^{-k-1} \binom{r-1}{k-1}. \quad \boxed{\mu \text{ is reversible.}}$$

Example (Birth-Death chain) We claim that $\mu(x) = \prod_{k=1}^x \frac{p_{k-1}}{q_k}$ is reversible.

$$\begin{aligned} \text{In fact, } \mu(x) p(x, x+1) &= \prod_{k=1}^x \frac{p_{k-1}}{q_k} \cdot p_x = \prod_{k=0}^{x-1} \frac{p_k}{q_{k+1}} = q_{x+1} \prod_{k=0}^x \frac{p_k}{q_{k+1}} = q_{x+1} \prod_{k=1}^{x+1} \frac{p_{k-1}}{q_k} \\ &= \mu(x+1) p(x+1, x). \end{aligned}$$

$\boxed{\text{MC that have a reversible measure is called a reversible MC}}$

Theorem (Time reversal for reversible Markov chains) Let μ be a stationary measure and suppose X_0 has distribution μ . Then, $Y_m = X_{n-m}$, $0 \leq m \leq n$, is a MC with initial measure μ and transition probability: $q(x, y) = \mu(y) p(y, x) / \mu(x)$. q is called "dual transition probability". $q = p$ if μ is reversible.

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$$\text{Pf: } \mathbb{P}_\mu(Y_{m+1}=y \mid Y_m=x, Y_{m-1}=x_{m-1}, \dots, Y_0=x_0)$$

$$= \mathbb{P}_\mu(Y_{m+1}=y, Y_m=y^*, Y_{m-1}=x_{m-1}, \dots, Y_0=x_0)$$

$$\mathbb{P}_\mu(Y_m=x, Y_{m-1}=x_{m-1}, \dots, Y_0=x_0)$$

$$\stackrel{?}{=} \mathbb{P}_\mu(X_n=x_0, \dots, X_{n-m+1}=x_{m-1}, X_{n-m}=x^*, X_{n-m-1}=y)$$

$$\mathbb{P}_\mu(X_n=x_0, \dots, X_{n-m+1}=x_{m-1}, X_{n-m}=x)$$

$$= \mathbb{P}_\mu(X_n=x_0, \dots, X_{n-m+1}=x_{m-1}, X_{n-m}=x \mid X_{n-m-1}=y) \mathbb{P}_\mu(X_{n-m-1}=y)$$

$$\mathbb{P}_\mu(X_n=x_0, \dots, X_{n-m}=x)$$

$$= \frac{\mu(y) p(y, x) p(x, x_{m-1}) \dots p(x_1, x_0)}{\mu(x) p(x, x_{m-1}) \dots p(x_1, x_0)} = \frac{\mu(y) p(y, x)}{\mu(x)} \quad \square$$

We now give a necessary and sufficient condition for a MC to be reversible.

Thm (Kolmogorov's cycle condition) Suppose p is irreducible. A necessary and sufficient condition for the existence of a reversible measure is: (i) $p(x, y) > 0$ implies $p(y, x) > 0$; (ii) for any loop $x_0, x_1, x_2, \dots, x_n=x_0$ with $\prod_{i \in S(n)} p(x_i, x_{i+1}) > 0$, $\prod_{i=1}^n \frac{p(x_{i-1}, x_i)}{p(x_i, x_{i+1})} = 1$.

Pf: $\boxed{\text{if}}$ $\forall x, y \in S$, $\exists n \in \mathbb{N}$ s.t. $p^n(x, y) > 0$, $p^n(y, x) > 0$.

Irreducibility implies that any stationary measure has $\mu(y) > 0 \forall y \in S$. Otherwise,

$$0 = \mu(y) = \sum_x \mu(x) p(x, y) \Rightarrow \mu(x) = 0 \forall x \in S \text{ s.t. } p(x, y) > 0 \Rightarrow \mu(x) = 0 \forall x \in S \text{ s.t. } p^2(x, y) > 0$$

$$\Rightarrow \dots \Rightarrow \mu(x) = 0 \forall x \in S. \text{ Then, the condition } \mu(x) p(x, y) = \mu(y) p(y, x)$$

implies (i) holds. For (ii), we notice that

$$\prod_{i=1}^n \frac{p(x_{i-1}, x_i)}{p(x_i, x_{i+1})} = \prod_{i=1}^n \frac{\mu(x_i)}{\mu(x_{i+1})} = 1.$$

For the sufficiency, fix $a \in S$, set $\mu(a)=1$. If $x_0=a, x_1, x_2, \dots, x_n=x$ is a sequence with $\prod_{i=1}^n p(x_{i-1}, x_i) > 0$ (such a sequence exists), we let

$$\mu(x) = \frac{\prod_{i=1}^n p(x_{i-1}, x_i)}{\prod_{i=1}^n p(x_i, x_{i+1})}. \quad \text{The cycle condition guarantees that this def. does not depend on the path.}$$

To check the reversibility condition, we add note that if $p(x, y) > 0$ ($\text{so } p(y, x) > 0$), then adding $x_{n+1}=y$ to the end of a path to x , we have

$$\mu(y) = \mu(x) \frac{p(x, y)}{p(y, x)} \Rightarrow \mu(y) p(y, x) = \mu(x) p(x, y). \quad \square$$

Theorem: If S is finite, then any MC has a stationary ~~measure~~ distribution. ~~If~~ In addition, if p is irreducible, the stationary distribution π is unique and satisfies $\pi(x) > 0$ let $|S|=n$.

Pf. An elementary proof: Note the rows of p add to 1, so $p - I$ has rank $\leq n-1$.

So there is a vector v so that $vp=v$. Let $q = (I+p)/2$ be the transition matrix

Let $q = (I+p)/2$ be the transition matrix of the lazy chain. By the Perron-Frobenius theorem, we can choose the signs of v such that

We claim that ~~when~~ ~~if~~ p is irreducible, then $\boxed{q^{n-1}(x, y) > 0}$ all entries of v are non-negative.

$\forall x, y \in S$. In fact, $\forall x \neq y$, there exists a minimum k s.t. $p^k(x, y) > 0$. Since k is

minimum, the shortest path does not visit any state more than once, so $k \leq n-1$. It

follows that $r(x, y) := q^{n-1}(x, y) > 0$. Then, we normalize v to get a stationary distribution π

with $\sum_x \pi(x) = 1$. From $\pi(y) = \sum_x \pi(x) r(x, y)$, we see that $\pi(y) > 0 \forall y \in S$. Now, suppose

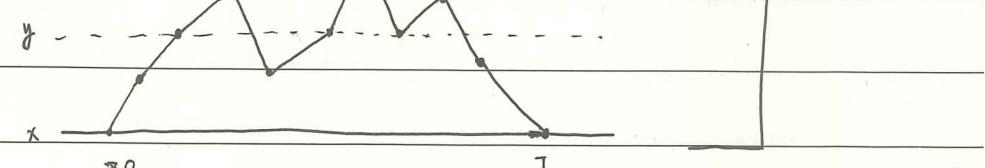
π' is another stationary distribution. We also have $\pi'(y) = \sum_x \pi(x) r(x, y)$ and $\pi'(y) > 0 \forall y \in S$.

$$\text{let } y_0 \in S \text{ s.t. } \frac{\pi'(y_0)}{\pi(y_0)} = \min_{x \in S} \frac{\pi'(x)}{\pi(x)}. \text{ Since } \sum_x \pi(x) = \sum_x \pi'(x) = 1, \text{ we have } \frac{\pi'(y_0)}{\pi(y_0)} = a \leq 1 \Rightarrow \frac{\sum_x \pi(x) r(x, y_0)}{\sum_x \pi'(x) r(x, y_0)} = a \leq 1.$$

Theorem (Existence of stationary measure) Let x be a recurrent state. Let $T = \inf\{n \geq 1 : X_n=x\}$. Then, μ_x defines a stationary measure if we let

$$\mu_x(y) = \mathbb{E}_x \left(\sum_{n=0}^{T-1} \mathbb{1}_{\{X_n=y\}} \right) = \sum_{n=0}^{\infty} \mathbb{P}_x(X_n=y, T>n)$$

Pf:



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$\mu_{x,y}$ is the expected # of visits ^{to} y in $\{0, 1, \dots, T-1\}$, and $\mu_x p(y)$
 $= \sum_y \mu_{x,y} p(y, z)$ is the expected # of visits to y in $\{1, \dots, T\}$, which is $= \mu_{x,y}$
 since $X_T = X_0 = x$. To give a rigorous proof:

$$\sum_y \mu_{x,y} p(y, z) = \sum_{y=0}^{\infty} \mathbb{P}_x(X_n=y, T>n) p(y, z) = \sum_{n=0}^{\infty} \sum_y \mathbb{P}_x(X_n=y, T>n) p(y, z).$$

① If $z \neq x$, then

$$\begin{aligned} \sum_y \mathbb{P}_x(X_n=y, T>n) p(y, z) &= \sum_y \mathbb{P}_x(X_n=y, X_{n+1}=z, T>n) \\ &= \mathbb{P}_x(X_{n+1}=z, T>n+1) \\ &= \mu_x(z). \end{aligned}$$

$$\text{Then, } \sum_y \mu_{x,y} p(y, z) = \sum_{n=0}^{\infty} \mathbb{P}_x(X_{n+1}=z, T>n+1) = \sum_{n=1}^{\infty} \mathbb{P}_x(X_n=z, T>n) = \sum_{n=0}^{\infty} \mathbb{P}_x(X_n=z, T>n)$$

② If $z=x$, then

$$\sum_y \mathbb{P}_x(X_n=y, T>n) p(y, x) = \sum_y \mathbb{P}_x(X_n=y, X_{n+1}=x, T>n) = \mathbb{P}(X_{n+1}=x, T>n) = \mathbb{P}(T=n+1).$$

$$\text{Thus, } \sum_{n=0}^{\infty} \sum_y \mathbb{P}_x(X_n=y, T>n) p(y, x) = \sum_{n=0}^{\infty} \mathbb{P}_x(T=n+1) = 1 = \mu_x(x) \text{ since } \mathbb{P}_x(T=0)=0 \text{ by def.}$$

$$[\mu_x(x) = \mathbb{E}_x \sum_{n=0}^{T-1} \mathbb{I}_{\{X_n=x\}}] = 1 \quad \square$$

Rmk: If x is transient, we have $\mu_x p(z) = \mu_x(z) \forall z \neq x$. For $z=x$, we have

$$\sum_{n=0}^{\infty} \mathbb{P}_x(T=n+1) < 1, \text{ so } \mu_x p(x) < \mu_x(x).$$

Rmk: To show that μ_x defines a measure, we still need to show that $\mu_x(y) < \infty$ for all y . Observe that $\mu_x p = \mu_x$ implies $\mu_x p^n = \mu_x \forall n \geq 1$, and $\mu_x(x) = 1$.

Thus $\sum_y \mu_x(y) p^n(y, x) = \mu_x(x) = 1 \Rightarrow$ if $p^n(y, x) > 0$, then $\mu_x(y) < \infty$. This is true

for all n , so we see that $p_{xy} > 0 \Rightarrow \mu_x(y) < \infty$. When x is recurrent,

$p_{xy} > 0$ implies $p_{yx} > 0$, and it follows that $\mu_x(y) < \infty$. On the other hand, if $p_{xy} = 0$, then $\mu_x(y) = 0$ by def.

The above theorem allows us to construct a stationary measure for each closed set

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of recurrent states. Conversely:

Thm: If p is irreducible and recurrent, then the stationary measure is unique up to constant multiples.

Pf: Let ν be an arbitrary stationary measure and let $a \in S$,

$$\begin{aligned} \nu(z) &= \sum_y \nu(y) p(y, z) = \nu(a) p(a, z) + \sum_{y \neq a} \nu(y) p(y, z) \\ &= \nu(a) p(a, z) + \sum_{y \neq a} \nu(a) p(a, y) p(y, z) + \sum_{x \neq a} \sum_{y \neq a} \nu(x) p(x, y) p(y, z) \\ &= \nu(a) \mathbb{P}_a(X_1=z) + \nu(a) \mathbb{P}_a(X_1 \neq a, X_2=z) + \cancel{\nu(a) \mathbb{P}_a(X_1 \neq a, X_2 \neq a)} \mathbb{P}_a(X_0 \neq a, X_1 \neq a, X_2=z) \end{aligned}$$

Continuing this expansion, we get

$$(*) \quad \nu(z) = \nu(a) \sum_{m=1}^n \mathbb{P}_a(X_{k+m-1}, X_m=z) + \mathbb{P}_a(X_j \neq a, 0 \leq j \leq n-1, X_n=z) \quad \downarrow$$

Let $n \rightarrow \infty$, $\nu(z) \geq \nu(a) \mu_a(z)$, where $\mu_a(z) = \sum_{n=0}^{\infty} \mathbb{P}_a(X_n=z, T>n)$. This converges to 0 if ν is a finite measure using a is

recurrent. But ν may be a ∞ measure.

$$\nu(a) = \sum_x \nu(x) p^n(x, a) \geq \nu(a) \sum_x \mu_a(x) p^n(x, a) = \nu(a) \mu_a(a) = \nu(a).$$

Thus, \geq must be $=$, and we ~~must~~ must have $\nu(x) = \nu(a) \mu_a(x)$

whenever $p^n(x, a) > 0$. Since p is irreducible, it follows that $\nu(x) = \nu(a) \mu_a(x) \forall x \in S$. \square

We now turn attention to stationary distributions. Stationary measures may exist for transient chains (e.g., SRW in $d \geq 3$). But:

Theorem: If there is a stationary distribution π^* , then all states y with $\pi(y) > 0$ are recurrent.

