

Homework 3

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Problem 1

(a) $\sum_{j=0}^{p-1} c_j \hat{\beta}_j$ can be denoted as $c^T \hat{\beta}$, where c and $\hat{\beta}$ are vectors. In least squares estimate, $\hat{\beta} = (X^T X)^{-1} X^T Y$, so

$$\text{Mean: } E(\sum_{j=0}^{p-1} c_j \hat{\beta}_j) = c^T \beta$$

$$\text{Variance: } \text{Var}(c^T \hat{\beta}) = c^T \sigma^2 (X^T X)^{-1} c$$

$$\text{Distribution: } c^T \hat{\beta} \sim N(c^T \beta, c^T \sigma^2 (X^T X)^{-1} c)$$

(b) Consider that $c^T \hat{\beta} \sim N(c^T \beta, c^T \sigma^2 (X^T X)^{-1} c)$. Then we know that $c^T \hat{\beta} - c^T \beta \sim N(0, c^T \sigma^2 (X^T X)^{-1} c)$.

Under H_0 , we use T test, and the test statistics is

$$T = \frac{c^T \hat{\beta} - h}{c^T \sigma^2 (X^T X)^{-1} c}$$

Under significant level α , the reject region is

$$|T| > t_{n-p, \alpha/2}$$

In this case, $\hat{\beta}$ is independent.

(c)

$$E(Y_{n+1} - \hat{Y}_{n+1}) = 0$$

$$\text{Var}(Y_{n+1} - \hat{Y}_{n+1}) = \sigma^2 (1 + Z^T (X^T X)^{-1} Z)$$

$$Y_{n+1} - \hat{Y}_{n+1} \sim N(0, \sigma^2 (1 + Z^T (X^T X)^{-1} Z))$$

(d) Proof:

$$\begin{aligned}MSE(\hat{Y}_{n+1}) &= E(\hat{Y}_{n+1} - Y_{n+1})^2 = \sigma^2(1 + Z^T(X^T X)^{-1}Z) \\Z^T(X^T X)^{-1}Z &= Z^T(X^T X)^{-1}X^T X(X^T X)^{-1}Z > 0 \\MSE(\hat{Y}_{n+1}) &= \sigma^2(1 + Z^T(X^T X)^{-1}Z) > \sigma^2\end{aligned}$$

(e) We consider the T test.

$$T = \frac{Y_{n+1} - \hat{Y}_{n+1}}{\sqrt{\sigma^2(1 + Z^T(X^T X)^{-1}Z)}} \sim t_{n-p}$$

So, the interval I

$$(\hat{Y}_{n+1} - t_{n-p, \alpha/2} \sqrt{\sigma^2(1 + Z^T(X^T X)^{-1}Z)}, \hat{Y}_{n+1} + t_{n-p, \alpha/2} \sqrt{\sigma^2(1 + Z^T(X^T X)^{-1}Z)})$$

Problem 2

(a)

The first model can be described as a single line because it is a SLR.

The second model can be described as a parallel lines because it contains indicator variables thus it can change the intercept when $l_{mass} = 0$, but cannot change the slope.

The third model can be described as a separate lines because it contains cross variables so it has both different intercepts and slopes.

(b)

F-test: whether there is a difference between the in-flight energy and body mass among birds, echolocating and non-echolocation bats. The second model is the reduced model and the first model is the full model.

Other test: whether there is a difference for birds, echolocating bats and non-echolocating bats after body mass is accounted for. The second model is the full model and the first model is the reduced model.

(c)

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> anova(lm1, lm2, lm3)
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Analysis of Variance Table

Model 1: lenergy ~ lmass

Model 2: lenergy ~ lmass + Type

Model 3: lenergy ~ Type * lmass

| | Res.Df | RSS | Df | Sum of Sq | F | Pr(>F) |
|---|--------|---------|----|-----------|--------|--------|
| 1 | 18 | 0.58289 | | | | |
| 2 | 16 | 0.55332 | 2 | 0.029574 | 0.4100 | 0.6713 |
| 3 | 14 | 0.50487 | 2 | 0.048450 | 0.6718 | 0.5265 |

Comparing the first, the second and third model, the p-values are 0.6713, 0.5265 separately. Hence, the first model is the best, which means that the full model is reduced to

