$$F_{XY}: \Re^2 \to \Re$$

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

$$F_{XY}: \Re^2 \to \Re$$

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

- It satisfies
  - 1.  $F_{XY}(-\infty, y) = F_{XY}(x, -\infty) = 0, \forall x, y;$  $F_{XY}(\infty, \infty) = 1$

$$F_{XY}: \Re^2 \to \Re$$

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

- It satisfies
  - 1.  $F_{XY}(-\infty, y) = F_{XY}(x, -\infty) = 0, \forall x, y;$  $F_{XY}(\infty, \infty) = 1$
  - 2.  $F_{XY}$  is non-decreasing in each of its arguments

$$F_{XY}: \Re^2 \to \Re$$

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

- It satisfies
  - 1.  $F_{XY}(-\infty, y) = F_{XY}(x, -\infty) = 0, \forall x, y;$  $F_{XY}(\infty, \infty) = 1$
  - 2.  $F_{XY}$  is non-decreasing in each of its arguments
  - 3.  $F_{XY}$  is right continuous and has left-hand limits in each of its arguments

$$F_{XY}: \Re^2 \to \Re$$

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

- It satisfies
  - 1.  $F_{XY}(-\infty, y) = F_{XY}(x, -\infty) = 0, \forall x, y;$  $F_{XY}(\infty, \infty) = 1$
  - 2.  $F_{XY}$  is non-decreasing in each of its arguments
  - 3.  $F_{XY}$  is right continuous and has left-hand limits in each of its arguments
  - 4. For all  $x_1 < x_2$  and  $y_1 < y_2$

$$F_{XY}(x_2, y_2) - F_{XY}(x_2, y_1) - F_{XY}(x_1, y_2) + F_{XY}(x_1, y_1) \ge 0$$

lacktriangle Given X,Y rv on same probability space, the joint distribution function is

$$F_{XY}: \Re^2 \to \Re$$

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

- It satisfies
  - 1.  $F_{XY}(-\infty, y) = F_{XY}(x, -\infty) = 0, \forall x, y;$  $F_{XY}(\infty, \infty) = 1$
  - 2.  $F_{XY}$  is non-decreasing in each of its arguments
  - 3.  $F_{XY}$  is right continuous and has left-hand limits in each of its arguments
  - 4. For all  $x_1 < x_2$  and  $y_1 < y_2$

$$F_{XY}(x_2, y_2) - F_{XY}(x_2, y_1) - F_{XY}(x_1, y_2) + F_{XY}(x_1, y_1) \ge 0$$

▶ Any  $F: \Re^2 \to \Re$  satisfying the above would be a joint distribution function.

 $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$ 

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .
- ► The joint pmf satisfies:

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .
- ► The joint pmf satisfies:

A1  $f_{XY}(x,y) \ge 0, \forall x,y$  and non-zero only for  $x_i,y_j$  pairs

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .
- ► The joint pmf satisfies:
  - A1  $f_{XY}(x,y) \geq 0, \forall x,y$  and non-zero only for  $x_i,y_j$  pairs
  - A2  $\sum_i \sum_j f_{XY}(x_i, y_j) = 1$

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .
- ▶ The joint pmf satisfies: A1  $f_{XY}(x,y) \ge 0, \forall x,y$  and non-zero only for  $x_i,y_j$  pairs A2  $\sum_i \sum_j f_{XY}(x_i,y_j) = 1$
- Given the joint pmf, we can get the joint df as

$$F_{XY}(x,y) = \sum_{\substack{i: \ x_i \le x}} \sum_{\substack{j: \ y_j \le y}} f_{XY}(x_i, y_j)$$

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .
- ▶ The joint pmf satisfies: A1  $f_{XY}(x,y) \ge 0, \forall x,y$  and non-zero only for  $x_i,y_j$  pairs A2  $\sum_i \sum_j f_{XY}(x_i,y_j) = 1$
- Given the joint pmf, we can get the joint df as

$$F_{XY}(x,y) = \sum_{\substack{i: \ x_i \le x}} \sum_{\substack{j: \ y_j \le y}} f_{XY}(x_i, y_j)$$

▶ Any  $f_{XY}: \Re^2 \to [0, 1]$  satisfying A1, A2 above is a joint pmf. (Because then  $F_{XY}$  satisfies all properties of df).

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .
- The joint pmf satisfies: A1  $f_{XY}(x,y) \geq 0, \forall x,y$  and non-zero only for  $x_i,y_j$  pairs A2  $\sum_i \sum_j f_{XY}(x_i,y_j) = 1$
- Given the joint pmf, we can get the joint df as

$$F_{XY}(x,y) = \sum_{\substack{i: \ x_i \le x}} \sum_{\substack{j: \ y_j \le y}} f_{XY}(x_i, y_j)$$

- ▶ Any  $f_{XY}: \Re^2 \to [0, 1]$  satisfying A1, A2 above is a joint pmf. (Because then  $F_{XY}$  satisfies all properties of df).
- Given the joint pmf, we can (in principle) compute the probability of any event involving the two discrete random variables.

- $X \in \{x_1, x_2, \cdots\}, Y \in \{y_1, y_2, \cdots\}$
- ▶ The joint pmf:  $f_{XY}(x,y) = P[X = x, Y = y]$ .
- The joint pmf satisfies:

  A1  $f_{XY}(x,y) \ge 0, \forall x,y$  and non-zero only for  $x_i,y_j$  pairs

  A2  $\sum_i \sum_j f_{XY}(x_i,y_j) = 1$
- Given the joint pmf, we can get the joint df as

$$F_{XY}(x,y) = \sum_{\substack{i: \ x_i \le x}} \sum_{\substack{j: \ y_i < y}} f_{XY}(x_i, y_j)$$

- ▶ Any  $f_{XY}: \Re^2 \to [0, 1]$  satisfying A1, A2 above is a joint pmf. (Because then  $F_{XY}$  satisfies all properties of df).
- ► Given the joint pmf, we can (in principle) compute the probability of any event involving the two discrete random variables.

$$P[(X,Y) \in B] = \sum_{i,j} f_{XY}(x_i, y_j)$$

▶ Two cont rv X, Y have a joint density  $f_{XY}$  if

$$F_{XY}(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{XY}(x',y') dy' dx', \ \forall x,y$$

▶ The joint density  $f_{XY}$  satisfies the following

▶ Two cont rv X,Y have a joint density  $f_{XY}$  if

$$F_{XY}(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{XY}(x',y') dy' dx', \ \forall x,y$$

- ▶ The joint density  $f_{XY}$  satisfies the following
  - 1.  $f_{XY}(x,y) \ge 0, \ \forall x,y$

▶ Two cont rv X,Y have a joint density  $f_{XY}$  if

$$F_{XY}(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{XY}(x',y') dy' dx', \ \forall x,y$$

- ▶ The joint density  $f_{XY}$  satisfies the following
  - 1.  $f_{XY}(x,y) \ge 0, \ \forall x,y$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x', y') dy' dx' = 1$

lacktriangle Two cont rv X,Y have a joint density  $f_{XY}$  if

$$F_{XY}(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{XY}(x',y') dy' dx', \ \forall x,y$$

- ▶ The joint density  $f_{XY}$  satisfies the following
  - 1.  $f_{XY}(x,y) \ge 0, \ \forall x,y$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x', y') dy' dx' = 1$
- ▶ Any function  $f_{XY}: \Re^2 \to \Re$  satisfying the above two is a joint density function. (Then the above  $F_{XY}$  can be shown to be a joint df).

ightharpoonup Two cont rv X, Y have a joint density  $f_{XY}$  if

$$F_{XY}(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f_{XY}(x',y') dy' dx', \quad \forall x, y$$

- ightharpoonup The joint density  $f_{XY}$  satisfies the following
  - 1.  $f_{XY}(x,y) > 0, \ \forall x,y$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x', y') dy' dx' = 1$
- Any function  $f_{XY}: \Re^2 \to \Re$  satisfying the above two is a joint density function. (Then the above  $F_{XY}$  can be shown to be a joint df).
- We also have

$$P[x_1 \le X \le x_2, y_1 \le Y \le y_2] = \int_{0}^{x_2} \int_{0}^{y_2} f_{XY} \, dy \, dx$$

and, in general,

, in general, 
$$P[(X,Y)\in B]=\int_{\mathbb{B}^*}f_{XY}(x,y)\;dx\;dy,\;\;\forall B\in\mathcal{B}^2$$
 P S Sastry, IISc, E1 222 Aug 2021 3/53

### Recap Marginals

lacktriangle Marginal distribution functions of X, Y are

$$F_X(x) = F_{XY}(x, \infty); \quad F_Y(y) = F_{XY}(\infty, y)$$

### Recap Marginals

lacktriangle Marginal distribution functions of X,Y are

$$F_X(x) = F_{XY}(x, \infty); \quad F_Y(y) = F_{XY}(\infty, y)$$

lacksquare X,Y discrete with joint pmf  $f_{XY}$ . The marginal pmfs are

$$f_X(x) = \sum_{y} f_{XY}(x, y); \quad f_Y(y) = \sum_{x} f_{XY}(x, y)$$

### Recap Marginals

ightharpoonup Marginal distribution functions of X, Y are

$$F_X(x) = F_{XY}(x, \infty); \quad F_Y(y) = F_{XY}(\infty, y)$$

lackbox X,Y discrete with joint pmf  $f_{XY}$ . The marginal pmfs are

$$f_X(x) = \sum_{y} f_{XY}(x, y); \quad f_Y(y) = \sum_{x} f_{XY}(x, y)$$

▶ If X,Y have joint pdf  $f_{XY}$  then the marginal pdf are

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x,y) \ dy \quad f_Y(y) = \int_{-\infty}^{\infty} f_{XY}(x,y) \ dx$$

▶ Let:  $X \in \{x_1, x_2, \dots\}$  and  $Y \in \{y_1, y_2, \dots\}$ . Then

▶ Let:  $X \in \{x_1, x_2, \dots\}$  and  $Y \in \{y_1, y_2, \dots\}$ . Then

$$F_{X|Y}(x|y_j) = P[X \le x|Y = y_j]$$

▶ Let:  $X \in \{x_1, x_2, \dots\}$  and  $Y \in \{y_1, y_2, \dots\}$ . Then

$$F_{X|Y}(x|y_j) = P[X \le x|Y = y_j] = \frac{P[X \le x, Y = y_j]}{P[Y = y_j]}$$

▶ Let:  $X \in \{x_1, x_2, \dots\}$  and  $Y \in \{y_1, y_2, \dots\}$ . Then

$$F_{X|Y}(x|y_j) = P[X \le x|Y = y_j] = \frac{P[X \le x, Y = y_j]}{P[Y = y_j]}$$

(We define  $F_{X|Y}(x|y)$  only when  $y = y_j$  for some j).

▶ Let:  $X \in \{x_1, x_2, \dots\}$  and  $Y \in \{y_1, y_2, \dots\}$ . Then

$$F_{X|Y}(x|y_j) = P[X \le x|Y = y_j] = \frac{P[X \le x, Y = y_j]}{P[Y = y_j]}$$

(We define  $F_{X|Y}(x|y)$  only when  $y = y_j$  for some j).

▶ For each  $y_i$ ,  $F_{X|Y}(x|y_i)$  is a df of a discrete rv in x.

▶ Let:  $X \in \{x_1, x_2, \dots\}$  and  $Y \in \{y_1, y_2, \dots\}$ . Then

$$F_{X|Y}(x|y_j) = P[X \le x|Y = y_j] = \frac{P[X \le x, Y = y_j]}{P[Y = y_j]}$$

(We define  $F_{X|Y}(x|y)$  only when  $y = y_j$  for some j).

- ▶ For each  $y_i$ ,  $F_{X|Y}(x|y_i)$  is a df of a discrete rv in x.
- ▶ The pmf corresponding to this df is called conditional pmf

$$f_{X|Y}(x_i|y_j) = P[X = x_i|Y = y_j]$$

▶ Let:  $X \in \{x_1, x_2, \dots\}$  and  $Y \in \{y_1, y_2, \dots\}$ . Then

$$F_{X|Y}(x|y_j) = P[X \le x|Y = y_j] = \frac{P[X \le x, Y = y_j]}{P[Y = y_j]}$$

(We define  $F_{X|Y}(x|y)$  only when  $y = y_j$  for some j).

- ▶ For each  $y_i$ ,  $F_{X|Y}(x|y_i)$  is a df of a discrete rv in x.
- ▶ The pmf corresponding to this df is called conditional pmf

$$f_{X|Y}(x_i|y_j) = P[X = x_i|Y = y_j] = \frac{f_{XY}(x_i, y_j)}{f_Y(y_j)}$$

▶ The conditional mass function is

$$f_{X|Y}(x_i|y_j) = P[X = x_i|Y = y_j] = \frac{f_{XY}(x_i, y_j)}{f_Y(y_j)}$$

▶ The conditional mass function is

$$f_{X|Y}(x_i|y_j) = P[X = x_i|Y = y_j] = \frac{f_{XY}(x_i, y_j)}{f_Y(y_i)}$$

► This gives us the useful identity

$$f_{XY}(x_i, y_j) = f_{X|Y}(x_i|y_j) f_Y(y_j)$$

▶ The conditional mass function is

$$f_{X|Y}(x_i|y_j) = P[X = x_i|Y = y_j] = \frac{f_{XY}(x_i, y_j)}{f_Y(y_i)}$$

► This gives us the useful identity

$$f_{XY}(x_i, y_j) = f_{X|Y}(x_i|y_j) f_Y(y_j)$$

This gives us the total probability rule for rv's

$$f_X(x_i) = \sum_j f_{XY}(x_i, y_j) = \sum_j f_{X|Y}(x_i|y_j) f_Y(y_j)$$

▶ The conditional mass function is

$$f_{X|Y}(x_i|y_j) = P[X = x_i|Y = y_j] = \frac{f_{XY}(x_i, y_j)}{f_{Y}(y_i)}$$

► This gives us the useful identity

$$f_{XY}(x_i, y_j) = f_{X|Y}(x_i|y_j) f_Y(y_j)$$

This gives us the total probability rule for rv's

$$f_X(x_i) = \sum_{i} f_{XY}(x_i, y_j) = \sum_{i} f_{X|Y}(x_i|y_j) f_Y(y_j)$$

Also gives us Bayes rule for discrete rv

$$f_{X|Y}(x_i|y_j) = \frac{f_{Y|X}(y_j|x_i)f_X(x_i)}{\sum_i f_{Y|X}(y_j|x_i)f_X(x_i)}$$

### Recap: Conditional densities

▶ Let X, Y have joint density  $f_{XY}$ .

