Recap: Multi-dimensional Gaussian density

 $ightharpoonup \mathbf{X} = (X_1, \cdots, X_n)^T$ are said to be jointly Gaussian if

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{(2\pi)^{\frac{n}{2}} |\Sigma|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1}(\mathbf{x} - \boldsymbol{\mu})}$$

- $ightharpoonup E\mathbf{X} = \boldsymbol{\mu}$ and $\Sigma_X = \Sigma$.
- ▶ The moment generating function is given by

$$M_{\mathbf{x}}(\mathbf{s}) = e^{\mathbf{s}^T \boldsymbol{\mu} + \frac{1}{2} \mathbf{s}^T \boldsymbol{\Sigma} \mathbf{s}}$$

 $\begin{tabular}{ll} \hline \end{tabular} \begin{tabular}{ll} When X,Y are jointly Gaussian, the joint density is given by \\ \end{tabular}$

$$f_{XY}(x,y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)} \left(\frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}\right)}$$

Recap

- ▶ If X_1, \dots, X_n are jointly Gaussian then they are independent if they are uncorrelated.
- Mhen X_1, \dots, X_n be jointly Gaussian (with zero means), there is an orthogonal transform $\mathbf{Y} = A\mathbf{X}$ such that Y_1, \dots, Y_n are jointly Gaussian and independent.
- X_1, \dots, X_n are jointly Gaussian if and only if $\mathbf{t}^T \mathbf{X}$ is Gaussian for for all non-zero $\mathbf{t} \in \mathbb{R}^n$.
- ▶ If X_1, \dots, X_n are jointly Gaussian and A is a $k \times n$ matrix of rank k, then, $\mathbf{Y} = A\mathbf{X}$ is jointly gaussian

Recap: Moment inequalities

▶ **Jensen's Inequality**: If g is convex and EX and E[g(X)] exist

$$g(EX) \le E[g(X)]$$

▶ Holder Inequality: For p,q>1 and $\frac{1}{p}+\frac{1}{q}=1$

$$E|XY| \le (E|X|^p)^{\frac{1}{p}} (E|Y|^q)^{\frac{1}{q}}$$

(assuming all expectations exist)

- For p = q = 2, the above is Cauchy-Schwartz inequality
- ▶ This implies $|\rho_{XY}| \leq 1$
- ► Minkowski's Inequality:

$$(E|X+Y|^r)^{\frac{1}{r}} \le (E|X|^r)^{\frac{1}{r}} + (E|Y|^r)^{\frac{1}{r}}$$

Recap

Chernoff Bounds

$$P[X > a] \le \frac{E[e^{sX}]}{e^{sa}} = \frac{M_X(s)}{e^{sa}}, \forall s > 0$$

▶ Hoeffding Inequality X_i iid, $X_i \in [a, b], \forall i$ and $EX_i = \mu$

$$P\left[\left|\sum_{i=1}^{n} X_i - n\mu\right| \ge \epsilon\right] \le 2e^{-\frac{2\epsilon^2}{n(b-a)}}, \ \epsilon > 0$$

Recap: Weak Law of large numbers

 $ightharpoonup X_i$ are iid, $EX_i = \mu$, $Var(X_i) = \sigma^2$, $S_n = \sum_{i=1}^n X_i$

$$E\left[\frac{S_n}{n}\right] = \mu; \quad \text{ and } \quad \operatorname{Var}\left(\frac{S_n}{n}\right) = \frac{\sigma^2}{n}$$

► By Chebyshev Inequality

$$P\left[\left|\frac{S_n}{n} - \mu\right| \ge \epsilon\right] \le \frac{\operatorname{Var}\left(\frac{S_n}{n}\right)}{\epsilon^2} = \frac{\sigma^2}{n\epsilon^2}, \quad \forall \epsilon > 0$$

$$\Rightarrow \lim_{n \to \infty} P\left[\left|\frac{S_n}{n} - \mu\right| \ge \epsilon\right] = 0, \quad \forall \epsilon > 0$$

Recap: Convergence in Probability

A sequence of random variables, X_n , is said to **converge** in **probability** to a random variable X_0 is

$$\lim_{n \to \infty} P[|X_n - X_0| > \epsilon] = 0, \ \forall \epsilon > 0$$

This is denoted as $X_n \stackrel{P}{\to} X_0$

▶ By the definition of limit, the above means

$$\forall \delta > 0, \ \exists N < \infty, \ s.t. \ P[|X_n - X_0| > \epsilon] < \delta, \ \forall n > N$$

ightharpoonup We only need marginal distributions of individual X_n to decide whether a sequence converges to a constant in probability

