

Notes

STA302H1 - Fall 2020

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Contents

| | | |
|----------|---|----------|
| 1 | Module 10 - Diagnostics in MLR | 2 |
| 1.1 | Inference for a Single Regression Coefficient | 2 |
| 1.1.1 | Hypothesis Testing | 2 |
| 1.1.2 | Confidence Interval | 2 |
| 1.1.3 | Global F-test vs. Individual t-tests | 2 |
| 1.2 | Multicollinearity | 4 |
| 1.3 | Diagnostics and Remedies | 5 |
| 1.3.1 | Leverage and Influential Points | 5 |
| 1.3.2 | The Box-Cox Transformation | 5 |
| 1.4 | Added Variable Plot (Need to Review Lecture Part 3) | 6 |
| 1.5 | Data Analysis Flow | 7 |

1 Module 10 - Diagnostics in MLR

1.1 Inference for a Single Regression Coefficient

1.1.1 Hypothesis Testing

As in SLR, we are interested in testing

$$H_0 : \beta_j = 0 \text{ vs. } H_a : \beta_j \neq 0. \quad (1.1)$$

Test statistic:

$$t_{obs} = \frac{b_j}{SE(b_j)} \quad (1.2)$$

Under H_0 , t_{obs} is an observation from the T distribution with $df = n - p - 1$. This test gives an indication of whether or not the j th predictor, $X_j, j = 1 \dots p$, contributes to the prediction of the response variable **over and above** all the other predictors.

Special case of the Partial F-test with $k = 1$

1.1.2 Confidence Interval

Confidence interval for $\beta_j, j = 1 \dots p$, assuming all the other predictors are in the model, is

$$\beta_j \pm t_{\alpha/2, n-p-1} SE(b_j), \quad (1.3)$$

(i.e. unbiased estimate \pm Margin of Error, where MOE is critical value times std error).

where

- b_j : unbiased estimator of β_j
- $SE(b_j)$: standard error of estimator
- $t_{\alpha/2, n-p-1}$: critical value of $100(1 - \alpha/2)$ th quantile from the T distribution with $df = n - p - 1$.

1.1.3 Global F-test vs. Individual t-tests

- In SLR, these tests are equivalent
- In MLR, the global F-test is designed to test the *overall model*, while the t -tests are designed to test *individual coefficients*.

Case A. If the Global F-test is significant and:

- A.1: All or some of the t -tests are significant, \implies there are some useful explanatory variables for predicting Y .

- A.2: All the t -tests are not significant, \implies this is an indication of "multicollinearity" - i.e., strongly correlated X 's. This implies that individual X 's do not contribute to the prediction of Y over and above other X 's.

Case B. If the Global F-test is NOT significant and:

- B.1: All the t -tests are not significant, \implies none of the X 's contributes to the prediction of Y .
- B.2: Some of the t -tests are significant, \implies
 - The model has no predictive ability. Likely, if there are many predictors, there are type I errors in the t -tests.
 - The predictors are poorly chosen. The contribution of one useful predictor among many poor ones may not be enough for the model (Global F-test) to be significant.

1.2 Multicollinearity

Definition 1.1. Multicollinearity occurs when explanatory variables are highly correlated.

In this case, it is difficult to measure the individual influence of one of the predictors on the response.

- The fitted equation is unstable
- The estimated regression coefficients vary widely from data set to data set (even if the data sets are similar) and depending on which predictor is included in the model.
- The estimated regression coefficients may even have opposite sign than what is expected (*e.g. Simpson's Paradox*).

Remark 1.1. When some X 's are perfectly correlated, we cannot estimate β because $X'X$ is singular. Even if $X'X$ is close to singular, its determinant will be close to zero and the standard errors of estimated coefficients will be large.

Remark 1.2. For the general multiple regression model

$$Y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_p x_p + e \quad (1.4)$$

$$Var(\hat{\beta}_j) = \frac{1}{1 - R_j^2} \times \frac{\sigma^2}{(n - 1)S_{x_j}^2}, j = 1 \dots p \quad (1.5)$$

where R_j^2 is the value of R^2 from the regression of x_j on the other x 's.

The j th Variance Inflation Factor (VIF) = $\frac{1}{1 - R_j^2}$

A commonly used cut-off is 5.

1.3 Diagnostics and Remedies

1.3.1 Leverage and Influential Points

Definition 1.2 (Identifying Leverage Points). Classify the i th point as a point of high leverage (i.e. a lvg point) in MLR model with p predictors if

$$h_{ii} > 2 \times \text{average}(h_{ii}) = 2 \times \frac{p+1}{n} \quad (1.6)$$

■ In SLR, $p = 1$, $h_{ii} > 2 \left(\frac{2}{n}\right) = \frac{4}{n}$.

Remark 1.3. When a **valid** model has been fit, a plot of standardized residuals against any predictor or any linear combination of the predictors (e.g. the fitted values) will have the following features:

1. A random scatter of points around the horizontal axis, since the mean function of e_i is zero when a *correct model has been fit* (linearity)
2. Constant variability as we look along the horizontal axis, i.e.

$$\text{Var}(e) = \sigma^2 I = \begin{pmatrix} \sigma^2 & 0 & \dots & 0 \\ 0 & \sigma^2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \sigma^2 \end{pmatrix} \quad (1.7)$$

Hence, **any pattern** in plot of standardized residuals is indicative that an **invalid** model has been fit to the data. Any nonrandom pattern itself does not provide direct information on how the model is misspecified.

1.3.2 The Box-Cox Transformation

Definition 1.3. The Box-Cox transformation is a general method for transforming a strictly positive (response or predictor) variable.

■ It aims to find transformation that makes the transformed variable close to normally distributed.

It considers a family of *power transformations*.

Suppose the power to be λ :

- $\lambda = 0$: Natural log
- $\lambda = 1$: No transformation
- $\lambda = 0.5$: Square root transformation
- $\lambda = -1$: Inverse transformation

■ It is based on maximizing a likelihood function.

1.4 Added Variable Plot (Need to Review Lecture Part 3)

Definition 1.4. Suppose our current model is

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e} \quad (\text{model } YX) \quad (1.8)$$

and we are considering the introduction of an additional predictor variable \mathbf{Z} , that is, our new model is

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\boldsymbol{\alpha} + \mathbf{e} \quad (\text{model } YXZ) \quad (1.9)$$

The added-variable plot is obtained by plotting the residuals from model YX against the residuals from the model

$$\mathbf{Z} = \mathbf{X}\boldsymbol{\delta} + \mathbf{e} \quad (\text{model } ZX) \quad (1.10)$$

Remark 1.4 (Why Added Variable Plot). • To visually assess the effect of each predictor, having adjusted for the effects of the other predictors

- To visually estimate α
- Can be used to identify points which have undue influence on the least squares estimate of α

1.5 Data Analysis Flow

