\* Bigram models of language

$$(\omega) \longrightarrow (\omega_{2}) \longrightarrow \cdots \longrightarrow (\omega_{c})$$

We € {1,2,..., V} V=# words in vocabular,

$$P_{ML}(w_{e+1}=j|w_{e}=i) = \frac{C_{ij}}{C_{i}} = \frac{count(i\rightarrow j)}{count(i)}$$

\* ML Estimation from incomplete data

Examples t=1,2,--,T

Hilden Nodes Hut

Visible Nodes V(\*)

Choos CPTs to maximize log-likelihood I = E log P(V)

How?

\* EM Algorithm

Iterative procedure to maximize & log P(V\*) in terms of CPTs

E-Step: compute posterior probabilities

M-Step: update CPTs

$$P(X_i = x | pa_i = \pi)$$
  $\leftarrow \frac{\sum P(X_i = x, pa_i = \pi | V = v(t))}{\sum P(pa_i = \pi | V = v(t))}$ 

Intuition: expected statistics under P(H|V) are filling in "missing values"

Iterate E&M Steps until convergence. Why iterate? RHS depends on current CPTs.

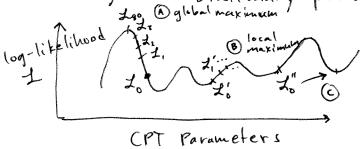
\* Key Properties

- monotonic convergence

Each iteration of EM improves the log-likelihood

If Ik is log-likelihood at kth iteration, then Ik = Ik+

- Converges to stationary point Log @global maximum



A global maximum: most desirable, but not guaranteed

B local maximum: usual outcome

© local minimum; possible in theory but never occurs in practice.

- No tuning parameters: no step sizes, learning rates, back-tracking, ...

## Example

$$\textcircled{B} \longrightarrow \textcircled{B} \longrightarrow \textcircled{B}$$

A and C are observed (visible nodes). B is hidden.

\* Posterior probability P(C=c|B=b,A=a)P(B=b|A=a)P(B=b | A=a, (=c) = Bayer rule E P(C=c| B=b', A=a) P(B=b'|A=a)  $= \frac{P(C=c|B=b)P(B=b|A=a)}{\sum_{b'} P(C=c|B=b')P(B=b'|A=a)}$ conditional independence

Shorthand: P(b|a,c) & P(B=b|A=a, C=c)

\* Incomplete data set 
$$\{(a_{t}, c_{t})\}_{t=1}^{T}$$
 (I.I.D.)

 $J = \{\{\{a_{t}, c_{t}\}\}_{t=1}^{T}\}$  (I.I.D.)

 $J = \{\{a_{t}, c_{t}\}\}$  (I.I.D.)

 $J$ 

General EM Algorithm:
$$P(X_i=x \mid pa_i=\pi) \neq \underbrace{\qquad \qquad }_{\substack{t \in P(X_i=x, pa_i=\pi \mid V^{tt})}}$$

$$\underbrace{P(X_i=x \mid pa_i=\pi \mid V^{tt})}_{\substack{t \in P(pa_i=\pi \mid V^{tt})}}$$

Now apply to this example:  

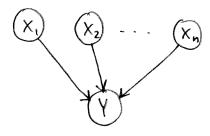
$$M-step: P(B=b|A=a) \leftarrow \frac{\sum P(A=a,B=b|A=a_{t,l}(=c_{t}))}{\sum P(A=a|A=a_{t,l}(=c_{t}))}$$

Simplify RHS: 
$$\frac{\mathcal{E}}{t} I(a_1 a_t) P(\mathbf{s} b | a_t, c_t)$$

$$P(C=c|B=b) \leftarrow \frac{\sum_{t}^{p} P(B=b)}{\sum_{t}^{q} P(B=b|A=a_{t}, C=c_{t})}$$
  
Simplify:

$$P(C=c|B=b) \leftarrow \frac{\sum I(c_1c_1)P(b|a_1,c_1)}{\sum P(b|a_1,c_1)}$$

## Noisy - OR Model



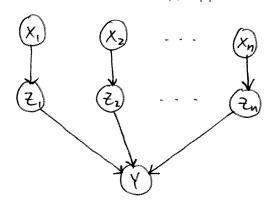
disease X; € {0,1} symptom Y ∈ {0,1}

$$P(Y=1|X_{1,1}X_{2,1-1},X_{n}) = 1 - \prod_{i=1}^{n} (1-p_{i})^{X_{i}}$$
 with  $p_{i} \in [0,1]$ 

- \* From complete data  $\{(\vec{x}_t, Y_t)\}_{t=1}^T$ , how do we estimate  $p_i \in [0,1]$ ?

  Note: Noisy-or is a "parametric" model of CPT.

  No simply, closed-form ML estimate for  $p_i \in [0,1]$ .
- \* Alternative formulation:



 $P(Y|Z_1,Z_2,...,Z_n) = OR(Z_1,Z_2,...,Z_n)$  logical-OR (deterministic)

$$P(z_{i}=1 | x=1) = p_{i}$$
  
 $P(z_{i}=1 | x=0) = 0$ 

Equivalently:  $P(Z_i=0|X_i)=(1-p_i)^{X_i}=\begin{cases} 1-p_i & X=1\\ 1 & X=0 \end{cases}$ What is  $P(Y=1|\overline{X})$  in this new model?

$$P(Y=1|\vec{X}) = \sum_{\vec{z} \in \{0,1\}^n} P(Y=1,\vec{z}|\vec{X}) \quad \text{marginalization}$$

$$= \sum_{\vec{z}} P(Y=1|\vec{z}|\vec{X}) P(\vec{z}|\vec{X}) \quad \text{product rule}$$

 $= \sum_{\vec{z} \neq \vec{\sigma}} P(\vec{z} | \vec{X})$ 

product rule conditional independence

because V=OR(2)

$$= |-P(\vec{z} = \vec{0} | \vec{x})$$
 from normalization
$$= |-TT| P(\vec{z} = 0 | \vec{x}_i)$$
 conditional Independence
$$= |-TT| (1-p_i)^{X_i}$$

Same as original Noisy-OR BN!

\* Posterior Probability Oify=0 (pi if  $\vec{x}_i = 1$ )
$$P(\vec{z}_i = 1 | \vec{x}_i \vec{y}_i) = P(\vec{y} | \vec{x}_i \vec{z}_i = 1) P(\vec{z}_i = 1 | \vec{x}_i)$$

$$P(\vec{y} | \vec{x}_i) = \frac{Y P(\vec{x}_i)}{1-T(1-p_i)^{X_i}}$$

$$= \frac{Y P(\vec{x}_i)}{1-T(1-p_i)^{X_i}}$$