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Probability axioms

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$$P:\mathcal{F} \to \Re, \ \mathcal{F} \subset 2^{\Omega}$$

A1
$$P(A) > 0$$
, $\forall A \in \mathcal{F}$

A2
$$P(\Omega) = 1$$

A3 If
$$A_i \cap A_j = \phi, \forall i \neq j$$
 then $P(\bigcup_{i=1}^{\infty} A_i) = \sum_{i=1}^{\infty} P(A_i)$

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- Thus, 'equally likely' is one way of specifying the probability function (in case of finite Ω).
- ► An obvious point worth remembering: specifying *P* for singleton events fixes it for all other events.

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- \triangleright This can be done for finite Ω too.

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- Consider a random experiment of tossing a biased coin repeatedly till we get a head. We take the outcome of the experiment to be the number of tails we had before the first head.
- ► A (reasonable) probability assignment is:

$$P({k}) = (1-p)^k p, k = 0, 1, \cdots$$

where p is the probability of head and 0 . (We assume you understand the idea of 'independent' tosses here).

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 - (There are many issues that need more attention here).

Problem: A rod of unit length is broken at two random points. What is the probability that the three pieces so formed would make a triangle.

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- ► For the pieces to make a triangle, sum of lengths of any two should be more than the third.

▶ The lengths are: x, (y - x), (1 - y).

$$x + (y - x) > (1 - y) \Rightarrow y > 0.5$$

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► So the event of interest is:

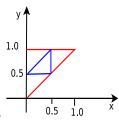
$$A = \{(x, y) : y > 0.5; x < 0.5; y < x + 0.5, 0 < x, y < 1\}$$

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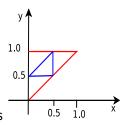
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- ► We can visualize it as follows
- ► The required probability is area of A divided by area of Ω which gives the answer as 0.25

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Let B be an event with P(B) > 0. We define conditional probability, conditioned on B, of any event, A, as

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- ▶ If A_1 , A_2 are mutually exclusive then A_1B and A_2B are also mutually exclusive and hence

$$P_B(A_1 + A_2) = \frac{P((A_1 + A_2)B)}{P(B)}$$

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Once we understand condional probability is a new probability assignment, we go back to the 'standard notation'

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- This is a useful intuition as long as we understand it properly.
- ▶ It is not as if we talk about conditional probability only for subsets of B. Conditional probability is also with respect to the original probability space. Every element of F has conditional probability defined.

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- ► Hence, conditional probabilities cannot actually capture causal influences.
- ► There are probabilistic methods to capture causation (but far beyond the scope of this course!)

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► This is a very useful in many situations. ("arguing by cases")

An urn contains r red balls and b black balls. We draw a ball at random, note its color, and put back that ball along with c balls of the same color. We keep repeating this process. Let R_n (B_n) denote the event of drawing a red (black) ball at the n^{th} draw. We want to calculate the probabilities of all these events.

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- ► This does not depend on the value of c!

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- ▶ $P(T_+|D)$ is called the true positive rate and $P(T_+|D^c)$ is called false positive rate.
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► These different cases are important in understanding the role of false positives rate.

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- ▶ Bayes rule essentially transforms the prior probability to posterior probability.

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 - ▶ The knowledge we need is $P(T_+|D)$, $P(T_+|D^c)$ which can be determined through experiment or modelling of channel.

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► Example: I have three coins with probability of heads being 0.1, 0.5, 0.8. I choose one at random and toss it twice and see heads both times. What is the probability it is the fair coin?

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- ▶ This gives an intuitive feel for independence.
- ▶ Independence is an important (often confusing!) concept.

A class has 20 female and 30 male course (MTech) students and 6 female and 9 male research (PhD) students. Are gender and degree independent?

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- Once that is given, the probabilities of all events are fixed.
- ► Hence whether or not two events are independent is a matter of 'calculation'

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- For example, in the previous problem, once we saw that F and C are independent, we can conclude M and C are also independent (because in this example we are taking F^c = M).

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- $\Omega = \{HH, HT, TH, TT\}.$ Suppose we employ 'equally likely idea'.
- ► That is, $P({HH}) = \frac{1}{4}$, $P({HT}) = \frac{1}{4}$ and so on
- ► Let $A = \text{`H on 1st toss'} = \{HH, HT\} \ (P(A) = \frac{1}{2})$ Let $B = \text{`T on second toss'} = \{HT, TT\} \ (P(B) = \frac{1}{2})$
- We have $P(AB) = P(\{HT\}) = 0.25$
- Since $P(A)P(B) = \frac{1}{2} \frac{1}{2} = \frac{1}{4} = P(AB)$, A, B are independent.
- ► Hence, in multiple tosses, assuming all outcomes are equally likely implies outcome of one toss is independent of another.

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- ▶ If we assume tosses are independent then we can assign probabilities easily.

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 - (We will look at it more formally when we consider multiple random variables).

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- As we saw, if A and B are independent, then P(A|B) = P(A)
- ► This is often used, at an intuitive level, to justify assumption of independence.

▶ Events A_1, A_2, \dots, A_n are said to be (totally) independent if for any k, $1 \le k \le n$, and any indices i_1, \dots, i_k , we have

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- ▶ But, $P(E_1E_2E_3) = 0.25 \neq (0.5)^3$

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- Events may be conditionally independent but not independent. (e.g., 'independent' multiple tests for confirming a disease)
- ▶ It is also possible that *A*, *B* are independent but are not conditionally independent given some other event *C*.

$$P(A|BC) = \frac{P(BC|A)P(A)}{P(BC|A)P(A) + P(BC|A^c)P(A^c)}$$

We can write Bayes rule with multiple conditioning events.

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- ▶ Take: A = D, $B = T_+^1$, $C = T_+^2$.
- Assuming conditional independence we can calculate the new posterior probability using the same information we had about true positive and false positive rate.

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$$= \frac{0.99 * 0.99 * 0.1}{0.99 * 0.99 * 0.1 + 0.05 * 0.05 * 0.9} = 0.97$$