

# Recap: Random Processes

- ▶ A random process or a stochastic process is a collection of random variables:  $\{X(t), t \in T\}$

# Recap: Random Processes

- ▶ A random process or a stochastic process is a collection of random variables:  $\{X(t), t \in T\}$
- ▶ The index set  $T$  can be countable or uncountable (discrete-time or continuous-time processes)

# Recap: Random Processes

- ▶ A random process or a stochastic process is a collection of random variables:  $\{X(t), t \in T\}$
- ▶ The index set  $T$  can be countable or uncountable (discrete-time or continuous-time processes)
- ▶ The  $X(t)$  can be discrete or continuous (discrete-state or continuous-state processes)

# Recap: Random Processes

- ▶ A random process or a stochastic process is a collection of random variables:  $\{X(t), t \in T\}$
- ▶ The index set  $T$  can be countable or uncountable (discrete-time or continuous-time processes)
- ▶ The  $X(t)$  can be discrete or continuous (discrete-state or continuous-state processes)
- ▶ We can view the process as  $X : \Omega \times T \rightarrow \Re$   
A collection of time functions.

# Recap: Distributions of a random process

- ▶ The first order distributions:

$$F_X(x; t) = Pr[X(t) \leq x] = F_{X(t)}(x)$$

# Recap: Distributions of a random process

- ▶ The first order distributions:

$$F_X(x; t) = Pr[X(t) \leq x] = F_{X(t)}(x)$$

- ▶ The second order distributions:

$$F_X(x_1, x_2; t_1, t_2) = Pr[X(t_1) \leq x_1, X(t_2) \leq x_2]$$

# Recap: Distributions of a random process

- ▶ The first order distributions:

$$F_X(x; t) = Pr[X(t) \leq x] = F_{X(t)}(x)$$

- ▶ The second order distributions:

$$F_X(x_1, x_2; t_1, t_2) = Pr[X(t_1) \leq x_1, X(t_2) \leq x_2]$$

- ▶ The  $n^{th}$  order distributions:

$$F_X(x_1, \dots, x_n; t_1, \dots, t_n) = Pr[X(t_i) \leq x_i, i = 1, \dots, n]$$

# Recap: Distributions of a random process

- ▶ The first order distributions:

$$F_X(x; t) = Pr[X(t) \leq x] = F_{X(t)}(x)$$

- ▶ The second order distributions:

$$F_X(x_1, x_2; t_1, t_2) = Pr[X(t_1) \leq x_1, X(t_2) \leq x_2]$$

- ▶ The  $n^{th}$  order distributions:

$$F_X(x_1, \dots, x_n; t_1, \dots, t_n) = Pr[X(t_i) \leq x_i, i = 1, \dots, n]$$

- ▶ The distributions can be specified through joint mass or density functions too.



# Recap

- ▶ One often makes some assumptions on the process so that all  $n^{th}$  order distributions are easily specified implicitly.

# Recap

- ▶ One often makes some assumptions on the process so that all  $n^{th}$  order distributions are easily specified implicitly.
- ▶ One example is the Markovian dependence

# Recap

- ▶ One often makes some assumptions on the process so that all  $n^{th}$  order distributions are easily specified implicitly.
- ▶ One example is the Markovian dependence
- ▶ Other examples: process with independent increments, Gaussian processes

# Recap: Mean, Autocorrelation, autocovariance

# Recap: Mean, Autocorrelation, autocovariance

- ▶ The mean or mean function is

$$\eta_X(t) = E[X(t)], \quad t \in T$$

# Recap: Mean, Autocorrelation, autocovariance

- ▶ The mean or mean function is

$$\eta_X(t) = E[X(t)], \quad t \in T$$

- ▶ The autocorrelation of the process is

$$R_X(t_1, t_2) = E[X(t_1)X(t_2)]$$

# Recap: Mean, Autocorrelation, autocovariance

- ▶ The mean or mean function is

$$\eta_X(t) = E[X(t)], \quad t \in T$$

- ▶ The autocorrelation of the process is

$$R_X(t_1, t_2) = E[X(t_1)X(t_2)]$$

- ▶ The autocovariance of the process is

$$C_X(t_1, t_2) = E[(X(t_1) - E[X(t_1)])(X(t_2) - E[X(t_2)])]$$

# Recap: Mean, Autocorrelation, autocovariance

- ▶ The mean or mean function is

$$\eta_X(t) = E[X(t)], \quad t \in T$$

- ▶ The autocorrelation of the process is

$$R_X(t_1, t_2) = E[X(t_1)X(t_2)]$$

- ▶ The autocovariance of the process is

$$\begin{aligned} C_X(t_1, t_2) &= E[(X(t_1) - E[X(t_1)])(X(t_2) - E[X(t_2)])] \\ &= R_X(t_1, t_2) - \eta_X(t_1)\eta_X(t_2) \end{aligned}$$



# Recap: Stationarity

# Recap: Stationarity

- ▶ A random process  $\{X(t), t \in T\}$  is said to be stationary if

# Recap: Stationarity

- ▶ A random process  $\{X(t), t \in T\}$  is said to be stationary if

*for all  $n$ , for all  $t_1, \dots, t_n$ , for all  $x_1, \dots, x_n$  and for all  $\tau$  we have*

$$F_X(x_1, \dots, x_n; t_1, \dots, t_n) = F_X(x_1, \dots, x_n; t_1 + \tau, \dots, t_n + \tau)$$

# Recap: Stationarity

- ▶ A random process  $\{X(t), t \in T\}$  is said to be stationary if

*for all  $n$ , for all  $t_1, \dots, t_n$ , for all  $x_1, \dots, x_n$  and for all  $\tau$  we have*

$$F_X(x_1, \dots, x_n; t_1, \dots, t_n) = F_X(x_1, \dots, x_n; t_1 + \tau, \dots, t_n + \tau)$$

- ▶ For a stationary process, the distributions are unaffected by translation of the time axis.

# Recap: Stationarity

- ▶ A random process  $\{X(t), t \in T\}$  is said to be stationary if

*for all  $n$ , for all  $t_1, \dots, t_n$ , for all  $x_1, \dots, x_n$  and for all  $\tau$  we have*

$$F_X(x_1, \dots, x_n; t_1, \dots, t_n) = F_X(x_1, \dots, x_n; t_1 + \tau, \dots, t_n + \tau)$$

- ▶ For a stationary process, the distributions are unaffected by translation of the time axis.
- ▶ This stringent condition is often referred to as strict-sense stationarity

## Recap: Wide-sense stationarity

- ▶  $\{X(t), t \in T\}$  is said to be wide-sense stationary if

# Recap: Wide-sense stationarity

- ▶  $\{X(t), t \in T\}$  is said to be wide-sense stationary if
  1.  $\eta_X(t) = \eta_X$ , a constant

# Recap: Wide-sense stationarity

- ▶  $\{X(t), t \in T\}$  is said to be wide-sense stationary if
  1.  $\eta_X(t) = \eta_X$ , a constant
  2.  $R_X(t_1, t_2)$  depends only on  $t_1 - t_2$



## Recap: Wide-sense stationarity

- ▶  $\{X(t), t \in T\}$  is said to be wide-sense stationary if
  1.  $\eta_X(t) = \eta_X$ , a constant
  2.  $R_X(t_1, t_2)$  depends only on  $t_1 - t_2$
- ▶ This would be so if the first and second order distributions are invariant to change of time origin.

## Recap: Wide-sense stationarity

- ▶  $\{X(t), t \in T\}$  is said to be wide-sense stationary if
  1.  $\eta_X(t) = \eta_X$ , a constant
  2.  $R_X(t_1, t_2)$  depends only on  $t_1 - t_2$
- ▶ This would be so if the first and second order distributions are invariant to change of time origin.
- ▶ For a wide-sense stationary process the autocorrelation is a symmetric function and its Fourier transform is the power spectral density

# Recap: Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary

# Recap: Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary
- ▶ Ergodicity is a question of whether time-averages converge to ensemble-averages

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X(i) \stackrel{?}{=} E[X(n)] = \eta_X$$

# Recap: Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary
- ▶ Ergodicity is a question of whether time-averages converge to ensemble-averages

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X(i) \stackrel{?}{=} E[X(n)] = \eta_X$$

Or, more generally

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n g(X(i)) \stackrel{?}{=} E[g(X(n))]$$

# Recap: Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary
- ▶ Ergodicity is a question of whether time-averages converge to ensemble-averages

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X(i) \stackrel{?}{=} E[X(n)] = \eta_X$$

Or, more generally

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n g(X(i)) \stackrel{?}{=} E[g(X(n))]$$

For a continuous time process we can write this as

$$\lim_{\tau \rightarrow \infty} \frac{1}{2\tau} \int_{-\tau}^{\tau} X(t) dt \stackrel{?}{=} E[X(t)] = \eta_X$$

# Recap: Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary
- ▶ Ergodicity is a question of whether time-averages converge to ensemble-averages

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X(i) \stackrel{?}{=} E[X(n)] = \eta_X$$

Or, more generally

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n g(X(i)) \stackrel{?}{=} E[g(X(n))]$$

For a continuous time process we can write this as

$$\lim_{\tau \rightarrow \infty} \frac{1}{2\tau} \int_{-\tau}^{\tau} X(t) dt \stackrel{?}{=} E[X(t)] = \eta_X$$

- ▶ One sufficient condition is that covariance between  $X(t)$  and  $X(t + \tau)$  decreases fast with increasing  $\tau$ .