Example

- Let $\Omega = [0, 1]$ with the usual probability measure and let $X_n = I_{[0, 1/n]}$.
- $P[X_n = 1] = \frac{1}{n} = 1 P[X_n = 0]$
- ▶ The probability of X_n taking value 1 is decreasing with n
- A good guess is that it converges to zero

$$P[|X_n - 0| > \epsilon] = P[X_n = 1] = \frac{1}{n}$$

which goes to zero as $n \to \infty$.

▶ Hence, $X_n \stackrel{P}{\rightarrow} 0$

Example

- Let X_1, X_2, \cdots be a sequence of iid random variable which are uniform over (0, 1).
- $\blacktriangleright \text{ Let } M_n = \max(X_1, X_2, \cdots, X_n)$
- ▶ Does M_n converge in probability?
- A reasonable guess for the limit is 1

$$P[|M_n - 1| \ge \epsilon] = P[M_n \le 1 - \epsilon] = (1 - \epsilon)^n$$

- ▶ This implies $M_n \stackrel{P}{\rightarrow} 1$
- Suppose $Z_n = \min(X_1, X_2, \cdots, X_n)$. Then $Z_n \stackrel{P}{\to} 0$

Some properties of convergence in probability

- $ightharpoonup X_n \stackrel{P}{\to} X \text{ and } X_n \stackrel{P}{\to} Y \Rightarrow P[X=Y] = 1$
- $ightharpoonup X_n \stackrel{P}{\to} X \Rightarrow P[|X_n X_m| > \epsilon] \to 0 \text{ as } n, m \to \infty$
- ▶ Suppose $X_n \stackrel{P}{\to} X$ and $Y_n \stackrel{P}{\to} Y$ Then the following hold
 - 1. $aX_n \stackrel{P}{\to} aX$
 - 2. $X_n + Y_n \stackrel{P}{\rightarrow} X + Y$
 - 3. $X_n Y_n \stackrel{P}{\to} XY$
 - 4. $g(X_n) \stackrel{P}{\to} g(X)$ where g is a continuous function from \Re to \Re .

- $\blacktriangleright X_n \stackrel{P}{\to} X$ and $X_n \stackrel{P}{\to} Y \Rightarrow P[X = Y] = 1$
- Proof: We have $|X - Y| = |X - X_n + X_n - Y| < |X - X_n| + |Y - X_n|.$

$$\Rightarrow \left(\left[|X - X_n| \le \frac{c}{2} \right] \cap \left[|Y - X_n| \le \frac{c}{2} \right] \right) \subset [|X - Y| \le c]$$

$$\Rightarrow [|X - Y| > c] \subset \left(\left\lceil |X - X_n| > \frac{c}{2} \right\rceil \cup \left\lceil |Y - X_n| > \frac{c}{2} \right\rceil \right)$$

$$\Rightarrow P[|X - Y| > c] \le P\left[|X - X_n| > \frac{c}{2}\right] + P\left[|X - X_n| > \frac{c}{2}\right], \forall n$$

- ightharpoonup Hence, we can conclude $P[|X-Y|>c]=0, \forall c>0$ and
- hence P[X = Y] = 1.

- $\blacktriangleright X_n \stackrel{P}{\to} X \text{ and } X_n \stackrel{P}{\to} Y \implies X_n + Y_n \stackrel{P}{\to} X + Y$
- ▶ We can prove this in the same manner

$$|X_n + Y_n - (X + Y)| = |(X_n - X) + (Y_n - Y)| \le |X_n - X| + |Y_n - Y|$$

► All the other properties listed can be proved by following similar arguments

We mentioned point-wise convergence of a sequence of functions

$$g_n \to g_0$$
 if $g_n(x) \to g_0(x)$, $\forall x$

- ➤ Since random variables are also functions we can define convergence like this.
- ▶ We can demand $X_n(\omega) \rightarrow X_0(\omega), \forall \omega$
- ► Such pointise convergence is too restrictive.
- \blacktriangleright But we can demand that it should be satisfied for almost all ω

A sequence of random variables, X_n , is said to converge almost surely or with probability one to X if

$$P(\{\omega : X_n(\omega) \to X(\omega)\}) = 1$$

or equivalently

$$P(\{\omega : X_n(\omega) \rightarrow X(\omega)\}) = 0$$

- ▶ Denoted as $X_n \stackrel{a.s.}{\to} X$ or $X_n \stackrel{w.p.1}{\to} X$ or $X_n \to X_0$ (w.p.1)
- ▶ We are saying that for 'almost all' ω , $X_n(\omega)$ converges to $X(\omega)$
- ▶ Given a sequence X_n , how does one establish almost sure convergence?

