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p-value

In <u>statistical hypothesis</u> is true, the <u>statistical summary</u> (such as the absolute value of the sample mean difference between two compared groups) would be greater than or equal to the <u>actual observed results</u>.^[1] In the case of a composite null hypothesis, the *p*-value is defined as the worst case probability. The use of *p*-values in statistical hypothesis testing is common in many fields of research^[2] such as <u>physics</u>, <u>economics</u>, <u>finance</u>, <u>political science</u>, <u>psychology</u>, <u>biology</u>, <u>criminal justice</u>, <u>criminology</u>, and sociology.^[4] The misuse of *p*-values has been a matter of considerable controversy

Italicisation, capitalisation and hyphenation of the term varies. For example, \underline{AMA} style uses "P value," \underline{APA} style uses "p value," and the American Statistical Associationuses "p-value." [5]

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Basic concepts

In statistics, every conjecture concerning the unknown distribution \mathbf{F} of a random variable \mathbf{X} is called a *statistical hypothesis*. If we state one hypothesis only and the aim of the statistical test is to verify whether this hypothesis is not false, but not, at the same time, to investigate other hypotheses, then such a test is called a *significance test*. A statistical hypothesis that refers only to the numerical values of unknown parameters of a distribution is called a *parametric hypothesis*. Methods of verifying statistical hypotheses are called *statistical tests*. Tests of parametric hypotheses are called *parametric tests*. We can likewise also have *non-parametric hypotheses* and *non-parametric tests*.

The *p*-value is used in the context of <u>null hypothesis</u> testing in order to quantify the idea of <u>statistical significance</u> of evidence. ^[note 1] Null hypothesis testing is a <u>reductio ad absurdum</u> argument adapted to statistics. In essence, a claim is assumed valid if its counterclaim is improbable.

As such, the only hypothesis that needs to be specified in this test and which embodies the counter-claim is referred to as the <u>null</u> <u>hypothesis</u> (that is, the hypothesis to be nullified). A result is said to be *statistically significant* if it allows us to reject the null hypothesis. That is, as per the reductio ad absurdum reasoning, the statistically significant result should be highly improbable if the null hypothesis is assumed to be true. The rejection of the null hypothesis implies that the correct hypothesis lies in the logical complement of the null hypothesis. However, unless there is a single alternative to the null hypothesis, the rejection of null hypothesis does not tell us which of the alternatives might be the correct one.

As a general example, if a null hypothesis is assumed to follow the standard $\underline{\text{normal distribution}}$ N(0,1), then the rejection of this null hypothesis can either mean (i) the mean is not zero, or (ii) the $\underline{\text{variance}}$ is not unity, or (iii) the distribution is not normal, depending on the type of test performed. However, supposing we manage to reject the zero mean hypothesis, even if we know the distribution is normal and variance is unity the null hypothesis test does not tell us which non-zero value we should adopt as the new mean.

If X is a <u>random variable</u> representing the observed data and H is the statistical hypothesis under consideration, then the notion of statistical significance can be naively quantified by the <u>conditional probability</u> Pr(X|H), which gives the likelihood of a certain observation event X if the hypothesis is *assumed* to be correct. However, if X is a continuous random variable, the probability of observing a specific instance x is zero, that is, Pr(X = x|H) = 0. Thus, this naive definition is inadequate and needs to be changed so as to accommodate the continuous random variables.

Nonetheless, it helps to clarify that p-values should **not** be confused with probability on hypothesis (as is done in <u>Bayesian</u> <u>hypothesis testing</u>) such as $\mathbf{Pr}(H|X)$, the probability of the hypothesis given the data, or $\mathbf{Pr}(H)$, the probability of the hypothesis being true, or $\mathbf{Pr}(X)$, the probability of observing the given data.

Definition and interpretation

The p-value is defined as the probability, under the <u>null hypothesis</u> H (at times denoted H_0 as opposed to H_a denoting the alternative hypothesis) on a population <u>variate</u>, for the variate to be observed as a value equal to or more extreme than the value observed. Depending on how it is looked at, the "equal to or more extreme than what was actually observed" can mean $\{X \geq x\}$ (right-tail event), $\{X \leq x\}$ (left-tail event) or the event giving the smallest probability among $\{X \leq x\}$ and $\{X \geq x\}$ (double-tailed event). Thus, the p-value is given by

- $\Pr(X \ge x | H)$ for right tail event,
- $Pr(X \le x|H)$ for left tail event,
- $2\min\{\Pr(X \le x|H), \Pr(X \ge x|H)\}$ for double tail event.