# Recap: Mean Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary with  $E[X(t)] = \eta$ .



# Recap: Mean Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary with  $E[X(t)] = \eta$ .
- ▶ Define

$$\eta_\tau = \frac{1}{2\tau} \int_{-\tau}^{\tau} X(t) dt \quad (\tau > 0)$$

# Recap: Mean Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary with  $E[X(t)] = \eta$ .
- ▶ Define

$$\eta_\tau = \frac{1}{2\tau} \int_{-\tau}^{\tau} X(t) dt \quad (\tau > 0)$$

- ▶ We say the process is mean-ergodic if

$$\eta_\tau \xrightarrow{P} \eta, \quad \text{as } \tau \rightarrow \infty$$

# Recap: Mean Ergodicity

- ▶ Let  $X(t)$  be wide-sense stationary with  $E[X(t)] = \eta$ .
- ▶ Define

$$\eta_\tau = \frac{1}{2\tau} \int_{-\tau}^{\tau} X(t) dt \quad (\tau > 0)$$

- ▶ We say the process is mean-ergodic if

$$\eta_\tau \xrightarrow{P} \eta, \quad \text{as } \tau \rightarrow \infty$$

- ▶ We showed this holds if

$$\int_{-\infty}^{\infty} |C_X(z)| dz < \infty$$

- ▶ Similar sufficient condition holds in case of discrete time processes also.

- ▶ Similar sufficient condition holds in case of discrete time processes also.
- ▶ A wide-sense stationary process  $\{X(n), n = 0, 1, \dots\}$  is said to be mean ergodic if

$$\frac{1}{n} \sum_{i=0}^{n-1} X(i) \xrightarrow{P} EX(n) = \eta$$

- ▶ Similar sufficient condition holds in case of discrete time processes also.
- ▶ A wide-sense stationary process  $\{X(n), n = 0, 1, \dots\}$  is said to be mean ergodic if

$$\frac{1}{n} \sum_{i=0}^{n-1} X(i) \xrightarrow{P} EX(n) = \eta$$

- ▶ Note that this is a generalization of (weak) law of large numbers to the case where the random variables may not be uncorrelated.

- ▶ Similar sufficient condition holds in case of discrete time processes also.
- ▶ A wide-sense stationary process  $\{X(n), n = 0, 1, \dots\}$  is said to be mean ergodic if

$$\frac{1}{n} \sum_{i=0}^{n-1} X(i) \xrightarrow{P} EX(n) = \eta$$

- ▶ Note that this is a generalization of (weak) law of large numbers to the case where the random variables may not be uncorrelated.
- ▶ The above holds if

$$\sum_{k=0}^{\infty} |C_X(k)| \leq \infty, \quad \text{where} \quad C_X(k) = \text{Cov}(X(n), X(n+k))$$

- ▶ Similar sufficient condition holds in case of discrete time processes also.
- ▶ A wide-sense stationary process  $\{X(n), n = 0, 1, \dots\}$  is said to be mean ergodic if

$$\frac{1}{n} \sum_{i=0}^{n-1} X(i) \xrightarrow{P} EX(n) = \eta$$

- ▶ Note that this is a generalization of (weak) law of large numbers to the case where the random variables may not be uncorrelated.
- ▶ The above holds if

$$\sum_{k=0}^{\infty} |C_X(k)| \leq \infty, \quad \text{where } C_X(k) = \text{Cov}(X(n), X(n+k))$$

- ▶ When the above holds we say the process is asymptotically uncorrelated.



- ▶ The proof in the discrete case is similar to that in the continuous case.

- ▶ The proof in the discrete case is similar to that in the continuous case.
- ▶ Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

- ▶ The proof in the discrete case is similar to that in the continuous case.
- ▶ Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var}(S_n) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} \text{Cov}(X(i), X(j))$$

- ▶ The proof in the discrete case is similar to that in the continuous case.
- ▶ Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var}(S_n) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} \text{Cov}(X(i), X(j)) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} C_X(i-j)$$

- ▶ The proof in the discrete case is similar to that in the continuous case.
- ▶ Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var}(S_n) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} \text{Cov}(X(i), X(j)) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} C_X(i - j)$$

- ▶ The above sum can be viewed as summing all the elements in a (Toeplitz) matrix where each (left-right) diagonal has all entries same.

- ▶ The proof in the discrete case is similar to that in the continuous case.
- ▶ Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var}(S_n) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} \text{Cov}(X(i), X(j)) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} C_X(i - j)$$

- ▶ The above sum can be viewed as summing all the elements in a (Toeplitz) matrix where each (left-right) diagonal has all entries same.
- ▶ Thus the sum can be rewritten as

$$\text{Var}(S_n) = \sum_{i=0}^{n-1} \sum_{j=0}^{n-1} C_X(i - j) = \sum_{k=-(n-1)}^{n-1} (n - |k|) C_X(k)$$



► Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .



► Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var} \left( \frac{S_n}{n} \right) = \frac{1}{n^2} \sum_{k=-(n-1)}^{n-1} (n-|k|) C_X(k)$$

► Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var} \left( \frac{S_n}{n} \right) = \frac{1}{n^2} \sum_{k=-(n-1)}^{n-1} (n-|k|) C_X(k) = \frac{1}{n} \sum_{k=-(n-1)}^{n-1} \left( 1 - \frac{|k|}{n} \right) C_X(k)$$

► Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var} \left( \frac{S_n}{n} \right) = \frac{1}{n^2} \sum_{k=-(n-1)}^{n-1} (n-|k|) C_X(k) = \frac{1}{n} \sum_{k=-(n-1)}^{n-1} \left( 1 - \frac{|k|}{n} \right) C_X(k)$$

► Now, using  $\sum_k |C_X(k)| < \infty$  we can show that

$$\lim_{n \rightarrow \infty} \text{Var} \left( \frac{S_n}{n} \right) = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=-(n-1)}^{n-1} \left( 1 - \frac{|k|}{n} \right) C_X(k) = 0$$

► Let  $S_n = \sum_{i=0}^{n-1} X(i)$ .

$$\text{Var} \left( \frac{S_n}{n} \right) = \frac{1}{n^2} \sum_{k=-(n-1)}^{n-1} (n-|k|) C_X(k) = \frac{1}{n} \sum_{k=-(n-1)}^{n-1} \left( 1 - \frac{|k|}{n} \right) C_X(k)$$

► Now, using  $\sum_k |C_X(k)| < \infty$  we can show that

$$\lim_{n \rightarrow \infty} \text{Var} \left( \frac{S_n}{n} \right) = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=-(n-1)}^{n-1} \left( 1 - \frac{|k|}{n} \right) C_X(k) = 0$$

► This shows that  $\frac{S_n}{n}$  converges in probability to  $\eta = E[X(n)]$ .



# Poisson Process

- ▶ This is the next process we study

# Poisson Process

- ▶ This is the next process we study
- ▶ This is a discrete-state continuous-time process

# Poisson Process

- ▶ This is the next process we study
- ▶ This is a discrete-state continuous-time process
- ▶ The index set is the interval  $[0, \infty)$  and all random variables are discrete and take non-negative integer values.



- ▶ A random process  $\{N(t), t \geq 0\}$  is called a counting process if

- ▶ A random process  $\{N(t), t \geq 0\}$  is called a counting process if
  1.  $N(t) \geq 0$  and is integer-valued

- A random process  $\{N(t), t \geq 0\}$  is called a counting process if
1.  $N(t) \geq 0$  and is integer-valued
  2. If  $s < t$  then,  $N(s) \leq N(t)$

- A random process  $\{N(t), t \geq 0\}$  is called a counting process if
1.  $N(t) \geq 0$  and is integer-valued
  2. If  $s < t$  then,  $N(s) \leq N(t)$

*Generally,  $N(t)$  represents number of 'events' till  $t$*

- ▶ A random process  $\{N(t), t \geq 0\}$  is called a counting process if
  1.  $N(t) \geq 0$  and is integer-valued
  2. If  $s < t$  then,  $N(s) \leq N(t)$

*Generally,  $N(t)$  represents number of 'events' till  $t$*

- ▶ The counting process has independent increments if for all  $t_1 < t_2 \leq t_3 < t_4$ ,  $N(t_2) - N(t_1)$  is independent of  $N(t_4) - N(t_3)$

- ▶ A random process  $\{N(t), t \geq 0\}$  is called a counting process if
  1.  $N(t) \geq 0$  and is integer-valued
  2. If  $s < t$  then,  $N(s) \leq N(t)$

