## Recap: Random Variables

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- It essentially results in an induced probability space

$$(\Omega, \mathcal{F}, P) \stackrel{X}{\to} (\Re, \mathcal{B}, P_X)$$

where  ${\cal B}$  is the Borel  $\sigma$ -algebra and

$$P_X(B) = P[X \in B] = P(\{\omega \in \Omega : X(\omega) \in B\})$$



- ▶ An  $\mathcal{F} \subset 2^{\Omega}$  is called a  $\sigma$ -algebra (also called  $\sigma$ -field) on  $\Omega$  if it satisfies
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  - 2.  $A \in \mathcal{F} \Rightarrow A^c \in \mathcal{F}$
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- ▶ We also have  $\mathcal{B} = \sigma(\{(-\infty, x] : x \in \Re\})$

# Recap: Distribution function of a random variable

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- ▶ The distribution function,  $F_X$ , completely specifies the probability measure,  $P_X$ .
- ▶ Thus, given the distribution function of X, one can (in principle) calculate probability of any event  $[X \in B]$ .

- ► The distribution function satisfies
  - 1.  $0 \le F_X(x) \le 1, \ \forall x$
  - 2.  $F_X(-\infty) = 0$ ;  $F_X(\infty) = 1$
  - 3.  $F_X$  is non-decreasing:  $x_1 \le x_2 \implies F_X(x_1) \le F_X(x_2)$
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 $P[a < X \le b] = F_X(b) - F_X(a).$ 



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- Note that the distribution function is defined for all random variables.

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$$\begin{split} [X \leq 1.57] &= \{\omega \ : \ X(\omega) \leq 1.57\} \\ &= \{\omega \ : \ X(\omega) = 0\} \cup \{\omega \ : \ X(\omega) = 1\} = [X = 0 \text{ or } 1] \end{split}$$



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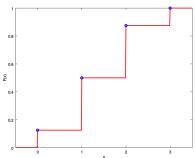
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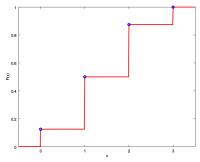
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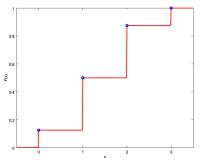


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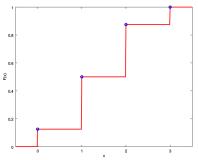
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- ▶ The jump at, e.g., x=2 is 3/8 which is the probability of X taking that value.

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- ► All discrete random variables would have this general form of distribution function.

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- ▶ If  $a_2 \le x < a_3$  then  $[X \le x] = [X = a_1] \cup [X = a_2] = B_1 + B_2$

▶ Hence we can write the distribution function as

$$F_X(x) = \begin{cases} 0 & x < a_1 \\ P(B_1) & a_1 \le x < a_2 \\ P(B_1) + P(B_2) & a_2 \le x < a_3 \end{cases}$$

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We can get pmf from df and df from pmf.



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- Please remember that we have defined distribution function for any random variable. But pmf is defined only for discrete random variables



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- As we saw this is how we can specify a probability assignment on any countable sample space.

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We next look at some standard discrete random variable models



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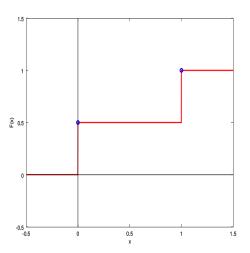
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- $P[I_B = 1] = P(\{\omega : I_B(\omega) = 1\}) = P(B)$
- ▶ Thus, this indicator rv has Bernoulli distribution with p = P(B)



#### One of the df examples we saw earlier is that of Bernoulli



 $ightharpoonup X \in \{0, 1, \dots, n\}$  with pmf

$$f_X(k) = {}^{n}C_k p^k (1-p)^{n-k}, k = 0, 1, \dots, n$$

where n, p are parameters (n is a +ve integer and 0 ).

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► Consider *n* independent tosses of coin whose probability of heads is *p*. If *X* is the number of heads then *X* has the above binomial distribution.

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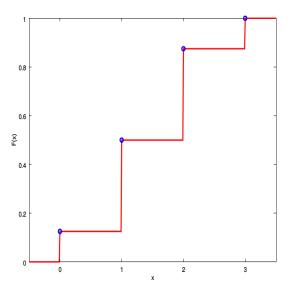
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- Consider n independent tosses of coin whose probability of heads is p. If X is the number of heads then X has the above binomial distribution.
  - (Number of successes in n bernoulli trials)
- Any one outcome (a seq of length n) with k heads would have probability  $p^k(1-p)^{n-k}$ . There are  ${}^nC_k$  outcomes with exactly k heads.



