

Recap: Expectation

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- ▶ We mostly consider rv with finite expectation where the sum or integral above is absolutely convergent
- ▶ Note that expectation is defined for all random variables

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- ▶ This result is true for all rv though we consider only discrete or continuous.

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- ▶ $E[ag_1(X) + bg_2(X)] = aE[g_1(X)] + bE[g_2(X)]$
- ▶ $E[(X - c)^2] \geq E[(X - EX)^2], \forall c$

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- $\text{Var}(cX) = c^2 \text{Var}(X)$

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- ▶ The k^{th} (order) moment of X is

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- ▶ If moment of order k is finite then so is moment of order s for $s < k$.

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- ▶ We say the mgf exists if $E[e^{tX}] < \infty$ for t in some interval around zero. When it exists we have

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- ▶ The mgf uniquely determines the df

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- ▶ We have

$$f_X(0) = P_X(0); \quad f_X(1) = \frac{P'_X(0)}{1!}; \quad f_X(2) = \frac{P''_X(0)}{2!} \dots$$

$$P'_X(1) = EX; \quad P''_X(1) = E[X(X-1)] \dots$$

► Let $p \in (0, 1)$. The number $x \in \Re$ that satisfies

$$P[X \leq x] \geq p \quad \text{and} \quad P[X \geq x] \geq 1 - p$$

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$$\begin{aligned} \text{► } p &\leq P[X \leq x] = F_X(x) \\ \text{► } 1 - p &\leq 1 - P[X < x] = 1 - (P[X \leq x] - P[X = x]) \\ &\Rightarrow 1 - p \leq 1 - F_X(x) + P[X = x] \\ &\Rightarrow F_X(x) \leq p + P[X = x] \end{aligned}$$

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- ▶ Note that for a given p there can be multiple values for x to satisfy the above.

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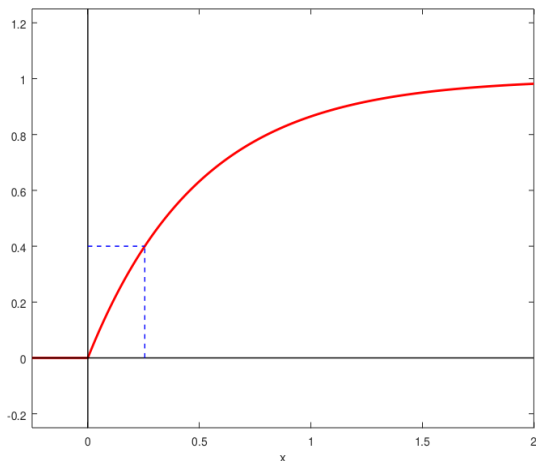
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- ▶ If X is continuous rv, we need to satisfy $p = F_X(x)$.
- ▶ In general, for a given p , there may be multiple x that satisfy the above.
- ▶ Let us see some examples.

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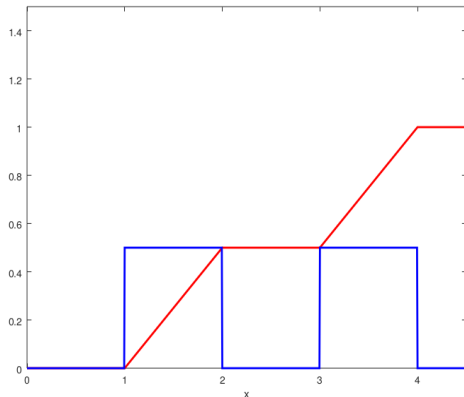


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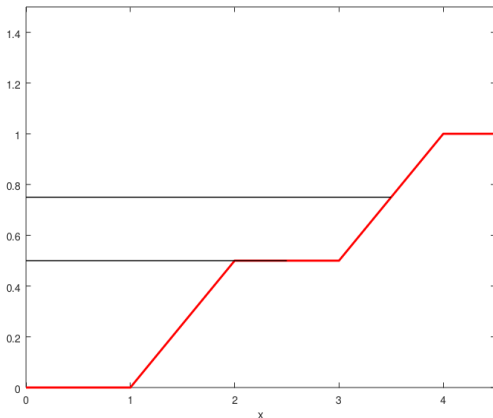
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But for $p = 0.75$ it is unique.

