Today's lecture

Approximate inference in graphical models.

- Forward and Backward KL divergence
- Variational Inference
- Mean Field: Naive and Structured
- Marginal Polytope
- Local Polytope
- Relaxation methods
- Loopy BP
- LP relaxations for MAP inference

Figures from D. Sontag, Murphy's book

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- Almost all approximate inference algorithms in practice are
 - Variational algorithms (e.g., mean-field, loopy belief propagation)
 - Sampling methods (e.g., Gibbs sampling, MCMC)

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- The KL-divergence is asymmetric

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- \bullet · · · computing $p(\mathbf{x})$ is still hard, due to the partition function
- What can we do?

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Let's look at the unnormalized distribution

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- Since Z is constant, by minimizing J(q), we will force q to become close to p
- The KL is always non-negative, so we see that J(q) is an upper bound on the negative log likelihood (NLL)

$$J(q) = KL(q||p) - \log Z \ge -\log Z = -\log p(\mathcal{D})$$

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$$J(q) = \mathbb{E}_q[\log q(\mathsf{x})] + \mathbb{E}_q[-\log \tilde{p}(\mathsf{x})]$$

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which is the expected energy minus the entropy.

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 This is the expected NLL plus a penalty term that measures how far apart the two distributions are

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$$\arg\min_{q} \mathit{KL}(p||q)$$

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What is the difference between the solution to

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• They differ only when q is minimized over a restricted set of probability distribution $Q = \{q_1, \dots\}$, and $p \neq q$. Why?

Forward or Reverse KL

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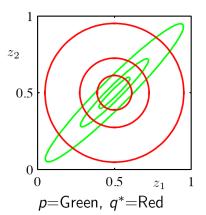
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KL divergence - M projection

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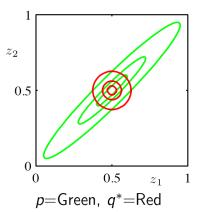
 $p(\mathbf{x})$ is a 2D Gaussian and Q is the set of all Gaussian distributions with diagonal covariance matrices



KL divergence - I projection

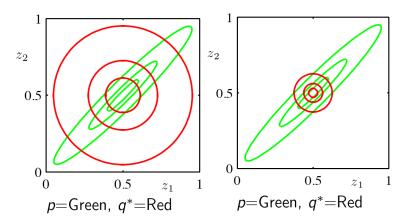
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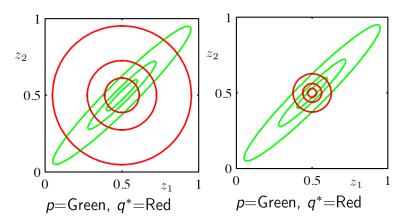
KL Divergence (single Gaussian)

• In this example, both the M-projection and I-projection find an approximate $q(\mathbf{x})$ that has the correct mean (i.e., $\mathbb{E}_p(\mathbf{z}) = \mathbb{E}_q(\mathbf{x})$)



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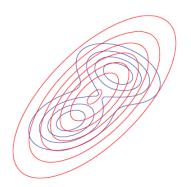


What if $p(\mathbf{x})$ is multimodal?

M projection (Mixture of Gaussians)

$$q^* = arg \min_{q \in Q} KL(p||q) = \sum_{\mathbf{x}} p(\mathbf{x}) \log \frac{p(\mathbf{x})}{q(\mathbf{x})}$$

 $p(\mathbf{x})$ is a mixture of two 2D Gaussians and Q is the set of all 2D Gaussian distributions (with arbitrary covariance matrices)

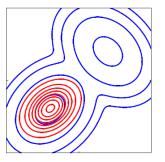


M-projection yields a distribution $q(\mathbf{x})$ with the correct mean and covariance.

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p=Blue, $q^*=$ Red (two local minima!)

