

Lec 12: Wavelet Thresholding

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Last lecture: for nonparametric regression, we may find basis functions $\{\phi_i(x)\}_{i=1}^m$ such that $f(x) \approx \sum_{i=1}^m \theta_i \phi_i(x)$
(choice of basis: polynomials, splines, ...)

This lecture: how about orthonormal basis?

$$\left(\int \phi_i(x) \phi_j(x) dx = \begin{cases} 0 & \text{if } i \neq j \\ 1 & \text{if } i = j \end{cases} \right)$$

Gaussian sequence model: throughout this lecture we assume that $x_i = \frac{i}{n}$. $y_i \sim N(f(x_i), \sigma^2)$.

Let $\{\phi_i(x)\}_{i=1}^\infty$ be a **complete** orthonormal basis of $L_2[0,1]$, i.e.

$$f(x) = \sum_{i=1}^\infty \theta_i \phi_i(x).$$

where

$$\theta_i = \int_0^1 f(x) \phi_i(x) dx.$$

(Pf :

$$\begin{aligned} \int_0^1 f(x) \phi_i(x) dx &= \int_0^1 \left(\sum_{j=1}^\infty \theta_j \phi_j(x) \right) \phi_i(x) dx \\ &= \sum_{j=1}^\infty \theta_j \int_0^1 \phi_i(x) \phi_j(x) dx \\ &= \sum_{j=1}^\infty \begin{cases} \theta_j & \text{if } j=i \\ 0 & \text{if } j \neq i \end{cases} = \theta_i \quad \square \end{aligned}$$

How to estimate θ_i ? Try

$$\begin{aligned} z_i &= \frac{1}{n} \sum_{j=1}^n \phi_i(x_j) y_j = \frac{1}{n} \sum_{j=1}^n \phi_i(x_j) (f(x_j) + \sigma_0 \mathfrak{z}_j) \quad (\mathfrak{z}_j \sim N(0,1)) \\ &= \frac{1}{n} \sum_{j=1}^n \phi_i(x_j) f(x_j) + \sigma_0 \cdot \frac{1}{n} \sum_{j=1}^n \phi_i(x_j) \mathfrak{z}_j \\ &\approx \int_0^1 \phi_i(x) f(x) dx + \sigma_0 \cdot N(0, \frac{1}{n} \int_0^1 \phi_i(x)^2 dx) \\ &= \theta_i + N(0, \frac{\sigma_0^2}{n}) \end{aligned}$$

This approx. error is often negligible; we do not consider it in this lecture.

Therefore, instead of observing $(x_i, y_i)_{i=1}^n$, we can equivalently assume that we observe (z_1, z_2, \dots) s.t.

$$z_i \overset{i.i.d.}{\sim} N(\theta_i, \frac{\sigma^2}{n}), \quad i=1, 2, \dots$$

Remark: 1) z_i 's are (approx.) independent as

$$\begin{aligned} \text{Cov}(z_i, z_j) &= \text{Cov}\left(\frac{1}{n} \sum_{k=1}^n \phi_i(x_k) y_k, \frac{1}{n} \sum_{k=1}^n \phi_j(x_k) y_k\right) \\ &= \frac{\sigma^2}{n^2} \sum_{k=1}^n \phi_i(x_k) \phi_j(x_k) \\ &\approx \frac{\sigma^2}{n} \int_0^1 \phi_i(x) \phi_j(x) dx = 0 \quad \text{for } i \neq j; \end{aligned}$$