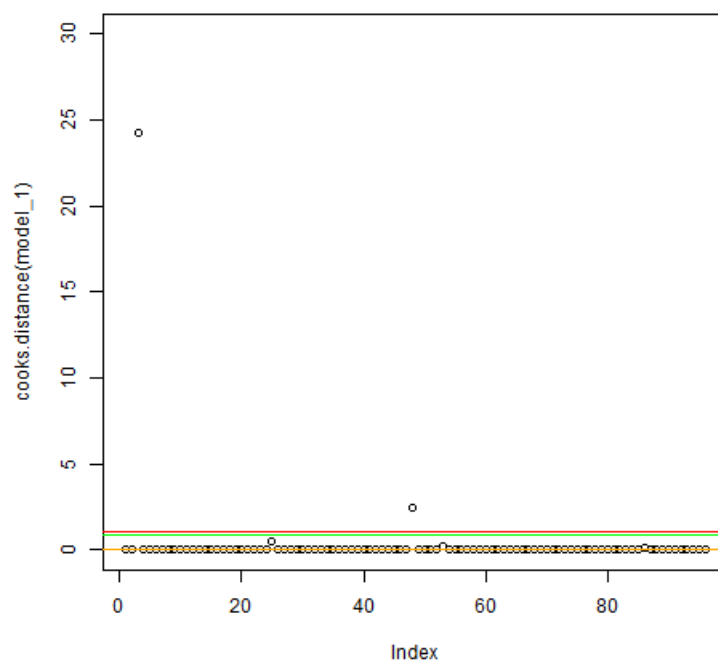
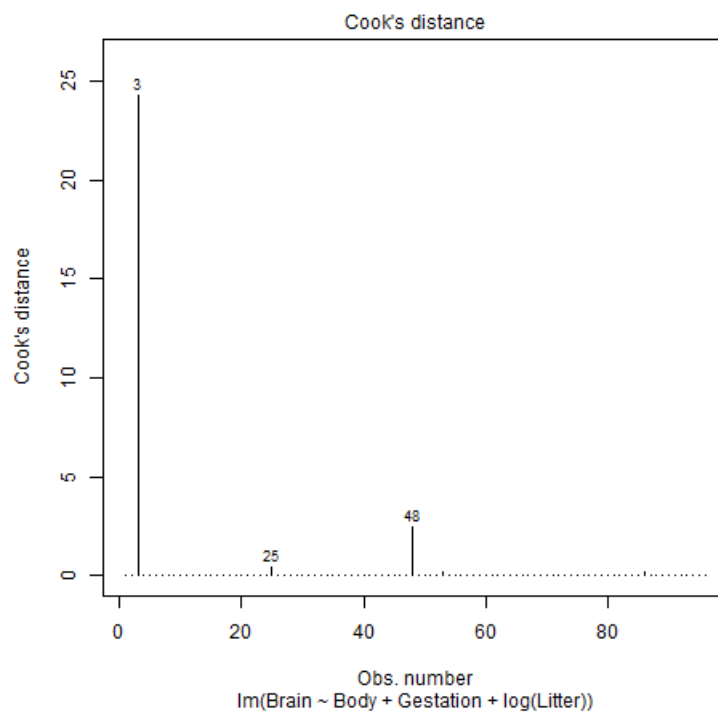
$$lenergy \sim lmass$$

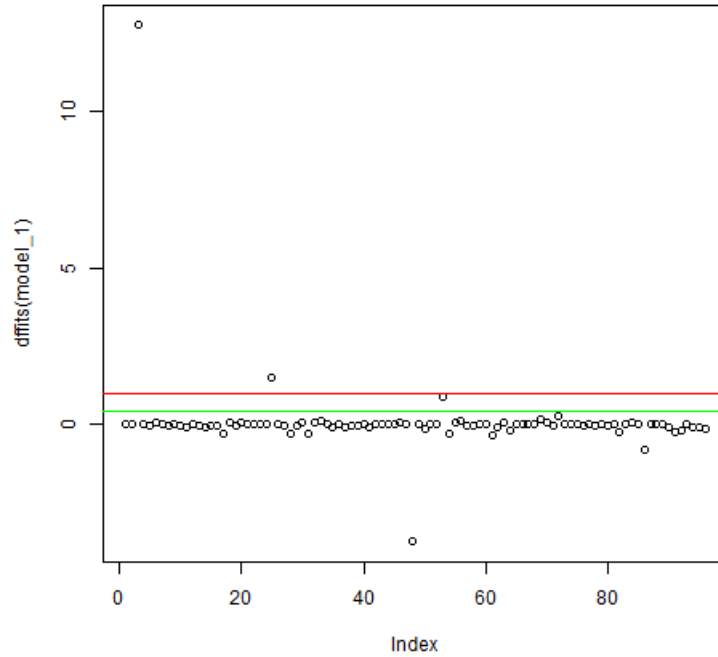
Problem 3

(a) By using R, we do the regression:

$$Y = -231.58 + 0.97X_1 + 1.93X_2 + 89\ln(X_3)$$

To find influential points, we compute the cook's distance and fiffits value and the graphs are shown below:



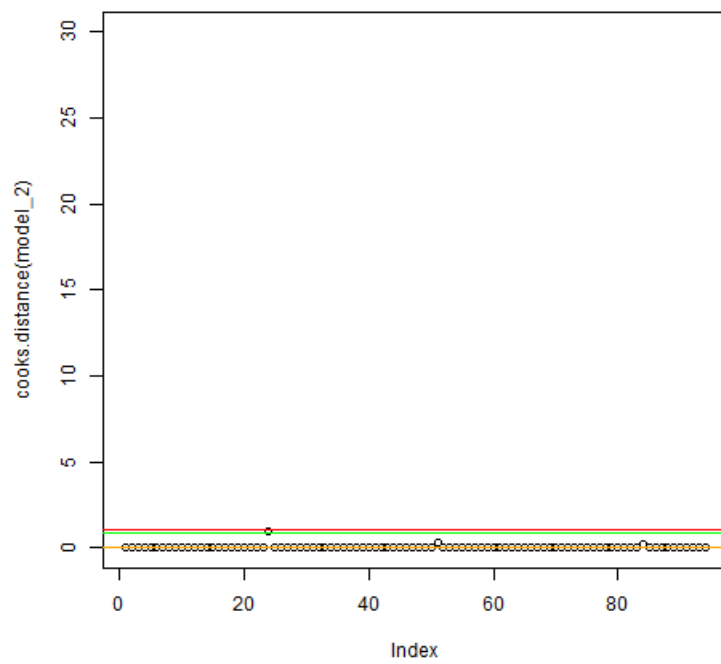
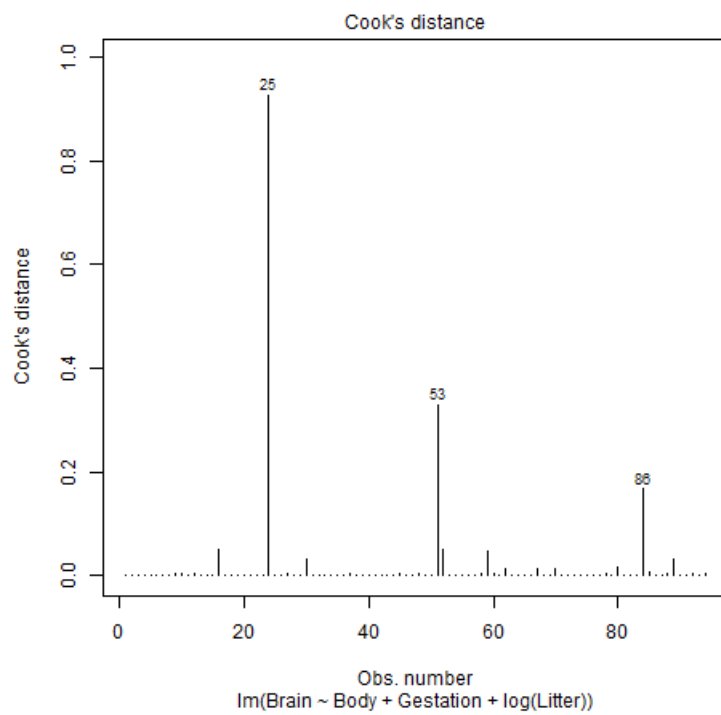