The smaller the *p*-value, the higher the significance because it tells the investigator that the hypothesis under consideration may not adequately explain the observation. The null hypothesis \boldsymbol{H} is rejected if any of these probabilities is less than or equal to a small, fixed but arbitrarily pre-defined threshold value $\boldsymbol{\alpha}$, which is referred to as the <u>level of significance</u>. Unlike the *p*-value, the $\boldsymbol{\alpha}$ level is not derived from any observational data and does not depend on the underlying hypothesis; the value of $\boldsymbol{\alpha}$ is instead set by the researcher before examining the data. The setting of $\boldsymbol{\alpha}$ is arbitrary. By convention, $\boldsymbol{\alpha}$ is commonly set to 0.05, 0.01, 0.005, or 0.001.

Since the value of \boldsymbol{x} that defines the left tail or right tail event is a random variable, this makes the p-value a function of \boldsymbol{x} and a random variable in itself; under the null hypothesis, the p-value is defined uniformly over [0,1] interval, assuming \boldsymbol{x} is continuous. Thus, the p-value is not fixed. This implies that p-value cannot be given a frequency counting interpretation since the probability has to be fixed for the frequency counting interpretation to hold. In other words, if the same test is repeated independently bearing upon the same overall null hypothesis, it will yield different p-values at every repetition. Nevertheless, these different p-values can be combined, for instance using Fisher's combined probability test It should further be noted that an instantiation of this random p-value can still be given a frequency counting interpretation with respect to the number of observations taken during a given test, as per the definition, as the percentage of observations more extreme than the one observed under the assumption that the null hypothesis is

true.

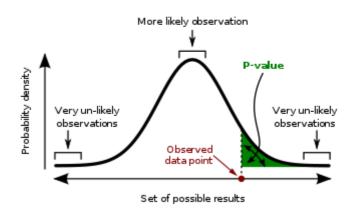
There is a problem with this definition (and indeed with the ASA^[1] informal definition: "a *p*value is the probability under a specified statistical model that a statistical summary of the data ... would be equal to or more extreme than its observed value"). In many situations, for example when testing the null hypothesis that a normal mean is less than or equal to zero against the alternative that the mean is greater than zero (variance known), the null hypothesis does not specify the probability distribution of the appropriate test statistic. In the just mentioned example that would be the Z-statistic belonging to the one-sided one-sample Z-test. For each possible value of the theoretical mean, the Z-test statistic has a different probability distribution. In these circumstances (the case of a so-called composite null hypothesis) the *p*-value is defined by taking the least favourable null-hypothesis case, which is typically on the border between null and alternative.

Important:

Pr (observation | hypothesis) ≠ Pr (hypothesis | observation)

The probability of observing a result given that some hypothesis is true is not equivalent to the probability that a hypothesis is true given that some result has been observed.

Using the p-value as a "score" is committing an egregious logical error: the transposed conditional fallacy.



A p-value (shaded green area) is the probability of an observed (or more extreme) result assuming that the null hypothesis is true.

Example of a *p*-value computation. The vertical coordinate is the probability density of each outcome, computed under the null hypothesis. The *p*-value is the area under the curve past the observed data point.

Misconceptions

There is widespread agreement that p-values are often misused and misinterpreted [1]. One practice that has been particularly criticized is accepting the alternative hypothesis for any p-value nominally less than .05 without other supporting evidence. Although p-values are helpful in assessing how incompatible the data are with a specified statistical model, contextual factors must also be considered, such as "the design of a study, the quality of the measurements, the external evidence for the phenomenon under study, and the validity of assumptions that underlie the data analysis [1]. Another concern is that the p-value is often misunderstood as being the probability that the null hypothesis is true. [1][7] Some statisticians have proposed replacing p-values with alternative measures of evidence, [1] such as confidence intervals [8][9] likelihood ratios [10][11] or Bayes factors, [12][13][14] but there is heated debate on the feasibility of these alternatives. [15][16] Others have suggested to remove fixed significance thresholds and to interpret p-values as continuous indices of the strength of evidence against the null hypothesis. [17][18] Yet others suggested to report alongside p-values the prior probability of a real effect that would be required to obtain a false positive risk (i.e. the probability that there is no real effect) below a pre-specified threshold (e.g. 5%). [19]