*Generally,  $N(t)$  represents number of 'events' till  $t$*

- ▶ The counting process has independent increments if for all  $t_1 < t_2 \leq t_3 < t_4$ ,  $N(t_2) - N(t_1)$  is independent of  $N(t_4) - N(t_3)$
- ▶ In particular, for all  $s > t > 0$ ,  $N(s) - N(t)$  is independent of  $N(t) - N(0)$

- ▶ A random process  $\{N(t), t \geq 0\}$  is called a counting process if
  1.  $N(t) \geq 0$  and is integer-valued
  2. If  $s < t$  then,  $N(s) \leq N(t)$

*Generally,  $N(t)$  represents number of 'events' till  $t$*

- ▶ The counting process has independent increments if for all  $t_1 < t_2 \leq t_3 < t_4$ ,  $N(t_2) - N(t_1)$  is independent of  $N(t_4) - N(t_3)$
- ▶ In particular, for all  $s > t > 0$ ,  $N(s) - N(t)$  is independent of  $N(t) - N(0)$
- ▶ The process is said to have stationary increments if  $N(t_2) - N(t_1)$  has the same distribution as  $N(t_2 + \tau) - N(t_1 + \tau)$ ,  $\forall \tau, \forall t_2 > t_1$

- ▶ We start with two definitions of Poisson process



- ▶ We start with two definitions of Poisson process
- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if

- ▶ We start with two definitions of Poisson process
- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$

- ▶ We start with two definitions of Poisson process
- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$
  2. The process has stationary and independent increments

- ▶ We start with two definitions of Poisson process
- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}, n = 0, 1, \dots$

- ▶ We start with two definitions of Poisson process
- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}$ ,  $n = 0, 1, \dots$
- ▶  $N(t)$  is Poisson with parameter  $\lambda t$

- ▶ We start with two definitions of Poisson process
- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}$ ,  $n = 0, 1, \dots$
- ▶  $N(t)$  is Poisson with parameter  $\lambda t$
- ▶  $E[N(t)] = \lambda t$  and hence  $\lambda$  is called rate

- ▶ We start with two definitions of Poisson process
- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}$ ,  $n = 0, 1, \dots$
- ▶  $N(t)$  is Poisson with parameter  $\lambda t$
- ▶  $E[N(t)] = \lambda t$  and hence  $\lambda$  is called rate
- ▶ Since the process has stationary increments and  $N(0) = 0$ ,  $(N(t + s) - N(s))$  would be Poisson with parameter  $\lambda t$  for all  $s, t > 0$ .

- **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if



► **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if

1.  $N(0) = 0$

- **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
1.  $N(0) = 0$
  2. The process has stationary and independent increments

► **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if

1.  $N(0) = 0$
2. The process has stationary and independent increments
3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  
 $Pr[N(h) \geq 2] = o(h)$

- **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  
 $Pr[N(h) \geq 2] = o(h)$
- We say  $g(h)$  is  $o(h)$  if

$$\lim_{h \rightarrow 0} \frac{g(h)}{h} = 0$$

► **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if

1.  $N(0) = 0$
2. The process has stationary and independent increments
3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  
 $Pr[N(h) \geq 2] = o(h)$

► We say  $g(h)$  is  $o(h)$  if

$$\lim_{h \rightarrow 0} \frac{g(h)}{h} = 0$$

► This definition tells us when Poisson process may be a good model

- **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  
 $Pr[N(h) \geq 2] = o(h)$
- We say  $g(h)$  is  $o(h)$  if

$$\lim_{h \rightarrow 0} \frac{g(h)}{h} = 0$$

- This definition tells us when Poisson process may be a good model
- We will show that both definitions are equivalent

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ **Definition 1**  $\{N(t), t \geq 0\}$  is a counting process with
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}, n = 0, 1, \dots$
- ▶ **Definition 2**  $\{N(t), t \geq 0\}$  is a counting process with
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  
 $Pr[N(h) \geq 2] = o(h)$



- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ **Definition 1**  $\{N(t), t \geq 0\}$  is a counting process with
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}, n = 0, 1, \dots$
- ▶ **Definition 2**  $\{N(t), t \geq 0\}$  is a counting process with
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  
 $Pr[N(h) \geq 2] = o(h)$
- ▶ For this we need to calculate distribution of  $N(t)$

► We first show Definition 2  $\Rightarrow$  Definition 1

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = \Pr[N(t) = n]$

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = Pr[N(t) = n]$

$$P_0(t+h) = Pr[N(t+h) = 0]$$

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = Pr[N(t) = n]$

$$\begin{aligned} P_0(t+h) &= Pr[N(t+h) = 0] \\ &= Pr[N(t) = 0, N(t+h) - N(t) = 0] \end{aligned}$$

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = Pr[N(t) = n]$

$$\begin{aligned}P_0(t+h) &= Pr[N(t+h) = 0] \\&= Pr[N(t) = 0, N(t+h) - N(t) = 0] \\&= Pr[N(t) = 0] Pr[N(t+h) - N(t) = 0] \\&\quad \text{(because of independent increments)}\end{aligned}$$

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = Pr[N(t) = n]$

$$\begin{aligned}P_0(t+h) &= Pr[N(t+h) = 0] \\&= Pr[N(t) = 0, N(t+h) - N(t) = 0] \\&= Pr[N(t) = 0] Pr[N(t+h) - N(t) = 0] \\&\quad \text{(because of independent increments)} \\&= Pr[N(t) = 0] Pr[N(h) = 0] \quad \text{(stationary increments)}\end{aligned}$$



- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = Pr[N(t) = n]$

$$\begin{aligned}P_0(t+h) &= Pr[N(t+h) = 0] \\&= Pr[N(t) = 0, N(t+h) - N(t) = 0] \\&= Pr[N(t) = 0] Pr[N(t+h) - N(t) = 0] \\&\quad \text{(because of independent increments)} \\&= Pr[N(t) = 0] Pr[N(h) = 0] \quad \text{(stationary increments)} \\&= P_0(t)(1 - \lambda h + o(h))\end{aligned}$$

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = Pr[N(t) = n]$

$$\begin{aligned}
 P_0(t+h) &= Pr[N(t+h) = 0] \\
 &= Pr[N(t) = 0, N(t+h) - N(t) = 0] \\
 &= Pr[N(t) = 0] Pr[N(t+h) - N(t) = 0] \\
 &\quad \text{(because of independent increments)} \\
 &= Pr[N(t) = 0] Pr[N(h) = 0] \quad \text{(stationary increments)} \\
 &= P_0(t)(1 - \lambda h + o(h))
 \end{aligned}$$

$$\Rightarrow \frac{P_0(t+h) - P_0(t)}{h} = -\lambda P_0(t) + \frac{o(h)}{h}$$

- ▶ We first show Definition 2  $\Rightarrow$  Definition 1
- ▶ For this we need to calculate distribution of  $N(t)$
- ▶ Let  $P_n(t) = Pr[N(t) = n]$

$$\begin{aligned}
 P_0(t+h) &= Pr[N(t+h) = 0] \\
 &= Pr[N(t) = 0, N(t+h) - N(t) = 0] \\
 &= Pr[N(t) = 0] Pr[N(t+h) - N(t) = 0] \\
 &\quad \text{(because of independent increments)} \\
 &= Pr[N(t) = 0] Pr[N(h) = 0] \quad \text{(stationary increments)} \\
 &= P_0(t)(1 - \lambda h + o(h))
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow \frac{P_0(t+h) - P_0(t)}{h} &= -\lambda P_0(t) + \frac{o(h)}{h} \\
 \Rightarrow \frac{d}{dt} P_0(t) &= -\lambda P_0(t)
 \end{aligned}$$

- Now we can solve this differential equation to get  $P_0(t)$

- Now we can solve this differential equation to get  $P_0(t)$

$$\frac{d}{dt}P_0(t) = -\lambda P_0(t)$$

- Now we can solve this differential equation to get  $P_0(t)$

$$\begin{aligned}\frac{d}{dt}P_0(t) &= -\lambda P_0(t) \\ \Rightarrow \frac{1}{P_0(t)} \frac{d}{dt}P_0(t) &= -\lambda\end{aligned}$$

- Now we can solve this differential equation to get  $P_0(t)$

$$\begin{aligned}\frac{d}{dt}P_0(t) &= -\lambda P_0(t) \\ \Rightarrow \frac{1}{P_0(t)} \frac{d}{dt}P_0(t) &= -\lambda \\ \Rightarrow \ln(P_0(t)) &= -\lambda t + c\end{aligned}$$

- Now we can solve this differential equation to get  $P_0(t)$

$$\begin{aligned}\frac{d}{dt}P_0(t) &= -\lambda P_0(t) \\ \Rightarrow \frac{1}{P_0(t)} \frac{d}{dt}P_0(t) &= -\lambda \\ \Rightarrow \ln(P_0(t)) &= -\lambda t + c \\ \Rightarrow P_0(t) &= Ke^{-\lambda t}\end{aligned}$$