### Recap: Conditional densities

- Let X, Y have joint density  $f_{XY}$ .
- ightharpoonup The conditional df of X given Y is

$$F_{X|Y}(x|y) = \lim_{\delta \to 0} P[X \le x|Y \in [y, \ y + \delta]]$$

- Let X, Y have joint density  $f_{XY}$ .
- ightharpoonup The conditional df of X given Y is

$$F_{X|Y}(x|y) = \lim_{\delta \to 0} P[X \le x | Y \in [y, y + \delta]]$$

▶ This exists if  $f_Y(y) > 0$  and then it has a density:

$$F_{X|Y}(x|y) = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

- Let X, Y have joint density  $f_{XY}$ .
- ightharpoonup The conditional df of X given Y is

$$F_{X|Y}(x|y) = \lim_{\delta \to 0} P[X \le x | Y \in [y, y + \delta]]$$

▶ This exists if  $f_Y(y) > 0$  and then it has a density:

$$F_{X|Y}(x|y) = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

► This conditional density is given by

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

- Let X, Y have joint density  $f_{XY}$ .
- ightharpoonup The conditional df of X given Y is

$$F_{X|Y}(x|y) = \lim_{\delta \to 0} P[X \le x | Y \in [y, y + \delta]]$$

▶ This exists if  $f_Y(y) > 0$  and then it has a density:

$$F_{X|Y}(x|y) = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

This conditional density is given by

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

▶ We (once again) have the useful identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) f_Y(y)$$

- Let X, Y have joint density  $f_{XY}$ .
- ightharpoonup The conditional df of X given Y is

$$F_{X|Y}(x|y) = \lim_{\delta \to 0} P[X \le x|Y \in [y, y + \delta]]$$

▶ This exists if  $f_Y(y) > 0$  and then it has a density:

$$F_{X|Y}(x|y) = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

This conditional density is given by

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

▶ We (once again) have the useful identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) \ f_Y(y) = f_{Y|X}(y|x) f_X(x)$$



► We have the identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) f_Y(y)$$

► We have the identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) f_Y(y)$$

▶ This gives us continuous analogue of total probability rule:

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) \ dy = \int_{-\infty}^{\infty} f_{X|Y}(x|y) \ f_Y(y) \ dy$$

► We have the identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) f_Y(y)$$

▶ This gives us continuous analogue of total probability rule:

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) \ dy = \int_{-\infty}^{\infty} f_{X|Y}(x|y) \ f_Y(y) \ dy$$

This also gives us Bayes rule for continuous rv

► We have the identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) f_Y(y)$$

▶ This gives us continuous analogue of total probability rule:

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) \ dy = \int_{-\infty}^{\infty} f_{X|Y}(x|y) \ f_Y(y) \ dy$$

► This also gives us Bayes rule for continuous rv

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)}$$

► We have the identity

$$f_{XY}(x,y) = f_{X|Y}(x|y) f_Y(y)$$

▶ This gives us continuous analogue of total probability rule:

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x,y) \ dy = \int_{-\infty}^{\infty} f_{X|Y}(x|y) \ f_Y(y) \ dy$$

This also gives us Bayes rule for continuous rv

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)}$$
$$= \frac{f_{Y|X}(y|x)f_X(x)}{\int_{-\infty}^{\infty} f_{Y|X}(y|x)f_X(x) dx}$$

▶ Now, let *X* be a continuous rv and let *Y* be discrete rv.

- ▶ Now, let *X* be a continuous rv and let *Y* be discrete rv.
- ightharpoonup We can define  $F_{X|Y}$  as

$$F_{X|Y}(x|y) = P[X \le x|Y = y]$$

- Now, let X be a continuous rv and let Y be discrete rv.
- ightharpoonup We can define  $F_{X|Y}$  as

$$F_{X|Y}(x|y) = P[X \le x|Y = y]$$

▶ Since *X* is continuous rv, this df would have a density

$$F_{X|Y}(x|y) = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

- Now, let *X* be a continuous rv and let *Y* be discrete rv.
- ightharpoonup We can define  $F_{X|Y}$  as

$$F_{X|Y}(x|y) = P[X \le x|Y = y]$$

▶ Since *X* is continuous rv, this df would have a density

$$F_{X|Y}(x|y) = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

► Hence we can write

$$P[X \le x, Y = y] = F_{X|Y}(x|y)P[Y = y]$$

- Now, let X be a continuous rv and let Y be discrete rv.
- ightharpoonup We can define  $F_{X|Y}$  as

$$F_{X|Y}(x|y) = P[X \le x|Y = y]$$

▶ Since *X* is continuous rv, this df would have a density

$$F_{X|Y}(x|y) = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

Hence we can write

$$P[X \le x, Y = y] = F_{X|Y}(x|y)P[Y = y]$$

$$= \int_{-x}^{x} f_{X|Y}(x'|y) f_{Y}(y) dx'$$

$$F_X(x) = P[X \le x]$$

$$F_X(x) = P[X \le x] = \sum_{x} P[X \le x, Y = y]$$

$$F_X(x) = P[X \le x] = \sum_y P[X \le x, Y = y]$$
  
=  $\sum_y \int_{-\infty}^x f_{X|Y}(x'|y) f_Y(y) dx'$ 

$$F_X(x) = P[X \le x] = \sum_{y} P[X \le x, Y = y]$$

$$= \sum_{y} \int_{-\infty}^{x} f_{X|Y}(x'|y) f_Y(y) dx'$$

$$= \int_{-\infty}^{x} \sum_{y} f_{X|Y}(x'|y) f_Y(y) dx'$$

$$F_X(x) = P[X \le x] = \sum_{y} P[X \le x, Y = y]$$

$$= \sum_{y} \int_{-\infty}^{x} f_{X|Y}(x'|y) f_Y(y) dx'$$

$$= \int_{-\infty}^{x} \sum_{x} f_{X|Y}(x'|y) f_Y(y) dx'$$

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

$$F_X(x) = P[X \le x] = \sum_{y} P[X \le x, Y = y]$$

$$= \sum_{y} \int_{-\infty}^{x} f_{X|Y}(x'|y) f_Y(y) dx'$$

$$= \int_{-\infty}^{x} \sum_{y} f_{X|Y}(x'|y) f_Y(y) dx'$$

► This gives us

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

▶ This is another version of total probability rule.

$$F_X(x) = P[X \le x] = \sum_y P[X \le x, Y = y]$$

$$= \sum_y \int_{-\infty}^x f_{X|Y}(x'|y) f_Y(y) dx'$$

$$= \int_{-\infty}^x \sum_y f_{X|Y}(x'|y) f_Y(y) dx'$$

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

- ▶ This is another version of total probability rule.
- ightharpoonup Earlier we derived this when X, Y are discrete.

$$F_X(x) = P[X \le x] = \sum_y P[X \le x, Y = y]$$

$$= \sum_y \int_{-\infty}^x f_{X|Y}(x'|y) f_Y(y) dx'$$

$$= \int_{-\infty}^x \sum_y f_{X|Y}(x'|y) f_Y(y) dx'$$

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

- ▶ This is another version of total probability rule.
- ▶ Earlier we derived this when *X,Y* are discrete.
- ▶ The formula is true even when X is continuous

$$F_X(x) = P[X \le x] = \sum_y P[X \le x, Y = y]$$

$$= \sum_y \int_{-\infty}^x f_{X|Y}(x'|y) f_Y(y) dx'$$

$$= \int_{-\infty}^x \sum_y f_{X|Y}(x'|y) f_Y(y) dx'$$

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

- ▶ This is another version of total probability rule.
- ▶ Earlier we derived this when *X,Y* are discrete.
- ► The formula is true even when X is continuous Only difference is we need to take f<sub>X</sub> as the density of X.

ightharpoonup When X,Y are discrete we have

$$f_X(x) = \sum_{x} f_{X|Y}(x|y) f_Y(y)$$

$$lacktriangle$$
 When  $X,Y$  are discrete we have

 $f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_{y} P[X = x|Y = y] P[Y = y]$ 

▶ When *X,Y* are discrete we have

$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_y P[X = x|Y = y] P[Y = y]$$

▶ When X is continuous and Y is discrete, we defined  $f_{X|Y}(x|y)$  to be the density corresponding to  $F_{X|Y}(x|y)$ 

▶ When *X,Y* are discrete we have

$$f_X(x) = \sum_{x} f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_{x} P[X = x|Y = y] P[Y = y]$$

▶ When X is continuous and Y is discrete, we defined  $f_{X|Y}(x|y)$  to be the density corresponding to  $F_{X|Y}(x|y) = P[X \le x|Y = y]$ 

$$f_X(x) = \sum_{x} f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_{x} P[X = x|Y = y] P[Y = y]$$

- When X is continuous and Y is discrete, we defined  $f_{X|Y}(x|y)$  to be the density corresponding to  $F_{X|Y}(x|y) = P[X \le x|Y = y]$
- ► Then we once again get

$$f_X(x) = \sum_{x} f_{X|Y}(x|y) f_Y(y)$$

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_{y} P[X = x|Y = y] P[Y = y]$$

- ▶ When X is continuous and Y is discrete, we defined  $f_{X|Y}(x|y)$  to be the density corresponding to  $F_{X|Y}(x|y) = P[X \le x|Y = y]$
- ► Then we once again get

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

Now,  $f_X$  is density (and not a mass function).

$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_y P[X = x|Y = y] P[Y = y]$$

- When X is continuous and Y is discrete, we defined  $f_{X|Y}(x|y)$  to be the density corresponding to  $F_{X|Y}(x|y) = P[X \le x|Y = y]$
- ► Then we once again get

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

Now,  $f_X$  is density (and not a mass function).

Suppose  $Y \in \{1, 2, 3\}$  and  $f_Y(i) = \lambda_i$ . Let  $f_{X|Y}(x|i) = f_i(x)$ . Then

$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_y P[X = x|Y = y] P[Y = y]$$

- When X is continuous and Y is discrete, we defined  $f_{X|Y}(x|y)$  to be the density corresponding to  $F_{X|Y}(x|y) = P[X \le x|Y = y]$
- ► Then we once again get

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

Now,  $f_X$  is density (and not a mass function).

Suppose  $Y \in \{1, 2, 3\}$  and  $f_Y(i) = \lambda_i$ . Let  $f_{X|Y}(x|i) = f_i(x)$ . Then

$$f_X(x) = \lambda_1 f_1(x) + \lambda_2 f_2(x) + \lambda_3 f_3(x)$$

ightharpoonup When X,Y are discrete we have

$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y) \ (P[X = x] = \sum_y P[X = x|Y = y] P[Y = y]$$

When X is continuous and Y is discrete, we defined  $f_{X|Y}(x|y)$  to be the density corresponding to  $F_{X|Y}(x|y) = P[X \le x|Y = y]$ 

► Then we once again get

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

Now,  $f_X$  is density (and not a mass function).

Now,  $f_X$  is defisity (and not a mass function

Suppose  $Y \in \{1, 2, 3\}$  and  $f_Y(i) = \lambda_i$ . Let  $f_{X|Y}(x|i) = f_i(x)$ . Then

$$f_X(x) = \lambda_1 f_1(x) + \lambda_2 f_2(x) + \lambda_3 f_3(x)$$

Called a mixture density model

$$F_{X|Y}(x|y) = P[X \le x|Y = y]$$

$$F_{X|Y}(x|y) = P[X \le x|Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

$$F_{X|Y}(x|y) = P[X \le x|Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

▶ We also have

$$P[X \le x, Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) f_{Y}(y) dx'$$

$$F_{X|Y}(x|y) = P[X \le x|Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) \ dx'$$

▶ We also have

$$P[X \le x, Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) f_{Y}(y) dx'$$

► Hence we can define a 'joint density'

$$f_{XY}(x,y) = f_{X|Y}(x|y)f_Y(y)$$

lackbox Continuing with X continuous rv and Y discrete. We have

$$F_{X|Y}(x|y) = P[X \le x|Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) dx'$$

▶ We also have

$$P[X \le x, Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) f_Y(y) dx'$$

► Hence we can define a 'joint density'

$$f_{XY}(x,y) = f_{X|Y}(x|y)f_Y(y)$$

▶ This is a kind of mixed density and mass function.

lackbox Continuing with X continuous rv and Y discrete. We have

$$F_{X|Y}(x|y) = P[X \le x|Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) dx'$$

We also have

$$P[X \le x, Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) f_{Y}(y) dx'$$

► Hence we can define a 'joint density'

$$f_{XY}(x,y) = f_{X|Y}(x|y)f_Y(y)$$

- ▶ This is a kind of mixed density and mass function.
- ▶ We will not be using such 'joint densities' here

ightharpoonup Continuing with X continuous rv and Y discrete

- ► Continuing with *X* continuous rv and *Y* discrete
- ► Can we define  $f_{Y|X}(y|x)$ ?

