

## Lec 6: Missing Data & EM Algorithm

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## Missing data in exponential families.

$(x_1, y_1), \dots, (x_n, y_n) \stackrel{i.i.d.}{\sim} p_\theta(x, y) = \exp(\langle \theta, T(x, y) \rangle - A(\theta)) h(x, y)$   
with  $(x_1, \dots, x_n)$ : **unobserved** variables  $(y_1, \dots, y_n)$ : observed variables

Goal: Find the MLE for  $\theta$ .

Incomplete log-likelihood:

$$\begin{aligned} \ell_\theta(y_1, \dots, y_n) &= \sum_{i=1}^n \log p_\theta(y_i) \\ &= \sum_{i=1}^n \log \int_{\mathcal{X}} p_\theta(x, y_i) dx \\ &= \sum_{i=1}^n \log \int_{\mathcal{X}} \exp(\langle \theta, T(x, y_i) \rangle - A(\theta)) h(x, y_i) dx \\ &= \sum_{i=1}^n \underbrace{\left( \log \int_{\mathcal{X}} \exp(\langle \theta, T(x, y_i) \rangle) h(x, y_i) dx \right)}_{=: A_{y_i}(\theta)} - A(\theta) \end{aligned}$$

The conditional distribution  $p_\theta(x|y_i)$  also belongs to an exponential family, with log-partition function  $A_{y_i}(\theta) \Rightarrow A_{y_i}(\theta)$  is convex in  $\theta$

$\ell_\theta(y_1, \dots, y_n) = \sum_{i=1}^n (A_{y_i}(\theta) - A(\theta))$  is the difference of two convex functions, which may not be concave in  $\theta$ !!

## Detour: a short introduction of convex duality

Def (convex conjugate) The convex conjugate of a function  $f$  on  $\mathbb{R}^d$  is

$$f^*(t) = \max_{x \in \mathbb{R}^d} \langle t, x \rangle - f(x)$$

Properties. 1) The maximizer  $x^* = (\nabla f)^{-1}(t) = \nabla f^*(t)$ , for convex  $f$ .

Pf. differentiation gives  $t = \nabla f(x^*) \Rightarrow x^* = (\nabla f)^{-1}(t)$ .

$$\begin{aligned} \nabla f^*(t) &= \nabla \left( \max_{x \in \mathbb{R}^d} \langle t, x \rangle - f(x) \right) \\ &= \left\{ \nabla_t (\langle t, x^* \rangle - f(x^*)) : x^* \in \operatorname{argmax}_x \langle t, x \rangle - f(x) \right\} = x^* \end{aligned}$$

2)  $f^*(t)$  is convex in  $t$

Pf. Because  $f^*(t)$  is a maximum over linear functions of  $t$ .

3) For convex  $f$ ,  $f(x) = \max_{t \in \mathbb{R}^d} \langle x, t \rangle - f^*(t)$  (in other words,  $f^{**} = f$ )

Pf. Definition of  $f^*$   $\Rightarrow f^*(t) + f(x) \geq \langle x, t \rangle \quad \forall x, t$   
 $\Rightarrow f(x) \geq \max_t \langle x, t \rangle - f^*(t)$ .

Property 1)  $\Rightarrow f^*(t) = \langle t, (\nabla f)^{-1}(t) \rangle - f((\nabla f)^{-1}(t))$

$\Rightarrow f^*(\nabla f(x)) = \langle \nabla f(x), x \rangle - f(x)$

$\Rightarrow f(x) = \langle \nabla f(x), x \rangle - f^*(\nabla f(x)) \leq \max_t \langle t, x \rangle - f^*(t)$

EM algorithm. Using convex duality,

$$\begin{aligned} \max_{\theta} \ell_{\theta}(\gamma_1, \dots, \gamma_n) &= \max_{\theta} \sum_{i=1}^n (A_{\gamma_i}(\theta) - A(\theta)) \\ &= \max_{\theta} \sum_{i=1}^n \left( \max_{\mu_i} \langle \mu_i, \theta \rangle - A_{\gamma_i}^*(\mu_i) - A(\theta) \right) \\ &= \max_{\theta} \max_{\mu_1, \dots, \mu_n} \underbrace{\sum_{i=1}^n (\langle \mu_i, \theta \rangle - A_{\gamma_i}^*(\mu_i) - A(\theta))}_{=: f(\theta, \mu_1, \dots, \mu_n)} \end{aligned}$$

Key intuitions: 1. for fixed  $\theta$ ,  $f(\theta, \mu_1, \dots, \mu_n)$  is concave in  $(\mu_1, \dots, \mu_n)$

2. for fixed  $(\mu_1, \dots, \mu_n)$ ,  $f(\theta, \mu_1, \dots, \mu_n)$  is concave in  $\theta$

(Warning:  $f(\theta, \mu_1, \dots, \mu_n)$  is NOT "jointly" concave in  $\theta$  and  $(\mu_1, \dots, \mu_n)$ !)

