

# Recap: Brownian Motion

- ▶  $\{X(t), t \geq 0\}$  is called a Brownian motion if
  1.  $X(0) = 0$
  2. The process has stationary and independent increments
  3. For every  $t > 0$ ,  $X(t)$  is Gaussian with mean 0 and variance  $\sigma^2 t$
- ▶ If  $\sigma^2 = 1$ , it is called standard Brownian Motion
- ▶  $Y(t) = X(t) + \mu t$  is called Brownian motion with a drift

## Recap: Some properties

- ▶ An important result is that Brownian motion paths are continuous
- ▶ The autocorrelation and autocovariance of the process is

$$R(t_1, t_2) = \sigma^2 \min(t_1, t_2) = \text{Cov}(X(t_1), X(t_2))$$

- ▶ All  $n^{\text{th}}$  order distributions are Gaussian:  
 $X(t_1), \dots, X(t_n)$  are jointly Gaussian.

# Recap: Gaussian Processes

- ▶ A continuous-time continuous-state process  $\{X(t), t \geq 0\}$  is said to be a Gaussian process if for all  $n$  and all  $t_1, t_2, \dots, t_n$ , we have that  $X(t_1), \dots, X(t_n)$  are jointly Gaussian.
- ▶ A general Gaussian process is specified by the mean function and the variance and covariance functions
- ▶ The Brownian motion is an example of a Gaussian Process
- ▶ The Brownian motion is a Gaussian process with

$$E[X(t)] = 0, \quad \text{Var}(X(t)) = \sigma^2 t, \quad \text{Cov}(X(s), X(t)) = \sigma^2 \min(s, t)$$

# Recap: Conditional Densities

- For  $s < t$ ,  $f_{X(s)|X(t)}$  is Gaussian and

$$E[X(s)|X(t)] = \frac{s}{t} X(t); \quad \text{Var}(X(s)|X(t)) = \frac{s}{t} (t - s)$$

- For  $s < t$ , we also showed  $f_{X(t)|X(s)}$  is Gaussian with

$$E[X(t) | X(s)] = X(s); \quad \text{Var}(X(t)|X(s)) = (t - s)$$

## Recap: Hitting Times

- ▶ Let  $T_a$  denote the first time Brownian motion hits  $a$ .

$$Pr[T_a \leq t] = \frac{2}{\sqrt{2\pi}} \int_{|a|/\sqrt{t}}^{\infty} e^{-\frac{y^2}{2}} dy$$

- ▶ By continuity of sample paths,

$$Pr\left[\max_{0 \leq s \leq t} X(s) \geq a\right] = Pr[T_a \leq t] = 2Pr[X(t) \geq a]$$

# Recap: Geometric Brownian Motion

- ▶ Let  $\{Y(t), t \geq 0\}$  be a Brownian motion with drift. Define

$$X(t) = e^{Y(t)}$$

- ▶ Then,  $\{X(t), t \geq 0\}$  is called geometric Brownian motion. It is useful in mathematical finance

- ▶ We have considered three random processes
- ▶ (discrete-time) Markov Chain
  - Example of Discrete-time discrete-state process
- ▶ Continuous time Markov Chains (e.g., Poisson Process)
  - Example of continuous-time discrete-state process
- ▶ Brownian Motion
  - Example of continuous-time continuous-state process
- ▶ We need an example of discrete-time continuous-state process!
- ▶ Any sequence of continuous random variables would be a discrete-time continuous-state process

- ▶ In general, any ‘stochastic’ algorithm would generate discrete-state continuous time process.
- ▶ If an algorithm uses a random step, then the algorithm would be like

$$X(n+1) = X(n) + \eta_n G(X(n), \xi(n))$$

where  $\xi(n)$  would be some random variable which may be dependent on  $X(n)$ .

