

(3) For every data point,  $x = \begin{pmatrix} x \\ y \end{pmatrix}$  &  
 $x^T x = r^2$ , or  $x = r \begin{pmatrix} \cos \theta \\ \sin \theta \end{pmatrix}$ , where  
 $\theta$  is uniformly distributed on  $(0, 2\pi)$ .

$$P(x_i | \mu, C) = \frac{1}{2\pi\sqrt{|C|}} \exp\left(-\frac{(x_i - \mu)^T C^{-1} (x_i - \mu)}{2}\right)$$

$$\text{Log Likelihood} = -\frac{n}{2} \log |C| - \frac{1}{2} \sum_{i=1}^N (x_i - \mu)^T C^{-1} (x_i - \mu) + C$$

$$\frac{\partial LL}{\partial \mu} = 0 \Rightarrow \left( \sum (x_i - \hat{\mu})^T C^{-T} + \sum (x_i - \hat{\mu})^T C^{-1} \right) = 0$$

$$\Rightarrow \text{[using } d(A^T B C) = C^T B^T d(A^T) + A^T B dC]$$

$$\Rightarrow \sum (x_i - \hat{\mu})^T (C^{-1} + C^{-T}) = 0$$

$$(C^{-1} + C^{-T}) \sum (x_i - \hat{\mu}) = 0$$

$$\Rightarrow \sum x_i - N \hat{\mu} = 0$$

$$\Rightarrow \hat{\mu} = \frac{\sum x_i}{N} = \frac{r \sum \begin{pmatrix} \cos \theta_i \\ \sin \theta_i \end{pmatrix}}{N}$$

Now, to calculate  $\hat{C} \rightarrow$

$$LL = -\frac{n}{2} \log |C| - \frac{1}{2} \sum_{i=1}^N (x_i - \mu)^T C^{-1} (x_i - \mu)$$

$$= \frac{n}{2} \log |C^{-1}| - \frac{1}{2} \sum_{i=1}^N (x_i - \mu)^T C^{-1} (x_i - \mu)$$

Since inverse of a matrix is unique, we can differentiate the expression wrt  $C^{-1}$ , & set it to 0 to obtain MLE for  $C$ .

$$\Rightarrow \frac{\partial \mathcal{L}}{\partial C^{-1}} = 0 \Rightarrow \frac{N}{2} \hat{C} - \frac{1}{2} \sum_{i=1}^N (x_i - \mu) \cdot (x_i - \mu)^T = 0$$

The first term is obtained by the identity

~~$$\frac{\partial}{\partial A} \log |A| = A^{-1}$$~~

$$\frac{\partial}{\partial x} \log |x| = x^{-1},$$

& the second by the identity

$$\frac{d}{da} [\text{tr}(XX^T(Ya + z))] = XX^T Y$$

$$\begin{aligned} [\because (x_i - \mu)^T C^{-1} (x_i - \mu) &= \text{tr}((x_i - \mu)^T C^{-1} (x_i - \mu)), \\ \text{since both are } 1 \times 1 \text{ L.H.I is } |X| \text{ in dimension.} \\ \text{further, since } \text{tr}(XY) &= \text{tr}(YX), \\ \text{tr}((x_i - \mu)^T C^{-1} (x_i - \mu)) &= \text{tr}(C^{-1} (x_i - \mu) \cdot (x_i - \mu)^T)] \end{aligned}$$

$$\Rightarrow \hat{C} = \frac{1}{N} \left( \sum_{i=1}^N (x_i - \hat{\mu}) \cdot (x_i - \hat{\mu})^T \right)$$

$$\hat{C} = \frac{1}{N} \sum \begin{pmatrix} (\cos \theta_i - \sum_N \cos \theta_j)^2 & (\cos \theta_i - \sum_N \cos \theta_j) \cdot (\sin \theta_i - \sum_N \sin \theta_j) \\ (\cos \theta_i - \sum_N \cos \theta_j) \cdot (\sin \theta_i - \sum_N \sin \theta_j) & (\sin \theta_i - \sum_N \sin \theta_j)^2 \end{pmatrix}$$

Mode of Gaussian  $\rightarrow x$  for which  $P(x|\mu, C)$  is max.

$$\Rightarrow \frac{dP}{dx} = 0$$

$$2) P(x|\mu, C) \cdot \left[ \frac{-(x-\mu)^T C^{-T} - (x-\mu)^T C^{-1}}{2} \right] = 0$$

$$\Rightarrow x_{\text{mode}} - \mu = 0 \quad [\because (C^{-1} + C^{-T}) \neq 0]$$

$$\Rightarrow \underline{x_{\text{mode}} = \mu}$$

Now, as  $N \rightarrow \infty$ ,

$$\hat{\mu} = \frac{r}{N} \begin{pmatrix} \sum_{i=1}^N \cos \theta_i \\ \sum_{i=1}^N \sin \theta_i \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \end{pmatrix} \quad [\because \theta_i \text{ are uniformly distributed RV over } (0, 2\pi) \text{ so by law of large numbers, means of } \cos \theta_i \text{ \& } \sin \theta_i \rightarrow 0]$$

$$\hat{C} = \frac{r^2}{N} \begin{pmatrix} (\cos \theta_i)^2 & \cos \theta_i \sin \theta_i \\ \cos \theta_i \sin \theta_i & (\sin \theta_i)^2 \end{pmatrix}$$

$$= r^2 \begin{pmatrix} \frac{1}{2} & 0 \\ 0 & \frac{1}{2} \end{pmatrix}$$

[similar arguments]



Hence, mode =  $\begin{pmatrix} 0 \\ 0 \end{pmatrix}$ .

This means that Gaussian is not a good fit for data, as none of data points lie inside circle, where the probability is highest.

Experimental Values-

$N=1000000$ .

$r=2$

$\text{Mu} = (1e-5 \quad 2e-5)$

$C = \begin{matrix} 2.0000 & 1e-6 \\ 1e-6 & 2.0001 \end{matrix}$

These are very close to the theoretical values predicted by the Gaussian kernel.