2) if we find an estimator $\hat{\theta} = (\hat{\theta}_1, \hat{\theta}_2, \dots)$ for θ , then we should estimate f by

$$\begin{aligned} \hat{f}(x) &= \sum_{i=1}^{\infty} \hat{\theta}_i \phi_i(x), \\ \text{and } \|\hat{f} - f\|_2^2 &= \int_0^1 (\hat{f}(x) - f(x))^2 dx \\ &= \int_0^1 \left(\sum_{i=1}^{\infty} (\hat{\theta}_i - \theta_i) \phi_i(x) \right)^2 dx \\ &= \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} (\hat{\theta}_i - \theta_i)(\hat{\theta}_j - \theta_j) \underbrace{\int_0^1 \phi_i(x) \phi_j(x) dx}_{=1(i=j)} \\ &= \sum_{i=1}^{\infty} (\hat{\theta}_i - \theta_i)^2 \\ &= \|\hat{\theta} - \theta\|_2^2 \quad (\text{Plancherel/Parseval identity}) \end{aligned}$$

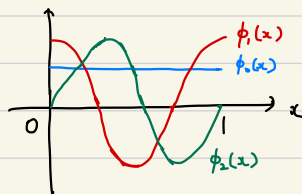
(estimation of $f \iff$ estimation of θ)

Choice I of $\{\phi_i(x)\}_{i=1}^{\infty}$: Fourier basis

Fourier basis of $L_2[0,1]$:

$$\phi_0(x) = 1, \quad \phi_{2j-1}(x) = \sqrt{2} \cos(2\pi j t), \quad \phi_{2j}(x) = \sqrt{2} \sin(2\pi j t)$$

One can check $\{\phi_i(x)\}_{i=0}^{\infty}$ are indeed orthonormal.



Here, $\theta_i = \int_0^1 f(x) \phi_i(x) dx$: Fourier coefficients of f
 $z_i = \frac{1}{n} \sum_{j=1}^n y_j \phi_i(x_j)$: discrete Fourier transform of $(y_i)_{i=1}^n$

Estimation in Sobolev space.

$$H^k(L) = \left\{ f \in L^2[0,1] : \int_0^1 |f^{(k)}(x)|^2 dx \leq L^2 \right\}$$

↓
average notion of smoothness

Theorem. $f \in H^k(L)$ if and only if its Fourier coefficients $(\theta_0, \theta_1, \dots)$ satisfies

$$\sum_{j=1}^{\infty} (2\pi j)^{2k} (\theta_{2j-1}^2 + \theta_{2j}^2) \leq L^2$$

Intuition: smoothness in time domain \Leftrightarrow tail in frequency domain

Estimator I: Fourier projection estimator.

$$\hat{\theta}_i = \begin{cases} z_i & \text{if } i \leq m \\ 0 & \text{if } i > m \end{cases}$$

$$\begin{aligned} \text{Analysis: } \mathbb{E} \|\hat{\theta} - \theta\|_2^2 &= \sum_{i=0}^m \mathbb{E} (z_i - \theta_i)^2 + \sum_{i>m} (0 - \theta_i)^2 \\ &= (m+1) \frac{\sigma_0^2}{n} + \sum_{i>m} \theta_i^2 \\ &\leq \frac{(m+1)\sigma_0^2}{n} + \frac{1}{(\pi m)^{2k}} \underbrace{\sum_{i>m} (\pi i)^{2k} \theta_i^2}_{\leq L^2} \\ &= O\left(\frac{m}{n} + \frac{1}{m^{2k}}\right) \end{aligned}$$

Choosing $m \asymp n^{\frac{1}{2k+1}}$ gives (bias, var)

$$\mathbb{E} \|\hat{f} - f\|_2^2 = \mathbb{E} \|\hat{\theta} - \theta\|_2^2 = O\left(n^{-\frac{2k}{2k+1}}\right).$$

Estimator II: optimal linear estimator (optional)

Set $\hat{\theta}_i = c_i z_i$ with $c_i \in [0, 1]$

$$\begin{aligned} \text{Then } \mathbb{E} \|\hat{\theta} - \theta\|_2^2 &= \sum_{i=0}^{\infty} \mathbb{E} (c_i z_i - \theta_i)^2 \\ &= \sum_{i=0}^{\infty} [(1-c_i)^2 \theta_i^2 + c_i^2 \cdot \frac{\sigma_0^2}{n}] \end{aligned}$$