Pf: Using $\pi p^n = \pi$, we get $\sum_x \pi(x) \sum_{n=1}^{\infty} p^n(x, y) = \sum_{n=1}^{\infty} \pi(y) @=\infty$, when $\pi(y) > 0$.

Thus, On the other hand, we have seen that $\sum_{n=1}^{\infty} p^n(x, y) = \frac{p_{xy}}{1-p_{yy}}$. Thus,

$$\infty = \sum_x \pi(x) \frac{p_{xy}}{1-p_{yy}} \leq \frac{1}{1-p_{yy}} \Rightarrow p_{yy} = 1. \quad \square$$

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Theorem: If p is irreducible and has stationary distribution π , then $\pi(x) = \frac{1}{E_x T_x}$.

Pf: Irreducibility implies that π is equal to μ_x up constant multiples.

Since $\mu_x(x) = 1 > 0$, we have $\pi(x) > 0$. Then, by the previous theorem, all states

are recurrent. Irreducibility implies that $\pi(x) > 0$ (otherwise, $\sum_y \pi(y) p^n(y, x) = \pi(x) = 0$

$\Rightarrow \pi(y) = 0 \forall y \in S$). Then, the previous theorem implies that all states are recurrent.

Now, μ_x defines a stationary measure with $\mu_x(x) = 1$ and

$$\mu_x(y) = \sum_{n=0}^{\infty} P_x(T_x=n, T_x=y)$$

Using Fubini's thm, we get $\sum_y \mu_x(y) = \sum_{n=0}^{\infty} P_x(T_x=n) = E_x T_x$. Since the stationary measure is unique up to constant multiples, we must have

$$\pi(x) = \frac{\mu_x(x)}{\sum_y \mu_x(y)} = \frac{1}{E_x T_x} \quad \square$$

Rmk: Another consequence is the following identity: $\sum_x \frac{1}{E_x T_x} p(x, y) = \frac{1}{E_y T_y}$.

A state is said to be positive recurrent if $E_x T_x < \infty$. A recurrent state with $E_x T_x = \infty$ is said to be null recurrent.

Note: positive recurrence implies recurrence: if $E_x T_x < \infty$, then $P_x(T_x=\infty) = 0$.

Theorem: If p is irreducible, then TFAE:

(i) Some $x \in S$ is positive recurrent.

(ii) There is a stationary distribution.

(iii) All states are positive recurrent.

In other words, being positive recurrent is a class property. If it holds for one state in an irreducible set, then it is true for all.

Pf: (i) \Rightarrow (ii). If x is positive recurrent, then

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$\pi(y) = \frac{\mu_x(y)}{E_x T_x}$ defines a stationary distribution.

(ii) \Rightarrow (iii) Irreducibility tells that $\pi(y) > 0 \forall y \in S$. Then, by previous thm,

$E_y T_y = \frac{1}{\pi(y)} < \infty$. Hence, $\forall y \in S$ is positive recurrent.

(iii) \Rightarrow (i) Trivial. \square

Example (Birth-Death chain) We have seen that $\mu(x) = \prod_{k=1}^x \frac{p_{k-1}}{q_k}$ is a reversible measure. Hence, it has a stationary distribution if and only if

$\sum_{x=1}^{\infty} \prod_{k=1}^{x-1} \frac{p_{k-1}}{q_k} < \infty$. On the other hand, we have seen that the chain is recurrent

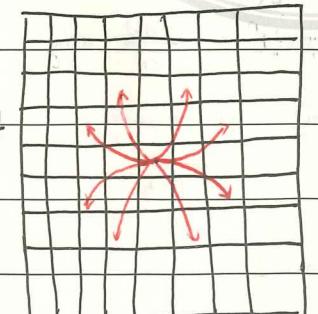
if and only if $\sum_{x=0}^{\infty} \prod_{j=1}^x \frac{q_j}{p_j} = \infty$.

$$b(x) = \frac{p_0}{p_x} \frac{1}{a(x)} \geq \frac{p_0}{a(x)}$$

$G = (V, E)$

Example: We have seen that for the SRW on a connected graph, there is a reversible measure $\mu(x) = \deg(x)$. Moreover, the SRW is irreducible. When the graph is finite, there is a stationary distribution $\pi(x) = \frac{\deg(x)}{2|E|}$. Hence, $\forall x \in V$, we have $E_x T_x = \frac{2|E|}{\deg(x)}$.

Example: (Random walk of a knight on a chess board) A chess board is an 8×8 grid of squares.



We regard it as a simple RW on a graph with 64 vertices.

A knight randomly moves to another grid starting from the lower right corner.

On average, how many does it take the knight to return to this corner?

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	2	3	4	4	
Total degree = 336	3	4	6	6	
	4	6	8	8	
	4	6	8	8	

The degree of the lower right corner x is $\deg(x) = 2$. Hence, $\pi(x) = \frac{2}{336} = 1/168 \Rightarrow E_x T_x = \frac{1}{\pi(x)} = 168$. \square

4. Recurrence or Transience of SRW on \mathbb{Z}^d

Theorem: Simple random walk is recurrent in $d \leq 2$ and transient in $d \geq 3$.

$$\text{SRW: } \mathbb{P}(X_i = e_j) = \mathbb{P}(X_i = -e_j) = \frac{1}{2d}, \quad S_n = X_1 + X_2 + \dots + X_n.$$

Let $T_0=0$ and $T_n = \inf \{m > T_{n-1} : S_m = 0\}$ be the time of the n -th return to 0.

By the Strong MP, we have $\mathbb{P}(T_n < \infty) = \mathbb{P}(T_1 < \infty)^n$.

[Theorem: TFAE: (i) $\mathbb{P}(T_1 < \infty) = 1$. (ii) $\mathbb{P}(S_m = 0 \text{ i.o.}) = 1$. (iii) $\sum_{m=0}^{\infty} \mathbb{P}(S_m = 0) = \infty$]

"A drunk man will eventually find his way home, but a drunk bird will get lost forever."

In $d=1$, we have the classical result.

$$\mathbb{P}(S_{2n}=0) = \frac{1+o(1)}{\sqrt{\pi n}} \text{ as } n \rightarrow \infty.$$

In fact, $\mathbb{P}(S_{2n}=0) = \binom{2n}{n} \left(\frac{1}{2}\right)^{2n} = \frac{(2n)!}{n! n!} \left(\frac{1}{2}\right)^n = \frac{1+o(1)}{\sqrt{\pi n}}$ using Stirling's formula:

$$n! \sim n^n e^{-n} \sqrt{2\pi n} \text{ as } n \rightarrow \infty.$$

Hence, we have $\sum_{m=0}^{\infty} \mathbb{P}(S_m = 0) = \infty$, showing that 0 is recurrent. By irreducibility, all states are recurrent for the SRW on \mathbb{Z} .

Let $p_d(n) = \mathbb{P}(S_n = 0)$ in \mathbb{Z}^d

Next, we show 2D SRW is recurrent. In order for $S_n = 0$, we must have that for some $0 \leq m \leq n$, there are m up steps, m down steps, $n-m$ to the left, and $n-m$

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to the right. Hence,

$$\mathbb{P}(S_n = 0) = 4^{-2n} \sum_{m=0}^n \frac{(2n)!}{m! m! (n-m)! (n-m)!} = 4^{-2n} \binom{2n}{n} \sum_{m=0}^n \binom{n}{m} \binom{n}{n-m}$$

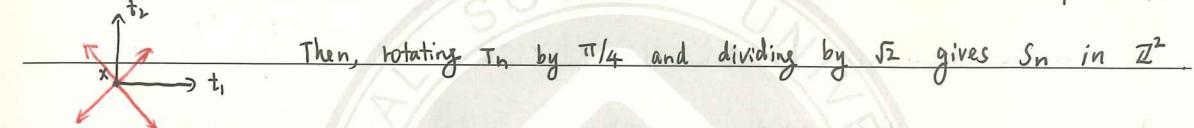
$$= 4^{-2n} \binom{2n}{n}^2 = \left[2^{2n} \binom{2n}{n} \right]^2 = [p_1(2n)]^2 = \frac{1+o(1)}{\pi n}.$$

Consider choosing n balls from a box with n black balls and n white balls

$$\Rightarrow \sum_{m=0}^{\infty} \mathbb{P}(S_m = 0) = \infty, \text{ so the SRW is recurrent.}$$

There is a direct proof of $p_2(2n) = [p_1(2n)]^2$: Let T_n^1 and T_n^2 be two ind.

1D SRWs. Then, $T_n = (T_n^1, T_n^2)$ jumps from x to $x + (\pm 1, \pm 1)$ with prob. $\frac{1}{4}$.



Then, rotating T_n by $\pi/4$ and dividing by $\sqrt{2}$ gives S_n in \mathbb{Z}^2 .

Finally, we consider the SRW in 3D. Intuitively, $p_3(2n)$ is of order $O(n^{-3/2})$,

which is summable, so the SRW is transient. We can calculate that

$$p_3(2n) = 6^{-2n} \sum_{j,k} \frac{(2n)!}{[j! k! (n-j-k)!]^2} = 2^{-2n} \binom{2n}{n} \sum_{j,k} \left[3^{-n} \frac{n!}{j! k! (n-j-k)!} \right]^2$$

$$\leq 2^{-2n} p_1(2n) \max_{j,k} 3^{-n} \frac{n!}{j! k! (n-j-k)!}$$

where we used $\sum_{j,k} 3^{-n} \frac{n!}{j! k! (n-j-k)!} = 1$. Next, we claim that

$$\max_{j,k} 3^{-n} \frac{n!}{j! k! (n-j-k)!} \leq \frac{C}{n} \text{ for a constant } C > 1.$$

Note the maximum is achieved when all three numbers are as close as possible

to $n/3$ (the difference between them are at most 1). In fact, suppose one of

$j, k, n-j-k$ is $< \lfloor n/3 \rfloor$. Then, there is another one that is $> \lfloor n/3 \rfloor$.

WLOG, suppose $j < \lfloor n/3 \rfloor$ and $k > \lfloor n/3 \rfloor$, then $(j+1)! (k-1)! \leq j! k!$, and so

$$\text{that } 3^{-n} \frac{n!}{j! k! (n-j-k)!} \leq 3^{-n} \frac{n!}{(j+1)! (k-1)! (n-j-k)!}.$$

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Then, using Stirlings formula, we get

$$\begin{aligned} \frac{3^{-n}}{\frac{n!}{j! k! (n-j-k)!}} &= [1+o(1)] \frac{\frac{3^{-n} n^n}{j^j k^k (n-j-k)^{n-j-k}} \sqrt{\frac{n}{j! k! (n-j-k)!}}}{\frac{1}{n!}} \frac{1}{\frac{1}{n!}} \\ &= [1+o(1)] \left[1 + \frac{3(j-k)n}{n} \right]^j \left[1 + \frac{3(k-h)}{n} \right]^k \left[1 + \frac{2n-3j-3k}{n} \right]^{n-j-k} \frac{3^{\frac{3}{2}}}{n!} \\ &\leq C/n. \end{aligned}$$

Hence, we get $p_3(2n) \leq \frac{C}{n^{3/2}}$, which is summable.

This shows that the 3D SRW is transient.

In dimensions $d \geq 4$: $S_n = (S_n^1, S_n^2, S_n^3, S_n^4, \dots, S_n^d)$. Let $T_n = (S_n^1, S_n^2, S_n^3)$.

Then, we have $\text{IP}(S_n=0) \leq \text{IP}(T_n=0) \Rightarrow p_{d-3}(n) \leq p_3(n) \Rightarrow p_d(n)$ is also summable \Rightarrow Then SRW on \mathbb{Z}^d , $d \geq 4$, is transient.

Note that the SRW on \mathbb{Z}^d has a stationary measure (even if SRW on \mathbb{Z}^d , $d \geq 3$, is recurrent): the uniform measure $\pi(x) = \frac{1}{d}$.

In $d=1, 2$, we know the SRW has a unique stationary measure. If there exists a stationary distribution π , then by uniqueness, $\pi(x) = a > 0$ for all $x \in \mathbb{Z}$.

Moreover, $1 = \sum_x \pi(x) = \sum_x a = \infty$, a contradiction. Hence, the SRW on \mathbb{Z}^d admits no stationary distribution. As a consequence, 1D and 2D SRW are null recurrent.

5. Asymptotic behavior of Markov chains, Periodicity

We now study the asymptotic behavior of the Markov chain, i.e., $p^n(x, y) \forall x, y \in S$, as $n \rightarrow \infty$. Note $p^n(x, y) \rightarrow 0$ as $n \rightarrow \infty$ if y is transient since $\sum_y p^n(x, y) < \infty$.

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[For recurrent states, our goal is to show that $p^n(x, y) \rightarrow \pi(y)$ asymptotically under certain conditions if there exists a stationary distribution]

Let $N_n(y) = \sum_{m=1}^n \mathbb{1}(X_m=y)$: # of visits to y up to time n .

Thm: Suppose y is recurrent. For any $x \in S$, as $n \rightarrow \infty$, we have

$$\frac{N_n(y)}{n} \rightarrow \frac{1}{\text{E}_y T_y} \mathbb{1}(T_y < \infty) \quad \text{IP}_x - \text{a.s.}$$

When $\text{E}_y T_y = \infty$, we let $\frac{1}{\infty} = 0$.

Pf: First, suppose $x=y$ (i.e., we start at y). Let $R_k = \min\{n \geq 1 : N_n(y)=k\}$ — the time of the k -th return to y . Let $t_k = R_k - R_{k-1}$, where $R_0=0$.