 \triangleright X_n converges to X almost surely or with probability one if

$$P(\{\omega : X_n(\omega) \to X(\omega)\}) = 1$$

By notation, this is same as

$$P[X_n \to X] = 1$$

Next we will try to characterize the event

$$\{\omega : X_n(\omega) \to X(\omega)\}$$

$$\{\omega : X_n(\omega) \nrightarrow X(\omega)\}$$

- ▶ Recall convergence of real number sequences.
- A sequence of real numbers x_n is said to converge to x_0 , $x_n \to x_0$, if

$$\forall \epsilon > 0, \ \exists N < \infty, \ s.t. \ |x_n - x_0| < \epsilon, \ \forall n > N$$

This is equivalent to

$$\forall \epsilon > 0, \ \exists N < \infty, \ \forall k > 0 \ |x_{N+k} - x_0| < \epsilon$$

ightharpoonup So, $x_n \rightarrow x_0$ means

$$\exists \epsilon \ \forall N \ \exists k \ |x_{N+k} - x_0| > \epsilon$$

- ▶ Note that given any ω , $X_n(\omega)$ is real number sequence.
- ▶ Hence $X_n(\omega) \to X(\omega)$ is same as

$$\forall \epsilon > 0 \ \exists N < \infty \ \forall k \ge 0 \ |X_{N+k}(\omega) - X(\omega)| \le \epsilon$$

This is equivalent to

$$\forall r>0, r \text{ integer} \quad \exists N<\infty \quad \forall k\geq 0 \ |X_{N+k}(\omega)-X(\omega)|\leq \frac{1}{r}$$

► Hence, $X_n(\omega) \rightarrow X(\omega)$ is same as

$$\exists r \ \forall N \ \exists k |X_{N+k}(\omega) - X(\omega)| > \frac{1}{r}$$

► Hence we can write this event as

$$\bigcup_{r=1}^{\infty} \cap_{N=1}^{\infty} \bigcup_{k=0}^{\infty} \left\{ \omega : |X_{N+k}(\omega) - X(\omega)| > \frac{1}{r} \right\}$$

▶ The event $\{\omega : X_n(\omega) \rightarrow X(\omega)\}\$ can be expressed as

$$\bigcup_{r=1}^{\infty} \cap_{N=1}^{\infty} \bigcup_{k=0}^{\infty} \left[|X_{N+k} - X| > \frac{1}{r} \right]$$

 \blacktriangleright Hence X_n converges almost surely to X iff

$$P\left(\bigcup_{r=1}^{\infty} \cap_{N=1}^{\infty} \bigcup_{k=0}^{\infty} \left[|X_{N+k} - X| > \frac{1}{r} \right] \right) = 0$$

This is same as

$$P\left(\cap_{N=1}^{\infty}\cup_{k=0}^{\infty}\left[\left|X_{N+k}-X\right|>\frac{1}{r}\right]\right)=0,\ \forall r>0,\ \mathrm{integer}$$

Same as

$$P\left(\bigcap_{N=1}^{\infty} \bigcup_{k=0}^{\infty} \left[|X_{N+k} - X| > \epsilon \right] \right) = 0, \quad \forall \epsilon > 0$$

► A sequence X_n is said to converge **almost surely** or with probability one to X if

$$P(\{\omega : X_n(\omega) \to X(\omega)\}) = 1$$

► We can also write it as

$$P[X_n \to X] = 1$$

▶ We showed that this is equivalent to

$$P\left(\bigcap_{N=1}^{\infty} \bigcup_{k=0}^{\infty} \left[|X_{N+k} - X| > \epsilon \right] \right) = 0, \ \forall \epsilon > 0$$

Same as

$$P\left(\bigcap_{N=1}^{\infty} \bigcup_{k=N}^{\infty} \left[|X_k - X| > \epsilon \right] \right) = 0, \ \forall \epsilon > 0$$