- ▶ Many algorithms can be written in this general form.
- ▶ The  $X(n), n = 0, 1, \dots$  would be a discrete-time continuous-state stochastic process
- ▶ We consider an important class of such processes



- ▶ Let  $\{X_n, n = 0, 1, \dots\}$  be a discrete-time continuous-state process.
- ▶ It is called a **martingale** if  $E|X_n| < \infty, \forall n$  and

$$E[X_{n+1} \mid X_n, \dots, X_0] = X_n, \forall n$$

- ▶ One can think of Martingale as a general fair gambling game.
- ▶ When  $X_n$  is a martingale, we have

$$E[X_{n+1}] = E[X_n], \forall n$$

## Example

- ▶ Sequence of sums of independent random variables forms a Martingale
- ▶ Suppose  $Z_i$  are iid with  $Pr[Z_i = +1] = Pr[Z_i = -1] = 0.5$ . Let

$$X_n = \sum_{i=1}^n Z_i \quad \Rightarrow \quad X_{n+1} = X_n + Z_{n+1}$$

- ▶ Since  $EZ_i = 0, \forall i$ ,

$$E[X_{n+1} \mid X_n, \dots] = E[X_n + Z_{n+1} \mid X_n, \dots] = X_n + E[Z_{n+1} \mid X_n, \dots] = X_n$$

- ▶ Hence,  $X_n$  is a martingale.

# Example

- ▶ Suppose  $X_1, X_2, \dots$  are iid with  $E[X_i] = 1$ .
- ▶ Let  $Z_n = \prod_{i=1}^n X_i$ .
- ▶ Since  $Z_{n+1} = Z_n X_{n+1}$ ,

$$\begin{aligned} E[Z_{n+1} | Z_n, \dots, Z_1] &= E[Z_n X_{n+1} | Z_n, \dots, Z_1] \\ &= Z_n E[X_{n+1} | Z_n, \dots, Z_1] \\ &= Z_n, \text{ since } X_i \text{ iid, } E[X_i] = 1 \end{aligned}$$

- ▶ Thus,  $Z_n$  is a martingale

- Given  $\{X_n, n = 1, 2, \dots\}$  and  $E|X_n| < \infty, \forall n$ , It is called a martingale if

$$E[X_{n+1} \mid X_n, \dots, X_1] = X_n, \forall n$$

- We may often actually show

$$E[X_{n+1} \mid Y_n, \dots, Y_1] = X_n, \forall n$$

Where  $Y_1, \dots, Y_n$  determine  $X_1, \dots, X_n$  for all  $n$ .

- This is justified by the following two identities

$$\begin{aligned} E[ E[X_{n+1} \mid X_n, \dots, X_1, Y_n, \dots, Y_1] \mid X_1, \dots, X_n] \\ = E[X_{n+1} \mid X_n, \dots, X_1] \end{aligned}$$

for any  $X_i, Y_i$  by properties of conditional expectation

$$E[X_{n+1} \mid X_n, \dots, X_1, Y_n, \dots, Y_1] = E[X_{n+1} \mid Y_n, \dots, Y_1]$$

when  $Y_1, \dots, Y_n$  determine  $X_1, \dots, X_n$  for all  $n$

# A property of Conditional Expectation

- ▶ Conditional expectation satisfies

$$E[ E[h(X)|Y, Z] | Y] = E[h(X)|Y]$$

Note that all these can be random vectors.

- ▶ Let

$$\begin{aligned} g_1(Y, Z) &= E[h(X)|Y, Z] \\ g_2(Y) &= E[g_1(Y, Z)|Y] \end{aligned}$$

We want to show  $g_2(Y) = E[h(X)|Y]$

► Recall:  $g_1(Y, Z) = E[h(X)|Y, Z]$ ,  $g_2(Y) = E[g_1(Y, Z)|Y]$

$$\begin{aligned} g_2(y) &= \int g_1(y, z) f_{Z|Y}(z|y) dz \\ &= \int \left[ \int h(x) f_{X|YZ}(x|y, z) dx \right] f_{Z|Y}(z|y) dz \\ &= \int h(x) \left[ \int f_{X|YZ}(x|y, z) f_{Z|Y}(z|y) dz \right] dx \\ &= \int h(x) \left[ \int f_{XZ|Y}(x, z|y) dz \right] dx \\ &= \int h(x) f_{X|Y}(x|y) dx \end{aligned}$$