When $X(0)=y$, t_1, t_2, t_3, \dots are i.i.d. by the strong MP and the strong LLN implies

$$R(k)/k = \frac{1}{k} \sum_{i=1}^k t_i \rightarrow \text{E}_y T_y \quad \text{IP}_y - \text{a.s.}$$

Note that $R(N_n(y)) \leq n \leq R(N_n(y)+1)$, we have

$$\frac{R(N_n(y))}{N_n(y)} \leq \frac{n}{N_n(y)} \leq \frac{R(N_n(y)+1)}{N_n(y)+1} \cdot \frac{N_n(y)+1}{N_n(y)}$$

Let $n \rightarrow \infty$. We have $N_n(y) \rightarrow \infty$ a.s. since y is recurrent. Then, we get

$$\frac{n}{N_n(y)} \rightarrow \frac{1}{\text{E}_y T_y} \quad \text{IP}_y - \text{a.s.}$$

Next, we extend it to $x \neq y$. Observe that if $T_y = \infty$, then $N_n(y) = 0$ for all n and hence $N_n(y)/n \rightarrow 0$ on $\{T_y = \infty\}$.

Conditional

on the events $\{T_y < \infty\}$, the strong MP implies that t_1, t_2, \dots are i.i.d. so

$$\text{P}_x(t_{k+1} = 1 | T_y = \infty) = \text{P}_y(T_y = \infty), \quad R(k)/k = t_1/k + (t_2 + t_3 + \dots + t_k)/k \rightarrow 0 + \text{E}_y T_y \quad \text{IP}_x - \text{a.s.}$$

Repeating the proof for the case $x=y$ shows that $N_n(y)/n \rightarrow 1/\text{E}_y T_y$ $\text{IP}_x - \text{a.s.}$ on $\{T_y < \infty\}$. \square

Rmk: This result tells that if y is positive / null recurrent, then the asymptotic fraction of time spent at x is positive and in the second case it is 0.

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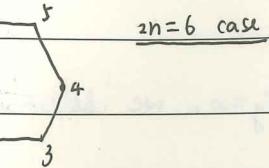
By the bounded convergence theorem, we have $\frac{1}{n} E_x \frac{N_{n,y}}{n} \rightarrow E_x \left(\frac{1}{E_y T_y} \mathbb{1}_{\{T_y < +\infty\}} \right)$, so

$$\frac{1}{n} \sum_{m=1}^n p^m(x, y) \rightarrow p_{xy} / E_y T_y$$

This also holds for transient y , since $E_y T_y = \infty$ and $\sum_{m=1}^{\infty} p^m(x, y) < \infty$. This shows that $p^n(x, y)$ converges in the Cesaro sense. But this is weaker than the convergence $p^n(x, y)$ itself.

Example: Consider the SRW on the $2n$ -cycle:

$$\begin{cases} p^{2k}(0, i) = 0 & \text{if } i \text{ is even} \\ p^{2k+1}(0, i) = 0 & \text{if } i \text{ is odd} \end{cases}$$

 $2n=6$ case

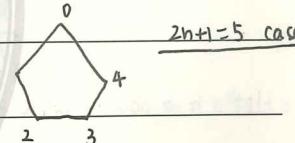
Hence, $p^{2k}(0, 0) + p^{2k}(0, 3) + p^{2k}(0, 5) = 0$, but

$$p^{2k+1}(0, 0) + p^{2k+1}(0, 3) + p^{2k+1}(0, 5) = 1.$$

Hence, $p^n(0, x)$ does not converge for this chain as $n \rightarrow \infty$.

On the other hand, if we consider the SRW on the $(2n+1)$ -cycle, we do not have the above issue.

Starting from 0, we can come back in both even and odd number (by going through the cycle) of steps.

 $2n+1=5$ case

We will show that this "periodicity" is the only issue that can prevent in an irreducible MC the convergence of $p^n(x, y)$.

Def: (Period) Let x be a recurrent state. Let $I_x = \{n \geq 1 : p^n(x, x) > 0\}$. Let d_x be the gcd (greatest common divisor) of I_x . d_x is called the period of x .

Example: For the SRW on the $2n$ -cycle, $I_0 = \{0, 2, 4, \dots\}$. Hence, $d_0 = 2$.

For the SRW on the $(2n+1)$ -cycle, we have $2 \in I_0$ and $2n+1 \in I_0$. Hence, $d_0 = 1$.

Period is a "class property".

Lemma: If $p_{xy} > 0$, then $d_x = d_y$.

Pf: Let $K, L \in \mathbb{N}$ such that $p^K(x, y) > 0$ and $p^L(y, x) > 0$. (x is recurrent, so $p_{yx} > 0$)

Then, $K+L \in I_x \cap I_y$ because $p^{K+L}(y, y) \geq p^L(y, x)p^K(x, y) > 0$, $p^{K+L}(x, x) \geq p^K(x, y)p^L(y, x) > 0$.

Thus, we have $d_y | (K+L)$. Let $m \in \mathbb{N}$ be such that $p^m(x, x) > 0$. Then,

$$p^{K+m}(y, y) \geq p^L(y, x)p^m(x, x)p^K(x, y) > 0, \text{ so } d_y | (K+m) \Rightarrow d_y | m.$$

Thus, $d_y | n \forall n \in I_x \Rightarrow d_y | d_x$. Similarly, $d_x | d_y$. Hence, $d_x = d_y$. \square

If a chain is irreducible and $d_x = 1$, then all states in the chain have period 1 and the chain is called aperiodic. Note if there exists a state x with $p(x, x) > 0$, then the chain is aperiodic.

Hence, give an irreducible and recurrent MC, people sometimes consider the "Lazy chain" with transition probability $q = (I+p)/2$. ~~The Lazy chain is always aperiodic.~~

$\exists m_0$ s.t.

Lemma: If $d_x = 1$, then $p^m(x, x) > 0$ for all $m \geq m_0$.

Pf: We first show that I_x contains two consecutive integers $k, k+1$.

A fact from number theory: if the gcd of a set I_x is 1, then there are

$i_1, i_2, \dots, i_m \in I_x$ and positive/negative integer coefficients $c_1, \dots, c_m \in \mathbb{Z}$ so that

$$c_1 i_1 + c_2 i_2 + \dots + c_m i_m = 1.$$

Let $a_i = c_i^+$ be the positive coefficients and $b_i = c_i^-$ be the absolute values of the negative ones. The above equation holds $a_1 i_1 + \dots + a_m i_m = (b_1 i_1 + \dots + b_m i_m) + 1$, $a_i \in \mathbb{N}, b_i \in \mathbb{N}$.

Then, we let $k = b_1 i_1 + \dots + b_m i_m$.

Given $k, k+1$, we get that $2k, 2k+1 = k+(k+1), 2k+2 = 2(k+1) \in I_x$

$$\Rightarrow 3k = 2k+k, 3k+1, 3k+2, 3k+3 = 3(k+1) \in I_x \Rightarrow \dots \Rightarrow (k-1)k, (k-1)k+1, \dots, (k-1)k+k-1$$

$\in I_x$ for k consecutive integers. For any $m \geq (k-1)k+k-1$, we can find $a \in \{(k-1)k, \dots, (k-1)k+k-1\}$ and $b \in \mathbb{N}$ s.t. $m = a+bk$. Hence, the lemma holds for ~~all~~ $m_0 = (k-1)k$. \square

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- (ii) Let n, m be such that $p^n(y, z), p^m(x, y) > 0$. Since $p^{m+n}(x, z) > 0$, by (i), we have $n+m \equiv j \pmod{d}$. Since $m \equiv i \pmod{d}$, we get $n \equiv j-i \pmod{d}$.
- (iii) $\forall x, y \in S$, $\exists n$ s.t. $p^n(x, y) > 0 \Rightarrow n \equiv 0 \pmod{d}$ by (ii)
 $\Rightarrow (p^d)^k(x, y) > 0$ for some $k \in \mathbb{N}$. This gives irreducibility. \square

A partition S_0, S_1, \dots, S_{d-1} satisfying the above lemma is called a cyclic decomposition of the state space. Up to the labels, this decomposition is unique.

Theorem: (Convergence theorem, periodic case) Suppose p is irreducible, has a stationary distribution π , and all states have period d . Let $x \in S_0$, and let S_0, S_1, \dots, S_{d-1} be the cyclic decomposition of the state space with $x \in S_0$. If $y \in S_r$, then

$$\lim_{m \rightarrow \infty} p^{md+r}(x, y) = \pi(y)d.$$

Pf: If $y \in S_0$, then by previous lemma (iii) and MC convergence theorem, we have that $\lim_{m \rightarrow \infty} p^{md}(x, y)$ exists. To find the limit, we notice that

$$\frac{1}{n} \sum_{m=1}^n p^m(x, y) \rightarrow \pi(y).$$

But, we know $p^m(x, y) = 0$ unless $d \mid m$ by the previous lemma (ii), so we get that $\frac{1}{n} \sum_{m=1}^n p^{md}(x, y) \rightarrow \pi(y) \Rightarrow p^{md}(x, y)$ must converge to $\pi(y)d$.

If $y \in S_r$ with $r \in \{1, 2, \dots, d-1\}$, then

$$p^{md+r}(x, y) = \sum_{z \in S_r} p^r(x, z) p^{md}(z, y).$$

By the previous case, $p^{md}(z, y) \rightarrow \pi(y)d$ as $m \rightarrow \infty$. Moreover, $\sum_z p^{md}(z, y) \leq 1$ and $\sum_z p^r(x, z) = 1$, so by DCT, we have

$$\lim_{m \rightarrow \infty} p^{md+r}(x, y) = \sum_{z \in S_r} p^r(x, z) \pi(y)d = \pi(y)d. \quad \square$$

Chapter 3. Brownian Motion

Brownian motion is a continuous time Gaussian Markov process with stationary independent increments. It is also a continuous time martingale.

1. Definition & Construction of BM

Def: A one-dimensional Brownian motion (BM) is a real-valued process $B_t, t \geq 0$, with the following properties.

(a) (Independent increments) If $t_0 < t_1 < \dots < t_n$, then $B(t_0), B(t_1) - B(t_0), \dots, B(t_n) - B(t_{n-1})$ are independent. (Stationary Gaussian increments)

(b) If $s, t \geq 0$, then $\mathbb{P}(B(s+t) - B(s) \in A) = \int_A \frac{1}{\sqrt{\pi t}} \exp(-\frac{x^2}{2t}) dx$, i.e., $B(s+t) - B(s) \sim N(0, t)$.

(c) $t \mapsto B_t$ is continuous a.s.

Rmk: Discuss why these three properties make sense physically.

Rmk: We say the BM is standard if $B_0 = 0$.

(Translation invariance)

Proposition: $\forall t_0 \geq 0$, $\{B_{t+t_0} - B_{t_0}\}_{t \geq 0}$ is a standard BM independent of B_{t_0} .

Pf: Let $A_1 = \{B_{t_0}\}$ and A_2 be the events of the form $\{B(t_1) - B(t_0), \dots, B(t_n) - B(t_{n-1}) \in A_1\}$.

Then, A_2 is a π -system generating $\sigma(\{B_{t+t_0} - B_{t_0}\}_{t \geq 0})$. By $\pi-\lambda$ theorem, $\{B_{t+t_0} - B_{t_0}\}_{t \geq 0}$ is independent of B_{t_0} . Then, we can show that $B_{t+t_0} - B_{t_0}$ satisfies

the properties (a)-(c) for standard BM. \square

Proposition (Brownian scaling relation): Let B_t be a standard BM. Then, $\forall t \geq 0$, $\{t^{-\frac{1}{2}} B_{st} : s \geq 0\}$ is a standard BM (with time s).

Pf: We check $t^{-\frac{1}{2}} B_{st}$ satisfies the properties (a)-(c) above. \square

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- An equivalent definition of BM with $B_0 = 0$: B_t , $t \geq 0$, is a real-valued process satisfying
- $B(t)$ is a Gaussian process (all its finite dimensional marginals are multivariate normal).
 - $\mathbb{E}B_t = 0$ and $\mathbb{E}(B_s B_t) = s t$
 - $t \mapsto B_t$ is continuous a.s.

Pf: (a) + (b) \Rightarrow (a'): for $t_0 < t_1 < t_2 < \dots < t_n$, the distribution of $(B_{t_0}, B_{t_1}, \dots, B_{t_n})$ is determined by a Gaussian vector of independent entries $(B_{t_0}, B_{t_1} - B_{t_0}, B_{t_2} - B_{t_1}, \dots, B_{t_n} - B_{t_{n-1}})$.

(a) + (b') \Rightarrow (b'): if $s < t$, then $\mathbb{E}(B_s B_t) = \mathbb{E}(B_s (B_s + B_t - B_s)) = \mathbb{E}B_s^2 + \mathbb{E}(B_s) \mathbb{E}(B_t - B_s) = s$.

(a') + (b') \Rightarrow (a) + (b): $B(s+t) - B(s)$ is Gaussian distributed and $\mathbb{E}(B(s+t) - B(s)) = 0$,

$$\mathbb{E}[(B(s+t) - B(s))^2] = \mathbb{E}[B(s+t)^2] - 2\mathbb{E}[B(s+t)B(s)] + \mathbb{E}[B(s)^2] = s+t - 2s + s = t.$$

$$\mathbb{E}[(B(t_0) - B(t_1))^2] = 0, \quad \mathbb{E}[(B(t) - B(s))(B(s) - B(r))] = 0 \quad \forall r < s < t$$

\Rightarrow Independent increments □

Q: Does BM exist?

The idea is straightforward: we first construct the finite dimensional distributions using multivariate normal distribution and then extend it using the Kolmogorov's extension thm. However, Kolmogorov's extension thm only gives a measure on a countable sequence of RVs, while the index set for BM is uncountable. To deal with this issue, we adopt the usual strategy, that is, we first construct a measure on a countable dense subset of \mathbb{R}^+ and then extend it to the whole \mathbb{R}^+ using the continuity requirement on BM.