Choose $\{c_i\}_{i=0}^{\infty}$ to solve the following min-max program:

$$\min_{\{c_i\}_{i=0}^{\infty}} \max_{\{\theta_i\}_{i=0}^{\infty} : \sum_{j=1}^{\infty} (2\pi j)^{2h} (\theta_{2j-1}^2 + \theta_{2j}^2) \leq L^2} \sum_{i=0}^{\infty} [(1-c_i)^2 \theta_i^2 + c_i^2 \frac{\sigma_0^2}{n}]$$

Pinsker's Theorem: the above min-max program exhibits an explicit solution, and the resulting estimator attains $(1+o(1)) \cdot \text{minimax risk}$.

Problem with Fourier:

1) estimators become suboptimal for other Sobolev balls

$$W^{k,p}(L) = \{f \in L^p[0,1] : \int_0^1 |f^{(k)}(x)|^p dx \leq L^p\}$$

for large p , or general Besov balls;

2) estimators become suboptimal when f has spatial inhomogeneity;

3) any linear estimator suffers from the same problem.

Solution: wavelets!

Choice II of $\{\phi_i(x)\}_{i=1}^{\infty}$: Wavelets

Definition: Idea: multiresolutional analysis

A wavelet basis consists of a father wavelet $\phi(x)$ and a mother wavelet $\psi(x)$ on $[0,1]$, s.t. if

$$V_j = \text{span} \{ \phi_{j,k}(x) = 2^{j/2} \phi(2^j x - k) : 0 \leq k \leq 2^j - 1 \}$$

$$W_j = \text{span} \{ \psi_{j,k}(x) = 2^{j/2} \psi(2^j x - k) : 0 \leq k \leq 2^j - 1 \}$$

then :

$$1) V_{j+1} = V_j \oplus W_j \quad (\Rightarrow V_{j+s} = V_j \oplus W_{j+1} \oplus W_{j+2} \oplus \dots \oplus W_{j+s-1})$$

$$2) L^2[0,1] = \overline{V_{j_0} \oplus W_{j_0} \oplus W_{j_0+1} \oplus \dots} \quad (\text{spans all function on } [0,1])$$

$$3) \{ \phi_{j_0,k} : 0 \leq k \leq 2^{j_0} - 1 \} \text{ \& \; } \{ \psi_{j,k} : j \geq j_0, 0 \leq k \leq 2^j - 1 \} \text{ are orthonormal.}$$

By 2) & 3), any $f \in L^2[0,1]$ can be written as

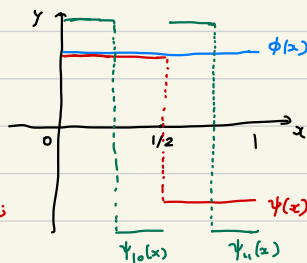
$$f(x) = \underbrace{\sum_{k=0}^{2^{j_0}-1} \alpha_{j_0,k} \phi_{j_0,k}(x)}_{\text{Gross information}} + \sum_{j \geq j_0} \underbrace{\sum_{k=0}^{2^j-1} \beta_{j,k} \psi_{j,k}(x)}_{\text{Detailed information at level } j}$$

Example of wavelets:

Haar wavelets. $\phi(x) = 1(x \in [0,1])$

$$\psi(x) = \begin{cases} 1, & \text{if } x \in [0, \frac{1}{2}) \\ -1, & \text{if } x \in [\frac{1}{2}, 1] \end{cases}$$

(j : resolution parameter, $|\text{supp}(\phi_{j,k})| = |\text{supp}(\psi_{j,k})| = 2^{-j}$;
 $k \in \{0, 1, \dots, 2^j - 1\}$: spatial location)



Meyer wavelets. All moments vanish, but infinite support in time domain

Daubchies wavelets. Vanishing moments up to desired order + compactly supported
 (most widely used wavelets)

Cohen-Daubechies-Fearn wavelet: used in JPEG 2000 standard.