- Continuing with X continuous rv and Y discrete
- ightharpoonup Can we define  $f_{Y|X}(y|x)$ ?
- ▶ Since *Y* is discrete, this (conditional) mass function is

$$f_{Y|X}(y|x) = P[Y = y|X = x]$$

- Continuing with X continuous rv and Y discrete
- ightharpoonup Can we define  $f_{Y|X}(y|x)$ ?
- ▶ Since *Y* is discrete, this (conditional) mass function is

$$f_{Y|X}(y|x) = P[Y = y|X = x]$$

But the conditioning event has zero prob

- ► Continuing with *X* continuous rv and *Y* discrete
- ightharpoonup Can we define  $f_{Y|X}(y|x)$ ?
- ▶ Since *Y* is discrete, this (conditional) mass function is

$$f_{Y|X}(y|x) = P[Y = y|X = x]$$

- Continuing with X continuous rv and Y discrete
- ightharpoonup Can we define  $f_{Y|X}(y|x)$ ?
- ► Since *Y* is discrete, this (conditional) mass function is

$$f_{Y|X}(y|x) = P[Y = y|X = x]$$

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y|X \in [x, x + \delta]]$$

- Continuing with X continuous rv and Y discrete
- ightharpoonup Can we define  $f_{Y|X}(y|x)$ ?
- ightharpoonup Since Y is discrete, this (conditional) mass function is

$$f_{Y|X}(y|x) = P[Y = y|X = x]$$

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y|X \in [x, x + \delta]]$$

► For simplifying this we note the following:

$$P[X \le x, Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) f_Y(y) dx'$$

- Continuing with X continuous rv and Y discrete
- ightharpoonup Can we define  $f_{Y|X}(y|x)$ ?
- ightharpoonup Since Y is discrete, this (conditional) mass function is

$$f_{Y|X}(y|x) = P[Y = y|X = x]$$

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y|X \in [x, x + \delta]]$$

► For simplifying this we note the following:

$$P[X \le x, Y = y] = \int_{-\infty}^{x} f_{X|Y}(x'|y) f_{Y}(y) dx'$$

$$\Rightarrow P[X \in [x, x+\delta], Y = y] = \int_{-\infty}^{x+\delta} f_{X|Y}(x'|y) f_Y(y) dx'$$

# ► We have

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y|X \in [x, x + \delta]]$$

### ► We have

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y | X \in [x, x + \delta]]$$
$$= \lim_{\delta \to 0} \frac{P[Y = y, X \in [x, x + \delta]]}{P[X \in [x, x + \delta]]}$$

#### ▶ We have

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y|X \in [x, x + \delta]]$$

$$= \lim_{\delta \to 0} \frac{P[Y = y, X \in [x, x + \delta]]}{P[X \in [x, x + \delta]]}$$

$$= \lim_{\delta \to 0} \frac{\int_{x}^{x+\delta} f_{X|Y}(x'|y) f_{Y}(y) dx'}{\int_{x}^{x+\delta} f_{X}(x') dx'}$$

### ► We have

$$\begin{split} f_{Y|X}(y|x) &= \lim_{\delta \to 0} P[Y = y | X \in [x, x + \delta]] \\ &= \lim_{\delta \to 0} \frac{P[Y = y, X \in [x, x + \delta]]}{P[X \in [x, x + \delta]]} \\ &= \lim_{\delta \to 0} \frac{\int_x^{x + \delta} f_{X|Y}(x'|y) f_Y(y) dx'}{\int_x^{x + \delta} f_X(x') dx'} \\ &= \lim_{\delta \to 0} \frac{f_{X|Y}(x|y)\delta f_Y(y)}{f_X(x) \delta} \end{split}$$

# ► We have

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y | X \in [x, x + \delta]]$$

$$= \lim_{\delta \to 0} \frac{P[Y = y, X \in [x, x + \delta]]}{P[X \in [x, x + \delta]]}$$

$$= \lim_{\delta \to 0} \frac{\int_{x}^{x + \delta} f_{X|Y}(x'|y) f_{Y}(y) dx'}{\int_{x}^{x + \delta} f_{X}(x') dx'}$$

$$= \lim_{\delta \to 0} \frac{f_{X|Y}(x|y) \delta f_{Y}(y)}{f_{X}(x) \delta}$$

$$= \frac{f_{X|Y}(x|y) f_{Y}(y)}{f_{Y}(x)}$$

We have

$$f_{Y|X}(y|x) = \lim_{\delta \to 0} P[Y = y | X \in [x, x + \delta]]$$

$$= \lim_{\delta \to 0} \frac{P[Y = y, X \in [x, x + \delta]]}{P[X \in [x, x + \delta]]}$$

$$= \lim_{\delta \to 0} \frac{\int_{x}^{x+\delta} f_{X|Y}(x'|y) f_{Y}(y) dx'}{\int_{x}^{x+\delta} f_{X}(x') dx'}$$

$$= \lim_{\delta \to 0} \frac{f_{X|Y}(x|y)\delta f_{Y}(y)}{f_{X}(x) \delta}$$

$$= \frac{f_{X|Y}(x|y) f_{Y}(y)}{f_{X}(x)}$$

► This gives us further versions of total probability rule and Bayes rule.

First let us look at the total probability rule possibilities

- First let us look at the total probability rule possibilities
- ▶ When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

- First let us look at the total probability rule possibilities
- lacktriangle When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

- First let us look at the total probability rule possibilities
- lacktriangle When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

▶ Since  $f_{X|Y}$  is a density (corresponding to  $F_{X|Y}$ ),

$$\int_{-\infty}^{\infty} f_{X|Y}(x|y) \ dx = 1$$

- First let us look at the total probability rule possibilities
- lacktriangle When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

▶ Since  $f_{X|Y}$  is a density (corresponding to  $F_{X|Y}$ ),

$$\int_{-\infty}^{\infty} f_{X|Y}(x|y) \ dx = 1$$

► Hence we get

$$f_Y(y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x) f_X(x) \ dx$$

- First let us look at the total probability rule possibilities
- ightharpoonup When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

▶ Since  $f_{X|Y}$  is a density (corresponding to  $F_{X|Y}$ ),

$$\int_{-\infty}^{\infty} f_{X|Y}(x|y) \ dx = 1$$

► Hence we get

$$f_Y(y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x) f_X(x) \ dx$$

► Earlier we derived the same formula when *X,Y* have a joint density.

**1.** 
$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

**1.** 
$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y)$$

 $\blacktriangleright$  We first derived this when X,Y are discrete.

**1.** 
$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y)$$

- ▶ We first derived this when *X,Y* are discrete.
- ▶ But now we proved this holds when Y is discrete

1. 
$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y)$$

- ▶ We first derived this when *X,Y* are discrete.
- ▶ But now we proved this holds when Y is discrete If X is continuous the  $f_X, f_{X|Y}$  are densities; If X is also discrete they are mass functions

1. 
$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y)$$

- ▶ We first derived this when *X,Y* are discrete.
- ▶ But now we proved this holds when Y is discrete If X is continuous the  $f_X, f_{X|Y}$  are densities; If X is also discrete they are mass functions

**2.** 
$$f_Y(y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x) f_X(x) \ dx$$

1. 
$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y)$$

- ▶ We first derived this when *X,Y* are discrete.
- ▶ But now we proved this holds when Y is discrete If X is continuous the  $f_X$ ,  $f_{X|Y}$  are densities; If X is also discrete they are mass functions

**2.** 
$$f_Y(y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x) f_X(x) \ dx$$

ightharpoonup We first proved it when X,Y have a joint density

1. 
$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y)$$

- ▶ We first derived this when *X,Y* are discrete.
- ▶ But now we proved this holds when Y is discrete If X is continuous the  $f_X, f_{X|Y}$  are densities; If X is also discrete they are mass functions

**2.** 
$$f_Y(y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x) f_X(x) \ dx$$

We first proved it when X, Y have a joint density We now know it holds also when X is cont and Y is discrete. In that case  $f_Y$  is a mass function

 $\blacktriangleright$  When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

 $\blacktriangleright$  When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

► This once again gives rise to Bayes rule:

$$f_{Y|X}(y|x) = \frac{f_{X|Y}(x|y) f_Y(y)}{f_X(x)} \quad f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)}$$

lacktriangle When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

▶ This once again gives rise to Bayes rule:

$$f_{Y|X}(y|x) = \frac{f_{X|Y}(x|y) f_Y(y)}{f_X(x)} \quad f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)}$$

► Earlier we showed this hold when *X,Y* are both discrete or both continuous.

lacktriangle When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

▶ This once again gives rise to Bayes rule:

$$f_{Y|X}(y|x) = \frac{f_{X|Y}(x|y) f_Y(y)}{f_X(x)} \quad f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)}$$

- ightharpoonup Earlier we showed this hold when X,Y are both discrete or both continuous.
- ► Thus Bayes rule holds in all four possible scenarios

▶ When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

▶ This once again gives rise to Bayes rule:

$$f_{Y|X}(y|x) = \frac{f_{X|Y}(x|y) f_Y(y)}{f_X(x)} \quad f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)}$$

- ightharpoonup Earlier we showed this hold when X,Y are both discrete or both continuous.
- ▶ Thus Bayes rule holds in all four possible scenarios
- lackbox Only difference is we need to interpret  $f_X$  or  $f_{X|Y}$  as mass functions when X is discrete and as densities when X is a continuous ry

▶ When X is continuous rv and Y is discrete rv, we derived

$$f_{Y|X}(y|x)f_X(x) = f_{X|Y}(x|y) f_Y(y)$$

► This once again gives rise to Bayes rule:

$$f_{Y|X}(y|x) = \frac{f_{X|Y}(x|y) \ f_{Y}(y)}{f_{X}(x)} \quad f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_{X}(x)}{f_{Y}(y)}$$

- ► Earlier we showed this hold when *X,Y* are both discrete or both continuous.
- Thus Bayes rule holds in all four possible scenarios
- lackbox Only difference is we need to interpret  $f_X$  or  $f_{X|Y}$  as mass functions when X is discrete and as densities when X is a continuous rv
- ► In general, one refers to these always as densities since the actual meaning would be clear from context.

# Example

Consider a communication system. The transmitter puts out 0 or 5 volts for the bits of 0 and 1, and, volage measured by the receiver is the sent voltage plus noise added by the channel.

- Consider a communication system. The transmitter puts out 0 or 5 volts for the bits of 0 and 1, and, volage measured by the receiver is the sent voltage plus noise added by the channel.
- We assume noise has Gaussian density with mean zero and variance  $\sigma^2$ .

- Consider a communication system. The transmitter puts out 0 or 5 volts for the bits of 0 and 1, and, volage measured by the receiver is the sent voltage plus noise added by the channel.
- We assume noise has Gaussian density with mean zero and variance  $\sigma^2$ .
- ▶ We may want the probability that the sent bit is 1 when measured voltage at the receiver is x to decide what is sent.

- Consider a communication system. The transmitter puts out 0 or 5 volts for the bits of 0 and 1, and, volage measured by the receiver is the sent voltage plus noise added by the channel.
- We assume noise has Gaussian density with mean zero and variance  $\sigma^2$ .
- ▶ We may want the probability that the sent bit is 1 when measured voltage at the receiver is x to decide what is sent.
- Let X be the measured voltage and let Y be sent bit.

- Consider a communication system. The transmitter puts out 0 or 5 volts for the bits of 0 and 1, and, volage measured by the receiver is the sent voltage plus noise added by the channel.
- We assume noise has Gaussian density with mean zero and variance  $\sigma^2$ .
- ▶ We may want the probability that the sent bit is 1 when measured voltage at the receiver is x to decide what is sent.
- Let X be the measured voltage and let Y be sent bit.
- ▶ We want to calculate  $f_{Y|X}(1|x)$ .

- Consider a communication system. The transmitter puts out 0 or 5 volts for the bits of 0 and 1, and, volage measured by the receiver is the sent voltage plus noise added by the channel.
- We assume noise has Gaussian density with mean zero and variance  $\sigma^2$ .
- ▶ We may want the probability that the sent bit is 1 when measured voltage at the receiver is *x* to decide what is sent.
- Let X be the measured voltage and let Y be sent bit.
- ▶ We want to calculate  $f_{Y|X}(1|x)$ .
- ▶ We want to use the Bayes rule to calculate this

ightharpoonup We need  $f_{X|Y}$ .

▶ We need  $f_{X|Y}$ . What does our model say?

- ▶ We need  $f_{X|Y}$ . What does our model say?
- $f_{X|Y}(x|1)$  is Gaussian with mean 5 and variance  $\sigma^2$  and  $f_{X|Y}(x|0)$  is Gaussian with mean zero and variance  $\sigma^2$

- ▶ We need  $f_{X|Y}$ . What does our model say?
- $f_{X|Y}(x|1)$  is Gaussian with mean 5 and variance  $\sigma^2$  and  $f_{X|Y}(x|0)$  is Gaussian with mean zero and variance  $\sigma^2$

$$P[Y = 1|X = x] = f_{Y|X}(1|x) = \frac{f_{X|Y}(x|1) f_Y(1)}{f_X(x)}$$

- ▶ We need  $f_{X|Y}$ . What does our model say?
- $f_{X|Y}(x|1)$  is Gaussian with mean 5 and variance  $\sigma^2$  and  $f_{X|Y}(x|0)$  is Gaussian with mean zero and variance  $\sigma^2$

$$P[Y = 1|X = x] = f_{Y|X}(1|x) = \frac{f_{X|Y}(x|1) f_Y(1)}{f_X(x)}$$

▶ We need  $f_Y(1), f_Y(0)$ . Let us take them to be same.