Step 1: (Construct finite-dimensional measures)

Fix $x \in \mathbb{R}$ and $0 < t_1 < t_2 < \dots < t_n$, we define a measure on \mathbb{R}^n as

$$\mu_{x, t_1, t_2, \dots, t_n}(A_1 \times A_2 \times \dots \times A_n) = \int_{A_1} dx_1 \int_{A_2} dx_2 \dots \int_{A_n} dx_n \prod_{m=1}^n p_{t_m - t_{m-1}}(x_{m-1}, x_m)$$

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where A_i are measurable, $x_0 = x$, $t_0 = t$ and p_t is the heat kernel:

$$p_t(a, b) = (2\pi t)^{-\frac{1}{2}} \exp(-|b-a|^2/2t)$$

We observe that for each fixed x , the family μ is a consistent set of finite dimensional distributions, that is, $\forall j \in \{1, \dots, n\}$,

$$(*) \mu_{x, t_1, \dots, t_j, t_{j+1}, \dots, t_n}(A_1 \times \dots \times A_{j-1} \times A_j \times A_{j+1} \times \dots \times A_n) = \mu_{x, t_1, \dots, t_n}(A_1 \times \dots \times A_{j-1} \times R \times A_{j+1} \times \dots \times A_n)$$

This is trivial when $j=n$, because $p_{t_n - t_{n-1}}$ is a PDF.

The key point is that (*) also holds for $1 \leq j < n$. We need to show that

$$\int_R p_{t_j - t_{j-1}}(x, y) p_{t_{j+1} - t_j}(y, z) dy = p_{t_{j+1} - t_{j-1}}(x, z).$$

The LHS is the PDF for $N(x, t_j - t_{j-1}) \oplus N(0, t_{j+1} - t_j)$, independent sum.

The RHS is the PDF for $N(x, t_{j+1} - t_{j-1})$.

Step 2: (Kolmogorov's extension) Let $\mathbb{Q}_2 = \{m2^{-n} : m, n \geq 0\}$ be the dyadic rationals.

Let $\Omega = \{w : \mathbb{Q}_2 \rightarrow \mathbb{R}\}$ and \mathcal{F}_2 is the σ -field generated by finite dimensional sets:

$$\{w_{t_i} \in A_i, \dots, w_{t_n} \in A_n : t_1, t_2, \dots, t_n \in \mathbb{Q}_2\}.$$

Enumerating the dyadic rationals in \mathbb{Q}_2 , we have constructed a sequence of consistent probability measures indexed by them. Then, applying the Kolmogorov's extension thm, we can construct a probability measure ν_x on $(\Omega_2, \mathcal{F}_2)$ so that $\nu_x\{w : w(t_0) = x\} = 1$ and

$$\nu_x\{w : w(t_i) \in A_i, i=1, \dots, n\} = \mu_{x, t_1, \dots, t_n}(A_1 \times \dots \times A_n).$$

Step 3: (Extension to a process defined on $\mathbb{R}^+ = [0, \infty)$)

Theorem: Let $T < \infty$ and $x \in \mathbb{R}$. Then, ν_x assigns probability 1 to paths $w : \mathbb{Q}_2 \rightarrow \mathbb{R}$ that are uniformly continuous on $\mathbb{Q}_2 \cap [0, T]$.

Before proving this key theorem, we first discuss a key consequence of this result.

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We can move v_x to (C^0, \mathbb{P}) why $C^0 = \{\omega \text{ continuous: } [0, \infty) \rightarrow \mathbb{R}\}$ and \mathbb{P} is the σ -field generated by the coordinate maps $t \mapsto \omega(t)$.

Let $\psi: \Omega_\theta \rightarrow C^0$ be the map that takes a uniformly continuous path/sample in Ω_θ to its unique continuous extension, i.e., let $w \in \Omega_\theta$ be a uniformly continuous function. Then, $\psi(w): \mathbb{R}_+ \rightarrow \mathbb{R}$ is defined as $\psi(w)(x) = \lim_{n \rightarrow \infty} w(x_n)$, where x_n is a sequence of points in \mathbb{Q}_2 such that $x_n \rightarrow x$. We can show that ψ is measurable. Then, we let $\mathbb{P}_\theta = \nu_x \circ \psi^{-1}$ be the prob. measure on C^0 .

Our construction guarantees that $B_t(w) = w_t$ has the right finite-dimensional distributions for $t \in \mathbb{Q}_2$. Then, using continuity of paths and a simple limiting argument, we find that properties (a) and (b) hold for $t \in [0, \infty)$.

Like the MC, we now have one set of RVS $B_t(w) = w_t$, and a family of probability measures \mathbb{P}_θ , $\theta \in \mathbb{R}$, so that under \mathbb{P}_θ , B_t is a BM with $\mathbb{P}_\theta(B_0 = x) = 1$.

Pf of the theorem: By translation invariance and scaling relation, we can take $T=1$ and $s=0$.

Using the definition, we get $\mathbb{E}_0[(B_t - B_s)^\alpha] = \mathbb{E}_0[B_t^\alpha] \leq C n t^{-\alpha/2}$ \forall fixed $n \geq 2$, $n \in \mathbb{N}$.

Then, the following result due to Kolmogorov gives the Hölder continuity of the ~~BM~~ paths of BM.

Theorem + Suppose $\mathbb{E}|X_s - X_t|^\beta \leq K|t-s|^{1+\alpha}$ where $\alpha, \beta > 0$, $t, s \in \mathbb{Q}_2$. If $\alpha < \beta/\alpha$, then with

probability 1, there is a constant $C(\omega)$ so that

$$|X_q - X_r| \leq C|q-r|^\alpha \text{ for all } q, r \in \mathbb{Q}_2 \cap [0, 1].$$

Proof of Thm +: Let $G_n = \{|X(i2^{-n}) - X((i-1)2^{-n})| \leq 2^{-\alpha n} \text{ for all } 0 < i \leq 2^n\}$

By Chebyshev's inequality, $\mathbb{P}(G_n^c) \leq 2^n \cdot (2^{-\alpha n})^\beta \mathbb{E}|X(i2^{-n}) - X((i-1)2^{-n})|^\beta$

$$\leq 2^{-n} \cdot 2^{n\beta} \cdot K 2^{-n(1+\alpha)} = K 2^{n(\beta-\alpha)} = K 2^{-n\lambda},$$

$$\lambda = \alpha - \beta > 0.$$

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Lemma: On $H_N = \bigcap_{n=N}^{\infty} G_n$, we have $|X(q) - X(r)| \leq \frac{2^{-\alpha}}{1-2^{-\alpha}} |q-r|^\alpha \quad \forall q, r \in \mathbb{Q}_2 \cap [0, 1]$ with $|q-r| \leq 2^{-N}$

If $I_i^k = [(i-1)2^{-k}, i2^{-k}]$ and let m be the smallest k for which $|q-r| \leq 2^{-N}$

q, r are in different intervals I_i^k . Since they were in the same interval on level $k-1$,

$q \in I_i^m$ and $r \in I_{i+1}^m$ for some $0 \leq i \leq 2^m$, we can write

$$r = i2^{-m} + 2^{-m+1} + \dots + 2^{-l+1}$$

$$q = i2^{-m} - 2^{-m+1} - \dots - 2^{-g(k)}$$

for some $0 < r(1) < \dots < r(l)$ and $0 < q(1) < \dots < q(k)$. Then, on H_N , we have

$$|X(q) - X(r)| \leq \sum_{h=1}^k (2^{-g(h)})^\alpha \leq \sum_{h=m}^{\infty} (2^{-h})^\alpha = \frac{2^{-\alpha m}}{1-2^{-\alpha}}$$

$$|X(r) - X(i2^{-m})| \leq \frac{2^{-\alpha m}}{1-2^{-\alpha}}.$$

Then, $|X(q) - X(r)| \leq \frac{2^{-\alpha(m+1)}}{1-2^{-\alpha}} \leq |q-r| \geq 2^{-m}$. This concludes the proof. \square

Now, to conclude the proof of Theorem +, we have

$$\mathbb{P}(H_N^c) \leq \sum_{n=N}^{\infty} \mathbb{P}(G_n^c) \leq K \sum_{n=N}^{\infty} 2^{-\alpha n} = \frac{K 2^{-N\lambda}}{1-2^{-\alpha}}$$

Thus, $\sum_{n=1}^{\infty} \mathbb{P}(H_N^c) < +\infty$, so by Borel-Cantelli, we get that $\mathbb{P}(H^c) = 0$ for almost every ω , $\exists \delta(\omega) > 0$

s.t. $|X(q) - X(r)| \leq A|q-r|^\alpha$ for $q, r \in \mathbb{Q}_2$ with $|q-r| < \delta(\omega)$.

It remains to extend this to $q, r \in \mathbb{Q}_2 \cap [0, 1]$. Let $s_0 = q < s_1 < s_2 < \dots < s_n = r$ with

$|s_i - s_{i-1}| < \delta(\omega)$ and $s_i \in \mathbb{Q}_2$. Using triangle inequality, we get that

$$|X(q) - X(r)| \leq \sum_{i=0}^{n-1} |X(s_{i+1}) - X(s_i)| \leq A \sum_{i=0}^{n-1} |s_{i+1} - s_i|^\alpha \leq \frac{A}{\delta(\omega)^{1-\alpha}} \sum_{i=1}^n |s_{i+1} - s_i|^\alpha$$

$$\leq \frac{A}{\delta(\omega)^{1-\alpha}} |q-r| \leq \frac{A}{\delta(\omega)^{1-\alpha}} |q-r|^\alpha.$$

By the scaling relation, $\mathbb{E}|B_t - B_s|^{2m} = C_m |t-s|^m$, $C_m = \mathbb{E}|B_1|^{2m}$.

Thus, the above thm shows that a.s., w is α -Hölder continuous $\forall \alpha < \frac{m-1}{2m}$.

Letting $m \rightarrow +\infty$, we get:

Thm: (Wiener 1923) Brownian paths are Hölder continuous \forall exponent $\alpha < 1/2$.

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Theorem With prob. one, Brownian paths are not Lipschitz continuous (and hence not differentiable) at any point.

Pf: Fix a constant $C > 0$. Let $A_n = \{w: \text{there is an } s \in [0, 1] \text{ so that } |B_t - B_s| \leq C(t-s)$

for when $|t-s| \leq \frac{3}{n}$. For $1 \leq k \leq n-2$, let

$$Y_{k,n} = \max \left\{ \left| B\left(\frac{k+j}{n}\right) - B\left(\frac{k+j-1}{n}\right) \right| : j=0,1,2 \right\}$$

$B_n = \{ \text{at least one } Y_{k,n} \leq 5C/n \}$. By triangle ineq, we have

The worst case is $s=1$. Picking $k=n-2$, we observe that

$$\left| B\left(\frac{n-3}{n}\right) - B\left(\frac{n-2}{n}\right) \right| \leq \left| B\left(\frac{n-3}{n}\right) - B(1) \right| + \left| B(1) - B\left(\frac{n-2}{n}\right) \right| \leq \frac{5C}{n}.$$

Using $A_n \subset B_n$, we get $\mathbb{P}(A_n) \leq \mathbb{P}(B_n) \leq n \mathbb{P}(|B|/n) \leq 5C/n^3$

$$= n \mathbb{P}(|B(1)| \leq 5C/n^3) \leq Cn^{-\frac{1}{2}}.$$

Letting $n \rightarrow \infty$, we get $\mathbb{P}(A_n) \rightarrow 0$. But $\overset{n}{\rightarrow} A_n$ is increasing. Hence, we must have

$\mathbb{P}(A_n) = 0$, which completes the proof. \square

Multi-dimensional BM: With 1D BM we just constructed, we can define d-dimensional BM

with $d \geq 2$. A d-dim BM starting at $\vec{x} \in \mathbb{R}^d$ is defined as $\vec{B}_t = (B_t^1, \dots, B_t^d)$, where

B_t^1, \dots, B_t^d are d independent BMs with $B_0^i = x^i$. In this case, the transition prob

is the heat kernel in \mathbb{R}^d : $p_t(x,y) = (2\pi t)^{-d/2} \exp(-\frac{|x-y|^2}{2t})$, i.e., for $0 < t_1 < \dots < t_n$,

$$\mathbb{P}_x(B_{t_0} \in A_0, \dots, B_{t_n} \in A_n) = \int_{A_0} dx_0 \int_{A_1} dx_1 \cdots \int_{A_n} dx_n \prod_{m=1}^n p_{t_m-t_{m-1}}(\vec{x}_{m-1}, \vec{x}_m),$$

with $\vec{x}_0 = \vec{x}$ and $t_0 = 0$.

2. Markov Property

MP says that " $B(t+s) - B(s)$ is a BM that is independent of what happened

before time s ". To make this statement rigorous, we first need to make sense of

"what happened before time s ". The most natural choice is $\mathcal{F}_s^0 = \sigma\{B_r : r \leq s\}$.

However, it is more convenient to consider a slightly larger σ -field:

$\mathcal{F}_s^+ = \bigcap_{t \geq s} \mathcal{F}_t^0 \supset \mathcal{F}_s^0$. This field is nicer because it is right continuous:

$$\bigcap_{t \geq s} \mathcal{F}_t^+ = \bigcap_{t \geq s} (\bigcap_{u \geq t} \mathcal{F}_u^0) = \bigcap_{u \geq s} \mathcal{F}_u^0 = \mathcal{F}_s^+.$$

Intuitively, \mathcal{F}_s^+ allows for an "infinitesimal peek at the future", i.e., $A \in \mathcal{F}_s^+$ if it is in $\mathcal{F}_{s+\epsilon}^0 \forall \epsilon > 0$. If $f_u > 0 \forall u > 0$, then $\limsup_{t \downarrow s} \frac{B_t - B_s}{t-s}$ is measurable w.r.t.