Usage

The p-value is widely used in <u>statistical hypothesis testing</u> specifically in <u>null hypothesis significance testing</u>. In this method, as part of <u>experimental design</u>, before performing the experiment, one first chooses a model (th<u>eull hypothesis</u>) and a threshold value for p, called the <u>significance level</u> of the test, traditionally 5% or $1\%^{20}$ and denoted as α . If the p-value is less than the chosen significance level (α), that suggests that the observed data is sufficiently inconsistent with the <u>null hypothesis</u> and that the null hypothesis may be rejected. However, that does not prove that the tested hypothesis is true. When the p-value is calculated correctly, this test guarantees that the <u>type I error rate</u> is at most α . For typical analysis, using the standard $\alpha = 0.05$ cutoff, the null hypothesis is rejected when p < 0.05 and not rejected when p > 0.05. The p-value does not, in itself, support reasoning about the probabilities of hypotheses but is only a tool for deciding whether to reject the null hypothesis.

Calculation

Usually, **X** is a <u>test statistic</u>, rather than any of the actual observations. A test statistic is the output of a <u>scalar</u> function of all the observations. This statistic provides a single number, such as the average or the <u>correlation coefficient</u>, that summarizes the characteristics of the data, in a way relevant to a particular inquiry. As such, the test statistic follows a distribution determined by the function used to define that test statistic and the distribution of the input observational data.

For the important case in which the data are hypothesized to follow the normal distribution, depending on the nature of the test statistic and thus the underlying hypothesis of the test statistic, different null hypothesis tests have been developed. Some such tests are <u>z-test</u> for <u>normal distribution</u>, <u>t-test</u> for <u>Student's t-distribution</u>, <u>f-test</u> for <u>f-distribution</u>. When the data do not follow a normal distribution, it can still be possible to approximate the distribution of these test statistics by a normal distribution by invoking the central limit theorem for large samples, as in the case of Pearson's chi-squared test

Thus computing a *p*-value requires a null hypothesis, a test statistic (together with deciding whether the researcher is performing a <u>one-tailed test</u> or a <u>two-tailed test</u>), and data. Even though computing the test statistic on given data may be easy, computing the sampling distribution under the null hypothesis, and then computing its <u>cumulative distribution function</u> (CDF) is often a difficult problem. Today, this computation is done using statistical software, often via numeric methods (rather than exact formulae), but, in the early and mid 20th century, this was instead done via tables of values, and one interpolated or extrapolated *p*-values from these discrete values. Rather than using a table of *p*-values, Fisher instead inverted the CDF, publishing a list of values of the test statistic for given fixed *p*-values; this corresponds to computing the quantile function (inverse CDF).

Distribution

When the null hypothesis is true, if it takes the form $H_0: \theta = \theta_0$, and the underlying random variable is continuous, then the <u>probability distribution</u> of the *p*-value is <u>uniform</u> on the interval [0,1]. By contrast, if the alternative hypothesis is true, the distribution is dependent on sample size and the true value of the parameter being studied.

The distribution of p-values for a group of studies is called a p-curve.^[22] The curve is affected by four factors: the proportion of studies that examined false null hypotheses, the <u>power</u> of the studies that investigated false null hypotheses, the alpha levels, and <u>publication bias</u> [23] A p-curve can be used to assess the reliability of scientific literature, such as by detecting publication bias or p-hacking.^{[22][24]}

Examples

Here a few simple examples follow each illustrating a potential pitfall.

One roll of a pair of dice

Suppose a researcher rolls a pair of <u>dice</u> once and assumes a null hypothesis that the dice are fair, not loaded or weighted toward any specific number/roll/result; uniform. The test statistic is "the sum of the rolled numbers" and is <u>one-tailed</u>. The researcher rolls the dice and observes that both dice show 6, yielding a test statistic of 12. The *p*-value of this outcome is 1/36 (because under the assumption of the null hypothesis, the test statistic is uniformly distributed) or about 0.028 (the highest test statistic out of $6\times6=36$ possible outcomes). If the researcher assumed a significance level of 0.05, this result would be deemed significant and the hypothesis that the dice are fair would be rejected.

