

► Auto covariance

$$C_{xx}(t_1, t_2) = E \left[\underbrace{(X(t_1) - \eta_X(t_1))}_{\bar{X}(t_1)} \underbrace{(X(t_2) - \eta_X(t_2))}_{\bar{X}(t_2)} \right] = R_{xx}(t_1, t_2) - \eta_X(t_1)\eta_X(t_2)$$

* for zero mean R.P.

$$R_{xx}(t_1, t_2) = C_{xx}(t_1, t_2)$$

► Correlation-coefficient

$$\rho_{X, X}(t_1, t_2) = \frac{C_{xx}(t_1, t_2)}{\sqrt{C_{xx}(t_1, t_1) \cdot C_{xx}(t_2, t_2)}} \quad \text{as } C_{xx}(t_1, t_1) = \text{var}(X(t_1))$$

Note if $X(t_1)$ and $X(t_2)$ are independent, then

$$R_{xx}(t_1, t_2) = E[X(t_1) \cdot X(t_2)] = \eta_X(t_1) \cdot \eta_X(t_2)$$

and thus,

$$C_{xx}(t_1, t_2) = R_{xx}(t_1, t_2) - \eta_X(t_1) \cdot \eta_X(t_2) = 0$$

[if covariance is zero \Rightarrow R.V. are un-correlated]

Let $X(t) = \underbrace{r_0}_{\text{R.V.}} \cos(\omega t + \underbrace{\phi}_{U[-\pi, \pi]})$

$$R_{xx}(t_1, t_2) = E[X(t_1) \cdot X(t_2)]$$

$$= E \left[\int_{-\pi}^{\pi} \int_{-\pi}^{\pi} \right]$$

$$= E[r^2 \cos(\omega t_1 + \phi) \cdot \cos(\omega t_2 + \phi)]$$

$$= E[r^2] \cdot E[\cos(\omega t_1 + \phi) \cdot \cos(\omega t_2 + \phi)]$$

$$= \frac{E[r^2]}{2\pi} \int_{-\pi}^{\pi} \cos(\omega t_1 + \phi) \cdot \cos(\omega t_2 + \phi) \cdot d\phi$$

$$R_{xx}(t_1, t_2) = \frac{E[Y^2]}{2\lambda} \cos(\omega(t_1 - t_2))$$

we see that autocorrelation does not depend on actual values of t_1, t_2 (only upon the difference $t_1 - t_2$).

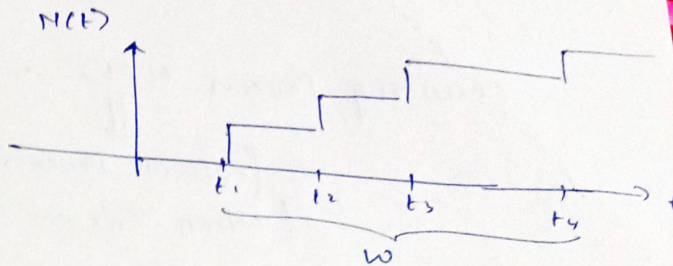
Let $\omega \in \Omega$

and $\omega = (t_1, t_2, t_3, t_4)$.

then $N(t)$ for this ω is

then $N(t; \omega)$ is R.P. function.

$N(t)$ is R.P.



Let $T(n)$ represent the arrival time of n th customer.

$$\Delta T(n) = T(n+1) - T(n)$$

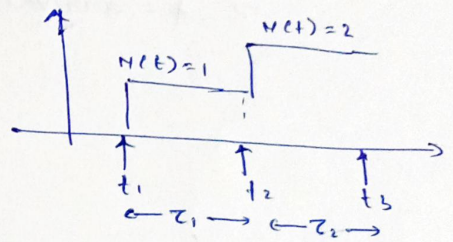
we can easily show that

\Rightarrow

$$T(n) = \sum_{i=1}^{n-1} \Delta T[i]$$

\uparrow
{sum of iid $\exp(\lambda)$ }

\hookrightarrow iid of $\exp(\lambda)$
(say)



$\Rightarrow T(n) \Rightarrow$ conv. of T_i s

we will get $f_T(\alpha, n) = \frac{(\lambda \alpha)^{n-1}}{(n-1)!} \exp(-\lambda \alpha)$; $\alpha > 0$

\uparrow
Erlang's Distribution.

[Sum of iid $\exp(\lambda)$ is Erlang's Dist]

$$f_T(t) = \lambda \exp(-\lambda t) \quad t \geq 0$$

$$P[N(t) = n] = P[T(n) \leq t \text{ and } T(n+1) > t].$$

Now since $\tau(n) = T(n+1) - T(n) \Rightarrow T(n+1) = \tau(n) + T(n)$

$$P[N(t) = n] = P[T(n) \leq t; \tau(n) > t - \overbrace{T(n)}^{\lambda t = \alpha}]$$

(we will say
own whole $T(n)$
later)

$$= \cancel{P[T(n) \leq t]} \cdot P[\tau(n) > t - T(n)]$$

$$P[N(t) = n] = \int_0^t f_T(\alpha, n) \cdot \int_{t-\alpha}^{\infty} f_{\tau}(\beta) d\beta d\alpha = \frac{(\lambda t)^n \exp(-\lambda t)}{n!}$$

Counting Process $N(t)$ ~~~~~ Poisson(λt)

\Downarrow
 \Rightarrow [Poisson Process]
 [when τ_n 's are iid & $\exp(\lambda)$]

$$E[N(t)] = \lambda t.$$

$\Rightarrow \lambda = \text{arrival rate of customers.}$