► Thus we get

$$E[ E[h(X)|Y, Z] | Y ] = E[h(X)|Y]$$

# Example

- ▶ We saw that partial sums of iid  $X_i$  with  $E[X_i] = 0$  is a martingale.
- ▶ We can generalize this.
- ▶ Let  $X_1, X_2, \dots$  any sequence of random variables.
- ▶ Let  $W_i = X_i - E[X_i \mid X_1, \dots, X_{i-1}]$ .
- ▶ Now  $E[W_i] = 0, \forall i$  though  $W_i$  may not be independent.
- ▶ Let  $Z_n = \sum_{i=1}^n W_i$ .
- ▶ We can show that  $Z_n, n = 1, 2, \dots$  is a martingale, assuming that  $E[|Z_n|] < \infty$ .

- ▶ We have

$$Z_n = \sum_{i=1}^n W_i = \sum_{i=1}^n X_i - E[X_i | X_1, \dots, X_{i-1}]$$

- ▶ Note that  $X_1, \dots, X_n$  determine  $Z_1, \dots, Z_n, \forall n$
- ▶ We have,  $Z_{n+1} = Z_n + X_{n+1} - E[X_{n+1} | X_1, \dots, X_n]$ .

$$\begin{aligned} E[Z_{n+1} | X_1, \dots, X_n] &= E[Z_n + W_{n+1} | X_1, \dots, X_n] \\ &= Z_n + E[X_{n+1} | X_1, \dots, X_n] - \\ &\quad E[X_{n+1} | X_1, \dots, X_n] \\ &= Z_n \end{aligned}$$

- ▶ This shows that  $Z_n$  is a martingale



- ▶ Let  $N$  be a positive integer valued random variable with  $P[N < \infty] = 1$ .
- ▶  $N$  is said to be a stopping time for the process  $Z_1, Z_2, \dots$  if the event  $[N = n]$  is determined by the random variable  $Z_1, \dots, Z_n$ .
- ▶ If we know the values of  $Z_1, \dots, Z_n$ , then we can say whether or not  $N = n$ .
- ▶ The idea is that we can decide to stop the process at  $N$  and the decision to stop cannot anticipate the future

- ▶ Let  $N$  be a stopping time for the process  $\{Z_n, n \geq 1\}$ .
- ▶ Define

$$\bar{Z}_n = \begin{cases} Z_n & \text{if } n \leq N \\ Z_N & \text{if } n > N \end{cases}$$

- ▶ We call  $\{\bar{Z}_n, n \geq 1\}$  the stopped process

- ▶ Let  $\{Z_n, n \geq 1\}$  be a martingale and let  $N$  a stopping time for it.
- ▶ **Theorem:** The stopped Process is a martingale  
That is, a stopped martingale is a martingale
- ▶ Note that, by definition,  $\bar{Z}_1 = Z_1$ .
- ▶ So, if the stopped process is a martingale, then,

$$E[\bar{Z}_n] = E[\bar{Z}_1] = E[Z_1], \forall n$$

- ▶ Recall that for a martingale,  $E[Z_n] = E[Z_1], \forall n$
- ▶ The theorem essentially says there is no strategy to have a positive expectation from a fair gambling game.

- ▶  $\{X_n, n = 0, 1, \dots\}$  and  $E|X_n| < \infty, \forall n$  is called a martingale if

$$E[X_{n+1} \mid X_n, \dots, X_0] = X_n, \forall n$$

- ▶ It is called a supermartingale if

$$E[X_{n+1} \mid X_n, \dots, X_0] \leq X_n, \forall n$$

- ▶ It is called a submartingale if

$$E[X_{n+1} \mid X_n, \dots, X_0] \geq X_n, \forall n$$

- ▶ Note that  $E[X_{n+1}] \leq E[X_n]$  for supermartingales and  $E[X_{n+1}] \geq E[X_n]$  for submartingales.
- ▶ In the above, the conditioning random variables can be another sequence  $Y_i$  if  $Y_1, \dots, Y_n$  determine  $X_1, \dots, X_n$

- ▶ A useful result is the martingale convergence theorem.