- ▶ We need  $f_{X|Y}$ . What does our model say?
- $f_{X|Y}(x|1)$  is Gaussian with mean 5 and variance  $\sigma^2$  and  $f_{X|Y}(x|0)$  is Gaussian with mean zero and variance  $\sigma^2$

$$P[Y = 1|X = x] = f_{Y|X}(1|x) = \frac{f_{X|Y}(x|1) f_Y(1)}{f_X(x)}$$

- ▶ We need  $f_Y(1), f_Y(0)$ . Let us take them to be same.
- ▶ In practice we only want to know whether  $f_{Y|X}(1|x) > f_{Y|X}(0|x)$

- ▶ We need  $f_{X|Y}$ . What does our model say?
- $f_{X|Y}(x|1)$  is Gaussian with mean 5 and variance  $\sigma^2$  and  $f_{X|Y}(x|0)$  is Gaussian with mean zero and variance  $\sigma^2$

$$P[Y = 1|X = x] = f_{Y|X}(1|x) = \frac{f_{X|Y}(x|1) f_Y(1)}{f_X(x)}$$

- ▶ We need  $f_Y(1), f_Y(0)$ . Let us take them to be same.
- ▶ In practice we only want to know whether  $f_{Y|X}(1|x) > f_{Y|X}(0|x)$
- ▶ Then we do not need to calculate  $f_X(x)$ .

- ▶ We need  $f_{X|Y}$ . What does our model say?
- $f_{X|Y}(x|1)$  is Gaussian with mean 5 and variance  $\sigma^2$  and  $f_{X|Y}(x|0)$  is Gaussian with mean zero and variance  $\sigma^2$

$$P[Y = 1|X = x] = f_{Y|X}(1|x) = \frac{f_{X|Y}(x|1) f_Y(1)}{f_X(x)}$$

- ▶ We need  $f_Y(1), f_Y(0)$ . Let us take them to be same.
- ▶ In practice we only want to know whether  $f_{Y|X}(1|x) > f_{Y|X}(0|x)$
- ► Then we do not need to calculate  $f_X(x)$ . We only need ratio of  $f_{Y|X}(1|x)$  and  $f_{Y|X}(0|x)$ .

$$\frac{f_{Y|X}(1|x)}{f_{Y|X}(0|x)} = \frac{f_{X|Y}(x|1) f_{Y}(1)}{f_{X|Y}(x|0) f_{Y}(0)}$$

$$\frac{f_{Y|X}(1|x)}{f_{Y|X}(0|x)} = \frac{f_{X|Y}(x|1) f_{Y}(1)}{f_{X|Y}(x|0) f_{Y}(0)}$$
$$= \frac{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-5)^{2}}}{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-0)^{2}}}$$

$$\frac{f_{Y|X}(1|x)}{f_{Y|X}(0|x)} = \frac{f_{X|Y}(x|1) f_{Y}(1)}{f_{X|Y}(x|0) f_{Y}(0)}$$

$$= \frac{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-5)^{2}}}{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-0)^{2}}}$$

$$= e^{-0.5\sigma^{-2}(x^{2}-10x+25-x^{2})}$$

$$\frac{f_{Y|X}(1|x)}{f_{Y|X}(0|x)} = \frac{f_{X|Y}(x|1) f_{Y}(1)}{f_{X|Y}(x|0) f_{Y}(0)}$$

$$= \frac{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-5)^{2}}}{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-0)^{2}}}$$

$$= e^{-0.5\sigma^{-2}(x^{2}-10x+25-x^{2})}$$

$$= e^{0.5\sigma^{-2}(10x-25)}$$

$$\frac{f_{Y|X}(1|x)}{f_{Y|X}(0|x)} = \frac{f_{X|Y}(x|1) f_{Y}(1)}{f_{X|Y}(x|0) f_{Y}(0)}$$

$$= \frac{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-5)^{2}}}{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-0)^{2}}}$$

$$= e^{-0.5\sigma^{-2}(x^{2}-10x+25-x^{2})}$$

$$= e^{0.5\sigma^{-2}(10x-25)}$$

We are only interested in whether the above is greater than 1 or not.

$$\frac{f_{Y|X}(1|x)}{f_{Y|X}(0|x)} = \frac{f_{X|Y}(x|1) f_{Y}(1)}{f_{X|Y}(x|0) f_{Y}(0)}$$

$$= \frac{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-5)^{2}}}{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-0)^{2}}}$$

$$= e^{-0.5\sigma^{-2}(x^{2}-10x+25-x^{2})}$$

$$= e^{0.5\sigma^{-2}(10x-25)}$$

- ▶ We are only interested in whether the above is greater than 1 or not.
- ▶ The ratio is greater than 1 if 10x > 25 or x > 2.5

$$\frac{f_{Y|X}(1|x)}{f_{Y|X}(0|x)} = \frac{f_{X|Y}(x|1) f_{Y}(1)}{f_{X|Y}(x|0) f_{Y}(0)}$$

$$= \frac{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-5)^{2}}}{\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2\sigma^{2}}(x-0)^{2}}}$$

$$= e^{-0.5\sigma^{-2}(x^{2}-10x+25-x^{2})}$$

$$= e^{0.5\sigma^{-2}(10x-25)}$$

- ▶ We are only interested in whether the above is greater than 1 or not.
- ▶ The ratio is greater than 1 if 10x > 25 or x > 2.5
- ▶ So, if X > 2.5 we will conclude bit 1 is sent. Intuitively obvious!

▶ We did not calculate  $f_X(x)$  in the above.

- ▶ We did not calculate  $f_X(x)$  in the above.
- ▶ We can calculate it if we want.

- ▶ We did not calculate  $f_X(x)$  in the above.
- ▶ We can calculate it if we want.
- Using total probability rule

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

- ▶ We did not calculate  $f_X(x)$  in the above.
- ▶ We can calculate it if we want.
- Using total probability rule

$$f_X(x) = \sum_y f_{X|Y}(x|y) f_Y(y)$$
  
=  $f_{X|Y}(x|1) f_Y(1) + f_{X|Y}(x|0) f_Y(0)$ 

- ▶ We did not calculate  $f_X(x)$  in the above.
- ▶ We can calculate it if we want.
- Using total probability rule

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

$$= f_{X|Y}(x|1) f_Y(1) + f_{X|Y}(x|0) f_Y(0)$$

$$= \frac{1}{2} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-5)^2}{2\sigma^2}} + \frac{1}{2} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{x^2}{2\sigma^2}}$$

- ▶ We did not calculate  $f_X(x)$  in the above.
- ▶ We can calculate it if we want.
- Using total probability rule

$$f_X(x) = \sum_{y} f_{X|Y}(x|y) f_Y(y)$$

$$= f_{X|Y}(x|1) f_Y(1) + f_{X|Y}(x|0) f_Y(0)$$

$$= \frac{1}{2} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{(x-5)^2}{2\sigma^2}} + \frac{1}{2} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{x^2}{2\sigma^2}}$$

▶ It is a mixture density

As we saw, given the joint distribution we can calculate all the marginals.

- As we saw, given the joint distribution we can calculate all the marginals.
- ► However, there can be many joint distributions with the same marginals.

- As we saw, given the joint distribution we can calculate all the marginals.
- However, there can be many joint distributions with the same marginals.
- Let  $F_1, F_2$  be one dimensional df's of continuous rv's with  $f_1, f_2$  being the corresponding densities.

- As we saw, given the joint distribution we can calculate all the marginals.
- However, there can be many joint distributions with the same marginals.
- Let  $F_1, F_2$  be one dimensional df's of continuous rv's with  $f_1, f_2$  being the corresponding densities.

Define a function  $f: \Re^2 \to \Re$  by

$$f(x,y) = f_1(x)f_2(y) [1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1)]$$
  
where  $\alpha \in (-1,1)$ .

- As we saw, given the joint distribution we can calculate all the marginals.
- ► However, there can be many joint distributions with the same marginals.
- Let  $F_1, F_2$  be one dimensional df's of continuous rv's with  $f_1, f_2$  being the corresponding densities. Define a function  $f: \Re^2 \to \Re$  by

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$

where  $\alpha \in (-1, 1)$ .

▶ First note that f(x,y) > 0,  $\forall \alpha \in (-1,1)$ .

- As we saw, given the joint distribution we can calculate all the marginals.
- ► However, there can be many joint distributions with the same marginals.
- Let  $F_1, F_2$  be one dimensional df's of continuous rv's with  $f_1, f_2$  being the corresponding densities. Define a function  $f: \Re^2 \to \Re$  by

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$

where  $\alpha \in (-1,1)$ .

First note that  $f(x,y) \ge 0$ ,  $\forall \alpha \in (-1,1)$ . For different  $\alpha$  we get different functions.

- As we saw, given the joint distribution we can calculate all the marginals.
- ► However, there can be many joint distributions with the same marginals.
- Let  $F_1, F_2$  be one dimensional df's of continuous rv's with  $f_1, f_2$  being the corresponding densities. Define a function  $f: \Re^2 \to \Re$  by

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$
 where  $\alpha \in (-1,1)$ .

- First note that  $f(x,y) \ge 0$ ,  $\forall \alpha \in (-1,1)$ . For different  $\alpha$  we get different functions.
- $\blacktriangleright$  We first show that f(x,y) is a joint density.

- As we saw, given the joint distribution we can calculate all the marginals.
- ► However, there can be many joint distributions with the same marginals.
- Let  $F_1, F_2$  be one dimensional df's of continuous rv's with  $f_1, f_2$  being the corresponding densities. Define a function  $f: \Re^2 \to \Re$  by

$$f(x,y) = f_1(x) f_2(y) \left[ 1 + \alpha (2F_1(x) - 1)(2F_2(y) - 1) \right]$$

where  $\alpha \in (-1, 1)$ .

- First note that  $f(x,y) \ge 0$ ,  $\forall \alpha \in (-1,1)$ . For different  $\alpha$  we get different functions.
- ▶ We first show that f(x,y) is a joint density.
- ► For this, we note the following

$$\int_{-\infty}^{\infty} f_1(x) \ F_1(x) \ dx = \left. \frac{(F_1(x))^2}{2} \right|^{\infty} = \frac{1}{2}$$

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$

$$f(x,y) = f_1(x)f_2(y)\left[1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1)\right]$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \ dx \ dy =$$

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) \ dx \ dy = \int_{-\infty}^{\infty} f_1(x) \ dx \int_{-\infty}^{\infty} f_2(y) \ dy$$

$$f(x,y) = f_1(x)f_2(y)\left[1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1)\right]$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) \, dx \, dy = \int_{-\infty}^{\infty} f_1(x) \, dx \int_{-\infty}^{\infty} f_2(y) \, dy + \alpha \int_{-\infty}^{\infty} (2f_1(x)F_1(x) - f_1(x)) \, dx \int_{-\infty}^{\infty} (2f_2(y)F_2(y) - f_2(y)) \, dy$$

$$f(x,y) = f_1(x)f_2(y)\left[1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1)\right]$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) \, dx \, dy = \int_{-\infty}^{\infty} f_1(x) \, dx \int_{-\infty}^{\infty} f_2(y) \, dy + \alpha \int_{-\infty}^{\infty} (2f_1(x)F_1(x) - f_1(x)) \, dx \int_{-\infty}^{\infty} (2f_2(y)F_2(y) - f_2(y)) \, dy$$

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \, dx \, dy = \int_{-\infty}^{\infty} f_1(x) \, dx \int_{-\infty}^{\infty} f_2(y) \, dy + \alpha \int_{-\infty}^{\infty} (2f_1(x)F_1(x) - f_1(x)) \, dx \int_{-\infty}^{\infty} (2f_2(y)F_2(y) - f_2(y)) \, dy = 1$$

because  $2 \int_{-\infty}^{\infty} f_1(x) F_1(x) dx = 1$ .

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) \, dx \, dy = \int_{-\infty}^{\infty} f_1(x) \, dx \int_{-\infty}^{\infty} f_2(y) \, dy + \alpha \int_{-\infty}^{\infty} (2f_1(x)F_1(x) - f_1(x)) \, dx \int_{-\infty}^{\infty} (2f_2(y)F_2(y) - f_2(y)) \, dy = 1$$

because  $2 \int_{-\infty}^{\infty} f_1(x) F_1(x) dx = 1$ . This also shows

$$\int_{-\infty}^{\infty} f(x,y)dx = f_2(y);$$

$$f(x,y) = f_1(x)f_2(y) \left[ 1 + \alpha(2F_1(x) - 1)(2F_2(y) - 1) \right]$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \, dx \, dy = \int_{-\infty}^{\infty} f_1(x) \, dx \int_{-\infty}^{\infty} f_2(y) \, dy + \alpha \int_{-\infty}^{\infty} (2f_1(x)F_1(x) - f_1(x)) \, dx \int_{-\infty}^{\infty} (2f_2(y)F_2(y) - f_2(y)) \, dy = 1$$

because  $2 \int_{-\infty}^{\infty} f_1(x) F_1(x) dx = 1$ . This also shows

$$\int_{-\infty}^{\infty} f(x,y)dx = f_2(y); \quad \int_{-\infty}^{\infty} f(x,y)dy = f_1(x)$$



► Thus infinitely many joint distributions can all have the same marginals.

- ► Thus infinitely many joint distributions can all have the same marginals.
- ► So, in general, the marginals cannot determine the joint distribution.

