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- ▶ The $X(t)$ is defined by piece-wise constant interpolation of $X(nT)$
- ▶ For large n , $\frac{X(nT)}{s\sqrt{n}}$ would be Gaussian

$$Pr \left[\frac{X(nT)}{s\sqrt{n}} \leq y \right] \approx \Phi(y)$$

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- ▶ We intuitively expect $W(t)$ to be a process with stationary and independent increments and for each t , $W(t)$ is Gaussian with zero mean and variance proportional to t

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- ▶ The mean can be a function of time

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- ▶ The paths are continuous but non-differentiable everywhere
- ▶ This is a deep result

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- ▶ Since $E[X(t)] = 0, \forall t$, we have

$$\text{Cov}(X(t_1), X(t_2)) = E[X(t_1)X(t_2)] = \sigma^2 \min(t_1, t_2)$$

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- ▶ Thus, $X(t_2) - X(t_1)$ is Gaussian with zero mean and variance $\sigma^2(t_2 - t_1)$
- ▶ Since increments are also independent, we can show that all n^{th} order distributions are Gaussian

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- ▶ Hence we can get joint density of $X(t_1), \dots, X(t_n)$ in terms of joint density of Y_1, \dots, Y_n
- ▶ This is how we can get n^{th} order density for any continuous-state process with independent increments

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- ▶ We can write their joint density because we know the means, variances and covariances
- ▶ We can also write the density using the transformation considered earlier

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► Hence we get

$$f_{X(t_1)X(t_2)\cdots X(t_n)}(x_1, \cdots, x_n) = f_{Y_1\cdots Y_n}(x_1, x_2 - x_1, \cdots, x_n - x_{n-1})$$

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► Take $t_1 < t_2 < \cdots < t_n$

$$f_X(x_1, \cdots, x_n; t_1, \cdots, t_n) = f_{Y_1}(x_1)f_{Y_2}(x_2 - x_1) \cdots f_{Y_n}(x_n - x_{n-1})$$

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- ▶ Hence the conditional density is Gaussian with mean bs/t and variance $s(t-s)/t$

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- Thus, conditioned on $X(s)$ the density of $X(t)$ is normal with mean $X(s)$ and variance $(t - s)$

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$$\lim_{t \rightarrow \infty} \mathbb{P}[T_a \leq t] = 2 \cdot \frac{1}{\sqrt{2\pi}} \int_0^{\infty} e^{-\frac{1}{2}y^2} dy = 1$$

f.d. every a

$$\mathbb{P}[T_a < \infty] = 1$$

Ex 1: $E[X(t_1)X(t_2)]$ $t_1 < t_2$ $\begin{matrix} \text{me } X(t_2) \\ X(t_1) \text{ and } X(t_2) \sim (t-1) \end{matrix}$

$$\begin{aligned}
 E[E[X(t_1)X(t_2) | X(t_1)]] &= E[X(t_1) \cdot E[X(t_2) | X(t_1)]] \\
 &= E[X(t_1) \cdot X(t_1)] \\
 &= E[X^2(t_1)]
 \end{aligned}$$

Ex 2: $X(t) \sim \text{BM with } \sigma^2 = 1$

$$\begin{aligned}
 E[X(t_1)X(t_2)X(t_3)] &= E[E[X(t_1)X(t_2)X(t_3) | X(t_1)X(t_2)]] \\
 &= E[X(t_1)X^2(t_2)] \\
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 &= E[X(t_1)E[X^2(t_2) | X(t_1)]] \\
 &= E[X(t_1)(X^2(t_1) + (t_2 - t_1))] \\
 &= E[X(t_1)(t_2 - t_1)] + E[X^3(t_1)] \\
 &= 0
 \end{aligned}$$

Ex 1: Prove $X(t) + X(1)$ $t > 1$, $\sigma^2 = 1$

$$X(t) + X(1) = \underbrace{2X(1)} + \underbrace{X(t) - X(1)}$$

$$2X(1) \sim \mathcal{N}(0, 4\sigma)$$

$$X(t) + X(1) \sim \mathcal{N}(0, 4\sigma + (t-1)) = \mathcal{N}(0, 3\sigma + t)$$

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- ▶ Since $\ln(Y_i)$ are iid, with suitable normalization, the interpolated process $\ln(X(t))$ would be Brownian motion and $X(t)$ would be geometric Brownian motion

Gaussian Processes

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- ▶ Recall that the multivariate Gaussian density is specified by the marginal means, variances and the covariances of the random variables
- ▶ Hence, a general Gaussian process is specified by the mean function and the variance and covariance functions

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Recall that, for $s < t$, conditional density of $X(s)$ conditioned on $X(t) = b$ is gaussian with mean bs/t and variance $s(t-s)/t$

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Thus, for $0 < t < 1$, conditioned on $X(1) = 0$, this process has mean 0 and covariance function $s(1 - t)$, $s < t$

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- ▶ The actual concept involved is rather deep

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- ▶ Any sequence of continuous random variables would be a discrete-time continuous-state process

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- ▶ We may want to know the conditions under which we can prove the sequence converges.