\mathcal{F}_s^+ but not \mathcal{F}_s^0 . We will see below that \mathcal{F}_s^+ and \mathcal{F}_s^0 are the same up to null sets.

We have defined a family of measures \mathbb{P}_x , $x \in \mathbb{R}^d$, on $(\mathcal{C}(\mathbb{R}^d), \mathcal{F})$ so that under \mathbb{P}_x ,

$B_t(w) = w(t)$ is a BM starting at x . For $s \geq 0$, we define the shift transformation

$\theta_s: \mathcal{C}(\mathbb{R}^d) \rightarrow \mathcal{C}(\mathbb{R}^d)$ by $(\theta_s w)(t) = w(t+s)$ for $t \geq 0$.

Then, we define the MP similar to what we have done for MCs.

Theorem (MP for BM) If $s \geq 0$ and Y is bounded and \mathcal{F}_s -measurable, then $\forall x \in \mathbb{R}^d$,

$$\mathbb{E}_x(Y \circ \theta_s | \mathcal{F}_s^+) = \mathbb{E}_{B_s} Y,$$

where the RHS is the function $\varphi(x) = \mathbb{E}_x Y$ evaluated at $x = B_s$.

Pf: We need to show that $\mathbb{E}_x(Y \circ \theta_s | A) = \mathbb{E}(\mathbb{E}_{B_s} Y; A)$ for all $A \in \mathcal{F}_s^+$.

We prove this equation for a carefully chosen special case and then use the monotone class

theorem to conclude the general case. Suppose $\mathbb{P}(Y(w) = \prod_{m=1}^n f_m(w(t_m)))$, $0 < t_1 < t_2 < \dots < t_n$ and the functions f_m are bounded and measurable.

Let $0 < s_0 < t_1$, $0 < s_1 < s_2 < \dots < s_k \leq s + t_n$, and $A = \{w: w(s_j) \in A_j, 1 \leq j \leq k\}$, A_j Borel measurable.

By def. of BM,

$$\mathbb{E}_x(Y \circ \theta_s; A) = \mathbb{E}\left(\prod_{m=1}^n f_m(w(t_m+s)) ; A\right) = \int_{A_1} dx_1 \int_{A_2} dx_2 \cdots \int_{A_k} dx_k$$

$$\int dy \int dx_{k+1} p_{s_k-s_{k-1}}(x_{k-1}, x_k) \int dy p_{s_{k+1}-s_k}(x_k, y) \cdot \varphi(y, h),$$

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$$\text{where } \varphi(y, h) = \int dy_1 P_{t_1-h}(y_1, y_1) f_1(y_1) \int dy_2 P_{t_2-h}(y_1, y_2) f_2(y_2) \cdots \int dy_n P_{t_n-h}(y_{n-1}, y_n) f_n(y_n).$$

We notice that the RHS can be written as $\mathbb{E}_x(\varphi(B_{s+h}, h); A)$.

Thus, we get $\mathbb{E}_x(Y_0 \theta_s; A) = \mathbb{E}_x(\varphi(B_{s+h}, h); A)$. This holds for all finite dimensional sets

A , so the π - λ thm implies that this identity holds for all $A \in \mathcal{F}_{s+h}^0 \supset \mathcal{F}_s^+$.

Denote $\psi(y) = f_1(y_1) \int dy_2 P_{t_2-h}(y_1, y_2) f_2(y_2) \cdots \int dy_n P_{t_n-h}(y_{n-1}, y_n) f_n(y_n)$. It is

easy to check that ψ is bdd measurable. Let $\phi(x, h) = \int dy_1 P_{t_1-h}(x, y_1) \psi(y_1)$.

Letting $h \rightarrow 0$ and using DCT, we get that if $x_n \rightarrow x$, then $\phi(x_n, h) \rightarrow \phi(x, 0)$ as $h \rightarrow 0$.

we take $\lim_{h \rightarrow 0}$ in $\mathbb{E}_x(Y_0 \theta_s; A) = \mathbb{E}_x(\varphi(B_{s+h}, h); A)$. Using DCT, we get that

$$\mathbb{E}_x(Y_0 \theta_s; A) = \underbrace{\mathbb{E}_x(\varphi(B_s, 0); A)}_{\mathbb{E}_{B_s} Y} \quad \forall A \in \mathcal{F}_s^+.$$

(+) $\Rightarrow \mathbb{E}_x(Y_0 \theta_s; A) = \mathbb{E}_x(\mathbb{E}_{B_s} Y; A) \quad \forall A \in \mathcal{F}_s^+$, for $Y = \prod_{m=1}^n f_m(w(t_m))$ with f_m being bdd measurable functions.

Next, we apply the MCT: $\mathcal{H} = \{ \text{bdd measurable fns for which (+) holds} \}$. If

satisfies (ii) and (iii). Let $\mathcal{A} = \{ \text{sets of the form } \{w : w(t_j) \in A_j\}, A_j - \text{Borel measurable} \}$

The special case treated above shows that $\mathbb{1}_A \in \mathcal{H} \quad \forall A \in \mathcal{A}$. The conclusion follows. \square

Note that $\mathbb{E}_x(Y_0 \theta_s | \mathcal{F}_s^+) = \mathbb{E}_{B(s)} Y \in \mathcal{F}_s^0 \Rightarrow \mathbb{E}_x(Y_0 \theta_s | \mathcal{F}_s^+) = \mathbb{E}(Y_0 \theta_s | \mathcal{F}_s^0)$. Then:

Theorem: If $Z \in \mathcal{G}$ is bounded, then $\forall s \geq 0$ and $x \in \mathbb{R}^d$,

$$\mathbb{E}_x(Z | \mathcal{F}_s^+) = \mathbb{E}_x(Z | \mathcal{F}_s^0).$$

Pf: Similar to the pf of the previous thm, it suffices to prove the result when $Z = \prod_{m=1}^n f_m(B(t_m))$.

Then, Z can be written as $Z = (Y_0 \theta_s) \cdot (X)$, where X is \mathcal{F}_s^0 measurable and $Y \in \mathcal{G}$.

Thus, $\mathbb{E}_x(Z | \mathcal{F}_s^+) = X \mathbb{E}_x(Y_0 \theta_s | \mathcal{F}_s^+) = X \mathbb{E}_{B_s} Y \in \mathcal{F}_s^0$. \square

As a special case, if $Z \in \mathcal{F}_s^+$, then this thm implies $Z = \mathbb{E}(Z | \mathcal{F}_s^+) = \mathbb{E}_x(Z | \mathcal{F}_s^0) \in \mathcal{F}_s^0$.

Hence, \mathcal{F}_s^0 and \mathcal{F}_s^+ are the same up to null sets (sets with measure zero).

Thm: (Blumenthal's 0-1 law) If $A \in \mathcal{F}_0^+$, then $\forall x \in \mathbb{R}^d$, $\mathbb{P}_x(A) \in \{0, 1\}$.

Pf: $\forall A \in \mathcal{F}_0^+$, we have $\mathbb{1}_A = \mathbb{E}_x(\mathbb{1}_A | \mathcal{F}_0^+) = \mathbb{E}_x(\mathbb{1}_A | \mathcal{F}_0^0) = \mathbb{P}_x(A)$ \mathbb{P}_x -a.s.

where we use that \mathcal{F}_0^0 is trivial under \mathbb{P}_x . \square

\mathcal{F}_0^+ : the germ field. This says that \mathcal{F}_0^+ is trivial. It is useful in studying the local behavior of Brownian paths.

Thm: For 1D BM, $\mathbb{P}_0(\tau_0 = 1) = 1$ where $\tau_0 := \inf\{t \geq 0 : B_t > 0\}$.

Pf: We have $\mathbb{P}_0(\tau_0 \leq t) \geq \mathbb{P}_0(B_t > 0) = 1/2$. Letting $t \rightarrow 0$, we get $\mathbb{P}_0(\tau_0 = 0) = \lim_{t \rightarrow 0} \mathbb{P}_0(\tau_0 \leq t) \geq \frac{1}{2}$. Since $\{\tau_0 = 0\}$ is in the germ field \mathcal{F}_0^+ , we have $\mathbb{P}_0(\tau_0 = 0) = 1$. \square

By symmetry, BM must also hit $(-\infty, 0)$ immediately, i.e., $\mathbb{P}_0(\tau_{-\infty} = 0) = 1$ for $\tau_{-\infty} := \inf\{t \geq 0 : B_t < 0\}$.

Since B_t is continuous in t , we get:

Thm: Let $T_0 := \inf\{t \geq 0 : B_t = 0\}$. Then, $\mathbb{P}_0(T_0 = 0) = 1$.

Using the following time inversion trick, we can use the above results to get information about the behavior as $t \rightarrow +\infty$.

Thm: If B_t is a BM starting at 0, then so is the process defined by $X_0 = 0$ and $X_t = +B(1/t)$ for $t > 0$.

Pf: We check the second definition of BM. (i) If $0 < t_1 < t_2 < \dots$, then $(X|t_1), \dots, X|t_n)$ has multivariate normal distribution with mean 0. (ii) $\mathbb{E}X_s = 0$ and if $s < t$, then

$$\mathbb{E}(X_t | X_s) = ts \mathbb{E}(B(1/t) B(1/s)) = ts \cdot \frac{1}{t} = s. \quad (\text{iii}) \quad X_t \text{ is continuous at } t \neq 0. \quad \text{We still}$$

need to check the continuity at $t=0$. We need to show that $B(t)/t \rightarrow 0$ a.s. as $t \rightarrow 0$.

By the SLLN, we have that $B_n/n \rightarrow 0$ a.s. as $n \rightarrow \infty$ through integers. To handle values in between integers, we use Kolmogorov's maximal inequality to conclude that

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$$P\left(\sup_{0 \leq k \leq m} |B_{n+k}-B_n| > n^{2/3}\right) \leq n^{-\frac{4}{3}} \text{Var}(B_{n+1}-B_n) = n^{-\frac{4}{3}}.$$

$$\text{Letting } m \rightarrow \infty, \text{ we have } P\left(\sup_{n \in [n, n+1]} |B_n - B_{n+1}| > n^{2/3}\right) \leq n^{-\frac{4}{3}}$$

Since $\sum n^{-\frac{4}{3}} < \infty$, by B-C, we get that $B_n/n \rightarrow 0$ a.s. as $n \rightarrow \infty$. \square

The above thm allows us to relate the behavior of B_t as $t \rightarrow \infty$ and as $t \rightarrow -\infty$.

Together with Blumenthal's 0-1 law, it gives some interesting results.

Let $F_t^1 = \sigma(B_s : s \geq t)$ - the future at time t , $\mathcal{F} = \bigcap_{t \geq 0} F_t^1$ = the tail σ -field

Thm: If $A \in \mathcal{F}$, then either $P_x(A) \in \{0, 1\} \forall x \in \mathbb{R}$, either $P_x(A) = 0$ or $P_x(A) = 1 \forall x \in \mathbb{R}$.

Rmk: This form is stronger than Blumenthal's 0-1 law. In that case, $P_x(A)$ may depend on x .

Pf: The tail σ -field of B_t is the same as the germ σ -field of X_t . Hence, we have $P_x(A) \in \{0, 1\}$. To improve this, observe that $A \in F_t^1$, so $1_A = 1_D \circ \theta_t$, for some event D .

Applying the Markov property, we get

$$\begin{aligned} P_x(A) &= E_x(1_D \circ \theta_t) = E_x(E_x(1_D \circ \theta_t | F_t^1)) = E_x(E_B 1_D) \\ &= \frac{1}{2\pi} \int P_y(D) e^{-\frac{(y-x)^2}{2}} dy. \end{aligned}$$

Taking $x=0$, we observe that if $P_0(A)=0$, then $P_y(D)=0$ for a.e. y w.r.t. to the Lebesgue measure. Using the above formula again, we get $P_x(A)=0$ for all x .

If $P_0(A)=1$, then $A^c \in \mathcal{F}$ with $P_0(A^c)=0$. The above argument shows that $P_x(A^c)=0$ for all x . \square

As an application of the above result:

Thm: Let B_t be a 1D BM starting at 0. Then with prob. 1,

$$\limsup_{t \rightarrow \infty} B_t/\sqrt{t} = \infty, \quad \liminf_{t \rightarrow \infty} B_t/\sqrt{t} = -\infty.$$

Pf: Let $0 < K < \infty$. We have

$$P_0(B_n/\sqrt{n} > K \text{ i.o.}) \geq \limsup_{n \rightarrow \infty} P_0(B_n > K\sqrt{n}) = P_0(B_1 > K) > 0.$$

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So the 0-1 law tells that $P_0(B_n > K\sqrt{n} \text{ i.o.}) = 1$. Since K is arbitrary, we get the first result. \square

Thm: Let B_t be a 1D BM and let $A = \bigcap_n \{B_t=0 \text{ for some } t \geq n\}$. Then, $P_x(A)=1$ for all x .

Pf: This follows from the previous thm, translation invariance, and continuity of Brownian paths. \square

In other words, this says that 1D BM is "recurrent". For any starting point x , it will return to 0 i.o., i.e., there is a sequence of times $t_n \uparrow \infty$ so that $B_{t_n}=0$.

The Blumenthal's 0-1 law tells that F_t^+ and F_t^0 are identical up to null sets

Let $N_x = \{A : A \subseteq D \text{ with } P_x(A)=0\}$, $F_t^x = \sigma(F_t^+ \cup N_x)$, $F_\infty^x = \bigcap_{t \geq 0} F_t^x$.

N_x are the null sets and F_t^x are the completed σ -fields for P_x . Since we do not want the filtration to depend on the initial state, we take the intersection of all σ -fields.