The I-projection does not necessarily yield the correct moments

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We can do the maximization one node at a time, in an iterative fashion

• Focus on q_j (holding all other terms constant)

$$L(q_j) = \sum_{\mathbf{x}} \prod_{i} q_i(\mathbf{x}) \left[\log \tilde{p}(\mathbf{x}) - \sum_{k} \log q_k(\mathbf{x}_k) \right]$$

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$$\log f_{i}(\mathbf{x}_{i}) = \sum_{\mathbf{x}_{j}} \prod_{i \neq j} q_{i}(\mathbf{x}_{i}) \log \tilde{p}(\mathbf{x}) = \mathbb{E}_{\mathbf{x}_{j}} [\log \tilde{p}(\mathbf{x})]$$

where

$$\log f_j(\mathbf{x}_j) = \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \log \tilde{p}(\mathbf{x}) = \mathbb{E}_{-q_j}[\log \tilde{p}(\mathbf{x})]$$

• Focus on q_i (holding all other terms constant)

$$\begin{split} L(q_j) &= \sum_{\mathbf{x}} \prod_i q_i(\mathbf{x}) \left[\log \tilde{p}(\mathbf{x}) - \sum_k \log q_k(\mathbf{x}_k) \right] \\ &= \sum_{\mathbf{x}_j} \sum_{\mathbf{x}_{-j}} q_j(\mathbf{x}_j) \prod_{i \neq j} q_i(\mathbf{x}_i) \left[\log \tilde{p}(\mathbf{x}) - \sum_k \log q_k(\mathbf{x}_k) \right] \\ &= \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \log \tilde{p}(\mathbf{x}) - \\ &\sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \sum_{\mathbf{x}_{-j}} \prod_{i \neq j} q_i(\mathbf{x}_i) \left[\sum_{k \neq j} \log q_k(\mathbf{x}_k) + \log q_j(\mathbf{x}_j) \right] \\ &= \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \log f_j(\mathbf{x}_j) - \sum_{\mathbf{x}_j} q_j(\mathbf{x}_j) \log q_j(\mathbf{x}_j) + \text{const} \\ &\log f_j(\mathbf{x}_j) = \sum \prod_i q_i(\mathbf{x}_i) \log \tilde{p}(\mathbf{x}) = \mathbb{E}_{-q_j} [\log \tilde{p}(\mathbf{x})] \end{split}$$

where

• So we average out all the variables except x_i , and can rewrite $L(q_i)$ as

$$L(q_i) = -KL(q_i||f_i)$$

Suppose that we have an arbitrary graphical model

$$p(\mathbf{x}; \theta) = \frac{1}{Z(\theta)} \prod_{c \in C} \phi_c(\mathbf{x}_c) = \exp\left(\sum_{c \in C} \theta_c(\mathbf{x}_c) - \ln Z(\theta)\right)$$

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Mean Field for Variational Inference

$$\max_{q \in \mathcal{Q}} \sum_{c \in \mathcal{C}} \sum_{\mathbf{x}_c} q(\mathbf{x}_c) \theta_c(\mathbf{x}_c) + H(q(\mathbf{x}))$$

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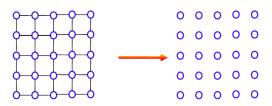
$$\max_{q \in \mathcal{Q}} \sum_{c \in \mathcal{C}} \sum_{\mathbf{x}_c} q(\mathbf{x}_c) \theta_c(\mathbf{x}_c) + H(q(\mathbf{x}))$$

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- Although this function is concave and thus in theory should be easy to optimize, we need some compact way of representing $q(\mathbf{x})$
- Mean field: assume a factored representation of the joint distribution



$$q(\mathbf{x}) = \prod_{i \in V} q_i(x_i)$$

This is called "naive" mean field

 Suppose that Q consists of all fully factorized distributions, then we can simplify

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subject to the constraints

$$q_i(x_i) \geq 0 \quad \forall i \in V, x_i$$

$$\sum_{x_i} q_i(x_i) = 1 \qquad \forall i \in V$$

For pairwise MRFs we have

$$\max_{q} \sum_{ij \in E} \sum_{x_i, x_j} \theta_{ij}(x_i, x_j) q_i(x_i) q_j(x_j) - \sum_{i \in V} \sum_{x_i} q_i(x_i) \ln q_i(x_i)$$
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$$q_i(x_i) \leftarrow \frac{1}{Z_i} \exp \left\{ \theta_i(x_i) + \sum_{j \in N(i)} \sum_{x_j} q_j(x_j) \theta_{ij}(x_i, x_j) \right\}$$

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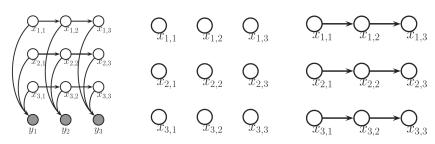
• See Mean field example for the Ising Model, Murphy 21.3.2