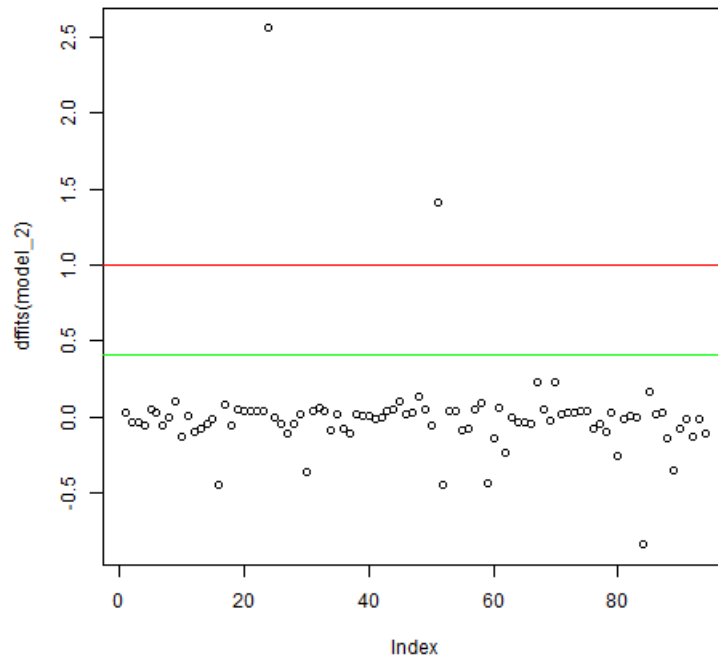
Based on the graphs, we can see that there are two influential points: 3, 25, 48, 53 and 86, which corresponding to African elephant, dolpin, hippopotamus, human being and Tapir respectively.

(b) After deleting the influential points, we do the regression again and get

$$Y = -135.44 + 0.47X_1 + 1.972X_2 + 39.38\ln(X_3)$$

To find influential points, we compute the cook's distance and dffits value and the graphs are shown below:





Based on the graphs, there are still two influential points; 25 and 53.

Main differences: The coefficient of body weight and litter size and interception changed, F-value and t-value are decreasing.

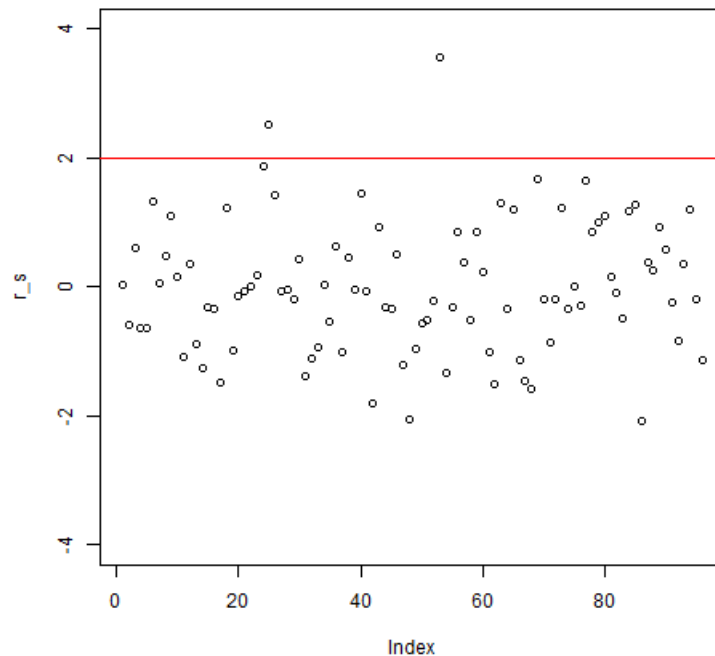
(c) Consider all the data and fit the regression of log brain weight on log body weight, log gestation, and log litter size:

$$\ln(Y) = 0.85482 + 0.57507\ln(X_1) + 0.41794\ln(X_2) - 0.31007\ln(X_3)$$

Then, we can calculate the studentized residuals:

| > r_s | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 |
|-------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|
| | 0.03743244 | -0.59640636 | 0.60328414 | -0.62769966 | -0.63113474 | 1.32746311 | 0.04998926 | 0.47261422 |
| | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 |
| | 1.09307100 | 0.16232902 | -1.07699483 | 0.35226078 | -0.88133921 | -1.25638075 | -0.32026894 | -0.33613386 |
| | 17 | 18 | 19 | 20 | 21 | 22 | 23 | 24 |
| | -1.47463460 | 1.21438396 | -0.99798694 | -0.14673423 | -0.06347308 | 0.01125227 | 0.19359786 | 1.87603881 |
| | 25 | 26 | 27 | 28 | 29 | 30 | 31 | 32 |
| | 2.51898284 | 1.42291877 | -0.06192184 | -0.03992283 | -0.18511100 | 0.42982103 | -1.37284899 | -1.10691827 |
| | 33 | 34 | 35 | 36 | 37 | 38 | 39 | 40 |
| | -0.94276157 | 0.02146864 | -0.54734955 | 0.62567595 | -1.00847048 | 0.45726259 | -0.04718166 | 1.43773970 |
| | 41 | 42 | 43 | 44 | 45 | 46 | 47 | 48 |
| | -0.06989352 | -1.81623410 | 0.92124229 | -0.30751319 | -0.33994930 | 0.51483675 | -1.20725253 | -2.06358145 |
| | 49 | 50 | 51 | 52 | 53 | 54 | 55 | 56 |
| | -0.95073672 | -0.55595462 | -0.51873099 | -0.21630768 | 3.56200203 | -1.32504584 | -0.31638860 | 0.86161582 |
| | 57 | 58 | 59 | 60 | 61 | 62 | 63 | 64 |
| | 0.37144308 | -0.51348066 | 0.84246524 | 0.22297786 | -1.00275641 | -1.52085748 | 1.30755363 | -0.34956029 |
| | 65 | 66 | 67 | 68 | 69 | 70 | 71 | 72 |
| | 1.21232022 | -1.14859836 | -1.45193427 | -1.59083568 | 1.67086812 | -0.18567379 | -0.85392954 | -0.18195498 |
| | 73 | 74 | 75 | 76 | 77 | 78 | 79 | 80 |
| | 1.22084501 | -0.33484868 | 0.01087871 | -0.30203906 | 1.65498562 | 0.84042796 | 0.99887177 | 1.10146238 |
| | 81 | 82 | 83 | 84 | 85 | 86 | 87 | 88 |
| | 0.14542630 | -0.10334766 | -0.49922765 | 1.16613802 | 1.27758691 | -2.08175598 | 0.39041565 | 0.26424150 |
| | 89 | 90 | 91 | 92 | 93 | 94 | 95 | 96 |
| | 0.92300908 | 0.57397277 | -0.23412728 | -0.82806937 | 0.35813822 | 1.19458870 | -0.18847718 | -1.14395781 |

(d) To find influential points, we compute the cook's distance and fiffits value and the graphs are shown below:



we can conclude that at the points 25 and 53, which corresponding to dolphin and human, have substantially larger brain weights than were predicted by the model.

(e) According to the graphics shown in (d), the conclusion is — there is no mammals have substantially smaller brain weights than were predicted by the model.

(f)

1. After log transformation, F-value increased, which means the model fits better.
2. Removing influential points and refit the regression may not be a good way to make thing right, even make it worse.
3. The assumption of normality can hold by log transformation.

Problem 4

(a) $SSTO = 8100$ and $\hat{\sigma}^2 = \frac{SSE}{df}$. The estimates of σ^2 are shown below:

| Model variables | $\hat{\sigma}^2$ |
|-----------------|------------------|
| None | 300 |
| A | 240 |
| B | 230 |
| C | 260 |
| AB | 220 |
| AC | 210 |
| BC | 230 |
| ABC | 215 |

(b) Adjusted $R^2 = 1 - \frac{\frac{SSE_p}{n-p}}{\frac{SSTO}{N-1}}$. The adjusted R^2 are shown below:

| Model variables | Adjusted R^2 |
|-----------------|----------------|
| None | 0 |
| A | 0.2 |
| B | 0.233 |
| C | 0.133 |
| AB | 0.267 |
| AC | 0.3 |
| BC | 0.233 |
| ABC | 0.283 |

(c) $C_p(M) = \frac{SSE(M)}{\hat{\sigma}^2} - n + 2p(M)$, the C_p statistics are shown below:

| Model variables | C_p |
|-----------------|--------|
| None | 11.674 |
| A | 5.023 |
| B | 3.814 |
| C | 7.442 |
| AB | 3.851 |
| AC | 2.419 |
| BC | 4.744 |
| ABC | 4 |

(d) $BIC_p = n \ln(SSE_p) - n \ln(n) + p \ln(n)$. The BIC are shown below:

| Model variables | BIC_p |
|-----------------|----------|
| None | 162.0198 |
| A | 158.0473 |
| B | 156.8556 |
| C | 160.2885 |
| AB | 157.8450 |
| AC | 156.5424 |
| BC | 159.0896 |
| ABC | 159.3905 |

(e)

(i) The last model with variables A, B and C has the smallest $\hat{\sigma}^2$, 198.46.