- Thus infinitely many joint distributions can all have the same marginals.
- ► So, in general, the marginals cannot determine the joint distribution.
- ► An important special case where this is possible is that of independent random variables

▶ Two random variable X, Y are said to be independent if for all Borel sets  $B_1, B_2$ , the events  $[X \in B_1]$  and  $[Y \in B_2]$  are independent.

- ▶ Two random variable X, Y are said to be independent if for all Borel sets  $B_1, B_2$ , the events  $[X \in B_1]$  and  $[Y \in B_2]$  are independent.
- ▶ If *X*, *Y* are independent then

$$P[X \in B_1, Y \in B_2] = P[X \in B_1] P[Y \in B_2], \ \forall B_1, B_2 \in \mathcal{B}$$

- ▶ Two random variable X, Y are said to be independent if for all Borel sets  $B_1, B_2$ , the events  $[X \in B_1]$  and  $[Y \in B_2]$  are independent.
- ▶ If *X,Y* are independent then

$$P[X \in B_1, Y \in B_2] = P[X \in B_1] P[Y \in B_2], \ \forall B_1, B_2 \in \mathcal{B}$$

► In particular

$$F_{XY}(x,y) = P[X \le x, Y \le y]$$

- ▶ Two random variable X, Y are said to be independent if for all Borel sets  $B_1, B_2$ , the events  $[X \in B_1]$  and  $[Y \in B_2]$  are independent.
- ▶ If *X,Y* are independent then

$$P[X \in B_1, Y \in B_2] = P[X \in B_1] P[Y \in B_2], \ \forall B_1, B_2 \in \mathcal{B}$$

► In particular

$$F_{XY}(x,y) = P[X \le x, Y \le y] = P[X \le x]P[Y \le y]$$

- ▶ Two random variable X, Y are said to be independent if for all Borel sets  $B_1, B_2$ , the events  $[X \in B_1]$  and  $[Y \in B_2]$  are independent.
- ▶ If *X*, *Y* are independent then

$$P[X \in B_1, Y \in B_2] = P[X \in B_1] P[Y \in B_2], \ \forall B_1, B_2 \in \mathcal{B}$$

► In particular

$$F_{XY}(x,y) = P[X \leq x, Y \leq y] = P[X \leq x]P[Y \leq y] = F_X(x) \ F_Y(y)$$



- ▶ Two random variable X, Y are said to be independent if for all Borel sets  $B_1, B_2$ , the events  $[X \in B_1]$  and  $[Y \in B_2]$  are independent.
- ▶ If *X*, *Y* are independent then

$$P[X \in B_1, Y \in B_2] = P[X \in B_1] P[Y \in B_2], \ \forall B_1, B_2 \in \mathcal{B}$$

► In particular

$$F_{XY}(x,y) = P[X \le x, Y \le y] = P[X \le x]P[Y \le y] = F_X(x) \ F_Y(y)$$

▶ **Theorem**: X, Y are independent if and only if  $F_{XY}(x, y) = F_X(x)F_Y(y)$ .



 $\triangleright$  Suppose X, Y are independent discrete rv's

$$f_{XY}(x,y) = P[X = x, Y = y]$$

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y]$$

► Suppose *X,Y* are independent discrete rv's

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y] = f_X(x)f_Y(y)$$

The joint mass function is a product of marginals.

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y] = f_X(x)f_Y(y)$$

The joint mass function is a product of marginals.

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y] = f_X(x)f_Y(y)$$

The joint mass function is a product of marginals.

$$F_{XY}(x,y) = \sum_{x_i < x, y_i < y} f_{XY}(x_i, y_j)$$

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y] = f_X(x)f_Y(y)$$

The joint mass function is a product of marginals.

$$F_{XY}(x,y) = \sum_{x_i \le x, y_j \le y} f_{XY}(x_i, y_j) = \sum_{x_i \le x, y_j \le y} f_X(x_i) f_Y(y_j)$$

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y] = f_X(x)f_Y(y)$$

The joint mass function is a product of marginals.

$$F_{XY}(x,y) = \sum_{x_i \le x, y_j \le y} f_{XY}(x_i, y_j) = \sum_{x_i \le x, y_j \le y} f_X(x_i) f_Y(y_j)$$
$$= \sum_{x_i \le x} f_X(x_i) \sum_{y_j \le y} f_Y(y_j)$$

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y] = f_X(x)f_Y(y)$$

The joint mass function is a product of marginals.

$$F_{XY}(x,y) = \sum_{x_i \le x, y_j \le y} f_{XY}(x_i, y_j) = \sum_{x_i \le x, y_j \le y} f_X(x_i) f_Y(y_j)$$
$$= \sum_{x_i \le x} f_X(x_i) \sum_{y_i \le y} f_Y(y_j) = F_X(x) F_Y(y)$$

$$f_{XY}(x,y) = P[X = x, Y = y] = P[X = x]P[Y = y] = f_X(x)f_Y(y)$$

The joint mass function is a product of marginals.

▶ Suppose  $f_{XY}(x,y) = f_X(x)f_Y(y)$ . Then

$$F_{XY}(x,y) = \sum_{x_i \le x, y_j \le y} f_{XY}(x_i, y_j) = \sum_{x_i \le x, y_j \le y} f_X(x_i) f_Y(y_j)$$
$$= \sum_{x_i \le x} f_X(x_i) \sum_{y_j \le y} f_Y(y_j) = F_X(x) F_Y(y)$$

So, X, Y are independent if and only if  $f_{XY}(x, y) = f_X(x) f_Y(y)$ 

Let X, Y be independent continuous rv

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') dx' \int_{-\infty}^y f_Y(y') dy'$$

 $\blacktriangleright$  Let X,Y be independent continuous rv

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') \, dx' \int_{-\infty}^y f_Y(y') \, dy'$$
$$= \int_{-\infty}^y \int_{-\infty}^x (f_X(x')f_Y(y')) \, dx' \, dy'$$

 $\blacktriangleright$  Let X,Y be independent continuous rv

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') \, dx' \int_{-\infty}^y f_Y(y') \, dy'$$
$$= \int_{-\infty}^y \int_{-\infty}^x (f_X(x')f_Y(y')) \, dx' \, dy'$$

This implies joint density is product of marginals.

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') \, dx' \int_{-\infty}^y f_Y(y') \, dy'$$
$$= \int_{-\infty}^y \int_{-\infty}^x (f_X(x')f_Y(y')) \, dx' \, dy'$$

- This implies joint density is product of marginals.
- Now, suppose  $f_{XY}(x,y) = f_X(x)f_Y(y)$

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') \, dx' \int_{-\infty}^y f_Y(y') \, dy'$$
$$= \int_{-\infty}^y \int_{-\infty}^x (f_X(x')f_Y(y')) \, dx' \, dy'$$

- This implies joint density is product of marginals.
- Now, suppose  $f_{XY}(x,y) = f_X(x)f_Y(y)$

$$F_{XY}(x,y) = \int_{-\infty}^{y} \int_{-\infty}^{x} f_{XY}(x',y') dx' dy'$$

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') \, dx' \int_{-\infty}^y f_Y(y') \, dy'$$
$$= \int_{-\infty}^y \int_{-\infty}^x (f_X(x')f_Y(y')) \, dx' \, dy'$$

- This implies joint density is product of marginals.
- Now, suppose  $f_{XY}(x,y) = f_X(x)f_Y(y)$

$$F_{XY}(x,y) = \int_{-\infty}^{y} \int_{-\infty}^{x} f_{XY}(x',y') \, dx' \, dy'$$
$$= \int_{-\infty}^{y} \int_{-\infty}^{x} f_{X}(x') f_{Y}(y') \, dx' \, dy'$$

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') \, dx' \int_{-\infty}^y f_Y(y') \, dy'$$
$$= \int_{-\infty}^y \int_{-\infty}^x (f_X(x')f_Y(y')) \, dx' \, dy'$$

- ▶ This implies joint density is product of marginals.
- Now, suppose  $f_{XY}(x,y) = f_X(x)f_Y(y)$

$$F_{XY}(x,y) = \int_{-\infty}^{y} \int_{-\infty}^{x} f_{XY}(x',y') dx' dy'$$

$$= \int_{-\infty}^{y} \int_{-\infty}^{x} f_{X}(x') f_{Y}(y') dx' dy'$$

$$= \int_{-\infty}^{x} f_{X}(x') dx' \int_{-\infty}^{y} f_{Y}(y') dy' = F_{X}(x) F_{Y}(y)$$

 $\blacktriangleright$  Let X,Y be independent continuous rv

$$F_{XY}(x,y) = F_X(x)F_Y(y) = \int_{-\infty}^x f_X(x') \, dx' \int_{-\infty}^y f_Y(y') \, dy'$$
$$= \int_{-\infty}^y \int_{-\infty}^x (f_X(x')f_Y(y')) \, dx' \, dy'$$

- ► This implies joint density is product of marginals.
- Now, suppose  $f_{XY}(x,y) = f_X(x)f_Y(y)$

$$F_{XY}(x,y) = \int_{-\infty}^{y} \int_{-\infty}^{x} f_{XY}(x',y') \, dx' \, dy'$$

$$= \int_{-\infty}^{y} \int_{-\infty}^{x} f_{X}(x') f_{Y}(y') \, dx' \, dy'$$

$$= \int_{-\infty}^{x} f_{X}(x') \, dx' \int_{-\infty}^{y} f_{Y}(y') \, dy' = F_{X}(x) F_{Y}(y)$$

So, X, Y are independent if and only if  $f_{XY}(x, y) = f_X(x) f_Y(y)$ 

ightharpoonup Let X, Y be independent.

- Let *X*, *Y* be independent.
- ▶ Then  $P[X \in B_1 | Y \in B_2] = P[X \in B_1]$ .

- ightharpoonup Let X, Y be independent.
- ▶ Then  $P[X \in B_1 | Y \in B_2] = P[X \in B_1]$ .
- ▶ Hence, we get  $F_{X|Y}(x|y) = F_X(x)$ .

- Let *X*, *Y* be independent.
- ▶ Then  $P[X \in B_1 | Y \in B_2] = P[X \in B_1]$ .
- ▶ Hence, we get  $F_{X|Y}(x|y) = F_X(x)$ .
- ▶ This also implies  $f_{X|Y}(x|y) = f_X(x)$ .

- Let *X*, *Y* be independent.
- ▶ Then  $P[X \in B_1 | Y \in B_2] = P[X \in B_1]$ .
- ▶ Hence, we get  $F_{X|Y}(x|y) = F_X(x)$ .
- ▶ This also implies  $f_{X|Y}(x|y) = f_X(x)$ .
- ightharpoonup This is true for all the four possibilities of X,Y being continuous/discrete.

Everything we have done so far is easily extended to multiple random variables.

- Everything we have done so far is easily extended to multiple random variables.
- ightharpoonup Let X, Y, Z be rv on the same probability space.

- Everything we have done so far is easily extended to multiple random variables.
- $\blacktriangleright$  Let X, Y, Z be rv on the same probability space.
- We define joint distribution function by

$$F_{XYZ}(x, y, z) = P[X \le x, Y \le y, Z \le z]$$

- Everything we have done so far is easily extended to multiple random variables.
- ightharpoonup Let X, Y, Z be rv on the same probability space.
- ▶ We define joint distribution function by

$$F_{XYZ}(x, y, z) = P[X \le x, Y \le y, Z \le z]$$

▶ If all three are discrete then the joint mass function is

$$f_{XYZ}(x, y, z) = P[X = x, Y = y, Z = z]$$



- Everything we have done so far is easily extended to multiple random variables.
- ightharpoonup Let X, Y, Z be rv on the same probability space.
- We define joint distribution function by

$$F_{XYZ}(x, y, z) = P[X \le x, Y \le y, Z \le z]$$

▶ If all three are discrete then the joint mass function is

$$f_{XYZ}(x, y, z) = P[X = x, Y = y, Z = z]$$

If they are continuous, they have a joint density if

$$F_{XYZ}(x, y, z) = \int_{-\infty}^{z} \int_{-\infty}^{y} \int_{-\infty}^{x} f_{XYZ}(x', y', z') dx' dy' dz'$$



► Easy to see that joint mass function satisfies

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$
- ► Similarly the joint density satisfies

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$
- ► Similarly the joint density satisfies
  - 1.  $f_{XYZ}(x, y, z) \ge 0$

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$
- ► Similarly the joint density satisfies
  - 1.  $f_{XYZ}(x, y, z) \ge 0$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dx dy dz = 1$

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$
- ► Similarly the joint density satisfies
  - 1.  $f_{XYZ}(x, y, z) \ge 0$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dx dy dz = 1$
- ► These are straight-forward generalizations

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$
- ► Similarly the joint density satisfies
  - 1.  $f_{XYZ}(x, y, z) \ge 0$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dx dy dz = 1$
- ▶ These are straight-forward generalizations
- ► The properties of joint distribution function such as it being non-decreasing in each argument etc are easily seen to hold here too.

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$
- ► Similarly the joint density satisfies
  - 1.  $f_{XYZ}(x, y, z) \ge 0$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dx dy dz = 1$
- ▶ These are straight-forward generalizations
- ► The properties of joint distribution function such as it being non-decreasing in each argument etc are easily seen to hold here too.
- ► Generalizing the special property of the df (relating to probability of cylindrical sets) is a little more complicated. (An exercise for you!)