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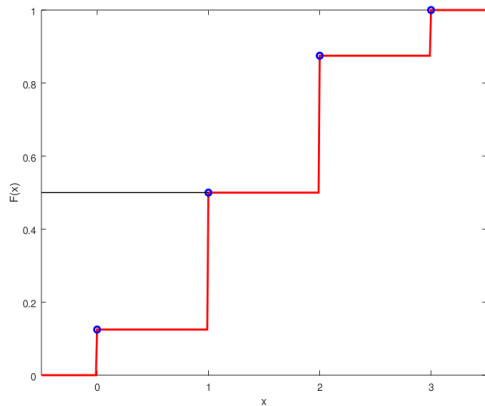
- ▶ Let $X \in \{x_1, x_2, \dots\}$
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- ▶ For $x_i \leq x \leq x_{i+1}$, we have $p \leq F_X(x) \leq p + P[X = x]$

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- ▶ For $x_i \leq x \leq x_{i+1}$, we have $p \leq F_X(x) \leq p + P[X = x]$
- ▶ So, quantile of order p is not unique and all such x qualify.

► This situation is illustrated below



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- ▶ Hence we have

$$p < p + \delta_2 = F_X(x_i) < p + \delta_2 + \delta_1 = p + P[X = x_i]$$

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- ▶ Let $F_X(x_{i-1}) = p - \delta_1$ and $F_X(x_i) = p + \delta_2$. (Note that $\delta_1, \delta_2 > 0$)
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- ▶ Hence we have

$$p < p + \delta_2 = F_X(x_i) < p + \delta_2 + \delta_1 = p + P[X = x_i]$$

- ▶ Hence, x_i is quantile of order p .

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- ▶ Let $F_X(x_{i-1}) = p - \delta_1$ and $F_X(x_i) = p + \delta_2$. (Note that $\delta_1, \delta_2 > 0$)
- ▶ Then $P[X = x_i] = F_X(x_i) - F_X(x_{i-1}) = \delta_2 + \delta_1$
- ▶ Hence we have

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- ▶ Hence, x_i is quantile of order p .
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- ▶ Now suppose p is such that $F_X(x_{i-1}) < p < F_X(x_i)$.
- ▶ Let $F_X(x_{i-1}) = p - \delta_1$ and $F_X(x_i) = p + \delta_2$. (Note that $\delta_1, \delta_2 > 0$)
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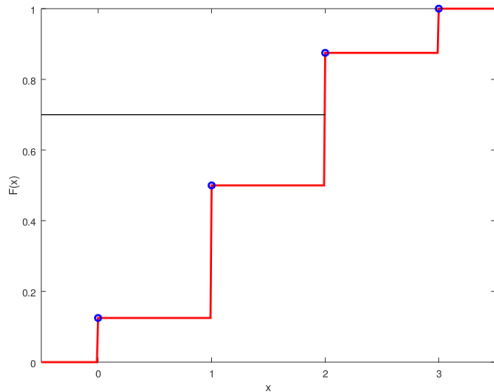
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- ▶ For any x , with $x_i < x < x_{i+1}$ we have $p + P[X = x] = p < F_X(x) = p + \delta_2$.
- ▶ Similarly, for $x \geq x_{i+1}$ we have $F_X(x) > p + P[X = x]$.
- ▶ Thus quantile of order p is unique here.

► This situation is illustrated below



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- ▶ One can show that $\int_{-\infty}^0 f_X(x) dx = 0.5$ and hence the median is at the origin.

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(Exercise: Show this for discrete and continuous rv)

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- ▶ For the Gaussian density, the mode, the median and the mean are all same.

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- ▶ These help us bound the probabilities of some important events in terms of the moments.

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- ▶ Markov inequality is often used in this form.

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- ▶ An example of what are called concentration inequalities.

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- ▶ This is true for all random variables and the RHS above does not depend on the distribution of X .