Estimation: wavelet thresholding

Soft and hard thresholding: consider Gaussian sequence model

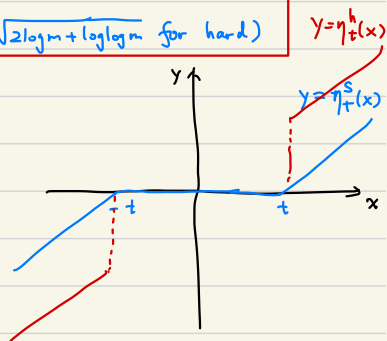
$$z_i \sim N(\theta_i, \sigma^2), \quad i=1, \dots, m.$$

Soft-thresholding estimator: $\hat{\theta}_i^s = \eta_t^s(z_i) = \text{sign}(z_i) \cdot (|z_i| - t)_+$

Hard-thresholding estimator: $\hat{\theta}_i^h = \eta_t^h(z_i) = z_i \cdot \mathbf{1}(|z_i| \geq t)$

(choice of thresholds: $t = \sigma\sqrt{2\log m}$ for soft, $t = \sigma\sqrt{2\log m + \log\log m}$ for hard)

Intuition: when z is small, think of $\theta \approx 0$,
when z is large, think of $\theta \approx z$.



Property (pf omitted): thresholding estimators are optimal when θ has "sparse" structures

Wavelet thresholding: choose $j_0 \sim 1$, $j_1 \sim \log n$, use wavelet transform to obtain

$$\begin{cases} \tilde{\alpha}_{j_0 k} \sim N(\alpha_{j_0 k}, \frac{\sigma_\alpha^2}{n}), & 0 \leq k \leq 2^{j_0} - 1 \\ \hat{\beta}_{j_1 k} \sim N(\beta_{j_1 k}, \frac{\sigma_\beta^2}{n}), & j_0 \leq j_1 < j_1, \quad 0 \leq k \leq 2^{j_1} - 1. \end{cases}$$

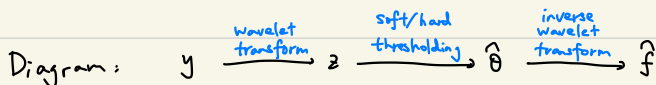
Wavelet thresholding estimator:

$$\tilde{\alpha} = \hat{\alpha}, \quad \tilde{\beta} = \eta_t^s(\hat{\beta}) \text{ or } \eta_t^h(\hat{\beta}),$$

and estimate f by

$$\tilde{f}(x) = \sum_{k=0}^{2^{j_0}-1} \tilde{\alpha}_{j_0 k} \phi_{j_0 k}(x) + \sum_{j=j_0}^{j_1-1} \sum_{k=0}^{2^j-1} \tilde{\beta}_{j k} \psi_{j k}(x)$$

(inverse wavelet transform)



Choice of threshold t :

Option I: estimate noise level σ_0 , use the theory prediction (VisuShrink)

Option II: use cross validation or unbiased risk estimate to choose t
(SureShrink and others)

Properties: 1) optimal in all Sobolev and Besov classes;

2) adaptive to smoothness and local inhomogeneity of f ;
(non-linearity of estimator plays a key role here)

3) easy to implement using fast wavelet transforms.

Why wavelet thresholding (option-I)?

1) why wavelets? \rightarrow representation power of wavelets

Wavelet is an unconditional basis for Sobolev or Besov norms $\|\cdot\|$,
i.e. for every $\varepsilon_i \in [-1, 1]$,

$$\left\| \sum_i \varepsilon_i \phi_i \right\| \leq C \left\| \sum_i \phi_i \right\|$$

(Fourier basis does not satisfy this property)

2) why thresholding? \rightarrow idea of "shrinkage"

Similar to the James-Stein estimator, thresholding introduces a small bias to significantly reduce the variance

(HW 8: mimic the performance of the "ideal truncation estimator")

$$\hat{\theta}_i^{\text{ITE}} = z_i \cdot 1(|\theta_i| \geq \sigma) \quad)$$