Structured mean-field approximations

• Rather than assuming a fully-factored distribution for q, we can use a structured approximation, such as a spanning tree

Structured mean-field approximations

- Rather than assuming a fully-factored distribution for q, we can use a structured approximation, such as a spanning tree
- For example, for a factorial HMM, a good approximation may be a product of chain-structured models (see Murphy 21.4.1)



Approximate Inference via Loopy BP

- Mean field inference approximates posterior as product of marginal distributions
- Allows use of different forms for each variable: useful when inferring statistical parameters of models, or regression weights
- An alternative approximate inference algorithm is loopy belief propagation
- Same algorithm shown to do exact inference in trees last class
- In loopy graphs, BP not guaranteed to give correct results, may not converge, but often works well in practice

Loopy BP on Pairwise Models

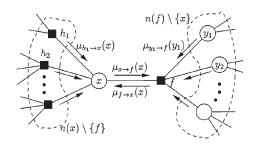
Algorithm 22.1: Loopy belief propagation for a pairwise MRF

- 1 Input: node potentials $\psi_s(x_s)$, edge potentials $\psi_{st}(x_s, x_t)$;
- 2 Initialize messages $m_{s\to t}(x_t)=1$ for all edges s-t;
- 3 Initialize beliefs $bel_s(x_s) = 1$ for all nodes s;
- 4 repeat
- 5 Send message on each edge

$$m_{s \to t}(x_t) = \sum_{x_s} \left(\psi_s(x_s) \psi_{st}(x_s, x_t) \prod_{u \in \text{nbr}_s \setminus t} m_{u \to s}(x_s) \right);$$

- 6 Update belief of each node $\operatorname{bel}_s(x_s) \propto \psi_s(x_s) \prod_{t \in \operatorname{nbr}_s} m_{t \to s}(x_s);$
- 7 **until** beliefs don't change significantly;
- 8 Return marginal beliefs $bel_s(x_s)$;

Loopy BP for Factor Graph



$$m_{i \to f}(x_i) = \prod_{h \in M(i) \setminus f} m_{h \to i}(x_i)$$

$$m_{f \to i}(x_i) = \sum_{\mathbf{x}_c \setminus x_i} f(\mathbf{x}_c) \prod_{j \in N(f) \setminus i} m_{j \to f}(x_j)$$

$$\mu_i(x_i) \propto \prod_{f \in M(i)} m_{f \to i}(x_i)$$

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- Change from synchronous to asynchronous updates
 - Update sets of nodes at a time, e.g., spanning trees (tree reparameterization)

LBP as Variational Inference

 More theoretical analysis of LBP from variational point of view: (Wainwright & Jordan, 2008)

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LBP as Variational Inference

- More theoretical analysis of LBP from variational point of view: (Wainwright & Jordan, 2008)
- Dense tome
- Simplify by considering pairwise UGMs, discrete variables

Variational Inference for Graphical Models

Suppose that we have an arbitrary graphical model

$$p(\mathbf{x}; \theta) = \frac{1}{Z(\theta)} \prod_{c \in C} \phi_c(\mathbf{x}_c) = \exp\left(\sum_{c \in C} \theta_c(\mathbf{x}_c) - \ln Z(\theta)\right)$$

We can compute the KL

$$\begin{aligned} KL(q||p) &= \sum_{\mathbf{x}} q(\mathbf{x}) \ln \frac{q(\mathbf{x})}{p(\mathbf{x})} \\ &= -\sum_{\mathbf{x}} q(\mathbf{x}) \ln p(\mathbf{x}) - \sum_{\mathbf{x}} q(\mathbf{x}) \ln \frac{1}{q(\mathbf{x})} \\ &= -\sum_{\mathbf{x}} q(\mathbf{x}) \left(\sum_{c \in C} \theta_c(\mathbf{x}_c) - \ln Z(\theta) \right) - H(q(\mathbf{x})) \\ &= -\sum_{c \in C} \sum_{\mathbf{x}} q(\mathbf{x}) \theta_c(\mathbf{x}_c) + \sum_{\mathbf{x}} q(\mathbf{x}) \ln Z(\theta) - H(q(\mathbf{x})) \\ &= -\sum_{c \in C} \mathbb{E}_q[\theta_c(\mathbf{x}_c)] + \ln Z(\theta) - H(q(\mathbf{x})) \end{aligned}$$