- ► Easy to see that joint mass function satisfies
  - 1.  $f_{XYZ}(x,y,z) \ge 0$  and is non-zero only for countably many tuples.
  - 2.  $\sum_{x,y,z} f_{XYZ}(x,y,z) = 1$
- ► Similarly the joint density satisfies
  - 1.  $f_{XYZ}(x, y, z) \ge 0$
  - 2.  $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dx dy dz = 1$
- ► These are straight-forward generalizations
- ► The properties of joint distribution function such as it being non-decreasing in each argument etc are easily seen to hold here too.
- Generalizing the special property of the df (relating to probability of cylindrical sets) is a little more complicated. (An exercise for you!)
- ► We specify multiple random variables either through joint mass function or joint density function.

$$F_{XY}(x,y) = F_{XYZ}(x,y,\infty);$$

$$F_{XY}(x,y) = F_{XYZ}(x,y,\infty); \ F_{Z}(z) = F_{XYZ}(\infty,\infty,z)$$
 and so on

$$F_{XY}(x,y) = F_{XYZ}(x,y,\infty); \ \ F_Z(z) = F_{XYZ}(\infty,\infty,z) \quad \text{and so on}$$

► Similarly we get

$$f_{YZ}(y,z) = \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dx;$$
  
$$f_{X}(x) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dy dz$$

$$F_{XY}(x,y) = F_{XYZ}(x,y,\infty); \ \ F_Z(z) = F_{XYZ}(\infty,\infty,z) \quad \text{and so on}$$

► Similarly we get

$$f_{YZ}(y,z) = \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dx;$$
  
$$f_{X}(x) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dy dz$$

Any marginal is a joint density of a subset of these rv's and we obtain it by integrating the (full) joint density with respect to the remaining variables.

$$F_{XY}(x,y) = F_{XYZ}(x,y,\infty); \ \ F_Z(z) = F_{XYZ}(\infty,\infty,z) \quad \text{and so on}$$

► Similarly we get

$$f_{YZ}(y,z) = \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dx;$$
  
$$f_{X}(x) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dy dz$$

- ► Any marginal is a joint density of a subset of these rv's and we obtain it by integrating the (full) joint density with respect to the remaining variables.
- We obtain the marginal mass functions for a subset of the rv's also similarly where we sum over the remaining variables.

▶ We have to be a little careful in dealing with these when some random variables are discrete and others are continuous.

- ▶ We have to be a little careful in dealing with these when some random variables are discrete and others are continuous.
- ► Suppose *X* is continuous and *Y*, *Z* are discrete. We do not have any joint density or mass function as such.

- ▶ We have to be a little careful in dealing with these when some random variables are discrete and others are continuous.
- ► Suppose *X* is continuous and *Y*, *Z* are discrete. We do not have any joint density or mass function as such.
- ► However, the joint df is always well defined.

- ▶ We have to be a little careful in dealing with these when some random variables are discrete and others are continuous.
- ► Suppose *X* is continuous and *Y*, *Z* are discrete. We do not have any joint density or mass function as such.
- ► However, the joint df is always well defined.
- ▶ Suppose we want marginal joint distribution of X, Y. We know how to get  $F_{XY}$  by marginalization.

- ▶ We have to be a little careful in dealing with these when some random variables are discrete and others are continuous.
- ► Suppose *X* is continuous and *Y*, *Z* are discrete. We do not have any joint density or mass function as such.
- ► However, the joint df is always well defined.
- ▶ Suppose we want marginal joint distribution of X, Y. We know how to get  $F_{XY}$  by marginalization.
- ▶ Then we can get  $f_X$  (a density),  $f_Y$  (a mass fn),  $f_{X|Y}$  (conditional density) and  $f_{Y|X}$  (conditional mass fn)

- ▶ We have to be a little careful in dealing with these when some random variables are discrete and others are continuous.
- ► Suppose *X* is continuous and *Y*, *Z* are discrete. We do not have any joint density or mass function as such.
- ► However, the joint df is always well defined.
- ▶ Suppose we want marginal joint distribution of X, Y. We know how to get  $F_{XY}$  by marginalization.
- ▶ Then we can get  $f_X$  (a density),  $f_Y$  (a mass fn),  $f_{X|Y}$  (conditional density) and  $f_{Y|X}$  (conditional mass fn)
- With these we can generally calculate most quantities of interest.

Like in case of marginals, there are different types of conditional distributions now.

- Like in case of marginals, there are different types of conditional distributions now.
- We can always define conditional distribution functions like

- Like in case of marginals, there are different types of conditional distributions now.
- We can always define conditional distribution functions like

$$F_{XY|Z}(x,y|z) = P[X \le x, Y \le y|Z=z]$$

- Like in case of marginals, there are different types of conditional distributions now.
- We can always define conditional distribution functions like

$$F_{XY|Z}(x,y|z) = P[X \le x, Y \le y|Z = z]$$
  
$$F_{X|YZ}(x|y,z) = P[X \le x|Y = y, Z = z]$$

- Like in case of marginals, there are different types of conditional distributions now.
- We can always define conditional distribution functions like

$$\begin{aligned} F_{XY|Z}(x,y|z) &=& P[X \leq x, Y \leq y | Z = z] \\ F_{X|YZ}(x|y,z) &=& P[X \leq x | Y = y, Z = z] \end{aligned}$$

► In all such cases, if the conditioning random variables are continuous, we define the above as a limit.

- Like in case of marginals, there are different types of conditional distributions now.
- We can always define conditional distribution functions like

$$F_{XY|Z}(x,y|z) = P[X \le x, Y \le y|Z = z]$$
  
$$F_{X|YZ}(x|y,z) = P[X \le x|Y = y, Z = z]$$

- ► In all such cases, if the conditioning random variables are continuous, we define the above as a limit.
- $\triangleright$  For example when Z is continuous

$$F_{XY|Z}(x,y|z) = \lim_{\delta \to 0} P[X \le x, Y \le y | Z \in [z,z+\delta]]$$

$$f_{X|YZ}(x|y,z) = P[X = x|Y = y, Z = z]$$

$$f_{X|YZ}(x|y,z) = P[X = x|Y = y, Z = z]$$

$$f_{X|YZ}(x|y,z) = P[X = x|Y = y, Z = z]$$

$$f_{XY|Z}(x,y|z) = \frac{f_{XYZ}(x,y,z)}{f_{Z}(z)}$$

$$f_{X|YZ}(x|y,z) = P[X = x|Y = y, Z = z]$$

$$f_{XY|Z}(x,y|z) = \frac{f_{XYZ}(x,y,z)}{f_Z(z)}$$
$$f_{X|YZ}(x|y,z) = \frac{f_{XYZ}(x,y,z)}{f_{YZ}(y,z)}$$

$$f_{X|YZ}(x|y,z) = P[X = x|Y = y, Z = z]$$

$$\begin{array}{lcl} f_{XY|Z}(x,y|z) & = & \frac{f_{XYZ}(x,y,z)}{f_{Z}(z)} \\ \\ f_{X|YZ}(x|y,z) & = & \frac{f_{XYZ}(x,y,z)}{f_{YZ}(y,z)} \\ \\ f_{XYZ}(x,y,z) & = & f_{Z|YX}(z|y,x)f_{Y|X}(y|x)f_{X}(x) \end{array}$$

$$f_{X|YZ}(x|y,z) = P[X = x|Y = y, Z = z]$$

► Thus the following are obvious

$$f_{XY|Z}(x, y|z) = \frac{f_{XYZ}(x, y, z)}{f_{Z}(z)}$$

$$f_{X|YZ}(x|y, z) = \frac{f_{XYZ}(x, y, z)}{f_{YZ}(y, z)}$$

$$f_{XYZ}(x, y, z) = f_{Z|YX}(z|y, x)f_{Y|X}(y|x)f_{X}(x)$$

▶ For example, the first one above follows from

$$P[X = x, Y = y | Z = z] = \frac{P[X = x, Y = y, Z = z]}{P[Z = z]}$$



$$F_{Z|XY}(z|x,y) = \lim_{\delta \to 0} \frac{P[Z \le z, X \in [x, x + \delta], Y \in [y, y + \delta]]}{P[X \in [x, x + \delta, Y \in [y, y + \delta]]}$$

$$F_{Z|XY}(z|x,y) = \lim_{\delta \to 0} \frac{P[Z \le z, X \in [x, x + \delta], Y \in [y, y + \delta]]}{P[X \in [x, x + \delta, Y \in [y, y + \delta]]}$$
$$= \lim_{\delta \to 0} \frac{\int_{-\infty}^{z} \int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XYZ}(x', y', z') \, dy' \, dx' \, dz'}{\int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XY}(x', y') \, dy' \, dx'}$$

$$F_{Z|XY}(z|x,y) = \lim_{\delta \to 0} \frac{P[Z \le z, X \in [x, x + \delta], Y \in [y, y + \delta]]}{P[X \in [x, x + \delta, Y \in [y, y + \delta]]}$$

$$= \lim_{\delta \to 0} \frac{\int_{-\infty}^{z} \int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XYZ}(x', y', z') \, dy' \, dx' \, dz'}{\int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XY}(x', y') \, dy' \, dx'}$$

$$= \int_{-\infty}^{z} \frac{f_{XYZ}(x, y, z')}{f_{XY}(x, y)} \, dz'$$

$$F_{Z|XY}(z|x,y) = \lim_{\delta \to 0} \frac{P[Z \le z, X \in [x, x + \delta], Y \in [y, y + \delta]]}{P[X \in [x, x + \delta, Y \in [y, y + \delta]]}$$

$$= \lim_{\delta \to 0} \frac{\int_{-\infty}^{z} \int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XYZ}(x', y', z') \ dy' \ dx' \ dz'}{\int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XY}(x', y') \ dy' \ dx'}$$

$$= \int_{-\infty}^{z} \frac{f_{XYZ}(x, y, z')}{f_{XY}(x, y)} \ dz' = \int_{-\infty}^{z} f_{Z|XY}(z'|x, y) \ dz'$$

$$\begin{split} F_{Z|XY}(z|x,y) &= \lim_{\delta \to 0} \frac{P[Z \leq z, X \in [x, x + \delta], Y \in [y, y + \delta]]}{P[X \in [x, x + \delta, Y \in [y, y + \delta]]} \\ &= \lim_{\delta \to 0} \frac{\int_{-\infty}^{z} \int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XYZ}(x', y', z') \; dy' \; dx' \; dz'}{\int_{x}^{x + \delta} \int_{y}^{y + \delta} f_{XY}(x', y') \; dy' \; dx'} \\ &= \int_{-\infty}^{z} \frac{f_{XYZ}(x, y, z')}{f_{XY}(x, y)} \; dz' = \int_{-\infty}^{z} f_{Z|XY}(z'|x, y) \; dz' \end{split}$$

► Thus we get

$$f_{XYZ}(x, y, z) = f_{Z|XY}(z|x, y)f_{XY}(x, y) = f_{Z|XY}(z|x, y)f_{Y|X}(y|x)f_{X}(x)$$

We can similarly talk about the joint distribution of any finite number of rv's

- We can similarly talk about the joint distribution of any finite number of rv's
- Let  $X_1, X_2, \dots, X_n$  be rv's on the same probability space.

- We can similarly talk about the joint distribution of any finite number of ry's
- Let  $X_1, X_2, \dots, X_n$  be rv's on the same probability space.
- ▶ We denote it as a vector  $\mathbf{X}$  or  $\underline{X}$ . We can think of it as a mapping,  $\mathbf{X}: \Omega \to \Re^n$ .

- We can similarly talk about the joint distribution of any finite number of rv's
- Let  $X_1, X_2, \dots, X_n$  be rv's on the same probability space.
- ▶ We denote it as a vector  $\mathbf{X}$  or  $\underline{X}$ . We can think of it as a mapping,  $\mathbf{X}: \Omega \to \Re^n$ .
- ▶ We can write the joint distribution as

$$F_{\mathbf{X}}(\mathbf{x}) = P[\mathbf{X} \le \mathbf{x}] = P[X_i \le x_i, i = 1, \dots, n]$$

- We can similarly talk about the joint distribution of any finite number of rv's
- Let  $X_1, X_2, \dots, X_n$  be rv's on the same probability space.
- ▶ We denote it as a vector  $\mathbf{X}$  or  $\underline{X}$ . We can think of it as a mapping,  $\mathbf{X}: \Omega \to \Re^n$ .
- ▶ We can write the joint distribution as

$$F_{\mathbf{X}}(\mathbf{x}) = P[\mathbf{X} \le \mathbf{x}] = P[X_i \le x_i, i = 1, \dots, n]$$

▶ We represent by  $f_{\mathbf{X}}(\mathbf{x})$  the joint density or mass function. Sometimes we also write it as  $f_{X_1 \cdots X_n}(x_1, \cdots, x_n)$ 

- We can similarly talk about the joint distribution of any finite number of rv's
- Let  $X_1, X_2, \dots, X_n$  be rv's on the same probability space.
- ▶ We denote it as a vector  $\mathbf{X}$  or  $\underline{X}$ . We can think of it as a mapping,  $\mathbf{X}: \Omega \to \Re^n$ .
- ▶ We can write the joint distribution as

$$F_{\mathbf{X}}(\mathbf{x}) = P[\mathbf{X} \le \mathbf{x}] = P[X_i \le x_i, i = 1, \dots, n]$$

- ▶ We represent by  $f_{\mathbf{X}}(\mathbf{x})$  the joint density or mass function. Sometimes we also write it as  $f_{X_1 \cdots X_n}(x_1, \cdots, x_n)$
- We use similar notation for marginal and conditional distributions

▶ Random variables  $X_1, X_2, \cdots, X_n$  are said to be independent if the the events  $[X_i \in B_i], i = 1, \cdots, n$  are independent.

