

1. (a) The parameter space Θ is the possible demand of the pairs of skis. $\therefore \Theta = \{30, 40, 50, 60\}$

The possible actions A includes buying the following pairs of skis (right at the boundary of each price change).

$$A = \{25, 50, 75\}$$

(b) The prior distribution is $f(\theta)$

$$P_f(\theta = 30) = 0.2 \quad P_f(\theta = 40) = 0.4 \quad P_f(\theta = 50) = 0.2 \\ P_f(\theta = 60) = 0.2$$

(c)

$L(g_i, a_j)$	$a_1 = 25$	$a_2 = 50$	$a_3 = 75$
$\theta_1 = 30$	1275	1375	1500
$\theta_2 = 40$	1325	875	1000
$\theta_3 = 50$	1375	375	500
$\theta_4 = 60$	1425	425	0

The scenario with the smallest loss is when $\theta_4 = 60$ $a_3 = 75$. Consider, the loss now is 0. Then the losses of other scenarios are shown above.

(d) Bayes risk

$$a_1: r(f, a_1) = 1275 \times 0.2 + 1325 \times 0.4 + 1375 \times 0.2 + 1425 \times 0.2 = 1345$$

$$a_2: r(f, a_2) = 1375 \times 0.2 + 875 \times 0.4 + 375 \times 0.2 + 425 \times 0.2 = 785$$

$$a_3: r(f, a_3) = 1500 \times 0.2 + 1000 \times 0.4 + 500 \times 0.2 = 800$$

$\therefore a_2$ has the smallest Bayes risk.

$\therefore a_2 = 50$ is the Bayes rule.

$$2. (a) f(p|x) \propto f(x|p) f(p) \propto p^x (1-p)^{n-x} p^{\alpha-1} (1-p)^{\beta-1} \\ = p^{\alpha+x-1} (1-p)^{\beta+n-x-1}$$

$$\frac{d}{d\beta} \frac{p^{\alpha+x-1} (1-p)^{\beta+n-x-1}}{(1-p)^{\alpha+\beta-1}} = \frac{\alpha(\alpha+1)}{(\alpha+\beta)(\alpha+\beta+1)}$$

$$\therefore f(p|x) \sim \text{Beta}(\alpha+x, \beta+n-x)$$

$$r(\hat{p}|x) = E_{p|x}(L(p, \hat{p}(x))) = E_{p|x}((p - \hat{p}(x))^2)$$

$$= E_{p|x}(p^2 - 2p\hat{p}(x) + \hat{p}(x)^2) = E_{p|x}(p^2 - 2p\hat{p} + \hat{p}^2)$$

$$= E_{p|x}(p^2) - 2\hat{p} E_{p|x}(p) + \hat{p}^2$$

$$= \frac{(\alpha+x)(\alpha+x+1)}{(\alpha+\beta+n)(\alpha+\beta+n+1)} - \frac{2(\alpha+x)}{(\alpha+\beta+n)} \hat{p} + \hat{p}^2$$

(b) Take the first derivative of $r(\hat{p}|x)$ to \hat{p} .

$$\frac{d}{d\hat{p}} r(\hat{p}|x) = \frac{-2(\alpha+x)}{(\alpha+\beta+n)} + 2\hat{p} = 0$$

$$\hat{p} = \frac{\alpha+x}{\alpha+\beta+n}$$

To prove it is the global minimum value, take the second derivative.

$$\frac{d^2}{d\hat{p}^2} r(\hat{p}|x) = 2 > 0$$

$\therefore \hat{p} = \frac{\alpha+x}{\alpha+\beta+n}$ minimizes $r(\hat{p}|x)$

$\therefore \hat{p} = \frac{\alpha+x}{\alpha+\beta+n}$ is the posterior mean

The Bayes estimator is also $\frac{\alpha+x}{\alpha+\beta+n}$ because of the property of $L(p, \hat{p}(x)) = (p - \hat{p}(x))^2$

(c) $\therefore \hat{p} = \frac{\alpha+x}{\alpha+\beta+n}$

$$\therefore r(\hat{p}|x) = \frac{(\alpha+x)(\alpha+x+1)}{(\alpha+\beta+n)(\alpha+\beta+n+1)} - \frac{2(\alpha+x)}{(\alpha+\beta+n)} \times \frac{\alpha+x}{\alpha+\beta+n} + \frac{(\alpha+x)^2}{(\alpha+\beta+n)^2}$$

$$= \frac{(\alpha+x)(\alpha+x+1)(\alpha+\beta+n) - 2(\alpha+x)^2(\alpha+\beta+n+1) + (\alpha+x)^2(\alpha+\beta+n+1)}{(\alpha+\beta+n)^2(\alpha+\beta+n+1)}$$

$$= \frac{(\alpha+x)((\alpha+x+1)(\alpha+\beta+n) - (\alpha+x)(\alpha+\beta+n+1))}{(\alpha+\beta+n)^2(\alpha+\beta+n+1)}$$

$$= \frac{(\alpha+x)(\beta+n-x)}{(\alpha+\beta+n)^2(\alpha+\beta+n+1)} \Rightarrow \text{which is the variance of the posterior distribution}$$

