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- It essentially results in an induced probability space

$$(\Omega, \mathcal{F}, P) \stackrel{X}{\to} (\Re, \mathcal{B}, P_X)$$

where ${\cal B}$ is the Borel σ -algebra and

$$P_X(B) = P[X \in B] = P(\{\omega \in \Omega : X(\omega) \in B\})$$

Recap: Distribution function of a random variable

▶ Let X be a random variable. It distribution function, $F_X: \Re \to \Re$, is defined by

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▶ The distribution function, F_X , completely specifies the probability measure, P_X .

- ► The distribution function satisfies
 - 1. $0 \le F_X(x) \le 1, \ \forall x$
 - 2. $F_X(-\infty) = 0$; $F_X(\infty) = 1$
 - 3. F_X is non-decreasing: $x_1 \le x_2 \implies F_X(x_1) \le F_X(x_2)$
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 $P[a < X \le b] = F_X(b) - F_X(a).$



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- ▶ Let $X \in \{x_1, x_2, \dots\}$
- Its distribution function, F_X is a stair-case function with jump discontinuities at each x_i and the magnitude of the jump at x_i is equal to $P[X=x_i]$

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We can calculate the probability of any event as

$$P[X \in B] = \sum_{\substack{i: \\ x: \in B}} f_X(x_i)$$

▶ X is said to be a continuous random variable if there exists a function $f_X: \Re \to \Re$ satisfying

$$F_X(x) = \int_{-\infty}^x f_X(x) \ dx$$

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A continuous rv takes uncountably many distinct values. However, not every rv that takes uncountably many values is a continuous rv

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In particular,

$$P[a \le X \le b] = \int_a^b f_X(t) dt$$



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- lackbox We can determine distribution of Y given the function g and the distribution of X

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- ▶ Then Y is also discrete and $Y \in \{g(x_1), g(x_2), \dots\}$.
- ▶ We can find the pmf of Y as

$$f_Y(y) = p[Y = y] = P[g(X) = y]$$

$$= P[X \in \{x_i : g(x_i) = y\}]$$

$$= \sum_{\substack{i: \\ g(x_i) = y}} f_X(x_i)$$

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$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right|, \ a \le y \le b$$

where $a = \min(g(\infty), \ g(-\infty))$ and $b = \max(g(\infty), \ g(-\infty))$

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► This theorem is useful in some cases to find the densities of functions of continuous random variables

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- ▶ Note that expectation is defined for all random variables



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$$E[g(X)] = \sum_{i=1}^{\infty} g(x_i) f_X(x_i)$$
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▶ If $X \ge 0$ then $EX \ge 0$

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- ► $E[(X-c)^2] \ge E[(X-EX)^2], \forall c$



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 - $ightharpoonup Var(cX) = c^2 Var(X)$

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▶ If moment of order k is finite then so is moment of order s for s < k.



Recap: Moment Generating function

▶ The moment generating function – $M_X: \Re \to \Re$

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- ▶ We say the mgf exists if $E[e^{tX}] < \infty$ for t in some interval around zero
- ▶ If $M_X(t)$ exists (for $t \in [-a, a]$ for some a > 0) then all its derivatives also exist and

$$\left. \frac{d^k M_X(t)}{dt^k} \right|_{t=0} = E[X^k]$$

Generating function

▶ For $X \in \{0, 1, 2, \cdots\}$ the (probability) generating function of X is defined by

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▶ We can also get the moments:

$$P'_X(1) = EX, \quad P''_X(1) = E[X(X-1)]$$



quantiles of a distribution

▶ Let $p \in (0, 1)$. The number $x \in \Re$ that satisfies

$$P[X \le x] \ge p \quad \text{and} \quad p[X \ge x] \ge 1 - p$$

is called the quantile of order p or the $100p^{th}$ percentile of rv X.

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- ▶ For p = 0.5, it is called the median.

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► A specific instance of this is

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▶ With $EX = \mu$ and $Var(X) = \sigma^2$, we get

$$P[|X - \mu| > k\sigma] \le \frac{1}{k^2}$$

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- ► Given the joint distribution function, probability of any event involving the pair of random variables can be (in principle) calculated.

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▶ Joint distribution function: $F_{XY}: \Re^2 \to \Re$

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▶ Any $F: \Re^2 \to \Re$ satisfying the above would be a joint distribution function.

