

Q4.

$$a) (a) \text{ Likelihood} = 1/\sigma^n$$

for likelihood to be max^m σ has to min^m

but $\sigma > x_1, x_2, \dots, x_n$

$$\therefore \hat{\sigma}_{ml} = \max\{x_1, x_2, \dots, x_n\}$$

$$\text{Prior: } P(\sigma) \propto \left(\frac{\sigma_m}{\sigma}\right)^2 \quad \sigma > \sigma_m$$

$$= 0 \quad \text{otherwise}$$

$$\text{Posterior} \propto 1/\sigma^n \cdot \left(\frac{\sigma_m}{\sigma}\right)^2 \propto \frac{\sigma_m^2}{\sigma^{n+2}}$$

for this to be max^m σ has to min^m

but $\sigma > \{x_1, x_2, \dots, x_n\}$

$\sigma > \sigma_m$

$$\therefore \hat{\sigma}_{MAP} = \max\{x_1, x_2, \dots, x_n, \sigma_m\}$$

(b) No, because σ_{MAP} ^{depends} ~~depends~~ on σ_m but σ_{ml} doesn't
this is not desirable.

Also, θ_{ml} asymptotically converges to the true value, but θ_{map} may or may not.

Q4)

$$(c) \text{ Posterior } \propto \frac{\theta_m^\alpha \left(\prod_{i=1}^n \frac{1}{\theta} \right)}{\theta^\alpha} \text{ if } \theta \geq \theta_m \text{ \& } \theta \geq \max_i(x_i) \\ = 0 \text{ otherwise}$$

$$\text{Posterior } \propto \frac{\theta_m^{n+\alpha}}{\theta^{n+\alpha}}$$

$$\propto \frac{\theta_m^{n+\alpha}}{\theta^{n+\alpha}} \left(\frac{C}{\theta} \right)^{n+\alpha}, \text{ where}$$

$$C = \max(\theta_m, \max_i(x_i))$$

So, this is a pareto distribution with parameters $C, n+\alpha$

$$\Rightarrow \text{Posterior mean} = \frac{\int_C^\infty \frac{C^{n+\alpha}}{\theta^{n+\alpha}} \cdot \theta d\theta}{\int_C^\infty \frac{C^{n+\alpha}}{\theta^{n+\alpha}} d\theta} \\ = \left(\frac{n+\alpha-1}{n+\alpha-2} \right) \cdot C$$

$$\hat{\theta}_{\text{Posterior mean}} = \left(\frac{n + \alpha - 1}{n + \alpha - 2} \right) \cdot \max(\theta_m, \max_i x_i)$$

(d) As is clear, $\lim_{n \rightarrow \infty} \hat{\theta}_{\text{Posterior mean}} = \max(\theta_m, \max_i x_i)$

$$= \hat{\theta}_{\text{map}}$$

$$\neq \hat{\theta}_{\text{MLE}}$$

This is not desirable, because an infinite sample size should mean that likelihood is ~~not~~ much more informative than prior.