• The partition function is a constant when minimizing over q

• Since $KL(q||p) \ge 0$ we have

$$-\sum_{c \in C} \mathbb{E}_q[\theta_c(\mathbf{x}_c)] + \ln Z(\theta) - H(q(\mathbf{x})) \ge 0$$

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which implies

$$\ln Z(heta) \geq \sum_{c \in C} \mathbb{E}_q[heta_c(\mathbf{x}_c)] + H(q(\mathbf{x}))$$

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This casts exact inference as a variational optimization problem

$$\ln Z(\theta) = \max_{q} \sum_{c \in C} \mathbb{E}_{q}[\theta_{c}(\mathbf{x}_{c})] + H(q(\mathbf{x}))$$

$$\begin{aligned} & \ln Z(\theta) &= & \max_{q} \sum_{c \in \mathcal{C}} \mathbb{E}_{q}[\theta_{c}(\mathbf{x}_{c})] + H(q(\mathbf{x})) \\ &= & \max_{q} \sum_{c \in \mathcal{C}} \sum_{\mathbf{x}} q(\mathbf{x})\theta_{c}(\mathbf{x}_{c}) + H(q(\mathbf{x})) \end{aligned}$$

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• Assume that p(x) is in the exponential family, and let f(x) be its sufficient statistic vector

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where M is the marginal polytope, having all valid marginal vectors

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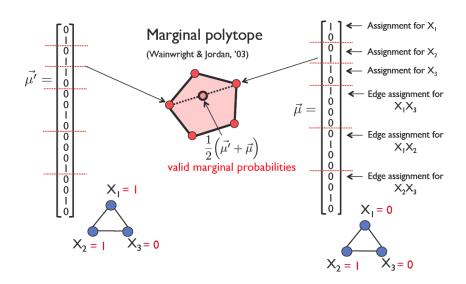
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- ullet For a discrete-variable MRF, the sufficient statistic vector $\mathbf{f}(\mathbf{x})$ is simply the concatenation of indicator functions for each clique of variables that appear together in a potential function
- For example, if we have a pairwise MRF on binary variables with m = |V| variables and |E| edges, d = 2m + 4|E|

Marginal Polytope for Discrete MRFs



$$\ln Z(\theta) = \max_{\mu \in M} \sum_{c \in C} \sum_{\mathbf{x}_c} \theta_c(\mathbf{x}_c) \mu_c(\mathbf{x}_c) + H(\mu)$$

We still haven't achieved anything, because:

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- We replace $H(\mu)$ with a function $\tilde{H}(\mu)$ which approximates $H(\mu)$

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- The μ_i and μ_{ii} are called pseudo-marginals

polytope for a tree-structured MRF, and the pseudomarginals are the marginals. marginal polytope, i.e., $M \subseteq M_L$

Mean-field vs relaxation

$$\max_{q \in \mathcal{Q}} \sum_{c \in \mathcal{C}} \sum_{\mathbf{x}_c} q(\mathbf{x}_c) \theta_c(\mathbf{x}_c) + H(q(\mathbf{x}))$$

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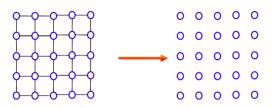
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- Relaxation algorithms work directly with pseudo-marginals which may not be consistent with any joint distribution
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$$q(\mathbf{x}) = \prod_{i \in V} q_i(x_i)$$

Naive Mean-Field

• Using the same notation naive mean-field is:

$$(*) \max_{\mu} \sum_{c \in C} \sum_{\mathbf{x}_c} \mu_c(\mathbf{x}_c) \theta_c(\mathbf{x}_c) + \sum_{i \in V} H(\mu_i) \quad \text{subject to}$$

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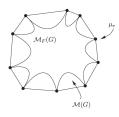
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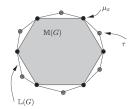
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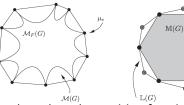
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Corresponds to optimizing over an inner bound on the marginal polytope:



• We obtain a lower bound on the partition function, i.e., $(*) \leq \ln Z(\theta)$