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- ▶ Given the joint pmf, we can get the joint df as

$$F_{XY}(x,y) = \sum_{\substack{i: \ x_i \le x}} \sum_{\substack{j: \ y_j \le y}} f_{XY}(x_i, y_j)$$



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- ► These are very similar to the properties of the density of a single rv

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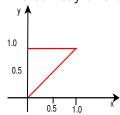
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▶ We can say this density is uniform over the region



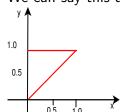
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The figure is not a plot of the density function!!

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- ▶ Given f_{XY} satisfying the above, define

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properties of joint density

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▶ Then we can show F_{XY} is a joint distribution.



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- ▶ The only property left is the special property of F_{XY} we mentioned earlier.

 $\Delta \triangleq F_{XY}(x_2, y_2) - F_{XY}(x_1, y_2) - F_{XY}(x_2, y_1) + F_{XY}(x_1, y_1).$

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- ▶ We need to show $\Delta \ge 0$ if $x_1 < x_2$ and $y_1 < y_2$.
- We have

$$\Delta = \int_{-\infty}^{x_2} \int_{-\infty}^{y_2} f_{XY} \, dy \, dx - \int_{-\infty}^{x_1} \int_{-\infty}^{y_2} f_{XY} \, dy \, dx$$
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$$\Delta \triangleq F_{XY}(x_2, y_2) - F_{XY}(x_1, y_2) - F_{XY}(x_2, y_1) + F_{XY}(x_1, y_1).$$

- ▶ We need to show $\Delta \ge 0$ if $x_1 < x_2$ and $y_1 < y_2$.
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► This actually shows

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► In general

$$P[(X,Y) \in B] = \int_{B} f_{XY}(x,y) \ dx \ dy, \ \forall B \in \mathcal{B}^{2}$$



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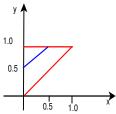
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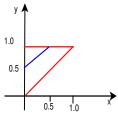
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$$= 2 \frac{y^{2}}{2} \Big|_{0.5}^{1} - y \Big|_{0.5}^{1} = 1 - 0.25 - 1 + 0.5 = 0.25$$

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► The probability of the event we want is the area of the small triangle divided by that of the big triangle.

Marginal Distributions

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► These are simply distribution functions of *X* and *Y* obtained from the joint distribution function.



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- For every value of y, $F_{X|Y}(x|y)$ is a distribution function in the variable x.
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$$\begin{split} F_{X|Y}(x|y_j) &= \frac{P[X \leq x, Y = y_j]}{P[Y = y_j]} = \frac{\sum_{i: x_i \leq x} P[X = x_i, Y = y_j]}{P[Y = y_j]} \\ &= \sum_{i: x_i \leq x} \frac{P[X = x_i, Y = y_j]}{P[Y = y_j]} \end{split}$$

▶ Let: $X \in \{x_1, x_2, \dots\}$ and $Y \in \{y_1, y_2, \dots\}$. Then

$$F_{X|Y}(x|y_j) = P[X \le x|Y = y_j] = \frac{P[X \le x, Y = y_j]}{P[Y = y_j]}$$

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Note that

$$\sum_i f_{X|Y}(x_i|y_j) = 1, \ \forall y_j; \quad \text{and} \quad F_{X|Y}(x|y_j) = \sum_{i:x_i \leq x} f_{X|Y}(x_i|y_j)$$

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Given there is only one head, it is equally likely to occur on any toss.



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 $(P(A) = \sum_{j} P(A|B_j)P(B_j)$ when B_1, \cdots form a partition)

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- ▶ The limit exists for all y where $f_Y(y) > 0$ (and for all x)

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▶ We (once again) have the useful identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) \ f_Y(y) = f_{Y|X}(y|x) f_X(x)$$

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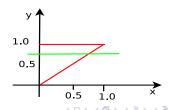
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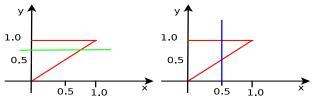
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- ► We can specify the marginal density of one and the conditional density of the other given the first.
- ► This may actually be the model of how the the rv's are generated.

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► We can verify it to be a density

$$-\int_{0}^{1} \ln(y) \ dy = -y \ln(y) \Big|_{0}^{1} + \int_{0}^{1} y \frac{1}{y} \ dy = 1$$



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- While often that gives the right result, one needs to be very careful

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► This is essentially identical to Bayes rule for discrete rv's. We have essentially put the pdf wherever there was pmf

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- ► This gives total probability rule and Bayes rule for random

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