- Now we can solve this differential equation to get  $P_0(t)$

$$\begin{aligned}\frac{d}{dt}P_0(t) &= -\lambda P_0(t) \\ \Rightarrow \frac{1}{P_0(t)} \frac{d}{dt}P_0(t) &= -\lambda \\ \Rightarrow \ln(P_0(t)) &= -\lambda t + c \\ \Rightarrow P_0(t) &= Ke^{-\lambda t}\end{aligned}$$

- Since  $P_0(0) = Pr[N(0) = 0] = 1$ , we get  $K = 1$  and hence

$$P_0(t) = Pr[N(t) = 0] = e^{-\lambda t}$$

- Now we can solve this differential equation to get  $P_0(t)$

$$\begin{aligned}\frac{d}{dt}P_0(t) &= -\lambda P_0(t) \\ \Rightarrow \frac{1}{P_0(t)} \frac{d}{dt}P_0(t) &= -\lambda \\ \Rightarrow \ln(P_0(t)) &= -\lambda t + c \\ \Rightarrow P_0(t) &= Ke^{-\lambda t}\end{aligned}$$

- Since  $P_0(0) = Pr[N(0) = 0] = 1$ , we get  $K = 1$  and hence

$$P_0(t) = Pr[N(t) = 0] = e^{-\lambda t}$$

- Next we consider  $P_n(t)$  for  $n > 0$

$$P_n(t+h) = Pr[N(t+h) = n]$$

$$\begin{aligned}
P_n(t+h) &= \Pr[N(t+h) = n] \\
&= \Pr[N(t) = n, N(t+h) - N(t) = 0] +
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] +
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n Pr[N(t) = n-k, N(t+h) - N(t) = k]
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n Pr[N(t) = n-k, N(t+h) - N(t) = k]
\end{aligned}$$

We have

$$\begin{aligned}
&Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(t+h) - N(t) = k]
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n Pr[N(t) = n-k, N(t+h) - N(t) = k]
\end{aligned}$$

We have

$$\begin{aligned}
&Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(h) = k]
\end{aligned}$$



$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n Pr[N(t) = n-k, N(t+h) - N(t) = k]
\end{aligned}$$

We have

$$\begin{aligned}
&Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(h) = k] = o(h), \forall k \geq 2
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n Pr[N(t) = n-k, N(t+h) - N(t) = k]
\end{aligned}$$

We have

$$\begin{aligned}
&Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(h) = k] = o(h), \forall k \geq 2
\end{aligned}$$

$$Pr[N(t) = n, N(t+h) - N(t) = 0] = P_n(t)P_0(h)$$

$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n Pr[N(t) = n-k, N(t+h) - N(t) = k]
\end{aligned}$$

We have

$$\begin{aligned}
&Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(t+h) - N(t) = k] \\
&= Pr[N(t) = n-k] P[N(h) = k] = o(h), \forall k \geq 2
\end{aligned}$$

$$Pr[N(t) = n, N(t+h) - N(t) = 0] = P_n(t)P_0(h)$$

$$Pr[N(t) = n-1, N(t+h) - N(t) = 1] = P_{n-1}(t)P_1(h)$$

$$\begin{aligned}
P_n(t+h) &= \Pr[N(t+h) = n] \\
&= \Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad \Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n \Pr[N(t) = n-k, N(t+h) - N(t) = k]
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= \Pr[N(t+h) = n] \\
&= \Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad \Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n \Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= P_n(t)P_0(h) + P_{n-1}(t)P_1(h) + o(h)
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= \Pr[N(t+h) = n] \\
&= \Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad \Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n \Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= P_n(t)P_0(h) + P_{n-1}(t)P_1(h) + o(h) \\
&= P_n(t)(1 - \lambda h + o(h))
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= \Pr[N(t+h) = n] \\
&= \Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad \Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n \Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= P_n(t)P_0(h) + P_{n-1}(t)P_1(h) + o(h) \\
&= P_n(t)(1 - \lambda h + o(h)) + P_{n-1}(t)(\lambda h + o(h)) + o(h)
\end{aligned}$$

$$\begin{aligned}
P_n(t+h) &= \Pr[N(t+h) = n] \\
&= \Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad \Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n \Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= P_n(t)P_0(h) + P_{n-1}(t)P_1(h) + o(h) \\
&= P_n(t)(1 - \lambda h + o(h)) + P_{n-1}(t)(\lambda h + o(h)) + o(h) \\
&= P_n(t) - \lambda h P_n(t) + \lambda h P_{n-1}(t) + o(h)
\end{aligned}$$



$$\begin{aligned}
P_n(t+h) &= \Pr[N(t+h) = n] \\
&= \Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad \Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n \Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= P_n(t)P_0(h) + P_{n-1}(t)P_1(h) + o(h) \\
&= P_n(t)(1 - \lambda h + o(h)) + P_{n-1}(t)(\lambda h + o(h)) + o(h) \\
&= P_n(t) - \lambda h P_n(t) + \lambda h P_{n-1}(t) + o(h)
\end{aligned}$$

$$\Rightarrow \frac{P_n(t+h) - P_n(t)}{h} = -\lambda P_n(t) + \lambda P_{n-1}(t) + \frac{o(h)}{h}$$

$$\begin{aligned}
P_n(t+h) &= Pr[N(t+h) = n] \\
&= Pr[N(t) = n, N(t+h) - N(t) = 0] + \\
&\quad Pr[N(t) = n-1, N(t+h) - N(t) = 1] + \\
&\quad \sum_{k=2}^n Pr[N(t) = n-k, N(t+h) - N(t) = k] \\
&= P_n(t)P_0(h) + P_{n-1}(t)P_1(h) + o(h) \\
&= P_n(t)(1 - \lambda h + o(h)) + P_{n-1}(t)(\lambda h + o(h)) + o(h) \\
&= P_n(t) - \lambda h P_n(t) + \lambda h P_{n-1}(t) + o(h)
\end{aligned}$$

$$\Rightarrow \frac{P_n(t+h) - P_n(t)}{h} = -\lambda P_n(t) + \lambda P_{n-1}(t) + \frac{o(h)}{h}$$

$$\Rightarrow \frac{d}{dt}P_n(t) = -\lambda P_n(t) + \lambda P_{n-1}(t)$$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- We need to solve this linear ODE to obtain  $P_n$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ .

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$e^{\lambda t} (P'_n(t) + \lambda P_n(t)) = e^{\lambda t} \lambda P_{n-1}(t)$$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ .



$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ . Take  $n = 1$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ . Take  $n = 1$

$$\frac{d}{dt} (P_1(t) e^{\lambda t}) = \lambda e^{\lambda t} P_0(t)$$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ . Take  $n = 1$

$$\frac{d}{dt} (P_1(t) e^{\lambda t}) = \lambda e^{\lambda t} P_0(t) = \lambda e^{\lambda t} e^{-\lambda t} = \lambda$$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ . Take  $n = 1$

$$\begin{aligned} \frac{d}{dt} (P_1(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_0(t) = \lambda e^{\lambda t} e^{-\lambda t} = \lambda \\ \Rightarrow e^{\lambda t} P_1(t) &= \lambda t + c \end{aligned}$$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ . Take  $n = 1$

$$\begin{aligned} \frac{d}{dt} (P_1(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_0(t) = \lambda e^{\lambda t} e^{-\lambda t} = \lambda \\ \Rightarrow e^{\lambda t} P_1(t) &= \lambda t + c \Rightarrow P_1(t) = e^{-\lambda t} (\lambda t + c) \end{aligned}$$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ . Take  $n = 1$

$$\begin{aligned} \frac{d}{dt} (P_1(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_0(t) = \lambda e^{\lambda t} e^{-\lambda t} = \lambda \\ \Rightarrow e^{\lambda t} P_1(t) &= \lambda t + c \Rightarrow P_1(t) = e^{-\lambda t} (\lambda t + c) \end{aligned}$$

- ▶  $P_1(0) = Pr[N(0) = 1] = 0 \Rightarrow c = 0$

$$\frac{d}{dt}P_n(t) + \lambda P_n(t) = \lambda P_{n-1}(t)$$

- ▶ We need to solve this linear ODE to obtain  $P_n$
- ▶ The integrating factor is  $e^{\lambda t}$ . Let  $P'_n(t) = \frac{d}{dt}P_n(t)$

$$\begin{aligned} e^{\lambda t} (P'_n(t) + \lambda P_n(t)) &= e^{\lambda t} \lambda P_{n-1}(t) \\ \Rightarrow \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) \end{aligned}$$

- ▶ We need  $P_{n-1}$  to solve for  $P_n$ . Take  $n = 1$

$$\begin{aligned} \frac{d}{dt} (P_1(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_0(t) = \lambda e^{\lambda t} e^{-\lambda t} = \lambda \\ \Rightarrow e^{\lambda t} P_1(t) &= \lambda t + c \Rightarrow P_1(t) = e^{-\lambda t} (\lambda t + c) \end{aligned}$$

- ▶  $P_1(0) = Pr[N(0) = 1] = 0 \Rightarrow c = 0$   
Hence  $P_1(t) = \lambda t e^{-\lambda t}$