Idea. Successive maximization starting from some  $\theta^{(0)}$ :  $\theta^{(0)} \rightarrow \mu^{(1)} \rightarrow \theta^{(1)} \rightarrow \mu^{(2)} \rightarrow \dots$

1) E-step: fix  $\theta^{(t)}$ , find the maximizer  $\mu^{(t+1)}$

$$\mu_i^{(t+1)} = \nabla A_{\gamma_i}(\theta^{(t)}) = \mathbb{E}_{x \sim p_{\theta^{(t)}}(x|y_i)} [T(x, y_i)] \quad (\text{"expectation" step})$$

2) M-step: fix  $\mu^{(t+1)}$ , find the maximizer  $\theta^{(t+1)}$

$$\nabla A(\theta^{(t+1)}) = \frac{1}{n} \sum_{i=1}^n \mu_i^{(t+1)} \quad (\text{"maximization" step})$$

### EM Intuition:

1. E-step: for each sample  $i$  with missing data  $x_i$ , think of a fake  $\tilde{x}_i \sim p_\theta(x_i | y_i)$  and compute sufficient statistic  $\mu_i = \mathbb{E}[T(\tilde{x}_i, y_i)]$ .
2. M-step: aggregate the sufficient statistics "as if" there were no missing data problem:  $\frac{1}{n} \sum_{i=1}^n \mu_i = \nabla A(\theta)$
3. Iterate the above process.

### Example: Gaussian mixture model

$$p_\theta(x, y): \quad P(x = j) = \pi_j, \quad j = 1, 2, \dots, k$$

$$y | x = j \sim N(\mu_j, 1), \quad j = 1, 2, \dots, k.$$

$$\text{Unknown parameter: } \theta = (\pi_1, \dots, \pi_k, \mu_1, \dots, \mu_k)$$

$$\text{Unobserved variable: } x_1, \dots, x_n$$

$$\text{Observed variable: } y_1, \dots, y_n.$$

$\varphi$ : pdf of  $N(0, 1)$

E-step: Given  $\theta = \theta^{(t)}$ , understand  $p_\theta(x | y)$ :

$$p_\theta(x = j | y) = \frac{p_\theta(x = j, y)}{p_\theta(y)} = \frac{p_\theta(y | x = j) p_\theta(x = j)}{\sum_{i=1}^k p_\theta(y | x = i) p_\theta(x = i)} = \frac{\pi_j \varphi(y - \mu_j)}{\sum_i \pi_i \varphi(y - \mu_i)}$$

$$\Rightarrow p_{i,j}^{(t+1)} := p_{\theta^{(t+1)}}(x_i = j | y_i) = \frac{\pi_j^{(t)} \varphi(y_i - \mu_j^{(t)})}{\sum_{\ell=1}^k \pi_\ell^{(t)} \varphi(y_i - \mu_\ell^{(t)})}, \quad \begin{matrix} i = 1, \dots, n \\ j = 1, \dots, k \end{matrix}$$

M-step: Pretend that  $x_i \sim p_{\theta^{(t+1)}}(\cdot | y_i)$ , maximize the log-likelihood

$$\begin{aligned} \sum_{i=1}^n \mathbb{E}_{x_i \sim p_{\theta^{(t+1)}}(\cdot | y_i)} [\log p_\theta(x_i, y_i)] &= \sum_{i=1}^n \sum_{j=1}^k p_{i,j}^{(t+1)} \log p_\theta(x_i = j, y_i) \\ &= \sum_{i=1}^n \sum_{j=1}^k p_{i,j}^{(t+1)} \left( \log \pi_j - \frac{(y_i - \mu_j)^2}{2} - \log \sqrt{2\pi} \right) \\ &= \sum_{j=1}^k \left[ \left( \sum_{i=1}^n p_{i,j}^{(t+1)} \right) \log \pi_j - \frac{1}{2} \sum_{i=1}^n (y_i - \mu_j)^2 p_{i,j}^{(t+1)} \right] + \text{Const.} \end{aligned}$$