**Martingale convergence theorem:** If  $X_n$  is a martingale with  $\sup_n E|X_n| < \infty$  then  $X_n$  converges almost surely to a rv  $X$  which will have finite expectation. A positive supermartingale also converges almost surely

- ▶ If  $X_n$  are bounded, then the condition is always true and the almost sure convergence implies convergence in the mean.
- ▶ This is often useful in dealing with many sequences of random variables such as a stochastic algorithm.

# Example

- ▶ We consider a simple algorithm for a two armed bandit problem.
- ▶ 2 arms. Response is binary (1 for reward).
- ▶  $d_i$  - prob of reward for arm- $i$ . Do not know  $d_i$
- ▶ need to play and find which is better arm.
- ▶ We choose arm-1 with prob  $p(k)$  (and hence arm-2 with prob  $(1 - p(k))$ ) at iteration  $k$  and update  $p(k)$  based on the outcome.

- ▶  $p(k)$  denotes probability of choosing arm-1 at  $k$  and  $b(k)$  denotes the response.
- ▶ Our algorithm for the 2-armed bandit problem is

$$\begin{aligned} p(k+1) &= p(k) + \lambda(1 - p(k)) \quad \text{if arm 1 chosen, } b(k) = 1 \\ &= p(k) - \lambda p(k) \quad \text{if arm 2 is chosen and } b(k) = 1 \\ &= p(k) \quad \text{if } b(k) = 0 \end{aligned}$$

- ▶ We want to know whether the algorithm converges.

- The algorithm is

$$\begin{aligned} p(k+1) &= p(k) + \lambda(1 - p(k)) \quad \text{if arm 1 chosen, } b(k) = 1 \\ &= p(k) - \lambda p(k) \quad \text{if arm 2 is chosen and } b(k) = 1 \\ &= p(k) \quad \text{if } b(k) = 0 \end{aligned}$$

- We get

$$\begin{aligned} &E[p(k+1) - p(k) | p(k)] \\ &= \lambda(1 - p(k)) Pr[b(k) = 1, \text{arm 1} | p(k)] \\ &\quad - \lambda p(k) Pr[b(k) = 1, \text{arm 2} | p(k)] \\ &= \lambda(1 - p(k)) Pr[b(k) = 1 | \text{arm 1}, p(k)] Pr[\text{arm 1} | p(k)] \\ &\quad - \lambda p(k) Pr[b(k) = 1 | \text{arm 2}, p(k)] Pr[\text{arm 2} | p(k)] \end{aligned}$$



- ▶ This gives us

$$\begin{aligned} E[p(k+1) - p(k) | p(k)] &= \lambda(1 - p(k)) d_1 p(k) \\ &\quad - \lambda p(k) d_2 (1 - p(k)) \\ &= \lambda p(k)(1 - p(k)) (d_1 - d_2) \\ &\geq 0, \quad \text{if } d_1 > d_2 \end{aligned}$$

$$\Rightarrow E[p(k+1) | p(k)] \geq p(k) \Rightarrow E[p(k+1)] \geq E[p(k)], \forall k$$

- ▶ This also shows  $p(k)$  is a submartingale.
- ▶ Here,  $p(k)$  is bounded and  $1 - p(k)$  is a supermartingale.
- ▶ So, we can conclude, the algorithm converges almost surely

- ▶ We considered martingales as an example of discrete-time continuous-state processes
- ▶ Stochastic iterative algorithms generate such processes.
- ▶ Martingales can be useful in analyzing convergence of many stochastic algorithms
- ▶ While we mentioned only discrete-time martingales, one can similarly have continuous-time martingales that satisfy

$$E[X(t)|(X(s'), 0 \leq s' \leq s < t] = X(s)$$