► We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$



- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$

- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

$$\frac{d}{dt} (P_n(t) e^{\lambda t}) = \lambda e^{\lambda t} P_{n-1}(t)$$

- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

$$\frac{d}{dt} (P_n(t) e^{\lambda t}) = \lambda e^{\lambda t} P_{n-1}(t) = \lambda e^{\lambda t} e^{-\lambda t} \frac{(\lambda t)^{n-1}}{(n-1)!}$$

- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

$$\frac{d}{dt} (P_n(t) e^{\lambda t}) = \lambda e^{\lambda t} P_{n-1}(t) = \lambda e^{\lambda t} e^{-\lambda t} \frac{(\lambda t)^{n-1}}{(n-1)!} = \lambda^n \frac{t^{n-1}}{(n-1)!}$$

- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

$$\frac{d}{dt} (P_n(t) e^{\lambda t}) = \lambda e^{\lambda t} P_{n-1}(t) = \lambda e^{\lambda t} e^{-\lambda t} \frac{(\lambda t)^{n-1}}{(n-1)!} = \lambda^n \frac{t^{n-1}}{(n-1)!}$$

$$\Rightarrow e^{\lambda t} P_n(t) = \lambda^n \frac{t^n}{n (n-1)!} + c$$

- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

$$\begin{aligned} \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) = \lambda e^{\lambda t} e^{-\lambda t} \frac{(\lambda t)^{n-1}}{(n-1)!} = \lambda^n \frac{t^{n-1}}{(n-1)!} \\ \Rightarrow e^{\lambda t} P_n(t) &= \lambda^n \frac{t^n}{n (n-1)!} + c \Rightarrow P_n(t) = e^{-\lambda t} \frac{(\lambda t)^n}{n!} \end{aligned}$$

- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

$$\begin{aligned} \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) = \lambda e^{\lambda t} e^{-\lambda t} \frac{(\lambda t)^{n-1}}{(n-1)!} = \lambda^n \frac{t^{n-1}}{(n-1)!} \\ \Rightarrow e^{\lambda t} P_n(t) &= \lambda^n \frac{t^n}{n (n-1)!} + c \Rightarrow P_n(t) = e^{-\lambda t} \frac{(\lambda t)^n}{n!} \end{aligned}$$

where  $c = 0$  because  $P_n(0) = 0$ .



- ▶ We showed:  $P_0(t) = e^{-\lambda t}$  and  $P_1(t) = \lambda t e^{-\lambda t}$
- ▶ We need to show:  $P_k(t) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$
- ▶ Assume it is true till  $k = n - 1$

$$\begin{aligned} \frac{d}{dt} (P_n(t) e^{\lambda t}) &= \lambda e^{\lambda t} P_{n-1}(t) = \lambda e^{\lambda t} e^{-\lambda t} \frac{(\lambda t)^{n-1}}{(n-1)!} = \lambda^n \frac{t^{n-1}}{(n-1)!} \\ \Rightarrow e^{\lambda t} P_n(t) &= \lambda^n \frac{t^n}{n (n-1)!} + c \Rightarrow P_n(t) = e^{-\lambda t} \frac{(\lambda t)^n}{n!} \end{aligned}$$

where  $c = 0$  because  $P_n(0) = 0$ .

- ▶ This completes the proof that Definition 2 implies Definition 1

► **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if

1.  $N(0) = 0$
2. The process has stationary and independent increments
3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}, n = 0, 1, \dots$

► **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if

1.  $N(0) = 0$
2. The process has stationary and independent increments
3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  
 $Pr[N(h) \geq 2] = o(h)$

► Now we prove Definition 1 implies Definition 2

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$Pr[N(h) = 1] = \lambda h e^{-\lambda h}$$

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$Pr[N(h) = 1] = \lambda h e^{-\lambda h} = \lambda h + \lambda h (e^{-\lambda h} - 1)$$

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$Pr[N(h) = 1] = \lambda h e^{-\lambda h} = \lambda h + \lambda h (e^{-\lambda h} - 1) = \lambda h + o(h)$$



- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$Pr[N(h) = 1] = \lambda h e^{-\lambda h} = \lambda h + \lambda h (e^{-\lambda h} - 1) = \lambda h + o(h)$$

because

$$\lim_{h \rightarrow 0} \frac{\lambda h (e^{-\lambda h} - 1)}{h}$$

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$Pr[N(h) = 1] = \lambda h e^{-\lambda h} = \lambda h + \lambda h (e^{-\lambda h} - 1) = \lambda h + o(h)$$

because

$$\lim_{h \rightarrow 0} \frac{\lambda h (e^{-\lambda h} - 1)}{h} = \lim_{h \rightarrow 0} \lambda (e^{-\lambda h} - 1)$$

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$Pr[N(h) = 1] = \lambda h e^{-\lambda h} = \lambda h + \lambda h (e^{-\lambda h} - 1) = \lambda h + o(h)$$

because

$$\lim_{h \rightarrow 0} \frac{\lambda h (e^{-\lambda h} - 1)}{h} = \lim_{h \rightarrow 0} \lambda (e^{-\lambda h} - 1) = 0$$

- ▶ Now we prove Definition 1 implies Definition 2
- ▶ We need to only show point(3) of Definition 2 using point (3) of Definition 1

$$\text{Let } Pr[N(t) = k] = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

$$Pr[N(h) = 1] = \lambda h e^{-\lambda h} = \lambda h + \lambda h (e^{-\lambda h} - 1) = \lambda h + o(h)$$

because

$$\lim_{h \rightarrow 0} \frac{\lambda h (e^{-\lambda h} - 1)}{h} = \lim_{h \rightarrow 0} \lambda (e^{-\lambda h} - 1) = 0$$

- ▶ We showed  $Pr[N(h) = 1] = \lambda h + o(h)$

- Now we need to show  $Pr[N(h) \geq 2] = o(h)$

- Now we need to show  $Pr[N(h) \geq 2] = o(h)$

$$Pr[N(h) \geq 2] = 1 - Pr[N(h) = 0] - Pr[N(h) = 1]$$

► Now we need to show  $Pr[N(h) \geq 2] = o(h)$

$$\begin{aligned} Pr[N(h) \geq 2] &= 1 - Pr[N(h) = 0] - Pr[N(h) = 1] \\ &= 1 - e^{-\lambda h} - \lambda h e^{-\lambda h} \end{aligned}$$

- Now we need to show  $Pr[N(h) \geq 2] = o(h)$

$$\begin{aligned} Pr[N(h) \geq 2] &= 1 - Pr[N(h) = 0] - Pr[N(h) = 1] \\ &= 1 - e^{-\lambda h} - \lambda h e^{-\lambda h} \end{aligned}$$

- This goes to zero as  $h \rightarrow 0$



- ▶ Now we need to show  $Pr[N(h) \geq 2] = o(h)$

$$\begin{aligned} Pr[N(h) \geq 2] &= 1 - Pr[N(h) = 0] - Pr[N(h) = 1] \\ &= 1 - e^{-\lambda h} - \lambda h e^{-\lambda h} \end{aligned}$$

- ▶ This goes to zero as  $h \rightarrow 0$
- ▶ We can use L'Hospital rule

- Now we need to show  $Pr[N(h) \geq 2] = o(h)$

$$\begin{aligned} Pr[N(h) \geq 2] &= 1 - Pr[N(h) = 0] - Pr[N(h) = 1] \\ &= 1 - e^{-\lambda h} - \lambda h e^{-\lambda h} \end{aligned}$$

- This goes to zero as  $h \rightarrow 0$
- We can use L'Hospital rule

$$\lim_{h \rightarrow 0} \frac{1 - e^{-\lambda h} - \lambda h e^{-\lambda h}}{h} = \lim_{h \rightarrow 0} \frac{\lambda e^{-\lambda h} - \lambda e^{-\lambda h} + \lambda^2 h e^{-\lambda h}}{1} = 0$$

- Now we need to show  $Pr[N(h) \geq 2] = o(h)$

$$\begin{aligned} Pr[N(h) \geq 2] &= 1 - Pr[N(h) = 0] - Pr[N(h) = 1] \\ &= 1 - e^{-\lambda h} - \lambda h e^{-\lambda h} \end{aligned}$$

- This goes to zero as  $h \rightarrow 0$
- We can use L'Hospital rule

$$\lim_{h \rightarrow 0} \frac{1 - e^{-\lambda h} - \lambda h e^{-\lambda h}}{h} = \lim_{h \rightarrow 0} \frac{\lambda e^{-\lambda h} - \lambda e^{-\lambda h} + \lambda^2 h e^{-\lambda h}}{1} = 0$$

- This completes the proof that Definition 2 implies Definition 1

# These two definitions are equivalent

- ▶ **Definition 1** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(t) = n] = e^{-\lambda t} \frac{(\lambda t)^n}{n!}$ ,  $n = 0, 1, \dots$
- ▶ **Definition 2** A counting process  $\{N(t), t \geq 0\}$  is said to be a Poisson process with rate  $\lambda > 0$  if
  1.  $N(0) = 0$
  2. The process has stationary and independent increments
  3.  $Pr[N(h) = 1] = \lambda h + o(h)$  and  $Pr[N(h) \geq 2] = o(h)$