In this case, a single roll provides a very weak basis (that is, insufficient data) to draw a meaningful conclusion about the dice. This illustrates the danger with blindly applying value without considering the experiment design

Five heads in a row

Suppose a researcher flips a coin five times in a row and assumes a null hypothesis that the coin is fair. The test statistic of "total number of heads" can be one-tailed or two-tailed: a one-tailed test corresponds to seeing if the coin is biased towards heads, but a two-tailed test corresponds to seeing if the coin is biased either way. The researcher flips the coin five times and observes heads each time (HHHHHH), yielding a test statistic of 5. In a one-tailed test, this is the upper extreme of all possible outcomes, and yields a *p*-value of $(1/2)^5 = 1/32 \approx 0.03$. If the researcher assumed a significance level of 0.05, this result would be deemed significant and the hypothesis that the coin is fair would be rejected. In a two-tailed test, a test statistic of zero heads (TTTTT) is just as extreme and thu the data of HHHHHH would yield a*p*-value of $2 \times (1/2)^5 = 1/16 \approx 0.06$, which is not significant at the 0.05 level.

This demonstrates that specifying a direction (on a symmetric test statistic) halves the *p*-value (increases the significance) and can mean the difference between data being considered significant or not.

Sample size dependence

Suppose a researcher flips a coin some arbitrary number of times (n) and assumes a null hypothesis that the coin is fair. The test statistic is the total number of heads and is a two-tailed test. Suppose the researcher observes heads for each flip, yielding a test statistic of n and a p-value of $2/2^n$. If the coin was flipped only 5 times, the p-value would be 2/32 = 0.0625, which is not significant at the 0.05 level. But if the coin was flipped 10 times, the p-value would be $2/1024 \approx 0.002$, which is significant at the 0.05 level.

In both cases the data suggest that the null hypothesis is false (that is, the coin is not fair somehow), but changing the sample size changes the *p*-value. In the first case, the sample size is not large enough to allow the null hypothesis to be rejected at the 0.05 level (in fact, the *p*-value can never be below 0.05 for the coin example).

This demonstrates that in interpretingp-values, one must also know the sample size, which complicates the analysis.

Alternating coin flips

Suppose a researcher flips a coin ten times and assumes a null hypothesis that the coin is fair. The test statistic is the total number of heads and is two-tailed. Suppose the researcher observes alternating heads and tails with every flip (HTHTHTHTHT). This yields a test statistic of 5 and a*p*-value of **1**, which is exactly what is expected.

Suppose instead that the test statistic for this experiment was the "number of alternations" (that is, the number of times when H followed T or T followed H), which is one-tailed. That would yield a test statistic of 9, which is extreme and has a p-value of $2/2^{10} = 1/512 \approx 0.0019$. That would be considered extremely significant, well beyond the 0.05 level. These data indicate that, in terms of one test statistic, the data set is extremely unlikely to have occurred by chance, but it does not suggest that the coin is biased towards heads or tails.

By the first test statistic, the data yield a high *p*-value, suggesting that the number of heads observed is not unlikely. By the second test statistic, the data yield a low *p*-value, suggesting that the pattern of flips observed is very, very unlikely. There is no "alternative hypothesis" (so only rejection of the null hypothesis is possible) and such data could have many causes. The data may instead be forged, or the coin may be flipped by a magicianwho intentionally alternated outcomes.

This example demonstrates that the *p*-value depends completely on the test statistic used and illustrates that *p*-values can only help researchers to reject a null hypothesis, not consider other hypotheses.

Coin flipping

As an example of a statistical test, an experiment is performed to determine whether a <u>coin flip</u> is <u>fair</u> (equal chance of landing heads or tails) or unfairly biased (one outcome being more likely than the other).