MAP Inference

Recall the MAP inference task

$$\arg\max_{\mathbf{x}} p(\mathbf{x}), \qquad p(\mathbf{x}) = \frac{1}{Z} \prod_{c \in C} \phi_c(\mathbf{x}_c)$$

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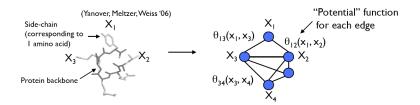
• Since the log is monotonic, let $\theta_c(\mathbf{x}_c) = \log \phi_c(\mathbf{x}_c)$

$$\arg\max_{\mathbf{x}} \sum_{c \in C} \theta_c(\mathbf{x}_c)$$

This is called the max-sum

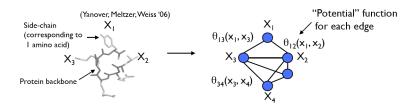
Application: protein side-chain placement

 Find "minimum energy" configuration of amino acid side-chains along fixed carbon backbone:



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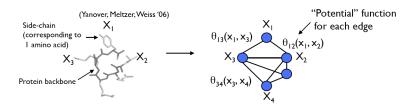
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Application: protein side-chain placement

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- Orientations of the side-chains are represented by discretized angles called rotamers
- Rotamer choices for nearby amino acids are energetically coupled (attractive and repulsive forces)

• Given a sentence, predict the dependency tree that relates the words



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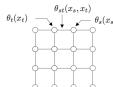
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- We represent the problem as

$$\max_{\mathbf{x}} \theta_{T}(\mathbf{x}) + \sum_{ij} \theta_{ij}(x_{ij}) + \sum_{i} \theta_{i|}(\mathbf{x}_{i|})$$

with $\mathbf{x}_{|i} = \{x_{ij}\}_{j \neq i}$ (all outgoing edges)

Application: Semantic Segmentation

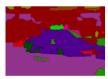
 Use Potts to encode that neighboring pixels are likely to have the same discrete label and hence belong to the same segment

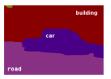


$$p(\mathbf{x}, \theta) = \max_{\mathbf{x}} \sum_{i} \theta_{i}(x_{i}) + \sum_{i,j} \theta_{i,j}(x_{i}, x_{j})$$



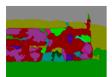


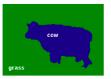












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MAP as a discrete optimization problem is

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• What is the dimension of μ , if binary variables?

$$\mathbf{x}^* = arg \max_{\mathbf{x}} \sum_{i \in V} \theta_i(x_i) + \sum_{ij \in E} \theta_{ij}(x_i, x_j)$$

- To turn this into an integer linear program (ILP) we introduce indicator variables
 - \bullet $\mu_i(x_i)$, one for each $i \in V$ and state x_i
 - ② $\mu_{ij}(x_i, x_j)$, one for each edge $ij \in E$ and pair of states x_i, x_j
- The objective function is then

$$\max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} \theta_{ij}(x_i, x_j) \mu_{ij}(x_i, x_j)$$

- What is the dimension of μ , if binary variables?
- Are these two problems equivalent?

Constraints

$$\max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} \theta_{ij}(x_i, x_j) \mu_{ij}(x_i, x_j)$$

Constraints

$$\max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} \theta_{ij}(x_i, x_j) \mu_{ij}(x_i, x_j)$$

For every "cluster" of variables to choose a local assignment

$$\mu_i(x_i) \in \{0,1\} \quad \forall i \in V, x_i$$

$$\sum_{x_i} \mu_i(x_i) = 1 \quad \forall i \in V$$

$$\mu_{ij}(x_i, x_j) \in \{0,1\} \quad \forall i, j \in E, x_i, x_j$$

$$\sum_{x_i, x_j} \mu_{ij}(x_i, x_j) = 1 \quad \forall i, j \in E$$

Constraints

$$\max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} \theta_{ij}(x_i, x_j) \mu_{ij}(x_i, x_j)$$