- $\blacktriangleright \text{ let } A_k = [|X_k X| \ge \epsilon]$
- $\blacktriangleright \text{ Let } B_N = \cup_{k=N}^{\infty} A_k.$
- ▶ Then, $B_{N+1} \subset B_N$ and hence $B_N \downarrow$.
- $\blacktriangleright \text{ Hence, } \lim B_N = \cap_{N=1}^{\infty} B_N.$
- ightharpoonup We saw that $X_n \stackrel{a.s.}{\to} X$ is same as

$$P\left(\bigcap_{N=1}^{\infty} \cup_{k=N}^{\infty} \left[|X_k - X| \ge \epsilon \right] \right) = 0, \ \forall \epsilon > 0$$

$$\Leftrightarrow P\left(\lim_{N\to\infty} \bigcup_{k=N}^{\infty} [|X_k - X| \ge \epsilon]\right) = 0, \ \forall \epsilon > 0$$

$$\Leftrightarrow \lim_{N \to \infty} P\left(\bigcup_{k=N}^{\infty} [|X_k - X| \ge \epsilon]\right) = 0, \ \forall \epsilon > 0$$

 $ightharpoonup X_n$ converges to X almost surely iff

$$\lim_{n \to \infty} P\left(\bigcup_{k=n}^{\infty} [|X_k - X| \ge \epsilon]\right) = 0, \quad \forall \epsilon > 0$$

- ▶ To show convergence with probability one using this one needs to know the joint distribution of X_n, X_{n+1}, \cdots
- ▶ Contrast this with $X_n \stackrel{P}{\to} X$ which is

$$\lim_{n \to \infty} P[|X_n - X| > \epsilon] = 0, \ \forall \epsilon > 0$$

► This also shows that

$$X_n \stackrel{a.s.}{\to} X \implies X_n \stackrel{P}{\to} X$$

► Almost sure convergence is a stronger mode of convergence

simple example: almost sure convergence

Let $\Omega = [0, 1]$ with the usual probability measure and let $X_n = I_{[0, 1/n]}$.

$$X_n(\omega) = \left\{ \begin{array}{ll} 1 & \omega \leq \frac{1}{n} \\ 0 & \text{otherwise} \end{array} \right.$$

- ▶ Since $X_n \stackrel{P}{\to} 0$, zero is the only candidate for limit
- $ightharpoonup X_n(\omega) = 1$ only when $n < 1/\omega$.
- ▶ Given any ω , for all $n > 1/\omega$, $X_n(\omega) = 0$
- ► Hence, $\{\omega : X_n(\omega) \to 0\} = (0, 1]$

$$P[X_n \to X_0] = P(\{\omega : X_n(\omega) \to 0\}) = P((0,1]) = 1$$

 \blacktriangleright Hence $X_n \stackrel{a.s}{\rightarrow} 0$

 $ightharpoonup X_n$ converges to X almost surely iff

$$P\left(\bigcap_{N=1}^{\infty} \bigcup_{k=N}^{\infty} \left[|X_k - X| \ge \epsilon \right] \right) = 0, \ \forall \epsilon > 0$$

$$\Leftrightarrow \lim_{n \to \infty} P\left(\bigcup_{k=n}^{\infty} [|X_k - X| \ge \epsilon]\right) = 0, \quad \forall \epsilon > 0$$

- \blacktriangleright We normally do not specify X_n as functions over Ω
- ► We are only given the distributions
- ▶ How do we establish convergence almost surely?
- ▶ We will first look at limits of general sequences of events.

- ▶ Let A_1, A_2, \cdots be a sequence of events.
- ► How do we define limit of this sequence ?
- ▶ Define sequences

$$B_n = \bigcup_{k=n}^{\infty} A_k$$
 $C_n = \bigcap_{k=n}^{\infty} A_k$

- ▶ These are monotone: $B_n \downarrow$, $C_n \uparrow$. They have limits.
- Define

$$\lim \sup A_n \triangleq \lim B_n = \bigcap_{n=1}^{\infty} \bigcup_{k=n}^{\infty} A_k$$
$$\lim \inf A_n \triangleq \lim C_n = \bigcup_{n=1}^{\infty} \bigcap_{k=n}^{\infty} A_k$$