Suppose that the experimental results show the coin turning up heads 14 times out of 20 total flips. The null hypothesis is that the coin is fair, and the test statistic is the number of heads. If a right-tailed test is considered, the *p*-value of this result is the chance of a fair coin landing on heads*at least* 14 times out of 20 flips. That probability can be computed from binomial coefficients as

$$ext{Prob}(14 ext{ heads}) + ext{Prob}(15 ext{ heads}) + \cdots + ext{Prob}(20 ext{ heads}) \\ = rac{1}{2^{20}} \left[inom{20}{14} + inom{20}{15} + \cdots + inom{20}{20}
ight] = rac{60,460}{1,048,576} pprox 0.058$$

This probability is the *p*-value, considering only extreme results that favor heads. This is called a <u>one-tailed test</u>. However, the deviation can be in either direction, favoring either heads or tails. The two-tailed *p*-value, which considers deviations favoring either heads or tails, may instead be calculated. As the <u>binomial distribution</u> is symmetrical for a fair coin, the two-sided *p*-value is simply twice the above calculated single-sided*p*-value: the two-sided*p*-value is 0.115.

In the above example:

■ Null hypothesis (H₀): The coin is fair, with Prob(heads) = 0.5

• Test statistic: Number of heads

Level of significance: 0.05

Observation O: 14 heads out of 20 flips; and

■ Two-tailed p-value of observation O given $H_0 = 2*min(Prob(no. of heads <math>\geq 14 \text{ heads})$, Prob(no. of heads $\leq 14 \text{ heads}$) = 2*min(0.058, 0.978) = 2*0.058 = 0.115.

Note that the Prob(no. of heads ≤ 14 heads) = 1 - Prob(no. of heads ≥ 14 heads) + Prob(no. of head = 14) = 1 - 0.058 + 0.036 = 0.978; however, symmetry of the binomial distribution makes it an unnecessary computation to find the smaller of the two probabilities. Here, the calculated p-value exceeds 0.05, so the observation is consistent with the null hypothesis, as it falls within the range of what would happen 95% of the time were the coin in fact fair. Hence, the null hypothesis at the 5% level is not rejected. Although the coin did not fall evenly the deviation from the expected outcome is small enough to be consistent with chance.

However, had one more head been obtained, the resulting *p*-value (two-tailed) would have been 0.0414 (4.14%). The null hypothesis is rejected when a 5% cut-o**f** is used.

History

Computations of p-values date back to the 1700s, where they were computed for the <u>human sex ratio</u> at birth, and used to compute statistical significance compared to the null hypothesis of equal probability of male and female births. [25] <u>John Arbuthnot</u> studied this question in 1710, [26][27][28][29] and examined birth records in London for each of the 82 years from 1629 to 1710. In every year, the number of males born in London exceeded the number of females. Considering more male or more female births as equally likely, the probability of the observed outcome is 0.5^{82} , or about 1 in 4,836,000,000,000,000,000,000,000; in modern terms, the p-value. This is vanishingly small, leading Arbuthnot that this was not due to chance, but to divine providence: "From whence it follows, that it is Art, not Chance, that governs." In modern terms, he rejected the null hypothesis of equally likely male and female births at the $p = 1/2^{82}$ significance level. This is and other work by Arbuthnot is credited as "... the first use of significance tests ..." [30] the first example of reasoning about statistical significance, [31] and "... perhaps the first published report of a <u>nonparametric test</u> ...", [27] specifically the sign test; see details at Sign test § History.



John Arbuthnot

The same question was later addressed by <u>Pierre-Simon Laplace</u>, who instead used a *parametric* test, modeling the number of male births with a binomial distribution^[32]

In the 1770s Laplace considered the statistics of almost half a million births. The statistics showed an excess of boys compared to girls. He concluded by calculation of *p*-value that the excess was a real, but unexplained, effect.

The *p*-value was first formally introduced by Karl Pearson, in his Pearson's chi-squared test, [33] using the chi-squared distribution and notated as capital P. [33] The *p*-values for the chi-squared distribution (for various values of χ^2 and degrees of freedom), now notated as *P*, was calculated in (Elderton 1902), collected in (Pearson 1914, pp. xxxi–xxxiii, 26–28, Table XII).