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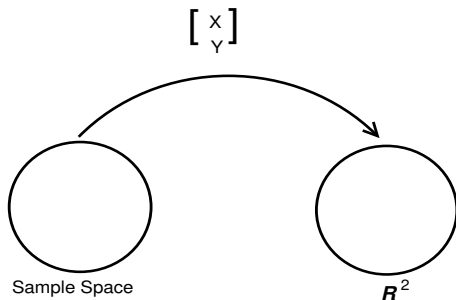
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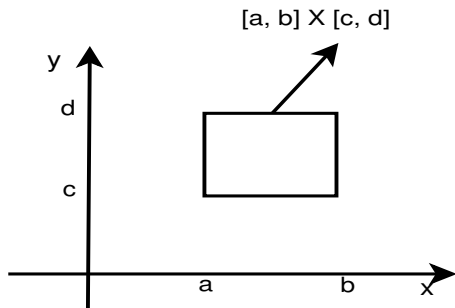
$$\mathcal{B}^2 = \sigma(\{B_1 \times B_2 : B_1, B_2 \in \mathcal{B}\})$$

where \mathcal{B} is the Borel σ -algebra we considered earlier, and \mathcal{B}^2 is the set of Borel sets of \mathbb{R}^2 .

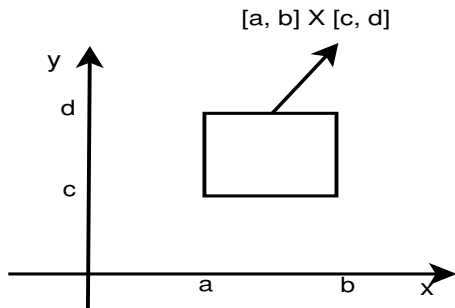
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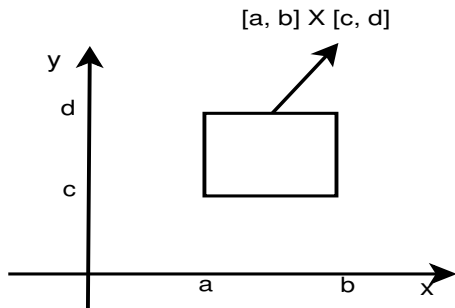


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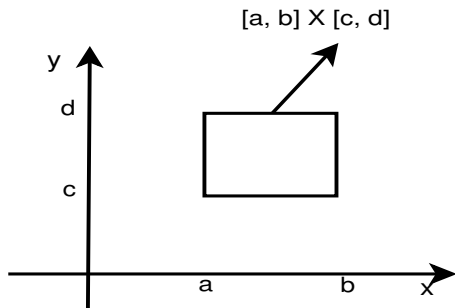
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- ▶ The joint distribution function is the probability of the intersection of the events $[X \leq x]$ and $[Y \leq y]$.

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- ▶ But there is another crucial property satisfied by F_{XY} .

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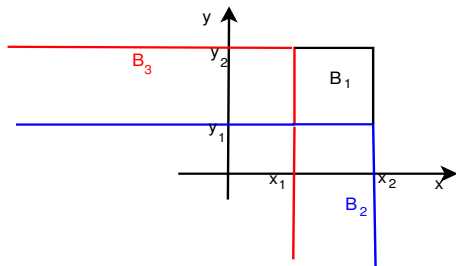
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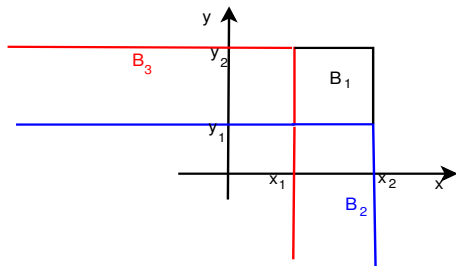
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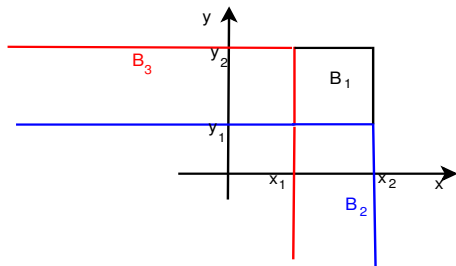