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• Enforce that these local assignments are globally consistent

$$\mu_i(x_i) = \sum_{x_j} \mu_{ij}(x_i, x_j) \quad \forall ij \in E, x_i$$

$$\mu_j(x_j) = \sum_{x_j} \mu_{ij}(x_i, x_j) \quad \forall ij \in E, x_j$$

$$\max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} \theta_{ij}(x_i, x_j) \mu_{ij}(x_i, x_j)$$

subject to:

$$\mu_{i}(x_{i}) \in \{0,1\} \quad \forall i \in V, x_{i}$$

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• Many extremely good off-the-shelf solvers, such as CPLEX and Gurobi

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- Many extremely good off-the-shelf solvers, such as CPLEX and Gurobi
- But it might be too slow...

$$MAP(\theta) = \max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} \theta_{ij}(x_i, x_j) \mu_{ij}(x_i, x_j)$$

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• Relax integrality constraints, allowing the variables to be between 0 and 1

$$\mu_i(x_i) \in [0,1] \ \forall i \in V, x_i \qquad \mu_{ii}(x_i, x_i) \in [0,1] \ \forall ij \in E, x_i, x_i$$

$$\begin{split} LP(\theta) &= \max_{\mu} \sum_{i \in V} \sum_{x_i} \theta_i(x_i) \mu_i(x_i) + \sum_{ij} (x_i, x_j) \mu_{ij}(x_i, x_j) \\ & \mu_i(x_i) \quad \in \quad [0, 1] \quad \forall i \in V, x_i \\ & \mu_{ij}(x_i, x_j) \quad \in \quad [0, 1] \quad \forall i, j \in E, x_i, x_j \\ & \sum_{x_i} \mu_i(x_i) \quad = \quad 1 \quad \forall i \in V \\ & \mu_i(x_i) \quad = \quad \sum_{x_j} \mu_{ij}(x_i, x_j) \quad \forall ij \in E, x_i \\ & \mu_j(x_j) \quad = \quad \sum_{x_j} \mu_{ij}(x_i, x_j) \quad \forall ij \in E, x_j \end{split}$$

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 Linear programs can be solved relatively efficient via Simplex method, interior point, ellipsoid algorithm

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- Since the LP relaxation maximizes over a larger set of solutions, its value can only be higher

$$MAP(\theta) \leq LP(\theta)$$

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LP relaxation is tight for tree-structured MRFs

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- Zero limit temperature of the variational inference for Marginals

Zemel & Urtasun (UofT)

• Introducing Lagrange multipliers and solving we get (see Murphy 22.3.5.4)

$$M_{i o j}(x_i) \propto \max_{x_j} \left[\exp\{\theta_{ij}(x_i, x_j) + \theta_j(x_j)\} \prod_{u \in N(j) \setminus i} M_{u o j}(x_j) \right]$$

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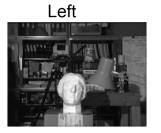
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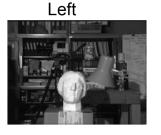
- We then compute the maximal value of $\mu_s(x_s)$
- What if two solutions that have the same score?

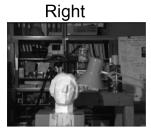
• Tsukuba images from Middlebury stereo database





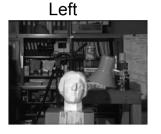
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• MRF for each pixel, with states the disparity

Tsukuba images from Middlebury stereo database



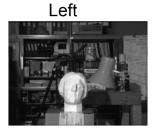


- MRF for each pixel, with states the disparity
- Our unary is the matching term

$$\theta_i(d_i) = |L(x + d_i, y) - R(x, y)|$$

where pixel $p_i = (x, y)$

Tsukuba images from Middlebury stereo database





- MRF for each pixel, with states the disparity
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ullet The pairwise factor $heta_{ij}$ between neighboring pixels favor smoothness

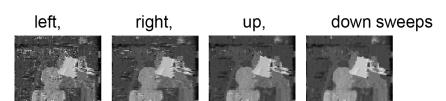
• If we only use the unary terms. How would you do inference in this case?



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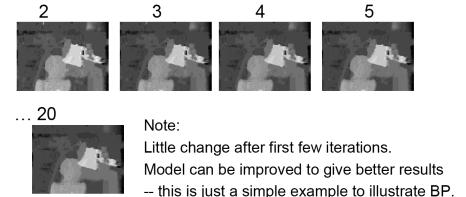
If full graphical model



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[Credit: Coughlan BP Tutorial]

Subsequent iterations:



[Credit: Coughlan BP Tutorial]