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- ► As we shall see, Poisson process is a special case of continuous time Markov chain

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- ▶ It is called homogeneous chain if

$$Pr[X(t+s)=j\mid X(s)=i]=Pr[X(t)=j\mid X(0)=i],\;\forall s$$



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- ▶ Analogous to transition probabilities in the discrete case
- ► Like in the discrete case, we can show that the Markov condition implies

$$Pr[X(s) \in B_s, s \in (t, t + \tau] \mid X(s) = i, X(s'), \ 0 \le s' < s < t]$$

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- Note that there is no concept of a transition from i to i.
- ► Let us look at the distribution of time spent in a state before leaving it

$$Pr[X(s) = i, \ s \in [t, \ t + \tau] \mid X(s') = i, \ 0 \le s' \le t]$$

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 $\Rightarrow T_i$ is memoryless and hence exponential

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- Note that $P_{ij}(t)$ is different from these z_{ij}

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- ▶ Before we calculate these, we need to recall some properties of exponential random variables

$$Pr[X < Y] = \int_0^\infty \int_0^y \lambda_1 e^{-\lambda_1 x} \lambda_2 e^{-\lambda_2 y} dx dy$$

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We can also get this using

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▶ Let $Z = \min(X, Y)$. Then

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Thus, $\min(X, Y) \sim \exp(\lambda_1 + \lambda_2)$. 4□ > 4回 > 4回 > 4 = > = 900 P S Sastry, IISc, E1 222, Lecture 27, Aug. 2021 10/39 lackbox Consider the birth-death process with birth rate λ_n and death rate μ_n in state n

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- \blacktriangleright Hence, $\mu_0=0$ and $\nu_0=\lambda_0$ and $z_{01}=1$

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- ightharpoonup A variation: M/M/K queue

- Consider a queuing system
- \blacktriangleright Suppose people joining the queue is a Poisson process with rate λ
- ▶ Suppose the time to service each customer is independent and exponential with parmeter μ .
- ► We assume that the arrival and service processes are independent.
- ► Then this is a birth death process with

$$\lambda_n = \lambda, \ n \ge 0$$
 and $\mu_n = \mu, \ n \ge 1$

- ightharpoonup This is known as an M/M/1 queue
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$$\lambda_n = \lambda, \ n \ge 0 \quad \text{ and } \quad \mu_n = \left\{ \begin{array}{ll} n\mu & 1 \le n \le K \\ K\mu & n > K \end{array} \right.$$

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$$I_i = \left\{ \begin{array}{ll} 1 & \text{if first transition out of } i \text{ is to } i+1 \\ 0 & \text{if first transition out of } i \text{ is to } i-1 \end{array} \right.$$

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$$E[Y_i \mid I_i = 0] = \frac{1}{\lambda_i + \mu_i} + E[Y_{i-1}] + E[Y_i]$$

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lackbox We can also calculate expected time to go from i to j for i < j as

$$E[Y_i] + E[Y_{i+1}] + \cdots + E[Y_{i-1}]$$

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▶ Note that all these are only for birth-death processes

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- ▶ We need a technical assumption.
- ▶ We assume that the transition probability function is such that the probability of infinite number of transitions in a finite interval of time is zero. (Called the regularity condition).

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$$\Rightarrow \lim_{h \to 0} \frac{P_{ij}(t+h) - P_{ij}(t)}{h} = \sum_{k \neq i} \lim_{h \to 0} \frac{P_{ik}(h)}{h} P_{kj}(t) - \lim_{h \to 0} \frac{1 - P_{ii}(h)}{h} P_{ij}(t)$$

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(For an infinite chain there is the issue of whether we can take the limit inside the summation)

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▶ Hence we can write the earlier equation as

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by taking $q_{ii} = -\sum_{k \neq i} q_{ik}$

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- ▶ Hence we specify the transition functions of a continuous time Markov chain through the q_{ij} .
- ► These are referred to as the infinitesimal generator of the chain.

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▶ Thus $\sum_{i\neq i} q_{ij} = \nu_i$. It is rate of transition out of i

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▶ Hence we can calculate $P_{ij}(t)$ for any t and i, j

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► For a birth-death chain the equation becomes

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- Now one can see that the above ODE is what we got for Poisson process.

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- ▶ We may want to calculate $P_{00}(T)$, the probability that the machine would be working at a time T units later given it is in working condition now
- ▶ We can calculate it by solving the ODE's

$$P'_{ij}(t) = \lambda_i P_{i+1,j}(t) + \mu_i P_{i-1,j}(t) - (\lambda_i + \mu_i) P_{ij}(t)$$

For the two state chain, these equations are

$$P'_{00}(t) = \lambda_0 P_{10}(t) - \lambda_0 P_{00}(t)$$

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- ► Solving these we can show

$$P_{00}(t) = \frac{\lambda}{\lambda + \mu} e^{-(\lambda + \mu)t} + \frac{\mu}{\lambda + \mu}$$

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