- ▶ Since the process has stationary increments, for  $t_2 > t_1$ ,

- ▶ Since the process has stationary increments, for  $t_2 > t_1$ ,

$$Pr[N(t_2) - N(t_1) = k] = Pr[N(t_2 - t_1) - N(0) = k]$$

- Since the process has stationary increments, for  $t_2 > t_1$ ,

$$\begin{aligned} Pr[N(t_2) - N(t_1) = k] &= Pr[N(t_2 - t_1) - N(0) = k] \\ &= e^{-\lambda(t_2 - t_1)} \frac{(\lambda(t_2 - t_1))^k}{k!} \end{aligned}$$

- ▶ Since the process has stationary increments, for  $t_2 > t_1$ ,

$$\begin{aligned} Pr[N(t_2) - N(t_1) = k] &= Pr[N(t_2 - t_1) - N(0) = k] \\ &= e^{-\lambda(t_2 - t_1)} \frac{(\lambda(t_2 - t_1))^k}{k!} \end{aligned}$$

- ▶ The first order distribution of the process is:  
 $N(t) \sim \text{Poisson}(\lambda t)$



- ▶ Since the process has stationary increments, for  $t_2 > t_1$ ,

$$\begin{aligned} Pr[N(t_2) - N(t_1) = k] &= Pr[N(t_2 - t_1) - N(0) = k] \\ &= e^{-\lambda(t_2 - t_1)} \frac{(\lambda(t_2 - t_1))^k}{k!} \end{aligned}$$

- ▶ The first order distribution of the process is:  
 $N(t) \sim \text{Poisson}(\lambda t)$
- ▶ This, along with stationary and independent increments property determines all distributions

- ▶ Since the process has stationary increments, for  $t_2 > t_1$ ,

$$\begin{aligned} Pr[N(t_2) - N(t_1) = k] &= Pr[N(t_2 - t_1) - N(0) = k] \\ &= e^{-\lambda(t_2 - t_1)} \frac{(\lambda(t_2 - t_1))^k}{k!} \end{aligned}$$

- ▶ The first order distribution of the process is:  
 $N(t) \sim \text{Poisson}(\lambda t)$
- ▶ This, along with stationary and independent increments property determines all distributions

$$\begin{aligned} Pr[N(t_1) = n_1, N(t_2) = n_2, N(t_3) = n_3] \\ = Pr[N(t_1) = n_1] Pr[N(t_2) - N(t_1) = n_2 - n_1] \\ Pr[N(t_3) - N(t_2) = n_3 - n_2] \end{aligned}$$

- ▶ Since the process has stationary increments, for  $t_2 > t_1$ ,

$$\begin{aligned} Pr[N(t_2) - N(t_1) = k] &= Pr[N(t_2 - t_1) - N(0) = k] \\ &= e^{-\lambda(t_2 - t_1)} \frac{(\lambda(t_2 - t_1))^k}{k!} \end{aligned}$$

- ▶ The first order distribution of the process is:  
 $N(t) \sim \text{Poisson}(\lambda t)$
- ▶ This, along with stationary and independent increments property determines all distributions

$$\begin{aligned} &Pr[N(t_1) = n_1, N(t_2) = n_2, N(t_3) = n_3] \\ &= Pr[N(t_1) = n_1] Pr[N(t_2) - N(t_1) = n_2 - n_1] \\ &\quad Pr[N(t_3) - N(t_2) = n_3 - n_2] \\ &= Pr[N(t_1) = n_1] Pr[N(t_2 - t_1) = n_2 - n_1] Pr[N(t_3 - t_2) = n_3 - n_2] \end{aligned}$$

- ▶ Since the process has stationary increments, for  $t_2 > t_1$ ,

$$\begin{aligned} Pr[N(t_2) - N(t_1) = k] &= Pr[N(t_2 - t_1) - N(0) = k] \\ &= e^{-\lambda(t_2 - t_1)} \frac{(\lambda(t_2 - t_1))^k}{k!} \end{aligned}$$

- ▶ The first order distribution of the process is:  
 $N(t) \sim \text{Poisson}(\lambda t)$
- ▶ This, along with stationary and independent increments property determines all distributions

$$\begin{aligned} &Pr[N(t_1) = n_1, N(t_2) = n_2, N(t_3) = n_3] \\ &= Pr[N(t_1) = n_1] Pr[N(t_2) - N(t_1) = n_2 - n_1] \\ &\quad Pr[N(t_3) - N(t_2) = n_3 - n_2] \\ &= Pr[N(t_1) = n_1] Pr[N(t_2 - t_1) = n_2 - n_1] Pr[N(t_3 - t_2) = n_3 - n_2] \end{aligned}$$

where we assumed  $t_1 < t_2 < t_3$

- ▶ We can easily calculate mean and autocorrelation of the process

- ▶ We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t$$

- ▶ We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

- ▶ We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$R_N(t_1, t_2) = E[N(t_2)N(t_1)]$$



- We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$\begin{aligned} R_N(t_1, t_2) &= E[N(t_2)N(t_1)] \\ &= E[N(t_1)(N(t_2) - N(t_1) + N(t_1))] \end{aligned}$$

- We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$\begin{aligned} R_N(t_1, t_2) &= E[N(t_2)N(t_1)] \\ &= E[N(t_1)(N(t_2) - N(t_1) + N(t_1))] \\ &= E[N(t_1)(N(t_2) - N(t_1))] + E[N(t_1)^2] \end{aligned}$$

- We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$\begin{aligned} R_N(t_1, t_2) &= E[N(t_2)N(t_1)] \\ &= E[N(t_1)(N(t_2) - N(t_1) + N(t_1))] \\ &= E[N(t_1)(N(t_2) - N(t_1))] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2) - N(t_1)] + E[N(t_1)^2] \end{aligned}$$

- We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$\begin{aligned} R_N(t_1, t_2) &= E[N(t_2)N(t_1)] \\ &= E[N(t_1)(N(t_2) - N(t_1) + N(t_1))] \\ &= E[N(t_1)(N(t_2) - N(t_1))] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2) - N(t_1)] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2 - t_1)] + E[N(t_1)^2] \end{aligned}$$

- We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$\begin{aligned} R_N(t_1, t_2) &= E[N(t_2)N(t_1)] \\ &= E[N(t_1)(N(t_2) - N(t_1) + N(t_1))] \\ &= E[N(t_1)(N(t_2) - N(t_1))] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2) - N(t_1)] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2 - t_1)] + E[N(t_1)^2] \\ &= \lambda t_1 (\lambda(t_2 - t_1)) + (\lambda t_1 + \lambda^2 t_1^2) \end{aligned}$$

- We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$\begin{aligned} R_N(t_1, t_2) &= E[N(t_2)N(t_1)] \\ &= E[N(t_1)(N(t_2) - N(t_1) + N(t_1))] \\ &= E[N(t_1)(N(t_2) - N(t_1))] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2) - N(t_1)] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2 - t_1)] + E[N(t_1)^2] \\ &= \lambda t_1 (\lambda(t_2 - t_1)) + (\lambda t_1 + \lambda^2 t_1^2) \\ &= \lambda t_1 + \lambda^2 t_1 t_2 \end{aligned}$$

- ▶ We can easily calculate mean and autocorrelation of the process

$$\eta_N(t) = E[N(t)] = \lambda t \quad \Rightarrow \quad \text{not stationary}$$

With  $t_2 > t_1$ , we have

$$\begin{aligned} R_N(t_1, t_2) &= E[N(t_2)N(t_1)] \\ &= E[N(t_1)(N(t_2) - N(t_1) + N(t_1))] \\ &= E[N(t_1)(N(t_2) - N(t_1))] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2) - N(t_1)] + E[N(t_1)^2] \\ &= E[N(t_1)] E[N(t_2 - t_1)] + E[N(t_1)^2] \\ &= \lambda t_1 (\lambda(t_2 - t_1)) + (\lambda t_1 + \lambda^2 t_1^2) \\ &= \lambda t_1 + \lambda^2 t_1 t_2 \end{aligned}$$

$$\Rightarrow R_N(t_1, t_2) = \lambda^2 t_1 t_2 + \lambda \min(t_1, t_2)$$

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n - 1)^{st}$  events.



# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n - 1)^{st}$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  – time of  $n^{th}$  event

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n - 1)st$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  – time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0]$$

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n-1)^{st}$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  – time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0] = e^{-\lambda t}$$

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n-1)^{st}$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  – time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0] = e^{-\lambda t}$$

$$\Rightarrow T_1 \sim \text{exponential}(\lambda)$$

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n-1)^{st}$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  - time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0] = e^{-\lambda t}$$

$$\Rightarrow T_1 \sim \text{exponential}(\lambda)$$

$$Pr[T_2 > t | T_1 = s] = Pr[0 \text{ events in } (s, s+t] \mid T_1 = s]$$

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n-1)^{st}$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  - time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0] = e^{-\lambda t}$$

$$\Rightarrow T_1 \sim \text{exponential}(\lambda)$$

$$\begin{aligned} Pr[T_2 > t | T_1 = s] &= Pr[0 \text{ events in } (s, s+t] \mid T_1 = s] \\ &= Pr[0 \text{ events in } (s, s+t)] \end{aligned}$$

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n-1)st$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  - time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0] = e^{-\lambda t}$$

$$\Rightarrow T_1 \sim \text{exponential}(\lambda)$$

$$\begin{aligned} Pr[T_2 > t | T_1 = s] &= Pr[0 \text{ events in } (s, s+t] \mid T_1 = s] \\ &= Pr[0 \text{ events in } (s, s+t)] = e^{-\lambda t} \end{aligned}$$

# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n-1)^{st}$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  – time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0] = e^{-\lambda t}$$

$$\Rightarrow T_1 \sim \text{exponential}(\lambda)$$

$$\begin{aligned} Pr[T_2 > t | T_1 = s] &= Pr[0 \text{ events in } (s, s+t] \mid T_1 = s] \\ &= Pr[0 \text{ events in } (s, s+t)] = e^{-\lambda t} \end{aligned}$$

$$\Rightarrow Pr[T_2 > t] = \int Pr[T_2 > t | T_1 = s] f_{T_1}(s) ds = e^{-\lambda t}$$



# Inter-arrival or waiting times

- ▶ Let  $T_1$  denote the time of first event and let  $T_n$  denote the time between  $n^{th}$  and  $(n-1)^{st}$  events.
- ▶ Let  $S_n = \sum_{i=1}^n T_i$  - time of  $n^{th}$  event

$$Pr[T_1 > t] = Pr[N(t) = 0] = e^{-\lambda t}$$

$$\Rightarrow T_1 \sim \text{exponential}(\lambda)$$

$$\begin{aligned} Pr[T_2 > t | T_1 = s] &= Pr[0 \text{ events in } (s, s+t] \mid T_1 = s] \\ &= Pr[0 \text{ events in } (s, s+t)] = e^{-\lambda t} \end{aligned}$$

$$\Rightarrow Pr[T_2 > t] = \int Pr[T_2 > t | T_1 = s] f_{T_1}(s) ds = e^{-\lambda t}$$

- ▶  $T_n$  are iid exponential with parameter  $\lambda$

- The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

- The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

Since  $T_i$  are iid, exponential,  $S_n$  is Gamma with parameters  $n, \lambda$

- ▶ The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

Since  $T_i$  are iid, exponential,  $S_n$  is Gamma with parameters  $n, \lambda$

- ▶ Let  $0 < s < t$ .

$$Pr[S_1 \leq s | N(t) = 1]$$

- The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

Since  $T_i$  are iid, exponential,  $S_n$  is Gamma with parameters  $n, \lambda$

- Let  $0 < s < t$ .

$$Pr[S_1 \leq s | N(t) = 1] = \frac{Pr[S_1 \leq s, N(t) = 1]}{Pr[N(t) = 1]}$$

- The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

Since  $T_i$  are iid, exponential,  $S_n$  is Gamma with parameters  $n, \lambda$

- Let  $0 < s < t$ .

$$\begin{aligned} Pr[S_1 \leq s | N(t) = 1] &= \frac{Pr[S_1 \leq s, N(t) = 1]}{Pr[N(t) = 1]} \\ &= \frac{Pr[1 \text{ event in } [0, s], 0 \text{ in } (s, t)]}{Pr[N(t) = 1]} \end{aligned}$$

- The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

Since  $T_i$  are iid, exponential,  $S_n$  is Gamma with parameters  $n, \lambda$

- Let  $0 < s < t$ .

$$\begin{aligned} Pr[S_1 \leq s | N(t) = 1] &= \frac{Pr[S_1 \leq s, N(t) = 1]}{Pr[N(t) = 1]} \\ &= \frac{Pr[1 \text{ event in } [0, s], 0 \text{ in } (s, t)]}{Pr[N(t) = 1]} \\ &= \frac{\lambda s e^{-\lambda s} e^{-\lambda(t-s)}}{\lambda t e^{-\lambda t}} \end{aligned}$$

- The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

Since  $T_i$  are iid, exponential,  $S_n$  is Gamma with parameters  $n, \lambda$

- Let  $0 < s < t$ .

$$\begin{aligned} Pr[S_1 \leq s | N(t) = 1] &= \frac{Pr[S_1 \leq s, N(t) = 1]}{Pr[N(t) = 1]} \\ &= \frac{Pr[1 \text{ event in } [0, s], 0 \text{ in } (s, t)]}{Pr[N(t) = 1]} \\ &= \frac{\lambda s e^{-\lambda s} e^{-\lambda(t-s)}}{\lambda t e^{-\lambda t}} \\ &= \frac{s}{t} \end{aligned}$$



- ▶ The time of  $n^{th}$  event is

$$S_n = \sum_{i=1}^n T_i$$

Since  $T_i$  are iid, exponential,  $S_n$  is Gamma with parameters  $n, \lambda$

- ▶ Let  $0 < s < t$ .

$$\begin{aligned} Pr[S_1 \leq s | N(t) = 1] &= \frac{Pr[S_1 \leq s, N(t) = 1]}{Pr[N(t) = 1]} \\ &= \frac{Pr[1 \text{ event in } [0, s], 0 \text{ in } (s, t)]}{Pr[N(t) = 1]} \\ &= \frac{\lambda s e^{-\lambda s} e^{-\lambda(t-s)}}{\lambda t e^{-\lambda t}} \\ &= \frac{s}{t} \end{aligned}$$

- ▶ Conditioned on  $N(t) = 1$ ,  $S_1$  is uniform over  $[0, t]$

- ▶ This can be used, e.g., in simulating Poisson process

- ▶ This can be used, e.g., in simulating Poisson process
- ▶ We can cut time axis into small intervals of length  $h$ .

- ▶ This can be used, e.g., in simulating Poisson process
- ▶ We can cut time axis into small intervals of length  $h$ .
- ▶ In each interval we can decide whether or not there is an event, with prob  $\lambda h$ .

- ▶ This can be used, e.g., in simulating Poisson process
- ▶ We can cut time axis into small intervals of length  $h$ .
- ▶ In each interval we can decide whether or not there is an event, with prob  $\lambda h$ .
- ▶ If there is an event, we choose its time uniformly in the interval.

- ▶ This can be used, e.g., in simulating Poisson process
- ▶ We can cut time axis into small intervals of length  $h$ .
- ▶ In each interval we can decide whether or not there is an event, with prob  $\lambda h$ .
- ▶ If there is an event, we choose its time uniformly in the interval.
- ▶ Called Bernoulli approximation of Poisson process

- ▶ This can be used, e.g., in simulating Poisson process
- ▶ We can cut time axis into small intervals of length  $h$ .
- ▶ In each interval we can decide whether or not there is an event, with prob  $\lambda h$ .
- ▶ If there is an event, we choose its time uniformly in the interval.
- ▶ Called Bernoulli approximation of Poisson process
- ▶ We could also generate Poisson process by generating independent exponential random variables

- ▶ Let  $S_1, \dots, S_n$  be the times of the first  $n$  events.



- ▶ Let  $S_1, \dots, S_n$  be the times of the first  $n$  events.
- ▶ We calculated conditional density of  $S_1$  conditioned on  $N(t) = 1$ .

- ▶ Let  $S_1, \dots, S_n$  be the times of the first  $n$  events.
- ▶ We calculated conditional density of  $S_1$  conditioned on  $N(t) = 1$ .
- ▶ Suppose we want to calculate the conditional joint density of  $S_1, \dots, S_n$  conditioned on  $N(t) = n$ .

- ▶ Let  $S_1, \dots, S_n$  be the times of the first  $n$  events.
- ▶ We calculated conditional density of  $S_1$  conditioned on  $N(t) = 1$ .
- ▶ Suppose we want to calculate the conditional joint density of  $S_1, \dots, S_n$  conditioned on  $N(t) = n$ .
- ▶ Note that the  $S_i$  have to satisfy  $S_1 < S_2 < \dots < S_n$ .

- ▶ Let  $S_1, \dots, S_n$  be the times of the first  $n$  events.
- ▶ We calculated conditional density of  $S_1$  conditioned on  $N(t) = 1$ .
- ▶ Suppose we want to calculate the conditional joint density of  $S_1, \dots, S_n$  conditioned on  $N(t) = n$ .
- ▶ Note that the  $S_i$  have to satisfy  $S_1 < S_2 < \dots < S_n$ .
- ▶ We can show that the conditional joint density of  $S_1, \dots, S_n$  conditioned on  $N(t) = n$ , would be same as the order statistics of  $n$  iid random variables uniform over  $[0, t]$ .

- Take  $t_i, 1 \leq i \leq n$  satisfying  $0 < t_1 < t_2 < \cdots < t_n < t$ .

- ▶ Take  $t_i, 1 \leq i \leq n$  satisfying  $0 < t_1 < t_2 < \cdots < t_n < t$ .
- ▶ Let  $h_i$  be small positive numbers such that  $t_i + h_i < t_{i+1}, \forall i$ .