By def, F_t are right-continuous, i.e., $F_t = \bigcap_{s \geq t} F_s$.

3. Strong Markov Property

A ~~RV~~ T taking values in $[0, \infty]$ is called a stopping time if $\{T \geq t\} \in \mathcal{F}_t$,

Another $\{T < t\} \in \mathcal{F}_t$. An equivalent def. is to require that $\{T \leq t\} \in \mathcal{F}_t$. These two

defs are equivalent in continuous time for a right continuous filtration \mathcal{F}_t :

(1) If $\{T \leq t\} \in \mathcal{F}_t$, then $\{T \geq t\} = \bigcup_n \{t \leq T < t + \frac{1}{n}\} \in \mathcal{F}_t$.

(2) If $\{T < t\} \in \mathcal{F}_t$, then $\{T \leq t\} = \bigcap_n \{T < t + \frac{1}{n}\} \in \mathcal{F}_t$ (by using that \mathcal{F}_t is right

Thm: (Hitting time) If G is an open set and $T = \inf\{t \geq 0 : B_t \in G\}$, then T is a stopping time.

Pf: Since G is open and $t \mapsto B_t$ is continuous, we have $\{T < t\} = \bigcup_{g \in G} \{B_g \in G\} \in \mathcal{F}_t$. \square

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Thm: If T_n is a sequence of stopping times and $T_n \downarrow T$, then T is a stopping time.

Pf: $\{T < t\} = \bigcup_n \{T_n < t\}$ \square

Thm: If T_n is a sequence of stopping times and $T_n \uparrow T$, then T is a stopping time.

Pf: $\{T \leq t\} = \bigcap_n \{T_n \leq t\}$ \square

Thm: (Hitting time 2) If K is a closed set and $T = \inf\{t \geq 0 : B_t \in K\}$, then T is a stopping time.

Pf: Let $B(x, r) = \{y : |y - x| \leq r\}$ be the ball of radius r around x . Let $G_n = \bigcup_{x \in K} B(x, 1/n)$ and

$T_n := \inf\{t \geq 0 : B_t \in G_n\}$. Note G_n is open, so T_n is a stopping time. We claim that

$T_n \uparrow T$. First, notice that $T_n \leq T \forall n$, so $\lim_{n \rightarrow \infty} T_n \leq T$. On the other hand, suppose

$T_n \uparrow S < \infty$. Since $B(T_n) \in G_n$ for all n and $B(S) \in G_n$, we get that

$B(S) \in K$ (since K is closed) and $T \leq S$. \square

Combining the above results, we see that the hitting time of a countable union of closed / countable intersection of closed / open sets are stopping times.

In fact, the hitting time of any Borel set is a stopping time, but proving that is difficult.

Thm: Let S be a stopping time and let $S_n = \lfloor 2^n S \rfloor + 1 / 2^n$. Then, S_n are stopping times.

$$S_n = \begin{cases} (m+1)2^{-n} & \text{if } m2^{-n} \leq S < (m+1)2^{-n}, \\ & \quad (\text{integer part}) \end{cases}$$

i.e., we stop at the first time of the form $k2^{-n}$ after S (i.e., $> S$)

Pf: If $m2^{-n} < t \leq (m+1)2^{-n}$, then $\{S_n < t\} = \{S \leq m2^{-n}\} \in \mathcal{F}_{m2^{-n}} \in \mathcal{F}_t$. \square

Given a non-negative RV $S(w)$, we define the random shift operator θ_S , which

"cuts off the part of w before $S(w)$ and shift the ~~rest~~ ^{time axis} to the ~~right~~ ^{right} by $S(w)$ ".

In formulas, we write $(\theta_S(w))(t) = \begin{cases} w(S(w)+t) & \text{on } \{S < \infty\} \\ \Delta & \text{on } \{S = \infty\} \end{cases}$

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where Δ is an extra point added to $\mathbb{C}(IR^d)$. But, we will always explicitly restrict our attention to $\{S < \infty\}$.

Next, we define the σ -field \mathcal{F}_S - "the information known up to time S ".

$\mathcal{F}_S = \{A : A \cap \{S \leq t\} \in \mathcal{F}_t \text{ for all } t \geq 0\}$

When \mathcal{F}_t is right-continuous, this def. is unchanged if we replace $\{S \leq t\}$ by $\{S < t\}$.

Thm: If $T_n \downarrow T$ are stopping times, then $\mathcal{F}_T = \bigcap_n \mathcal{F}_{T_n}$.

Pf: If $A \in \mathcal{F}_S$, then $A \cap \{T \leq t\} = (A \cap \{S \leq t\}) \cap \{T \leq t\} \in \mathcal{F}_t$. \square

Thm: If $T_n \uparrow T$ are stopping times, then $\mathcal{F}_T = \bigcap_n \mathcal{F}_{T_n}$.

Pf: The previous thm tells that $\mathcal{F}_T \subset \mathcal{F}_{T_n} \forall n \Rightarrow \mathcal{F}_T \subset \bigcap_n \mathcal{F}_{T_n}$. On the other hand, let $A \in \bigcap_n \mathcal{F}_{T_n}$. Since $A \cap \{T_n < t\} \in \mathcal{F}_t$ and $T_n \uparrow T$, $A \cap \{T < t\} = \bigcup_n (A \cap \{T_n < t\}) \in \mathcal{F}_t$. \square

Thm: Let S be a stopping time. Then, $B_S \in \mathcal{F}_S$, i.e., the value B_S is measurable w.r.t. the information known up to time S .

Pf: Let $S_n = (\lfloor 2^n S \rfloor + 1) / 2^n$ be stopping times defined before. If A is a Borel set, we have $\{B(S_n) \in A\} = \bigcup_{m=1}^{\infty} \{S_n = m/2^n, B(m/2^n) \in A\} \in \mathcal{F}_{S_n}$.

Let $n \rightarrow \infty$, we have $S_n \downarrow S$ and $\{B(S) \in A\} = \lim_{n \rightarrow \infty} \{B(S_n) \in A\} \in \bigcap_n \mathcal{F}_{S_n} = \mathcal{F}_S$. \square

Now, we state the Strong MP, which says that MP holds at stopping times.

* Thm (Strong MP): Let $(s, w) \mapsto Y_s(w)$ be bounded and $\mathbb{R} \times \mathbb{R}^d$ measurable. If S is a stopping time, then $\forall x \in \mathbb{R}^d$, $\mathbb{E}_x(Y_S \circ \theta_S | \mathcal{F}_S) = \mathbb{E}_{B_S} Y_S$ on $\{S < \infty\}$, where the RHS is the function $y(x, t) = \mathbb{E}_x Y_t$ evaluated at $x = B_S$ and $t = S$.

Rmk: The proof ~~only~~ uses the following two properties of BM: (i) it is a Markov process; (ii) (the Feller property) if f is a bounded continuous function, then $x \mapsto \mathbb{E}_x f(B_t)$ is continuous.

Hence, the SMP holds for general Markov processes satisfying the Feller property.

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Pf. We first "discrete". The basic idea is to "discretize" the stopping time.

We first consider a simplified assumption that there is a sequence of times $t_n \uparrow \infty$, so that

$\mathbb{P}_x(S < \infty) = \sum_x \mathbb{P}_x(S = t_n)$. Then, the SMP follows directly from the MP.

Let $Z_n = Y_{t_n}(w)$ and $A \in \mathcal{F}_S$, then $\mathbb{E}_x(Y_S \circ \theta_S; A \cap \{S < \infty\}) = \sum_{n=1}^{\infty} \mathbb{E}_x(Z_n \circ \theta_{t_n}; A \cap \{S = t_n\})$ (*)

If $A \in \mathcal{F}_S$, then $A \cap \{S = t_n\} = A \cap \{S \leq t_n\} \in \mathcal{F}_{t_n}$. Hence, using the MP, we get that

(*) = $\sum_{n=1}^{\infty} \mathbb{E}_x(\mathbb{E}_{B(t_n)} Z_n; A \cap \{S = t_n\}) = \mathbb{E}_x(\mathbb{E}_{B(S)} Y_S; A \cap \{S < \infty\})$.

To prove the result in general, we let $S_n = ([\tau_n S] + 1)/2^n$ be the stopping time defined before.

To let $n \rightarrow \infty$, we restrict our attention to Y of the form $Y_s(w) = f_0(s) \prod_{m=1}^n f_m(w(t_m))$,

where $0 < t_1 < \dots < t_n$ and f_0, \dots, f_n are bounded continuous. Then, we can check that

$\varphi(x, s) = \mathbb{E}_x Y_s = f_0(s) \int dy_1 p_{t_0}(x, y_1) f_1(y_1) \dots \int dy_n p_{t_{n-1}, t_n}(y_{n-1}, y_n) f_n(y_n)$ is bdd. continuous.

Next, we complete the proof. Let $A \in \mathcal{F}_S$. Since $s \leq S_n$, we have $A \in \mathcal{F}_S \subset \mathcal{F}_{S_n}$. Applying

the special case proved above to S_n and observing that $\{S < \infty\} = \{S_n < \infty\}$, we get that

$\mathbb{E}_x(Y_{S_n} \circ \theta_{S_n}; A \cap \{S < \infty\}) = \mathbb{E}_x(\varphi(B(S_n), S_n); A \cap \{S < \infty\})$

As $n \rightarrow \infty$, $S_n \downarrow S$, $B(S_n) \rightarrow B(S)$, $\varphi(B(S_n), S_n) \rightarrow \varphi(B(S), S)$ and $Y_{S_n} \circ \theta_{S_n} \rightarrow Y_S \circ \theta_S$.

Then, by the DCT, we get that $\mathbb{E}_x(Y_S \circ \theta_S; A \cap \{S < \infty\}) = \mathbb{E}_x(\varphi(B(S), S); A \cap \{S < \infty\})$

Holds when Y takes the form $Y_s(w) = f_0(s) \prod_{m=1}^n f_m(w(t_m))$.

Finally, we apply the Monotone class theorem. Let $\mathcal{H} = \text{collection of } Y \text{ for which}$

$\mathbb{E}_x(Y_S \circ \theta_S; A) = \mathbb{E}_x(\mathbb{E}_{B(S)} Y_S; A) \quad \forall A \in \mathcal{F}_S$.

It is easy to see that (ii) If $f, g \in \mathcal{H}$, then $f+g \in \mathcal{H}$ and $c f \in \mathcal{H} \quad \forall c \in \mathbb{R}$.

(iii) If $f_n \in \mathcal{H}$ are monotone and $f_n \uparrow f$, then $f \in \mathcal{H}$.

To show (i) for some \mathcal{A} that generates the σ -field $\mathcal{B} \times \mathcal{P}_x$, we let \mathcal{A} to be

the collection of sets of the form $A = G_0 \times \{w: w(s_j) \in G_j, 1 \leq j \leq k\}$, where

$s_1 < s_2 < \dots$ and G_j are open sets. We let $K_j = G_j^c$ and $f_j^n(x) = 1 \wedge np(x, K_j)$, where $p(x, K) = \inf \{ |x-y| : y \in K \}$. Then, f_j^n are bounded continuous functions with $f_j^n \uparrow 1_{G_j}$. But, we have shown that $Y_S^n(w) = \sum_{j=0}^n f_j^n(s_j) \prod_{i=1}^{j-1} f_i^n(w(s_i)) \in \mathcal{H}$ and (iii) holds. This implies that $1_A \in \mathcal{H}$, and verifies (i) in the MCT. \square

4. Path Properties

We will use the SMP to derive some path properties for 1D BM.

① Zeros of BM

Let $R_t = \inf \{u > t : B_u = 0\}$ and let $T_0 = \inf \{u > 0 : B_u = 0\}$. We already know that $\mathbb{P}_x(R_t < \infty) = 1$, so $B(R_t) = 0$ a.s. and the SMP implies that

$\mathbb{P}_x(T_0 > 0 \mid \mathcal{F}_{R_t}) = \mathbb{P}_x(T_0 > 0) = 0 \Rightarrow \mathbb{P}_x(T_0 > R_t > 0 \text{ for some rational } t) = 0$ (*)

Let $Z(w) = \{t : B_t(w) = 0\}$ be the set of zeros of w .

Then, $Z(w)$ is a closed set. Moreover, it has no isolated point (i.e., $Z(w)$ is a perfect set).

To see this, we notice that (*) implies: if a point $\overset{\text{u}}{t} \in Z(w)$ is isolated from the left (i.e., there is a rational $t < u$ so that $(t, u) \cap Z(w) = \emptyset$), then a.s., it is a decreasing limit of points in $Z(w)$.

We know from the point topology that a perfect set is uncountable. Hence, $Z(w)$ is uncountable.

Moreover, by Fubini's thm, $\forall T > 0$,

$$\mathbb{E}_x([1_{Z(w)}] \cap [0, T]) = \int_0^T \mathbb{P}_x(B_t = 0) dt = 0.$$

So, $Z(w)$ is also a measure zero set. (In fact, Z is a fractal set of dimension $1/2$.)

② Hitting times: $T_a = \inf \{t \geq 0 : B_t = a\}$. Our goal is to derive the distribution \mathbb{P}_a .

By translation symmetry, it suffices to study this problem for $a = 0$.

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Thm: Under P_0 , $\{T_a : a \geq 0\}$ has stationary independent increments.

Pf: Notice that if $0 < a < b$, then $T_a < T_b$ and $T_b \circ \theta_{T_a} = T_b - T_a$. Hence, w/ bdd measurable function f , using the SMP and translation invariance, we get

$$E_0(f(T_b - T_a) | F_{T_a}) = E_0(f(T_b) \circ \theta_{T_a} | F_{T_a}) = E_0 f(T_b) = E_0 f(T_b-a).$$

This gives stationarity of increments. It remains to prove independence.