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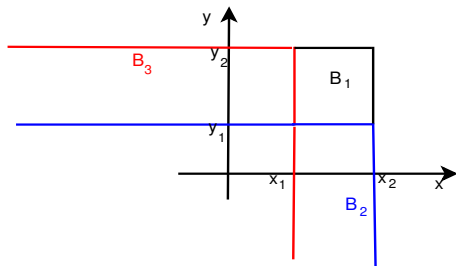
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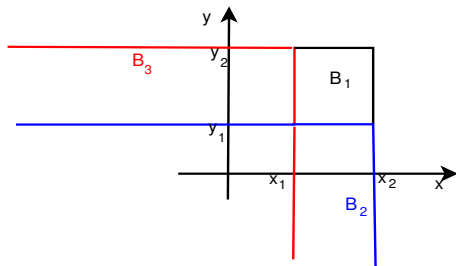
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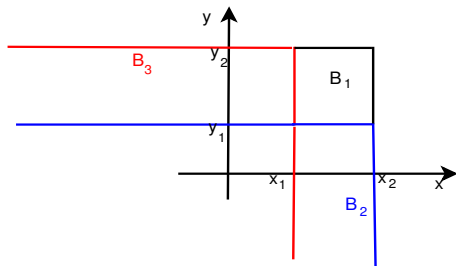
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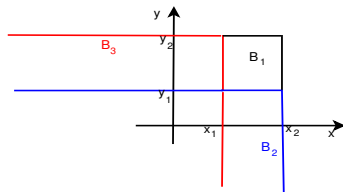
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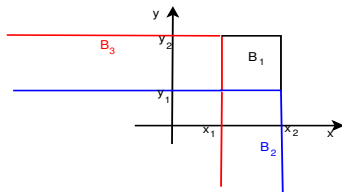
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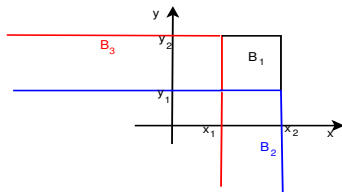
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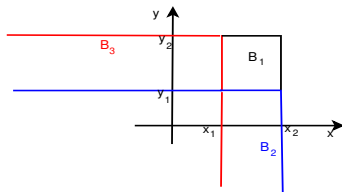




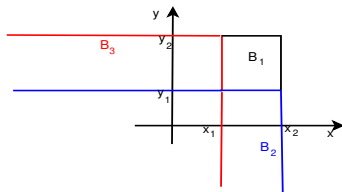
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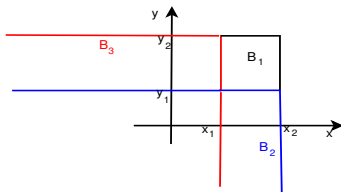


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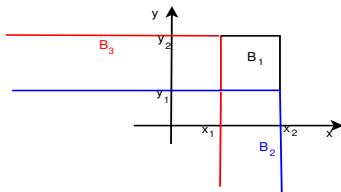
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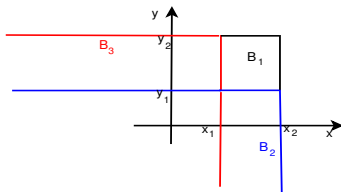


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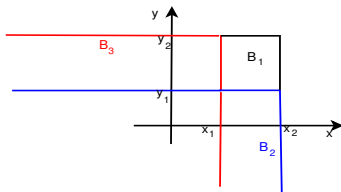
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- ▶ This is an additional condition that a function has to satisfy to be the joint distribution function of a pair of random variables

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- ▶ Any $F : \mathbb{R}^2 \rightarrow \mathbb{R}$ satisfying the above would be a joint distribution function.

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 - ▶ $f_{XY}(x, y) \geq 0, \forall x, y$ and $\sum_i \sum_j f_{XY}(x_i, y_j) = 1$
- ▶ This is a straight-forward extension of the pmf of a single discrete rv.

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- ▶ If $\omega = 0.2576$ then $X(\omega) = 2$ and $Y(\omega) = 5$
- ▶ Easy to see that $X, Y \in \{0, 1, \dots, 9\}$.
- ▶ We want to calculate the joint pmf of X and Y

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- ▶ Hence the joint pmf of X and Y is

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- ▶ So, $P[X = m, Y = n]$ is either $2/36$ or $1/36$ (assuming m, n satisfy other requirements)

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Joint Probability mass function

- ▶ Let $X \in \{x_1, x_2, \dots\}$ and $Y \in \{y_1, y_2, \dots\}$ be discrete random variables.
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- ▶ Given the joint pmf, we can get the joint df as

$$F_{XY}(x, y) = \sum_{\substack{i: \\ x_i \leq x}} \sum_{\substack{j: \\ y_j \leq y}} f_{XY}(x_i, y_j)$$

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