Random variables X<sub>1</sub>, X<sub>2</sub>, · · · , X<sub>n</sub> are said to be independent if the the events [X<sub>i</sub> ∈ B<sub>i</sub>], i = 1, · · · , n are independent.
(Recall definition of independence of a set of events)

- Random variables X<sub>1</sub>, X<sub>2</sub>, · · · , X<sub>n</sub> are said to be independent if the the events [X<sub>i</sub> ∈ B<sub>i</sub>], i = 1, · · · , n are independent.
  (Recall definition of independence of a set of events)
- ▶ Independence implies that the marginals would determine the joint distribution.

- ▶ Random variables  $X_1, X_2, \dots, X_n$  are said to be independent if the the events  $[X_i \in B_i], i = 1, \dots, n$  are independent.
  - (Recall definition of independence of a set of events)
- ► Independence implies that the marginals would determine the joint distribution.
- ▶ If X, Y, Z are independent then  $f_{XYZ}(x, y, z) = f_X(x) f_Y(y) f_Z(z)$

- Random variables X<sub>1</sub>, X<sub>2</sub>, · · · , X<sub>n</sub> are said to be independent if the the events [X<sub>i</sub> ∈ B<sub>i</sub>], i = 1, · · · , n are independent.
  (Recall definition of independence of a set of events)
- ► Independence implies that the marginals would determine the joint distribution.
- ► If X, Y, Z are independent then  $f_{XYZ}(x, y, z) = f_X(x) f_Y(y) f_Z(z)$
- ► For independent random variables, the joint mass function (or density function) is product of individual mass functions (or density functions)

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) \ dz \ dy \ dx =$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) \ dz \ dy \ dx = \int_{0}^{1}$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) \ dz \ dy \ dx = \int_{0}^{1} \int_{0}^{x}$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) \ dz \ dy \ dx = \int_{0}^{1} \int_{0}^{x} \int_{0}^{y} K \ dz \ dy \ dx$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dz dy dx = \int_{0}^{1} \int_{0}^{x} \int_{0}^{y} K dz dy dx$$
$$= K \int_{0}^{1} \int_{0}^{x} y dy dx$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dz dy dx = \int_{0}^{1} \int_{0}^{x} \int_{0}^{y} K dz dy dx$$
$$= K \int_{0}^{1} \int_{0}^{x} y dy dx$$
$$= K \int_{0}^{1} \frac{x^{2}}{2} dx$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dz dy dx = \int_{0}^{1} \int_{0}^{x} \int_{0}^{y} K dz dy dx$$
$$= K \int_{0}^{1} \int_{0}^{x} y dy dx$$
$$= K \int_{0}^{1} \frac{x^{2}}{2} dx$$
$$= K \frac{1}{6}$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dz dy dx = \int_{0}^{1} \int_{0}^{x} \int_{0}^{y} K dz dy dx$$
$$= K \int_{0}^{1} \int_{0}^{x} y dy dx$$
$$= K \int_{0}^{1} \frac{x^{2}}{2} dx$$
$$= K \frac{1}{6} \Rightarrow K = 6$$

Let a joint density be given by

$$f_{XYZ}(x, y, z) = K, \quad 0 < z < y < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XYZ}(x, y, z) dz dy dx = \int_{x=0}^{1} \int_{y=0}^{x} \int_{z=0}^{y} K dz dy dx$$
$$= K \int_{x=0}^{1} \int_{y=0}^{x} y dy dx$$
$$= K \int_{0}^{1} \frac{x^{2}}{2} dx$$
$$= K \frac{1}{6} \Rightarrow K = 6$$

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

Suppose we want to find the (marginal) joint distribution of X and Z.

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

Suppose we want to find the (marginal) joint distribution of X and Z.

$$f_{XZ}(x,z) = \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dy$$

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

Suppose we want to find the (marginal) joint distribution of X and Z.

$$f_{XZ}(x,z) = \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) dy$$
$$= \int_{z}^{x} 6 dy, \quad 0 < z < x < 1$$

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

▶ Suppose we want to find the (marginal) joint distribution of X and Z.

$$f_{XZ}(x,z) = \int_{-\infty}^{\infty} f_{XYZ}(x,y,z) \, dy$$
$$= \int_{z}^{x} 6 \, dy, \quad 0 < z < x < 1$$
$$= 6(x-z), \quad 0 < z < x < 1$$

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XZ}(x,z) \ dz \ dx = \int_{0}^{1} \int_{0}^{x} 6(x-z) \ dz \ dx$$

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XZ}(x,z) dz dx = \int_{0}^{1} \int_{0}^{x} 6(x-z) dz dx$$
$$= \int_{0}^{1} \left( 6x z \Big|_{0}^{x} - 6 \frac{z^{2}}{2} \Big|_{0}^{x} \right) dx$$

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XZ}(x,z) \, dz \, dx = \int_{0}^{1} \int_{0}^{x} 6(x-z) \, dz \, dx$$
$$= \int_{0}^{1} \left( 6x \, z \big|_{0}^{x} - 6 \, \frac{z^{2}}{2} \, \Big|_{0}^{x} \right) \, dx$$
$$= \int_{0}^{1} \left( 6x^{2} - 6 \, \frac{x^{2}}{2} \, \right) \, dx$$

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XZ}(x,z) \, dz \, dx = \int_{0}^{1} \int_{0}^{x} 6(x-z) \, dz \, dx$$

$$= \int_{0}^{1} \left( 6x \, z \big|_{0}^{x} - 6 \, \frac{z^{2}}{2} \Big|_{0}^{x} \right) \, dx$$

$$= \int_{0}^{1} \left( 6x^{2} - 6 \, \frac{x^{2}}{2} \right) \, dx$$

$$= 3 \, \frac{x^{3}}{3} \Big|_{0}^{1} = 1$$

ightharpoonup The joint density of X, Y, Z is

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

▶ The joint density of X, Y, Z is

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

▶ The joint density of X, Z is

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

▶ The joint density of X, Y, Z is

$$f_{XYZ}(x, y, z) = 6, \quad 0 < z < y < x < 1$$

ightharpoonup The joint density of X, Z is

$$f_{XZ}(x,z) = 6(x-z), \quad 0 < z < x < 1$$

► Hence,

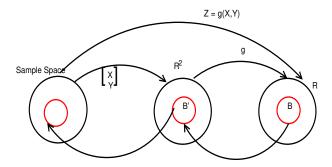
$$f_{Y|XZ}(y|x,z) = \frac{f_{XYZ}(x,y,z)}{f_{XZ}(x,z)} = \frac{1}{x-z}, \quad 0 < z < y < x < 1$$

► Let *X,Y* be random variables on the same probability space.

- ► Let *X,Y* be random variables on the same probability space.
- ▶ Let  $g: \Re^2 \to \Re$ .

- ► Let *X,Y* be random variables on the same probability space.
- ightharpoonup Let  $q: \Re^2 \to \Re$ .
- ▶ Let Z = g(X, Y). Then Z is a rv

- ► Let *X,Y* be random variables on the same probability space.
- ightharpoonup Let  $q: \Re^2 \to \Re$ .
- ▶ Let Z = g(X, Y). Then Z is a rv
- ► This is analogous to functions of a single rv



 $\blacktriangleright \text{ let } Z = g(X,Y)$ 

- $\blacktriangleright \text{ let } Z = g(X,Y)$
- ightharpoonup We can determine distribution of Z from the joint distribution of X,Y

- $\blacktriangleright \text{ let } Z = g(X,Y)$
- lackbox We can determine distribution of Z from the joint distribution of X,Y

$$F_Z(z) = P[Z \le z] = P[g(X, Y) \le z]$$

- $\blacktriangleright \text{ let } Z = g(X,Y)$
- lackbox We can determine distribution of Z from the joint distribution of X,Y

$$F_Z(z) = P[Z \le z] = P[g(X, Y) \le z]$$

 $\blacktriangleright$  For example, if X, Y are discrete, then

$$f_Z(z) = P[Z = z] = P[g(X, Y) = z] = \sum_{\substack{x_i, y_j: \\ g(x_i, y_j) = z}} f_{XY}(x_i, y_j)$$

$$f_Z(z) = P[\min(X, Y) = z]$$

$$f_Z(z) = P[\min(X, Y) = z]$$
  
=  $P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$ 

$$f_{Z}(z) = P[\min(X, Y) = z]$$

$$= P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= \sum_{y>z} P[X = z, Y = y] + \sum_{x>z} P[X = x, Y = z]$$

$$+P[X = z, Y = z]$$

$$f_{Z}(z) = P[\min(X, Y) = z]$$

$$= P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= \sum_{y>z} P[X = z, Y = y] + \sum_{x>z} P[X = x, Y = z]$$

$$+ P[X = z, Y = z]$$

$$= \sum_{y>z} f_{XY}(z, y) + \sum_{x>z} f_{XY}(x, z) + f_{XY}(z, z)$$

$$f_{Z}(z) = P[\min(X, Y) = z]$$

$$= P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= \sum_{y>z} P[X = z, Y = y] + \sum_{x>z} P[X = x, Y = z]$$

$$+ P[X = z, Y = z]$$

$$= \sum_{y>z} f_{XY}(z, y) + \sum_{x>z} f_{XY}(x, z) + f_{XY}(z, z)$$

Now suppose X, Y are independent and both of them have geometric distribution with the same parameter, p.

$$f_{Z}(z) = P[\min(X, Y) = z]$$

$$= P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= \sum_{y>z} P[X = z, Y = y] + \sum_{x>z} P[X = x, Y = z]$$

$$+ P[X = z, Y = z]$$

$$= \sum_{y>z} f_{XY}(z, y) + \sum_{x>z} f_{XY}(x, z) + f_{XY}(z, z)$$

- Now suppose X, Y are independent and both of them have geometric distribution with the same parameter, p.
- Such random variables are called independent and identically distributed or iid random variables.

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$
  
=  $P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$ 

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$
  
=  $P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$   
=  $p(1-p)^{z-1}$ 

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z * 2 +$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z * 2 + (p(1-p)^{z-1})^2$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z * 2 + (p(1-p)^{z-1})^2$$

$$= 2p(1-p)^{z-1}(1-p)^z + (p(1-p)^{z-1})^2$$

$$f_Z(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^z * 2 + (p(1-p)^{z-1})^2$$

$$= 2p(1-p)^{z-1}(1-p)^z + (p(1-p)^{z-1})^2$$

$$= 2p(1-p)^{2z-1} + p^2(1-p)^{2z-2}$$

$$f_{Z}(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^{z} * 2 + (p(1-p)^{z-1})^{2}$$

$$= 2p(1-p)^{z-1}(1-p)^{z} + (p(1-p)^{z-1})^{2}$$

$$= 2p(1-p)^{2z-1} + p^{2}(1-p)^{2z-2}$$

$$= p(1-p)^{2z-2}(2(1-p) + p)$$

$$f_{Z}(z) = P[X = z, Y > z] + P[Y = z, X > z] + P[X = Y = z]$$

$$= P[X = z]P[Y > z] + P[Y = z]P[X > z] + P[X = z]P[Y = z]$$

$$= p(1-p)^{z-1}(1-p)^{z} * 2 + (p(1-p)^{z-1})^{2}$$

$$= 2p(1-p)^{z-1}(1-p)^{z} + (p(1-p)^{z-1})^{2}$$

$$= 2p(1-p)^{2z-1} + p^{2}(1-p)^{2z-2}$$

$$= p(1-p)^{2z-2}(2(1-p) + p)$$

 $= (2-p)p(1-p)^{2z-2}$ 

$$\sum_{z=0}^{\infty} f_{z}(z) = \sum_{z=0}^{\infty} (2-p)p(1-p)^{2z-2}$$

$$\sum_{z=1}^{\infty} f_Z(z) = \sum_{z=1}^{\infty} (2-p)p(1-p)^{2z-2}$$
$$= (2-p)p \sum_{z=1}^{\infty} (1-p)^{2z-2}$$

$$\sum_{z=1}^{\infty} f_Z(z) = \sum_{z=1}^{\infty} (2-p)p(1-p)^{2z-2}$$
$$= (2-p)p \sum_{z=1}^{\infty} (1-p)^{2z-2}$$
$$= (2-p)p \frac{1}{1-(1-p)^2}$$

$$\sum_{z=1}^{\infty} f_Z(z) = \sum_{z=1}^{\infty} (2-p)p(1-p)^{2z-2}$$

$$= (2-p)p \sum_{z=1}^{\infty} (1-p)^{2z-2}$$

$$= (2-p)p \frac{1}{1-(1-p)^2}$$

$$= (2-p)p \frac{1}{2p-p^2} = 1$$

Let us consider the max and min functions, in general.

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$F_Z(z) = P[Z \le z]$$

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$F_Z(z) = P[Z \le z] = P[\max(X, Y) \le z]$$

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$F_Z(z)$$
 =  $P[Z \le z] = P[\max(X, Y) \le z]$   
 =  $P[X \le z, Y \le z]$ 

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$F_Z(z) = P[Z \le z] = P[\max(X, Y) \le z]$$

$$= P[X \le z, Y \le z]$$

$$= F_{XY}(z, z)$$

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$F_Z(z) = P[Z \le z] = P[\max(X, Y) \le z]$$

$$= P[X \le z, Y \le z]$$

$$= F_{XY}(z, z)$$

$$= F_X(z)F_Y(z), \text{ if } X, Y \text{ are independent}$$

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$\begin{split} F_Z(z) &= P[Z \leq z] = P[\max(X,Y) \leq z] \\ &= P[X \leq z, Y \leq z] \\ &= F_{XY}(z,z) \\ &= F_X(z)F_Y(z), \quad \text{if } X,Y \text{ are independent} \\ &= (F_X(z))^2, \quad \text{if they are iid} \end{split}$$

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$\begin{split} F_Z(z) &= P[Z \leq z] = P[\max(X,Y) \leq z] \\ &= P[X \leq z, Y \leq z] \\ &= F_{XY}(z,z) \\ &= F_X(z)F_Y(z), \quad \text{if } X,Y \text{ are independent} \\ &= (F_X(z))^2, \quad \text{if they are iid} \end{split}$$

This is true of all random variables.

- Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$\begin{split} F_Z(z) &= P[Z \leq z] = P[\max(X,Y) \leq z] \\ &= P[X \leq z, Y \leq z] \\ &= F_{XY}(z,z) \\ &= F_X(z)F_Y(z), \quad \text{if } X,Y \text{ are independent} \\ &= (F_X(z))^2, \quad \text{if they are iid} \end{split}$$

- ► This is true of all random variables.
- ightharpoonup Suppose X,Y are iid continuous rv. Then density of Z is

- ightharpoonup Let us consider the  $\max$  and  $\min$  functions, in general.
- ▶ Let  $Z = \max(X, Y)$ . Then we have

$$\begin{split} F_Z(z) &= P[Z \leq z] = P[\max(X,Y) \leq z] \\ &= P[X \leq z, Y \leq z] \\ &= F_{XY}(z,z) \\ &= F_X(z)F_Y(z), \quad \text{if } X,Y \text{ are independent} \\ &= (F_X(z))^2, \quad \text{if they are iid} \end{split}$$

- ► This is true of all random variables.
- ightharpoonup Suppose X,Y are iid continuous rv. Then density of Z is

$$f_Z(z) = 2F_X(z)f_X(z)$$

▶ Suppose X, Y are iid uniform over (0, 1)

- ightharpoonup Suppose X, Y are iid uniform over (0, 1)
- ▶ Then we get df and pdf of Z = max(X, Y) as

- $\triangleright$  Suppose X, Y are iid uniform over (0, 1)
- ▶ Then we get df and pdf of Z = max(X, Y) as

$$F_Z(z) = z^2, 0 < z < 1;$$
 and  $f_Z(z) = 2z, 0 < z < 1$ 

- ightharpoonup Suppose X, Y are iid uniform over (0, 1)
- ▶ Then we get df and pdf of  $Z = \max(X, Y)$  as

$$F_Z(z) = z^2, 0 < z < 1;$$
 and  $f_Z(z) = 2z, 0 < z < 1$ 

$$F_Z(z)=0$$
 for  $z\leq 0$  and  $F_Z(z)=1$  for  $z\geq 1$ 



- ightharpoonup Suppose X, Y are iid uniform over (0, 1)
- ▶ Then we get df and pdf of  $Z = \max(X, Y)$  as

$$F_Z(z) = z^2, 0 < z < 1;$$
 and  $f_Z(z) = 2z, 0 < z < 1$ 

$$F_Z(z)=0$$
 for  $z\leq 0$  and  $F_Z(z)=1$  for  $z\geq 1$  and  $f_Z(z)=0$  outside  $(0,1)$ 

ightharpoonup This is easily generalized to n radom variables.

- ightharpoonup This is easily generalized to n radom variables.
- $\blacktriangleright \text{ Let } Z = \max(X_1, \cdots, X_n)$

- ightharpoonup This is easily generalized to n radom variables.
- $\blacktriangleright \text{ Let } Z = \max(X_1, \cdots, X_n)$

$$F_Z(z) = P[Z \le z]$$

- $\triangleright$  This is easily generalized to n radom variables.
- $\blacktriangleright \text{ Let } Z = \max(X_1, \cdots, X_n)$

$$F_Z(z) = P[Z \le z] = P[\max(X_1, X_2, \dots, X_n) \le z]$$

- $\triangleright$  This is easily generalized to n radom variables.
- ightharpoonup Let  $Z = \max(X_1, \cdots, X_n)$

$$F_Z(z) = P[Z \le z] = P[\max(X_1, X_2, \dots, X_n) \le z]$$
  
=  $P[X_1 \le z, X_2 \le z, \dots, X_n \le z]$ 

- ▶ This is easily generalized to *n* radom variables.
- ightharpoonup Let  $Z = \max(X_1, \cdots, X_n)$

$$F_{Z}(z) = P[Z \le z] = P[\max(X_{1}, X_{2}, \dots, X_{n}) \le z]$$

$$= P[X_{1} \le z, X_{2} \le z, \dots, X_{n} \le z]$$

$$= F_{X_{1} \dots X_{n}}(z, \dots, z)$$

- ▶ This is easily generalized to *n* radom variables.
- $\blacktriangleright \text{ Let } Z = \max(X_1, \cdots, X_n)$

$$F_Z(z) = P[Z \le z] = P[\max(X_1, X_2, \cdots, X_n) \le z]$$

$$= P[X_1 \le z, X_2 \le z, \cdots, X_n \le z]$$

$$= F_{X_1 \cdots X_n}(z, \cdots, z)$$

$$= F_{X_1}(z) \cdots F_{X_n}(z), \text{ if they are independent}$$

- ▶ This is easily generalized to *n* radom variables.
- $\blacktriangleright \text{ Let } Z = \max(X_1, \cdots, X_n)$

$$\begin{split} F_Z(z) &= P[Z \leq z] = P[\max(X_1, X_2, \cdots, X_n) \leq z] \\ &= P[X_1 \leq z, X_2 \leq z, \cdots, X_n \leq z] \\ &= F_{X_1 \cdots X_n}(z, \cdots, z) \\ &= F_{X_1}(z) \cdots F_{X_n}(z), \quad \text{if they are independent} \\ &= (F_X(z))^n, \quad \text{if they are iid} \\ &\qquad \qquad \text{where we take } F_X \text{ as the common df} \end{split}$$

- ightharpoonup This is easily generalized to n radom variables.
- $\blacktriangleright \text{ Let } Z = \max(X_1, \cdots, X_n)$

$$\begin{split} F_Z(z) &= P[Z \leq z] = P[\max(X_1, X_2, \cdots, X_n) \leq z] \\ &= P[X_1 \leq z, X_2 \leq z, \cdots, X_n \leq z] \\ &= F_{X_1 \cdots X_n}(z, \cdots, z) \\ &= F_{X_1}(z) \cdots F_{X_n}(z), \quad \text{if they are independent} \\ &= (F_X(z))^n, \quad \text{if they are iid} \\ &\qquad \qquad \text{where we take } F_X \text{ as the common df} \end{split}$$

▶ For example if all  $X_i$  are uniform over (0,1) and ind, then  $F_Z(z) = z^n, \ 0 < z < 1$ 

▶ Consider  $Z = \min(X, Y)$  and X, Y independent

▶ Consider  $Z = \min(X, Y)$  and X, Y independent

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

 $\blacktriangleright$  It is difficult to write this in terms of joint df of X,Y.

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X,Y*.
- So, we consider the following

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X*, *Y*.
- ► So, we consider the following

$$P[Z > z] = P[\min(X, Y) > z]$$

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X*, *Y*.
- ► So, we consider the following

$$P[Z > z] = P[\min(X, Y) > z]$$
$$= P[X > z, Y > z]$$

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X*, *Y*.
- ► So, we consider the following

$$\begin{split} P[Z>z] &= P[\min(X,Y)>z] \\ &= P[X>z,Y>z] \\ &= P[X>z]P[Y>z], \quad \text{using independence} \end{split}$$

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X*, *Y*.
- ► So, we consider the following

$$P[Z>z] = P[\min(X,Y)>z]$$
  
=  $P[X>z,Y>z]$   
=  $P[X>z]P[Y>z]$ , using independence  
=  $(1-F_X(z))(1-F_Y(z))$ 

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X*, *Y*.
- ► So, we consider the following

$$\begin{split} P[Z>z] &= P[\min(X,Y)>z] \\ &= P[X>z,Y>z] \\ &= P[X>z]P[Y>z], \quad \text{using independence} \\ &= (1-F_X(z))(1-F_Y(z)) \\ &= (1-F_X(z))^2, \quad \text{if they are iid} \end{split}$$

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X*, *Y*.
- ► So, we consider the following

$$\begin{split} P[Z>z] &= P[\min(X,Y)>z] \\ &= P[X>z,Y>z] \\ &= P[X>z]P[Y>z], \quad \text{using independence} \\ &= (1-F_X(z))(1-F_Y(z)) \\ &= (1-F_X(z))^2, \quad \text{if they are iid} \end{split}$$

Hence, 
$$F_Z(z) = 1 - (1 - F_X(z))(1 - F_Y(z))$$

$$F_Z(z) = P[Z \le z] = P[\min(X, Y) \le z]$$

- ▶ It is difficult to write this in terms of joint df of *X*, *Y*.
- ► So, we consider the following

$$\begin{split} P[Z>z] &= P[\min(X,Y)>z] \\ &= P[X>z,Y>z] \\ &= P[X>z]P[Y>z], \quad \text{using independence} \\ &= (1-F_X(z))(1-F_Y(z)) \\ &= (1-F_X(z))^2, \quad \text{if they are iid} \end{split}$$

Hence, 
$$F_Z(z) = 1 - (1 - F_X(z))(1 - F_Y(z))$$

lackbox We can once again find density of Z if X,Y are continuous

▶ Suppose X, Y are iid uniform (0, 1).

- ▶ Suppose X, Y are iid uniform (0, 1).
- $ightharpoonup Z = \min(X, Y)$

- ightharpoonup Suppose X, Y are iid uniform (0, 1).
- $ightharpoonup Z = \min(X, Y)$

$$F_Z(z) = 1 - (1 - F_X(z))^2 = 1 - (1 - z)^2, 0 < z < 1$$

- ightharpoonup Suppose X, Y are iid uniform (0, 1).
- $ightharpoonup Z = \min(X, Y)$

$$F_Z(z) = 1 - (1 - F_X(z))^2 = 1 - (1 - z)^2, 0 < z < 1$$

Notice that P[X > z] = (1 - z).

- ▶ Suppose X, Y are iid uniform (0, 1).
- $ightharpoonup Z = \min(X, Y)$

$$F_Z(z) = 1 - (1 - F_X(z))^2 = 1 - (1 - z)^2, 0 < z < 1$$

- Notice that P[X > z] = (1 z).
- ightharpoonup We get the density of Z as

$$f_Z(z) = 2(1-z), 0 < z < 1$$

 $ightharpoonup \min$  fn is also easily generalized to n random variables

- $ightharpoonup \min$  fn is also easily generalized to n random variables
- $\blacktriangleright \text{ Let } Z = \min(X_1, X_2, \cdots, X_n)$

- $ightharpoonup \min$  fn is also easily generalized to n random variables
- $\blacktriangleright \text{ Let } Z = \min(X_1, X_2, \cdots, X_n)$

$$P[Z > z] = P[\min(X_1, X_2, \cdots, X_n) > z]$$

- $ightharpoonup \min$  fn is also easily generalized to n random variables
- $\blacktriangleright \text{ Let } Z = \min(X_1, X_2, \cdots, X_n)$

$$P[Z > z] = P[\min(X_1, X_2, \dots, X_n) > z]$$
  
=  $P[X_1 > z, \dots, X_n > z]$ 

- $ightharpoonup \min$  fn is also easily generalized to n random variables
- $\blacktriangleright \text{ Let } Z = \min(X_1, X_2, \cdots, X_n)$

$$\begin{split} P[Z>z] &= P[\min(X_1,X_2,\cdots,X_n)>z] \\ &= P[X_1>z,\cdots,X_n>z] \\ &= P[X_1>z]\cdots P[X_n>z], \quad \text{using independence} \end{split}$$

- $ightharpoonup \min$  fn is also easily generalized to n random variables
- $\blacktriangleright \text{ Let } Z = \min(X_1, X_2, \cdots, X_n)$

$$\begin{split} P[Z>z] &= P[\min(X_1, X_2, \cdots, X_n) > z] \\ &= P[X_1>z, \cdots, X_n>z] \\ &= P[X_1>z] \cdots P[X_n>z], \quad \text{using independence} \\ &= (1-F_{X_1}(z)) \cdots (1-F_{X_n}(z)) \end{split}$$

- $ightharpoonup \min$  fn is also easily generalized to n random variables
- $\blacktriangleright \text{ Let } Z = \min(X_1, X_2, \cdots, X_n)$

$$\begin{split} P[Z>z] &= P[\min(X_1, X_2, \cdots, X_n) > z] \\ &= P[X_1 > z, \cdots, X_n > z] \\ &= P[X_1 > z] \cdots P[X_n > z], \quad \text{using independence} \\ &= (1 - F_{X_1}(z)) \cdots (1 - F_{X_n}(z)) \\ &= (1 - F_X(z))^n, \quad \text{if they are iid} \end{split}$$

- $ightharpoonup \min$  fn is also easily generalized to n random variables
- $\blacktriangleright \text{ Let } Z = \min(X_1, X_2, \cdots, X_n)$

$$\begin{split} P[Z>z] &= P[\min(X_1, X_2, \cdots, X_n) > z] \\ &= P[X_1 > z, \cdots, X_n > z] \\ &= P[X_1 > z] \cdots P[X_n > z], \quad \text{using independence} \\ &= (1 - F_{X_1}(z)) \cdots (1 - F_{X_n}(z)) \\ &= (1 - F_X(z))^n, \quad \text{if they are iid} \end{split}$$

 $\blacktriangleright$  Hence, when  $X_i$  are iid, the df of Z is

$$F_Z(z) = 1 - (1 - F_X(z))^n$$

where  $F_X$  is the common df