Let $a_0 < a_1 < \dots < a_n$, and let f_1, f_2, \dots, f_n be bounded measurable. Define $F_i = f_i(T_{a_i} - T_{a_{i-1}})$.

Conditioning on $F_{T_{a_{n-1}}}$, we get

$$\begin{aligned} E_0 f_n(T_{a_n} - a_{n-1}) &= E_0 f_n(T_{a_n} - T_{a_{n-1}}) \\ &\stackrel{\text{def}}{=} E_0 F_n \\ E_0\left(\prod_{i=1}^n F_i\right) &= E_0\left(\prod_{i=1}^n F_i | F_{T_{a_{n-1}}}\right) = E_0\left(\prod_{i=1}^{n-1} F_i \cdot E_0(F_n | F_{T_{a_{n-1}}})\right) \\ &= E_0\left(\prod_{i=1}^{n-1} F_i\right) E_0 F_n. \end{aligned}$$

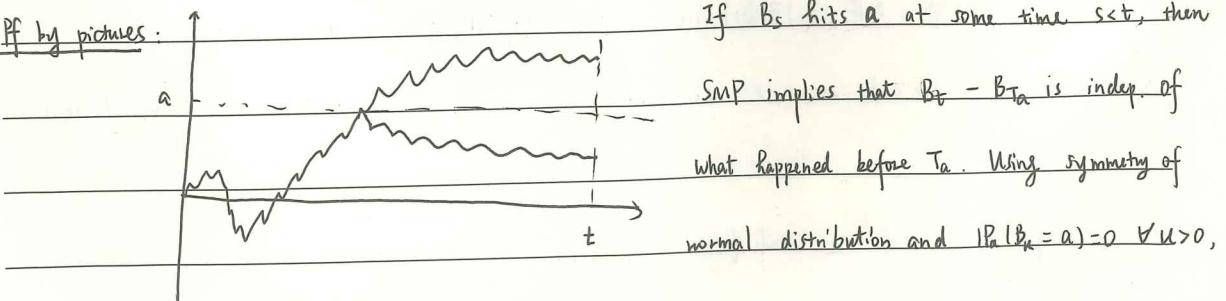
By induction, we get $E_0 \prod_{i=1}^n F_i = \prod_{i=1}^n E_0 F_i$, which implies the independence. \square

Let $t_k = T_{a_k} - T_{a_{k-1}}$. The above results implies that t_1, t_2, t_3, \dots are i.i.d. and each of them has the same distribution as T_1 . On the other hand, by scaling relation, we have $T_n \stackrel{d}{=} n^2 T_1$. Hence, we have $\frac{t_1 + t_2 + \dots + t_n}{n^2} \rightarrow T_1$ in law. This shows that T_1 has stable law (and hence any T_a has stable law). Using the above scaling and the fact that $T_a \geq 0$, we can determine the law of T_1 uniquely. Here we give a direct calculation of the distribution of T_a using the reflection principle.

Theorem: If $a > 0$, we have $P_0(T_a < t) = 2P_0(B_t > a)$.

Rank: Since BM is continuous, in contrast to the reflection principle for MC, we have " $=$ " here.

Pf by pictures:



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we get $P_0(T_a < t, B_t > a) = \frac{1}{2} P_0(T_a < t) \Rightarrow P_0(T_a < t) = 2P_0(B_t > a)$

$$P_0(T_a < t, B_t > a)$$

Mathematical proof: We let $Y_s(w) = \begin{cases} 1, & \text{if } s < t \text{ and } w(t-s) > a \\ 0, & \text{otherwise} \end{cases}$

Define $S = \inf\{s < t : B_s = a\}$ with $\inf \emptyset = \infty$. Then, we have $Y_s \circ \theta_S(w) = \begin{cases} 1, & \text{if } S < t, B_t(w) > a \\ 0, & \text{otherwise} \end{cases}$

Then, using the SMP, we get

$$E_0(Y_s \circ \theta_S | F_S) = E_{B_S} Y_s \text{ on } \{S < \infty\} = \{T_a < t\}$$

$$= E_0 Y_S = \frac{1}{2} \text{ on } \{S < \infty\} = \{T_a < t\}$$

$\Rightarrow E_0(E_0(Y_s \circ \theta_S | F_S)) = E_0(Y_s \circ \theta_S)$ Taking expectation over $\{S < \infty\}$ of both sides.

$$\text{LHS} = E_0(E_0(Y_s \circ \theta_S | F_S) \mathbf{1}_{S < \infty}) = E_0(Y_s \circ \theta_S ; S < \infty) = P_0(T_a < t, B_t > a)$$

$$\text{RHS} = \frac{1}{2} E_0(\mathbf{1}_{S < \infty}) = \frac{1}{2} P_0(T_a < t)$$

We have calculated that $P_0(T_a < t) = 2P_0(B_t > a) = \int_{a/\sqrt{t}}^{\infty} \frac{2}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}} dx$
 $= \int_{a/\sqrt{t}}^{\infty} \frac{2}{\sqrt{2\pi t}} e^{-\frac{s^2}{2t}} ds = \int_0^{\sqrt{t}/a} \frac{2}{\sqrt{2\pi t}} e^{-\frac{1}{2s^2}} ds$
 $\stackrel{x=as^2}{=} \int_0^{\sqrt{t}/a} \frac{1}{\sqrt{2\pi t^3}} \exp(-\frac{a^2}{2t}) ds. \text{ Hence, the PDF is } \frac{a}{\sqrt{2\pi t^3}} \exp(-\frac{a^2}{2t}).$

Example: $L = \sup\{t \leq 1 : B_t = 0\}$ the last visit time of 0. $P_0(T_x > 1-s)$

$$P_0(L \leq s) = \int_{-\infty}^{\infty} P_0(0, x) P_0(T_0 > 1-s) ds = \int_{-\infty}^{+\infty} P_0(0, x) P_0(T-x > 1-s) ds$$

$$= 2 \int_0^{\infty} \frac{1}{\sqrt{2\pi s}} \exp(-\frac{x^2}{2s}) \int_{1-s}^{\infty} \frac{x}{\sqrt{2\pi r^3}} \exp(-\frac{x^2}{2r}) dr ds$$

$$= \frac{1}{\pi} \int_{1-s}^{\infty} (sr^3)^{-\frac{1}{2}} \int_0^{\infty} x \exp(-\frac{x^2}{2}(\frac{1}{r} + \frac{1}{s})) dx dr$$

$$= \frac{1}{\pi} \int_{1-s}^{\infty} (sr^3)^{-\frac{1}{2}} \frac{rs}{r+s} dr = \frac{1}{\pi} \int_{1-s}^{\infty} (\frac{s}{r})^{\frac{1}{2}} \frac{1}{r+s} dr \quad t = \frac{s}{r+s}$$

$$= \frac{1}{\pi} \int_0^s \frac{1}{\sqrt{t(1-t)}} dt = \frac{2}{\pi} \int_0^{\sqrt{s}} \frac{dt}{\sqrt{1-t^2}} = \frac{2}{\pi} \arcsin(\sqrt{s}).$$

\Rightarrow The arcsin law.

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5. Martingales

X_t is a continuous time martingale if (i) $\mathbb{E}|X_t| < \infty$; (ii) X_t is adapted to \mathcal{F}_t ; (iii) $\mathbb{E}(X_t | \mathcal{F}_s) = X_s \quad \forall s < t$.

Suppose $\text{Thm: } X_t$ is a right-continuous martingale adapted a right-continuous filtration. If T is a bounded stopping time, then $\mathbb{E}X_T = \mathbb{E}X_0$. Let n be an integer s.t. $T \leq n-1$ a.s.

Pf: We again discretize T as $T_m = (\lceil 2^m T \rceil + 1)/2^m$. Then, Note $Y_k^m = X(\lfloor k2^{-m} \rfloor)$ is a martingale w.r.t. $\mathcal{F}_k^m = \mathcal{F}(\lfloor k2^{-m} \rfloor)$ and $S_m = \sum_{k=0}^{m-1} T_m$ is a stopping time for (Y_k^m, \mathcal{F}_k^m) . Hence,

$$X(T_m) = Y_{S_m}^m = \mathbb{E}(Y_{S_m}^m | \mathcal{F}_{T_m}^m) = \mathbb{E}(X_m | \mathcal{F}(T_m)).$$

As $m \uparrow \infty$, $X(T_m) \rightarrow X(T)$ and $\mathcal{F}(T_m) \downarrow \mathcal{F}(T)$ by right-continuity of the path and filtration. Thus, we get $X(T) = \lim_{m \uparrow \infty} \mathbb{E}(X_m | \mathcal{F}(T_m)) = \mathbb{E}(X_m | \mathcal{F}(T))$. Taking expectation gives

$\mathbb{E}X_T = \mathbb{E}X_0 = \mathbb{E}X$ since X is a martingale.

Thm: B_t is a martingale w.r.t. the filtration \mathcal{F}_t .

Pf: By MP, $\mathbb{E}_x(B_t | \mathcal{F}_s) = \mathbb{E}_{B_s}(B_{t-s}) = B_s$ a.s., where we used $\mathbb{E}_y B_u = y \quad \forall u \geq 0$ \square

Thm: If $a < x < b$, then $\mathbb{P}_x(T_a < T_b) = \frac{b-x}{b-a}$

Pf: Let $T = T_a \wedge T_b$. we have seen that $T < \infty$ a.s. Then $\forall t > 0$, T_{nt} is a bounded stopping time, so we have $x = \mathbb{E}_x B_{nt}$. Letting $t \rightarrow \infty$ and using the DCT (since $a \leq B_{nt} \leq b$), we get

$$x = \mathbb{E}_x B_T \Rightarrow x = a \mathbb{P}_x(T_a < T_b) + b \mathbb{P}_x(T_a > T_b) = a \mathbb{P}_x(T_a < T_b) + b(1 - \mathbb{P}_x(T_a < T_b)).$$

Thm: $B_t^2 - t$ is a martingale.

Pf: $\mathbb{E}_x(B_t^2 | \mathcal{F}_s) = \mathbb{E}_x[B_s^2 + 2(B_t - B_s)B_s + (B_t - B_s)^2 | \mathcal{F}_s] = B_s^2 + B_s \mathbb{E}_x(B_t - B_s) + \mathbb{E}_x[(B_t - B_s)^2]$
 $= B_s^2 + t - s \quad \square$

Thm: If $a < x < b$, let $T = \inf\{t: B_t \notin [a, b]\}$. Then, $\mathbb{E}_x T = (b-x)(x-a)$.

Pf: $\forall t > 0$, $\mathbb{P}_x(T_{nt} < t)$ is a bounded stopping time. So, $\mathbb{E}_x(B_{T_{nt}}^2 - T_{nt}) = x^2$

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Letting $t \rightarrow \infty$, using the MCT, we get $\mathbb{E}(B_{T_{nt}}^2) \rightarrow \mathbb{E}B_T^2$; using the DCT, we get

$$\mathbb{E}B_{T_{nt}}^2 \rightarrow \mathbb{E}B_T^2. \text{ Thus, we get } x^2 = \mathbb{E}_x(B_T^2 - T) = a^2 \mathbb{P}(T_a < T_b) + b^2 \mathbb{P}(T_b \leq T_a) - \mathbb{E}_x T$$

$$\Rightarrow \mathbb{E}_x T = a^2 \frac{b-x}{b-a} + b^2 \frac{x-a}{b-a} - x^2 = -ab + \frac{\theta}{2}(b+a)x - x^2 = (b-x)(x-a). \quad \square$$

Thm: $\exp(\theta B_b - \frac{\theta^2}{2}t)$ is a martingale.

$$\text{Pf: } \mathbb{E}_x(\exp(\theta B_t - \frac{\theta^2}{2}t) | \mathcal{F}_s) = \exp(\theta B_s) \mathbb{E}_x(\exp(\theta(B_t - B_s)) | \mathcal{F}_s)$$

$$= \exp(\theta B_s) \mathbb{E}_x \exp(\theta(B_t - B_s)) = \exp(\theta B_s) \exp(\frac{\theta^2}{2}(t-s)). \quad \square$$

Thm: $\forall t > 0$, we have $\mathbb{E}_0 \exp(-\lambda T_a) = e^{-\lambda \sqrt{2}a}$.

Pf: $\forall t > 0$, we have $1 = \mathbb{E}_0 \exp(\theta B(T_a t) - \frac{\theta^2}{2}T_a t)$. Taking $t \rightarrow \infty$ and using

DCT, we get $1 = \mathbb{E}_0 \exp(\theta B(T_a) - \frac{\theta^2}{2}T_a)$. Taking $\theta = \sqrt{2}\lambda$, we get $\mathbb{E}_0 e^{-\lambda T_a} = e^{-\lambda \sqrt{2}a}$. \square

Rmk: Note $\mathbb{E}_0 \exp(-\lambda T_a)$ is the Laplace transform of T_a . Taking the inverse Laplace transform, we can calculate the distribution of T_a .

The following result gives a general way to construct martingales for 1D BM. It also can be derived from Ito's formula.

Thm: If $u(t, x)$ is a polynomial in t and x satisfying the heat equation:

$$\frac{\partial u}{\partial t} + \frac{1}{2} \frac{\partial^2 u}{\partial x^2} = 0, \text{ then } u(t, B_b) \text{ is a martingale.}$$

Pf: Recall the heat kernel $p_t(x, y) = \frac{1}{\sqrt{4\pi t}} \exp\left(-\frac{(y-x)^2}{4t}\right)$. we can check directly that it satisfies the heat equation $\frac{\partial p_t}{\partial t} + \frac{1}{2} \frac{\partial^2 p_t}{\partial x^2} = 0$. Then,

$$\frac{\partial}{\partial t} \mathbb{E}_x u(t, B_b) = \int \frac{\partial}{\partial t} (p_t(x, y) u(t, y)) dy = \frac{1}{2} \int \frac{\partial^2}{\partial y^2} p_t(x, y) \cdot u(t, y) dy$$

$$+ \int p_t(x, y) \frac{\partial}{\partial t} u(t, y) dy \quad \square \quad (\text{Integrating by parts twice})$$

$$= \int p_t(x, y) \left(\frac{1}{2} \frac{\partial^2}{\partial y^2} u(t, y) + \frac{\partial}{\partial t} u(t, y) \right) dy = 0. \quad (\text{Since } u(t, y) \text{ is polynomial})$$

there are no boundary terms in the integration by parts argument.)