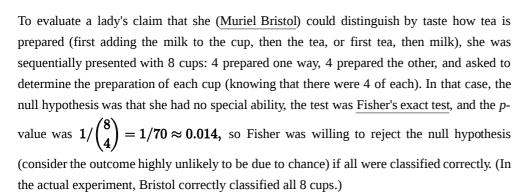
The use of the p-value in statistics was popularized by Ronald Fisher, [34] and it plays a central role in his approach to the subject. In his influential book Statistical Methods for Research Workers (1925), Fisher proposed the level p = 0.05, or a 1 in 20 chance of being exceeded by chance, as a limit for statistical significance, and applied this to a normal distribution (as a two-tailed test), thus yielding the rule of two standard deviations (on a normal distribution) for statistical significance (see68–95–99.7 rule). [36] [note 2] [37]



Pierre-Simon Laplace

He then computed a table of values, similar to Elderton but, importantly, reversed the roles of χ^2 and p. That is, rather than computing p for different values of χ^2 (and degrees of freedom n), he computed values of χ^2 that yield specified p-values, specifically 0.99, 0.98, 0.95, 0,90, 0.80, 0.70, 0.50, 0.30, 0.20, 0.10, 0.05, 0.02, and 0.01. That allowed computed values of χ^2 to be compared against cutoffs and encouraged the use of p-values (especially 0.05, 0.02, and 0.01) as cutoffs, instead of computing and reporting p-values themselves. The same type of tables were then compiled in (Fisher & Yates 1938), which cemented the approach [37]

As an illustration of the application of p-values to the design and interpretation of experiments, in his following book $\underline{The\ Design\ of\ Experiments}$ (1935), Fisher presented the $\underline{lady\ tasting\ tea}$ experiment, $\underline{lady\ tasting\ tea}$ experiment, $\underline{lady\ tasting\ tea}$ which is the archetypal example of the p-value.



Fisher reiterated the p = 0.05 threshold and explained its rationale, stating.

It is usual and convenient for experimenters to take 5 per cent as a standard level of significance, in the sense that they are prepared to ignore all results which fail to reach this standard, and, by this means, to eliminate from further discussion the greater part of the fluctuations which chance causes have introduced into their experimental results.



Karl Pearson



Ronald Fisher

He also applies this threshold to the design of experiments, noting that had only 6 cups been presented (3 of each), a perfect classification would have only yielded a p-value of $1/\binom{6}{3} = 1/20 = 0.05$, which would not have met this level of significance. Fisher also underlined the interpretation of p, as the long-run proportion of values at least as extreme as the data, assuming the null hypothesis is true.

In later editions, Fisher explicitly contrasted the use of the p-value for statistical inference in science with the Neyman–Pearson method, which he terms "Acceptance Procedures". Fisher emphasizes that while fixed levels such as 5%, 2%, and 1% are convenient, the exact p-value can be used, and the strength of evidence can and will be revised with further experimentation. In contrast, decision procedures require a clear-cut decision, yielding an irreversible action, and the procedure is based on costs of error, which, he argues, are inapplicable to scientific regarch.

Related quantities

A closely related concept is the E-value, ^{42]} which is the <u>expected</u> number of times in <u>multiple testing</u> that one expects to obtain a test statistic at least as extreme as the one that was actually observed if one assumes that the null hypothesis is true. The E-value is the product of the number of tests and the p-value.

The q-value is the analog of the p-value with respect to the positive false discovery rate [43] It is used in multiple hypothesis testing to maintain statistical power while minimizing the false positive rate [44]

See also

- Bonferroni correction
- Confidence interval
- Counternull
- False discovery rate
- Fisher's method of combiningp-values
- Generalized p-value
- Holm–Bonferroni method
- Multiple comparisons
- Null hypothesis
- p-rep
- p-value fallacy
- Statistical hypothesis testing
- q-value

Notes

- 1. Note that the statistical significance of a result does not imply that the result is scientifically significant as well.
- 2. To be precise the p=0.05 corresponds to about 1.96 standard deviations for a normal distribution (two-tailed test), and 2 standard deviations corresponds to about a 1 in 22 chance of being exceeded by chance, $pr \approx 0.045$; Fisher notes these approximations.

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External links

- Free online p-values calculators for various specific tests (chi-square, Fisher's F-test, etc.).
- Understanding *p*-values, including a Java applet that illustrates how the numerical values q6-values can give quite misleading impressions about the truth or falsity of the hypothesis under test.
- StatQuest: P Values, clearly explained on YouTube
- StatQuest: P-value pitfalls and power calculationson YouTube

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