Carrying out the maximization over  $\theta = (\pi_1, \dots, \pi_k, \mu_1, \dots, \mu_k)$  gives

$$\begin{cases} \pi_j^{(t+1)} = \frac{1}{n} \sum_{i=1}^n p_{i,j}^{(t+1)} \\ \mu_j^{(t+1)} = \frac{\sum_{i=1}^n p_{i,j}^{(t+1)} y_i}{\sum_{i=1}^n p_{i,j}^{(t+1)}} \end{cases}$$

## General EM via evidence lower bound

Def. For probability distributions  $P, Q$  over  $X$ , the Kullback-Leibler (KL) divergence is

$$D_{KL}(P \parallel Q) = \begin{cases} \sum_{x \in X} p(x) \log \frac{p(x)}{q(x)}, & \text{for pmfs} \\ \int_X p(x) \log \frac{p(x)}{q(x)} dx, & \text{for pdfs} \end{cases}$$

Thm.  $D_{KL}(P \parallel Q) \geq 0$ .

Pf. Since  $\log t \geq 1 - \frac{1}{t}$  for all  $t \geq 0$ ,

$$\begin{aligned} D_{KL}(P \parallel Q) &= \sum_x p(x) \log \frac{p(x)}{q(x)} \geq \sum_x p(x) \left(1 - \frac{q(x)}{p(x)}\right) \\ &= \sum_x p(x) - \sum_x q(x) = 1 - 1 = 0 \end{aligned}$$

## Evidence lower bound (ELBO)

$$\log p_\theta(y^n) = \max_{q(\cdot)} \underbrace{\mathbb{E}_{x^n \sim q(\cdot)} \left[ \log \frac{p_\theta(x^n, y^n)}{q(x^n)} \right]}_{\text{ELBO}}$$

$$\begin{aligned} \text{Pf. } \log p_\theta(y^n) - \text{ELBO} &= \mathbb{E}_{x^n \sim q(\cdot)} \left[ \log \frac{p_\theta(y^n) q(x^n)}{p_\theta(x^n, y^n)} \right] \\ &= \mathbb{E}_{x^n \sim q(\cdot)} \left[ \log \frac{q(x^n)}{p_\theta(x^n | y^n)} \right] = D_{KL}(q(x^n) \parallel p_\theta(x^n | y^n)). \end{aligned}$$

So  $\log p_\theta(y^n) \geq \text{ELBO}$ , with equality iff  $q(x^n) = p_\theta(x^n | y^n)$ .

General EM:  $\max_{\theta} \log p_{\theta}(y^n) = \max_{\theta} \max_{q(x^n)} \mathbb{E}_{x^n \sim q(\cdot)} \left[ \log \frac{p_{\theta}(x^n, y^n)}{q(x^n)} \right]$

- Fix  $\theta$ : the maximizer is  $q^*(x^n) = p_{\theta}(x^n | y^n)$
- Fix  $q$ : solve the maximization  $\theta \mapsto Q(\theta | q) := \mathbb{E}_{x^n \sim q(\cdot)} [\log p_{\theta}(x^n, y^n)]$ , which is often tractable

Example 1: exponential family  $p_{\theta}(x, y) = \exp(\langle \theta, T(x, y) \rangle - A(\theta)) h(x, y)$ .

If  $q(x^n) = p_{\theta^{(t)}}(x^n | y^n)$ , then

$$\begin{aligned} Q(\theta | q) &:= Q(\theta | \theta^{(t)}) = \mathbb{E}_{x^n \sim p_{\theta^{(t)}}(\cdot | y^n)} [\log p_{\theta}(x^n, y^n)] \\ &= \sum_{i=1}^n \langle \theta, \mathbb{E}_{x_i \sim p_{\theta^{(t)}}(\cdot | y_i)} [T(x_i, y_i)] \rangle - n A(\theta) + \text{const} \end{aligned}$$

So maximizing  $\theta \mapsto Q(\theta | \theta^{(t)})$  requires

- evaluation of  $\mathbb{E}_{x_i \sim p_{\theta^{(t)}}(\cdot | y_i)} [T(x_i, y_i)]$  (E-step)
- solving the equation  $\frac{1}{n} \sum_{i=1}^n \mathbb{E}_{x_i \sim p_{\theta^{(t)}}(\cdot | y_i)} [T(x_i, y_i)] = \nabla A(\theta^{(t+1)})$  (M-step)

Example 2: gradient descent

$$Q(\theta | \theta^{(t)}) = \mathbb{E}_{x^n \sim p_{\theta^{(t)}}(\cdot | y^n)} [\log p_{\theta}(x^n, y^n)]$$