- ▶ If $\limsup A_n = \liminf A_n$ then we define that as $\lim A_n$. Otherwise we say the sequence does not have a limit
- Note that $\lim \sup A_n$ and $\lim \inf A_n$ are events
- \blacktriangleright Note that $X_n \stackrel{a.s.}{\to} X$ iff

$$P\left(\bigcap_{N=1}^{\infty} \cup_{k=N}^{\infty} \left[|X_k - X| \ge \epsilon \right] \right) = 0, \ \forall \epsilon > 0$$

$$\Leftrightarrow P\left(\lim \sup \left[|X_n - X| \ge \epsilon \right] \right) = 0, \ \forall \epsilon > 0$$

▶ We can show $\lim \inf A_n \subset \lim \sup A_n$

$$\omega \in \lim \inf A_n \implies \omega \in \bigcup_{n=1}^{\infty} \cap_{k=n}^{\infty} A_k$$

$$\Rightarrow \exists m, \ \omega \in A_k, \ \forall k \ge m$$

$$\Rightarrow \omega \in \bigcup_{j=n}^{\infty} A_j, \ \forall n$$

$$\Rightarrow \omega \in \bigcap_{n=1}^{\infty} \bigcup_{j=n}^{\infty} A_j$$

$$\Rightarrow \omega \in \lim \sup A_n$$

ightharpoonup We can characterize $\lim \inf A_n$ as follows

$$\begin{array}{ll} \omega \in \lim \, \inf A_n & \Rightarrow & \omega \in \cup_{n=1}^\infty \cap_{k=n}^\infty A_k \\ & \Rightarrow & \exists m, \; \omega \in A_k, \; \forall k \geq m \\ & \Rightarrow & \omega \; \text{belongs to all but finitely many of} \; \; A_n \end{array}$$

Thus, $\lim \inf A_n$ consists of all points that are there in all but finitely many A_n .

 \blacktriangleright We can characterize $\lim \sup A_n$ as follows

$$\begin{array}{lll} \omega \in \lim \, \sup A_n & \Rightarrow & \omega \in \cap_{n=1}^\infty \cup_{k=n}^\infty A_k \\ & \Rightarrow & \omega \in \cup_{k=n}^\infty A_k, \, \forall n \\ & \Rightarrow & \omega \text{ belongs to infinitely many of } \, A_n \end{array}$$

Thus $\limsup A_n$ consists of points that are in infinitely many A_n

One refers to $\limsup A_n$ also as ' A_n infinitely often' or ' A_n i.o.'

- lackbox What is the difference between Points that belong to all but finitely many A_n and Points that belong to infinitely many A_n
- ▶ There can be ω that are there in infinitely many of A_n and are also not there in infinitely many of the A_n

Example

- \blacktriangleright Consider the following sequence of sets: A, B, A, B, \cdots
- ► Recall

$$\lim \sup A_n = \bigcap_{n=1}^{\infty} \bigcup_{k=n}^{\infty} A_k \quad \lim \inf A_n = \bigcup_{n=1}^{\infty} \bigcap_{k=n}^{\infty} A_k$$

$$\bigcup_{k=n}^{\infty} A_k = A \cup B, \ \forall n \quad \Rightarrow \quad \lim \sup A_n = A \cup B$$

$$\bigcap_{k=n}^{\infty} A_k = A \cap B, \ \forall n \ \Rightarrow \ \lim \inf A_n = A \cap B$$

example

Consider the sets $A_n = [0, 1 + \frac{(-1)^n}{n})$ The sequence is

$$[0, 0), [0, 1+\frac{1}{2}), [0, 1-\frac{1}{3}), [0, 1+\frac{1}{4})\cdots$$

- ▶ Guess: $\limsup A_n = [0, 1]$ and $\liminf A_n = [0, 1)$
- First note that $[0, 1 + \frac{1}{n+1}) \subset \bigcup_{k=n}^{\infty} A_k \subset [0, 1 + \frac{1}{n}).$

First note that
$$[0, 1 + \frac{1}{n+1}) \subset \bigcup_{k=n}^{\infty} A_k \subset [0, 1 + \frac{1}{n})$$
. Hence

$$x \in [0, 1] \Rightarrow x \in \bigcup_{k=n}^{\infty} A_k, \ \forall n \ \Rightarrow x \in \bigcap_{n=1}^{\infty} \bigcup_{k=n}^{\infty} A_k \ \Rightarrow x \in \lim \sup A_n$$