#### One of the df examples we considered was that of Binomial



 $ightharpoonup X \in \{0, 1, 2, \cdots\}$  with pmf

$$f_X(k) = \frac{\lambda^k e^{-\lambda}}{k!}, \ k = 0, 1, 2, \dots$$

where  $\lambda > 0$  is a parameter.

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Poisson distribution is also useful in many applications



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- ▶ In general waiting for 'success' in independent Bernoulli trials.



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(Does this also tell us what is df of geometric rv?)



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▶ Suppose  $X \in \{0, 1, \cdots\}$  is a discrete rv satisfying, for all non-negative integers, m, n

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- ▶ Let us take P[X > 0] = 1 (and hence P[X = 0] = 0).



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=  $P[X > m-1]P[X > 1]$ 

 $\blacktriangleright$  We have, for any m,

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A function  $g:\Re\to\Re$  is absolutely continuous on an interval, I, if given any  $\epsilon>0$  there is a  $\delta>0$  such that for any finite sequence of pair-wise disjoint subintervals,  $(x_k,y_k)$ , with  $x_k,y_k\in I,\ \forall k$ , satisfying  $\sum_k(y_k-x_k)<\delta$ , we have  $\sum_k|f(y_k)-f(x_k)|<\epsilon$ 

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▶ In the above, g would be differentiable almost everywhere and h would be its derivative (wherever g is differentiable).



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- ▶ By the fundamental theorem of calculus, we have

$$\frac{dF_x(x)}{dx} = f_X(x), \ \forall x \text{ where } f_X \text{ is continuous}$$



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- This implies  $F_X(x) = F_X(x^+) = F_X(x^-)$
- ▶ Hence, if *X* is a continuous random variable then

$$P[X = x] = F_X(x) - F_X(x^-) = 0, \ \forall x$$



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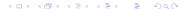
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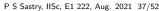
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- $\blacktriangleright$  Let us find  $F_Y$  and  $f_Y$ .



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▶ We have  $F_X = F_Y$  and thus  $f_X = f_Y$ . (However, note that  $X(\omega) \neq Y(\omega)$  except at  $\omega = 0.5$ ).

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Recall that for a general rv

$$F_X(b) - F_X(a) = P[a < X \le b]$$



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 $\blacktriangleright$  That is why  $f_X$  is called probability density function.

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- ▶ Writing  $f_X(X=5)$  when  $f_X$  is a pdf, is particularly bad. Note that for a continuous rv, P[X=5]=0 and  $f_X(5) \neq P[X=5]$ .

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- We next consider a few standard continuous random variables.

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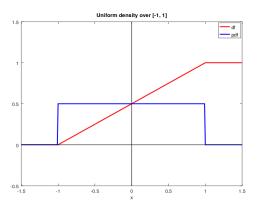
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- ► The earlier examples we considered are uniform over [0, 1].

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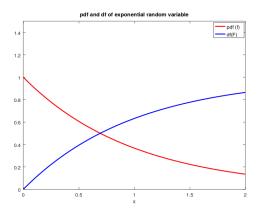
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This also gives us:  $P[X > x] = 1 - F_X(x) = e^{-\lambda x}$  for x > 0.



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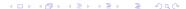
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- ► Exponential distribution is a useful model for, e.g., life-time of components.
- ► If the distribution of a non-negative continuous random variable is memory less then it must be exponential.



▶ The pdf of Gaussian distribution is given by

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty$$

where  $\sigma > 0$  and  $\mu \in \Re$  are parameters.

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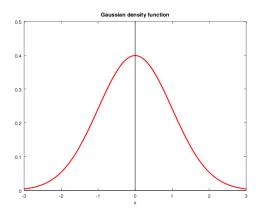
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- The special case where  $\mu=0$  and  $\sigma^2=1$  is called standard Gaussian (or standard Normal) distribution.



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- ► Take  $\mu = 0, \sigma = 1$ . Let  $I = \int_{-\infty}^{\infty} f_X(x) \ dx$ . Then

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Now converting the above integral into polar coordinates would allow you to show I=1.

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Now converting the above integral into polar coordinates would allow you to show I=1. (Left as an exercise for you!)