GD update for maximizing  $\theta \mapsto Q(\theta | \theta^{(t)})$ :

$$\theta^{(t+1)} = \theta^{(t)} + \alpha \mathbb{E}_{x^n \sim p_{\theta^{(t)}}(\cdot | y^n)} [\nabla_{\theta} \log p_{\theta}(x^n, y^n) |_{\theta = \theta^{(t)}}]$$

## Case study: variational autoencoders (VAE)

Target: given  $y_1, \dots, y_n$  (e.g. images), find  $\theta$  (e.g. params of a deep net) s.t.

$$y|x \sim N(\mu_\theta(x), \sigma_\theta^2(x)I), \text{ with } x \sim N(0, I)$$

(once we learn  $\theta$ , we can generate new images by first sampling  $x \sim N(0, I)$  and then drawing  $y|x \sim N(\mu_\theta(x), \sigma_\theta^2(x)I)$ )

$$\text{MLE: } \max_{\theta} p_{\theta}(y^n) \approx \max_{\theta} \max_{\phi} \mathbb{E}_{x \sim q_{\phi}(\cdot|y^n)} \left[ \log \frac{p_{\theta}(x^n, y^n)}{q_{\phi}(x^n|y^n)} \right]$$

$x|y: \sim N(\mu_{\phi}(y), \sigma_{\phi}^2(y)I)$  parametrized by another neural network  $\phi$

Aim to perform SGD jointly over  $(\theta, \phi)$

$$\begin{aligned} \text{Idea of VAE: } \mathbb{E}_{x \sim q_{\phi}(\cdot|y^n)} \left[ \log \frac{p_{\theta}(x^n, y^n)}{q_{\phi}(x^n|y^n)} \right] \\ = -D_{KL}(q_{\phi}(x^n|y^n) \| p_{\theta}(x^n)) + \mathbb{E}_{x \sim q_{\phi}(\cdot|y^n)} [\log p_{\theta}(y^n|x^n)] \end{aligned}$$

- First term: as  $q_{\phi}(x_i|y_i) = N(\mu_{\phi}(y_i), \sigma_{\phi}^2(y_i)I)$  and  $p_{\theta}(x_i) = N(0, I)$ , the KL divergence has an explicit form in  $(\theta, \phi)$ , so easy to compute the gradient.

- Second term:  $\nabla_{\theta}$ : easy as  $\nabla_{\theta} \log p_{\theta}(y|x)$  is quite simple, and

$$\begin{aligned} \nabla_{\theta} \mathbb{E}_{x \sim q_{\phi}(\cdot|y)} [\log p_{\theta}(y|x)] \\ \approx \nabla_{\theta} \left( \frac{1}{L} \sum_{e=1}^L \log p_{\theta}(y|x_e) \right) \\ = \frac{1}{L} \sum_{e=1}^L \nabla_{\theta} \log p_{\theta}(y|x_e) \quad \text{for } x_1, \dots, x_L \sim q_{\phi}(\cdot|y) \end{aligned}$$

$\nabla_{\phi}$ : 1) Approach 1 (REINFORCE):

$$\begin{aligned} \nabla_{\phi} \mathbb{E}_{x \sim q_{\phi}(\cdot|y)} [f(x)] &= \mathbb{E}_{x \sim q_{\phi}(\cdot|y)} [f(x) \nabla_{\phi} \log q_{\phi}(x|y)] \\ &\approx \frac{1}{L} \sum_{e=1}^L f(x_e) \nabla_{\phi} \log q_{\phi}(x_e|y) \end{aligned}$$

simple expression due to Gaussian

2) Approach II (reparametrization):

$$\begin{aligned}\nabla_{\phi} \mathbb{E}_{x \sim N(\mu_{\phi}(y), \sigma_{\phi}^2(y)I)}[f(x)] &= \nabla_{\phi} \mathbb{E}_{\varepsilon \sim N(0, I)}[f(\mu_{\phi}(y) + \sigma_{\phi}(y)\varepsilon)] \\ &= \mathbb{E}_{\varepsilon \sim N(0, I)}[\nabla_{\phi} f(\mu_{\phi}(y) + \sigma_{\phi}(y)\varepsilon)] \\ &\approx \frac{1}{L} \sum_{\ell=1}^L \nabla_{\phi} f(\mu_{\phi}(y) + \sigma_{\phi}(y)\varepsilon_{\ell}) \\ &\quad \text{for } \varepsilon_1, \dots, \varepsilon_L \sim N(0, I)\end{aligned}$$