So far, we have seen that $\mathbb{E}_x u(t, B_b)$ is constant in t . For the mg property,

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We need to prove that $E_x(u(t, B_t) | F_s) = u(s, B_s)$, $\forall 0 < s < t$.

We use the MP to get $E_x[u(t, B_t) | F_s] = E_x[v(t-s, B_{t-s}) \circ B_s | F_s]$

$= E_{B(s)} v(t-s, B_{t-s}) = v(0, B_s) = u(s, B_s)$, where $v(r, x) = u(r+s, x)$ also satisfies

\rightarrow it is constant in t ; so we take $t=s$. the heat equation w.r.t. r and x . \square

Examples of polynomials: $x, x^2-t, x^3-3tx, x^4-6x^2t+3t^2, \dots$

The result can also be extended to $\exp(\theta x - \frac{\theta^2}{2}t)$, since it ~~decreases~~ increases only exponentially.

Example: Let $T = \inf\{t : B_t \notin (-a, a)\}$. Then, $B_T^4 - 6B_{T \wedge T}^2 + 3(T \wedge t)^2$ is a martingale,

so we have $E_x[B_T^4 - 6E_x[B_{T \wedge T}^2] + 3E_x(T \wedge t)^2] = 0$. We know $T < \infty$ a.s.

Taking $t \rightarrow \infty$, we get $E_x[B_T^4] - 6E_x[B_{T \wedge T}^2] + 3E_x[T^2] = 0$

$$a^4 - 6a^2 \underbrace{E_x[T^2]}_{\Theta a^2} + 3E_x[T^2] = 0 \Rightarrow E_x[T^2] = \frac{5}{3}a^4. \quad \square$$

6. Donsker's Theorem (RW converges to BM)

Thm (Skorokhod's representation thm) If $E_x=0$, $E_x^2 < \infty$, then there is a stopping time

T for BM so that $B_T \stackrel{d}{=} X$ and $E_x T = E_x^2$.

Pf: Suppose X is supported on $[a, b]$ with $a < 0 < b$. Since $E_x=0$, we have

$$P(X=a) = \frac{b}{b-a}, \quad P(X=b) = \frac{-a}{b-a}$$

We let $T = T_{a,b} = \inf\{t : B_t \notin (a, b)\}$. The previous thm implies $B_T \stackrel{d}{=} X$ and

$E_x T = -ab = E_x B_T^2$. For the general case, we write $F(x) = P(X \leq x)$ as a mixture of

two point distributions with mean 0. Let $C = - \int_{-\infty}^0 u dF(u) = \int_0^\infty v dF(v)$.

Let φ be a bounded function with $\varphi(0)=0$. Then,

$$C = \int \varphi(x) dF(x) = \left(\int_0^\infty \varphi(v) dF(v) \right) \left(\int_{-\infty}^0 (-u) dF(u) \right)$$

$$+ \left(\int_{-\infty}^0 \varphi(u) dF(u) \right) \left(\int_0^\infty v dF(v) \right)$$

$$= \int_0^\infty dF(v) \int_{-\infty}^0 dF(u) (v\varphi(u) - u\varphi(v))$$



Thus, we have $\int \varphi(x) dF(x) = C^{-1} \int_0^\infty dF(v) \int_{-\infty}^0 dF(u) (v-u) \left(\frac{v}{v-u} \varphi(u) + \frac{-u}{v-u} \varphi(v) \right)$ (*)
A two-point distribution

We let $(u, v) \in \mathbb{R}^2$ such that $P((u, v) = (0, 0)) = F(\{0\})$,

$$(+) P((u, v) \in A) = C^{-1} \iint_{(u,v) \in A} dF(u) dF(v) (v-u), \quad A \subset (-\infty, 0] \times [0, \infty)$$

Then, define a family of two-point probability measures: $\mu_{0,0}(\{0\}) = 1$ and

$$\mu_{u,v}(\{u\}) = \frac{v}{v-u}, \quad \mu_{u,v}(\{v\}) = \frac{-u}{v-u} \quad \text{for } u < 0 < v.$$

Now, we write (*) as $\int \varphi(x) dF(x) = E \int \varphi(x) d\mu_{u,v}(x)$.

We can show that the above holds without requiring $\varphi(0)=0$.

(Taking $\varphi \equiv 1$ shows that (+) defines a prob. measure.)

From the above calculations, it follows that if we have (u, v) with distributions given in (+) and an independent BM defined on the same space, then $B(T_{u,v}) \stackrel{d}{=} X$.

To make $T_{u,v}$ a stopping time, we let $T_0 = 0$, u, v .

Next, we check that: $E(T_{u,v}) = E \{ E[T_{u,v} | u, v] \} = -E(UV)$.

On the other hand, with (+) we get that

$$\begin{aligned} E(-UV) &= C^{-1} \int_0^\infty dF(v) \int_{-\infty}^0 dF(u) (-uv)(v-u) \\ &= \cancel{\int_{-\infty}^0 dF(u) (-u)} \left\{ -u + C^{-1} \int_0^\infty dF(v) v^2 \right\} \\ &= \int_{-\infty}^0 u^2 dF(u) + \int_0^\infty v^2 dF(v) = E(X^2). \end{aligned}$$

A simple extension of the above result gives:

Thm: Let X_1, X_2, \dots be i.i.d. with a distribution F , which has mean 0 and variance 1.

Let $S_n = X_1 + X_2 + \dots + X_n$. There is a sequence of stopping times $T_0 = 0, T_1, T_2, \dots$ such that $S_n \stackrel{d}{=} B(T_n)$ and $T_n - T_{n-1}$ are i.i.d.

Pf: Let $(u_1, v_1), (u_2, v_2), \dots, (u_n, v_n)$ be i.i.d. and have distribution given in (+).

Let B_t be an independent BM. Let $T_0 = 0$, and, for $n \geq 1$, let

$$T_n = \inf \{ t \geq T_{n-1} : B_t - B(T_{n-1}) \notin (u_n, v_n) \}.$$

With this result, we immediately get another way of showing the CLT of S_n .

Thm (CLT) Under the setting of the above thm, $S_n / \sqrt{n} \rightarrow N(0, 1)$ in law.



Pf: By Brownian scaling, then $S_n/\sqrt{n} \xrightarrow{d} B(T_n)/\sqrt{n} \xrightarrow{d} B(T_n/n)$.

By WLLN, $T_n/n \rightarrow 1$ in probability, which would imply that $B(T_n/n) \rightarrow B(1)$, in law.

To fill in the details, $\forall \varepsilon > 0$, pick $\delta > 0$ so that $\mathbb{P}(|B_t - B_1| > \varepsilon \text{ for some } t \in (1-\delta, 1+\delta)) < \varepsilon/2$.

Then, pick N large enough so that $\mathbb{P}(|T_n/n - 1| > \delta) < \varepsilon/2 \quad \forall n \geq N$. The last two estimates give $\forall n \geq N$, $\mathbb{P}(|B(T_n/n) - B(1)| > \varepsilon) < \varepsilon$. Since ε is arbitrary, it follows that $B(T_n/n) \rightarrow B(1)$ in prob. \square

The main goal is to prove the celebrated Donsker's thm, showing that RW paths converge to BM paths. Given a RW S_n , let

$$S(u) = \begin{cases} S_k, & \text{if } u \in [k, k+1] \\ \text{linear on } [k, k+1] \text{ for } k \in \mathbb{N} \end{cases}$$

Thm: (Donsker's thm) Under the setting of the above thm, we get $\mathbb{P}(S(n \cdot)/\sqrt{n} \rightarrow B(\cdot))$ in law, i.e., the associated measures on $C([0, 1])$ converge weakly.

Pf: For simplicity of notations, let $X_{n,m}$, $1 \leq m \leq n$, be a triangular array of RVs,

$S_{n,m} = X_{n,1} + \dots + X_{n,m}$, and suppose $S_{n,m} = B(T_m^n)$. Then, let

$$S_{n,u} = \begin{cases} S_{n,m} & \text{if } u = m \in \{0, 1, \dots, n\} \\ \text{linear for } u \in [m-1, m] \text{ where } m \in \{1, \dots, n\} \end{cases}$$

Lemma: If $T_{[ns]}^n \rightarrow s$ in probability for each $s \in [0, 1]$, then $\|S_{n,(n \cdot)} - B(\cdot)\| \rightarrow 0$ in prob.

In the original setting, $X_{n,m} = X_m/\sqrt{n}$ and define T_1^n, \dots, T_n^n so that

$(S_{n,1}, \dots, S_{n,n}) \stackrel{d}{=} (B(T_1^n), \dots, B(T_n^n))$. By Brownian scaling, we have $T_m^n \xrightarrow{d} T_m/n$.

By LLN, the hypothesis of the lemma is satisfied.

To conclude the proof, we need to show that: \forall bold continuous function φ on $C([0, 1])$,

$\mathbb{E} \varphi(S_{n,(n \cdot)}) \rightarrow \mathbb{E} \varphi(B(\cdot))$. In fact,

$\forall \varepsilon > 0$, let $G_\delta := \{w \in C([0, 1]): \text{if } \|w - w'\| < \delta, \text{ then } |\varphi(w) - \varphi(w')| < \varepsilon\}$.

Since φ is continuous, $G_\delta \uparrow C([0, 1])$ as $\delta \downarrow 0$. Let $A = \|S_{n,(n \cdot)} - B(\cdot)\|$. Then,

$$|\mathbb{E} \varphi(S_{n,(n \cdot)}) - \mathbb{E} \varphi(B(\cdot))| \leq \varepsilon + (2 \sup_{w \in G_\delta} |\varphi(w)|) \left\{ \mathbb{P}(G_\delta^c) + \mathbb{P}(A \geq \delta) \right\}$$

↓ as $n \rightarrow \infty$

for large enough n . \square



Pf: By ^{unif.} continuity of B , if $\varepsilon > 0$ then there is a $\delta > 0$ so that $1/\delta \in \mathbb{N}$ and $(*) \quad \mathbb{P}(|B_t - B_s| < \varepsilon \quad \forall 0 \leq s \leq 1, |t-s| < 2\delta) > 1-\varepsilon$.

By the hypothesis of the lemma, there exists large enough N_5 so that if $n \geq N_5$ then

$$\mathbb{P}(|T_{[ns]}^n - ks| < \delta \text{ for } k=1, 2, \dots, \lfloor 1/\delta \rfloor) \geq 1-\varepsilon$$

Since $m \rightarrow T_m^n$ is increasing, we see that if $s \in ((k-1)\delta, k\delta)$,

$$\mathbb{P}(T_{[ns]}^n - s \geq 0) \geq \mathbb{P}(T_{[nk-1]\delta}^n - k\delta > -2\delta \text{ on the good event } E)$$

$$T_{[ns]}^n - s \leq T_{[nk]\delta}^n - (k-1)\delta < 2\delta \text{ on } E$$

$$\text{Hence, } \forall n \geq N_5, \quad \mathbb{P}(\sup_{0 \leq s \leq 1} |T_{[ns]}^n - s| < 2\delta) \geq 1-\varepsilon. \quad (**)$$

Combining (*) with (**), we get $\mathbb{P}(\sup_{0 \leq s \leq 1} |T_{[ns]}^n - s| < \varepsilon) \geq$

$$(***) \quad \mathbb{P}(|S_{n,m} - B_m/n| < \varepsilon \quad \forall m \leq n) \geq 1-2\varepsilon \text{ for large enough } n$$

To deal with $t = (m+\theta)/n$ with $0 < \theta < 1$, we observe that

$$\mathbb{P}(|S_{n,(n \cdot)} - B_t| \leq (1-\theta) |S_{n,m} - B_m/n| + \theta |S_{n,m+1} - B_{(m+1)/n}| + (1-\theta) |B_m/n - B_t| + \theta |B_{(m+1)/n} - B_t|)$$

We use (**) to bound the first two terms and (*) to bound the last two. Then, we get if $n \geq N_5$ and $1/n < 2\delta$, then $\|S_{n,(n \cdot)} - B(\cdot)\| < 2\varepsilon$ with prob. $\geq 1-2\varepsilon$. \square

With Donsker's thm, the following simple fact allows us to calculate some probabilities regarding RW using techniques of BMs:

If $\psi: C([0, 1]) \rightarrow \mathbb{R}$ is continuous \mathbb{P}_0 -a.s., then $\psi(S(n \cdot)/\sqrt{n}) \rightarrow \psi(B(\cdot))$ in law.

Example: Find the law of the maxima: $\max_{0 \leq m \leq n} S_m/\sqrt{n}$.

Let $\psi(w) = \max\{w(t): 0 \leq t \leq 1\}$. $\psi: C([0, 1]) \rightarrow \mathbb{R}$ is continuous. Then, by Donsker's thm, $\max_{0 \leq m \leq n} S_m/\sqrt{n} \Rightarrow M_n = \max_{0 \leq t \leq 1} B_t$. Recall that

$$\mathbb{P}_0(M_n \geq a) = \mathbb{P}_0(T_a \leq 1) = 2\mathbb{P}_0(B_1 > a). \quad \square$$