- ▶ Given any $\epsilon > 0$, $1 + \epsilon \notin [0, 1 + \frac{1}{n})$ if $\epsilon > \frac{1}{n}$ or $n > \frac{1}{\epsilon}$.
- ► Hence, given any $\epsilon > 0$, $\exists n \text{ such that } 1 + \epsilon \notin \bigcup_{k=-n}^{\infty} A_k$.
- ▶ This proves $\limsup A_n = [0, 1]$

P S Sastry, IISc, E1 222 Lecture 16, Aug 2021 28/38

- Now let us consider: $\lim \inf A_n = \bigcup_{n=1}^{\infty} \bigcap_{k=n}^{\infty} A_k$.
- ▶ Recall $A_n = [0, 1 + \frac{(-1)^n}{n})$
- First note that $[0, 1-\frac{1}{n}) \subset \bigcap_{k=n}^{\infty} A_k \subset [0, 1-\frac{1}{n+1})$
- ▶ Given any $\epsilon > 0$, $1 \epsilon < 1 \frac{1}{n}$ if $n > \frac{1}{\epsilon}$
- ▶ Hence, given any $\epsilon > 0$, $\exists n$ such that $1 \epsilon \in \bigcap_{k=n}^{\infty} A_k$
- ▶ Hence $1 \epsilon \in \lim \inf A_n$
- ▶ It is easy to see $1 \notin \bigcap_{k=n}^{\infty} A_k$ for ay n.
- ▶ Hence $1 \notin \lim \inf A_n$
- ▶ This proves $\lim \inf A_n = [0, 1)$
- ▶ Since $\limsup A_n \neq \liminf A_n$, this sequence does not have a limit

 $\blacktriangleright X_n \stackrel{a.s.}{\to} X$ iff

$$P\left(\bigcap_{N=1}^{\infty} \cup_{k=N}^{\infty} \left[|X_k - X| \ge \epsilon \right] \right) = 0, \ \forall \epsilon > 0$$

- $\blacktriangleright \text{ Let } A_n^{\epsilon} = [|X_n X| \ge \epsilon]$
- ▶ Then $X_n \stackrel{a.s.}{\to} X$ iff

$$P(\lim \sup A_n^{\epsilon}) = 0, \ \forall \epsilon > 0$$

- The question now is: can we get probability of $\limsup A_n$
- ▶ We look at an important result that allows us to do this

Borel-Cantelli Lemma

- **Borel-Cantelli lemma**: Given sequence of events, A_1, A_2, \cdots
 - 1. If $\sum_{i=1}^{\infty} P(A_i) < \infty$, then, $P(\limsup A_n) = 0$
 - 2. If $\sum_{i=1}^{\infty} P(A_i) = \infty$ and A_i are independent, $P(\limsup A_n) = 1$

Proof:

- ▶ We will first show: $P(\bigcup_{i=n}^{\infty} A_i) \leq \sum_{i=n}^{\infty} P(A_i), \ \forall n$
- ▶ We have the result: $P(\bigcup_{i=n}^{N} A_i) \leq \sum_{i=n}^{N} P(A_i), n \leq N$
- For any n, let $B_N = \bigcup_{i=n}^N A_i$. Then $B_N \subset B_{N+1}$.

$$P(\bigcup_{i=n}^{\infty} A_i) = P(\lim_{N \to \infty} \bigcup_{i=n}^{N} A_i) = \lim_{N \to \infty} P(\bigcup_{i=n}^{N} A_i)$$

$$\leq \lim_{N \to \infty} \sum_{i=n}^{N} P(A_i) = \sum_{i=n}^{\infty} P(A_i)$$

▶ If
$$\sum_{k=1}^{\infty} P(A_k) < \infty$$
, then, $\lim_{n\to\infty} \sum_{k=n}^{\infty} P(A_k) = 0$

$$0 \le P (\limsup A_n) = P \left(\bigcap_{n=1}^{\infty} \bigcup_{k=n}^{\infty} A_k \right)$$

$$= P \left(\lim_{n \to \infty} \bigcup_{k=n}^{\infty} A_k \right)$$

$$= \lim_{n \to \infty} P \left(\bigcup_{k=n}^{\infty} A_k \right)$$

$$\le \lim_{n \to \infty} \sum_{k=n}^{\infty} P(A_k)$$

$$= 0, \quad \text{if} \quad \sum_{k=n}^{\infty} P(A_k) < \infty$$

▶ This completes proof of first part of Borel-Cantelli lemma

- ▶ Let $\sum_{k=1}^{\infty} P(A_k) = C < \infty$
- ▶ It means given any $\epsilon > 0$, $\exists n$

$$\left| \sum_{k=1}^{n} P(A_k) - C \right| < \epsilon \quad \Rightarrow \quad \left| \sum_{k=n}^{\infty} P(A_k) \right| < \epsilon$$