- ▶ Take  $t_i, 1 \leq i \leq n$  satisfying  $0 < t_1 < t_2 < \cdots < t_n < t$ .
- ▶ Let  $h_i$  be small positive numbers such that  $t_i + h_i < t_{i+1}, \forall i$ .

$$Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \cdots, n \mid N(t) = n]$$

- ▶ Take  $t_i, 1 \leq i \leq n$  satisfying  $0 < t_1 < t_2 < \cdots < t_n < t$ .
- ▶ Let  $h_i$  be small positive numbers such that  $t_i + h_i < t_{i+1}, \forall i$ .

$$\begin{aligned} & Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \cdots, n \mid N(t) = n] \\ = & \frac{Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \cdots, n, N(t) = n]}{Pr[N(t) = n]} \end{aligned}$$



- ▶ Take  $t_i, 1 \leq i \leq n$  satisfying  $0 < t_1 < t_2 < \dots < t_n < t$ .
- ▶ Let  $h_i$  be small positive numbers such that  $t_i + h_i < t_{i+1}, \forall i$ .

$$\begin{aligned}
 & Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \dots, n \mid N(t) = n] \\
 = & \frac{Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \dots, n, N(t) = n]}{Pr[N(t) = n]} \\
 = & \frac{Pr[1 \text{ event in each } [t_i, t_i + h_i], 1 \leq i \leq n, 0 \text{ in rest of } [0, t]]}{Pr[N(t) = n]}
 \end{aligned}$$

- ▶ Take  $t_i, 1 \leq i \leq n$  satisfying  $0 < t_1 < t_2 < \dots < t_n < t$ .
- ▶ Let  $h_i$  be small positive numbers such that  $t_i + h_i < t_{i+1}, \forall i$ .

$$\begin{aligned}
 & Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \dots, n \mid N(t) = n] \\
 = & \frac{Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \dots, n, N(t) = n]}{Pr[N(t) = n]} \\
 = & \frac{Pr[1 \text{ event in each } [t_i, t_i + h_i], 1 \leq i \leq n, 0 \text{ in rest of } [0, t]]}{Pr[N(t) = n]} \\
 = & \frac{(\prod_{i=1}^n \lambda h_i e^{-\lambda h_i}) e^{-\lambda(t-h_1-\dots-h_n)}}{((\lambda t)^n / n!) e^{-\lambda t}}
 \end{aligned}$$

- ▶ Take  $t_i, 1 \leq i \leq n$  satisfying  $0 < t_1 < t_2 < \dots < t_n < t$ .
- ▶ Let  $h_i$  be small positive numbers such that  $t_i + h_i < t_{i+1}, \forall i$ .

$$\begin{aligned}
 & Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \dots, n \mid N(t) = n] \\
 = & \frac{Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \dots, n, N(t) = n]}{Pr[N(t) = n]} \\
 = & \frac{Pr[1 \text{ event in each } [t_i, t_i + h_i], 1 \leq i \leq n, 0 \text{ in rest of } [0, t]]}{Pr[N(t) = n]} \\
 = & \frac{(\prod_{i=1}^n \lambda h_i e^{-\lambda h_i}) e^{-\lambda(t-h_1-\dots-h_n)}}{((\lambda t)^n / n!) e^{-\lambda t}} \\
 = & \frac{n! h_1 \dots h_n}{t^n}
 \end{aligned}$$

► Thus we have for  $0 < t_1 < \cdots < t_n < t$ ,

$$\frac{\Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \cdots, n \mid N(t) = n]}{h_1 \cdots h_n} = \frac{n!}{t^n}$$

- ▶ Thus we have for  $0 < t_1 < \cdots < t_n < t$ ,

$$\frac{\Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \cdots, n \mid N(t) = n]}{h_1 \cdots h_n} = \frac{n!}{t^n}$$

- ▶ If we now take limit as all  $h_i$  go to zero, the LHS above would be the conditional joint density of  $S_1, \cdots, S_n$  conditioned on  $N(t) = n$ .

- ▶ Thus we have for  $0 < t_1 < \cdots < t_n < t$ ,

$$\frac{\Pr[t_i \leq S_i \leq t_i + h_i, i = 1, \cdots, n \mid N(t) = n]}{h_1 \cdots h_n} = \frac{n!}{t^n}$$

- ▶ If we now take limit as all  $h_i$  go to zero, the LHS above would be the conditional joint density of  $S_1, \cdots, S_n$  conditioned on  $N(t) = n$ .
- ▶ Thus

$$f_{S_1 \cdots S_n | N(t)}(t_1, \cdots, t_n | n) = \frac{n!}{t^n}, \quad 0 < t_1 < \cdots < t_n < t$$

- ▶ Let  $X_1, \dots, X_n$  be iid continuous random variables with common density  $f_x$ .

- ▶ Let  $X_1, \dots, X_n$  be iid continuous random variables with common density  $f_x$ .
- ▶ Recall that  $X_{(k)}$  denotes the  $k^{th}$  smallest of them.



- ▶ Let  $X_1, \dots, X_n$  be iid continuous random variables with common density  $f_x$ .
- ▶ Recall that  $X_{(k)}$  denotes the  $k^{th}$  smallest of them.
- ▶ Then the joint density of  $X_{(1)}, \dots, X_{(n)}$  is given by

$$f(x_1, \dots, x_n) = n! \prod_{i=1}^n f_x(x_i), \quad x_1 < \dots < x_n$$

- ▶ Let  $X_1, \dots, X_n$  be iid continuous random variables with common density  $f_x$ .
- ▶ Recall that  $X_{(k)}$  denotes the  $k^{th}$  smallest of them.
- ▶ Then the joint density of  $X_{(1)}, \dots, X_{(n)}$  is given by

$$f(x_1, \dots, x_n) = n! \prod_{i=1}^n f_x(x_i), \quad x_1 < \dots < x_n$$

- ▶ If  $X_i$  are uniform over  $[0, t]$

$$f(t_1, \dots, t_n) = \frac{n!}{t^n}, \quad 0 < t_1 < \dots < t_n < t$$

# Examples

- ▶ We look at a few simple example problems using Poisson process.

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$E[N(4) - N(2) | N(1) = 3] =$$

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$E[N(4) - N(2) | N(1) = 3] = E[N(4) - N(2)]$$

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$\begin{aligned} E[N(4) - N(2) | N(1) = 3] &= E[N(4) - N(2)] \\ &= E[N(2) - 0] = 2\lambda \end{aligned}$$

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$\begin{aligned} E[N(4) - N(2) | N(1) = 3] &= E[N(4) - N(2)] \\ &= E[N(2) - 0] = 2\lambda \end{aligned}$$

- ▶ Another example;

$$E[S_4] =$$

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$\begin{aligned} E[N(4) - N(2) | N(1) = 3] &= E[N(4) - N(2)] \\ &= E[N(2) - 0] = 2\lambda \end{aligned}$$

- ▶ Another example;

$$E[S_4] = E \left[ \sum_{i=1}^4 T_i \right] = \frac{4}{\lambda}$$



# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$\begin{aligned} E[N(4) - N(2) | N(1) = 3] &= E[N(4) - N(2)] \\ &= E[N(2) - 0] = 2\lambda \end{aligned}$$

- ▶ Another example;

$$E[S_4] = E\left[\sum_{i=1}^4 T_i\right] = \frac{4}{\lambda}$$

- ▶ The memoryless property of exponential rv can be useful

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$\begin{aligned} E[N(4) - N(2) | N(1) = 3] &= E[N(4) - N(2)] \\ &= E[N(2) - 0] = 2\lambda \end{aligned}$$

- ▶ Another example;

$$E[S_4] = E\left[\sum_{i=1}^4 T_i\right] = \frac{4}{\lambda}$$

- ▶ The memoryless property of exponential rv can be useful

$$Pr[S_3 > t | N(1) = 2]$$

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$\begin{aligned} E[N(4) - N(2) | N(1) = 3] &= E[N(4) - N(2)] \\ &= E[N(2) - 0] = 2\lambda \end{aligned}$$

- ▶ Another example;

$$E[S_4] = E\left[\sum_{i=1}^4 T_i\right] = \frac{4}{\lambda}$$

- ▶ The memoryless property of exponential rv can be useful

$$Pr[S_3 > t | N(1) = 2] = \begin{cases} 1 & \text{if } t < 1 \\ \end{cases}$$

# Examples

- ▶ We look at a few simple example problems using Poisson process.

$$\begin{aligned} E[N(4) - N(2) | N(1) = 3] &= E[N(4) - N(2)] \\ &= E[N(2) - 0] = 2\lambda \end{aligned}$$

- ▶ Another example;

$$E[S_4] = E\left[\sum_{i=1}^4 T_i\right] = \frac{4}{\lambda}$$

- ▶ The memoryless property of exponential rv can be useful

$$Pr[S_3 > t | N(1) = 2] = \begin{cases} 1 & \text{if } t < 1 \\ e^{-\lambda(t-1)} & \text{if } t \geq 1 \end{cases}$$