(ii) The model with variables A and C has the largest adjusted R^2 , 0.3.

(iii) The model with variables A and C has the smallest C_p statistic, 2.419.

(iv) The model with variables A and C has the smallest BIC , 156.54.

(f)

Starting with none variable, we get the model contains Bhas the smallest SSE. Then we perform an extra-sum-of-squares F -test to see whether that variable is significant.

Under $\alpha = 0.05$, $F = 7.75$. Thus, B is significant.

Then, we get AB has the smaller SSE. Samely, we perform an extra-sum-of-squares F -test to see whether the additional variable is significant.

Under $\alpha = 0.05$, $F = 2.182 < 4.24 = F_{1,n-4,0.05}$. So Ais insignificant.

Finally, by forward selecion, we have the model form:

$$Y\beta_0 + \beta_2B + \epsilon$$

(g) To find the posterior probability for these 8 models, we first subtract the smallest BIC from the other ones, and then calculate the difference of each model as below:

$$BIC_j^* = BIC_j - BIC_{min}$$

Then,we have the posterior probability:

$$p(M_j) = \frac{e^{-\frac{1}{2}BIC_j^*}}{\sum_{j=1}^k e^{-\frac{1}{2}BIC_j^*}}$$

The posterior distributions are shown below:

| Model variables | Posterior distribution |
|-----------------|------------------------|
| None | 0.01802701 |
| A | 0.13138647 |
| B | 0.23840663 |
| C | 0.04284313 |
| AB | 0.14537115 |
| AC | 0.27882110 |
| BC | 0.07802002 |
| ABC | 0.06712450 |

A Appendix: Problem 2

```
bats = read.csv( file = "bats.csv" , header = T)
bats = transform(bats , lenergy = log(Energy))
bats = transform(bats , lmass = log(Mass))
lm1 = lm(lenergy ~ lmass , data = bats)
lm2 = lm(lenergy ~ lmass + Type, data = bats)
lm3 = lm(lenergy ~ Type*lmass , data = bats)
anova(lm1 , lm2 , lm3)
```

B Appendix: Problem 3

```
data = read.csv( "brain.csv")
model_1=lm(Brain ~ Body+Gestation+log(Litter) ,data = data)
summary(model_1)
cooks.distance(model_1)
n = dim(model.matrix(model_1))[1]
p = dim(model.matrix(model_1))[2]
png("3_a_1.png")
plot(model_1, which = 4)
dev.off()
png("3_a_2.png")
plot(cooks.distance(model_1) ,ylim = c(0,30))
abline(h=1, col="red")
abline(h=qf(0.5 , p, n-p) , col="green")
abline(h=4/n, col="blue")
abline(h=4/(n-p-1-1), col="orange")
dev.off()
dffits(model_1)
png("3_a_3.png")
plot(dffits(model_1) )
abline(h=1, col="red")
abline(h=2*sqrt(p/n) , col="green")
```

```

dev.off ()
summary(influence.measures(model_1))
data_new = data[-48,][-3,]
model_2=lm(Brain ~ Body+Gestation+log(Litter) ,data = data_new)
summary(model_2)
cooks.distance(model_2)
n = dim(model.matrix(model_2))[1]
p = dim(model.matrix(model_2))[2]
png("3_b_1.png")
plot(model_2, which = 4)
dev.off ()
png("3_b_2.png")
plot(cooks.distance(model_2) ,ylim = c(0,30))
abline(h=1, col="red")
abline(h=qf(0.5 , p, n-p) , col="green")
abline(h=4/n, col="blue")
abline(h=4/(n-p-1-1), col="orange")
dev.off ()
dffits(model_2)
png("3_b_3.png")
plot(dffits(model_2) )
abline(h=1, col="red")
abline(h=2*sqrt(p/n) , col="green")
dev.off ()
summary(influence.measures(model_2))

```