► This implies

$$\lim_{n \to \infty} \sum_{k=n}^{\infty} P(A_k) = 0$$

► For the second part of the lemma:

$$\begin{split} P\left(\lim\sup A_n\right) &= P\left(\cap_{n=1}^{\infty} \cup_{k=n}^{\infty} A_k\right) \\ &= P\left(\lim_{n \to \infty} \cup_{k=n}^{\infty} A_k\right) \\ &= \lim_{n \to \infty} P\left(\cup_{k=n}^{\infty} A_k\right) \\ &= \lim_{n \to \infty} \left(1 - P\left(\cap_{k=n}^{\infty} A_k^c\right)\right) \\ &= \lim_{n \to \infty} \left(1 - \prod_{k=n}^{\infty} \left(1 - P(A_k)\right)\right) \\ &= \operatorname{because} A_k \text{ are independent} \\ &= 1 - \lim_{n \to \infty} \prod_{k=n}^{\infty} \left(1 - P(A_k)\right) \end{split}$$

► We can compute that limit as follows

$$\lim_{n \to \infty} \prod_{k=n}^{\infty} (1 - P(A_k)) \leq \lim_{n \to \infty} \prod_{k=n}^{\infty} e^{-P(A_k)}, \text{ since } 1 - x \leq e^{-x}$$

$$= \lim_{n \to \infty} e^{-\sum_{k=n}^{\infty} P(A_k)}$$

$$= 0$$

because

$$\sum_{k=1}^{\infty} P(A_k) = \infty \quad \Rightarrow \quad \lim_{n \to \infty} \sum_{k=n}^{\infty} P(A_k) = \infty$$

► This finally gives us

$$P(\lim \sup A_n) = 1 - \lim_{n \to \infty} \prod^{\infty} (1 - P(A_k)) = 1$$

- Given a sequence X_n we want to know whether it converges to X
- $\blacktriangleright \text{ Let } A_k^{\epsilon} = [|X_k X| \ge \epsilon]$
- $X_n \stackrel{P}{\to} X$ if

k=1

$$\lim_{k\to\infty}P[|X_k-X|\geq\epsilon]=0\quad\text{ same as }\ \lim_{k\to\infty}P(A_k)=0,\ \ \forall\epsilon>0$$

► By Borel-Cantelli lemma

$$\sum_{k=0}^{\infty} P(A_k) < \infty \implies P(\lim \sup A_k) = 0 \implies X_k \stackrel{a.s.}{\to} X$$

ightharpoonup Consider a sequence X_n with

$$P[X_n = 0] = 1 - a_n; \quad P[X_n = c_n] = a_n$$

- ▶ We want to investigate convergence to 0.
- $A_n^{\epsilon} = [|X_n 0| > \epsilon] = [X_n = c_n], \ \forall \epsilon > 0$
- ightharpoonup Hence $P(A_n^{\epsilon}) = a_n, \ \forall \epsilon > 0.$
- ▶ If $a_n \to 0$ then $X_n \stackrel{P}{\to} 0$. (e.g., $a_n = \frac{1}{n}, \frac{1}{n^2}$)
- ightharpoonup If $\sum a_n < \infty$, $X_n \stackrel{a.s.}{\to} 0$ (e.g., $a_n = \frac{1}{n^2}$)

ightharpoonup Consider a sequence X_n with

$$P[X_n = 0] = 1 - \frac{1}{n}; \quad P[X_n = 1] = \frac{1}{n}$$

- ▶ We can easily conclude $X_n \stackrel{P}{\to} 0$.
- ▶ But since, $\sum_{n} \frac{1}{n} = \infty$, Borel-Cantelli lemma is not really useful here
- ightharpoonup We saw one example where such X_n converge almost surely.
- ▶ But, if X_n are independent, then by Borel-Cantelli lemma, they do not converge
- Convergence (to a constant) in probability depends only on distribution of individual X_n .
- Convergence almost surely depends on the joint distribution