3. Proof: The zero-one loss function is:

$$L(\theta, \hat{\theta}) = \begin{cases} 1 & \theta \neq \hat{\theta} \\ 0 & \theta = \hat{\theta} \end{cases}$$

$$\begin{aligned} \text{posterior risk is: } r(\hat{\theta}|x) &= \sum_{i=1}^n L(\theta_i, \hat{\theta}) f(\theta_i|x) \\ &= \sum_{\theta \neq \hat{\theta}} f(\theta|x) = 1 - f(\hat{\theta}|x) \end{aligned}$$

\therefore Bayes estimator $\hat{\theta}$ minimizes posterior risk, so $f(\hat{\theta}|x)$ is the maximum value, so $\hat{\theta}$ is the posterior mode.

$$4. (a) R(\theta_1, a_1) = E_{X|\theta_1}[L(\theta_1, a_1)]$$

$$= P(X=0|\theta_1) L(\theta_1, a_1(0)) + P(X=1|\theta_1) L(\theta_1, a_1(1))$$

$$= 0.2 \times 1 + 0.8 \times 2 = 1.8$$

$$R(\theta_1, a_2) = E_{X|\theta_1}[L(\theta_1, a_2)]$$

$$= P(X=0|\theta_1) L(\theta_1, a_2(0)) + P(X=1|\theta_1) L(\theta_1, a_2(1))$$

$$= 0.2 \times 4 + 0.8 \times 0 = 0.8$$

$$R(\theta_2, a_1) = E_{X|\theta_2}[L(\theta_2, a_1)]$$

$$= P(X=0|\theta_2) L(\theta_2, a_1(0)) + P(X=1|\theta_2) L(\theta_2, a_1(1))$$

$$= 0.4 \times 3 + 0.6 \times 1 = 1.8$$

$$R(\theta_2, a_2) = E_{X|\theta_2}[L(\theta_2, a_2)]$$

$$= P(X=0|\theta_2) L(\theta_2, a_2(0)) + P(X=1|\theta_2) L(\theta_2, a_2(1))$$

$$= 0.4 \times 1 + 0.6 \times 4 = 2.8$$

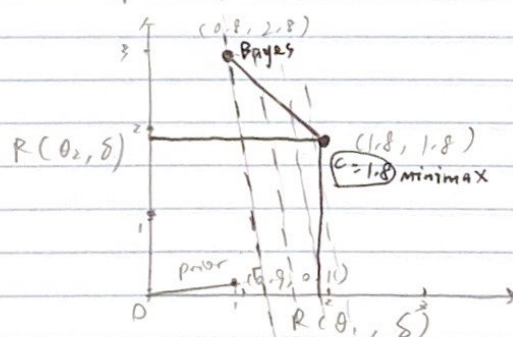
$$\therefore \lambda \in [0, 1]$$

$$\therefore \text{When } \lambda = 0 \quad (r_1, r_2) = (R(\theta_1, a_2), R(\theta_2, a_2)) = (0.8, 2.8)$$

$$\text{When } \lambda = 1 \quad (r_1, r_2) = (R(\theta_1, a_1), R(\theta_2, a_1)) = (1.8, 1.8)$$

$\therefore S$ is clearly a linear function

\therefore The plot of S is



$$r(f, a_1) = 0.9 \times 1.8 + 0.1 \times 1.8 = 1.8$$

$$r(f, a_2) = 0.9 \times 0.8 + 0.1 \times 2.8 = 1.0 \text{ smallest}$$

$\therefore a_2$ is the Bayes rule.

(b) The Bayes risk equals to the dot product of the prior vector and the point in the risk set. The y-intersection of the

Those dashed lines in the graph above is proportional to the Bayes and perpendicular risk. So the line with the smallest Bayes risk is the to prior vector. line through point $(0.8, 2.8)$. So the Bayes rule is

$$(0.8, 2.8) \Rightarrow (R(\theta_1, a_2), R(\theta_2, a_2)) \rightarrow \text{action } a_2$$

(c) As for the minimax rule.

$$a_1: \max_{\theta} R(\theta, a_1) = 1.8 \text{ smallest} \quad a_2: \max_{\theta} R(\theta, a_2) = 2.8$$

Then minimax rule will be a_1 . The first such of square $A(c)$ to Risk Set.