

Recap: Convergence in Probability

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► By the definition of limit, the above means

$$\forall \delta > 0, \exists N < \infty, \text{ s.t. } P[|X_n - X_0| > \epsilon] < \delta, \forall n > N$$

Recap: Almost sure convergence

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Recap: Convergence in r^{th} mean

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- ▶ $X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$
- ▶ $X_n \xrightarrow{d} k \Rightarrow X_n \xrightarrow{P} k$, where k is a constant

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- ▶ Hence $N_n \xrightarrow{P} \theta$
- ▶ Does it converge almost surely? In the mean?

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- ▶ Hence, a sufficient condition is $\sigma_n^2 \rightarrow 0$.
- ▶ What is a sufficient condition for convergence almost surely?

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- ▶ All the implications are one-way and we have seen counter examples
- ▶ In general, almost sure convergence does not imply convergence in r^{th} mean and vice versa

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$$\lim_{n \rightarrow \infty} P[\tilde{S}_n \leq a] = \Phi(a) \triangleq \int_{-\infty}^a \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

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- ▶ We use characteristic functions for proving CLT

Characteristic Function

- ▶ Given rv X , its characteristic function, ϕ_X , is defined by

$$\phi_X(u) = E[e^{iuX}] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

Characteristic Function

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$$\phi_X(u) = E[e^{iuX}] = \int e^{iux} dF_X(x) \quad (i = \sqrt{-1})$$

- ▶ Since $|e^{iux}| \leq 1$, ϕ_X exists for all random variables

Properties of characteristic function

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- ▶ If ν_r is finite, then

$$\phi_X(u) = \sum_{s=0}^{r-1} \mu_s \frac{(iu)^s}{s!} + \rho(u) \mu_r \frac{(iu)^r}{r!}$$

where $|\rho(u)| \leq 1$ and $\rho(u) \rightarrow 1$ as $u \rightarrow 0$

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- ▶ **Continuity theorem**
 - ▶ If $F_n \rightarrow F$ then $\phi_{F_n} \rightarrow \phi_F$
 - ▶ If $\phi_{F_n} \rightarrow \psi$ and ψ is continuous at zero, then ψ would be characteristic function of some df, say, F , and $F_n \rightarrow F$

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- ▶ Recall $M_X(t) = (pe^t + (1-p))^n$
- ▶ We have $\phi_X(u) = M_X(iu)$

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$$\lim_{n \rightarrow \infty} P \left[\tilde{S}_n \leq x \right] = \lim_{n \rightarrow \infty} P \left[\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq x \right] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad \forall x$$

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Proof:

- ▶ Without loss of generality let us assume $\mu = 0$.
- ▶ We use characteristic function of \tilde{S}_n for the proof.

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- Recall that we can expand ϕ in a Taylor series

$$\phi(t) = \sum_{s=0}^{r-1} \mu_s \frac{(it)^s}{s!} + \rho(t) \mu_r \frac{(it)^r}{r!}, \quad \rho(t) \rightarrow 1, \text{ as } t \rightarrow 0$$

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► By Continuity theorem, distribution function of \tilde{S}_n converges to that of standard Normal rv

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- ▶ Thus, S_n is well approximated by a normal rv with mean $n\mu$ and variance $n\sigma^2$, if n is large

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 &\approx \Phi\left(\frac{3}{\sqrt{\frac{5}{3}}}\right) - \Phi\left(\frac{-3}{\sqrt{\frac{5}{3}}}\right) \\
 &\approx \Phi(2.3) - \Phi(-2.3) \\
 &= 0.9893 - 0.0107 \approx 0.98
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- ▶ Hence probability that the sum differs from true sum by more than 3 is 0.02

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- To get a good approximation, to calculate $P[S_n \leq m]$ one uses $P[S_n \leq m + 0.5]$ in the above approximation formula

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(because $\Phi(-x) = (1 - \Phi(x))$)

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- If we change ϵ to 0.05, then at variance equal to 0.5 the probability becomes about 0.02 while the Chebyshev bound would be about 0.2